

# LDPC error floor estimation using Importance Sampling

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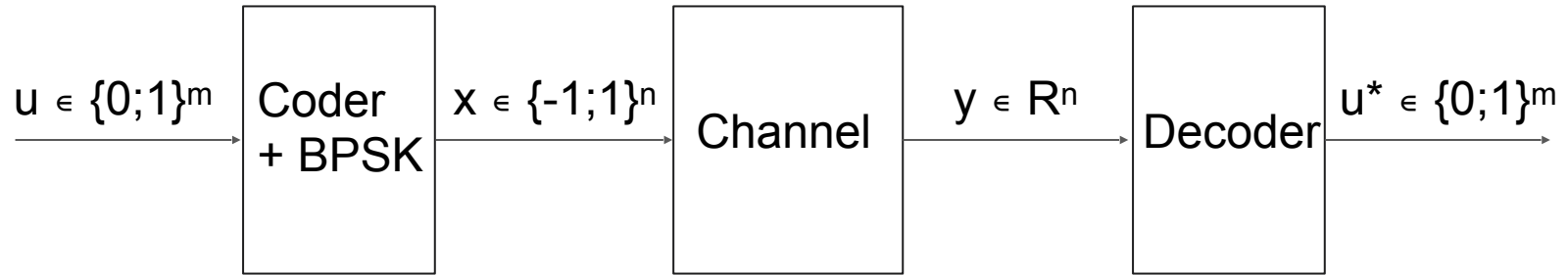
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# Agenda

- Introduction to LDPC codes
- Problem formulation
- Importance Sampling
- Quality metrics
- Numerical results
- Conclusion

# Intro to coding theory



# Simple codes

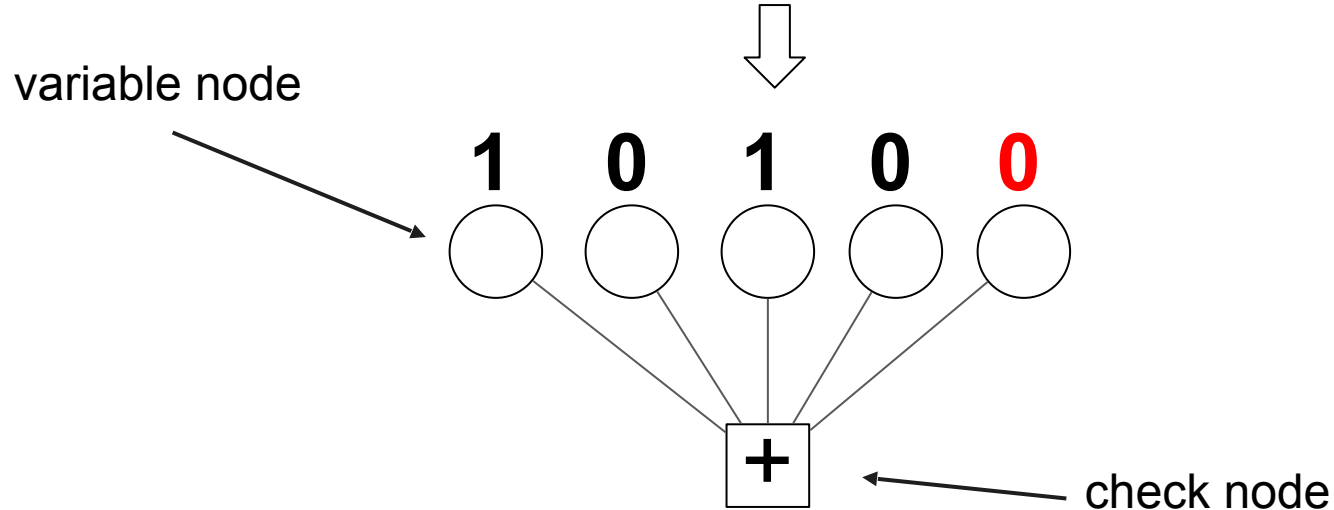
**1010**  $\longrightarrow$  **11...100...011...100...0**  
⏟  
n times

1010  $\longrightarrow$  10100

parity bit

# Parity check codes

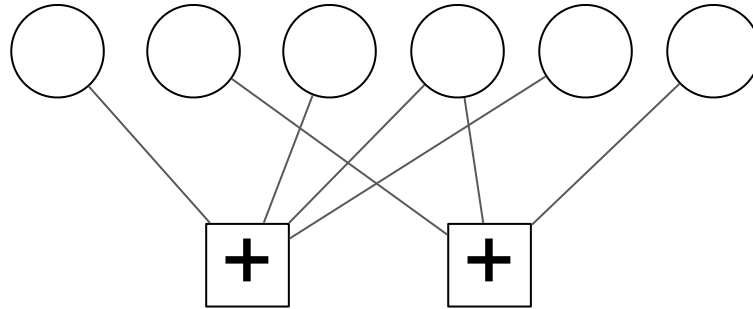
$$10100 \Rightarrow x_1 + x_2 + x_3 + x_4 + x_5 = 0$$



# Parity check codes

$$\begin{cases} x_1 + x_3 + x_4 + x_5 = 0 \\ x_2 + x_4 + x_6 = 0 \end{cases}$$

$$H = \begin{bmatrix} 1 & 0 & 1 & 1 & 1 & 0 \\ 0 & 1 & 0 & 1 & 0 & 1 \end{bmatrix}$$



$$H =$$

# LDPC codes

$$H_c = [ -P_{m \times (n-m)}^T \mid E_{m \times m} ]$$

$$G_c = [ E_{(n-m) \times (n-m)} \mid P_{(n-m) \times m} ]$$

A block code  $C$  with parity check matrix  $H$  can be defined in two ways:

$$1) \ C = \{ c \in \{0; 1\}^n \mid c * H^T = 0 \}$$

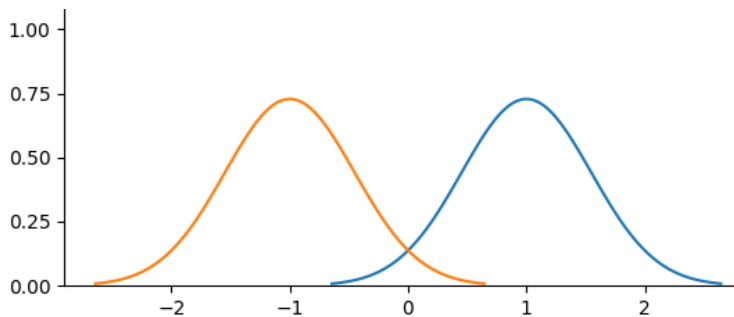
$$1) \ C = \{ u * G \text{ for every } u \in \{0; 1\}^{n-m} \}$$



# Decoding of LDPC codes

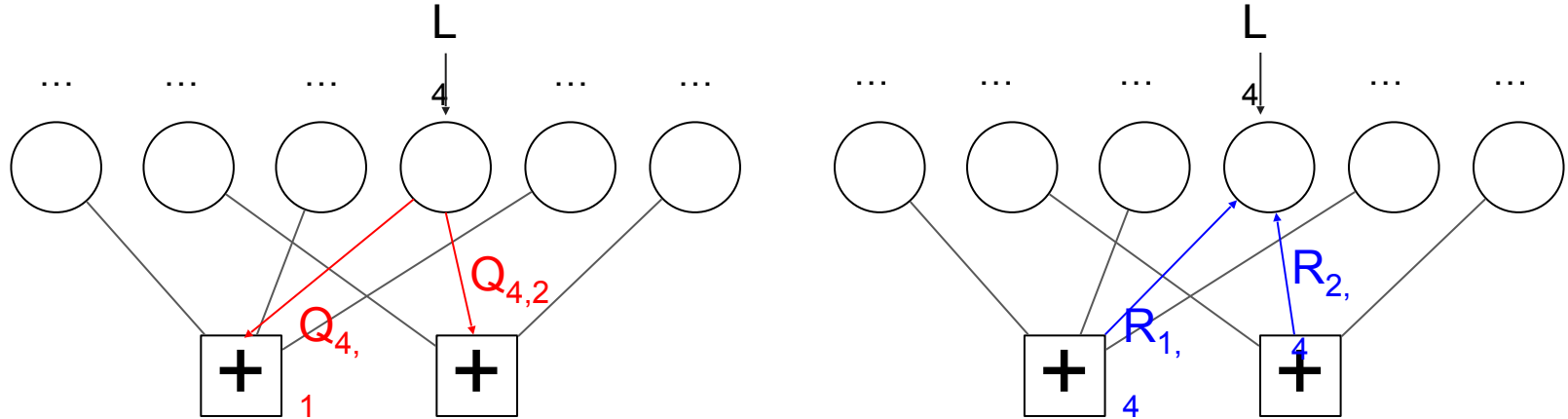
AWGN channel: noise  $\sim N(0, \sigma)$

$y = \pm 1 + \text{noise}$



$$L_j = \log [ \Pr(x_j=0 \mid y_j) / \Pr(x_j=1 \mid y_j) ] = 2y_j / \sigma^2$$

# Decoding of LDPC codes: MinSum



$$R_{j,i} = \prod_{k \in N(j) \setminus i} \text{sign}(Q_{k,j}) * \min_{k \in N(j) \setminus i} |Q_{k,j}|$$

$$Q_i = L_i + \sum_{k \in N(i)} R_{k,i}$$

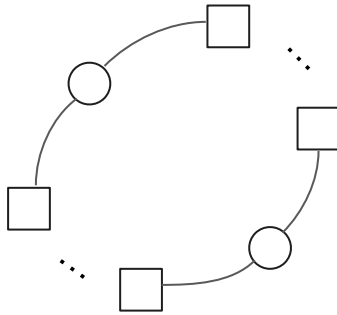
$$Q_{i,j} = L_i + \sum_{k \in N(i) \setminus j} R_{k,i}$$

# Trapping Sets

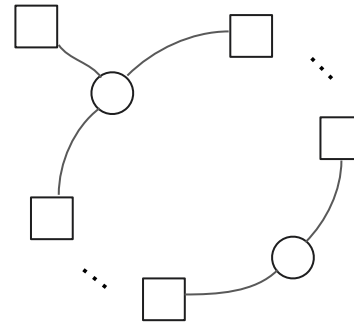
TS, denoted as  $(\mathbf{a}, \mathbf{b})$ , is the set of  $\mathbf{a}$  variable nodes and adjacent check nodes,  $\mathbf{b}$  of which have an odd number of occurrences in the given TS.

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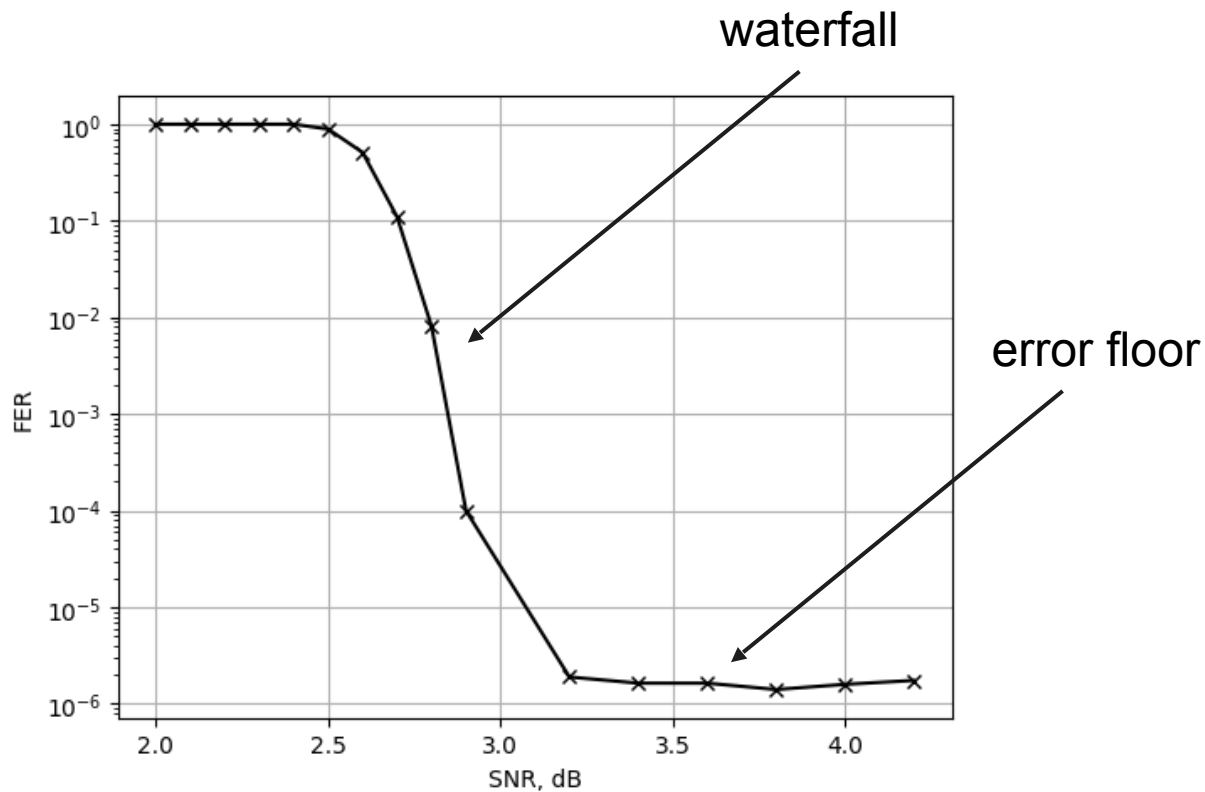


$(\mathbf{a}, 0)$



$(\mathbf{a}, 1)$

# Error floor



# Problem Statement

- Very low probability can hardly be estimated with Monte Carlo method.
- There is a need in the fast and accurate error floor estimation for LDPC code construction.
- There can be several different estimations, which one is better? There is a need in the universal criterion.

# Importance Sampling

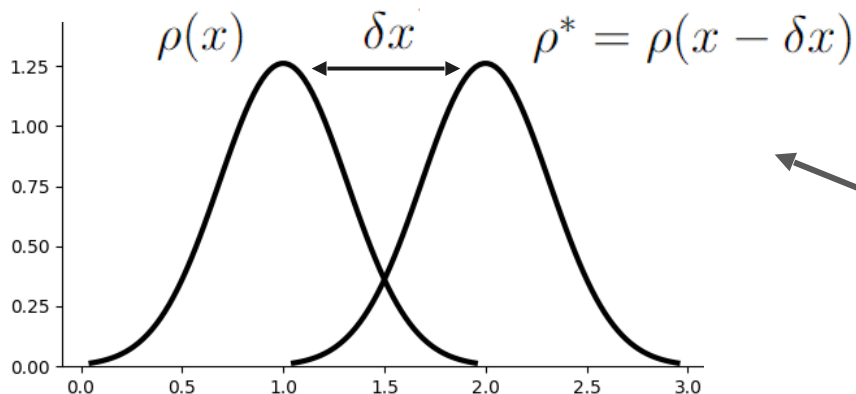
Let  $X$  be a random variable from  $\mathbb{R}^n$  with pdf  $\rho$  and let  $Y$  be a random variable with pdf  $\rho^*$ , such that  $\rho^* \neq 0$ , then

$$\mathbb{E}_{\rho} f(X) = \int_{\mathbb{R}^n} f(\vec{x}) \rho(\vec{x}) d\vec{x} = \int_{\mathbb{R}^n} f(\vec{y}) \frac{\rho(\vec{y})}{\rho^*(\vec{y})} \rho^*(\vec{y}) d\vec{y} = \mathbb{E}_{\rho^*} f(Y) w(Y)$$

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Mean Shift IS



# Goals and Tasks

## Goal:

Develop a criterion for comparing the results of different estimates of the LDPC code error floor and compare the IS estimates with uniform and shifted normal distributions.

## Tasks:

1. Identify the limitations of the existing criteria.
2. Develop a new criterion that eliminates these limitations.
3. Compare the estimates of IS with uniformly distribution and with the mean shifted normal one.

# Novelty

Most of paper use Mean-Shift IS (MS-IS) and the only criterion, which has no need in the MC values is gamma-criterion:

- M. Zhu, M. Jiang, and C. Zhao, “*Error floor estimation of qc-ldpc coded modulation with importance sampling*,” IEEE Communications Letters, vol. 25, no. 1, pp. 28–32, 2021
- Neshaastegaran, Peyman, Amir H. Banihashemi и Ramy H. Gohary (май 2021). “*Error Floor Estimation of LDPC Coded Modulation Systems Using Importance Sampling*”. B: IEEE Transactions on Communications 69.5, с. 2784—2799. doi: 10.1109/TCOMM.2021.3057625

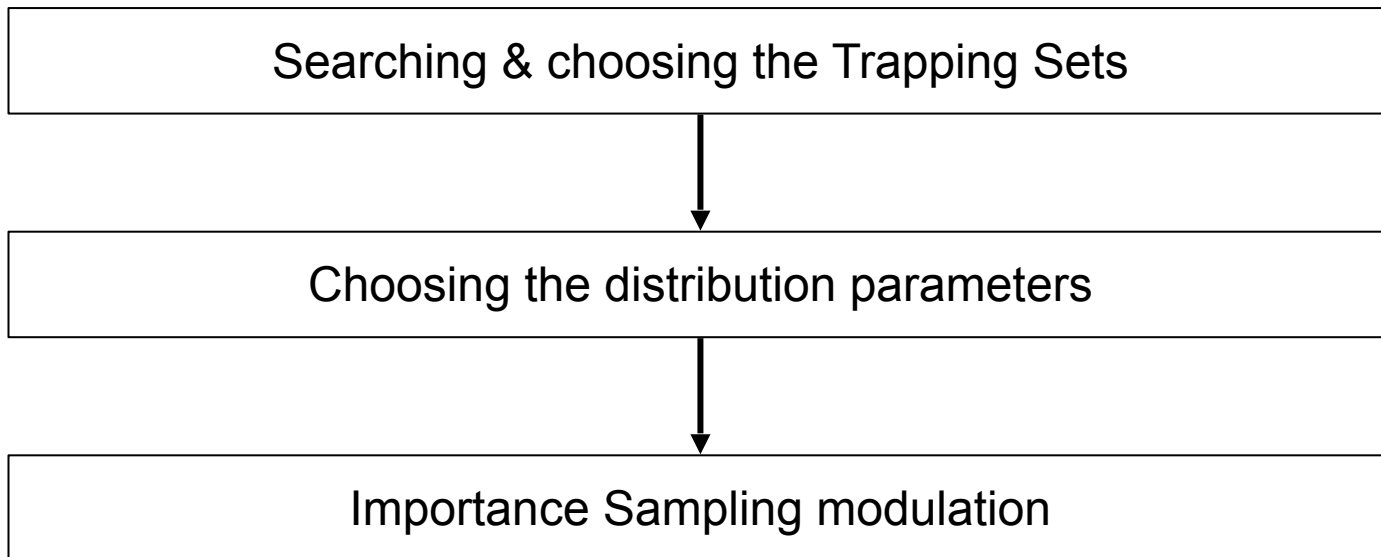
# IS applied to LDPC codes

- $f(\vec{x}) = \mathbb{I}_{\mathbb{A}}(\vec{x})$ , where  $A$  is a set of inputs causing the decoder failure  
 $\Pr(\vec{x} \in \mathbb{A}) \ll 1$

$\rho^*(\vec{y})$  should satisfy both:

- $\Pr(\vec{y} \in \mathbb{A}) \sim 1$   
 $\mathbb{E} \mathbb{I}_{\mathbb{A}}^2(Y) w^2(Y)$  is small enough

# Error floor estimation



# Choosing Trapping Sets

Searching all the possible TSs:

- Find all the cycles in Tanner graph
- Find unions of cycles

Choosing the dangerous ones:

- Apply Mean Shift IS with shifting E
- Choose those, which Bit Error Rate (BER) > ber threshold

# Choosing the Distribution parameters

For each Signal-to-Noise Ratio (SNR) and each pdf:

- Iterate through distribution parameters
- Choose the first one, which satisfy the condition:  $\Pr(\vec{y} \in \mathbb{A}) > \theta$  where  $\theta$  is some predefined parameter  $\in (0, 1)$ .

# Importance Sampling Modulation

$$\rho^*(\vec{y}) = \frac{1}{|T|} \sum_{t \in T} \prod_{i \in t} p_{IS}(\vec{y}_i) \prod_{j \notin t} p_{MC}(\vec{y}_j)$$

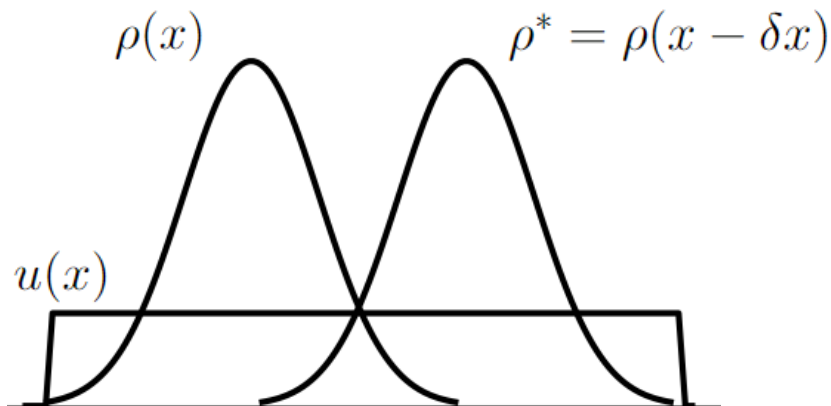
$T$  – set of chosen TSs.

$p_{IS}$  – sampling pdf (shifted normal or uniform).

$p_{MC}$  – pdf of AWGN channel.

$i, j$  – indexes of variable nodes.

# Uniform vs Normal



If shifted normal pdf is used for sampling, then weight function  $\omega$  has log-normal distribution with large variance. It yields a slow convergence.



# Uniform vs Normal

For unlimited domain (e.g. normal pdf):

$$\mathbb{E}_{\rho^*} w(Y) = 1 \implies \hat{\mathbb{E}}_{\rho^*} \mathbb{I}_{\mathbb{A}}(X) = \sum_{i=1}^N \mathbb{I}_{\mathbb{A}}(\vec{y}_i) w(\vec{y}_i)$$

For limited domain (e.g. uniform pdf):

$$\mathbb{E}_{\rho^*} w(Y) < 1 \implies \tilde{\mathbb{E}}_{\rho^*} \mathbb{I}_{\mathbb{A}}(X) = \frac{\sum_{i=1}^N \mathbb{I}_{\mathbb{A}}(\vec{y}_i) w(\vec{y}_i)}{\sum_{i=1}^N w(\vec{y}_i)}$$

# Metrics: MSE

$$\text{MSE} = \sum_{\text{SNR}} \left( \log \mathbb{E}_{\rho^*}^{IS} \mathbb{I}_{\mathbb{A}}(X) - \log \mathbb{E}_{\rho}^{MC} \mathbb{I}_{\mathbb{A}}(X) \right)^2$$

Pros:

- + independence from the estimator

Cons:

- need to calculate MC values

# Metrics: gamma

$$\gamma(\rho^*) = \frac{1}{N} \left( \frac{\widehat{\mathbb{E}}_{\rho^*} \mathbb{I}^2(X)}{\widehat{\mathbb{E}}_{\rho^*}^2 \mathbb{I}(X)} - 1 \right)$$

Pros:

- + no need to calculate MC values

Cons:

- dependence from the estimator

$$\mathbb{E}_{\rho} f(X) \in \left( \widehat{\mathbb{E}}_{\rho} f(X) - \frac{C}{\sqrt{N}}, \widehat{\mathbb{E}}_{\rho} f(X) + \frac{C}{\sqrt{N}} \right)$$

# New criterion

Most of the confidence intervals are as follows:

$$(p - CN^{-\alpha}, p + CN^{-\alpha})$$

$$C > 0,$$

$$\alpha > 0,$$

$p$  – predicted value

$N$  – number of trials

# New criterion

Suppose, we have two independent estimations of  $N$  variables  $F$  and  $F'$ , then

$$\alpha = -\frac{\ln|F - F'|}{\ln N}$$

$$r = \frac{\alpha(\text{uniform pdf})}{\alpha(\text{normal pdf})}$$

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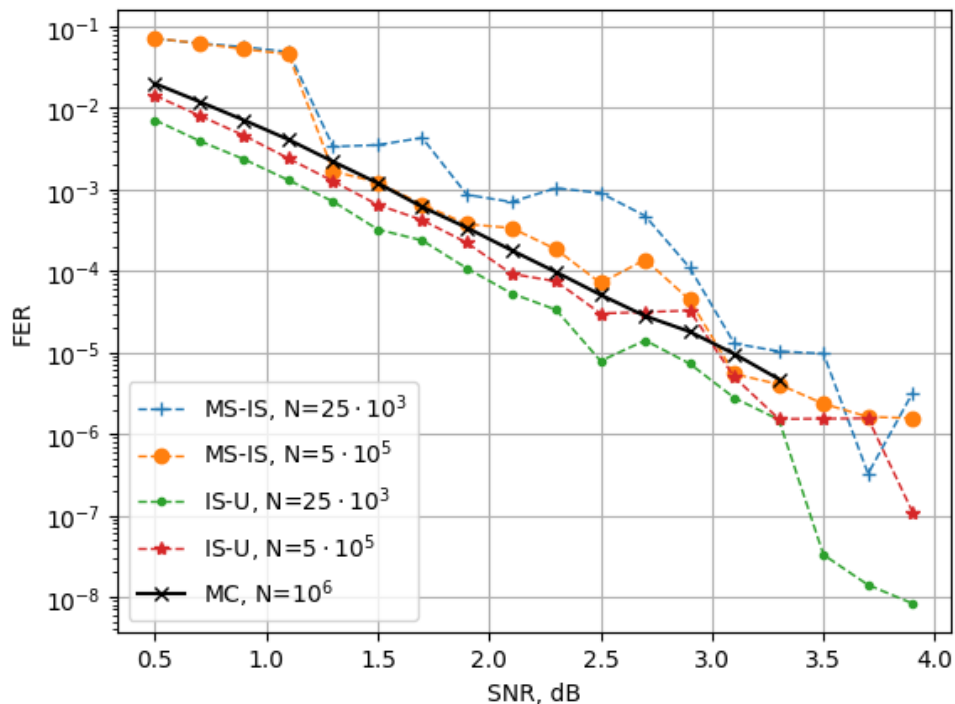
$$\alpha = -\frac{\ln|F - F'|}{\ln N}$$

$$r = \frac{\alpha(\text{uniform pdf})}{\alpha(\text{normal pdf})}$$

P.S.: Actually,  $F$  must satisfy the following condition

$$N^\alpha(F(X_1, \dots, X_N) - p) \xrightarrow[N \rightarrow \infty]{d} \eta$$

# Numerical Results: matrix (96, 48)

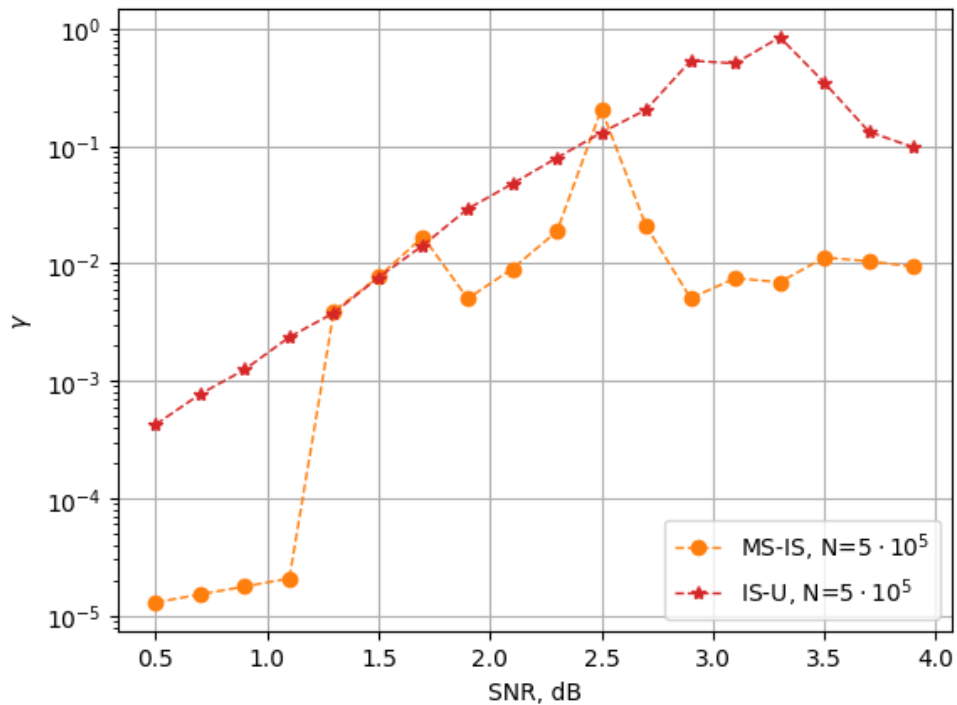


# MSE: matrix (96, 48)

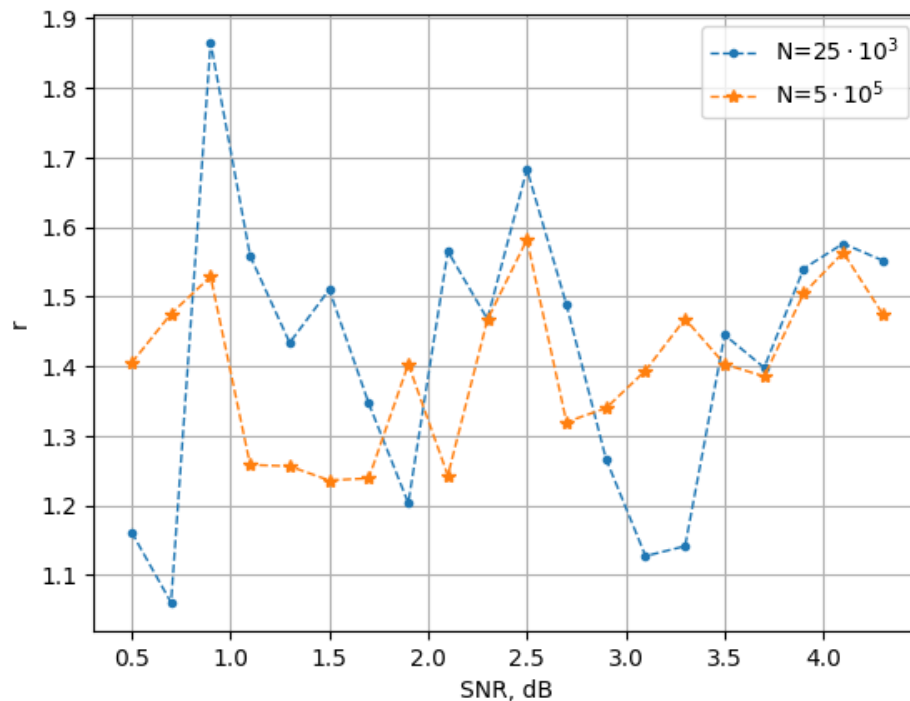
Estimation	MS-IS, $N = 25 \cdot 10^3$	IS-U, $N = 25 \cdot 10^3$	MS-IS, $N = 5 \cdot 10^5$	IS-U, $N = 5 \cdot 10^5$
MSE from 0.5 dB	9.15	3.83	3.56	0.87
MSE from 1.3 dB	6.37	2.92	0.78	0.64



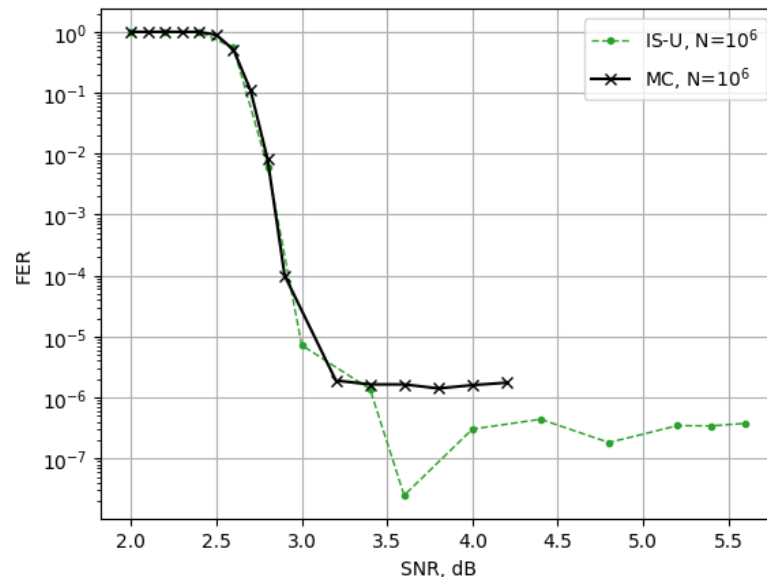
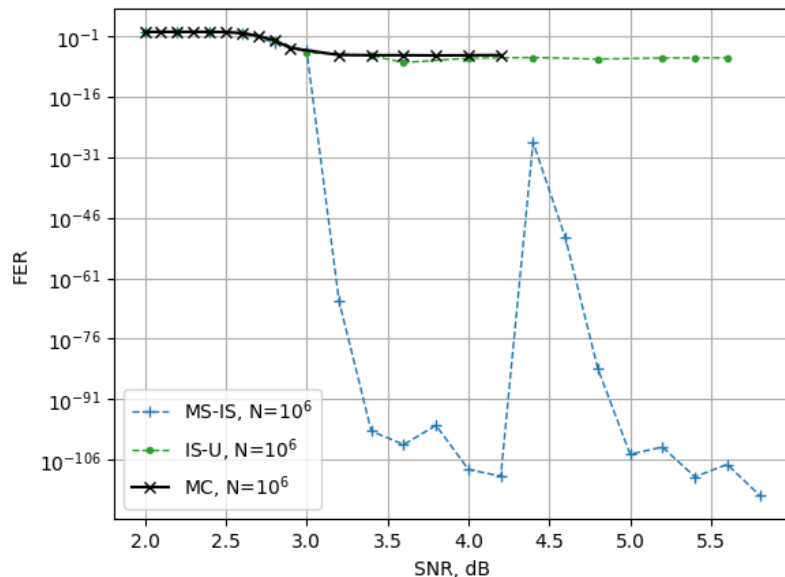
# Gamma: matrix (96, 48)



# New criterion: matrix (96, 48)



# Numerical Results: matrix (17k, 3k)



# Conclusion

- The comparison showed that IS-U provide more accurate estimations than MS-IS
- New criterion allows us to compare estimations without MC values and regardless of the estimator's structure

# Q/A



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