Methodology

Panel Data: a theoretical background

This study uses the panel data methodology. Panel is a common approach to address the CFP-CEP nexus [@Albertini2013]. Furthermore, it is considered to be one of the most efficient analytical methods for data analysis [@DimitriosAsteriou2006]. It usually contains more degrees of freedom, less collinearity among the variables, more efficiency and more sample variability than one-dimensional method (i.e.cross-sectional data and time series data) giving a more accurate inference of the parameters estimated in the model [@Hsiao2007]. @Roberts2013 also argued that using panel data offers a partial solution to the problem of omitted variables in econometric model, namely the most common causes of endogeneity in empirical corporate finance. Panel data takes the following econometric form:

$$Y_{it} = \alpha + \beta X_{it} + u_{it} \tag{1}$$

Panel data, also called longitudinal data, includes observations on i=1,...,n cross section units (e.g. firms) over t=1,...,T time-periods [@Hsiao2007]. Here Y_{it} is the dependent variable, X_{it} represents a K-dimensional row vectors of independent variables, α is the intercept, β is a K-dimensional column vectors of parameters and u_{it} is the random disturbance term of mean equals zero. The latter can be decomposed as $u_{it} = \mu_i + \epsilon_{it}$. The first term, μ_i , represents the individual error components and do not change over time. It can be considered as the unobserved effect model. The second term, ϵ_{it} , is the idiosyncratic error which is assumed well-behaved and independent of X_{it} and μ_i .

The starting point of all panel data is to determine if μ_i is correlated with X_{it} . In presence of correlation, then μ_i is considered as the *Fixed Effect* (i.e. FE) and the initial equation 1 is now described as the equation 2. Else, μ_i is considered as the *Random Effect* (i.e. RE) and the equation 1 becomes equation 3.

$$Y_{it} = (\alpha + \mu i) + \beta X_{it} + \epsilon_{it} \tag{2}$$

$$Y_{it} = \alpha + \beta X_{it} + (\epsilon_{it} + \mu i) \tag{3}$$

Fixed (i.e. Equation 2) and Random (i.e. Equation 3) Effect Model implies that the Ordinary Least Square (i.e. OLS) estimator of β are inconsistent. Five assumptions are required to produce consistent estimators with OLS: (i) a random sample of observations on y and $(x_1, ..., x_n)$, (ii) a random sample of n observations, (iii) no linear relationship among the explanatory variables, (iv) an error term that is uncorrelated with each explanatory variables and (v) an error term with zero mean conditional on the explanatory variables. FE Model violates the fourth assumption while RE model implies that the common error component over individuals induces correlation across the composite error terms making the third assumption violated [@Croissant2008].

While OLS is not consistent to estimate panel data model, the R package *plm* provides pertinent estimation methods. (i) The pooled ols estimation ignores the panel structure of the data and apply the same cofficients to each individual [@Schmidheiny2015]. (ii) The random effects estimation is the feasible Generalized Least Squares (i.e. GLS) estimator. (iii) The fixed effects estimation, also called within estimation, transforms the original equation 1 in substracting the time averages to every variables, such as:

$$Y_{it} - \bar{Y}_i = \beta (X_{itk} - \bar{X}_{ik}) + (u_{it} - \bar{u}_i)$$
(4)

The presence of RE model in a panel data can be tested using the Breusch-Pagan Lagrange Multiplier (i.e. BPLM) test [@Breusch1980]. This test is represented by the plmtest function in R. It examines if time and/or individual specific variance components equal zero [@Park2011]. If Ho is verified, then there is no RE model in the panel data. The presence of FE model is tested by a F test (i.e. the function pFtest in R). The

latter tests the individual and/or time effects based on the comparison of the within and the pooling model [@Croissant2008]. If Ho is verified, then there is no FE model in the panel data.

In case of the absence of both RE and FE model, namely $\mu_i = 0$, pooled ols estimation is the most efficient estimator [@Croissant2008]. Under the assumption of the FE model, the random effects estimators are biased and inconsistent given that μ_i is omitted and potentially correlated with other regressors. Therefore, the fixed effects estimation need to be used. Under the assumption of the RE model, both FE and RE estimators are unbiased and consistent. According to @Schmidheiny2015, scholars should prefer the RE estimator only and only if $E[\mu_i, X_i] = 0$. This precondition is tested by the Hausman test [@Hausman1981]. If Ho is verified, then scholars should use RE estimator.

Econometric Model

This study uses the equation 5 to analyze the link between outcome-based and process-based CEP and the equation 6 to test the combined effect of process and outcome-based CEP on CFP (short term and long term).

$$Y_{it} = \alpha + \beta_1 SPL_{it} + \beta_2 STC_{it} + \beta_3 A_{it} + u_{it}$$

$$\tag{5}$$

where Y_{it} is a proxy of outcome-based CEP measured as carbon productivity, water productivity and waste productivity, SPL_{it} is a proxy for a firm's sustainability pay link, STC_{it} is a proxy for a firm's sustainability themed commitment, A_{it} is a proxy for a firm's audit score and u_{it} which is the error term.

$$Y_{it+1} = \alpha + \beta_1 SPL_{it}$$

$$+\beta_2 STC_{it} + \beta_3 A_{it}$$

$$+\beta_4 CaP_{it} + \beta_5 WatP_{it}$$

$$+\beta_6 WastP_{it} + Controls_{it}$$

$$+u_{it}$$

$$(6)$$

where Y_{it+1} is a proxy of CFP measured as ROA or Tobin's Q, SPL_{it} is a proxy for a firm's sustainability pay link, STC_{it} is a proxy for a firm's sustainability themed commitment, A_{it} is a proxy for a firm's audit score, CP_{it} is a proxy for a firm's carbon productivity, $WatP_{it}$ is a proxy for a firm's water productivity, $WatP_{it}$ is a vector of control variables that includes firm size, industry sector, financial leverage and growth and lastly u_{it} which is the error term.

Recent meta-analysis provided evidences that the CFP-CEP nexus is characterized by a bidirectional causality [@Orlitzky2001, @Orlitzky2003, @Wu2006, @Albertini2013, @Dixon-Fowler2013, @EndrikatMakingsenseconflicting2014, @Ludecadedebatenexus2014, @WangMetaAnalyticReviewCorporate2016, @Busch2018]. This could cause simultaneous causality between the dependent and the independent variable and lead to endogeneity concern in my model [@Sanchez-Ballesta2007, @Biorn2008, @Roberts2013]. In order to adress this issue I have lagged observations in independent and control variables one year behind financial performance. This increases the confidence of the direction of the relationship [@Hart1996, @Delmas2015, @MiroshnychenkoGreenpracticesfinancial2017] and in fine reduce the potential simultaneity bias.