



Derivatives Pricing Course

Lecture 6 – Monte Carlo methods. Advanced topics.

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Agenda

- Discretization Schemes, Convergence, and Stability
- Variance Reduction Techniques
- Strong path-dependency
- C++ implementation

Monte Carlo methods

Convergence and stability

- For the simple Black-Scholes model, SDE state variables could be expressed analytically in terms of independent increments of a Brownian motion, making path generation straightforward.
- In practice however we are often working with SDEs that do not permit closed-form solution.

Monte Carlo methods

Convergence and stability

- Consider a vector-valued SDE:

$$dX(t) = \mu(t, X(t))dt + \sigma(t, X(t))dW(t)$$

- Consider an equidistant time grid $\{0, \Delta, 2\Delta, \dots, m\Delta\}$.

Set $\hat{X}_i \triangleq \hat{X}(i\Delta)$. We say that the underlying approximation is *weakly consistent* if there exists a function $c(\Delta)$ with

$$\lim_{\Delta \downarrow 0} c(\Delta) = 0$$

such that $E \left(\left| E(\Delta^{-1}(\hat{X}_{i+1} - \hat{X}_i)) - \mu(i\Delta, \hat{X}_i) \right|^2 \right) \leq c(\Delta)$, and

$$E \left(\left| E \left(\Delta^{-1}(\hat{X}_{i+1} - \hat{X}_i)(\hat{X}_{i+1} - \hat{X}_i)^T \right) - \sigma(i\Delta, \hat{X}_i)\sigma(i\Delta, \hat{X}_i)^T \right|^2 \right) \leq c(\Delta),$$

for all $i = 0, \dots, m - 1$

Monte Carlo methods

Convergence and stability

- A concept related to consistency is the notion of *weak convergence*.
- We say that an approximate solution converges weakly at time $T = m\Delta$ with respect to a class \mathcal{C} of test functions $g: \mathbb{R}^p \rightarrow \mathbb{R}$ if

$$\lim_{\Delta \downarrow 0} \left| E(g(X(T))) - E(g(\hat{X}(T))) \right| = 0$$

for all $g \in \mathcal{C}$.

- C_P^l - functions with polynomially bounded derivatives of order $0, 1, \dots, l$.
We say that a scheme converges with *weak order* β if, for all $g \in C_P^{2(\beta+1)}$:

$$\left| E(g(X(T))) - E(g(\hat{X}(T))) \right| \leq c\Delta^\beta$$

for all $\Delta \in (0, \Delta_0)$, where Δ_0 and c are constants.

Monte Carlo methods

Examples – Euler-Maryama

- The most straightforward. Starting from $\hat{X}_0 = X(0)$,
$$\hat{X}_{i+1} = \hat{X}_i + \mu(i\Delta, \hat{X}_i)\Delta + \sigma(i\Delta, \hat{X}_i)(W(i\Delta + \Delta) - W(i\Delta)),$$
$$i = 0, 1, \dots, m - 1$$
- The most widely used and easy to implement.
- The scheme is weakly consistent.
- Given that the Euler scheme is fully explicit, our experience from finite difference methods suggests that the scheme may have stability problems – there are thus restrictions on how big a time step Δ can be used.

Monte Carlo methods

Examples – Log-Euler

- One potential problem with the pure Euler scheme is the fact that all increments are locally Gaussian, thereby implying a non-zero probability of \hat{X} crossing zero and becoming negative.

- Many SDEs, however, are known to produce only non-negative solutions.
For instance:

$$dX(t) = \sqrt{X(t)}dW(t), X(0) > 0$$

- Using the log-transform to the initial SDE:

$$\hat{X}_{i+1} = \hat{X}_i \exp \left(\left(\frac{\mu(t, \hat{X}_i)}{\hat{X}_i} - \frac{1}{2} \frac{\sigma(t, \hat{X}_i)^2}{\hat{X}_i^2} \right) \Delta + \frac{\sigma(t, \hat{X}_i)}{\hat{X}_i} Z_i \sqrt{\Delta} \right)$$

Monte Carlo methods

Bias vs. Error

- When we use an m -step discretization scheme in an n -path Monte Carlo run, we are exposed to two types of errors on the expectation we are trying to estimate:
 - i. the statistical Monte Carlo error e_s
 - ii. a bias e_b , originating from the discretization scheme
- Raising n will reduce the standard error, but will not affect the bias which can only be reduced by increasing the number of steps m .
- Assume that the discretization scheme has weak order β . Informally:

$$e_b = c_b \Delta^\beta$$

Monte Carlo methods

Bias vs. Error

- Also, we know that the variance of e_s :

$$\text{Var}(e_s) = \frac{c_s}{n}$$

- Consider a problem of minimizing MSE:

$$c_b^2 \Delta^{2\beta} + \frac{c_s}{n}$$

- Omitting several assumptions – when we increase or decrease our computing budget, it is thus reasonable to allocate resources in such a way that we keep the factor $n^{1/2}m^{-\beta}$ constant.

Monte Carlo methods

Variance Reduction Techniques

- The convergence of the Monte Carlo method is quite slow, of order $O(n^{-1/2})$.
- While there is little that can be done to improve the order itself, the constant multiplying $n^{-1/2}$ can be affected by a careful choice of the Monte Carlo estimator.
- Recall that the goal of the Monte Carlo method is to estimate some quantity μ as the sample mean of n i.i.d. random variables Y_1, \dots, Y_n where each Y_i has expectation $E(Y_i) = \mu$ and variance $\text{Var}(Y_i) = \sigma^2$

Monte Carlo methods

Variance Reduction Techniques

- Suppose we have available two sets of i.i.d. sequences $Y_{1,i}$ and $Y_{2,i}$ where $E(Y_{1,i}) = E(Y_{2,i}) = \mu$, but $\text{Var}(Y_{1,i}) = \sigma_1^2$ and $\text{Var}(Y_{2,i}) = \sigma_2^2$.
- Also suppose that the time it takes on a computer to generate individual samples $Y_{1,i}$ and $Y_{2,i}$ is τ_1 and τ_2 . Fixed computing time budget - τ .
- Thus we need to compare
$$\sigma_1^2 \tau_1 \text{ vs } \sigma_2^2 \tau_2$$
- For instance, a high-variance estimator may, in fact, be preferable to a low-variance estimator, provided that the former takes less time to compute than the latter.

Monte Carlo methods

Antithetic variates

- Assume that we are interested in estimating the expected value of a random variable $Y = G(Z)$, Z q -dimensional vector of independent Gaussian random variables.
- Rather than using the regular sample average estimator for $E(Y)$, consider instead using

$$\bar{Y}_n^a = n^{-1} \sum_{i=1}^n \frac{G(Z_i) + G(-Z_i)}{2}$$

- In other words, in addition to set Z_1, \dots, Z_n of Gaussian samples, we also effectively include the set $-Z_1, \dots, -Z_n$ in the Monte Carlo trial. We still must have

$$E(\bar{Y}_n^a) = E(Y)$$

Monte Carlo methods

Antithetic variates

- Also, as $G(Z)$ and $G(-Z)$ have identical variance

$$\text{Var}(\bar{Y}_n^a) = n^{-1} \left[\frac{1}{2} \text{Var}(Y) + \frac{1}{2} \text{Cov}(G(Z), G(-Z)) \right] = \frac{\text{Var}(Y)}{n} \frac{(1+\rho)}{2}$$

- Recalling that the regular sample average has variance

$$\text{Var}(\bar{Y}_n) = \frac{\text{Var}(Y)}{n}$$

- While use of antithetic variates can always be expected to lower the standard error, it is not necessarily more efficient.

Monte Carlo methods

Control variates

- While we may need to use Monte Carlo simulation to estimate the unknown mean of a random variable Y , there may be random variables “close” to Y with means that can be computed analytically.
- Formally, vector of control variates, ideally with a strong negative or positive correlation with Y

$$Y^c = (Y_1^c, \dots, Y_q^c)^T$$

- The mean is known to be

$$E(Y^c) = \mu^c = (\mu_1^c, \dots, \mu_q^c)^T$$

- Arbitrary constant vector

$$\beta = (\beta_1, \dots, \beta_q)^T$$

Monte Carlo methods

Control variates

- And consider forming the linear combination

$$X = Y - \beta^T (Y^c - \mu^c)$$

- Clearly

$$E(X) = E(Y) - \beta^T (E(Y^c) - \mu^c) = E(Y)$$

- Σ_{Y^c} - $q \times q$ covariance matrix of the vector Y^c
- Σ_{Y,Y^c} - q -dimensional vector of covariances between Y and Y^c
- Variance of X

$$\text{Var}(X) = \text{Var}(Y) - 2\beta^T \Sigma_{Y,Y^c} + \beta^T \Sigma_{Y^c} \beta$$

- Lemma 6.1. The variance of X is minimized at

$$\beta^* = \Sigma_{Y^c}^{-1} \Sigma_{Y,Y^c}$$

with minimum value $(1 - R^2)\text{Var}(Y)$

Monte Carlo methods

Importance Sampling

- For a given measure P , consider estimating

$$\mu = E^P(Y)$$

- Let \hat{P} be a measure equivalent to P

$$\mu = E^{\hat{P}}(Y/R)$$

where R is the Radon-Nikodym derivative

$$R = \frac{d\hat{P}}{dP}$$

- It is possible that variance of Y/R under measure \hat{P} is lower than the variance of Y under P .

Monte Carlo methods

Importance Sampling – Density Formulation

- Specifically, let us assume that Y can be represented as $g(X)$, where $g: \mathbb{R}^p \rightarrow \mathbb{R}$ is a well-behaved function and X is p – dimensional with probability density $f: \mathbb{R}^p \rightarrow \mathbb{R}$

$$\mu = E^P(g(X)) = \int g(x)f(x)dx$$

to which corresponds an estimator

$$\bar{\mu}_n = \frac{1}{n} \sum_{i=1}^n g(X_i)$$

- Let h be another density, such that $h(x) > 0$ where $f(x) > 0$

$$\mu = \int g(x) \frac{f(x)}{h(x)} h(x) dx$$

Monte Carlo methods

Importance Sampling – Density Formulation

- So-called *likelihood ratio* $l(x) = f(x)/h(x)$. New estimator

$$\bar{\mu}_n^h = \frac{1}{n} \sum_{i=1}^n g(X_i) \frac{f(X_i)}{h(X_i)}$$

- Investigate the variances

$$\begin{aligned} \text{Var}(\bar{\mu}_n^h) &= \frac{1}{n} \left[E^{\hat{P}} \left(g(X)^2 \frac{f(X)^2}{h(X)^2} \right) - \mu^2 \right] \\ &= \frac{1}{n} \left[E^P \left(g(X)^2 \frac{f(X)}{h(X)} \right) - \mu^2 \right] \end{aligned}$$

and

$$\text{Var}(\bar{\mu}_n) = \frac{1}{n} [E^P(g(X)^2) - \mu^2]$$

Monte Carlo methods

Importance Sampling – Density Formulation

- A good choice of likelihood density will sample in proportion to f and g .
- That is, values of X where both density $f(X)$ and the payout $g(X)$ are high should be assigned a high value of $h(X)$ (high “importance”), and values of X where either $f(X)$ or $g(X)$ (or both) are low should be assigned a low value of $h(X)$.

Monte Carlo methods

Arithmetic Asian option

- Consider a dividend-free stock S , with BS dynamics:

$$dS(t)/S(t) = r dt + \sigma dW(t)$$

- Let there be given an increasing set of observations times $\{t_1, t_2, \dots, t_m\}$, with $t_m = T$, and define the stock average:

$$A(T) = \frac{1}{m} \sum_{i=1}^m S(t_i)$$

- An *Asian* call option with strike K is defined by the terminal payout

$$g(T) = (A(T) - K)^+$$

we wish to price this option by Monte Carlo simulation.

Monte Carlo methods

Arithmetic Asian option

- In this model stock price can be written explicitly:

$$S(t) = S(0)e^{rt - \frac{1}{2}\sigma^2 t + \sigma W(t)}, t > 0$$

whereby, with $\Delta_i \triangleq t_i - t_{i-1}$ and $t_0 = 0$

$$S(t_i) = S(t_{i-1})\exp\left(\left[r - \frac{1}{2}\sigma^2\right]\Delta_i + \sigma[W(t_i) - W(t_{i-1})]\right)$$

- Recall the properties of a Brownian motion:

$$S(t_i) = S(t_{i-1})\exp\left(\left(r - \frac{1}{2}\sigma^2\right)\Delta_i\right)\exp(\sigma\sqrt{\Delta_i}Z_i)$$

- To produce a single sample draw of $g(T)$, we thus
 1. Draw independent standard Gaussian samples Z_i .
 2. Starting from $S(0)$, generate $S(t_i)$ from the iteration.
 3. Compute $g(T)$

Monte Carlo methods

Arithmetic Asian option

- Repeating the procedure n times, we can generate random samples g_1, g_2, \dots, g_n of $g(T)$.

- Therefore our estimate of an option value:

$$\bar{V}(0) = e^{-rT} \frac{1}{n} \sum_{j=1}^n g_j$$

- Computational cost of the pricing algorithm - $O(mn)$.
- Now consider a p -dimensional basket of stocks S_1, S_2, \dots, S_p each following geometric Brownian motion.
- The Brownian motions W_k and W_j are assumed to be correlated with constant correlation coefficient $\rho_{j,k}, j \neq k$.
- Define a unit-weighted basket price as

$$B(t) = \sum_{k=1}^p S_k(t)$$

Monte Carlo methods

Basket Asian option

- And set the terminal Asian option payout to be

$$g(T) = \left(\frac{1}{m} \sum_{i=1}^m B(t_i) - K \right)^+$$

- We draw sample paths

$$S_k(t_i) = S_k(t_{i-1}) \exp \left(\left(r - \frac{1}{2} \sigma_k^2 \right) \Delta_i \right) \exp(\sigma_k \sqrt{\Delta_i} Z_{k,i})$$

- where $Z_{k,i}$ are Gaussian samples, independently drawn at each time step but correlated across k 's.
- Let C be the Cholesky decomposition of the correlation matrix, in which case we can generate the correlated sample vectors $Z_i = (Z_{1,i}, Z_{2,i}, \dots, Z_{p,i})^T$ as $Z_i = CX_i$ for vector X_i of independent Gaussian samples.

Monte Carlo methods

Basket Asian option

- The last steps are the same as for an Asian option on a single stock.
- The total computational effort of an n -sample Monte Carlo scheme is $O(nmp)$, with the convergence order $O(n^{-\frac{1}{2}})$ and dependent only on n .

C++ code

Random(.h) (1/2)

```
#ifndef RANDOM_H
#define RANDOM_H

//Abstract class - define the interface, not all member functions are pure virtual
class RandomGenerator{
public:
    //Simulate array of standard normal variables
    virtual void getStdNormalInverse(std::vector<double>& arrayToFillGauss);
    //Simulate array of uniform variables
    virtual void getUniform(std::vector<double>& arrayToFillUni) = 0;
    int getSize() { return Size; }
    void setSize(int newSize) { Size = newSize;}
    virtual RandomGenerator* clone() const = 0;
    RandomGenerator(unsigned long Size_): Size(Size_){}

private:
    //Dimension of the array
    int Size;
};

class BasicRandomGenerator : public RandomGenerator{
public:
    virtual void getUniform(std::vector<double>& arrayToFillUni);
    BasicRandomGenerator(unsigned long Seed_, int Size_ = 1) : RandomGenerator(Size_), Seed(Seed_){}
    virtual RandomGenerator* clone() const;
    void resetSeed();

private:
    //Seed of the random generator
    unsigned long Seed;
};
```

C++ code

Random(.cpp) (1/2)

```
//Using random engine from boost
#include <boost/math/distributions/inverse_gaussian.hpp>
#include <boost/random.hpp>
void RandomGenerator::getStdNormalInverse(std::vector<double>& arrayToFillGauss)
{
    //Gaussian inverse function from boost
    boost::math::normal distr(0.0, 1.0);
    std::vector<double> arrayToFillUni;
    arrayToFillUni.resize(arrayToFillGauss.size());

    getUniform(arrayToFillUni);
    for (int i = 0; i < arrayToFillGauss.size(); i++){
        arrayToFillGauss[i] = quantile(distr, arrayToFillUni[i]);
    }
    Size = arrayToFillGauss.size();
}

void BasicRandomGenerator::getUniform(std::vector<double>& arrayToFillUni)
{
    //What happens if they are not static?
    //Uniform engine (Mersenne twister) from boost
    static boost::mt19937 generator(Seed);
    static boost::uniform_real<> range(0, 1);
    static boost::variate_generator<boost::mt19937&, boost::uniform_real<> > uniform(generator, range);
    for (int i = 0; i < arrayToFillUni.size(); i++){
        arrayToFillUni[i] = uniform();
    }
}
```

C++ code

PathDependentAsian(.h)

```
#ifndef PATHDEPENDENTASIAN_H
#define PATHDEPENDENTASIAN_H

#include "Bridge.h"
#include <vector>

class PathDependentAsian{
public:
    PathDependentAsian(double Time_, const Bridge& ThePayOff_, std::vector<double> averagingTimes_);
    double optionPayoff(double Spot) const;
    double getTime() const;
    std::vector<double> getAveragingTimes() const;

private:
    double Time;
    Bridge thePayoff;
    std::vector<double> averagingTimes; //at these dates from today underlying values will be fixed
};
#endif
```

C++ code

PathDependentAsian(.cpp)

```
PathDependentAsian::PathDependentAsian(double Time_, const Bridge& ThePayOff_, std::vector<double> averagingTimes_):
Time(Time_), thePayoff(ThePayOff_)
{
    averagingTimes.swap(averagingTimes_);
}

double PathDependentAsian::optionPayoff(double Spot) const{
    return thePayoff(Spot);
}

double PathDependentAsian::getTime() const{
    return Time;
}

std::vector<double> PathDependentAsian::getAveragingTimes() const{
    return averagingTimes;
}
```

C++ code

ExoticMonteCarlo(.h)

```
#include "PathDependentAsian.h"
#include "Random.h"

double* ExoticMC(const PathDependentAsian& theOption,
                 RandomGenerator& theGenerator,
                 double Spot,
                 double Vol,
                 double Rate,
                 int numberOfPaths);
```

C++ code

ExoticMonteCarlo(.cpp) (1/2)

```
#include "ExoticMonteCarlo.h"
```

```
double* ExoticMC(const PathDependentAsian& theOption,  
    RandomGenerator& theGenerator,  
    double Spot,  
    double Vol,  
    double Rate,  
    int numberOfPaths)  
{  
    double* result = new double[1];  
    Vol = Vol / 100.0; // in %  
    Rate = Rate / 100.0; // in %  
    //Euler scheme  
    //choose time step equal to difference between averaging dates  
    int numberOfDates = theOption.getAveragingTimes().size();  
  
    double sum = 0.0;
```

C++ code

ExoticMonteCarlo(.cpp) (2/2)

```
//outer loop - Paths
for (int i = 0; i < numberOfPaths; i++){
    double averagePath = 0.0;
    double lastSpot = Spot;
    std::vector<double> gaussianVector;
    gaussianVector.resize(numberOfDates - 1);
    theGenerator.getStdNormalInverse(gaussianVector);
    //inner loop - sum underlying values on each date
    for (int j = 1; j < numberOfDates; j++){
        double precomputedSpot = lastSpot*exp((Rate - 0.5*Vol*Vol)*(theOption.getAveragingTimes()[j] -
                                                                    theOption.getAveragingTimes()[j - 1]));
        double stdDeviation = Vol * sqrt(theOption.getAveragingTimes()[j] - theOption.getAveragingTimes()[j - 1]);
        double thisSpot = precomputedSpot*exp(stdDeviation*gaussianVector[j - 1]);
        lastSpot = thisSpot;
        averagePath += thisSpot;
    }
    averagePath = averagePath / (numberOfDates - 1);
    double payoffCalc = theOption.optionPayoff(averagePath);
    sum += payoffCalc;
}

result[0] = (sum / numberOfPaths) * exp(-Rate * theOption.getTime());

return result;
}
```

C++ code

BS_Pricer(.h)

```
const int numberOfPaths = 10000000;

//Create option object, define payoff
PayoffCall thePayoff(Strike);

//State averaging times
//Daily - 252 business days

int length = round(Time * 252);

std::vector<double> averagingTimes;
averagingTimes.resize(length);

for (int i = 0; i < length; i++){
    averagingTimes[i] = i * (Time / (length - 1));
}

PathDependentAsian theOption(Time, thePayoff, averagingTimes);

BasicRandomGenerator theGenerator(121);
//Decorated generator
AntiThetic theDecoratedGenerator(&theGenerator);

double* result = ExoticMC(theOption, theDecoratedGenerator, Spot, Vol, Rate, numberOfPaths);
```


Monte Carlo methods

Antithetic variates results - variance

	N=100	N=400	N=900
Simple MC	0.9925	0.5520	0.2644
Antithetic	0.5263	0.2385	0.0983

Homework assignment 5*

- Modify program to price Geometric Weekly Asian Put option on a Basket of two stocks.