

Example on Pulmonary Disease

- Topic: passive smoking and pulmonary health
- Information on pulmonary function was collected in 6 groups:
 - 1) Nonsmokers (NS): did not smoke
 - 2) Passive smokers (PS): in enclosed working area routinely contained tobacco smoke
 - 3) Non-inhaling smokers (NI): smoked pipes, cigars, or cigarettes (did not inhale)
 - 4) Light smokers (LS): smoked and inhaled 1-10 cigarettes per day for 20+ years
 - 5) Moderate smokers (MS): ...11-39 cigarettes.....
 - 6) Heavy smokers (HS)....40+ cigarettes......
- Measured forced mid-expiratory flow (FEF) and compare mean FEF among the 6 groups
- Q: how can the means of these six groups be compared?

One-Way ANOVA—Fixed-Effects Model

Suppose: k groups of n_i observations in the ith group.

- y_{ij}: jth observation in the ith group
- Model: $y_{ij} = \mu + \alpha_i + e_{ij}$
 - μ: constant
 - \circ α_i : constant specific to the *i*th group
 - o e_{ij} : error term (normally distributed with mean 0 and variance σ^2)
- A typical observation from the *i*th group is normally distributed with mean μ + α_i and variance σ^2

Table 12.1 FEF data for smoking and nonsmoking males

Group number,		Mean FEF	sd FEF	
i	Group name	(L/s)	(L/s)	n_i
1	NS	3.78	0.79	200
2	PS	3.30	0.77	200
3	NI	3.32	0.86	50
4	LS	3.23	0.78	200
5	MS	2.73	0.81	200
6	HS	2.59	0.82	200

Source: Reprinted by permission of The New England Journal of Medicine, 302(13), 720-723, 1980.

Q: How do you compare the means of the six groups?

One-way analysis of variance (one-way ANOVA) model:

- means of an arbitrary number of groups
- each group follows a normal distribution with the same variance
- can determine if the variability in the data comes mostly from variability within groups or can truly be attributed to variability between groups

Interpretation of the parameters of a one-way ANOVA fixed-effects model

- 1.µ: underlying mean of all groups
- $2.\alpha_i$: difference between mean of the *i*th group and the overall mean
- 3.e_{ij}: random error about the mean μ + α _i for an individual observation from *i*th group

Hypothesis Testing in One-Way ANOVA— Fixed-Effects Model H₀: all $\alpha_i = 0$

F test for overall comparison of group means

$$y_{ij} - \overline{y} = (y_{ij} - \overline{y_i}) + (\overline{y_i} - \overline{y})$$

 H_0 : all $\alpha_i = 0$ vs. H_1 at least one $\alpha_i \neq 0$

 $(y_{ij} - y_i)$, within-group variability: deviation of an individual observation from the group mean for that observation

 $(\overline{y_i} - \overline{y})$, between-group variability: deviation of a group mean from the overall mean

- if between-group variability is large + within-group variability is small:
 - $\circ \rightarrow reject H_0$ underlying group means are significantly different
- if between-group variability is small + within-group variability is large:
 - \circ \rightarrow accept H_0 : underlying group means are the same

Squared and summation of the squared deviations:

$$\sum_{i=1}^{k} \sum_{j=1}^{n_i} \left(y_{ij} - \overline{\overline{y}} \right)^2 = \sum_{i=1}^{k} \sum_{j=1}^{n_i} \left(y_{ij} - \overline{y}_i \right)^2 + \sum_{i=1}^{k} \sum_{j=1}^{n_i} \left(\overline{y}_i - \overline{\overline{y}} \right)^2$$

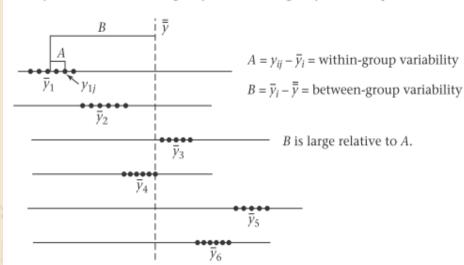
Total Sum of Squares (Total SS): $\sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \overline{y})^2$

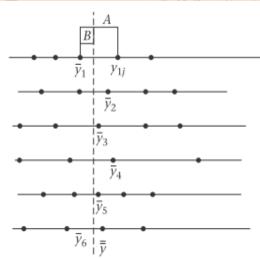
Within Sum of Squares (Within SS): $\sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2$

Between Sum of Squares (Between SS): $\sum_{i=1}^{k} \sum_{j=1}^{n_i} (\bar{y}_i - \bar{y})^2$

Total SS = Between SS + Within SS

Comparison of between-group and within-group variability





B is small relative to A.

Reject H₀

Accept H₀

Short computational form for the Between SS and Within SS

Between SS =
$$\sum_{i=1}^{k} n_i \overline{y}_i^2 - \frac{\left(\sum_{i=1}^{k} n_i \overline{y}_i\right)^2}{n} = \sum_{i=1}^{k} n_i \overline{y}_i^2 - \frac{y_{...}^2}{n}$$
Within SS =
$$\sum_{i=1}^{k} (n_i - 1) s_i^2$$

- y..= sum of the observations across all groups (grand total of all observations over all groups)
- n = total number of observations over all groups

Between Mean Square = Between MS = Between SS/(k-1) Within Mean Square = Within MS = Within SS/(n-k)

- Significance test: ratio of Between MS to Within MS
 - If this ratio is large \rightarrow reject H_0
 - If it is small \rightarrow accept H_0
 - Ounder H_0 : the ratio follows an F distribution with k-1 (numerator) and n-k (dominator) df

Overall F test for One-way ANOVA Procedure

 H_0 : $\alpha_i = 0$ for all i

 H_1 : at least one $\alpha_i \neq 0$

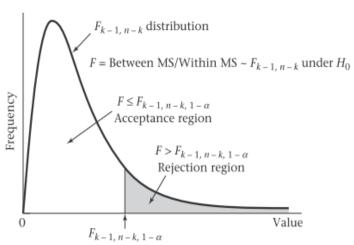
- 1. Compute Between SS, Between MS, Within SS, and Within MS
- 2. Compute test statistic F = Between MS/Within MS, which follows an F

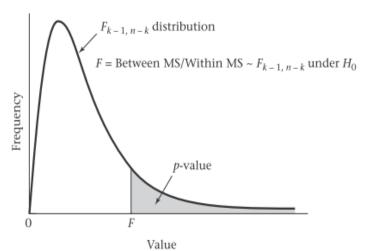
distribution with k - 1 and n - k df under H_0

- If $F > F_{k-1,n-k,1-\alpha} \rightarrow \text{reject } H_0$
- If $F \le F_{k-1,n-k,1-\alpha} \to \text{accept } H_1$
- 3. Exact *p*-value = area to the right of *F* under an $F_{k-1,n-k}$ distribution = $Pr(F_{k-1,n-k} > F)$

Acceptance and rejection regions for the overall F test for one-way ANOVA







Display of one-way ANOVA results

Source of variation	SS	df	MS	F statistic	<i>p</i> -value
Between	$\sum_{i=1}^k n_i \overline{y}_i^2 - \frac{y_{}^2}{n} = A$	k – 1	$\frac{A}{k-1}$	$\frac{A/(k-1)}{B/(n-k)} = F$	$Pr(F_{k-1,n-k} > F)$
Within	$\sum_{i=1}^{k} (n_i - 1)s_i^2 = B$	n – k	$\frac{B}{n-k}$		

Total Between SS + Within SS

To test whether the mean FEF scores differ significantly among the six groups:

- 1. Calculate the Between MS and Within MS.
- 2. Determine F = Between MS/Within MS
- 3. Find the F value (with the F-table / statistical software)
- 4. If $p < 0.05 \rightarrow \text{reject } H_0$ (all means are equal) \rightarrow not all means are equal
- 5. Conclusion: at least two of the means are significantly different

ANOVA tabl	e for FEF data i	n Table 12.1			
	SS	df	MS	F statistic	p-value
Between	184.38	5	36.875	58.0	p < .001
Within_ Total	663.87 848.25	1044	0.636		

R commands to perform one-way ANOVA

#use aov command

>model=aov(depvar ~ groupvar)

>summary(model)

Examples on One-way ANOVA – Pulmonary Disease

 Question: Compute the Within SS and Between SS for the FEF data in Table 12.1

TABLE 12.1 FEF data for smoking and nonsmoking males

Group number, <i>i</i>	Group name	Mean FEF (L/s)	sd FEF (L/s)	n_i
1	NS	3.78	0.79	200
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Examples on One-way ANOVA – Pulmonary Disease

Solution:

We calculate the following:

Between
$$SS = [200(3.78)^2 + 200(3.30)^2 + \dots + 200(2.59)^2]$$

$$\frac{[200(3.78) + 200(3.30) + \dots + 200(2.59)]^2}{1050}$$

$$= 10,505.58 - 3292^2/1050$$

$$= 10,505.58 - 13,321.20$$

$$= 184.38$$

Between SS =
$$\sum_{i=1}^{k} n_i \overline{y}_i^2 - \frac{\left(\sum_{i=1}^{k} n_i \overline{y}_i\right)^2}{n} = \sum_{i=1}^{k} n_i \overline{y}_i^2 - \frac{y_{\cdot \cdot}^2}{n}$$
Within SS =
$$\sum_{i=1}^{k} (n_i - 1) s_i^2$$

Within
$$SS = 199(0.79)^2 + 199(0.77)^2 + 49(0.86)^2 + 199(0.78)^2 + 199(0.81)^2 + 199(0.82)^2$$

$$= 124.20 + 117.99 + 36.24 + 121.07 + 130.56 + 133.81$$

$$= 663.87$$

Examples on One-way ANOVA – Pulmonary Disease

 Question: Test whether the mean FEF scores differ significantly among the six groups in Table 12.1

Solution:

Between Mean Square = Between MS = Between SS/(k-1)
Within Mean Square = Within MS = Within SS/(n-k)

Between SS = 184.38 and Within SS = 663.87.

Therefore, because there are 1050 observations combined over all 6 groups, it follows that

Between MS = 184.38 / 5 = 36.875

Within MS = 663.87 / (1050 - 6) = 663.87 / 1044 = 0.636

 $F = Between MS / Within MS = 36.875 / 0.636 = 58.0 \sim F_{5,1044} under H_0$

 $F = 58 > F_{5.1044} = 4.10 \rightarrow \text{reject H}_0$

Conclusion: at least two of the means are significantly different



	101	Percentage points of the F distribution ($F_{d_1d_2p}$)										
df for enominato						df for	numerato	r, d ₁				
d_2	", P	1	2	3	4	5	6	7	8	12	24	00
1	.90	39.86										
	.95	161.4	199.5	215.7	224.6 899.6	230.2	234.0	236.8	238.9	243.9	249.1	254.
	.975 .99	647.8 4052.	799.5 5000.	864.2 5403.	5625.	921.8 5764.	937.1 5859.	948.2 5928.	956.7 5981.	976.7 6106.	997.2 6235.	1018. 6366.
	.995	16211.	20000.	21615.	22500.	23056.	23437.	23715.	23925.	24426.	24940.	25464.
		405280.						592870.				
2	.90	8.53										
	.95 .975	18.51 38.51										
	.99	98.50										
	.995	198.5	199.0	199.2	199.2	199.3	199.3	199.4	199.4	199.4	199.5	199.
	.999	998.5	999.0	999.2	999.2	999.3	999.3	999.4	999.4	999.4	999.5	999.
3	.90	5.54										
	.95	10.13										
	.975	17.44										
	.99	34.12										
	.995	167.00) 47.47 141.1	46.20 137.1	134.6	9 44.84 132.8	4 44.43 131.6	3 44.13 130.6	3 43.39 128.3	9 42.62 125.9	2 41 123
4	.90	4.54										
4	.95	7.71										
	.975	12.22										
	.99	21.20										
	.995	31.33				22.46	3 21.9					
	.999	74.14	61.25	56.18	53.44	51.71	50.5	3 49.66	49.00	47.41	45.77	44
5	.90	4.06										
	.95	6.61										
	.975 .99	10.01										
	.995	16.26 22.78										
	.999	47.18										
6	.90	3.78	3.46			3.11			2.98	3 2.90	2.82	2 2
	.95	5.99										
	.975	8.81										
	.99 .995	13.75 18.64										
	.999	35.51										
7	.90	3.59	3.26	3.07	2.96	2.88	3 2.8	3 2.78	3 2.75	2.67	2.58	3 2
	.95	5.59	4.74	4.35	4.12	3.97					3.41	
	.975	8.07										
	.99	12.25										
	.995	16.24 29.25										
8	.90 .95	3.46 5.32										
	.975	7.57										
	.99	11.26										
	.995	14.69										
	.999	25.42										
9	.90	3.36										
	.95	5.12										
	.975	7.21										
	.99 .995	10.56 13.61										
	.999	22.86			12.56					7 9.57		
10	.90	3.29	2.92	2.73	2.61	2.52	2 2.4	8 2.41	2.38	3 2.28	3 2.18	3 2
	.95	4.96	4.10	3.71	3.48	3.33			3.07	7 2.91	2.74	1 2
	.975	6.94										
	.99	10.04										
	.995	12.83 21.04										
12	.90	3.18										
12	.95	4.75										
											2.0	_



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TABLE 8	Perce	ntage pol	nts of the	F distrib	ution (F _d	, _{dap}) (con	tinued)					
<i>df</i> for denominato	r. —					df for nun	nerator, o	l,				
d ₂	, b	1	2	3	4	5	6	7	8	12	24	00
	.99 .995 .999	9.33 11.75 18.64	6.93 8.51 12.97	5.95 7.23 10.80	5.41 6.52 9.63	5.06 6.07 8.89	4.82 5.76 8.38	4.64 5.52 8.00	4.50 5.35 7.71	4.16 4.91 7.00	3.78 4.43 6.25	3.36 3.90 5.42
14	.90 .95 .975 .99 .995	3.10 4.60 6.30 8.86 11.06 17.14	2.73 3.74 4.86 6.51 7.92 11.78	2.52 3.34 4.24 5.56 6.68 9.73	2.39 3.11 3.89 5.04 6.00 8.62	2.31 2.96 3.66 4.69 5.56 7.92	2.24 2.85 3.50 4.46 5.26 7.44	2.19 2.76 3.38 4.28 5.03 7.08	2.15 2.70 3.29 4.14 4.86 6.80	2.05 2.53 3.05 3.80 4.43 6.13	1.94 2.35 2.79 3.43 3.96 5.41	1.80 2.13 2.49 3.00 3.44 4.60
16	.90 .95 .975 .99 .995 .999	3.05 4.49 6.12 8.53 10.58 16.12	2.67 3.63 4.69 6.23 7.51 10.97	2.46 3.24 4.08 5.29 6.30 9.01	2.33 3.01 3.73 4.77 5.64 7.94	2.24 2.85 3.50 4.44 5.21 7.27	2.18 2.74 3.34 4.20 4.91 6.80	2.13 2.66 3.22 4.03 4.69 6.46	2.09 2.59 3.12 3.89 4.52 6.19	1.99 2.42 2.89 3.55 4.10 5.55	1.87 2.24 2.63 3.18 3.64 4.85	1.72 2.01 2.32 2.75 3.11 4.06
18	.90 .95 .975 .99 .995	3.01 4.41 5.98 8.29 10.22 15.38	2.62 3.55 4.56 6.01 7.21 10.39	2.42 3.16 3.95 5.09 6.03 8.49	2.29 2.93 3.61 4.58 5.37 7.46	2.20 2.77 3.38 4.25 4.96 6.81	2.13 2.66 3.22 4.01 4.66 6.35	2.08 2.58 3.10 3.84 4.44 6.02	2.04 2.51 3.01 3.71 4.28 5.76	1.93 2.34 2.77 3.37 3.86 5.13	1.81 2.15 2.50 3.00 3.40 4.45	1.66 1.92 2.19 2.57 2.87 3.67
20	.90 .95 .975 .99 .995	2.97 4.35 5.87 8.10 9.94 14.82	2.59 3.49 4.46 5.85 6.99 9.95	2.38 3.10 3.86 4.94 5.82 8.10	2.25 2.87 3.51 4.43 5.17 7.10	2.16 2.71 3.29 4.10 4.76 6.46	2.09 2.60 3.13 3.87 4.47 6.02	2.04 2.51 3.01 3.70 4.26 5.69	2.00 2.45 2.91 3.56 4.09 5.44	1.89 2.28 2.68 3.23 3.68 4.82	1.77 2.08 2.41 2.86 3.22 4.15	1.61 1.84 2.09 2.42 2.69 3.38
30	.90 .95 .975 .99 .995	2.88 4.17 5.57 7.56 9.18 13.29	2.49 3.32 4.18 5.39 6.35 8.77	2.28 2.92 3.59 4.51 5.24 7.05	2.14 2.69 3.25 4.02 4.62 6.12	2.05 2.53 3.03 3.70 4.23 5.53	1.98 2.42 2.87 3.47 3.95 5.12	1.93 2.33 2.75 3.30 3.74 4.82	1.88 2.27 2.65 3.17 3.58 4.58	1.77 2.09 2.41 2.84 3.18 4.00	1.64 1.89 2.14 2.47 2.73 3.36	1.46 1.62 1.79 2.01 2.18 2.59
40	.90 .95 .975 .99 .995	2.84 4.08 5.42 7.31 8.83 12.61	2.44 3.23 4.05 5.18 6.07 8.25	2.23 2.84 3.46 4.31 4.98 6.59	2.09 2.61 3.13 3.83 4.37 5.70	2.00 2.45 2.90 3.51 3.99 5.13	1.93 2.34 2.74 3.29 3.71 4.73	1.87 2.25 2.62 3.12 3.51 4.44	1.83 2.18 2.53 2.99 3.35 4.21	1.71 2.00 2.29 2.66 2.95 3.64	1.57 1.79 2.01 2.29 2.50 3.01	1.38 1.51 1.64 1.80 1.93 2.23
60	.90 .95 .975 .99 .995	2.79 4.00 5.29 7.08 8.49 11.97	2.39 3.15 3.93 4.98 5.80 7.77	2.18 2.76 3.34 4.13 4.73 6.17	2.04 2.53 3.01 3.65 4.14 5.31	1.95 2.37 2.79 3.34 3.76 4.76	1.87 2.25 2.63 3.12 3.49 4.37	1.82 2.17 2.51 2.95 3.29 4.09	1.77 2.10 2.41 2.82 3.13 3.86	1.66 1.92 2.17 2.50 2.74 3.32	1.51 1.70 1.88 2.12 2.29 2.69	1.29 1.39 1.48 1.60 1.69 1.89
120	.90 .95 .975 .99 .995 .999	2.75 3.92 5.15 6.85 8.18 11.38	2.35 3.07 3.80 4.79 5.54 7.32	2.13 2.68 3.23 3.95 4.50 5.78	1.99 2.45 2.89 3.48 3.92 4.95	1.90 2.29 2.67 3.17 3.55 4.42	1.82 2.17 2.52 2.96 3.28 4.04	1.77 2.09 2.39 2.79 3.09 3.77	1.72 2.02 2.30 2.66 2.93 3.55	1.60 1.83 2.05 2.34 2.54 3.02	1.45 1.61 1.76 1.95 2.09 2.40	1.19 1.25 1.31 1.38 1.43 1.54
00	.90 .95 .975 .99 .995	2.71 3.84 5.02 6.63 7.88 10.83	2.30 3.00 3.69 4.61 5.30 6.91	2.08 2.60 3.12 3.78 4.28 5.42	1.94 2.37 2.79 3.32 3.72 4.62	1.85 2.21 2.57 3.02 3.35 4.10	1.77 2.10 2.41 2.80 3.09 3.74	1.72 2.01 2.29 2.64 2.90 3.47	1.87 1.94 2.19 2.51 2.74 3.27	1.55 1.75 1.94 2.18 2.36 2.74	1.38 1.52 1.64 1.79 1.90 2.13	1.00 1.00 1.00 1.00 1.00 1.00



Comparisons of Specific Groups in One-Way ANOVA

- H₀: all group means are equal
 H₁: at least two group means are different
 - Do not know which of the groups have means that differ from each other
 - Overall F test: if reject H_0 → compare the specific groups

t Test for Comparison of Pairs of Groups

- test whether groups 1 and 2 have means that are significantly different from each other
- Under either hypothesis:
 - \circ Y₁ is normally distributed with mean μ+α₁ and variance σ^2/n_1
 - \circ \overline{Y}_2 is normally distributed with mean $\mu + \alpha_2$ and variance σ^2/n_2
- The difference of the sample means $(\overline{y_1} \overline{y_2})$ will be used as a test criterion:

$$\overline{Y}_1 - \overline{Y}_2 \sim N \left[\alpha_1 - \alpha_2, \sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2} \right) \right]$$

Reduce to:

$$\overline{Y}_1 - \overline{Y}_2 \sim N \left[0, \sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2} \right) \right]$$

• If σ^2 known, dividing by the standard error:

$$Z = \frac{\overline{Y}_1 - \overline{Y}_2}{\sqrt{\sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

- $Z \sim N(0,1)$ distribution under H_0
 - O Because σ² is unknown and can be estimated by s². The test statistic becomes:

$$S^2 = [(n_1-1)s_1^2 + (n_2-1)s_2^2]/(n_1+n_2-2)$$

- One-way ANOVA:
 - k sample variances
 - Estimate σ² by computing a weighted average of k individual sample variances
 - Weights are the df in each of k samples
 (degree of freedom: max. no. of logically independent values;
 values that have freedom to vary)

$$s^{2} = \sum_{i=1}^{k} (n_{i} - 1) s_{i}^{2} / \sum_{i=1}^{k} (n_{i} - 1) = \left[\sum_{i=1}^{k} (n_{i} - 1) s_{i}^{2} \right] / (n - k) = \text{Within MS}$$

Pooled estimate of the variance for one-way ANOVA

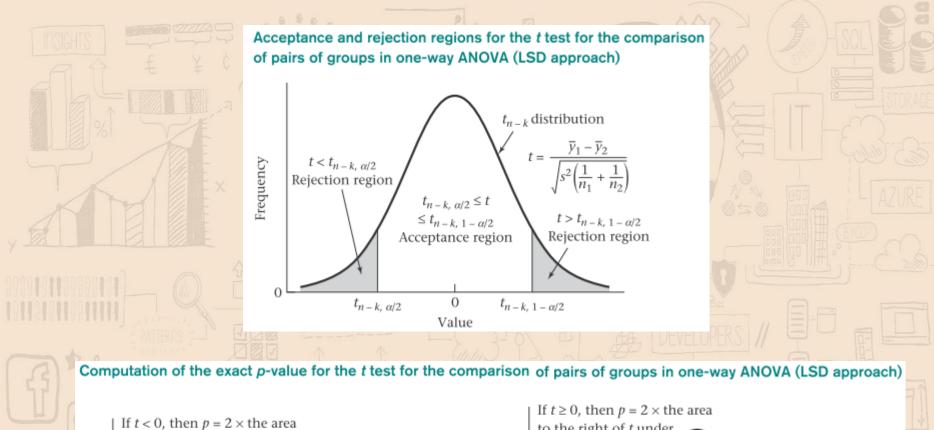
t Test for the Comparison of Pairs of Groups in One-Way ANOVA (LSD Procedure)

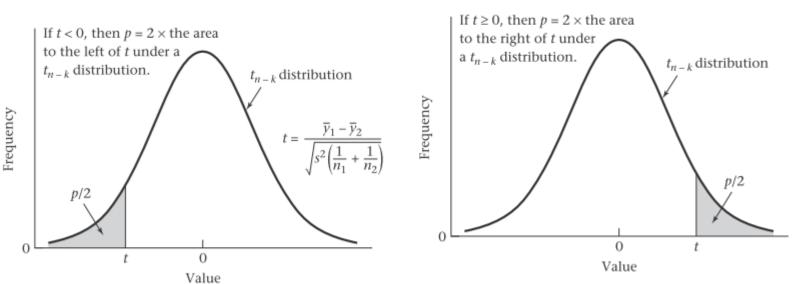
- Goal: suppose we wish to compare two specific groups (group 1 and group 2) among k groups
- H_0 : $\alpha_1 = \alpha_2$ vs. H_1 : $\alpha_1 \neq \alpha_2$

1. Compute the pooled estimate of variance s^2 = within MS from one way ANOVA

Compute the test statistic:
$$t = \sqrt{s^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

- 2. Compute the test statistic:
- which follows a t_{n-k} distribution under H_0
- 3. Two-sided level α test:
- reject H_0 : if $t > t_{n-k,1-\alpha/2}$ or $t < t_{n-k,\alpha/2}$
- accept H_0 if $t_{n-k,\omega/2} \le t \le t_{n-k,1-\omega/2}$
- 4. Exact *p*-value:
- $p = 2 \times \text{the area to the left of } t \text{ under a } t_{n-k} \text{ distribution if } t < 0$ $= 2 \times Pr(t_{n-k} < t)$
- $p = 2 \times \text{the areas to the right of } t \text{ under a } t_{n-k} \text{ distribution if } t \ge 0$ $= 2 \times Pr(t_{n-k} > t)$

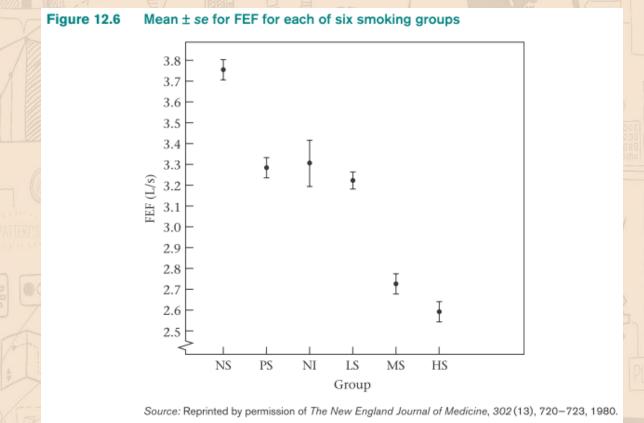




 Question: Compare each pair of groups for the FEF data in Table 12.1, and report any significant differences.

Solution:

- First plot the mean ± se of the FEF values for each of the six groups in Figure 12.6 to obtain some idea of the magnitude of the differences between groups.
- The standard error for an individual group mean is estimated by $s/\sqrt{n_i}$, where s^2 = within MS.
- Notice that the nonsmokers, and light smokers have about the same pulmonary function and are worse off than the nonsmokers; and the moderate and heavy smokers have the poorest pulmonary function.



Frequent error in performing the *t* test (compare groups 1 and 2):

- use sample variances to estimate σ^2 from these two groups (vs. from all k groups)
- If former situation: different estimates of σ^2 obtained for each pair of groups considered

correct

not reasonable (all the groups are assumed to have the same underlying variance σ^2)

- Note also that the standard error bars are wider for the non-inhaling smokers than for the other groups because this group has only 50 people compared with 200 for all other groups.
- Are the observer differences in the figure statistically significant as assessed by the LSD procedure in Equation 12.12? The results are presented in Table 12.4.

Table 12.4	Comparisons of using the LSD	of specific pairs of groups for the FEF data in Table 12 t test approach	.1
	Groups compared	Test statistic	p-value
	NS, PS	$t = \frac{3.78 - 3.30}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{200}\right)}} = \frac{0.48}{0.08} = 6.02^{a}$	< .001
	NS, NI	$t = \frac{3.78 - 3.32}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{50}\right)}} = \frac{0.46}{0.126} = 3.65$	< .001
	NS, LS	$t = \frac{3.78 - 3.23}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{200}\right)}} = \frac{0.55}{0.08} = 6.90$	< .001
	NS, MS	$t = \frac{3.78 - 2.73}{0.080} = \frac{1.05}{0.08} = 13.17$	< .001
	NS, HS	$t = \frac{3.78 - 2.59}{0.080} = \frac{1.19}{0.08} = 14.92$	< .001
	PS, NI	$t = \frac{3.30 - 3.32}{0.126} = \frac{-0.02}{0.126} = -0.16$	NS
	PS, LS	$t = \frac{3.30 - 3.23}{0.080} = \frac{0.07}{0.08} = 0.88$	NS
	PS, MS	$t = \frac{3.30 - 2.73}{0.080} = \frac{0.57}{0.08} = 7.15$	< .001
	PS, HS	$t = \frac{3.30 - 2.59}{0.080} = \frac{0.71}{0.08} = 8.90$	< .001
	NI, LS	$t = \frac{3.32 - 3.23}{0.126} = \frac{0.09}{0.126} = 0.71$	NS
	NI, MS	$t = \frac{3.32 - 2.73}{0.126} = \frac{0.59}{0.126} = 4.68$	< .001
	NI, HS	$t = \frac{3.32 - 2.59}{0.126} = \frac{0.73}{0.126} = 5.79$	< .001
	LS, MS	$t = \frac{3.23 - 2.73}{0.08} = \frac{0.50}{0.08} = 6.27$	< .001
	LS, HS	$t = \frac{3.23 - 2.59}{0.08} = \frac{0.64}{0.08} = 8.03$	< .001
	MS, HS	$t = \frac{2.73 - 2.59}{0.08} = \frac{0.14}{0.08} = 1.76$	NS
	*All test statistics follow	v a $t_{\rm 1044}$ distribution under $H_{\rm o}$,	

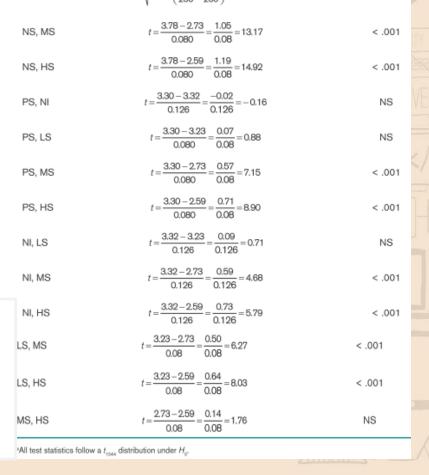


TABLE 5 Percentage points of the t distribution $(t_{du})^*$

Degrees of		и											
freedom, d	.75	.80	.85	.90	.95	.975	.99	.995	.9995				
1	1.000	1.376	1.963	3.078	6.314	12.706	31.821	63.657	636.619				
2	0.816	1.061	1.386	1.886	2.920	4.303	6.965	9.925	31.598				
3	0.765	0.978	1.250	1.638	2.353	3.182	4.541	5.841	12.924				
4	0.741	0.941	1.190	1.533	2.132	2.776	3.747	4.604	8.610				
5	0.727	0.920	1.156	1.476	2.015	2.571	3.365	4.032	6.869				
6	0.718	0.906	1.134	1.440	1.943	2.447	3.143	3.707	5.959				
7	0.711	0.896	1.119	1.415	1.895	2.365	2.998	3.499	5.408				
8	0.706	0.889	1.108	1.397	1.860	2.306	2.896	3.355	5.041				
9	0.703	0.883	1.100	1.383	1.833	2.262	2.821	3.250	4.781				
10	0.700	0.879	1.093	1.372	1.812	2.228	2.764	3.169	4.587				
11	0.697	0.876	1.088	1.363	1.796	2.201	2.718	3.106	4.437				
12	0.695	0.873	1.083	1.356	1.782	2.179	2.681	3.055	4.318				
13	0.694	0.870	1.079	1.350	1.771	2.160	2.650	3.012	4.221				
14	0.692	0.868	1.076	1.345	1.761	2.145	2.624	2.977	4.140				
15	0.691	0.866	1.074	1.341	1.753	2.131	2.602	2.947	4.073				
16	0.690	0.865	1.071	1.337	1.746	2.120	2.583	2.921	4.015				
17	0.689	0.863	1.069	1.333	1.740	2.110	2.567	2.898	3.965				
18	0.688	0.862	1.067	1.330	1.734	2.101	2.552	2.878	3.922				
19	0.688	0.861	1.066	1.328	1.729	2.093	2.539	2.861	3.883				
20	0.687	0.860	1.064	1.325	1.725	2.086	2.528	2.845	3.850				
21	0.686	0.859	1.063	1.323	1.721	2.080	2.518	2.831	3.819				
22	0.686	0.858	1.061	1.321	1.717	2.074	2.508	2.819	3.792				
23	0.685	0.858	1.060	1.319	1.714	2.069	2.500	2.807	3.767				
24	0.685	0.857	1.059	1.318	1.711	2.064	2.492	2.797	3.745				
25	0.684	0.856	1.058	1.316	1.708	2.060	2.485	2.787	3.725				
26	0.684	0.856	1.058	1.315	1.706	2.056	2.479	2.779	3.707				
27	0.684	0.855	1.057	1.314	1.703	2.052	2.473	2.771	3.690				
28	0.683	0.855	1.056	1.313	1.701	2.048	2.467	2.763	3.674				
29	0.683	0.854	1.055	1.311	1.699	2.045	2.462	2.756	3.659				
30	0.683	0.854	1.055	1.310	1.697	2.042	2.457	2.750	3.646				
40	0.681	0.851	1.050	1.303	1.684	2.021	2.423	2.704	3.551				
60	0.679	0.848	1.046	1.296	1.671	2.000	2.390	2.660	3.460				
120	0.677	0.845	1.041	1.289	1.658	1.980	2.358	2.617	3.373				
00	0.674	0.842	1.036	1.282	1.645	1.960	2.326	2.576	3.291				

Solution:

- There are very highly significant differences:
 - (1) between the nonsmokers and all other groups
 - (2) between the passive smokers and the moderate and heavy smokers
 - (3) between the non-inhaling and the moderate and heavy smokers
 - (4) between the light smokers and the moderate and heavy smokers.
- There are no significant differences between the passive smokers, non-inhalers, and light smokers and no significant differences between the moderate and heavy smokers
 - there is a trend toward significance with the latter comparison.

Multiple Comparisons

- Comparisons are made before looking at the data → t test procedure is appropriate
- Comparisons are made after looking at the data
 - → large no. of comparisons
- → some significant differences may be found by chance (false significant difference)

Multiple Comparisons—Bonferroni Approach

- Avoid too many falsely significant difference
- Overall probability of declaring any significant differences between all possible pairs of groups is maintained at some fixed significance level
- Simplest : Bonferroni adjustment

Suppose we want to compare two specific groups (group 1 and group 2) among k groups.

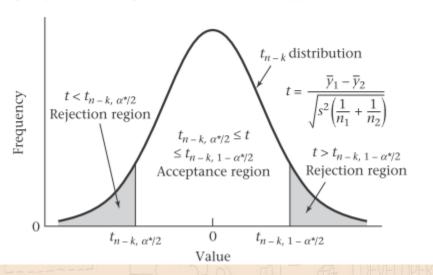
$$H_0$$
: $\alpha_1 = \alpha_2$ vs. H_1 : $\alpha_1 \neq \alpha_2$

Bonferroni multiple-comparisons procedure:

- 1. Compute pooled estimate of the variance s² = Within MS from the one-way ANOVA.
- 2. Compute test statistic: $t = \frac{\overline{y}_1 \overline{y}_2}{\sqrt{s^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$
- 3. Two-sided level α test, let $\alpha^* = \alpha/\binom{k}{2}$
 - reject H_0 : if $t > t_{n-k, 1-\alpha^*/2}$ or $t < t_{n-k, \alpha^*/2}$
 - accept H_0 : if $t > t_{n-k,\alpha^*/2} \le t \le t_{n-k,1-\alpha^*/2}$

no. of test =
$$\binom{k}{2} = \frac{k!}{2!(k-2)!}$$

Figure 12.7 Acceptance and rejection regions for the comparison of pairs of groups in one-way ANOVA (Bonferroni approach)



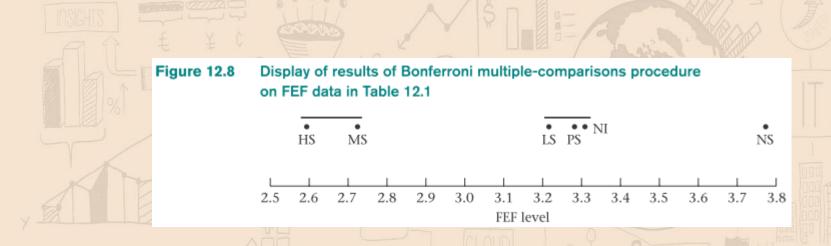
k groups:

- (^k₂) possible two-group comparisons
- each two-group comparison at the α^* level of significance.
- Let E be the event that at least one of the two-group comparisons is statistically significant

 $Pr(E) = Pr(none of the two-group comparisons is statistically significant) = 1- <math>\alpha$

• each of the two-group comparisons were independent from the multiplication law of probability: $Pr(E) = (1 - \alpha^*)^c$ where $c = \binom{k}{2}$

$$1 - \alpha = (1 - \alpha^*)^c$$



Line drawn between the names or numbers of each pair of means that is not significantly different

- Visually summarize the results of many comparisons of pairs of means
- Multiple-comparisons procedures are more strict than ordinary t tests (compare more than two means)
- As k increases, $c = \binom{k}{2}$ increases and therefore $\alpha^* = \alpha/c$ decreases
- The critical value, $t_{n-k,1-\alpha^*/2}$, therefore increases
 - \circ As k increases, the df (n-k) decreases and the percentile 1- $\alpha^*/2$ increases
 - Both lead to larger critical value

When is multiple-comparisons procedure used over LSD procedure?

- Multiple-comparisons procedures should be used if there are many groups and not all comparisons between individual groups are planned
- Relatively few groups and only specific comparisons of interest are intended, then use ordinary t tests (i.e., the LSD procedure)
- Multiple-comparisons procedure is applicable for comparing pairs of means

Example on Bonferroni Correction – Pulmonary Disease

 Question: Apply the Bonferroni multiple-comparisons procedure to the FEF data in Table 12.1

Solution:

- Experiment-wise type I error = .05
- n = 1050 subjects and k = 6 groups

$$\rightarrow n - k = 1044$$
 and $c = \binom{6}{2} = 15$

 $\alpha^* = .05/15 = .0033$ level of significance

- critical value for each of these t tests is $t_{1044,1-.0033/2} = t_{1044,.99833}$. We will approximate a t distribution with 1044 df by an N (0,1)distribution or, $t_{1044,.99833} \approx z_{.99833}$
- $z_{.99833} = 2.93$
- Table 12.4 which provides the t statistics for each two-group comparison

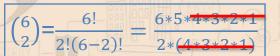




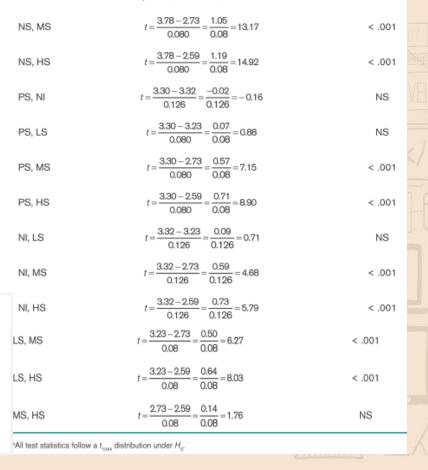
TABLE 3 The normal distribution (continued

X	A=	Bb	C=	D ^d	X	Α	В	С	D
1.82	.9656	.0344	.4656	.9312	2.39	.9916	.0084	.4916	.9832
1.83	.9664	.0336	.4664	.9327	2.40	.9918	.0082	.4918	.9836
1.84	.9671	.0329	.4671	.9342	2.41	.9920	.0080	.4920	.9840
1.85	.9678	.0322	.4678	.9357	2.42	.9922	.0078	.4922	.9845
1.86	.9686	.0314	4686	.9371	2.43	.9925	.0075	4925	.9849
1.87	.9693	.0307	.4693	.9385	2.44	.9927	.0073	.4927	.9853
1.88	.9699	.0301	.4699	.9399	2.45	.9929	.0071	.4929	.9857
1.89	.9706	.0294	.4708	.9412	2.46	.9931	.0069	.4931	.9861
	.9713	.0287	.4713	.9426	2.47	.9932	.0068	.4932	.9865
1.90 1.91				.9426	2.47	.9932	.0066	.4932	.9869
	.9719	.0281	.4719						
1.92	.9726	.0274	.4726	.9451	2.49	.9936	.0064	.4936	.9872
1.93	.9732	.0268	.4732	.9464	2.50	.9938	.0062	.4938	.9876
1.94	.9738	.0262	.4738	.9476	2.51	.9940	.0060	.4940	.9879
1.95	.9744	.0256	.4744	.9488	2.52	.9941	.0059	.4941	.9883
1.96	.9750	.0250	.4750	.9500	2.53	.9943	.0057	.4943	.9886
1.97	.9756	.0244	.4756	.9512	2.54	.9945	.0055	.4945	.9889
1.98	.9761	.0239	.4761	.9523	2.55	.9946	.0054	.4946	.9892
1.99	.9767	.0233	.4767	.9534	2.56	.9948	.0052	.4948	.9895
2.00	.9772	.0228	.4772	.9545	2.57	.9949	.0051	.4949	.9898
2.01	.9778	.0222	.4778	.9556	2.58	.9951	.0049	.4951	.9901
2.02	.9783	.0217	4783	.9566	2.59	.9952	.0048	.4952	.9904
2.02	.9788	.0212	.4788	.9576	2.60	.9953	.0047	.4953	.9907
									.9909
2.04	.9793	.0207	.4793	.9586	2.61	.9955	.0045	.4955	
2.05	.9798	.0202	.4798	.9596	2.62	.9956	.0044	.4956	.9912
2.06	.9803	.0197	.4803	.9606	2.63	.9957	.0043	.4957	.991
2.07	.9808	.0192	.4808	.9615	2.64	.9959	.0041	.4959	.9917
2.08	.9812	.0188	.4812	.9625	2.65	.9960	.0040	.4960	.9920
2.09	.9817	.0183	.4817	.9634	2.66	.9961	.0039	.4961	.9922
2.10	.9821	.0179	.4821	.9643	2.67	.9962	.0038	.4962	.9924
2.11	.9826	.0174	.4826	.9651	2.68	.9963	.0037	.4963	.9926
2.12	.9830	.0170	.4830	.9660	2.69	.9964	.0036	.4964	.9929
2.13	.9834	.0166	.4834	.9668	2.70	.9965	.0035	.4965	.9931
2.14	.9838	.0162	.4838	.9676	2.71	.9966	.0034	.4966	.9933
2.15	.9842	.0158	.4842	.9684	2.72	.9967	.0033	.4967	.9935
2.16	.9846	.0154	.4846	.9692	2.73	.9968	.0032	.4968	.9937
2.17	.9850	.0150	.4850	.9700	2.74	.9969	.0031	.4969	.9939
2.18	.9854	.0146	.4854	.9707	2.75	.9970	.0030	.4970	.9940
2.10	.9857	.0148	.4857	.9715	2.76	.9971	.0029	.4970	.9940
2.20	.9861	.0139	.4861	.9722	2.77	.9972	.0028	.4972	.9944
2.21	.9864	.0136	.4864	.9729	2.78	.9973	.0027	.4973	.9946
2.22	.9868	.0132	.4868	.9736	2.79	.9974	.0026	.4974	.9947
2.23	.9871	.0129	.4871	.9743	2.80	.9974	.0026	.4974	.9949
2.24	.9875	.0125	.4875	.9749	2.81	.9975	.0025	.4975	.9950
2.25	.9878	.0122	.4878	.9756	2.82	.9976	.0024	.4976	.9952
2.26	.9881	.0119	.4881	.9762	2.83	.9977	.0023	.4977	.9953
2.27	.9884	.0116	.4884	.9768	2.84	.9977	.0023	.4977	.9955
2.28	.9887	.0113	.4887	.9774	2.85	.9978	.0022	.4978	.9956
2.29	.9890	.0110	.4890	.9780	2.86	.9979	.0021	.4979	.9958
2.30	.9893	.0107	.4893	.9786	2.87	.9979	.0021	.4979	.9959
2.31	.9896	.0104	.4896	.9791	2.88	.9980	.0020	.4980	.9960
2.32	.9898	.0102	.4898	.9797	2.89	.9981	.0020	.4981	.9961
2.32	.9901	.0102	.4901	.9802	2.89	.9981	.0019	.4981	.9963
2.34	.9904	.0096	.4904	.9807	2.91	.9982	.0018	.4982	.9964
2.35	.9906	.0094	.4906	.9812			.0018	.4982	.9965
2.36	.9909	.0091	.4909	.9817	2.93	.9983	.0017	.4983	.9966
2.37	.9911	.0089	.4911	.9822	2.94	.9984	.0016	.4984	.9967
2.38	.9913	.0087	.4913	.9827	2.95	.9984	.0016	.4984	.9968



(continued on next page)

Table 12.4	Comparisons of using the LSD	of specific pairs of groups for the FEF data in Table 12 t test approach	.1
	Groups compared	Test statistic	p-value
	NS, PS	$t = \frac{3.78 - 3.30}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{200}\right)}} = \frac{0.48}{0.08} = 6.02^{a}$	< .001
	NS, NI	$t = \frac{3.78 - 3.32}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{50}\right)}} = \frac{0.46}{0.126} = 3.65$	< .001
	NS, LS	$t = \frac{3.78 - 3.23}{\sqrt{0.636 \left(\frac{1}{200} + \frac{1}{200}\right)}} = \frac{0.55}{0.08} = 6.90$	< .001
	NS, MS	$t = \frac{3.78 - 2.73}{0.080} = \frac{1.05}{0.08} = 13.17$	< .001
	NS, HS	$t = \frac{3.78 - 2.59}{0.080} = \frac{1.19}{0.08} = 14.92$	< .001
	PS, NI	$t = \frac{3.30 - 3.32}{0.126} = \frac{-0.02}{0.126} = -0.16$	NS
	PS, LS	$t = \frac{3.30 - 3.23}{0.080} = \frac{0.07}{0.08} = 0.88$	NS
	PS, MS	$t = \frac{3.30 - 2.73}{0.080} = \frac{0.57}{0.08} = 7.15$	< .001
	PS, HS	$t = \frac{3.30 - 2.59}{0.080} = \frac{0.71}{0.08} = 8.90$	< .001
	NI, LS	$t = \frac{3.32 - 3.23}{0.126} = \frac{0.09}{0.126} = 0.71$	NS
	NI, MS	$t = \frac{3.32 - 2.73}{0.126} = \frac{0.59}{0.126} = 4.68$	< .001
	NI, HS	$t = \frac{3.32 - 2.59}{0.126} = \frac{0.73}{0.126} = 5.79$	< .001
	LS, MS	$t = \frac{3.23 - 2.73}{0.08} = \frac{0.50}{0.08} = 6.27$	< .001
	LS, HS	$t = \frac{3.23 - 2.59}{0.08} = \frac{0.64}{0.08} = 8.03$	< .001
	MS, HS	$t = \frac{2.73 - 2.59}{0.08} = \frac{0.14}{0.08} = 1.76$	NS
	*All test statistics follow	v a $t_{\rm 1044}$ distribution under $H_{\rm o}$,	



Example on Bonferroni Correction – Pulmonary Disease

Solution:

- Absolute value of all t statistics for two-group comparisons that were statistically significant using the LSD approach are ≥ 3.65
 (3.65 ≥ 2.935 → remain statistically significant under the Bonferroni procedure)
- Comparisons that were not statistically significant with the LSD procedure are also not significant under the Bonferroni procedure
 *must be the case because the Bonferroni procedure is more conservative than the LSD procedure.
 - *critical region using the LSD procedure with a two-sided test (α = .05) is t < -1.96 or t > 1.96, whereas the comparable critical region using the Bonferroni procedure is t < -2.935 or t > 2.935
- The Bonferroni-corrected p-value for comparison of the NS vs. the PS group = 6(5) Pr[N(0,I) > 6.02]

=no. of test * 2 *Pr[N(0,I) > 6.02]
=
$$\frac{6*5}{2}$$
 * 2 *Pr[N(0,I) > 6.0

The False Discovery Rate

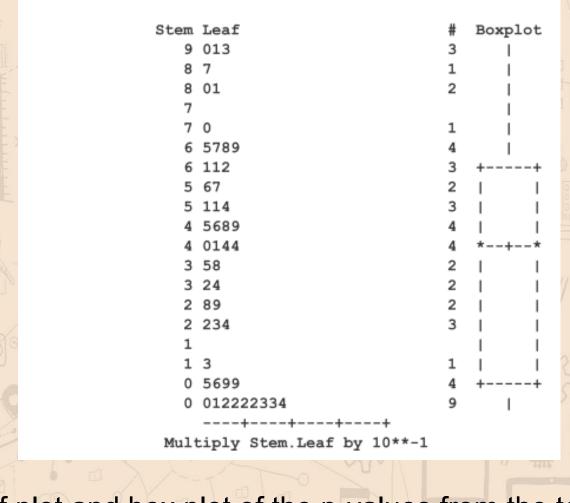
- Genetic studies with many hypotheses: control of the experiment-wise type I error does not seem a reasonable approach to control for multiple comparisons (very conservative inferential procedures)
- FDR: developed by Benjamini and Hochberg
 - Goal: control the proportion of false-positive results among reported statistically significant results

False-Discovery-Rate (FDR) Testing Procedure

- 1. k separate tests with p-values = $p_1, ..., p_k$
- 2. Renumber the tests : $p_1 \le p_2 \le ... \le p_k$
- 3. Define $q_i = kp_i/i$, i = 1,...,k
 - i = rank of the p-values among the k tests
- 4. Let FDRi = false-discovery rate for the ith test be defined by min $(q_i, ..., q_k)$
- 5. Find the largest i such that $FDR_i < FDR_0 = critical level for the FDR (usually 0.05).$
- 6. Reject H_0 for the hypotheses 1, ..., i, and accept H_0 for the remaining hypotheses.
- *no more than 5% of the reported positive results will be false positives*
- *less conservative*

Example on FDR – Cardiovascular Disease, Genetics

- A subsample of 520 cases of cardiovascular disease (CVD) and 1100 controls was obtained among men in a prospective cohort study *nested case-control study*
- Baseline blood samples were obtained from men in the subsample and analyzed for 50 candidate single-nucleotide polymorphisms (SNPs).
- Each SNP was coded as 0 if homozygous wild type (the most common), 1 if heterozygote, and 2 if homozygous mutant. The association of each SNP with CVD was assessed using contingencytable methods.
- A chi-square test for trend was run for each SNP. This yielded 50 separate p-values. If the Bonferroni approach in Equation 12.14 were used, then $\alpha^* = .05/50 = .001$. With such a low value for α^* it is likely that very few of the hypotheses would be rejected, resulting in a great loss in power.
- Instead, an alternative approach based on the false-discovery rate (FDR) was used to control for the problem of multiple testing.



p-Values from tests of 50 SNPs

Figure 12.10

Stem-and-leaf plot and box plot of the p-values from the tests of 50 SNPs.

Example on FDR – Cardiovascular Disease, Genetics

Question: Apply the FDR approach to the genetics data

Solution:

- Stem-and-leaf plot and box plot of the p-values from the tests of each of the 50 SNPs
- The p-values for the nominally significant genes:

Table 12.5 Ordered *p*-values for 10 most significant SNPs

	SNP	p-Value
1	gene30	<.0001
2	gene20	.011
3	gene48	.017
4	gene50	.017
5	gene4	.018
6	gene40	.019
7	gene7	.026
8	gene14	.034
9	gene26	.042
0	gene47	.048

Note that 10 of the genes are statistically significant, with nominal p-values ranging from <.0001 to .048

Example on FDR – Cardiovascular Disease, Genetics

Solution:

- The "Bonferroni p-value" = min {50 x nominal p-value, 1.0} (third column)
 - → the level of significance at which the results for a specific SNP would be just statistically significant if a Bonferroni correction were made
- q_i are not necessarily in the same order as the original nominal pvalues

Table 12.6 Use of the FDR approach to analyzing the CVD data

	SNP	Naïve p-value	Bonferroni <i>p</i> -value	q_{i}	FDR,
1	gene30	<.0001	.0035	.0035	.0035
2	gene20	.011	.54	.28	.16
3	gene48	.017	.86	.28	.16
4	gene50	.017	.87	.22	.16
5	gene4	.018	.92	.18	.16
6	gene40	.019	.94	.16	.16
7	gene7	.026	1.00	.18	.18
8	gene14	.034	1.00	.21	.21
9	gene26	.042	1.00	.23	.23
10	gene47	.048	1.00	.24	.24

Summary

- One-way ANOVA methods: relate a normally distributed outcome variable to the levels of a single categorical independent variable
 -Fixed-effects model: the levels of categorical variable are determined in advance
 - → test the hypothesis that the mean level of the dependent variable is different for different groups defined by the categorical variable
- Methods to adjust for multiple comparisons: Bonferroni correction and false-discovery rate