Testing Hypotheses about Multiple Population Parameters

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Today's plan

- 1. Review reading topics
 - 1.1 Hypothesis testing of multiple population parameters
 - Testing equivalence of two β_j 's
 - Testing multiple exclusion restrictions
 - The *F*-test
- 2. In-class activity: Practice conducting hypothesis tests of multiple population parameters

Test equivalence of two β_i 's

Testing single linear restrictions

- So far, we have tested hypotheses that involve only one parameter, eta_j
- But some hypotheses involve many parameters
- Example: Are returns to junior college same as university?

$$\log(wage) = \beta_0 + \beta_1 jc + \beta_2 univ + \beta_3 exper + u$$

$$H_0 : \beta_1 = \beta_2$$

$$H_a : \beta_1 \neq \beta_2$$

Testing single linear restrictions

- Recall the familiar t-statistic formula:

$$t = \frac{\text{estimate} - \text{null}}{\text{std. err.}}$$

- Can re-write the null hypothesis to be $H_0: \beta_1 \beta_2 = 0$
- Plugging in:

$$t = \frac{(\hat{\beta}_1 - \hat{\beta}_2) - o}{se(\hat{\beta}_1 - \hat{\beta}_2)}$$

Standard Error of $\hat{\beta}_1 - \hat{\beta}_2$

- OLS output gives us se $(\hat{\beta}_1)$ and se $(\hat{\beta}_2)$
- but that's not enough to get $se(\hat{\beta}_1 \hat{\beta}_2)$
- Why? Properties of variances

$$\begin{aligned} & Var(\hat{\beta}_1 - \hat{\beta}_2) = & Var(\hat{\beta}_1) + Var(\hat{\beta}_2) - 2Cov(\hat{\beta}_1, \hat{\beta}_2), \\ & Se(\hat{\beta}_1 - \hat{\beta}_2) = & \sqrt{Var(\hat{\beta}_1) + Var(\hat{\beta}_2) - 2Cov(\hat{\beta}_1, \hat{\beta}_2)} \end{aligned}$$

- Need to know $Cov(\hat{\beta}_1, \hat{\beta}_2)$ to complete the formula
- This number isn't readily reported by most regression packages

An easy way and a hard way

- **Easy way:** use the linearHypothesis() function in the car package
 - Syntax: linearHypothesis(est, "jc = univ")
 - est is the output of 'lm()'
- Hard way: re-run slightly different regression

$$\log(wage) = \beta_0 + \theta_1 jc + \beta_2 (jc + univ) + \beta_3 exper + u$$

- se
$$(\hat{ heta}_1)$$
 = se $(\hat{eta}_1 - \hat{eta}_2)$

Testing multiple exclusion restrictions

Testing multiple exclusion restrictions

- t test is for a **single hypothesis**, whether it involves 1 or 2+ parameters
- But often want to test more than one hypothesis
- We need a statistic that will allow us to test joint hypotheses
- Enter: the *F*-statistic

Example: MLB salaries

- Suppose we use the following model of salaries:

$$\log(\text{salary}) = \beta_0 + \beta_1 \text{years} + \beta_2 \text{gamesyr} + \beta_3 \text{bavg} + \beta_4 \text{hrunsyr} + \beta_5 \text{rbisyr} + u$$

- H_0 : Once we control for years and gamesyr, rbisyr, etc. have no effect

$$H_0: \beta_3 = 0, \beta_4 = 0, \beta_5 = 0$$

 $H_a: \beta_3 \neq 0 \text{ OR } \beta_4 \neq 0 \text{ OR } \beta_5 \neq 0$

- exclusion restrictions: bavg, hrunsyr, and rbisyr can be excluded
- To test Ho, we need a joint (multiple) hypotheses test

Regression output (using mlb1 data set)

```
tidv(est)
# A tibble: 6 x 5
          estimate std.error statistic
                                          p.value
 term
1 (Intercept)
              11.2
                                   38.8
                                          4.19e-128
                        0.289
2 years
               0.0689
                                    5.68
                                          2.79e-
                        0.0121
                        0.00265
                                                 6
3 gamesyr
               0.0126
                                          3.09e-
                                    4.74
4 bavg
                                    0.887 3.76e-
               0.000979 0.00110
                                                  1
5 rbisvr
               0.0108
                                          1.34e-
                        0.00717
                                    1.50
                                                  1
 hrunsvr
               0.0144
                        0.0161
                                    0.899 3.69e-
                                                  1
```

- Each t-stat for bavg, rbisyr, hrunsyr is insignificant (|t| < 2)
- Should we conclude that none of bavg, hrunsyr, and rbisyr affects salaries?



Multicollinearity and single hypothesis tests

- What's going on? Severe multicollinearity
- Corr (hrunsyr, rbisyr) \approx 0.9
- Theoretically, one can't hit a home run without getting at least +1 RBI
- Coeffs on hrunsyr and rbisyr are imprecisely estimated, so need joint test

Unrestricted vs. restricted model

Original model (with all variables) is the unrestricted model:

$$\log(salary) = \beta_0 + \beta_1 y ears + \beta_2 g ames yr + \beta_3 bavg + \beta_4 hruns yr + \beta_5 rbis yr + u$$

- Imposing $H_0: \beta_3 = 0, \beta_4 = 0, \beta_5 = 0$ gives the **restricted model**:

$$\log(salary) = \beta_0 + \beta_1 y ears + \beta_2 g a mes y r + u$$

- We want to see how the fit deteriorates as we remove the three variables
- Use SSR as the measure of fit (or lack thereof)

Intuition of the hypothesis test

- Let SSR_{ur} denote the SSR from the unrestricted model
- In this example, $df_{ur} = 353 6 = 347$
- Let SSR_r be the SSR from the restricted model; $df_{ur} = 353 3 = 350$
- Also recall that SSR must (weakly) decrease when more variables are added:

$$SSR_r \geq SSR_{ur}$$

- Want to know: Does SSR increase by enough to conclude that H_0 is false?

The F-statistic

- Let our general model be written as

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$$

- What to test if last q variables can be excluded:

$$H_0: \beta_{k-q+1} = 0, ..., \beta_k = 0$$

- Compute SSR_{ur} and SSR_r from each model
- F statistic formula:

$$F = \frac{(SSR_r - SSR_{ur})/(df_r - df_{ur})}{SSR_{ur}/df_{ur}} = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(N - k - 1)}$$

More on the F test

- While the t-stat has one degrees-of-freedom, the F-stat has two:
 - Numerator (q) and Denominator (N-k-1)
- We say "Our test statistic has an F distribution with (q, N k 1) d.f."

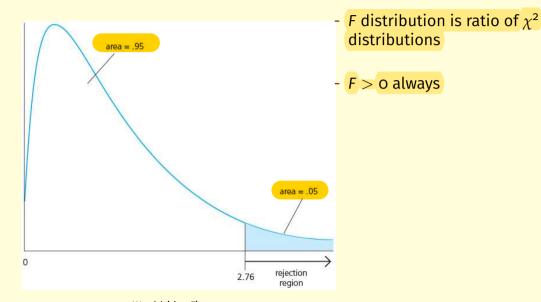
R^2 form of F test

- Can also compute the F-stat from the R^2 of the regression
- Simple algebra shows *F* can also be written as

$$F = \frac{(R_{ur}^2 - R_r^2)/q}{(1 - R_{ur}^2)/(n - k - 1)}$$

- Notice: R_{ur}^2 comes first in the numerator
- We know $R_{ur}^2 \ge R_r^2$ so this ensures $F \ge 0$.

Visualizing the *F* **distribution**



Performing the F test in R

The code to do the F test in R is below:

```
linearHypothesis(est,c("bavg=0", "rbisyr=0", "hrunsyr=0"))
```

Linear hypothesis test

The overall F test

- By default, R also reports an "overall" F test with every regression
- Conducts the following hypothesis test:

$$H_0: \beta_1 = 0, \ldots, \beta_k = 0$$

 $H_a: \text{any slope coefficient } \neq 0$

- From our MLB example:

```
>glance(est)
```

```
A tibble: 1 x 11
r.sq adj.r.sq sigma statistic p.value df logLik AIC BIC
0.628 0.622 0.727 117. 2.94e-72 6 -385. 784. 811.
```

Relationship between F and t tests

- In the F test setting, nothing rules out q = 1
- So, we can use the F statistic to test, say,

$$H_0: \beta_j = 0$$

 $H_a: \beta_i \neq 0$

- Of course, we can use a t statistic, too
- Does that mean there are now two ways to test a single restriction?
- No. It can be shown that, when q= 1, then $\mathit{F}=\mathit{t}^2$