POPULATION FILTER INFERENCE

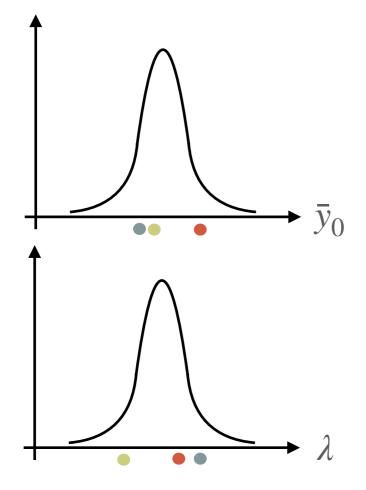
A scalable hierarchical inference framework

EXAMPLE: HIERARCHICAL EXP. GROWTH MODEL

Population model

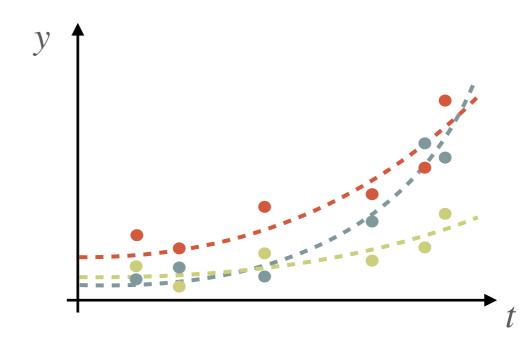
$$\bar{y}_0 \sim \mathcal{N}\left(\mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2\right)$$

$$\lambda \sim \mathcal{N}\left(\mu_{\lambda}, \sigma_{\lambda}^{2}\right)$$



Individual model

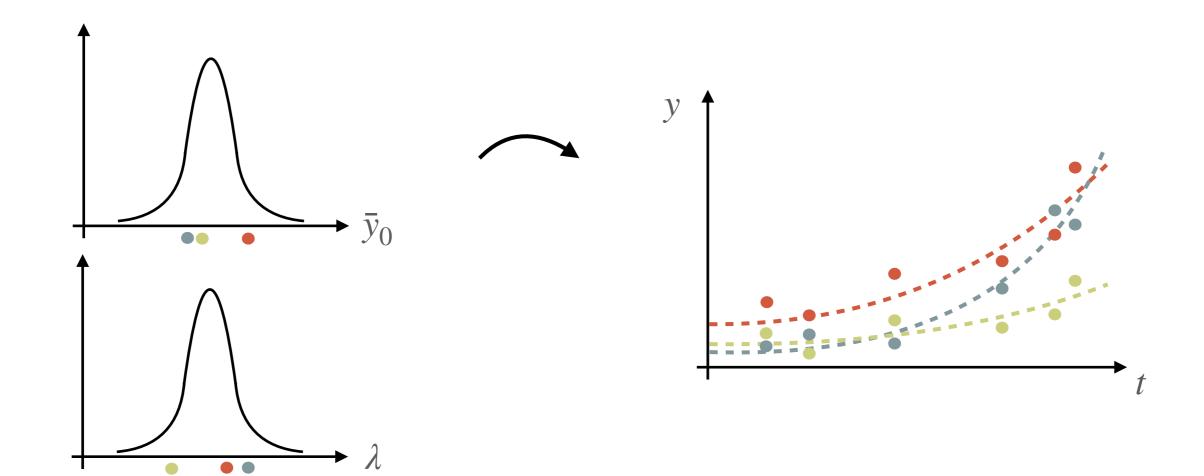
$$y = \bar{y}_0 e^{\lambda t} + \epsilon$$
$$\epsilon \sim \mathcal{N}(0, \sigma^2)$$



EXAMPLE: HIERARCHICAL EXP. GROWTH MODEL

Data-generating model

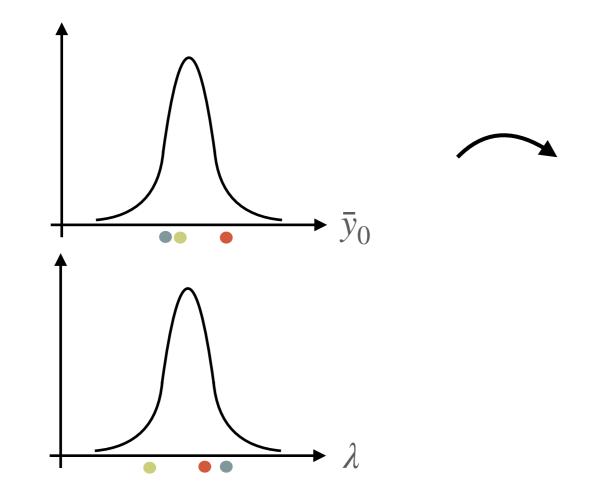
$$p(y | \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}, \mu_{\lambda}, \sigma_{\lambda}, \sigma, t) = \int d\bar{y}_0 d\lambda \, \mathcal{N}(y | \bar{y}_0 e^{\lambda t}, \sigma^2) \, \mathcal{N}(\bar{y}_0 | \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2) \, \mathcal{N}(\lambda | \mu_{\lambda}, \sigma_{\lambda}^2)$$

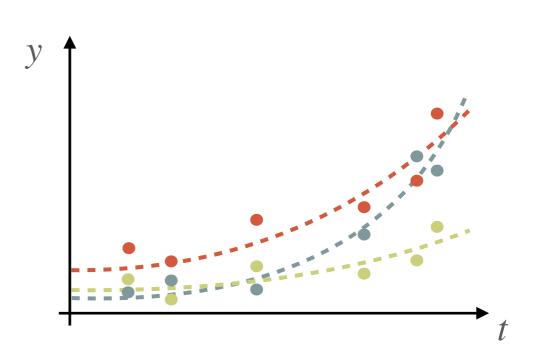


EXAMPLE: HIERARCHICAL EXP. GROWTH MODEL

Data-generating model

$$\underbrace{p(y \mid \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}, \mu_{\lambda}, \sigma_{\lambda}, \sigma, t)}_{p(y \mid \theta, t)} = \underbrace{\int d\bar{y}_0 d\lambda}_{d\psi} \underbrace{\mathcal{N}(y \mid \bar{y}_0 e^{\lambda t}, \sigma^2)}_{p(y \mid \psi, t)} \underbrace{\mathcal{N}(\bar{y}_0 \mid \mu_{\bar{y}_0}, \sigma^2_{\bar{y}_0}) \mathcal{N}(\lambda \mid \mu_{\lambda}, \sigma^2_{\lambda})}_{p(\psi \mid \theta)}$$





HOW CAN WE ESTIMATE POPULATION PARAMETERS FROM DATA?

1. NAÏVE APPROACH: POPULATION PARAMETER POSTERIOR

Log-likelihood & log-posterior

LOG-LIKELIHOOD:

$$\log p(\mathcal{D} \mid \theta) = \sum_{j} \log p(y_j^{\text{obs}} \mid \theta, t_j) \qquad \text{for data} \quad \mathcal{D} = \{(y_j^{\text{obs}}, t_j)\}$$

POSTERIOR:

$$p(\theta \mid \mathcal{D}) \propto p(\mathcal{D} \mid \theta) p(\theta)$$

Recall: Hierarchical exp. growth model

$$p(y \mid \theta, t) = \int d\bar{y}_0 d\lambda \,\mathcal{N}(y \mid \bar{y}_0 e^{\lambda t}, \sigma^2) \,\mathcal{N}(\bar{y}_0 \mid \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2) \,\mathcal{N}(\lambda \mid \mu_{\lambda}, \sigma_{\lambda}^2)$$

1. NAIVE APPROACH: POPULATION PARAMETER POSTERIOR

Log-likelihood & log-posterior

LOG-LIKELIHOOD:

$$\log p(\mathcal{D} \mid \theta) = \sum_{j} \log p(y_j^{\text{obs}} \mid \theta, t_j) \qquad \text{for data} \quad \mathcal{D} = \{(y_j^{\text{obs}}, t_j)\}$$

POSTERIOR:

$$p(\theta \mid \mathcal{D}) \propto p(\mathcal{D} \mid \theta) p(\theta)$$

Recall: Hierarchical exp. growth model

$$p(y \mid \theta, t) = \int d\bar{y}_0 d\lambda \, \mathcal{N}(y \mid \bar{y}_0 e^{\lambda t}, \sigma^2) \, \mathcal{N}(\bar{y}_0 \mid \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2) \, \mathcal{N}(\lambda \mid \mu_{\lambda}, \sigma_{\lambda}^2)$$

VERY HARD TO COMPUTE NUMERICALLY!

2. HIERARCHICAL APPROACH: JOINT POSTERIOR

$$p(y, \psi | \theta, t) = \mathcal{N}(y | \bar{y}_0 e^{\lambda t}, \sigma^2) \mathcal{N}(\bar{y}_0 | \mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2) \mathcal{N}(\lambda | \mu_{\lambda}, \sigma_{\lambda}^2)$$

Log-likelihood & log-posterior

LOG-LIKELIHOOD FOR AN INDIVIDUAL:

$$\log p(\mathcal{D}_i, \psi_i | \theta) = \sum_{i} \log p(y_{ij}^{\text{obs}}, \psi_i | \theta, t_{ij})$$

for data of individual i $\mathcal{D}_i = \{(y_{ij}^{\text{obs}}, t_{ij})\}$

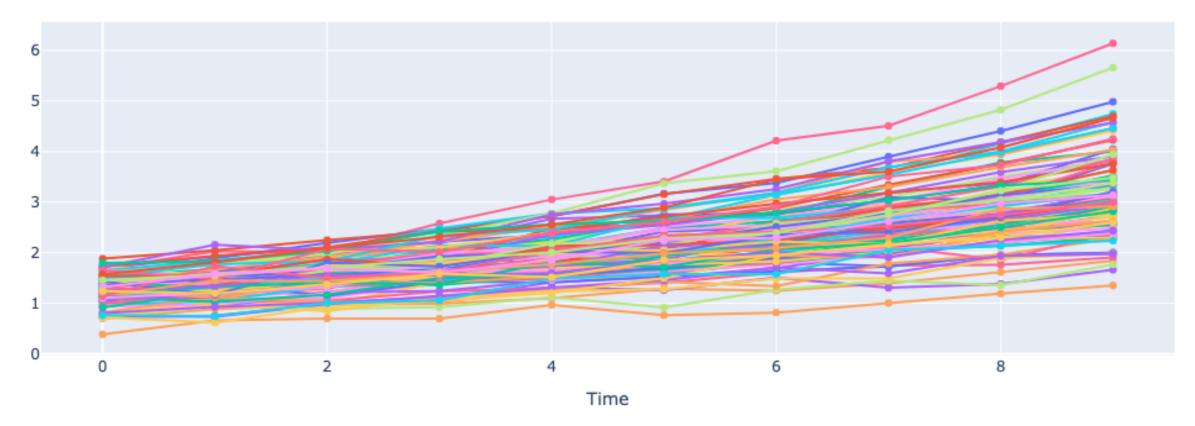
POPULATION LOG-LIKELIHOOD:

POSTERIOR:

$$\log p(\mathcal{D}, \Psi \mid \theta) = \sum_{i} \log p(\mathcal{D}_{i}, \psi_{i} \mid \theta) \qquad p(\theta, \Psi \mid \mathcal{D}) \propto p(\mathcal{D}, \Psi \mid \theta) p(\theta)$$
 with $\Psi = \{\psi_{i}\}$

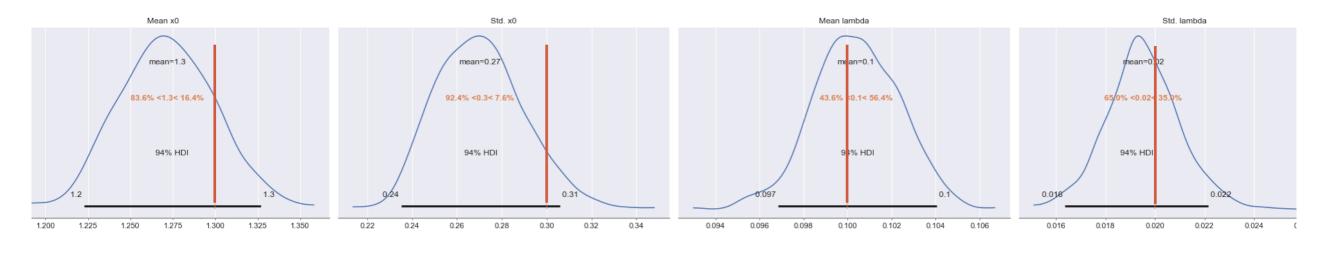
2. HIERARCHICAL APPROACH IN PRACTICE: EXAMPLE 1

DATA: 100 INDIVIDUALS WITH 10 OBSERVATIONS



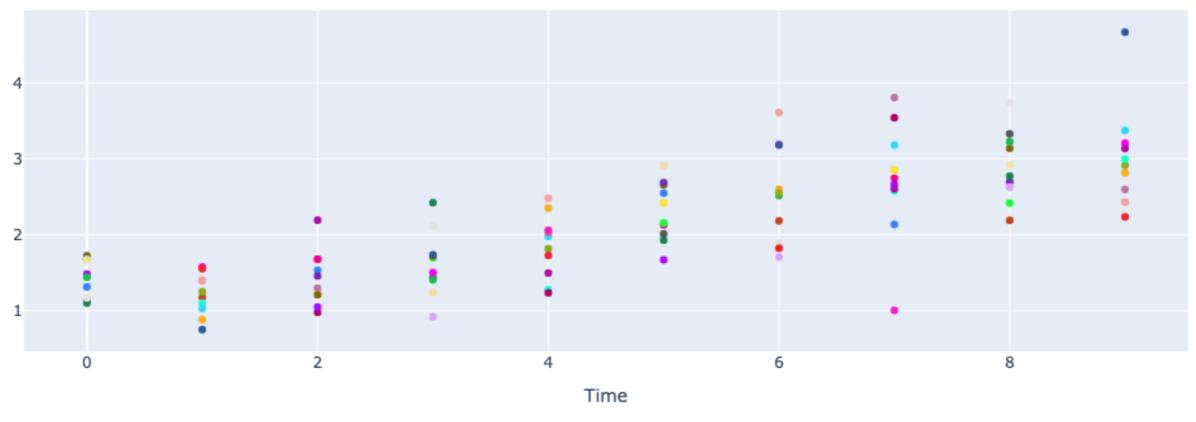
INFERENCE RESULTS:

RUN TIME ~ 13 MINUTES



2. HIERARCHICAL APPROACH IN PRACTICE: EXAMPLE 2

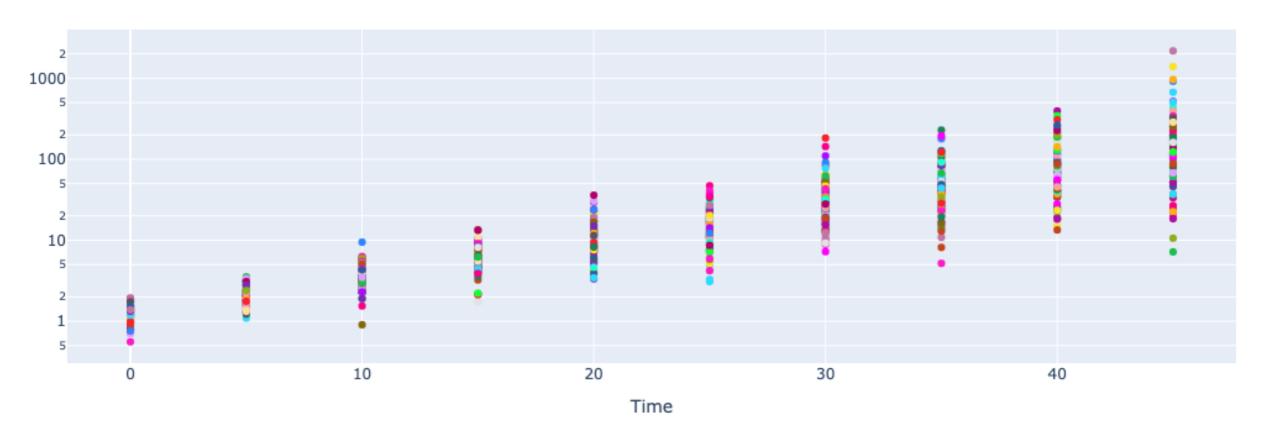
DATA: 100 INDIVIDUALS WITH 1 OBSERVATION EACH - SNAPSHOT DATA





2. HIERARCHICAL APPROACH IN PRACTICE: EXAMPLE 3

DATA: 1000 INDIVIDUALS WITH 1 OBSERVATION EACH - SNAPSHOT DATA



INFERENCE RESULTS: RUN TIME ~ DAYS

- ➤ HIERARCHICAL INFERENCE BECOMES INTRACTABLE.
- ➤ POPULATION VARIATION WILL BE SIGNIFICANTLY UNDERESTIMATED (SHRINKAGE).

WHAT ARE THE CHALLENGES WITH SNAPSHOT DATA?

WHAT ARE THE CHALLENGES WITH SNAPSHOT DATA?

Recall: Hierarchical exp. growth model

SOME INTUITION:

$$p(y,\psi \,|\, \theta,t) = p_n(y \,|\, \bar{y}_0 \mathrm{e}^{\lambda t},\sigma^2) \, p_n(\bar{y}_0 \,|\, \mu_{\bar{y}_0},\sigma^2_{\bar{y}_0}) \, p_n(\lambda \,|\, \mu_{\lambda},\sigma^2_{\lambda})$$

WITHOUT POPULATION MODEL: DATA IS FITTED EXACTLY.

$$\bar{y}_{0,i}e^{\lambda_i t} = y_i^{\text{obs}}$$
 and $\mathcal{N}(y_i^{\text{obs}} | \bar{y}_{0,i}e^{\lambda_i t}, \sigma^2) = 1$ for $\sigma \to 0$

DEFINES A FAMILY OF PARAMETERS FOR EACH INDIVIDUAL $\psi_i = (\bar{y}_{0,i}(\lambda_i), \lambda_i)$.

2. WITH POPULATION MODEL: VARIATION IN Ψ IS PENALISED.

$$\underbrace{\prod_{i=1}^{N} \mathcal{N}(y_{i}^{\text{obs}} | \bar{y}_{0,i} e^{\lambda_{i}t}, \sigma^{2})}_{N \, \text{factors}} \quad \mathbf{COMPETES \, AGAINST} \quad \underbrace{\prod_{i=1}^{N} \mathcal{N}(\bar{y}_{0,i} | \mu_{\bar{y}_{0}}, \sigma_{\bar{y}_{0}}^{2}) \, \mathcal{N}(\lambda_{i} | \mu_{\lambda}, \sigma_{\lambda}^{2})}_{2N \, \text{factors}}$$

> IT'S BENEFICIAL TO INCREASE NOISE AND REDUCE POPULATION VARIATION.

BUT: IMBALANCE (N VS. 2N) LEADS TO TOO MUCH REDUCTION IN POPULATION VARIATION (I.E. SHRINKAGE).

WHAT ARE THE PROBLEMS WITH SNAPSHOT DATA?

In general

$$\underbrace{\prod_{i=1}^{N} p(y_i^{\text{obs}} | \psi_i)}_{N \text{ factors}}$$

COMPETES AGAINST

$$\underbrace{\prod_{i=1}^{N} \prod_{k=1}^{K} p(\psi_{ik} | \theta_k)}_{KN \text{ factors}}$$

> THE MORE PARAMETERS THE MORE SHRINKAGE OF POPULATION VARIATION.

RESOLUTIONS:

- 1. STRONG PRIOR ON NOISE (OR FIX NOISE).
 - DOES NOT REMOVE SHRINKAGE.
 - + REDUCES MAGNITUDE OF SHRINKAGE.
- 2. IF POSSIBLE, AVOID FITTING TO INDIVIDUALS AND JUST FIT ON POPULATION LEVEL.
 - + REMOVES SHRINKAGE
 - + MAKES INFERENCE TIME INDEPENDENT OF NUMBER OF OBSERVED INDIVIDUALS
 - MAKES NOISE NON-IDENTIFIABLE
 - NEEDS APPROXIMATION OF POPULATION DISTRIBUTION

MOTIVATION 1: WANT TO FIT POPULATION DISTRIBUTION DIRECTLY TO DATA.

PROBLEM: CANNOT COMPUTE POPULATION DISTRIBUTION.

$$p(y \mid \theta, t) = \int d\psi \, p(y \mid \psi, t) \, p(\psi \mid \theta) \approx \mathcal{N}(y \mid \mu(\theta, t), \sigma^2(\theta, t))$$

WHERE

$$\mu(\theta, t) = \mathbb{E}[y | \theta, t]$$
 and $\sigma^2(\theta, t) = \text{Var}[y | \theta, t]$

AT FIRST GLANCE THIS LOOKS LIKE WE NOT ONLY HAD TO INTRODUCE AN APPROXIMATION, BUT WE ALSO INTRODUCED TWO FURTHER INTEGRALS THAT WE CANNOT SOLVE.

MOTIVATION 1: WANT TO FIT POPULATION DISTRIBUTION DIRECTLY TO DATA.

PROBLEM: CANNOT COMPUTE POPULATION DISTRIBUTION.

$$p(y \mid \theta, t) = \int d\psi \, p(y \mid \psi, t) \, p(\psi \mid \theta) \approx \mathcal{N}(y \mid \mu(\theta, t), \sigma^2(\theta, t))$$

WHERE

$$\mu(\theta, t) = \mathbb{E}[y | \theta, t]$$
 and $\sigma^2(\theta, t) = \text{Var}[y | \theta, t]$

BUT WE CAN USE SAMPLING TO ESTIMATE THEM: $y \sim p(\cdot \mid \theta, t)$

$$\hat{\mu}(\theta, t_j) = \frac{1}{N_s} \sum_{s=1}^{N_s} y_s(\theta, t_j) \quad \text{and} \quad \hat{\sigma}^2(\theta, t_j) = \frac{1}{N_s - 1} \sum_{s=1}^{N_s} \left(y_s(\theta, t_j) - \hat{\mu}(\theta, t_j) \right)^2$$

WE CAN NOW CONSTRUCT AN APPROXIMATE POPULATION POSTERIOR:

FILTER LOG-LIKELIHOOD:

$$\log p_f(\mathcal{D} \mid \theta) = \sum_{j} \log \mathcal{N}(y_j^{\text{obs}} \mid \hat{\mu}(\theta, t_j), \hat{\sigma}^2(\theta, t_j)) \qquad \text{for data} \quad \mathcal{D} = \{(y_j^{\text{obs}}, t_j)\}$$

FILTER POSTERIOR:

$$p_f(\theta \mid \mathcal{D}) \propto p_f(\mathcal{D} \mid \theta) p(\theta)$$

POUT -221.5

-222.5

0 100 200 300 400 500 600 700 800 900

Repeat

THE LOG-LIKELIHOOD IS NOW A STOCHASTIC FUNCTION OF THE POPULATION PARAMETERS.

THIS BREAKS ADAPTIVE AND GRADIENT-BASED MCMC ALGORITHMS (LIKE ACMC, NUTS)!

WE NEED TO CONTROL THE STOCHASTICITY WHILST SAMPLING.

STOCHASTICITY COMES FROM SAMPLING WITH A RANDOM NUMBER GENERATOR FROM

$$y \sim p(\cdot \mid \theta, t)$$

 \longrightarrow FOR THE SAME θ , WE GET DIFFERENT SAMPLES, AND THUS DIFFERENT APPROXIMATIONS

$$\mathcal{N}(y | \hat{\mu}(\theta, t), \hat{\sigma}^2(\theta, t))$$

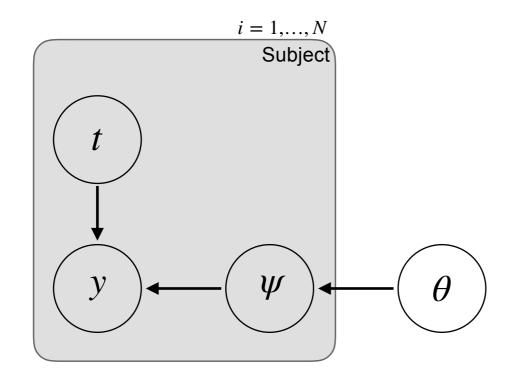
WE CAN OVERCOME THIS BY LETTING THE MCMC SAMPLER TAKE CHARGE OF THE SAMPLING:

$$y_s, \psi_s \sim p(y_s | \psi_s, t) p(\psi_s | \theta)$$
 AND $\mathcal{N}(y | \hat{\mu}(Y), \hat{\sigma}^2(Y))$

$$\log p_f(\mathcal{D}, Y, \Psi \mid \theta) = \sum_j \log \mathcal{N}(y_j^{\text{obs}} \mid \hat{\mu}(Y_j), \hat{\sigma}^2(Y_j)) + \log p(Y_j \mid \Psi, t_j) + \log p(\Psi \mid \theta)$$

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING INDIVIDUALS:

RECALL: HIERARCHAL LOG-LIKELIHOOD FOR SNAPSHOT DATA



$$\log p(\mathcal{D}, \Psi | \theta) = \sum_{i} \log p(y_i^{\text{obs}} | \psi_i, t_i) + \log p(\psi_i | \theta)$$

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING IND

RECALL: HIERARCHAL LOG-LIKELIHOOD FOR SNAPSHOT DATA



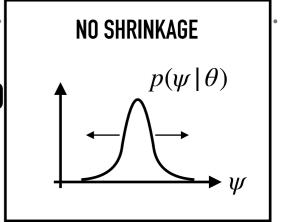
SHRINKAGE

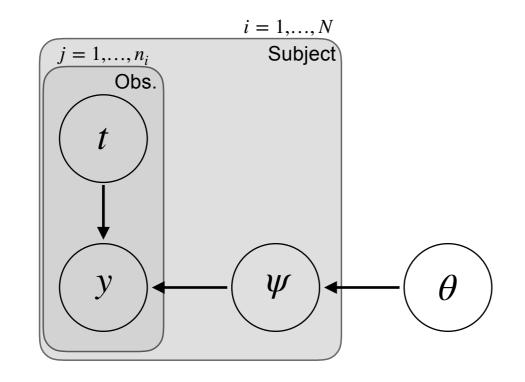
 $p(\psi | \theta)$

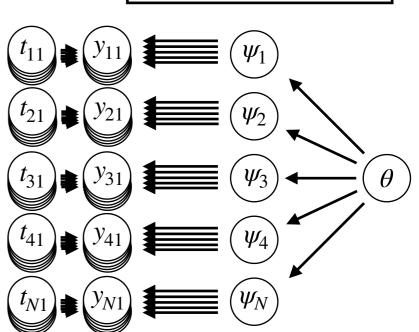
$$\log p(\mathcal{D}, \Psi | \theta) = \sum_{i} \log p(y_i^{\text{obs}} | \psi_i, t_i) + \log p(\psi_i | \theta)$$

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING IND

RECALL: HIERARCHAL LOG-LIKELIHOOD FOR NON-SNAPSHOT DATA





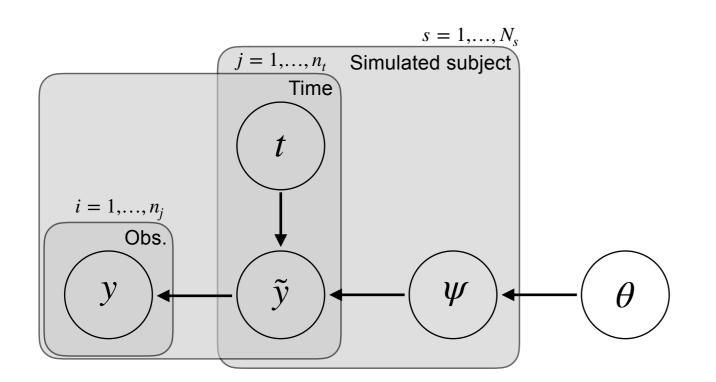


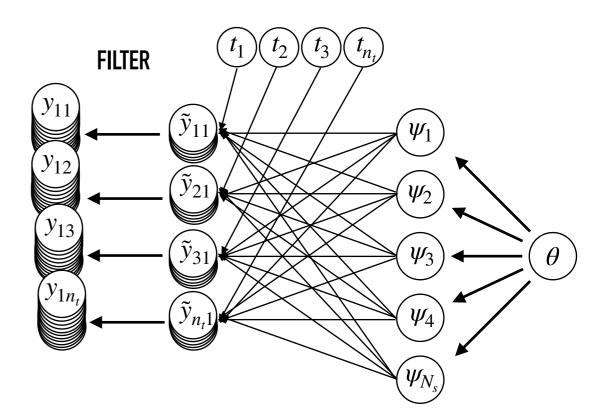
$$\log p(\mathcal{D}, \Psi | \theta) = \sum_{ij} \log p(y_{ij}^{\text{obs}} | \psi_i, t_{ij}) + \log p(\psi_i | \theta)$$

OBSERVATIONS DOMINATE \longrightarrow WEAKLY CORRELATED Ψ \longrightarrow INFERENCE TRACTABLE

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING INDIVIDUALS:

POPULATION FILTER INFERENCE:

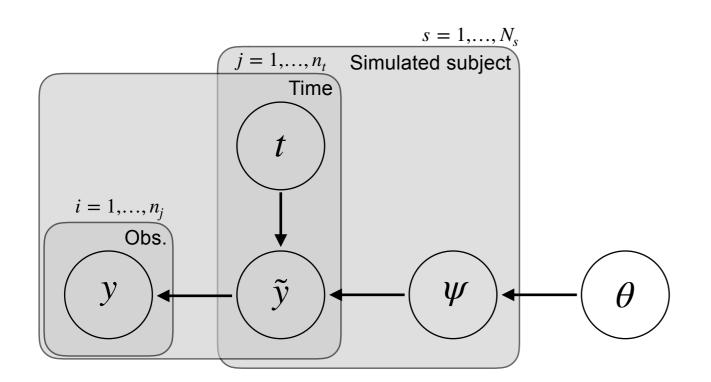


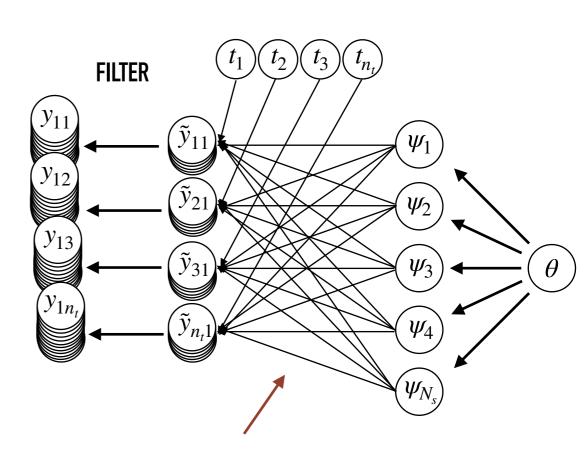


$$\log p(\mathcal{D}, \tilde{Y}, \Psi \mid \theta) = \sum_{ijs} \log p(y_{ij}^{\text{obs}} \mid \tilde{Y}_j) + \log p(\tilde{y}_{js} \mid \psi_s, t_j) + \log p(\psi_s \mid \theta)$$

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING INDIVIDUALS:

POPULATION FILTER INFERENCE:





AVOIDS SHRINKAGE / STRONG CORRELATIONS

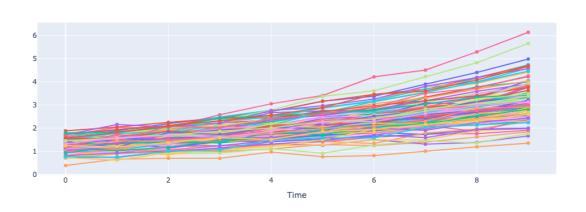
$$\log p(\mathcal{D}, \tilde{Y}, \Psi | \theta) = \sum_{ijs} \log p(y_{ij}^{\text{obs}} | \tilde{Y}_j) + \log p(\tilde{y}_{js} | \psi_s, t_j) + \log p(\psi_s | \theta)$$

POPULATION FILTER

HIERARCHICAL LOG-LIKELIHOOD OF SAMPLE

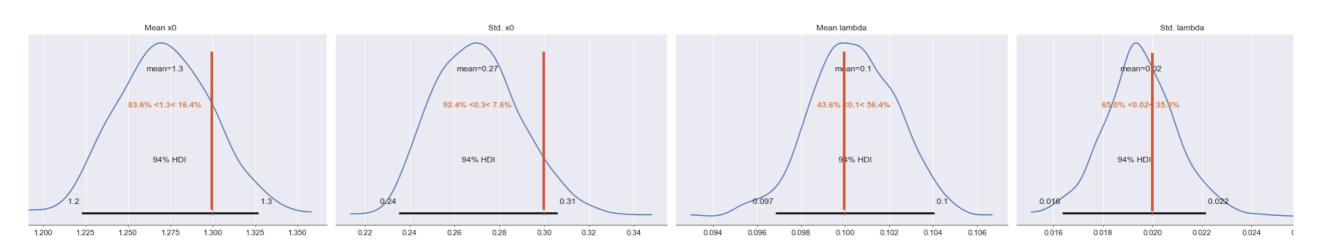
 $\log \mathcal{N}(y_{ij}^{\text{obs}} | \hat{\mu}(\tilde{Y}_j), \hat{\sigma}^2(\tilde{Y}_j)) \qquad \qquad \log p(\tilde{y}_s, \psi_s | \theta, t_j)$

DATA: 100 INDIVIDUALS WITH 10 OBSERVATIONS



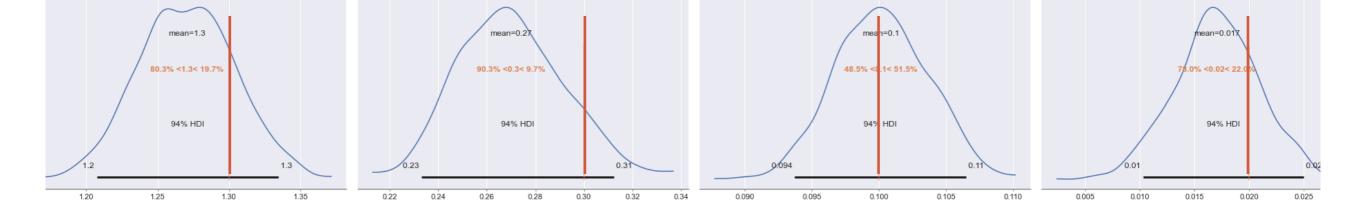
INFERENCE RESULTS FROM HIERARCHICAL INFERENCE (SLIDE 9):

RUN TIME ~ 13 MINUTES

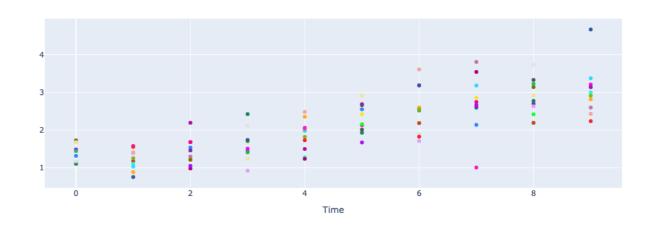


INFERENCE RESULTS FROM POPULATION KDE INFERENCE:

RUN TIME ~ 13 MINUTES

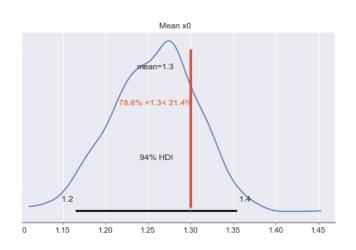


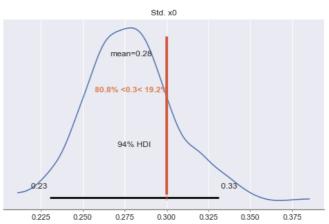
DATA: 100 INDIVIDUALS WITH 1 OBSERVATION EACH - SNAPSHOT DATA

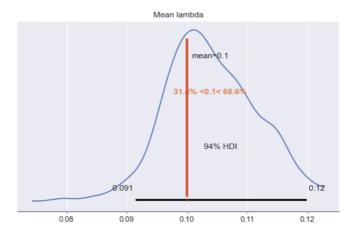


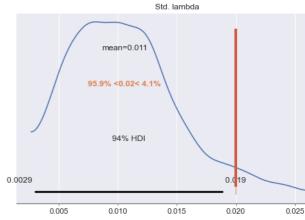
INFERENCE RESULTS FROM HIERARCHICAL INFERENCE (SLIDE 10):

RUN TIME ~ 23 MINUTES



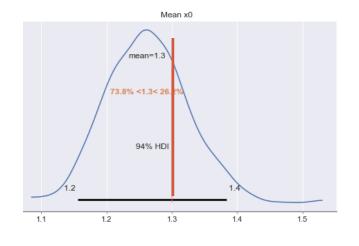


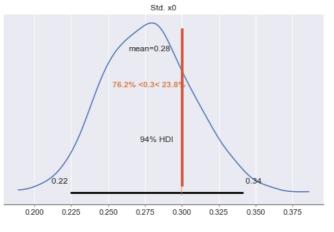


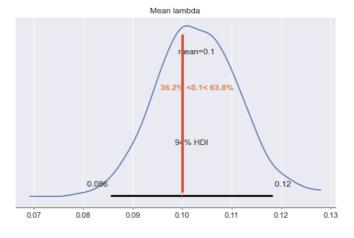


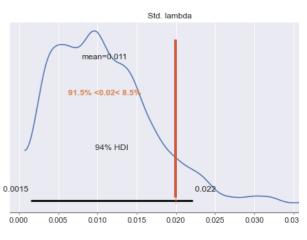
INFERENCE RESULTS FROM POPULATION KDE INFERENCE:

RUN TIME ~ 8 **MINUTES**

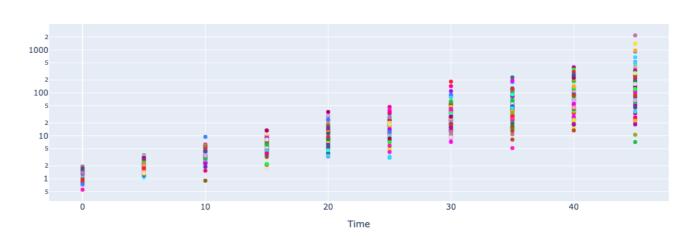








DATA: 1000 INDIVIDUALS WITH 1 OBSERVATION EACH - SNAPSHOT DATA

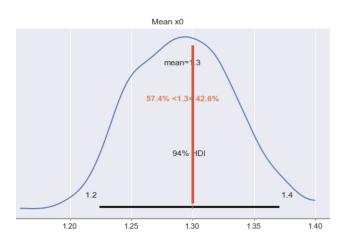


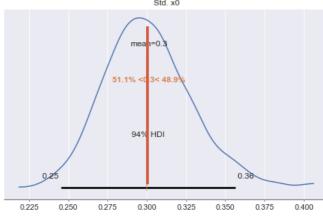
INFERENCE RESULTS FROM HIERARCHICAL INFERENCE (SLIDE 11):

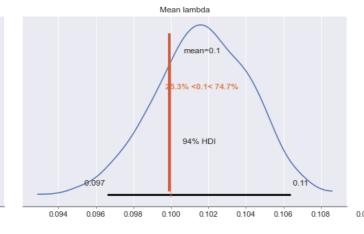
RUN TIME ~ DAYS

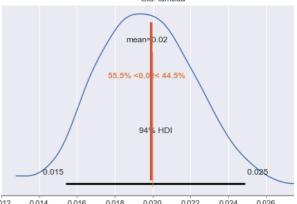
INFERENCE RESULTS FROM POPULATION KDE INFERENCE:

RUN TIME ~ 30 MINUTES





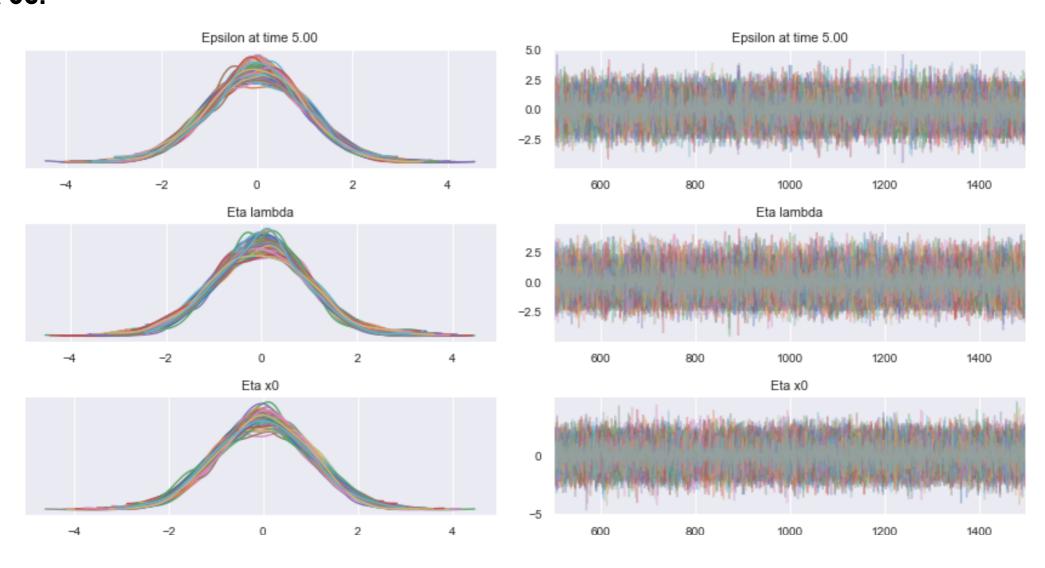




NO SIGNS OF SHRINKAGE AND SIGNIFICANTLY SUPERIOR EXECUTION TIME.

WHY DOES THIS SCALE SO WELL?

A RELATIVELY SMALL NUMBER OF SAMPLES FROM THE POPULATION DISTRIBUTION ARE SUFFICIENT TO APPROXIMATE THE POPULATION DISTRIBUTION WELL (~ 100 SAMPLES), BECAUSE THE MCMC SAMPLER EFFECTIVELY INTEGRATES OVER POPULATION DISTRIBUTION FOR US:



DIMENSIONALITY OF POSTERIOR DOES NOT SCALE WITH NUMBER OF OBSERVED INDIVIDUALS.

OPEN QUESTIONS

- 1. HOW OFTEN IS THE ASSUMPTION OF A GAUSSIAN POPULATION DISTRIBUTION SUFFICIENT FOR GOOD INFERENCE RESULTS?
- 2. HOW EASY CAN WE FIND GOOD KDE APPROXIMATIONS WHEN A GAUSSIAN KDE ISN'T A GOOD CHOICE?
- 3. IS IT STRAIGHTFORWARD TO VALIDATE WHETHER A GIVEN CHOICE OF POPULATION KDE IS SUFFICIENT?
- 4. IF A GOOD KDE CANNOT BE FOUND, HOW WELL DOES A MOMENT BASED APPROACH WORK THAT DOESN'T MAKE ASSUMPTIONS ABOUT THE DISTRIBUTIONAL SHAPE (LIKELY NEEDS AT LEAST 100 OBSERVATIONS PER TIME POINT TO BE ABLE TO GO BEYOND 2. MOMENT).

ADDRESSING Q1:

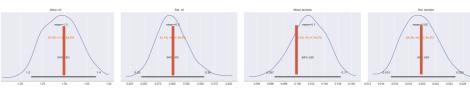
HOW OFTEN IS THE ASSUMPTION OF A GAUSSIAN POPULATION DISTRIBUTION SUFFICIENT FOR GOOD INFERENCE RESULTS?

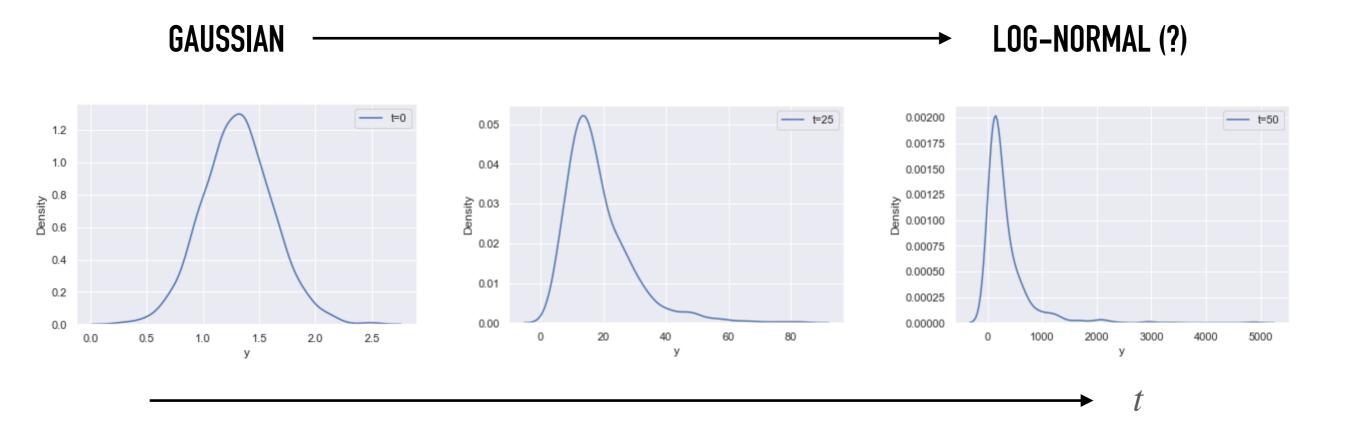
HOW JUSTIFIED IS GAUSSIAN POPULATION KDE?

Hierarchical exp. growth model

TIME SLICES OF DATA-GENERATING MODEL:

RECALL POSTERIORS FROM SLIDE 22:





GAUSSIAN POPULATION KDE APPROXIMATION IS NOT VERY ACCURATE.

BUT: INFERENCE RESULTS ARE NEVERTHELESS GOOD. WHY?

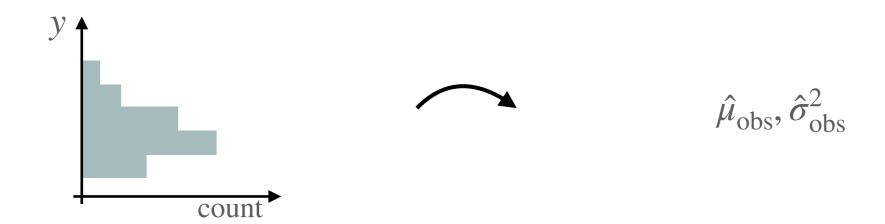
SHORT ANSWER Q1:

THE GAUSSIAN POPULATION KDE IS USED AS A 'FILTER' FOR BOTH THE OBSERVED DATA AND THE SAMPLES FOR A PROPOSED SET OF POPULATION PARAMETERS. THE POPULATION PARAMETERS ARE ACCEPTED RELATIVE TO THE FILTER LIKELIHOOD WITH RESPECT TO THE DATA. SO, SUBJECT TO THE APPROPRIATENESS OF THE FILTER, THE INFERENCE RESULTS WILL BE GOOD.

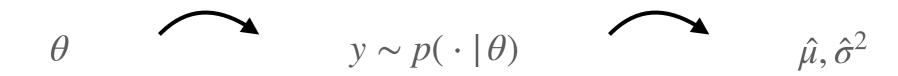
ANSWER Q1: ABC REJECTION SAMPLING

LET'S THINK ABOUT ABC REJECTION SAMPLING:

1. DATA IS SUMMARISED BY SUMMARY STATISTICS, E.G. MEAN AND VARIANCE (FILTERS).



2. PROPOSE PARAMETERS, SAMPLE FROM CANDIDATE MODEL AND COMPUTE SUMMARY STATISTICS.



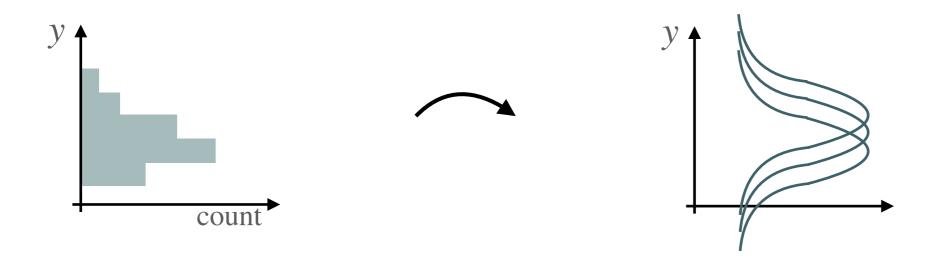
3. ACCEPT IF SUMMARY STATISTICS ARE CLOSE TO EACH OTHER, ELSE REJECT.

$$|\hat{\mu}_{\text{obs}} - \hat{\mu}| < \epsilon$$
 and $|\hat{\sigma}_{\text{obs}}^2 - \hat{\sigma}^2| < \epsilon$

ANSWER Q1: GAUSSIAN KERNEL FILTER

POPULATION FILTER AS 'CONTINUOUS ABC':

DATA IS SUMMARISED BY GAUSSIAN KERNELS WHICH ARE PROPOSED IN 2.:



PROPOSE PARAMETERS, SAMPLE FROM CANDIDATE MODEL AND ESTIMATE GAUSSIAN KERNEL:

$$\theta \qquad \qquad y \sim p(\cdot \mid \theta)$$
 RELATIVE TO KERNEL LIKELIHOOD

ACCEPT RELATIVE TO KERNEL LIKELIHOOD.

$$p_{\text{KDE}}(\theta, \{\epsilon_{ij}\}, \{\psi_i\} \mid \mathcal{D}) \propto p_{\text{KDE}}(\mathcal{D}, \{\epsilon_{ij}\}, \{\psi_i\} \mid \theta) p(\theta)$$

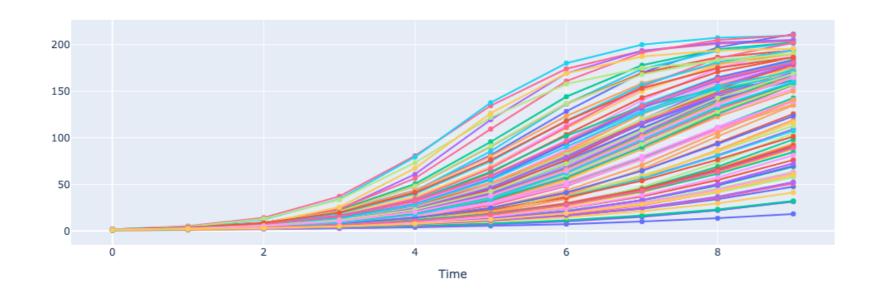
ANSWER Q1

- 1. GAUSSIAN POPULATION FILTER INFERENCE USES GAUSSIAN KERNELS AS FILTERS TO QUANTIFY THE CONSISTENCY OF THE OBSERVED DATA WITH THE PROPOSED SET OF POPULATION PARAMETERS.
 - → INFERRED POSTERIORS ARE CORRECT, BUT MAY BE WIDER THAN NECESSARY SUBJECT TO THE INFORMATION LOSS INCURRED BY FILTERING.
- 2. CAVEATS IN PRACTICE: ONLY SAMPLES CONTRIBUTE TO FILTER CONSTRUCTION WHICH INTRODUCES AN ASYMMETRY.

IF TRUE POPULATION DISTRIBUTION IS DIFFERENT FROM KERNEL:

- i. FEW OBSERVATIONS, FEW SAMPLES: NO PROBLEM.
- ii. FEW OBSERVATIONS, MANY SAMPLES: PARAMETER UNCERTAINTY MAY BE UNDERESTIMATED.
- iii. Many observations, few samples: parameter variance may be overestimated.
- iv. MANY OBSERVATIONS, MANY SAMPLES: NO PROBLEM.
- IF POSSIBLE, RUN WITH AS MANY SAMPLES AS OBSERVATIONS, BUT NEVER MORE SAMPLES THAN OBSERVATIONS IF TRUE POPULATION DISTRIBUTION IS DIFFERENT FROM KERNEL.

HIERARCHICAL LOGISTIC GROWTH MODEL:



$$y = \frac{c \,\overline{y}_0}{\overline{y}_0 + (c - \overline{y}_0)e^{-\lambda t}} + \epsilon$$

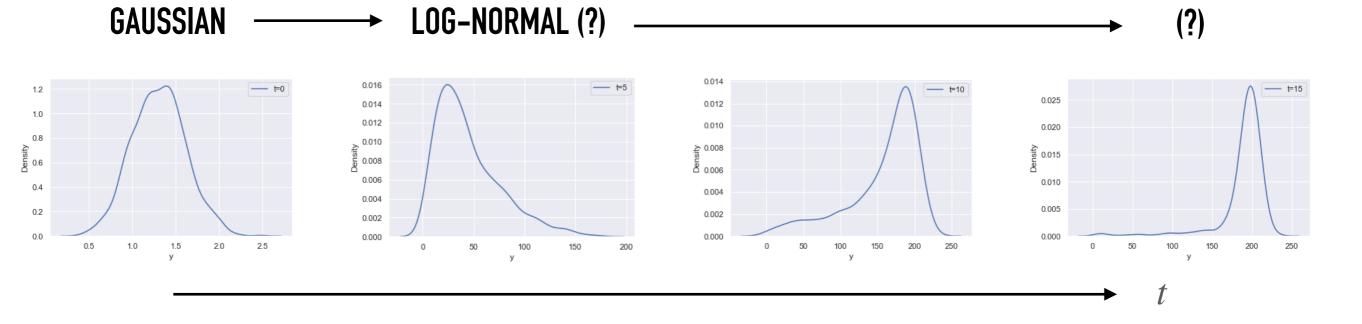
$$\bar{y}_0 \sim \mathcal{N}\left(\mu_{\bar{y}_0}, \sigma_{\bar{y}_0}^2\right)$$

$$\lambda \sim \mathcal{N}\left(\mu_{\lambda}, \sigma_{\lambda}^2\right)$$

$$c \sim \mathcal{N}\left(\mu_c, \sigma_c^2\right)$$

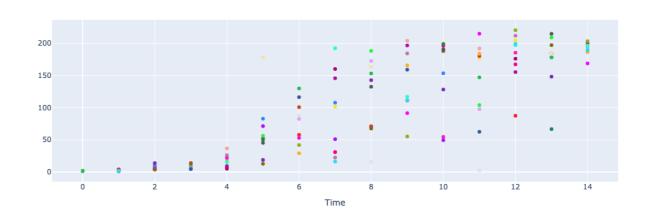
$$\epsilon \sim \mathcal{N}(0, \sigma^2)$$

TIME SLICES OF DATA-GENERATING MODEL:

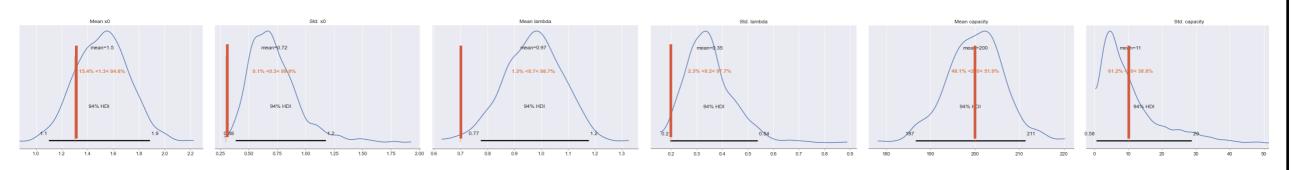


HIERARCHICAL LOGISTIC GROWTH: FEW OBSERVATIONS

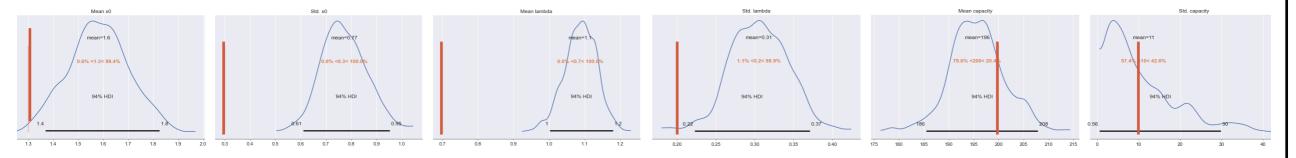
SNAPSHOT DATA: 10 INDIVIDUALS PER TIME POINT



10 SAMPLES:



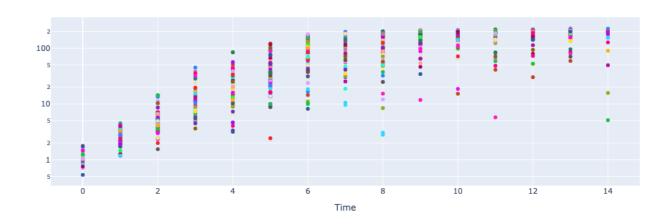
100 SAMPLES:



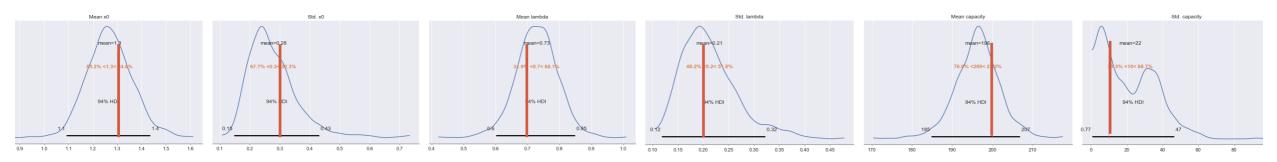
HEAVY TAILS OF TRUE DISTRIBUTION AND TOO LARGE SAMPLES-OBSERVATION RATIO LEAD TO UNDERESTIMATION OF UNCERTAINTY.

HIERARCHICAL LOGISTIC GROWTH: MORE OBSERVATIONS

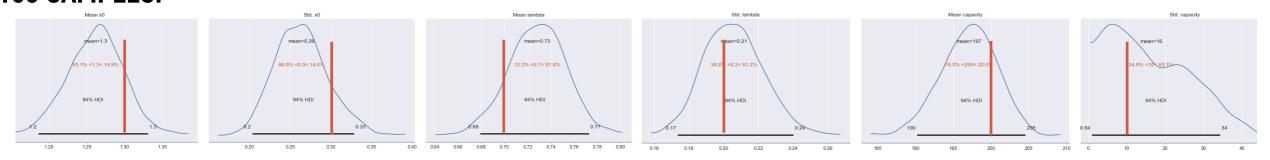
SNAPSHOT DATA: 100 INDIVIDUALS PER TIME POINT



10 SAMPLES:



100 SAMPLES:



HEAVY TAILS OF TRUE DISTRIBUTION AND TOO SMALL SAMPLES-OBSERVATION RATIO LEAD TO OVERESTIMATION OF UNCERTAINTY / SPURIOUS PEAKS IN POSTERIORS.

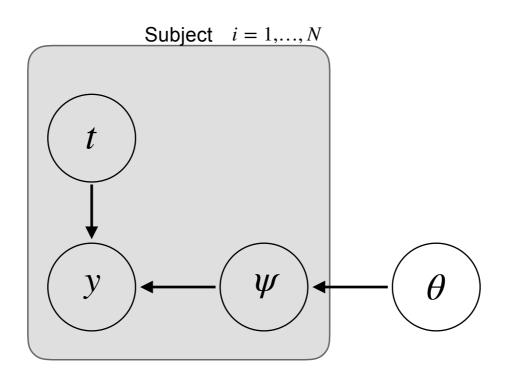
ADDRESSING Q2:

HOW EASY CAN WE FIND GOOD KDE APPROXIMATIONS WHEN A GAUSSIAN KDE ISN'T A GOOD CHOICE?

BACKUP SLIDES

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING INDIVIDUALS:

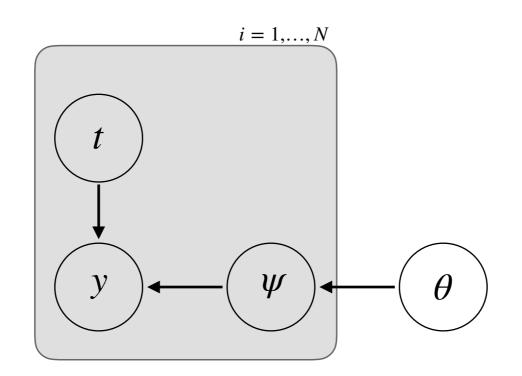
RECALL: HIERARCHAL LOG-LIKELIHOOD FOR SNAPSHOT DATA



$$\log p(\mathcal{D}, \Psi | \theta) = \sum_{i} \log p(y_i^{\text{obs}} | \psi_i, t_i) + \log p(\psi_i | \theta)$$

MOTIVATION 2: WE WANT TO DO HIERARCHICAL INFERENCE, BUT STOP FITTING INDIVIDUALS:

RECALL: HIERARCHAL LOG-LIKELIHOOD FOR SNAPSHOT DATA



SHRINKAGE

$$p(\psi | \theta)$$

$$\log p(\mathcal{D}, \Psi | \theta) = \sum_{i} \log p(y_i^{\text{obs}} | \psi_i, t_i) + \log p(\psi_i | \theta)$$