试题(A)参考答案评分标准

- 一、选择题(每小题 4 分, 共 20 分)
 - 1.A
 - 2.B
 - 3.A
 - 4.C
 - 5.C
- 二、填空题(每小题 4 分, 共 20 分)
 - 1. $\frac{7}{8}$
 - 2. $\frac{9}{64}$
 - $3.\frac{1}{2}$
 - 4. $\frac{2(n-1)}{n}\sigma^2$
 - 5. (5.616, 6.384)

三、解答题(每小题10分,共60分)

1.**解** 设 X 表示电源电压, B_1 表示事件"电源电压不超过 200V", B_2 表示事件"电源电压在 200~240V", B_3 表示事件"电源电压超过 240V",A表示事件"电子元件损坏",

则
$$X \sim N(200, 25^2)$$
, $B_1 = \{X \le 200\}, B_2 = \{200 < X \le 240\}, B_3 = \{X > 240\}$,且

$$P(B_1) = P(X \le 200) = \Phi\left(\frac{200 - 220}{25}\right) = \Phi(-0.8) = 1 - \Phi(0.8) = 0.212$$

$$P(B_2) = P(200 < X \le 240) = \mathcal{O}\left(\frac{240 - 220}{25}\right) - \mathcal{O}\left(\frac{200 - 220}{25}\right)$$

$$= \Phi(0.8) - \Phi(-0.8) = 2\Phi(0.8) - 1 = 0.576$$

$$P(B_3) = P(X > 240) = 1 - P(X \le 240) = 1 - \Phi\left(\frac{240 - 220}{25}\right) = 1 - \Phi\left(0.8\right) = 0.212$$

$$P(A|B_1) = 0.1, P(A|B_2) = 0.001, P(A|B_3) = 0.2 \dots 3$$

(1) 由全概率公式,得

$$P(A) = \sum_{i=1}^{3} P(B_i) P(A|B_i) = 0.212 \times 0.1 + 0.576 \times 0.001 + 0.212 \times 0.2$$

(2) 由 Bayes 公式, 得

$$P(B_2|A) = \frac{P(B_2)P(A|B_2)}{P(A)} = \frac{0.576 \times 0.001}{0.0642} = 0.009 \dots 3$$

(2) 由 (1) 知
$$X$$
 的概率密度为 $f(x) = \begin{cases} \frac{1}{6}x, & 0 < x < 2\sqrt{3} \\ 0, & 其他 \end{cases}$, 由于 X 是连续型随

机变量,因此所求的概率为
$$P\left(\frac{1}{2} < X < 2\right) = \int_{\frac{1}{2}}^{2} f(x) dx = \int_{\frac{1}{2}}^{2} \frac{1}{6} x dx = \frac{5}{16} \dots 2$$
 分

(3) 当
$$x < 0$$
 时, $F(x) = 0$; 当 $0 \le x < 2\sqrt{3}$ 时, $F(x) = \int_{-\infty}^{x} f(t)dt = \int_{0}^{x} \frac{1}{6}tdt = \frac{x^{2}}{12}$;

当 $x \ge 2\sqrt{3}$ 时, F(x) = 1, 即X的分布函数为

$$F(x) = \begin{cases} 0, & x < 0 \\ \frac{1}{12}x^2, & 0 \le x < 2\sqrt{3} \\ 1, & x \ge 2\sqrt{3} \end{cases}$$

(4) 当y < 0时, $F_y(y) = 0$; 当 $0 \le y < 12$ 时,

$$F_Y(y) = P(Y \le y = P(X^2 \le y) = P(X^2 < 0) + P(0 \le X^2 \le y)$$
$$= P(-\sqrt{y} \le X \le \sqrt{y}) = \int_0^{\sqrt{y}} \frac{1}{6} x dx = \frac{y}{12}$$

当 $y \ge 12$ 时, $F_Y(y) = 1$,故 Y 的分布函数为

$$F_{Y}(y) = \begin{cases} 0, & y < 0 \\ \frac{y}{12}, & 0 \le y < 12 \dots & 3 \text{ } \end{cases}$$

$$1, & x \ge 12$$

3.解 (1)
$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^{2x} dy = 2x, & 0 < x < 1 \\ 0, & 其他 \end{cases}$$
 2分

$$f_{Y}(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{\frac{y}{2}}^{1} dx = 1 - \frac{y}{2}, & 0 < y < 2 \\ 0, & \text{ 其他} \end{cases}$$

$$F_{Z}(z) = P(Z \le z) = P(2X - Y \le z) = \iint_{2x - y \le z} f(x, y) dxdy = z - \frac{z^{2}}{4}$$

当 $\frac{z}{2} \ge 1$, 即 $z \ge 2$ 时, $F_z(z) = 1$, 即Z的分布函数为

$$F_{z}(z) = \begin{cases} 0, & z < 0 \\ z - \frac{z^{2}}{4}, 0 \le z < 2 \\ 1, & z \ge 2 \end{cases}$$

再求Z的概率密度 $f_Z(z)$.

$$f_Z(z) = F_Z'(z) = \begin{cases} 1 - \frac{z}{2}, 0 < z < 2 \\ 0,$$
其它

方法二
$$f_Z(z) = \int_{-\infty}^{+\infty} f(x, 2x - z) dx$$

由
$$\begin{cases} 0 < x < 1 \\ 0 < 2x - z < 2x \end{cases}$$
, 得 $0 < z < 2x$, 从而 $0 < z < 2$.故当 $0 < z < 2$ 时, $f_z(z) > 0$,

在其他点, $f_z(z)=0$.

再由
$$\begin{cases} 0 < x < 1 \\ 0 < 2x - z < 2x \end{cases}$$
 , 得 $\begin{cases} 0 < x < 1 \\ \frac{z}{2} < x \end{cases}$, 即 $\frac{z}{2} < x < 1$,从而 Z 的概率密度为

$$f_{Z}(z) = \int_{-\infty}^{+\infty} f(x, z - x) dx = \begin{cases} \int_{\frac{\pi}{2}}^{1} 1 dx = 1 - \frac{z}{2}, & 0 < z < 2 \\ 0, & \text{ 其他} \end{cases}$$

(3)
$$P(Y \le \frac{1}{2} | X \le \frac{1}{2}) = \frac{P(X \le \frac{1}{2}, Y \le \frac{1}{2})}{P(X \le \frac{1}{2})} = \frac{\int_0^{\frac{1}{2}} dy \int_{\frac{y}{2}}^{\frac{1}{2}} dx}{\int_0^{\frac{1}{2}} 2x dx} = \frac{3}{4} \dots 2$$

4.解(1) X的可能取值为 0, 1, Y的可能取值为 0, 1, 2

$$P(X=0,Y=0) = \frac{C_3^2}{C_6^2} = \frac{1}{5}, \quad P(X=0,Y=1) = \frac{C_2^1 C_3^1}{C_6^2} = \frac{2}{5}$$

$$P(X=0,Y=2) = \frac{C_2^2}{C_6^2} = \frac{1}{15}, \quad P(X=1,Y=0) = \frac{C_1^1 C_3^1}{C_6^2} = \frac{1}{5}$$

$$P(X=1,Y=1) = \frac{C_1^1 C_2^1}{C_6^2} = \frac{2}{15}, P(X=1,Y=2) = 0$$

即(X,Y)的联合分布律为......5分

Y	0	1	2
0	<u>1</u> 5	<u>2</u> 5	1/15
1	<u>1</u> 5	2 15	0

(2) 由(X,Y)的联合分布律可得,X,Y,XY的分布律分别为

X	0	1
P	2/3	1/3

$$\begin{array}{|c|c|c|c|c|c|} \hline XY & 0 & 1 \\ \hline P & \frac{13}{15} & \frac{2}{15} \\ \hline \end{array}$$

$$EX = \frac{1}{3}$$
, $DX = \frac{2}{9}$, $EY = 0 \times \frac{2}{5} + 1 \times \frac{8}{15} + 2 \times \frac{1}{15} = \frac{2}{3}$, $EXY = \frac{2}{15}$

$$EY = 0^2 \times \frac{2}{5} + 1^2 \times \frac{8}{15} + 2^2 \times \frac{1}{15} = \frac{4}{5}, \quad EY = EY^2 - (EY)^2 = \frac{4}{5} - \frac{4}{9} = \frac{16}{45}$$

从而 $cov(X,Y) = EXY - EXEY = \frac{2}{15} - \frac{1}{3} \times \frac{2}{3} = -\frac{4}{45}$,故 X = Y 的相关系数为

5.解(1)
$$EX^2 = \int_{-\infty}^{+\infty} x f(x) dx = \int_{0}^{+\infty} x \cdot \frac{2}{\theta} x e^{-\frac{x^2}{\theta}} dx = \frac{\sqrt{\pi \theta}}{2}$$
,由矩估计法,得 $\frac{\sqrt{\pi \theta}}{2} = \bar{X}$,

(2) 对于样本 X_1, X_2, \cdots, X_n 的一组样本值 x_1, x_2, \cdots, x_n ,有

(i)似然函数为:
$$L(\theta) = \prod_{i=1}^{n} f(x_i) = \prod_{i=1}^{n} \frac{2}{\theta} x_i e^{-\frac{x_i^2}{\theta}} = \frac{2^n}{\theta^n} (\prod_{i=1}^{n} x_i) e^{-\frac{1}{\theta} \sum_{i=1}^{n} x_i^2}$$
, $x_i > 0, i = 1, 2, \dots, n$;

(ii)取自然对数:
$$\ln L(\theta) = n \ln 2 - n \ln \theta + \sum_{i=1}^{n} \ln x_i - \frac{1}{\theta} \sum_{i=1}^{n} x_i^2$$
;

(iii)令
$$\frac{d \ln L(\theta)}{d \theta} = -\frac{n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n x_i^2 = 0$$
,解之得 θ 的最大似然估计值为 $\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n x_i^2$,从

(3) 由于 $E(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_{0}^{+\infty} \frac{2}{\theta} x^3 e^{-\frac{x^2}{\theta}} dx = \theta$,因此
$E\hat{\theta}_2 = E\left[\frac{1}{n}\sum_{i=1}^n X_i^2\right] = \frac{1}{n}\sum_{i=1}^n EX_i^2 = \frac{1}{n}\sum_{i=1}^n EX^2 = \theta$
即 $\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n X_i^2$ 为 θ 的无偏估计量
从而 $D\hat{\theta}_2 = D[\frac{1}{n}\sum_{i=1}^n X_i^2] = \frac{1}{n^2}\sum_{i=1}^n DX_i^2 = \frac{1}{n^2}\sum_{i=1}^n DX^2 = \frac{\theta^2}{n}$,由切比雪夫不等式, $\forall \varepsilon > 0$,有
$0 \le P(\left \hat{\theta}_2 - \theta\right) = P(\left \hat{\theta}_2 - E\hat{\theta}_2\right \ge \varepsilon) \le \frac{D\hat{\theta}}{\varepsilon^2} = \frac{\theta^2}{n\varepsilon^2} \to 0, n \to \infty$
从而 $\lim_{n\to\infty} P(\left \hat{\theta}_2 - \theta\right \ge \varepsilon) = 0$,即 $\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n X_i^2 $ 为 θ 的相合估计量
6.解(i) 需检验
$H_0: \mu \le 225, H_1: \mu > 225.$ 25 25 25 25 25 25 25 25 25 25 25 25 25
(ii) 选择检验统计量
$T = \frac{\overline{X} - 225}{\frac{S}{\sqrt{n}}} \sim t(n-1) \dots 2 $
(iii) 由于 $\alpha = 0.05, n = 16$,因此临界点为 $t_{\alpha}(n-1) = t_{0.05}(15) = 1.7531$,从而接
受域为(-∞,1.7531)
(iv) 由于 $n = 16$, $s = 99$, $\bar{x} = 241.5$,因此检验统计量的样本值为
$t = \frac{241.5 - 225}{99 / \sqrt{16}} = 0.6667 \dots 2 \text{f}$
(v) 由于 $t=0.6667\in(-\infty,1.7531)$,因此接受 H_0 ,即可以认为元件的平均寿命不大
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