Probability and Statistics

Cookbook

This cookbook integrates a variety of topics in probability theory and statistics. It is based on literature [1, 6, 3] and in-class material from courses of the statistics department at the University of California in Berkeley but also influenced by other sources [4, 5]. If you find errors or have suggestions for further topics, I would appreciate if you send me an email. The most recent version of this document is available at http://matthias.vallentin.net/probability-and-statistics-cookbook/. To reproduce, please contact me.

Contents

1 Distribution Overview

1.1 Discrete Distributions

	Notation[1]	$F_X(x)$	$f_X(x)$	$\mathbb{E}[X]$	$\mathbb{V}[X]$	$M_X(s)$
Uniform	$\text{Unif}\{a,\ldots,b\}$	$\begin{cases} 0 & x < a \\ \frac{\lfloor x \rfloor - a + 1}{b-a} & a \le x \le b \\ 1 & x > b \end{cases}$	$\frac{I(a < x < b)}{b - a + 1}$	$\frac{a+b}{2}$	$\frac{(b-a+1)^2 - 1}{12}$	$\frac{e^{as} - e^{-(b+1)s}}{s(b-a)}$
Bernoulli	$\text{Bern}(p)$	$(1-p)^{1-x}$	$p^x(1-p)^{1-x}$	p	$p(1-p)$	$1 - p + pe^s$
Binomial	$\text{Bin}(n,p)$	$I_{1-p}(n-x, x+1)$	$\binom{n}{x} p^x (1-p)^{n-x}$	np	$np(1-p)$	$(1 - p + pe^s)^n$
Multinomial	$\text{Mult}(n,p)$		$\frac{n!}{x_1! \ldots x_k!} p_1^{x_1} \cdots p_k^{x_k} \quad \sum_{i=1}^{k} x_i = n$	np_i	$np_i(1-p_i)$	$\left(\sum_{i=0}^{k} p_i e^{s_i} \right)^n$
Hypergeometric	$\text{Hyp}(N,m,n)$	$\approx \Phi\left(\frac{x - np}{\sqrt{np(1-p)}} \right)$	$\frac{\binom{m}{x}\binom{m-x}{n-x}}{\binom{N}{x}}$	$\frac{nm}{N}$	$\frac{nm(N-n)(N-m)}{N^2(N-1)}$	
Negative Binomial	$\text{NBin}(r,p)$	$I_p(r, x+1)$	$\binom{x+r-1}{r-1} p^r (1-p)^x$	$r\frac{1-p}{p}$	$r\frac{1-p}{p^2}$	$\left(\frac{p}{1-(1-p)e^s} \right)^r$
Geometric	$\text{Geo}(p)$	$1 - (1-p)^x \quad x \in \mathbb{N}^+$	$p(1-p)^{x-1} \quad x \in \mathbb{N}^+$	$\frac{1}{p}$	$\frac{1-p}{p^2}$	$\frac{p}{1-(1-p)e^s}$
Poisson	$\text{Po}(\lambda)$	$e^{-\lambda} \sum_{i=0}^{x} \frac{\lambda^i}{i!}$	$\frac{\lambda^x e^{-\lambda}}{x!}$	λ	λ	$e^{\lambda(e^s - 1)}$
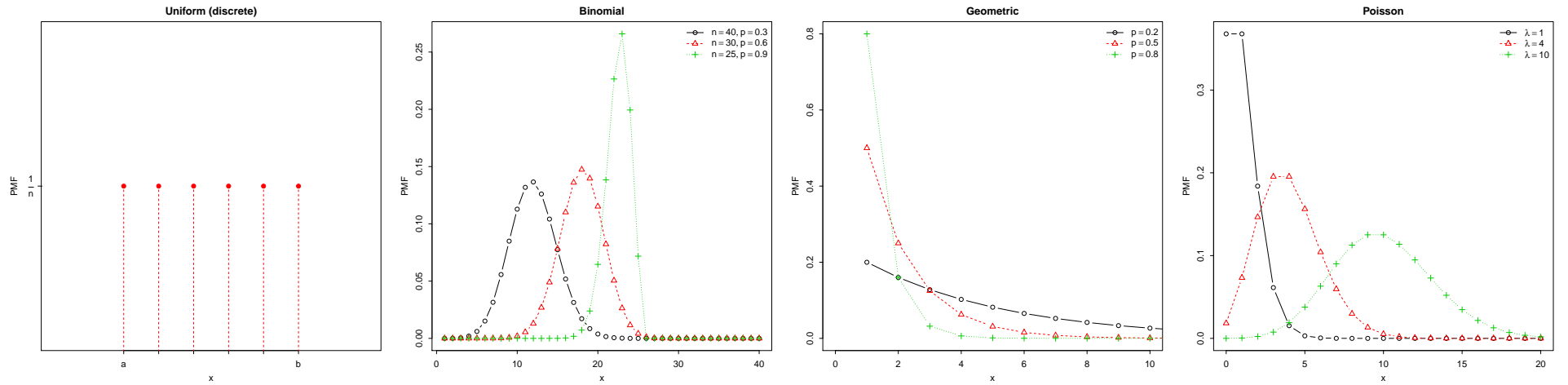
[1] We use the notation $\gamma(s,x)$ and $\Gamma(x)$ to refer to the Gamma functions (see §22.1), and use $\text{B}(x,y)$ and I_x to refer to the Beta functions (see §22.2).

1.2 Continuous Distributions

	Notation	$F_X(x)$	$f_X(x)$	$\mathbb{E}[X]$	$\mathbb{V}[X]$	$M_X(s)$
Uniform	$\text{Unif}(a,b)$	$\begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a < x < b \\ 1 & x > b \end{cases}$	$\dfrac{I(a < x < b)}{b-a}$	$\dfrac{a+b}{2}$	$\dfrac{(b-a)^2}{12}$	$\dfrac{e^{sb} - e^{sa}}{s(b-a)}$
Normal	$\mathcal{N}(\mu, \sigma^2)$	$\Phi(x) = \displaystyle\int_{-\infty}^{x} \phi(t)\,dt$	$\phi(x) = \dfrac{1}{\sigma\sqrt{2\pi}} \exp\left\{ -\dfrac{(x-\mu)^2}{2\sigma^2} \right\}$	μ	σ^2	$\exp\left\{ \mu s + \dfrac{\sigma^2 s^2}{2} \right\}$
Log-Normal	$\ln\mathcal{N}(\mu, \sigma^2)$	$\dfrac{1}{2} + \dfrac{1}{2}\,\text{erf}\left[\dfrac{\ln x - \mu}{\sqrt{2\sigma^2}} \right]$	$\dfrac{1}{x\sqrt{2\pi\sigma^2}} \exp\left\{ -\dfrac{(\ln x - \mu)^2}{2\sigma^2} \right\}$	$e^{\mu + \sigma^2/2}$	$(e^{\sigma^2} - 1)e^{2\mu + \sigma^2}$	
Multivariate Normal	$\text{MVN}(\mu, \Sigma)$		$(2\pi)^{-k/2}|\Sigma|^{-1/2} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$	μ	Σ	$\exp\left\{ \mu^T s + \dfrac{1}{2} s^T \Sigma s \right\}$
Student's t	$\text{Student}(\nu)$	$I_x\left(\dfrac{\nu}{2}, \dfrac{\nu}{2} \right)$	$\dfrac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi}\,\Gamma\left(\frac{\nu}{2}\right)} \left(1 + \dfrac{x^2}{\nu} \right)^{-(\nu+1)/2}$	0	0	
Chi-square	χ_k^2	$\dfrac{1}{\Gamma(k/2)} \gamma\left(\dfrac{k}{2}, \dfrac{x}{2} \right)$	$\dfrac{1}{2^{k/2}\Gamma(k/2)} x^{k/2} e^{-x/2}$	k	$2k$	$(1-2s)^{-k/2}\; s < 1/2$
F	$\text{F}(d_1, d_2)$	$I_{\frac{d_1 x}{d_1 x + d_2}}\left(\dfrac{d_1}{2}, \dfrac{d_1}{2} \right)$	$\dfrac{\sqrt{\frac{(d_1 x)^{d_1} d_2^{d_2}}{(d_1 x + d_2)^{d_1 + d_2}}}}{x\,\text{B}\left(\frac{d_1}{2}, \frac{d_1}{2} \right)}$	$\dfrac{d_2}{d_2 - 2}$	$\dfrac{2d_2^2(d_1 + d_2 - 2)}{d_1(d_2-2)^2(d_2-4)}$	
Exponential	$\text{Exp}(\beta)$	$1 - e^{-x/\beta}$	$\dfrac{1}{\beta} e^{-x/\beta}$	β	β^2	$\dfrac{1}{1-\beta s}\,(s < 1/\beta)$
Gamma	$\text{Gamma}(\alpha, \beta)$	$\dfrac{\gamma(\alpha, x/\beta)}{\Gamma(\alpha)}$	$\dfrac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}$	$\alpha\beta$	$\alpha\beta^2$	$\left(\dfrac{1}{1-\beta s} \right)^\alpha (s < 1/\beta)$
Inverse Gamma	$\text{InvGamma}(\alpha, \beta)$	$\dfrac{\Gamma\left(\alpha, \frac{\beta}{x} \right)}{\Gamma(\alpha)}$	$\dfrac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x}$	$\dfrac{\beta}{\alpha-1}\;\alpha > 1$	$\dfrac{\beta^2}{(\alpha-1)^2(\alpha-2)^2}\;\alpha > 2$	$\dfrac{2(-\beta s)^{\alpha/2}}{\Gamma(\alpha)} K_\alpha\left(\sqrt{-4\beta s} \right)$
Dirichlet	$\text{Dir}(\alpha)$		$\dfrac{\Gamma\left(\sum_{i=1}^k \alpha_i \right)}{\prod_{i=1}^k \Gamma(\alpha_i)} \prod_{i=1}^k x_i^{\alpha_i - 1}$	$\dfrac{\alpha_i}{\sum_{i=1}^k \alpha_i}$	$\dfrac{\mathbb{E}[X_i](1 - \mathbb{E}[X_i])}{\sum_{i=1}^k \alpha_i + 1}$	
Beta	$\text{Beta}(\alpha, \beta)$	$I_x(\alpha, \beta)$	$\dfrac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1}(1-x)^{\beta-1}$	$\dfrac{\alpha}{\alpha+\beta}$	$\dfrac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$	$1 + \displaystyle\sum_{k=1}^{\infty} \left(\prod_{r=0}^{k-1} \dfrac{\alpha+r}{\alpha+\beta+r} \right) \dfrac{s^k}{k!}$
Weibull	$\text{Weibull}(\lambda, k)$	$1 - e^{-(x/\lambda)^k}$	$\dfrac{k}{\lambda}\left(\dfrac{x}{\lambda} \right)^{k-1} e^{-(x/\lambda)^k}$	$\lambda\Gamma\left(1 + \dfrac{1}{k} \right)$	$\lambda^2\Gamma\left(1 + \dfrac{2}{k} \right) - \mu^2$	$\displaystyle\sum_{n=0}^{\infty} \dfrac{s^n \lambda^n}{n!} \Gamma\left(1 + \dfrac{n}{k} \right)$
Pareto	$\text{Pareto}(x_m, \alpha)$	$1 - \left(\dfrac{x_m}{x} \right)^\alpha\; x \geq x_m$	$\alpha \dfrac{x_m^\alpha}{x^{\alpha+1}}\quad x \geq x_m$	$\dfrac{\alpha x_m}{\alpha-1}\;\alpha > 1$	$\dfrac{x_m^\alpha}{(\alpha-1)^2(\alpha-2)}\;\alpha > 2$	$\alpha(-x_m s)^\alpha \Gamma(-\alpha, -x_m s)\; s < 0$
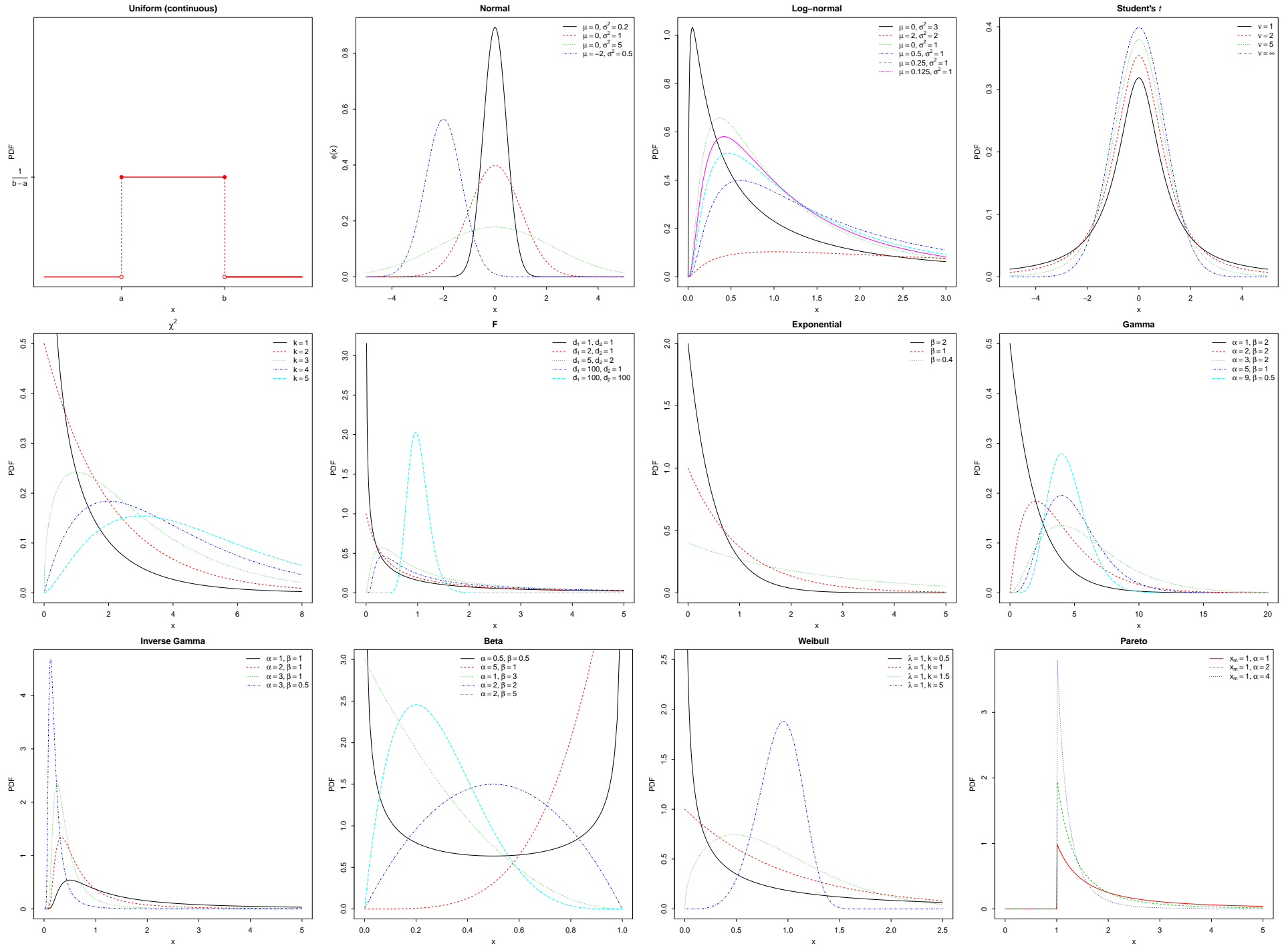
2 Probability Theory

Definitions

- Sample space Ω
- Outcome (point or element) $\omega \in \Omega$
- Event $A \subseteq \Omega$
- σ-algebra $\mathcal{A}$
  1. $\emptyset \in \mathcal{A}$
  2. $A_1, A_2, \ldots, \in \mathcal{A} \implies \bigcup_{i=1}^{\infty} A_i \in \mathcal{A}$
  3. $A \in \mathcal{A} \implies \neg A \in \mathcal{A}$
- Probability Distribution $\mathbb{P}$
  1. $\mathbb{P}[A] \geq 0 \quad \forall A$
  2. $\mathbb{P}[\Omega] = 1$
  3. $\mathbb{P}\left[\bigsqcup_{i=1}^{\infty} A_i\right] = \sum_{i=1}^{\infty} \mathbb{P}[A_i]$
- Probability space $(\Omega, \mathcal{A}, \mathbb{P})$

Properties

- $\mathbb{P}[\emptyset] = 0$
- $B = \Omega \cap B = (A \cup \neg A) \cap B = (A \cap B) \cup (\neg A \cap B)$
- $\mathbb{P}[\neg A] = 1 - \mathbb{P}[A]$
- $\mathbb{P}[B] = \mathbb{P}[A \cap B] + \mathbb{P}[\neg A \cap B]$
- $\mathbb{P}[\Omega] = 1 \qquad \mathbb{P}[\emptyset] = 0$
- $\neg(\bigcup_n A_n) = \bigcap_n \neg A_n \quad \neg(\bigcap_n A_n) = \bigcup_n \neg A_n \qquad$ DeMorgan
- $\mathbb{P}[\bigcup_n A_n] = 1 - \mathbb{P}[\bigcap_n \neg A_n]$
- $\mathbb{P}[A \cup B] = \mathbb{P}[A] + \mathbb{P}[B] - \mathbb{P}[A \cap B]$
  $\implies \mathbb{P}[A \cup B] \leq \mathbb{P}[A] + \mathbb{P}[B]$
- $\mathbb{P}[A \cup B] = \mathbb{P}[A \cap \neg B] + \mathbb{P}[\neg A \cap B] + \mathbb{P}[A \cap B]$
- $\mathbb{P}[A \cap \neg B] = \mathbb{P}[A] - \mathbb{P}[A \cap B]$

Continuity of Probabilities

- $A_1 \subset A_2 \subset \ldots \implies \lim_{n \to \infty} \mathbb{P}[A_n] = \mathbb{P}[A] \quad$ where $A = \bigcup_{i=1}^{\infty} A_i$
- $A_1 \supset A_2 \supset \ldots \implies \lim_{n \to \infty} \mathbb{P}[A_n] = \mathbb{P}[A] \quad$ where $A = \bigcap_{i=1}^{\infty} A_i$

Independence $\perp\!\!\!\perp$
$$A \perp\!\!\!\perp B \iff \mathbb{P}[A \cap B] = \mathbb{P}[A]\mathbb{P}[B]$$

Conditional Probability
$$\mathbb{P}[A \mid B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]} \qquad \mathbb{P}[B] > 0$$

Law of Total Probability
$$\mathbb{P}[B] = \sum_{i=1}^{n} \mathbb{P}[B|A_i]\mathbb{P}[A_i] \qquad \Omega = \bigsqcup_{i=1}^{n} A_i$$

BAYES' THEOREM
$$\mathbb{P}[A_i \mid B] = \frac{\mathbb{P}[B \mid A_i]\mathbb{P}[A_i]}{\sum_{j=1}^{n} \mathbb{P}[B \mid A_j]\mathbb{P}[A_j]} \qquad \Omega = \bigsqcup_{i=1}^{n} A_i$$

Inclusion-Exclusion Principle
$$\left|\bigcup_{i=1}^{n} A_i\right| = \sum_{r=1}^{n} (-1)^{r-1} \sum_{i \leq i_1 < \cdots < i_r \leq n} \left|\bigcap_{j=1}^{r} A_{i_j}\right|$$

3 Random Variables

Random Variable (RV)
$$X : \Omega \to \mathbb{R}$$

Probability Mass Function (PMF)
$$f_X(x) = \mathbb{P}[X = x] = \mathbb{P}[\{\omega \in \Omega : X(\omega) = x\}]$$

Probability Density Function (PDF)
$$\mathbb{P}[a \leq X \leq b] = \int_a^b f(x)\,dx$$

Cumulative Distribution Function (CDF)
$$F_X : \mathbb{R} \to [0,1] \qquad F_X(x) = \mathbb{P}[X \leq x]$$

1. Nondecreasing: $x_1 < x_2 \implies F(x_1) \leq F(x_2)$
2. Normalized: $\lim_{x \to -\infty} = 0$ and $\lim_{x \to \infty} = 1$
3. Right-Continuous: $\lim_{y \downarrow x} F(y) = F(x)$

$$\mathbb{P}[a \leq Y \leq b \mid X = x] = \int_a^b f_{Y|X}(y \mid x)dy \qquad a \leq b$$

$$f_{Y|X}(y \mid x) = \frac{f(x,y)}{f_X(x)}$$

Independence

1. $\mathbb{P}[X \leq x, Y \leq y] = \mathbb{P}[X \leq x]\mathbb{P}[Y \leq y]$
2. $f_{X,Y}(x,y) = f_X(x)f_Y(y)$

3.1 Transformations

Transformation function
$$Z = \varphi(X)$$

Discrete
$$f_Z(z) = \mathbb{P}\left[\varphi(X) = z\right] = \mathbb{P}\left[\{x : \varphi(x) = z\}\right] = \mathbb{P}\left[X \in \varphi^{-1}(z)\right] = \sum_{x \in \varphi^{-1}(z)} f(x)$$

Continuous
$$F_Z(z) = \mathbb{P}\left[\varphi(X) \le z\right] = \int_{A_z} f(x)\, dx \quad \text{with } A_z = \{x : \varphi(x) \le z\}$$

Special case if φ strictly monotone
$$f_Z(z) = f_X(\varphi^{-1}(z)) \left|\frac{d}{dz}\varphi^{-1}(z)\right| = f_X(x)\left|\frac{dx}{dz}\right| = f_X(x)\frac{1}{|J|}$$

The Rule of the Lazy Statistician
$$\mathbb{E}\left[Z\right] = \int \varphi(x)\, dF_X(x)$$

$$\mathbb{E}\left[I_A(x)\right] = \int I_A(x)\, dF_X(x) = \int_A dF_X(x) = \mathbb{P}\left[X \in A\right]$$

Convolution

- $Z := X + Y \qquad f_Z(z) = \int_{-\infty}^{\infty} f_{X,Y}(x, z-x)\, dx \overset{X,Y \ge 0}{=} \int_0^z f_{X,Y}(x, z-x)\, dx$

- $Z := |X - Y| \qquad f_Z(z) = 2\int_0^{\infty} f_{X,Y}(x, z+x)\, dx$

- $Z := \dfrac{X}{Y} \qquad f_Z(z) = \int_{-\infty}^{\infty} |x| f_{X,Y}(x, xz)\, dx \overset{\perp\!\!\!\perp}{=} \int_{-\infty}^{\infty} x f_x(x) f_X(x) f_Y(xz)\, dx$

4 Expectation

Definition and properties

- $\mathbb{E}\left[X\right] = \mu_X = \int x\, dF_X(x) = \begin{cases} \sum_x x f_X(x) & \text{X discrete} \\[2ex] \int x f_X(x) & \text{X continuous} \end{cases}$

- $\mathbb{P}\left[X = c\right] = 1 \implies \mathbb{E}\left[c\right] = c$
- $\mathbb{E}\left[cX\right] = c\,\mathbb{E}\left[X\right]$
- $\mathbb{E}\left[X + Y\right] = \mathbb{E}\left[X\right] + \mathbb{E}\left[Y\right]$

- $\mathbb{E}\left[XY\right] = \int_{X,Y} xy f_{X,Y}(x,y)\, dF_X(x)\, dF_Y(y)$
- $\mathbb{E}\left[\varphi(Y)\right] \ne \varphi(\mathbb{E}\left[X\right]) \qquad$ (cf. JENSEN inequality)
- $\mathbb{P}\left[X \ge Y\right] = 0 \implies \mathbb{E}\left[X\right] \ge \mathbb{E}\left[Y\right] \wedge \mathbb{P}\left[X = Y\right] = 1 \implies \mathbb{E}\left[X\right] = \mathbb{E}\left[Y\right]$
- $\mathbb{E}\left[X\right] = \sum_{x=1}^{\infty} \mathbb{P}\left[X \ge x\right]$

Sample mean
$$\bar{X}_n = \frac{1}{n}\sum_{i=1}^{n} X_i$$

Conditional expectation

- $\mathbb{E}\left[Y \mid X = x\right] = \int y f(y \mid x)\, dy$
- $\mathbb{E}\left[X\right] = \mathbb{E}\left[\mathbb{E}\left[X \mid Y\right]\right]$
- $E[\varphi(X,Y) \mid X = x] = \int_{-\infty}^{\infty} \varphi(x,y) f_{Y|X}(y \mid x)\, dx$
- $\mathbb{E}\left[\varphi(Y,Z) \mid X = x\right] = \int_{-\infty}^{\infty} \varphi(y,z) f_{(Y,Z)|X}(y,z \mid x)\, dy\, dz$
- $\mathbb{E}\left[Y + Z \mid X\right] = \mathbb{E}\left[Y \mid X\right] + \mathbb{E}\left[Z \mid X\right]$
- $\mathbb{E}\left[\varphi(X)Y \mid X\right] = \varphi(X)\mathbb{E}\left[Y \mid X\right]$
- $E[Y \mid X] = c \implies \mathrm{Cov}\left[X,Y\right] = 0$

5 Variance

Definition and properties

- $\mathbb{V}\left[X\right] = \sigma_X^2 = \mathbb{E}\left[(X - \mathbb{E}\left[X\right])^2\right] = \mathbb{E}\left[X^2\right] - \mathbb{E}\left[X\right]^2$
- $\mathbb{V}\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} \mathbb{V}\left[X_i\right] + 2\sum_{i \ne j} \mathrm{Cov}\left[X_i, Y_j\right]$
- $\mathbb{V}\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} \mathbb{V}\left[X_i\right] \quad \text{iff } X_i \perp\!\!\!\perp X_j$

Standard deviation
$$\mathsf{sd}[X] = \sqrt{\mathbb{V}\left[X\right]} = \sigma_X$$

Covariance

- $\mathrm{Cov}\left[X,Y\right] = \mathbb{E}\left[(X - \mathbb{E}\left[X\right])(Y - \mathbb{E}\left[Y\right])\right] = \mathbb{E}\left[XY\right] - \mathbb{E}\left[X\right]\mathbb{E}\left[Y\right]$
- $\mathrm{Cov}\left[X,a\right] = 0$
- $\mathrm{Cov}\left[X,X\right] = \mathbb{V}\left[X\right]$
- $\mathrm{Cov}\left[X,Y\right] = \mathrm{Cov}\left[Y,X\right]$
- $\mathrm{Cov}\left[aX, bY\right] = ab\,\mathrm{Cov}\left[X,Y\right]$

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- $\text{Cov}\left[X+a, Y+b\right] = \text{Cov}\left[X, Y\right]$
- $\text{Cov}\left[\sum_{i=1}^{n} X_i, \sum_{j=1}^{m} Y_j\right] = \sum_{i=1}^{n}\sum_{j=1}^{m} \text{Cov}\left[X_i, Y_j\right]$

Correlation

$$\rho\left[X, Y\right] = \frac{\text{Cov}\left[X, Y\right]}{\sqrt{\mathbb{V}\left[X\right]\mathbb{V}\left[Y\right]}}$$

Independence

$$X \perp\!\!\!\perp Y \implies \rho\left[X, Y\right] = 0 \iff \text{Cov}\left[X, Y\right] = 0 \iff \mathbb{E}\left[XY\right] = \mathbb{E}\left[X\right]\mathbb{E}\left[Y\right]$$

Sample variance

$$S^2 = \frac{1}{n-1}\sum_{i=1}^{n}(X_i - \bar{X}_n)^2$$

Conditional variance

- $\mathbb{V}\left[Y \mid X\right] = \mathbb{E}\left[(Y - \mathbb{E}\left[Y \mid X\right])^2 \mid X\right] = \mathbb{E}\left[Y^2 \mid X\right] - \mathbb{E}\left[Y \mid X\right]^2$
- $\mathbb{V}\left[Y\right] = \mathbb{E}\left[\mathbb{V}\left[Y \mid X\right]\right] + \mathbb{V}\left[\mathbb{E}\left[Y \mid X\right]\right]$

6 Inequalities

CAUCHY-SCHWARZ
$$\mathbb{E}\left[XY\right]^2 \le \mathbb{E}\left[X^2\right]\mathbb{E}\left[Y^2\right]$$

MARKOV
$$\mathbb{P}\left[\varphi(X) \ge t\right] \le \frac{\mathbb{E}\left[\varphi(X)\right]}{t}$$

CHEBYSHEV
$$\mathbb{P}\left[|X - \mathbb{E}\left[X\right]| \ge t\right] \le \frac{\mathbb{V}\left[X\right]}{t^2}$$

CHERNOFF
$$\mathbb{P}\left[X \ge (1+\delta)\mu\right] \le \left(\frac{e^\delta}{(1+\delta)^{1+\delta}}\right) \quad \delta > -1$$

JENSEN
$$\mathbb{E}\left[\varphi(X)\right] \ge \varphi(\mathbb{E}\left[X\right]) \quad \varphi \text{ convex}$$

7 Distribution Relationships

Binomial

- $X_i \sim \text{Bern}\left(p\right) \implies \sum_{i=1}^{n} X_i \sim \text{Bin}\left(n, p\right)$
- $X \sim \text{Bin}\left(n, p\right), Y \sim \text{Bin}\left(m, p\right) \implies X + Y \sim \text{Bin}\left(n+m, p\right)$

- $\lim_{n\to\infty} \text{Bin}\left(n, p\right) = \text{Po}\left(np\right) \qquad (n \text{ large}, p \text{ small})$
- $\lim_{n\to\infty} \text{Bin}\left(n, p\right) = \mathcal{N}\left(np, np(1-p)\right) \qquad (n \text{ large}, p \text{ far from 0 and 1})$

Negative Binomial

- $X \sim \text{NBin}\left(1, p\right) = \text{Geo}\left(p\right)$
- $X \sim \text{NBin}\left(r, p\right) = \sum_{i=1}^{r} \text{Geo}\left(p\right)$
- $X_i \sim \text{NBin}\left(r_i, p\right) \implies \sum X_i \sim \text{NBin}\left(\sum r_i, p\right)$
- $X \sim \text{NBin}\left(r, p\right). Y \sim \text{Bin}\left(s+r, p\right) \implies \mathbb{P}\left[X \le s\right] = \mathbb{P}\left[Y \ge r\right]$

Poisson

- $X_i \sim \text{Po}\left(\lambda_i\right) \wedge X_i \perp\!\!\!\perp X_j \implies \sum_{i=1}^{n} X_i \sim \text{Po}\left(\sum_{i=1}^{n}\lambda_i\right)$
- $X_i \sim \text{Po}\left(\lambda_i\right) \wedge X_i \perp\!\!\!\perp X_j \implies X_i \left|\sum_{j=1}^{n} X_j \sim \text{Bin}\left(\sum_{j=1}^{n} X_j, \frac{\lambda_i}{\sum_{j=1}^{n}\lambda_j}\right)\right.$

Exponential

- $X_i \sim \text{Exp}\left(\beta\right) \wedge X_i \perp\!\!\!\perp X_j \implies \sum_{i=1}^{n} X_i \sim \text{Gamma}\left(n, \beta\right)$
- Memoryless property: $\mathbb{P}\left[X > x+y \mid X > y\right] = \mathbb{P}\left[X > x\right]$

Normal

- $X \sim \mathcal{N}\left(\mu, \sigma^2\right) \implies \left(\frac{X-\mu}{\sigma}\right) \sim \mathcal{N}\left(0, 1\right)$
- $X \sim \mathcal{N}\left(\mu, \sigma^2\right) \wedge Z = aX + b \implies Z \sim \mathcal{N}\left(a\mu + b, a^2\sigma^2\right)$
- $X \sim \mathcal{N}\left(\mu_1, \sigma_1^2\right) \wedge Y \sim \mathcal{N}\left(\mu_2, \sigma_2^2\right) \implies X + Y \sim \mathcal{N}\left(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2\right)$
- $X_i \sim \mathcal{N}\left(\mu_i, \sigma_i^2\right) \implies \sum_i X_i \sim \mathcal{N}\left(\sum_i \mu_i, \sum_i \sigma_i^2\right)$
- $\mathbb{P}\left[a < X \le b\right] = \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$
- $\Phi(-x) = 1 - \Phi(x) \qquad \phi'(x) = -x\phi(x) \qquad \phi''(x) = (x^2 - 1)\phi(x)$
- Upper quantile of $\mathcal{N}\left(0, 1\right)$: $z_\alpha = \Phi^{-1}(1-\alpha)$

Gamma

- $X \sim \text{Gamma}\left(\alpha, \beta\right) \iff X/\beta \sim \text{Gamma}\left(\alpha, 1\right)$
- $\text{Gamma}\left(\alpha, \beta\right) \sim \sum_{i=1}^{\alpha} \text{Exp}\left(\beta\right)$
- $X_i \sim \text{Gamma}\left(\alpha_i, \beta\right) \wedge X_i \perp\!\!\!\perp X_j \implies \sum_i X_i \sim \text{Gamma}\left(\sum_i \alpha_i, \beta\right)$
- $\frac{\Gamma(\alpha)}{\lambda^\alpha} = \int_0^\infty x^{\alpha-1}e^{-\lambda x}\, dx$

Beta

- $\frac{1}{\text{B}(\alpha, \beta)}x^{\alpha-1}(1-x)^{\beta-1} = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}x^{\alpha-1}(1-x)^{\beta-1}$
- $\mathbb{E}\left[X^k\right] = \frac{\text{B}(\alpha+k, \beta)}{\text{B}(\alpha, \beta)} = \frac{\alpha+k-1}{\alpha+\beta+k-1}\mathbb{E}\left[X^{k-1}\right]$
- $\text{Beta}\left(1, 1\right) \sim \text{Unif}\left(0, 1\right)$

8 Probability and Moment Generating Functions

- $G_X(t) = \mathbb{E}\left[t^X\right] \qquad |t| < 1$

- $M_X(t) = G_X(e^t) = \mathbb{E}\left[e^{Xt}\right] = \mathbb{E}\left[\sum_{i=0}^{\infty} \frac{(Xt)^i}{i!}\right] = \sum_{i=0}^{\infty} \frac{\mathbb{E}\left[X^i\right]}{i!} \cdot t^i$

- $\mathbb{P}\left[X = 0\right] = G_X(0)$
- $\mathbb{P}\left[X = 1\right] = G'_X(0)$

- $\mathbb{P}\left[X = i\right] = \dfrac{G_X^{(i)}(0)}{i!}$

- $\mathbb{E}\left[X\right] = G'_X(1^-)$
- $\mathbb{E}\left[X^k\right] = M_X^{(k)}(0)$

- $\mathbb{E}\left[\dfrac{X!}{(X-k)!}\right] = G_X^{(k)}(1^-)$

- $\mathbb{V}\left[X\right] = G''_X(1^-) + G'_X(1^-) - \left(G'_X(1^-)\right)^2$
- $G_X(t) = G_Y(t) \implies X \overset{d}{=} Y$

9 Multivariate Distributions

9.1 Standard Bivariate Normal

Let $X, Y \sim \mathcal{N}(0, 1) \wedge X \perp\!\!\!\perp Z$ where $Y = \rho X + \sqrt{1 - \rho^2} Z$

Joint density
$$f(x, y) = \frac{1}{2\pi\sqrt{1 - \rho^2}} \exp\left\{-\frac{x^2 + y^2 - 2\rho xy}{2(1 - \rho^2)}\right\}$$

Conditionals
$$(Y \mid X = x) \sim \mathcal{N}\left(\rho x, 1 - \rho^2\right) \qquad \text{and} \qquad (X \mid Y = y) \sim \mathcal{N}\left(\rho y, 1 - \rho^2\right)$$

Independence
$$X \perp\!\!\!\perp Y \iff \rho = 0$$

9.2 Bivariate Normal

Let $X \sim \mathcal{N}\left(\mu_x, \sigma_x^2\right)$ and $Y \sim \mathcal{N}\left(\mu_y, \sigma_y^2\right)$.

$$f(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1 - \rho^2}} \exp\left\{-\frac{z}{2(1 - \rho^2)}\right\}$$

$$z = \left[\left(\frac{x - \mu_x}{\sigma_x}\right)^2 + \left(\frac{y - \mu_y}{\sigma_y}\right)^2 - 2\rho\left(\frac{x - \mu_x}{\sigma_x}\right)\left(\frac{y - \mu_y}{\sigma_y}\right)\right]$$

Conditional mean and variance

$$\mathbb{E}\left[X \mid Y\right] = \mathbb{E}\left[X\right] + \rho\frac{\sigma_X}{\sigma_Y}(Y - \mathbb{E}\left[Y\right])$$

$$\mathbb{V}\left[X \mid Y\right] = \sigma_X\sqrt{1 - \rho^2}$$

9.3 Multivariate Normal

Covariance matrix Σ \quad (Precision matrix Σ^{-1})

$$\Sigma = \begin{pmatrix} \mathbb{V}\left[X_1\right] & \cdots & \mathrm{Cov}\left[X_1, X_k\right] \\ \vdots & \ddots & \vdots \\ \mathrm{Cov}\left[X_k, X_1\right] & \cdots & \mathbb{V}\left[X_k\right] \end{pmatrix}$$

If $X \sim \mathcal{N}(\mu, \Sigma)$,

$$f_X(x) = (2\pi)^{-n/2}|\Sigma|^{-1/2}\exp\left\{-\frac{1}{2}(x - \mu)^T\Sigma^{-1}(x - \mu)\right\}$$

Properties

- $Z \sim \mathcal{N}(0, 1) \wedge X = \mu + \Sigma^{1/2}Z \implies X \sim \mathcal{N}(\mu, \Sigma)$
- $X \sim \mathcal{N}(\mu, \Sigma) \implies \Sigma^{-1/2}(X - \mu) \sim \mathcal{N}(0, 1)$
- $X \sim \mathcal{N}(\mu, \Sigma) \implies AX \sim \mathcal{N}\left(A\mu, A\Sigma A^T\right)$
- $X \sim \mathcal{N}(\mu, \Sigma) \wedge \|a\| = k \implies a^T X \sim \mathcal{N}\left(a^T\mu, a^T\Sigma a\right)$

10 Convergence

Let $\{X_1, X_2, \ldots\}$ be a sequence of RV's and let X be another RV. Let F_n denote the CDF of X_n and let F denote the CDF of X.

Types of convergence

1. In distribution (weakly, in law): $X_n \overset{\mathrm{D}}{\to} X$

$$\lim_{n \to \infty} F_n(t) = F(t) \qquad \forall t \text{ where } F \text{ continuous}$$

2. In probability: $X_n \overset{\mathrm{P}}{\to} X$

$$(\forall \varepsilon > 0) \lim_{n \to \infty} \mathbb{P}\left[|X_n - X| > \varepsilon\right] = 0$$

3. Almost surely (strongly): $X_n \overset{\mathrm{as}}{\to} X$

$$\mathbb{P}\left[\lim_{n \to \infty} X_n = X\right] = \mathbb{P}\left[\omega \in \Omega : \lim_{n \to \infty} X_n(\omega) = X(\omega)\right] = 1$$

4. In quadratic mean (L_2): $X_n \overset{\text{qm}}{\to} X$

$$\lim_{n\to\infty} \mathbb{E}\left[(X_n - X)^2\right] = 0$$

Relationships

- $X_n \overset{\text{qm}}{\to} X \implies X_n \overset{\text{P}}{\to} X \implies X_n \overset{\text{D}}{\to} X$
- $X_n \overset{\text{as}}{\to} X \implies X_n \overset{\text{P}}{\to} X$
- $X_n \overset{\text{D}}{\to} X \wedge (\exists c \in \mathbb{R})\, \mathbb{P}\left[X = c\right] = 1 \implies X_n \overset{\text{P}}{\to} X$
- $X_n \overset{\text{P}}{\to} X \wedge Y_n \overset{\text{P}}{\to} Y \implies X_n + Y_n \overset{\text{P}}{\to} X + Y$
- $X_n \overset{\text{qm}}{\to} X \wedge Y_n \overset{\text{qm}}{\to} Y \implies X_n + Y_n \overset{\text{qm}}{\to} X + Y$
- $X_n \overset{\text{P}}{\to} X \wedge Y_n \overset{\text{P}}{\to} Y \implies X_n Y_n \overset{\text{P}}{\to} XY$
- $X_n \overset{\text{P}}{\to} X \implies \varphi(X_n) \overset{\text{P}}{\to} \varphi(X)$
- $X_n \overset{\text{D}}{\to} X \implies \varphi(X_n) \overset{\text{D}}{\to} \varphi(X)$
- $X_n \overset{\text{qm}}{\to} b \iff \lim_{n\to\infty} \mathbb{E}\left[X_n\right] = b \wedge \lim_{n\to\infty} \mathbb{V}\left[X_n\right] = 0$
- $X_1, \ldots, X_n$ IID $\wedge \mathbb{E}\left[X\right] = \mu \wedge \mathbb{V}\left[X\right] < \infty \iff \bar{X}_n \overset{\text{qm}}{\to} \mu$

SLUTZKY'S THEOREM

- $X_n \overset{\text{D}}{\to} X$ and $Y_n \overset{\text{P}}{\to} c \implies X_n + Y_n \overset{\text{D}}{\to} X + c$
- $X_n \overset{\text{D}}{\to} X$ and $Y_n \overset{\text{P}}{\to} c \implies X_n Y_n \overset{\text{D}}{\to} cX$
- In general: $X_n \overset{\text{D}}{\to} X$ and $Y_n \overset{\text{D}}{\to} Y \not\implies X_n + Y_n \overset{\text{D}}{\to} X + Y$

10.1 Law of Large Numbers (LLN)

Let $\{X_1, \ldots, X_n\}$ be a sequence of IID RV's, $\mathbb{E}\left[X_1\right] = \mu$, and $\mathbb{V}\left[X_1\right] < \infty$.

Weak (WLLN)

$$\bar{X}_n \overset{\text{P}}{\to} \mu \qquad n \to \infty$$

Strong (WLLN)

$$\bar{X}_n \overset{\text{as}}{\to} \mu \qquad n \to \infty$$

10.2 Central Limit Theorem (CLT)

Let $\{X_1, \ldots, X_n\}$ be a sequence of IID RV's, $\mathbb{E}\left[X_1\right] = \mu$, and $\mathbb{V}\left[X_1\right] = \sigma^2$.

$$Z_n := \frac{\bar{X}_n - \mu}{\sqrt{\mathbb{V}\left[\bar{X}_n\right]}} = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \overset{\text{D}}{\to} Z \qquad \text{where } Z \sim \mathcal{N}(0, 1)$$

$$\lim_{n\to\infty} \mathbb{P}\left[Z_n \leq z\right] = \Phi(z) \qquad z \in \mathbb{R}$$

CLT notations

$$Z_n \approx \mathcal{N}(0, 1)$$
$$\bar{X}_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$
$$\bar{X}_n - \mu \approx \mathcal{N}\left(0, \frac{\sigma^2}{n}\right)$$
$$\sqrt{n}(\bar{X}_n - \mu) \approx \mathcal{N}(0, \sigma^2)$$
$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{n} \approx \mathcal{N}(0, 1)$$

Continuity correction

$$\mathbb{P}\left[\bar{X}_n \leq x\right] \approx \Phi\left(\frac{x + \frac{1}{2} - \mu}{\sigma/\sqrt{n}}\right)$$

$$\mathbb{P}\left[\bar{X}_n \geq x\right] \approx 1 - \Phi\left(\frac{x - \frac{1}{2} - \mu}{\sigma/\sqrt{n}}\right)$$

Delta method

$$Y_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right) \implies \varphi(Y_n) \approx \mathcal{N}\left(\varphi(\mu), (\varphi'(\mu))^2 \frac{\sigma^2}{n}\right)$$

11 Statistical Inference

Let $X_1, \cdots, X_n \overset{iid}{\sim} F$ if not otherwise noted.

11.1 Point Estimation

- Point estimator $\widehat{\theta}_n$ of θ is a RV: $\widehat{\theta}_n = g(X_1, \ldots, X_n)$
- $\text{bias}(\widehat{\theta}_n) = \mathbb{E}\left[\widehat{\theta}_n\right] - \theta$
- Consistency: $\widehat{\theta}_n \overset{\text{P}}{\to} \theta$
- Sampling distribution: $F(\widehat{\theta}_n)$
- Standard error: $\text{se}(\widehat{\theta}_n) = \sqrt{\mathbb{V}\left[\widehat{\theta}_n\right]}$
- Mean squared error: $\text{MSE} = \mathbb{E}\left[(\widehat{\theta}_n - \theta)^2\right] = \text{bias}(\widehat{\theta}_n)^2 + \mathbb{V}\left[\widehat{\theta}_n\right]$
- $\lim_{n\to\infty} \text{bias}(\widehat{\theta}_n) = 0 \wedge \lim_{n\to\infty} \text{se}(\widehat{\theta}_n) = 0 \implies \widehat{\theta}_n$ is consistent
- Asymptotic normality: $\dfrac{\widehat{\theta}_n - \theta}{\text{se}} \overset{\text{D}}{\to} \mathcal{N}(0, 1)$
- SLUTZKY'S THEOREM often lets us replace $\text{se}(\widehat{\theta}_n)$ by some (weakly) consistent estimator $\widehat{\sigma}_n$.

11.2 Normal-Based Confidence Interval

Suppose $\widehat{\theta}_n \approx \mathcal{N}\left(\theta, \widehat{\mathsf{se}}^2\right)$. Let $z_{\alpha/2} = \Phi^{-1}(1 - (\alpha/2))$, i.e., $\mathbb{P}\left[Z > z_{\alpha/2}\right] = \alpha/2$ and $\mathbb{P}\left[-z_{\alpha/2} < Z < z_{\alpha/2}\right] = 1 - \alpha$ where $Z \sim \mathcal{N}(0,1)$. Then

$$C_n = \widehat{\theta}_n \pm z_{\alpha/2}\widehat{\mathsf{se}}$$

11.3 Empirical distribution

Empirical Distribution Function (ECDF)

$$\widehat{F}_n(x) = \frac{\sum_{i=1}^n I(X_i \leq x)}{n}$$

$$I(X_i \leq x) = \begin{cases} 1 & X_i \leq x \\ 0 & X_i > x \end{cases}$$

Properties (for any fixed x)

- $\mathbb{E}\left[\widehat{F}_n\right] = F(x)$
- $\mathbb{V}\left[\widehat{F}_n\right] = \dfrac{F(x)(1 - F(x))}{n}$
- $\text{MSE} = \dfrac{F(x)(1 - F(x))}{n} \xrightarrow{\text{D}} 0$
- $\widehat{F}_n \xrightarrow{\text{P}} F(x)$

Dvoretzky-Kiefer-Wolfowitz (DKW) inequality $(X_1, \ldots, X_n \sim F)$

$$\mathbb{P}\left[\sup_x \left|F(x) - \widehat{F}_n(x)\right| > \varepsilon\right] = 2e^{-2n\varepsilon^2}$$

Nonparametric $1 - \alpha$ confidence band for F

$$L(x) = \max\{\widehat{F}_n - \epsilon_n, 0\}$$
$$U(x) = \min\{\widehat{F}_n + \epsilon_n, 1\}$$
$$\epsilon = \sqrt{\frac{1}{2n}\log\left(\frac{2}{\alpha}\right)}$$

$$\mathbb{P}\left[L(x) \leq F(x) \leq U(x) \,\forall x\right] \geq 1 - \alpha$$

11.4 Statistical Functionals

- Statistical functional: $T(F)$
- Plug-in estimator of $\theta = (F)$: $\widehat{\theta}_n = T(\widehat{F}_n)$
- Linear functional: $T(F) = \int \varphi(x)\, dF_X(x)$
- Plug-in estimator for linear functional:

$$T(\widehat{F}_n) = \int \varphi(x)\, d\widehat{F}_n(x) = \frac{1}{n}\sum_{i=1}^n \varphi(X_i)$$

- Often: $T(\widehat{F}_n) \approx \mathcal{N}\left(T(F), \widehat{\mathsf{se}}^2\right) \implies T(\widehat{F}_n) \pm z_{\alpha/2}\widehat{\mathsf{se}}$
- p^{th} quantile: $F^{-1}(p) = \inf\{x : F(x) \geq p\}$
- $\widehat{\mu} = \bar{X}_n$
- $\widehat{\sigma}^2 = \dfrac{1}{n-1}\sum_{i=1}^n (X_i - \bar{X}_n)^2$
- $\widehat{\kappa} = \dfrac{\frac{1}{n}\sum_{i=1}^n (X_i - \widehat{\mu})^3}{\widehat{\sigma}^3 j}$
- $\widehat{\rho} = \dfrac{\sum_{i=1}^n (X_i - \bar{X}_n)(Y_i - \bar{Y}_n)}{\sqrt{\sum_{i=1}^n (X_i - \bar{X}_n)^2}\sqrt{\sum_{i=1}^n (Y_i - \bar{Y}_n)^2}}$

12 Parametric Inference

Let $\mathfrak{F} = \left\{f(x; \theta) : \theta \in \Theta\right\}$ be a parametric model with parameter space $\Theta \subset \mathbb{R}^k$ and parameter $\theta = (\theta_1, \ldots, \theta_k)$.

12.1 Method of Moments

j^{th} moment

$$\alpha_j(\theta) = \mathbb{E}\left[X^j\right] = \int x^j\, dF_X(x)$$

j^{th} sample moment

$$\widehat{\alpha}_j = \frac{1}{n}\sum_{i=1}^n X_i^j$$

Method of moments estimator (MoM)

$$\alpha_1(\theta) = \widehat{\alpha}_1$$
$$\alpha_2(\theta) = \widehat{\alpha}_2$$
$$\vdots = \vdots$$
$$\alpha_k(\theta) = \widehat{\alpha}_k$$

Properties of the MoM estimator

- $\widehat{\theta}_n$ exists with probability tending to 1
- Consistency: $\widehat{\theta}_n \xrightarrow{\text{P}} \theta$
- Asymptotic normality:

$$\sqrt{n}(\widehat{\theta} - \theta) \xrightarrow{\text{D}} \mathcal{N}(0, \Sigma)$$

where $\Sigma = g\mathbb{E}\left[YY^T\right]g^T$, $Y = (X, X^2, \ldots, X^k)^T$, $g = (g_1, \ldots, g_k)$ and $g_j = \frac{\partial}{\partial\theta}\alpha_j^{-1}(\theta)$

12.2 Maximum Likelihood

Likelihood: $\mathcal{L}_n : \Theta \to [0, \infty)$

$$\mathcal{L}_n(\theta) = \prod_{i=1}^n f(X_i; \theta)$$

Log-likelihood

$$\ell_n(\theta) = \log \mathcal{L}_n(\theta) = \sum_{i=1}^n \log f(X_i; \theta)$$

Maximum likelihood estimator (MLE)

$$\mathcal{L}_n(\widehat{\theta}_n) = \sup_\theta \mathcal{L}_n(\theta)$$

Score function

$$s(X; \theta) = \frac{\partial}{\partial\theta}\log f(X; \theta)$$

Fisher information

$$I(\theta) = \mathbb{V}_\theta\left[s(X; \theta)\right]$$

$$I_n(\theta) = nI(\theta)$$

Fisher information (exponential family)

$$I(\theta) = \mathbb{E}_\theta\left[-\frac{\partial}{\partial\theta}s(X; \theta)\right]$$

Observed Fisher information

$$I_n^{obs}(\theta) = -\frac{\partial^2}{\partial\theta^2}\sum_{i=1}^n \log f(X_i; \theta)$$

Properties of the MLE

- Consistency: $\widehat{\theta}_n \xrightarrow{\text{P}} \theta$

- Equivariance: $\widehat{\theta}_n$ is the MLE $\implies \varphi(\widehat{\theta}_n)$ ist the MLE of $\varphi(\theta)$
- Asymptotic normality:

  1. $\mathsf{se} \approx \sqrt{1/I_n(\theta)}$

$$\frac{(\widehat{\theta}_n - \theta)}{\mathsf{se}} \xrightarrow{\text{D}} \mathcal{N}(0, 1)$$

  2. $\widehat{\mathsf{se}} \approx \sqrt{1/I_n(\widehat{\theta}_n)}$

$$\frac{(\widehat{\theta}_n - \theta)}{\widehat{\mathsf{se}}} \xrightarrow{\text{D}} \mathcal{N}(0, 1)$$

- Asymptotic optimality (or efficiency), i.e., smallest variance for large samples. If $\widetilde{\theta}_n$ is any other estimator, the asymptotic relative efficiency is

$$\text{ARE}(\widetilde{\theta}_n, \widehat{\theta}_n) = \frac{\mathbb{V}\left[\widehat{\theta}_n\right]}{\mathbb{V}\left[\widetilde{\theta}_n\right]} \leq 1$$

- Approximately the Bayes estimator

12.2.1 Delta Method

If $\tau = \varphi(\widehat{\theta})$ where φ is differentiable and $\varphi'(\theta) \neq 0$:

$$\frac{(\widehat{\tau}_n - \tau)}{\widehat{\mathsf{se}}(\widehat{\tau})} \xrightarrow{\text{D}} \mathcal{N}(0, 1)$$

where $\widehat{\tau} = \varphi(\widehat{\theta})$ is the MLE of τ and

$$\widehat{\mathsf{se}} = \left|\varphi'(\widehat{\theta})\right|\widehat{\mathsf{se}}(\widehat{\theta}_n)$$

12.3 Multiparameter Models

Let $\theta = (\theta_1, \ldots, \theta_k)$ and $\widehat{\theta} = (\widehat{\theta}_1, \ldots, \widehat{\theta}_k)$ be the MLE.

$$H_{jj} = \frac{\partial^2 \ell_n}{\partial\theta^2} \qquad H_{jk} = \frac{\partial^2 \ell_n}{\partial\theta_j \partial\theta_k}$$

Fisher information matrix

$$I_n(\theta) = -\begin{bmatrix} \mathbb{E}_\theta\left[H_{11}\right] & \cdots & \mathbb{E}_\theta\left[H_{1k}\right] \\ \vdots & \ddots & \vdots \\ \mathbb{E}_\theta\left[H_{k1}\right] & \cdots & \mathbb{E}_\theta\left[H_{kk}\right] \end{bmatrix}$$

Under appropriate regularity conditions

$$(\widehat{\theta} - \theta) \approx \mathcal{N}(0, J_n)$$

with $J_n(\theta) = I_n^{-1}$. Further, if $\widehat{\theta}_j$ is the jth component of θ, then

$$\frac{(\widehat{\theta}_j - \theta_j)}{\widehat{\mathsf{se}}_j} \xrightarrow{\text{D}} \mathcal{N}(0,1)$$

where $\widehat{\mathsf{se}}_j^2 = J_n(j,j)$ and $\text{Cov}\left[\widehat{\theta}_j, \widehat{\theta}_k\right] = J_n(j,k)$

12.3.1 Multiparameter delta method

Let $\tau = \varphi(\theta_1, \ldots, \theta_k)$ and let the gradient of φ be

$$\nabla\varphi = \begin{pmatrix} \dfrac{\partial\varphi}{\partial\theta_1} \\ \vdots \\ \dfrac{\partial\varphi}{\partial\theta_k} \end{pmatrix}$$

Suppose $\nabla\varphi\big|_{\theta=\widehat{\theta}} \neq 0$ and $\widehat{\tau} = \varphi(\widehat{\theta})$. Then,

$$\frac{(\widehat{\tau} - \tau)}{\widehat{\mathsf{se}}(\widehat{\tau})} \xrightarrow{\text{D}} \mathcal{N}(0,1)$$

where

$$\widehat{\mathsf{se}}(\widehat{\tau}) = \sqrt{\left(\widehat{\nabla\varphi}\right)^T \widehat{J}_n \left(\widehat{\nabla\varphi}\right)}$$

and $\widehat{J}_n = J_n(\widehat{\theta})$ and $\widehat{\nabla\varphi} = \nabla\varphi\big|_{\theta=\widehat{\theta}}$.

12.4 Parametric Bootstrap

Sample from $f(x;\widehat{\theta}_n)$ instead of from $\widehat{F}_n$, where $\widehat{\theta}_n$ could be the MLE or method of moments estimator.

13 Hypothesis Testing

$$H_0 : \theta \in \Theta_0 \qquad \text{versus} \qquad H_1 : \theta \in \Theta_1$$

Definitions

- Null hypothesis H_0
- Alternative hypothesis H_1
- Simple hypothesis $\theta = \theta_0$
- Composite hypothesis $\theta > \theta_0$ or $\theta < \theta_0$
- Two-sided test: $H_0 : \theta = \theta_0 \quad$ versus $\quad H_1 : \theta \neq \theta_0$
- One-sided test: $H_0 : \theta \leq \theta_0 \quad$ versus $\quad H_1 : \theta > \theta_0$
- Critical value c
- Test statistic T
- Rejection region $R = \{x : T(x) > c\}$
- Power function $\beta(\theta) = \mathbb{P}[X \in R]$
- Power of a test: $1 - \mathbb{P}[\text{Type II error}] = 1 - \beta = \inf_{\theta \in \Theta_1} \beta(\theta)$
- Test size: $\alpha = \mathbb{P}[\text{Type I error}] = \sup_{\theta \in \Theta_0} \beta(\theta)$

	Retain H_0	Reject H_0
H_0 true	$\checkmark$	Type I Error (α)
H_1 true	Type II Error (β)	$\checkmark$ (power)

p-value

- p-value $= \sup_{\theta \in \Theta_0} \mathbb{P}_\theta[T(X) \geq T(x)] = \inf\{\alpha : T(x) \in R_\alpha\}$
- p-value $= \sup_{\theta \in \Theta_0} \underbrace{\mathbb{P}_\theta[T(X^\star) \geq T(X)]}_{\substack{1 - F_\theta(T(X)) \quad \text{since } T(X^\star) \sim F_\theta}} = \inf\{\alpha : T(X) \in R_\alpha\}$

p-value	evidence
< 0.01	very strong evidence against H_0
$0.01 - 0.05$	strong evidence against H_0
$0.05 - 0.1$	weak evidence against H_0
> 0.1	little or no evidence against H_0

Wald test

- Two-sided test
- Reject H_0 when $|W| > z_{\alpha/2}$ where $W = \dfrac{\widehat{\theta} - \theta_0}{\widehat{\mathsf{se}}}$
- $\mathbb{P}\left[|W| > z_{\alpha/2}\right] \to \alpha$
- p-value $= \mathbb{P}_{\theta_0}[|W| > |w|] \approx \mathbb{P}[|Z| > |w|] = 2\Phi(-|w|)$

Likelihood ratio test (LRT)

- $T(X) = \dfrac{\sup_{\theta \in \Theta} \mathcal{L}_n(\theta)}{\sup_{\theta \in \Theta_0} \mathcal{L}_n(\theta)} = \dfrac{\mathcal{L}_n(\widehat{\theta}_n)}{\mathcal{L}_n(\widehat{\theta}_{n,0})}$

- $\lambda(X) = 2 \log T(X) \xrightarrow{\mathrm{D}} \chi^2_{r-q}$ where $\sum\limits_{i=1}^{k} Z_i^2 \sim \chi_k^2$ and $Z_1, \ldots, Z_k \overset{iid}{\sim} \mathcal{N}(0,1)$

- p-value $= \mathbb{P}_{\theta_0} [\lambda(X) > \lambda(x)] \approx \mathbb{P}\left[\chi^2_{r-q} > \lambda(x)\right]$

Multinomial LRT

- MLE: $\widehat{p}_n = \left(\dfrac{X_1}{n}, \ldots, \dfrac{X_k}{n}\right)$

- $T(X) = \dfrac{\mathcal{L}_n(\widehat{p}_n)}{\mathcal{L}_n(p_0)} = \prod\limits_{j=1}^{k} \left(\dfrac{\widehat{p}_j}{p_{0j}}\right)^{X_j}$

- $\lambda(X) = 2 \sum\limits_{j=1}^{k} X_j \log \left(\dfrac{\widehat{p}_j}{p_{0j}}\right) \xrightarrow{\mathrm{D}} \chi^2_{k-1}$

- The approximate size α LRT rejects H_0 when $\lambda(X) \geq \chi^2_{k-1,\alpha}$

Pearson Chi-square Test

- $T = \sum\limits_{j=1}^{k} \dfrac{(X_j - \mathbb{E}[X_j])^2}{\mathbb{E}[X_j]}$ where $\mathbb{E}[X_j] = np_{0j}$ under H_0

- $T \xrightarrow{\mathrm{D}} \chi^2_{k-1}$
- p-value $= \mathbb{P}\left[\chi^2_{k-1} > T(x)\right]$
- Faster $\xrightarrow{\mathrm{D}} X^2_{k-1}$ than LRT, hence preferable for small n

Independence testing

- I rows, J columns, $\mathbf{X}$ multinomial sample of size $n = I * J$
- MLEs unconstrained: $\widehat{p}_{ij} = \frac{X_{ij}}{n}$
- MLEs under H_0: $\widehat{p}_{0ij} = \widehat{p}_{i\cdot}\widehat{p}_{\cdot j} = \frac{X_{i\cdot}}{n}\frac{X_{\cdot j}}{n}$
- LRT: $\lambda = 2 \sum_{i=1}^{I} \sum_{j=1}^{J} X_{ij} \log \left(\frac{nX_{ij}}{X_{i\cdot}X_{\cdot j}}\right)$
- PearsonChiSq: $T = \sum_{i=1}^{I} \sum_{j=1}^{J} \frac{(X_{ij} - \mathbb{E}[X_{ij}])^2}{\mathbb{E}[X_{ij}]}$
- LRT and Pearson $\xrightarrow{\mathrm{D}} \chi_k^2 \nu$, where $\nu = (I-1)(J-1)$

14 Bayesian Inference

$$f(\theta \mid x) = \frac{f(x \mid \theta)f(\theta)}{f(x^n)} = \frac{f(x \mid \theta)f(\theta)}{\int f(x \mid \theta)f(\theta)\, d\theta} \propto \mathcal{L}_n(\theta)f(\theta)$$

Definitions

- $X^n = (X_1, \ldots, X_n)$

- $x^n = (x_1, \ldots, x_n)$
- Prior density $f(\theta)$
- Likelihood $f(x^n \mid \theta)$: joint density of the data

  In particular, X^n IID $\implies f(x^n \mid \theta) = \prod\limits_{i=1}^{n} f(x_i \mid \theta) = \mathcal{L}_n(\theta)$

- Posterior density $f(\theta \mid x^n)$
- Normalizing constant $c_n = f(x^n) = \int f(x \mid \theta)f(\theta)\, d\theta$
- Kernel: part of a density that depends on θ
- Posterior mean $\bar{\theta}_n = \int \theta f(\theta \mid x^n)\, d\theta = \frac{\int \theta \mathcal{L}_n(\theta)f(\theta)}{\int \mathcal{L}_n(\theta)f(\theta)\, d\theta}$

14.1 Credible Intervals

Posterior interval

$$\mathbb{P}[\theta \in (a,b) \mid x^n] = \int_a^b f(\theta \mid x^n)\, d\theta = 1 - \alpha$$

Equal-tail credible interval

$$\int_{-\infty}^{a} f(\theta \mid x^n)\, d\theta = \int_b^{\infty} f(\theta \mid x^n)\, d\theta = \alpha/2$$

Highest posterior density (HPD) region R_n

1. $\mathbb{P}[\theta \in R_n] = 1 - \alpha$
2. $R_n = \{\theta : f(\theta \mid x^n) > k\}$ for some k

R_n is unimodal $\implies R_n$ is an interval

14.2 Function of parameters

Let $\tau = \varphi(\theta)$ and $A = \{\theta : \varphi(\theta) \leq \tau\}$.
Posterior CDF for τ

$$H(r \mid x^n) = \mathbb{P}[\varphi(\theta) \leq \tau \mid x^n] = \int_A f(\theta \mid x^n)\, d\theta$$

Posterior density

$$h(\tau \mid x^n) = H'(\tau \mid x^n)$$

Bayesian delta method

$$\tau \mid X^n \approx \mathcal{N}\left(\varphi(\widehat{\theta}), \widehat{\mathsf{se}}\left|\varphi'(\widehat{\theta})\right|\right)$$

14.3   Priors

Choice

- Subjective bayesianism.
- Objective bayesianism.
- Robust bayesianism.

Types

- Flat: $f(\theta) \propto constant$
- Proper: $\int_{-\infty}^{\infty} f(\theta)\, d\theta = 1$
- Improper: $\int_{-\infty}^{\infty} f(\theta)\, d\theta = \infty$
- JEFFREY's prior (transformation-invariant):

$$f(\theta) \propto \sqrt{I(\theta)} \qquad f(\theta) \propto \sqrt{\det(I(\theta))}$$

- Conjugate: $f(\theta)$ and $f(\theta \,|\, x^n)$ belong to the same parametric family

14.3.1   Conjugate Priors

Discrete likelihood		
Likelihood	Conjugate prior	Posterior hyperparameters
Bern (p)	Beta (α, β)	$\alpha + \sum_{i=1}^{n} x_i, \beta + n - \sum_{i=1}^{n} x_i$
Bin (p)	Beta (α, β)	$\alpha + \sum_{i=1}^{n} x_i, \beta + \sum_{i=1}^{n} N_i - \sum_{i=1}^{n} x_i$
NBin (p)	Beta (α, β)	$\alpha + rn, \beta + \sum_{i=1}^{n} x_i$
Po (λ)	Gamma (α, β)	$\alpha + \sum_{i=1}^{n} x_i, \beta + n$
Multinomial(p)	Dir (α)	$\alpha + \sum_{i=1}^{n} x^{(i)}$
Geo (p)	Beta (α, β)	$\alpha + n, \beta + \sum_{i=1}^{n} x_i$

Continuous likelihood (subscript c denotes constant)		
Likelihood	Conjugate prior	Posterior hyperparameters
Unif $(0, \theta)$	Pareto(x_m, k)	$\max\left\{x_{(n)}, x_m\right\}, k + n$
Exp (λ)	Gamma (α, β)	$\alpha + n, \beta + \sum_{i=1}^{n} x_i$
$\mathcal{N}\left(\mu, \sigma_c^2\right)$	$\mathcal{N}\left(\mu_0, \sigma_0^2\right)$	$\left(\dfrac{\mu_0}{\sigma_0^2} + \dfrac{\sum_{i=1}^{n} x_i}{\sigma_c^2}\right) \Big/ \left(\dfrac{1}{\sigma_0^2} + \dfrac{n}{\sigma_c^2}\right), \left(\dfrac{1}{\sigma_0^2} + \dfrac{n}{\sigma_c^2}\right)^{-1}$
$\mathcal{N}\left(\mu_c, \sigma^2\right)$	Scaled Inverse Chi-square(ν, σ_0^2)	$\nu + n, \dfrac{\nu \sigma_0^2 + \sum_{i=1}^{n}(x_i - \mu)^2}{\nu + n}$
$\mathcal{N}\left(\mu, \sigma^2\right)$	Normal-scaled Inverse Gamma$(\lambda, \nu, \alpha, \beta)$	$\dfrac{\nu\lambda + n\bar{x}}{\nu + n}, \qquad \nu + n, \qquad \alpha + \dfrac{n}{2}, \quad \beta + \dfrac{1}{2}\sum_{i=1}^{n}(x_i - \bar{x})^2 + \dfrac{\gamma(\bar{x} - \lambda)^2}{2(n + \gamma)}$
MVN(μ, Σ_c)	MVN(μ_0, Σ_0)	$\left(\Sigma_0^{-1} + n\Sigma_c^{-1}\right)^{-1}\left(\Sigma_0^{-1}\mu_0 + n\Sigma^{-1}\bar{x}\right), \left(\Sigma_0^{-1} + n\Sigma_c^{-1}\right)^{-1}$
MVN(μ_c, Σ)	Inverse-Wishart(κ, Ψ)	$n + \kappa, \Psi + \sum_{i=1}^{n}(x_i - \mu_c)(x_i - \mu_c)^T$
Pareto(x_{m_c}, k)	Gamma (α, β)	$\alpha + n, \beta + \sum_{i=1}^{n}\log\dfrac{x_i}{x_{m_c}}$
Pareto(x_m, k_c)	Pareto(x_0, k_0)	$x_0, k_0 - kn$ where $k_0 > kn$
Gamma (α_c, β)	Gamma (α_0, β_0)	$\alpha_0 + n\alpha_c, \beta_0 + \sum_{i=1}^{n} x_i$

14.4   Bayesian Testing

If $H_0 : \theta \in \Theta_0$:

$$\text{Prior probability } \mathbb{P}\left[H_0\right] = \int_{\Theta_0} f(\theta)\, d\theta$$

$$\text{Posterior probability } \mathbb{P}\left[H_0 \,|\, x^n\right] = \int_{\Theta_0} f(\theta \,|\, x^n)\, d\theta$$

Let $H_0, \ldots, H_{K-1}$ be K hypotheses. Suppose $\theta \sim f(\theta \,|\, H_k)$,

$$\mathbb{P}\left[H_k \,|\, x^n\right] = \frac{f(x^n \,|\, H_k)\mathbb{P}\left[H_k\right]}{\sum_{k=1}^{K} f(x^n \,|\, H_k)\mathbb{P}\left[H_k\right]},$$

Marginal likelihood

$$f(x^n \mid H_i) = \int_\Theta f(x^n \mid \theta, H_i) f(\theta \mid H_i) \, d\theta$$

Posterior odds (of H_i relative to H_j)

$$\frac{\mathbb{P}[H_i \mid x^n]}{\mathbb{P}[H_j \mid x^n]} = \underbrace{\frac{f(x^n \mid H_i)}{f(x^n \mid H_j)}}_{\text{Bayes Factor } BF_{ij}} \times \underbrace{\frac{\mathbb{P}[H_i]}{\mathbb{P}[H_j]}}_{\text{prior odds}}$$

Bayes factor

$\log_{10} BF_{10}$	BF_{10}	evidence
$0 - 0.5$	$1 - 1.5$	Weak
$0.5 - 1$	$1.5 - 10$	Moderate
$1 - 2$	$10 - 100$	Strong
> 2	> 100	Decisive

$$p^* = \frac{\frac{p}{1-p} BF_{10}}{1 + \frac{p}{1-p} BF_{10}} \quad \text{where } p = \mathbb{P}[H_1] \text{ and } p^* = \mathbb{P}[H_1 \mid x^n]$$

15   Exponential Family

Scalar parameter

$$\begin{aligned}
f_X(x \mid \theta) &= h(x) \exp\{\eta(\theta) T(x) - A(\theta)\} \\
&= h(x) g(\theta) \exp\{\eta(\theta) T(x)\}
\end{aligned}$$

Vector parameter

$$\begin{aligned}
f_X(x \mid \theta) &= h(x) \exp\left\{\sum_{i=1}^s \eta_i(\theta) T_i(x) - A(\theta)\right\} \\
&= h(x) \exp\{\eta(\theta) \cdot T(x) - A(\theta)\} \\
&= h(x) g(\theta) \exp\{\eta(\theta) \cdot T(x)\}
\end{aligned}$$

Natural form

$$\begin{aligned}
f_X(x \mid \eta) &= h(x) \exp\{\eta \cdot \mathbf{T}(x) - A(\eta)\} \\
&= h(x) g(\eta) \exp\{\eta \cdot \mathbf{T}(x)\} \\
&= h(x) g(\eta) \exp\{\eta^T \mathbf{T}(x)\}
\end{aligned}$$

16   Sampling Methods

16.1   The Bootstrap

Let $T_n = g(X_1, \ldots, X_n)$ be a statistic.

1. Estimate $\mathbb{V}_F[T_n]$ with $\mathbb{V}_{\widehat{F}_n}[T_n]$.
2. Approximate $\mathbb{V}_{\widehat{F}_n}[T_n]$ using simulation:

   (a) Repeat the following B times to get $T_{n,1}^*, \ldots, T_{n,B}^*$, an IID sample from the sampling distribution implied by $\widehat{F}_n$

        i. Sample uniformly $X_1^*, \ldots, X_n^* \sim \widehat{F}_n$.
        ii. Compute $T_n^* = g(X_1^*, \ldots, X_n^*)$.

   (b) Then

$$v_{boot} = \widehat{\mathbb{V}}_{\widehat{F}_n} = \frac{1}{B} \sum_{b=1}^B \left(T_{n,b}^* - \frac{1}{B} \sum_{r=1}^B T_{n,r}^* \right)^2$$

16.1.1   Bootstrap Confidence Intervals

Normal-based interval

$$T_n \pm z_{\alpha/2} \widehat{se}_{boot}$$

Pivotal interval

1. Location parameter $\theta = T(F)$
2. Pivot $R_n = \widehat{\theta}_n - \theta$
3. Let $H(r) = \mathbb{P}[R_n \leq r]$ be the CDF of R_n
4. Let $R_{n,b}^* = \widehat{\theta}_{n,b}^* - \widehat{\theta}_n$. Approximate H using bootstrap:

$$\widehat{H}(r) = \frac{1}{B} \sum_{b=1}^B I(R_{n,b}^* \leq r)$$

5. $\theta_\beta^* = \beta$ sample quantile of $(\widehat{\theta}_{n,1}^*, \ldots, \widehat{\theta}_{n,B}^*)$
6. $r_\beta^* = \beta$ sample quantile of $(R_{n,1}^*, \ldots, R_{n,B}^*)$, i.e., $r_\beta^* = \theta_\beta^* - \widehat{\theta}_n$
7. Approximate $1 - \alpha$ confidence interval $C_n = \left(\hat{a}, \hat{b}\right)$ where

$$\begin{aligned}
\hat{a} &= \widehat{\theta}_n - \widehat{H}^{-1}\left(1 - \frac{\alpha}{2}\right) = \widehat{\theta}_n - r_{1-\alpha/2}^* = 2\widehat{\theta}_n - \theta_{1-\alpha/2}^* \\
\hat{b} &= \widehat{\theta}_n - \widehat{H}^{-1}\left(\frac{\alpha}{2}\right) = \widehat{\theta}_n - r_{\alpha/2}^* = 2\widehat{\theta}_n - \theta_{\alpha/2}^*
\end{aligned}$$

Percentile interval

$$C_n = \left(\theta_{\alpha/2}^*, \theta_{1-\alpha/2}^*\right)$$

16.2 Rejection Sampling

Setup

- We can easily sample from $g(\theta)$
- We want to sample from $h(\theta)$, but it is difficult
- We know $h(\theta)$ up to a proportional constant: $h(\theta) = \dfrac{k(\theta)}{\int k(\theta)\,d\theta}$
- Envelope condition: we can find $M > 0$ such that $k(\theta) \le Mg(\theta) \quad \forall \theta$

Algorithm

1. Draw $\theta^{cand} \sim g(\theta)$
2. Generate $u \sim \text{Unif}(0,1)$
3. Accept θ^{cand} if $u \le \dfrac{k(\theta^{cand})}{Mg(\theta^{cand})}$
4. Repeat until B values of θ^{cand} have been accepted

Example

- We can easily sample from the prior $g(\theta) = f(\theta)$
- Target is the posterior $h(\theta) \propto k(\theta) = f(x^n \mid \theta)f(\theta)$
- Envelope condition: $f(x^n \mid \theta) \le f(x^n \mid \widehat{\theta}_n) = \mathcal{L}_n(\widehat{\theta}_n) \equiv M$
- Algorithm
    1. Draw $\theta^{cand} \sim f(\theta)$
    2. Generate $u \sim \text{Unif}(0,1)$
    3. Accept θ^{cand} if $u \le \dfrac{\mathcal{L}_n(\theta^{cand})}{\mathcal{L}_n(\widehat{\theta}_n)}$

16.3 Importance Sampling

Sample from an importance function g rather than target density h.
Algorithm to obtain an approximation to $\mathbb{E}\left[q(\theta) \mid x^n\right]$:

1. Sample from the prior $\theta_1, \ldots, \theta_n \overset{iid}{\sim} f(\theta)$
2. $w_i = \dfrac{\mathcal{L}_n(\theta_i)}{\sum_{i=1}^{B} \mathcal{L}_n(\theta_i)} \quad \forall i = 1, \ldots, B$
3. $\mathbb{E}\left[q(\theta) \mid x^n\right] \approx \sum_{i=1}^{B} q(\theta_i)w_i$

17 Decision Theory

Definitions

- Unknown quantity affecting our decision: $\theta \in \Theta$

- Decision rule: synonymous for an estimator $\widehat{\theta}$
- Action $a \in \mathcal{A}$: possible value of the decision rule. In the estimation context, the action is just an estimate of θ, $\widehat{\theta}(x)$.
- Loss function L: consequences of taking action a when true state is θ or discrepancy between θ and $\widehat{\theta}$, $L : \Theta \times \mathcal{A} \to [-k, \infty)$.

Loss functions

- Squared error loss: $L(\theta, a) = (\theta - a)^2$
- Linear loss: $L(\theta, a) = \begin{cases} K_1(\theta - a) & a - \theta < 0 \\ K_2(a - \theta) & a - \theta \ge 0 \end{cases}$
- Absolute error loss: $L(\theta, a) = |\theta - a| \quad$ (linear loss with $K_1 = K_2$)
- L_p loss: $L(\theta, a) = |\theta - a|^p$
- Zero-one loss: $L(\theta, a) = \begin{cases} 0 & a = \theta \\ 1 & a \ne \theta \end{cases}$

17.1 Risk

Posterior risk

$$r(\widehat{\theta} \mid x) = \int L(\theta, \widehat{\theta}(x))f(\theta \mid x)\,d\theta = \mathbb{E}_{\theta \mid X}\left[L(\theta, \widehat{\theta}(x))\right]$$

(Frequentist) risk

$$R(\theta, \widehat{\theta}) = \int L(\theta, \widehat{\theta}(x))f(x \mid \theta)\,dx = \mathbb{E}_{X \mid \theta}\left[L(\theta, \widehat{\theta}(X))\right]$$

Bayes risk

$$r(f, \widehat{\theta}) = \iint L(\theta, \widehat{\theta}(x))f(x, \theta)\,dx\,d\theta = \mathbb{E}_{\theta, X}\left[L(\theta, \widehat{\theta}(X))\right]$$

$$r(f, \widehat{\theta}) = \mathbb{E}_{\theta}\left[\mathbb{E}_{X \mid \theta}\left[L(\theta, \widehat{\theta}(X))\right]\right] = \mathbb{E}_{\theta}\left[R(\theta, \widehat{\theta})\right]$$

$$r(f, \widehat{\theta}) = \mathbb{E}_{X}\left[\mathbb{E}_{\theta \mid X}\left[L(\theta, \widehat{\theta}(X))\right]\right] = \mathbb{E}_{X}\left[r(\widehat{\theta} \mid X)\right]$$

17.2 Admissibility

- $\widehat{\theta}'$ dominates $\widehat{\theta}$ if
$$\forall \theta : R(\theta, \widehat{\theta}') \le R(\theta, \widehat{\theta})$$
$$\exists \theta : R(\theta, \widehat{\theta}') < R(\theta, \widehat{\theta})$$

- $\widehat{\theta}$ is inadmissible if there is at least one other estimator $\widehat{\theta}'$ that dominates it. Otherwise it is called admissible.

17.3 Bayes Rule

Bayes rule (or Bayes estimator)

- $r(f, \widehat{\theta}) = \inf_{\widetilde{\theta}} r(f, \widetilde{\theta})$
- $\widehat{\theta}(x) = \inf r(\widehat{\theta} \,|\, x) \; \forall x \implies r(f, \widehat{\theta}) = \int r(\widehat{\theta} \,|\, x) f(x) \, dx$

Theorems

- Squared error loss: posterior mean
- Absolute error loss: posterior median
- Zero-one loss: posterior mode

17.4 Minimax Rules

Maximum risk

$$\bar{R}(\widehat{\theta}) = \sup_{\theta} R(\theta, \widehat{\theta}) \qquad \bar{R}(a) = \sup_{\theta} R(\theta, a)$$

Minimax rule

$$\sup_{\theta} R(\theta, \widehat{\theta}) = \inf_{\widetilde{\theta}} \bar{R}(\widetilde{\theta}) = \inf_{\widetilde{\theta}} \sup_{\theta} R(\theta, \widetilde{\theta})$$

$$\widehat{\theta} = \text{Bayes rule} \; \wedge \; \exists c : R(\theta, \widehat{\theta}) = c$$

Least favorable prior

$$\widehat{\theta}^f = \text{Bayes rule} \; \wedge \; R(\theta, \widehat{\theta}^f) \le r(f, \widehat{\theta}^f) \; \forall \theta$$

18 Linear Regression

Definitions

- Response variable Y
- Covariate X (aka predictor variable or feature)

18.1 Simple Linear Regression

Model

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i \qquad \mathbb{E}\left[\epsilon_i \,|\, X_i\right] = 0, \; \mathbb{V}\left[\epsilon_i \,|\, X_i\right] = \sigma^2$$

Fitted line

$$\widehat{r}(x) = \widehat{\beta}_0 + \widehat{\beta}_1 x$$

Predicted (fitted) values

$$\widehat{Y}_i = \widehat{r}(X_i)$$

Residuals

$$\widehat{\epsilon}_i = Y_i - \widehat{Y}_i = Y_i - \left(\widehat{\beta}_0 + \widehat{\beta}_1 X_i\right)$$

Residual sums of squares (RSS)

$$\text{RSS}(\widehat{\beta}_0, \widehat{\beta}_1) = \sum_{i=1}^{n} \widehat{\epsilon}_i^2$$

Least square estimates

$$\widehat{\beta}^T = (\widehat{\beta}_0, \widehat{\beta}_1)^T : \min_{\widehat{\beta}_0, \widehat{\beta}_1} \text{RSS}$$

$$\widehat{\beta}_0 = \bar{Y}_n - \widehat{\beta}_1 \bar{X}_n$$

$$\widehat{\beta}_1 = \frac{\sum_{i=1}^{n}(X_i - \bar{X}_n)(Y_i - \bar{Y}_n)}{\sum_{i=1}^{n}(X_i - \bar{X}_n)^2} = \frac{\sum_{i=1}^{n} X_i Y_i - n\bar{X}\bar{Y}}{\sum_{i=1}^{n} X_i^2 - n\overline{X^2}}$$

$$\mathbb{E}\left[\widehat{\beta} \,|\, X^n\right] = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}$$

$$\mathbb{V}\left[\widehat{\beta} \,|\, X^n\right] = \frac{\sigma^2}{n s_X} \begin{pmatrix} n^{-1} \sum_{i=1}^{n} X_i^2 & -\overline{X}_n \\ -\overline{X}_n & 1 \end{pmatrix}$$

$$\widehat{\text{se}}(\widehat{\beta}_0) = \frac{\widehat{\sigma}}{s_X \sqrt{n}} \sqrt{\frac{\sum_{i=1}^{n} X_i^2}{n}}$$

$$\widehat{\text{se}}(\widehat{\beta}_1) = \frac{\widehat{\sigma}}{s_X \sqrt{n}}$$

where $s_X^2 = n^{-1} \sum_{i=1}^{n}(X_i - \overline{X}_n)^2$ and $\widehat{\sigma}^2 = \frac{1}{n-2} \sum_{i=1}^{n} \widehat{\epsilon}_i^2$ (unbiased estimate). Further properties:

- Consistency: $\widehat{\beta}_0 \xrightarrow{\text{P}} \beta_0$ and $\widehat{\beta}_1 \xrightarrow{\text{P}} \beta_1$
- Asymptotic normality:

$$\frac{\widehat{\beta}_0 - \beta_0}{\widehat{\text{se}}(\widehat{\beta}_0)} \xrightarrow{\text{D}} \mathcal{N}(0, 1) \quad \text{and} \quad \frac{\widehat{\beta}_1 - \beta_1}{\widehat{\text{se}}(\widehat{\beta}_1)} \xrightarrow{\text{D}} \mathcal{N}(0, 1)$$

- Approximate $1 - \alpha$ confidence intervals for β_0 and β_1:

$$\widehat{\beta}_0 \pm z_{\alpha/2} \widehat{\text{se}}(\widehat{\beta}_0) \quad \text{and} \quad \widehat{\beta}_1 \pm z_{\alpha/2} \widehat{\text{se}}(\widehat{\beta}_1)$$

- Wald test for $H_0 : \beta_1 = 0$ vs. $H_1 : \beta_1 \ne 0$: reject H_0 if $|W| > z_{\alpha/2}$ where $W = \widehat{\beta}_1 / \widehat{\text{se}}(\widehat{\beta}_1)$.

R^2

$$R^2 = \frac{\sum_{i=1}^{n}(\widehat{Y}_i - \overline{Y})^2}{\sum_{i=1}^{n}(Y_i - \overline{Y})^2} = 1 - \frac{\sum_{i=1}^{n} \widehat{\epsilon}_i^2}{\sum_{i=1}^{n}(Y_i - \overline{Y})^2} = 1 - \frac{\text{RSS}}{\text{TSS}}$$

Likelihood

$$\mathcal{L} = \prod_{i=1}^{n} f(X_i, Y_i) = \prod_{i=1}^{n} f_X(X_i) \times \prod_{i=1}^{n} f_{Y|X}(Y_i \mid X_i) = \mathcal{L}_1 \times \mathcal{L}_2$$

$$\mathcal{L}_1 = \prod_{i=1}^{n} f_X(X_i)$$

$$\mathcal{L}_2 = \prod_{i=1}^{n} f_{Y|X}(Y_i \mid X_i) \propto \sigma^{-n} \exp\left\{ -\frac{1}{2\sigma^2} \sum_i \left(Y_i - (\beta_0 - \beta_1 X_i) \right)^2 \right\}$$

Under the assumption of Normality, the least squares estimator is also the MLE

$$\widehat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^{n} \widehat{\epsilon}_i^2$$

18.2 Prediction

Observe $X = x_*$ of the covariate and want to predict their outcome Y_*.

$$\widehat{Y}_* = \widehat{\beta}_0 + \widehat{\beta}_1 x_*$$

$$\mathbb{V}\left[\widehat{Y}_*\right] = \mathbb{V}\left[\widehat{\beta}_0\right] + x_*^2 \mathbb{V}\left[\widehat{\beta}_1\right] + 2x_* \mathrm{Cov}\left[\widehat{\beta}_0, \widehat{\beta}_1\right]$$

Prediction interval

$$\widehat{\xi}_n^2 = \widehat{\sigma}^2 \left(\frac{\sum_{i=1}^{n}(X_i - X_*)^2}{n \sum_i (X_i - \bar{X})^2 j} + 1 \right)$$

$$\widehat{Y}_* \pm z_{\alpha/2}\widehat{\xi}_n$$

18.3 Multiple Regression

$$Y = X\beta + \epsilon$$

where

$$X = \begin{pmatrix} X_{11} & \cdots & X_{1k} \\ \vdots & \ddots & \vdots \\ X_{n1} & \cdots & X_{nk} \end{pmatrix} \quad \beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_k \end{pmatrix} \quad \epsilon = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

Likelihood

$$\mathcal{L}(\mu, \Sigma) = (2\pi\sigma^2)^{-n/2} \exp\left\{ -\frac{1}{2\sigma^2} \mathrm{RSS} \right\}$$

$$\mathrm{RSS} = (y - X\beta)^T (y - X\beta) = \|Y - X\beta\|^2 = \sum_{i=1}^{N} (Y_i - x_i^T \beta)^2$$

If the $(k \times k)$ matrix $X^T X$ is invertible,

$$\widehat{\beta} = (X^T X)^{-1} X^T Y$$

$$\mathbb{V}\left[\widehat{\beta} \mid X^n\right] = \sigma^2 (X^T X)^{-1}$$

$$\widehat{\beta} \approx \mathcal{N}\left(\beta, \sigma^2 (X^T X)^{-1}\right)$$

Estimate regression function

$$\widehat{r}(x) = \sum_{j=1}^{k} \widehat{\beta}_j x_j$$

Unbiased estimate for σ^2

$$\widehat{\sigma}^2 = \frac{1}{n-k} \sum_{i=1}^{n} \widehat{\epsilon}_i^2 \qquad \widehat{\epsilon} = X\widehat{\beta} - Y$$

MLE

$$\widehat{\mu} = \bar{X} \qquad \widehat{\sigma}^2 = \frac{n-k}{n}\sigma^2$$

$1 - \alpha$ Confidence interval

$$\widehat{\beta}_j \pm z_{\alpha/2}\widehat{\mathrm{se}}(\widehat{\beta}_j)$$

18.4 Model Selection

Consider predicting a new observation Y^* for covariates X^* and let $S \subset J$ denote a subset of the covariates in the model, where $|S| = k$ and $|J| = n$. Issues

- Underfitting: too few covariates yields high bias
- Overfitting: too many covariates yields high variance

Procedure

1. Assign a score to each model
2. Search through all models to find the one with the highest score

Hypothesis testing

$$H_0 : \beta_j = 0 \text{ vs. } H_1 : \beta_j \neq 0 \quad \forall j \in J$$

Mean squared prediction error (MSPE)

$$\mathrm{MSPE} = \mathbb{E}\left[(\widehat{Y}(S) - Y^*)^2\right]$$

Prediction risk

$$R(S) = \sum_{i=1}^{n} \mathrm{MSPE}_i = \sum_{i=1}^{n} \mathbb{E}\left[(\widehat{Y}_i(S) - Y_i^*)^2\right]$$

Training error

$$\widehat{R}_{tr}(S) = \sum_{i=1}^{n} (\widehat{Y}_i(S) - Y_i)^2$$

R^2

$$R^2(S) = 1 - \frac{\text{RSS}(S)}{\text{TSS}} = 1 - \frac{\widehat{R}_{tr}(S)}{\text{TSS}} = 1 - \frac{\sum_{i=1}^{n}(\widehat{Y}_i(S) - \overline{Y})^2}{\sum_{i=1}^{n}(Y_i - \overline{Y})^2}$$

The training error is a downward-biased estimate of the prediction risk.

$$\mathbb{E}\left[\widehat{R}_{tr}(S)\right] < R(S)$$

$$\text{bias}(\widehat{R}_{tr}(S)) = \mathbb{E}\left[\widehat{R}_{tr}(S)\right] - R(S) = -2\sum_{i=1}^{n}\text{Cov}\left[\widehat{Y}_i, Y_i\right]$$

Adjusted R^2

$$R^2(S) = 1 - \frac{n-1}{n-k}\frac{\text{RSS}}{\text{TSS}}$$

MALLOW'S C_p statistic

$$\widehat{R}(S) = \widehat{R}_{tr}(S) + 2k\widehat{\sigma}^2 = \text{lack of fit} + \text{complexity penalty}$$

AKAIKE Information Criterion (AIC)

$$AIC(S) = \ell_n(\widehat{\beta}_S, \widehat{\sigma}_S^2) - k$$

Bayesian Information Criterion (BIC)

$$BIC(S) = \ell_n(\widehat{\beta}_S, \widehat{\sigma}_S^2) - \frac{k}{2}\log n$$

Validation and training

$$\widehat{R}_V(S) = \sum_{i=1}^{m}(\widehat{Y}_i^*(S) - Y_i^*)^2 \qquad m = |\{\text{validation data}\}|, \text{ often } \frac{n}{4} \text{ or } \frac{n}{2}$$

Leave-one-out cross-validation

$$\widehat{R}_{CV}(S) = \sum_{i=1}^{n}(Y_i - \widehat{Y}_{(i)}(S))^2 = \sum_{i=1}^{n}\left(\frac{Y_i - \widehat{Y}_i(S)}{1 - U_{ii}(S)}\right)^2$$

$$U(S) = X_S(X_S^T X_S)^{-1}X_S \text{ (``hat matrix'')}$$

19 Non-parametric Function Estimation

19.1 Density Estimation

Estimate $f(x)$, where $f(x) = \mathbb{P}\left[X \in A\right] = \int_A f(x)\,dx$.

Integrated square error (ISE)

$$L(f, \widehat{f}_n) = \int \left(f(x) - \widehat{f}_n(x)\right)^2 dx = J(h) + \int f^2(x)\,dx$$

Frequentist risk

$$R(f, \widehat{f}_n) = \mathbb{E}\left[L(f, \widehat{f}_n)\right] = \int b^2(x)\,dx + \int v(x)\,dx$$

$$b(x) = \mathbb{E}\left[\widehat{f}_n(x)\right] - f(x)$$

$$v(x) = \mathbb{V}\left[\widehat{f}_n(x)\right]$$

19.1.1 Histograms

Definitions

- Number of bins m
- Binwidth $h = \frac{1}{m}$
- Bin B_j has ν_j observations
- Define $\widehat{p}_j = \nu_j/n$ and $p_j = \int_{B_j} f(u)\,du$

Histogram estimator

$$\widehat{f}_n(x) = \sum_{j=1}^{m}\frac{\widehat{p}_j}{h}I(x \in B_j)$$

$$\mathbb{E}\left[\widehat{f}_n(x)\right] = \frac{p_j}{h}$$

$$\mathbb{V}\left[\widehat{f}_n(x)\right] = \frac{p_j(1 - p_j)}{nh^2}$$

$$R(\widehat{f}_n, f) \approx \frac{h^2}{12}\int (f'(u))^2\,du + \frac{1}{nh}$$

$$h^* = \frac{1}{n^{1/3}}\left(\frac{6}{\int (f'(u))^2\,du}\right)^{1/3}$$

$$R^*(\widehat{f}_n, f) \approx \frac{C}{n^{2/3}} \qquad C = \left(\frac{3}{4}\right)^{2/3}\left(\int (f'(u))^2\,du\right)^{1/3}$$

Cross-validation estimate of $\mathbb{E}\left[J(h)\right]$

$$\widehat{J}_{CV}(h) = \int \widehat{f}_n^2(x)\,dx - \frac{2}{n}\sum_{i=1}^{n}\widehat{f}_{(-i)}(X_i) = \frac{2}{(n-1)h} - \frac{n+1}{(n-1)h}\sum_{j=1}^{m}\widehat{p}_j^2$$

19.1.2 Kernel Density Estimator (KDE)

Kernel K

- $K(x) \geq 0$
- $\int K(x)\,dx = 1$
- $\int x K(x)\,dx = 0$
- $\int x^2 K(x)\,dx \equiv \sigma_K^2 > 0$

KDE

$$\widehat{f}_n(x) = \frac{1}{n}\sum_{i=1}^{n}\frac{1}{h}K\left(\frac{x-X_i}{h}\right)$$

$$R(f,\widehat{f}_n) \approx \frac{1}{4}(h\sigma_K)^4\int (f''(x))^2\,dx + \frac{1}{nh}\int K^2(x)\,dx$$

$$h^* = \frac{c_1^{-2/5}c_2^{-1/5}c_3^{-1/5}}{n^{1/5}} \qquad c_1 = \sigma_K^2,\ c_2 = \int K^2(x)\,dx,\ c_3 = \int (f''(x))^2\,dx$$

$$R^*(f,\widehat{f}_n) = \frac{c_4}{n^{4/5}} \qquad c_4 = \underbrace{\frac{5}{4}(\sigma_K^2)^{2/5}\left(\int K^2(x)\,dx\right)^{4/5}}_{C(K)}\left(\int (f'')^2\,dx\right)^{1/5}$$

EPANECHNIKOV Kernel

$$K(x) = \begin{cases} \frac{3}{4\sqrt{5}(1-x^2/5)} & |x| < \sqrt{5} \\ 0 & \text{otherwise} \end{cases}$$

Cross-validation estimate of $\mathbb{E}\left[J(h)\right]$

$$\widehat{J}_{CV}(h) = \int \widehat{f}_n^2(x)\,dx - \frac{2}{n}\sum_{i=1}^{n}\widehat{f}_{(-i)}(X_i) \approx \frac{1}{hn^2}\sum_{i=1}^{n}\sum_{j=1}^{n}K^*\left(\frac{X_i-X_j}{h}\right) + \frac{2}{nh}K(0)$$

$$K^*(x) = K^{(2)}(x) - 2K(x) \qquad K^{(2)}(x) = \int K(x-y)K(y)\,dy$$

19.2 Non-parametric Regression

Estimate $f(x)$ where $f(x) = \mathbb{E}\left[Y \mid X = x\right]$. Consider pairs of points $(x_1, Y_1), \ldots, (x_n, Y_n)$ related by

$$Y_i = r(x_i) + \epsilon_i$$
$$\mathbb{E}\left[\epsilon_i\right] = 0$$
$$\mathbb{V}\left[\epsilon_i\right] = \sigma^2$$

k-nearest Neighbor Estimator

$$\widehat{r}(x) = \frac{1}{k}\sum_{i:x_i \in N_k(x)} Y_i \qquad \text{where } N_k(x) = \{k \text{ values of } x_1, \ldots, x_n \text{ closest to } x\}$$

NADARAYA-WATSON Kernel Estimator

$$\widehat{r}(x) = \sum_{i=1}^{n} w_i(x)Y_i$$

$$w_i(x) = \frac{K\left(\frac{x-x_i}{h}\right)}{\sum_{j=1}^{n} K\left(\frac{x-x_j}{h}\right)} \quad \in [0,1]$$

$$R(\widehat{r}_n, r) \approx \frac{h^4}{4}\left(\int x^2 K^2(x)\,dx\right)^4 \int \left(r''(x) + 2r'(x)\frac{f'(x)}{f(x)}\right)^2 dx$$
$$+ \int \frac{\sigma^2 \int K^2(x)\,dx}{nhf(x)}\,dx$$

$$h^* \approx \frac{c_1}{n^{1/5}}$$

$$R^*(\widehat{r}_n, r) \approx \frac{c_2}{n^{4/5}}$$

Cross-validation estimate of $\mathbb{E}\left[J(h)\right]$

$$\widehat{J}_{CV}(h) = \sum_{i=1}^{n}(Y_i - \widehat{r}_{(-i)}(x_i))^2 = \sum_{i=1}^{n}\frac{(Y_i - \widehat{r}(x_i))^2}{\left(1 - \frac{K(0)}{\sum_{j=1}^{n} K\left(\frac{x-x_j}{h}\right)}\right)^2}$$

19.3 Smoothing Using Orthogonal Functions

Approximation

$$r(x) = \sum_{j=1}^{\infty}\beta_j\phi_j(x) \approx \sum_{i=1}^{J}\beta_j\phi_j(x)$$

Multivariate regression

$$Y = \Phi\beta + \eta$$

$$\text{where} \quad \eta_i = \epsilon_i \quad \text{and} \quad \Phi = \begin{pmatrix} \phi_0(x_1) & \cdots & \phi_J(x_1) \\ \vdots & \ddots & \vdots \\ \phi_0(x_n) & \cdots & \phi_J(x_n) \end{pmatrix}$$

Least squares estimator

$$\widehat{\beta} = (\Phi^T\Phi)^{-1}\Phi^T Y$$

$$\approx \frac{1}{n}\Phi^T Y \quad \text{(for equally spaced observations only)}$$

Cross-validation estimate of $\mathbb{E}\left[J(h)\right]$

$$\widehat{R}_{CV}(J) = \sum_{i=1}^{n} \left(Y_i - \sum_{j=1}^{J} \phi_j(x_i)\widehat{\beta}_{j,(-i)} \right)^2$$

20 Stochastic Processes

Stochastic Process

$$\{X_t : t \in T\} \qquad T = \begin{cases} \{0, \pm 1, \dots\} = \mathbb{Z} & \text{discrete} \\ [0, \infty) & \text{continuous} \end{cases}$$

- Notations X_t, $X(t)$
- State space $\mathcal{X}$
- Index set T

20.1 Markov Chains

Markov chain

$$\mathbb{P}\left[X_n = x \mid X_0, \dots, X_{n-1}\right] = \mathbb{P}\left[X_n = x \mid X_{n-1}\right] \quad \forall n \in T, x \in \mathcal{X}$$

Transition probabilities

$$p_{ij} \equiv \mathbb{P}\left[X_{n+1} = j \mid X_n = i\right]$$
$$p_{ij}(n) \equiv \mathbb{P}\left[X_{m+n} = j \mid X_m = i\right] \quad \text{n-step}$$

Transition matrix $\mathbf{P}$ (n-step: $\mathbf{P}_n$)

- (i,j) element is p_{ij}
- $p_{ij} > 0$
- $\sum_i p_{ij} = 1$

CHAPMAN-KOLMOGOROV

$$p_{ij}(m+n) = \sum_k p_{ij}(m)p_{kj}(n)$$

$$\mathbf{P}_{m+n} = \mathbf{P}_m \mathbf{P}_n$$
$$\mathbf{P}_n = \mathbf{P} \times \cdots \times \mathbf{P} = \mathbf{P}^n$$

Marginal probability

$$\mu_n = (\mu_n(1), \dots, \mu_n(N)) \quad \text{where} \quad \mu_i(i) = \mathbb{P}\left[X_n = i\right]$$
$$\mu_0 \triangleq \text{initial distribution}$$
$$\mu_n = \mu_0 \mathbf{P}^n$$

20.2 Poisson Processes

Poisson process

- $\{X_t : t \in [0, \infty)\}$ = number of events up to and including time t
- $X_0 = 0$
- Independent increments:

$$\forall t_0 < \cdots < t_n : X_{t_1} - X_{t_0} \perp\!\!\!\perp \cdots \perp\!\!\!\perp X_{t_n} - X_{t_{n-1}}$$

- Intensity function $\lambda(t)$

  - $\mathbb{P}\left[X_{t+h} - X_t = 1\right] = \lambda(t)h + o(h)$
  - $\mathbb{P}\left[X_{t+h} - X_t = 2\right] = o(h)$

- $X_{s+t} - X_s \sim \text{Po}\left(m(s+t) - m(s)\right)$ where $m(t) = \int_0^t \lambda(s)\,ds$

Homogeneous Poisson process

$$\lambda(t) \equiv \lambda \implies X_t \sim \text{Po}\left(\lambda t\right) \qquad \lambda > 0$$
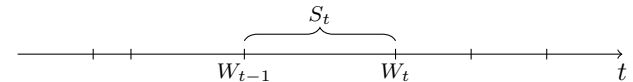
Waiting times

$$W_t := \text{time at which } X_t \text{ occurs}$$

$$W_t \sim \text{Gamma}\left(t, \frac{1}{\lambda}\right)$$

Interarrival times

$$S_t = W_{t+1} - W_t$$

$$S_t \sim \text{Exp}\left(\frac{1}{\lambda}\right)$$

21 Time Series

Mean function

$$\mu_{x_t} = \mathbb{E}\left[x_t\right] = \int_{-\infty}^{\infty} x f_t(x)\, dx$$

Autocovariance function

$$\gamma_x(s,t) = \mathbb{E}\left[(x_s - \mu_s)(x_t - \mu_t)\right] = \mathbb{E}\left[x_s x_t\right] - \mu_s \mu_t$$

$$\gamma_x(t,t) = \mathbb{E}\left[(x_t - \mu_t)^2\right] = \mathbb{V}\left[x_t\right]$$

Autocorrelation function (ACF)

$$\rho(s,t) = \frac{\mathrm{Cov}\left[x_s, x_t\right]}{\sqrt{\mathbb{V}\left[x_s\right]\mathbb{V}\left[x_t\right]}} = \frac{\gamma(s,t)}{\sqrt{\gamma(s,s)\gamma(t,t)}}$$

Cross-covariance function (CCV)

$$\gamma_{xy}(s,t) = \mathbb{E}\left[(x_s - \mu_{x_s})(y_t - \mu_{y_t})\right]$$

Cross-correlation function (CCF)

$$\rho_{xy}(s,t) = \frac{\gamma_{xy}(s,t)}{\sqrt{\gamma_x(s,s)\gamma_y(t,t)}}$$

Backshift operator

$$B^k(x_t) = x_{t-k}$$

Difference operator

$$\nabla^d = (1-B)^d$$

White noise

- $w_t \sim wn(0, \sigma_w^2)$
- Gaussian: $w_t \overset{iid}{\sim} \mathcal{N}\left(0, \sigma_w^2\right)$
- $\mathbb{E}\left[w_t\right] = 0 \quad t \in T$
- $\mathbb{V}\left[w_t\right] = \sigma^2 \quad t \in T$
- $\gamma_w(s,t) = 0 \quad s \neq t \,\wedge\, s,t \in T$

Random walk

- Drift δ
- $x_t = \delta t + \sum_{j=1}^{t} w_j$
- $\mathbb{E}\left[x_t\right] = \delta t$

Symmetric moving average

$$m_t = \sum_{j=-k}^{k} a_j x_{t-j} \qquad \text{where } a_j = a_{-j} \geq 0 \text{ and } \sum_{j=-k}^{k} a_j = 1$$

21.1 Stationary Time Series

Strictly stationary

$$\mathbb{P}\left[x_{t_1} \leq c_1, \ldots, x_{t_k} \leq c_k\right] = \mathbb{P}\left[x_{t_1+h} \leq c_1, \ldots, x_{t_k+h} \leq c_k\right]$$

$$\forall k \in \mathbb{N}, t_k, c_k, h \in \mathbb{Z}$$

Weakly stationary

- $\mathbb{E}\left[x_t^2\right] < \infty \qquad \forall t \in \mathbb{Z}$
- $\mathbb{E}\left[x_t^2\right] = m \qquad \forall t \in \mathbb{Z}$
- $\gamma_x(s,t) = \gamma_x(s+r, t+r) \qquad \forall r,s,t \in \mathbb{Z}$

Autocovariance function

- $\gamma(h) = \mathbb{E}\left[(x_{t+h} - \mu)(x_t - \mu)\right] \qquad \forall h \in \mathbb{Z}$
- $\gamma(0) = \mathbb{E}\left[(x_t - \mu)^2\right]$
- $\gamma(0) \geq 0$
- $\gamma(0) \geq |\gamma(h)|$
- $\gamma(h) = \gamma(-h)$

Autocorrelation function (ACF)

$$\rho_x(h) = \frac{\mathrm{Cov}\left[x_{t+h}, x_t\right]}{\sqrt{\mathbb{V}\left[x_{t+h}\right]\mathbb{V}\left[x_t\right]}} = \frac{\gamma(t+h, t)}{\sqrt{\gamma(t+h, t+h)\gamma(t,t)}} = \frac{\gamma(h)}{\gamma(0)}$$

Jointly stationary time series

$$\gamma_{xy}(h) = \mathbb{E}\left[(x_{t+h} - \mu_x)(y_t - \mu_y)\right]$$

$$\rho_{xy}(h) = \frac{\gamma_{xy}(h)}{\sqrt{\gamma_x(0)\gamma_y(h)}}$$

Linear process

$$x_t = \mu + \sum_{j=-\infty}^{\infty} \psi_j w_{t-j} \quad \text{where} \quad \sum_{j=-\infty}^{\infty} |\psi_j| < \infty$$

$$\gamma(h) = \sigma_w^2 \sum_{j=-\infty}^{\infty} \psi_{j+h} \psi_j$$

21.2 Estimation of Correlation

Sample mean

$$\bar{x} = \frac{1}{n} \sum_{t=1}^{n} x_t$$

Sample variance

$$\mathbb{V}[\bar{x}] = \frac{1}{n} \sum_{h=-n}^{n} \left(1 - \frac{|h|}{n}\right) \gamma_x(h)$$

Sample autocovariance function

$$\widehat{\gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x})$$

Sample autocorrelation function

$$\widehat{\rho}(h) = \frac{\widehat{\gamma}(h)}{\widehat{\gamma}(0)}$$

Sample cross-variance function

$$\widehat{\gamma}_{xy}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(y_t - \overline{y})$$

Sample cross-correlation function

$$\widehat{\rho}_{xy}(h) = \frac{\widehat{\gamma}_{xy}(h)}{\sqrt{\widehat{\gamma}_x(0)\widehat{\gamma}_y(0)}}$$

Properties

- $\sigma_{\widehat{\rho}_x(h)} = \dfrac{1}{\sqrt{n}}$ if x_t is white noise

- $\sigma_{\widehat{\rho}_{xy}(h)} = \dfrac{1}{\sqrt{n}}$ if x_t or y_t is white noise

21.3 Non-Stationary Time Series

Classical decomposition model

$$x_t = \mu_t + s_t + w_t$$

- μ_t = trend
- s_t = seasonal component
- w_t = random noise term

21.3.1 Detrending

Least squares

1. Choose trend model, e.g., $\mu_t = \beta_0 + \beta_1 t + \beta_2 t^2$
2. Minimize RSS to obtain trend estimate $\widehat{\mu}_t = \widehat{\beta}_0 + \widehat{\beta}_1 t + \widehat{\beta}_2 t^2$
3. Residuals $\triangleq$ noise w_t

Moving average

- The *low-pass* filter v_t is a symmetric moving average m_t with $a_j = \frac{1}{2k+1}$:

$$v_t = \frac{1}{2k+1} \sum_{i=-k}^{k} x_{t-1}$$

- If $\frac{1}{2k+1} \sum_{i=-k}^{k} w_{t-j} \approx 0$, a linear trend function $\mu_t = \beta_0 + \beta_1 t$ passes without distortion

Differencing

- $\mu_t = \beta_0 + \beta_1 t \implies \nabla x_t = \beta_1$

21.4 ARIMA models

Autoregressive polynomial

$$\phi(z) = 1 - \phi_1 z - \cdots - \phi_p z_p \qquad z \in \mathbb{C} \wedge \phi_p \neq 0$$

Autoregressive operator

$$\phi(B) = 1 - \phi_1 B - \cdots - \phi_p B^p$$

Autoregressive model order p, $\mathsf{AR}(p)$

$$x_t = \phi_1 x_{t-1} + \cdots + \phi_p x_{t-p} + w_t \iff \phi(B) x_t = w_t$$

$\mathsf{AR}(1)$

- $x_t = \phi^k(x_{t-k}) + \sum_{j=0}^{k-1} \phi^j (w_{t-j}) \overset{k \to \infty, |\phi| < 1}{=} \underbrace{\sum_{j=0}^{\infty} \phi^j (w_{t-j})}_{\text{linear process}}$

- $\mathbb{E}[x_t] = \sum_{j=0}^{\infty} \phi^j (\mathbb{E}[w_{t-j}]) = 0$
- $\gamma(h) = \mathrm{Cov}[x_{t+h}, x_t] = \frac{\sigma_w^2 \phi^h}{1 - \phi^2}$
- $\rho(h) = \frac{\gamma(h)}{\gamma(0)} = \phi^h$
- $\rho(h) = \phi\rho(h-1) \quad h = 1, 2, \ldots$

Moving average polynomial

$$\theta(z) = 1 + \theta_1 z + \cdots + \theta_q z_q \qquad z \in \mathbb{C} \wedge \theta_q \neq 0$$

Moving average operator

$$\theta(B) = 1 + \theta_1 B + \cdots + \theta_p B^p$$

MA (q) (moving average model order q)

$$x_t = w_t + \theta_1 w_{t-1} + \cdots + \theta_q w_{t-q} \iff x_t = \theta(B)w_t$$

$$\mathbb{E}[x_t] = \sum_{j=0}^{q} \theta_j \mathbb{E}[w_{t-j}] = 0$$

$$\gamma(h) = \mathrm{Cov}\,[x_{t+h}, x_t] = \begin{cases} \sigma_w^2 \sum_{j=0}^{q-h} \theta_j \theta_{j+h} & 0 \leq h \leq q \\ 0 & h > q \end{cases}$$

MA (1)

$$x_t = w_t + \theta w_{t-1}$$

$$\gamma(h) = \begin{cases} (1 + \theta^2)\sigma_w^2 & h = 0 \\ \theta \sigma_w^2 & h = 1 \\ 0 & h > 1 \end{cases}$$

$$\rho(h) = \begin{cases} \frac{\theta}{(1+\theta^2)} & h = 1 \\ 0 & h > 1 \end{cases}$$

ARMA (p, q)

$$x_t = \phi_1 x_{t-1} + \cdots + \phi_p x_{t-p} + w_t + \theta_1 w_{t-1} + \cdots + \theta_q w_{t-q}$$

$$\phi(B)x_t = \theta(B)w_t$$

Partial autocorrelation function (PACF)

- $x_i^{h-1} \triangleq$ regression of x_i on $\{x_{h-1}, x_{h-2}, \ldots, x_1\}$
- $\phi_{hh} = corr(x_h - x_h^{h-1}, x_0 - x_0^{h-1}) \quad h \geq 2$
- E.g., $\phi_{11} = corr(x_1, x_0) = \rho(1)$

ARIMA (p, d, q)

$$\nabla^d x_t = (1 - B)^d x_t \text{ is ARMA}\,(p, q)$$

$$\phi(B)(1 - B)^d x_t = \theta(B)w_t$$

Exponentially Weighted Moving Average (EWMA)

$$x_t = x_{t-1} + w_t - \lambda w_{t-1}$$

$$x_t = \sum_{j=1}^{\infty}(1 - \lambda)\lambda^{j-1} x_{t-j} + w_t \quad \text{when } |\lambda| < 1$$

$$\tilde{x}_{n+1} = (1 - \lambda)x_n + \lambda \tilde{x}_n$$

Seasonal ARIMA

- Denoted by $\mathsf{ARIMA}\,(p, d, q) \times (P, D, Q)_s$
- $\Phi_P(B^s)\phi(B)\nabla_s^D \nabla^d x_t = \delta + \Theta_Q(B^s)\theta(B)w_t$

21.4.1 Causality and Invertibility

ARMA (p, q) is causal (future-independent) $\iff \exists\{\psi_j\} : \sum_{j=0}^{\infty} \psi_j < \infty$ such that

$$x_t = \sum_{j=0}^{\infty} w_{t-j} = \psi(B)w_t$$

ARMA (p, q) is invertible $\iff \exists\{\pi_j\} : \sum_{j=0}^{\infty} \pi_j < \infty$ such that

$$\pi(B)x_t = \sum_{j=0}^{\infty} X_{t-j} = w_t$$

Properties

- ARMA (p, q) causal $\iff$ roots of $\phi(z)$ lie outside the unit circle

$$\psi(z) = \sum_{j=0}^{\infty} \psi_j z^j = \frac{\theta(z)}{\phi(z)} \quad |z| \leq 1$$

- ARMA (p, q) invertible $\iff$ roots of $\theta(z)$ lie outside the unit circle

$$\pi(z) = \sum_{j=0}^{\infty} \pi_j z^j = \frac{\phi(z)}{\theta(z)} \quad |z| \leq 1$$

Behavior of the ACF and PACF for causal and invertible ARMA models

	AR (p)	MA (q)	ARMA (p, q)
ACF	tails off	cuts off after lag q	tails off
PACF	cuts off after lag p	tails off q	tails off

21.5 Spectral Analysis

Periodic process

$$x_t = A\cos(2\pi\omega t + \phi)$$
$$= U_1 \cos(2\pi\omega t) + U_2 \sin(2\pi\omega t)$$

- Frequency index ω (cycles per unit time), period $1/\omega$

- Amplitude A
- Phase ϕ
- $U_1 = A\cos\phi$ and $U_2 = A\sin\phi$ often normally distributed RV's

Periodic mixture

$$x_t = \sum_{k=1}^{q} \left(U_{k1}\cos(2\pi\omega_k t) + U_{k2}\sin(2\pi\omega_k t) \right)$$

- U_{k1}, U_{k2}, for $k = 1, \ldots, q$, are independent zero-mean RV's with variances σ_k^2
- $\gamma(h) = \sum_{k=1}^{q} \sigma_k^2 \cos(2\pi\omega_k h)$
- $\gamma(0) = \mathbb{E}\left[x_t^2 \right] = \sum_{k=1}^{q} \sigma_k^2$

Spectral representation of a periodic process

$$\gamma(h) = \sigma^2 \cos(2\pi\omega_0 h)$$
$$= \frac{\sigma^2}{2} e^{-2\pi i\omega_0 h} + \frac{\sigma^2}{2} e^{2\pi i\omega_0 h}$$
$$= \int_{-1/2}^{1/2} e^{2\pi i\omega h}\, dF(\omega)$$

Spectral distribution function

$$F(\omega) = \begin{cases} 0 & \omega < -\omega_0 \\ \sigma^2/2 & -\omega \le \omega < \omega_0 \\ \sigma^2 & \omega \ge \omega_0 \end{cases}$$

- $F(-\infty) = F(-1/2) = 0$
- $F(\infty) = F(1/2) = \gamma(0)$

Spectral density

$$f(\omega) = \sum_{h=-\infty}^{\infty} \gamma(h) e^{-2\pi i\omega h} \quad -\frac{1}{2} \le \omega \le \frac{1}{2}$$

- Needs $\sum_{h=-\infty}^{\infty} |\gamma(h)| < \infty \implies \gamma(h) = \int_{-1/2}^{1/2} e^{2\pi i\omega h} f(\omega)\, d\omega \quad h = 0, \pm 1, \ldots$
- $f(\omega) \ge 0$
- $f(\omega) = f(-\omega)$
- $f(\omega) = f(1 - \omega)$
- $\gamma(0) = \mathbb{V}\left[x_t \right] = \int_{-1/2}^{1/2} f(\omega)\, d\omega$
- White noise: $f_w(\omega) = \sigma_w^2$
- ARMA (p, q), $\phi(B) x_t = \theta(B) w_t$:

$$f_x(\omega) = \sigma_w^2 \frac{|\theta(e^{-2\pi i\omega})|^2}{|\phi(e^{-2\pi i\omega})|^2}$$

where $\phi(z) = 1 - \sum_{k=1}^{p} \phi_k z^k$ and $\theta(z) = 1 + \sum_{k=1}^{q} \theta_k z^k$

Discrete Fourier Transform (DFT)

$$d(\omega_j) = n^{-1/2} \sum_{i=1}^{n} x_t e^{-2\pi i\omega_j t}$$

Fourier/Fundamental frequencies

$$\omega_j = j/n$$

Inverse DFT

$$x_t = n^{-1/2} \sum_{j=0}^{n-1} d(\omega_j) e^{2\pi i\omega_j t}$$

Periodogram

$$I(j/n) = |d(j/n)|^2$$

Scaled Periodogram

$$P(j/n) = \frac{4}{n} I(j/n)$$
$$= \left(\frac{2}{n} \sum_{t=1}^{n} x_t \cos(2\pi t j/n) \right)^2 + \left(\frac{2}{n} \sum_{t=1}^{n} x_t \sin(2\pi t j/n) \right)^2$$

22 Math

22.1 Gamma Function

- Ordinary: $\Gamma(s) = \int_0^{\infty} t^{s-1} e^{-t} dt$
- Upper incomplete: $\Gamma(s, x) = \int_x^{\infty} t^{s-1} e^{-t} dt$
- Lower incomplete: $\gamma(s, x) = \int_0^x t^{s-1} e^{-t} dt$
- $\Gamma(\alpha + 1) = \alpha \Gamma(\alpha) \qquad \alpha > 1$
- $\Gamma(n) = (n - 1)! \qquad n \in \mathbb{N}$
- $\Gamma(1/2) = \sqrt{\pi}$

22.2 Beta Function

- Ordinary: $\mathrm{B}(x, y) = \mathrm{B}(y, x) = \int_0^1 t^{x-1} (1 - t)^{y-1}\, dt = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x + y)}$
- Incomplete: $\mathrm{B}(x; a, b) = \int_0^x t^{a-1} (1 - t)^{b-1}\, dt$
- Regularized incomplete:
$$I_x(a, b) = \frac{\mathrm{B}(x; a, b)}{\mathrm{B}(a, b)} \overset{a, b \in \mathbb{N}}{=} \sum_{j=a}^{a+b-1} \frac{(a + b - 1)!}{j!(a + b - 1 - j)!} x^j (1 - x)^{a+b-1-j}$$

- $I_0(a, b) = 0 \qquad I_1(a, b) = 1$
- $I_x(a, b) = 1 - I_{1-x}(b, a)$

22.3 Series

Finite

- $\displaystyle\sum_{k=1}^{n} k = \frac{n(n+1)}{2}$
- $\displaystyle\sum_{k=1}^{n} (2k - 1) = n^2$
- $\displaystyle\sum_{k=1}^{n} k^2 = \frac{n(n+1)(2n+1)}{6}$
- $\displaystyle\sum_{k=1}^{n} k^3 = \left(\frac{n(n+1)}{2}\right)^2$
- $\displaystyle\sum_{k=0}^{n} c^k = \frac{c^{n+1} - 1}{c - 1} \quad c \neq 1$

Binomial

- $\displaystyle\sum_{k=0}^{n} \binom{n}{k} = 2^n$
- $\displaystyle\sum_{k=0}^{n} \binom{r+k}{k} = \binom{r+n+1}{n}$
- $\displaystyle\sum_{k=0}^{n} \binom{k}{m} = \binom{n+1}{m+1}$
- VANDERMONDE's Identity:
$$\sum_{k=0}^{r} \binom{m}{k}\binom{n}{r-k} = \binom{m+n}{r}$$
- Binomial Theorem:
$$\sum_{k=0}^{n} \binom{n}{k} a^{n-k} b^k = (a+b)^n$$

Infinite

- $\displaystyle\sum_{k=0}^{\infty} p^k = \frac{1}{1-p}, \quad \sum_{k=1}^{\infty} p^k = \frac{p}{1-p} \quad |p| < 1$
- $\displaystyle\sum_{k=0}^{\infty} k p^{k-1} = \frac{d}{dp}\left(\sum_{k=0}^{\infty} p^k\right) = \frac{d}{dp}\left(\frac{1}{1-p}\right) = \frac{1}{(1-p)^2} \quad |p| < 1$
- $\displaystyle\sum_{k=0}^{\infty} \binom{r+k-1}{k} x^k = (1-x)^{-r} \quad r \in \mathbb{N}^+$
- $\displaystyle\sum_{k=0}^{\infty} \binom{\alpha}{k} p^k = (1+p)^\alpha \quad |p| < 1, \, \alpha \in \mathbb{C}$

22.4 Combinatorics

Sampling

k out of n	w/o replacement	w/ replacement
ordered	$n^{\underline{k}} = \displaystyle\prod_{i=0}^{k-1}(n-i) = \frac{n!}{(n-k)!}$	n^k
unordered	$\binom{n}{k} = \frac{n^{\underline{k}}}{k!} = \frac{n!}{k!(n-k)!}$	$\binom{n-1+r}{r} = \binom{n-1+r}{n-1}$

Stirling numbers, 2nd kind

$$\left\{ n \atop k \right\} = k \left\{ n-1 \atop k \right\} + \left\{ n-1 \atop k-1 \right\} \qquad 1 \leq k \leq n \qquad \left\{ n \atop 0 \right\} = \begin{cases} 1 & n = 0 \\ 0 & \text{else} \end{cases}$$

Partitions

$$P_{n+k,k} = \sum_{i=1}^{n} P_{n,i} \qquad k > n: \; P_{n,k} = 0 \qquad n \geq 1: \; P_{n,0} = 0, \; P_{0,0} = 1$$

Balls and Urns $\qquad f : B \to U \qquad D = \text{distinguishable}, \neg D = \text{indistinguishable}.$

$|B| = n, |U| = m$	f arbitrary	f injective	f surjective	f bijective
$B : D, \; U : \neg D$	m^n	$\begin{cases} m^{\underline{n}} & m \geq n \\ 0 & \text{else} \end{cases}$	$m! \left\{ n \atop m \right\}$	$\begin{cases} n! & m = n \\ 0 & \text{else} \end{cases}$
$B : \neg D, \; U : D$	$\binom{n+n-1}{n}$	$\binom{m}{n}$	$\binom{n-1}{m-1}$	$\begin{cases} 1 & m = n \\ 0 & \text{else} \end{cases}$
$B : D, \; U : \neg D$	$\displaystyle\sum_{k=1}^{m} \left\{ n \atop k \right\}$	$\begin{cases} 1 & m \geq n \\ 0 & \text{else} \end{cases}$	$\left\{ n \atop m \right\}$	$\begin{cases} 1 & m = n \\ 0 & \text{else} \end{cases}$
$B : \neg D, \; U : \neg D$	$\displaystyle\sum_{k=1}^{m} P_{n,k}$	$\begin{cases} 1 & m \geq n \\ 0 & \text{else} \end{cases}$	$P_{n,m}$	$\begin{cases} 1 & m = n \\ 0 & \text{else} \end{cases}$

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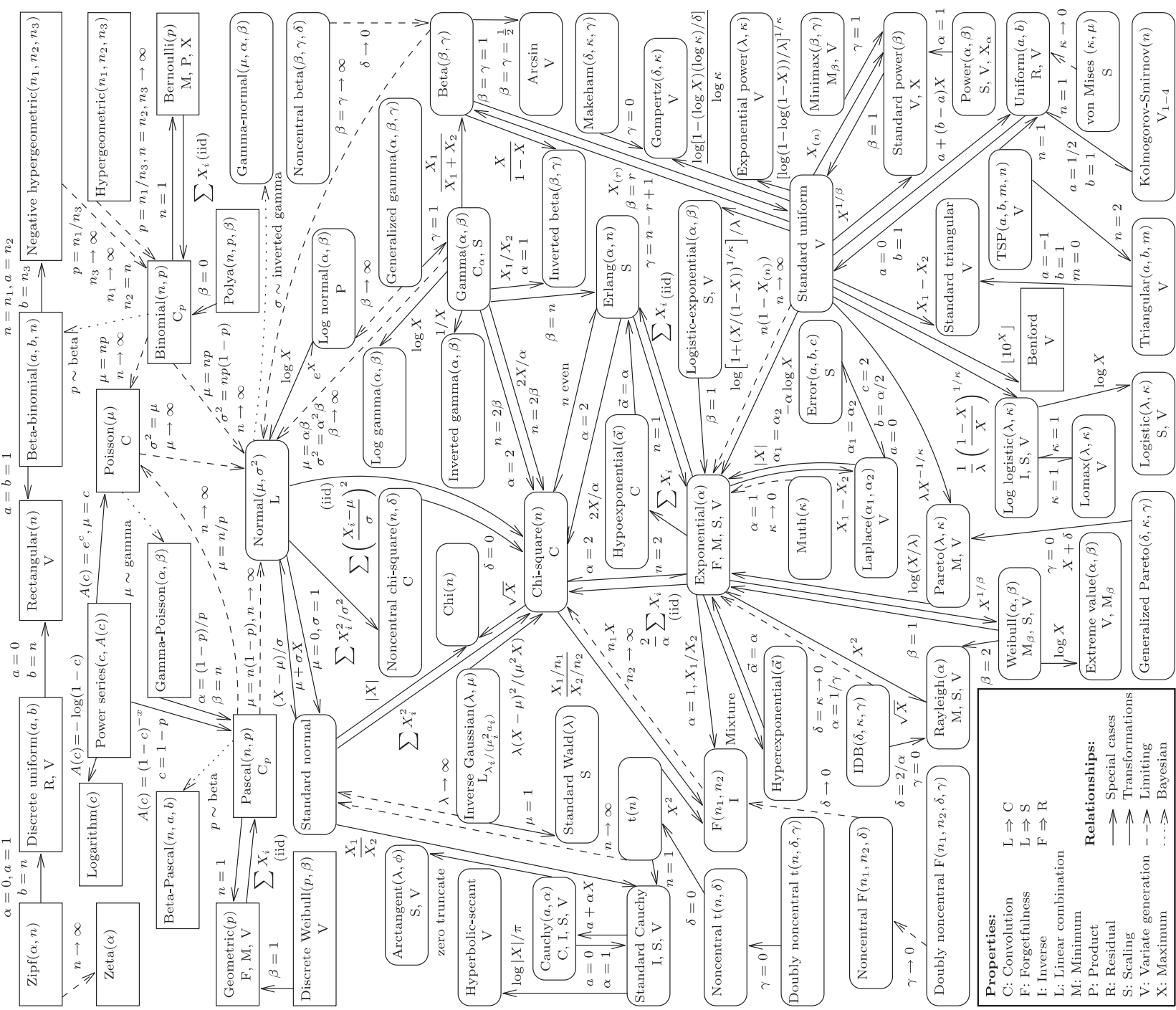
Univariate distribution relationships, courtesy Leemis and McQueston [2].