

Properties of Wishart Distribution

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1. Definition

Model: $W \sim W_p(r, \Sigma) \iff W \stackrel{d}{=} \sum_{j=1}^n Z_j Z_j^\top$

- Z_j : multivariate standard normal distribution
 $Z_j \stackrel{iid}{\sim} N_p(0, \Sigma), (j = 1, \dots, r).$

2. Basic Properties of Wishart Distribution

- Assuming $Z_j \stackrel{iid}{\sim} N_p(0, \Sigma)$, $(j = 1, \dots, r)$, $X_i \stackrel{iid}{\sim} N_p(0, \Sigma)$, $(i = 1, \dots, n)$, $X : n \times p$ data matrix.

- (quadratic form)

$$X^\top A X = \sum_{j=1}^r Z_j Z_j^\top \sim W_p(r, \Sigma) \iff A^2 = A,$$

where $r = \text{rank}(A) = \text{tr}(A)$

- (closure under convolution)

$$W_1 \sim W_p(r_1, \Sigma) \stackrel{d}{=} \sum_{j=1}^{r_1} Z_j Z_j^\top, W_2 \sim W_p(r_2, \Sigma) \stackrel{d}{=} \sum_{j=1}^{r_2} Z_j Z_j^\top$$

$$W_1 \perp W_2, W_1 + W_2 \sim W_p(r_1 + r_2, \Sigma)$$

- (scaled form)

$$W \sim W_p(r, \Sigma) \Rightarrow C W C^\top \sim W_q(r, C \Sigma C^\top) \forall C : q \times p$$

2. Basic Properties of Wishart Distribution : proof(1)

- (quadratic form)

$$(\Leftarrow) A^2 = A, X^T A X = \sum_{j=1}^r X^T P_j P_j^T X = \sum_{j=1}^r Z_j Z_j^T \sim W_p(r, \Sigma)$$

$$(\Rightarrow) X^T A X \sim W_p(r, \Sigma) \stackrel{d}{=} \sum_{j=1}^r Z_j Z_j^T \rightarrow \forall d \in R^p, d^T X^T A X d =$$

$$d^T \left(\sum_{j=1}^r Z_j Z_j^T \right) d = \sum_{j=1}^r d^T Z_j Z_j^T d \stackrel{d}{=} (d^T \Sigma d) \chi_r^2$$

- (closure under convolution)

$$W_1 + W_2 = \sum_{j=1}^{r_1} Z_{1j} Z_{1j}^T + \sum_{k=1}^{r_2} Z_{2k} Z_{2k}^T \stackrel{d}{=}$$

$$\sum_{j=1}^{r_1} Z_{1j} Z_{1j}^T + \sum_{j=r_1+1}^{r_1+r_2} Z_{1j} Z_{1j}^T = \sum_{j=1}^{r_1+r_2} Z_{1j} Z_{1j}^T \sim W_p(r_1 + r_2, \Sigma)$$

2. Basic Properties of Wishart Distribution : proof(2)

- (scaled form)

$$CW C^{\top} = C \left[\sum_{j=1}^r Z_j Z_j^{\top} \right] C^{\top} = \sum_{j=1}^r C Z_j C Z_j^{\top} = \sum_{j=1}^r Y_j Y_j^{\top} \sim W_q(r, C \Sigma C^{\top}), \because Y_j = C Z_j \sim N_q(0, C \Sigma C^{\top})$$

3. Additional properties of Wishart Distribution(1)

- Assume $S \sim W_p(r, \Sigma) (r > p)$

- $$S = \begin{pmatrix} S_{11} & s_{12} \\ s_{21} & s_{22} \end{pmatrix}$$

- $$\Sigma = \begin{pmatrix} \Sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix}$$

- $$S^{-1} = \begin{pmatrix} S^{11} & s^{12} \\ s^{21} & s^{22} \end{pmatrix}$$

- $$\Sigma^{-1} = \begin{pmatrix} \Sigma^{11} & \sigma^{12} \\ \sigma^{21} & \sigma^{22} \end{pmatrix}$$

3. Additional properties of Wishart Distribution (2)

- Then following can be shown

- (a) $\sigma^{22}/s^{22} \sim \chi^2(r - (p - 1))$,

where $S \stackrel{d}{=} \sum_{i=1}^r Z_i Z_i^\top$, $Z_i = [Z_{i1}^\top, z_{i2}^\top]^\top$, $Z_i \stackrel{iid}{\sim} N_p(0, \Sigma)$

- (b) $\forall d, (d^\top \Sigma^{-1} d) / (d^\top S^{-1} d) \sim \chi^2(r - (p - 1))$
- (c) $S_{22 \cdot 1} = S_{22} - S_{21} S_{11}^{-1} S_{12} \sim W_{p_2}(r - p_1, \Sigma_{22 \cdot 1})$

3. Additional properties of Wishart Distribution (proof of (a)-1)

- $s^{22} = (s_{22} - s_{21}S_{11}^{-1}s_{12})^{-1}$, $\sigma^{22} = (\sigma_{22} - \sigma_{21}\Sigma_{11}^{-1}\sigma_{12})^{-1}$
- $Y = \begin{bmatrix} Z_{12} \\ \vdots \\ Z_{r2} \end{bmatrix}$, $X = \begin{bmatrix} Z_{11}^\top \\ \vdots \\ Z_{r1}^\top \end{bmatrix}$
- put $\epsilon = Y - X\beta$, where $\beta = \Sigma_{11}^{-1}\sigma_{12}$
- $\epsilon = Y - X\beta = Y - X\Sigma_{11}^{-1}\sigma_{12} \sim N(0, \sigma_{22\cdot 1}I_r)$, $\epsilon \perp X$
- keypoint : $Z_{i1} \perp Z_{i2} - \sigma_{21}\Sigma_{11}^{-1}Z_{i1} \sim N(0, \sigma_{22\cdot 1})$

3. Additional properties of Wishart Distribution (proof of (a)-2)



$$S = \sum_{i=1}^r Z_i Z_i^\top \quad (1)$$

$$= \sum_{i=1}^r \begin{bmatrix} Z_{i1} \\ z_{i2} \end{bmatrix} [Z_{i1}^\top, z_{i2}^\top] \quad (2)$$

$$= \begin{bmatrix} \sum_{i=1}^r Z_{i1} Z_{i1}^\top & \sum_{i=1}^r Z_{i1} z_{i2}^\top \\ \sum_{i=1}^r z_{i2} Z_{i1}^\top & \sum_{i=1}^r z_{i2} z_{i2}^\top \end{bmatrix} \quad (3)$$

- $(s^{22})^{-1} = 1/s^{22} = (s_{22} - s_{21}S_{11}^{-1}s_{12})$

3. Additional properties of Wishart Distribution (proof of (a)-3)



$$(s^{22})^{-1} = (s_{22} - s_{21}S_{11}^{-1}s_{12}) \quad (4)$$

$$= \left[\sum_{i=1}^r z_{i2}z_{i2}^{\top} - \sum_{i=1}^r z_{i2}Z_{i1}^{\top} \left(\sum_{i=1}^r Z_{i1}Z_{i1}^{\top} \right)^{-1} \sum_{i=1}^r Z_{i1}z_{i2}^{\top} \right] \quad (5)$$

$$= [Y^{\top}Y - Y^{\top}X(X^{\top}X)^{-1}X^{\top}Y] \quad (6)$$

$$= Y^{\top}[I_r - X(X^{\top}X)^{-1}X^{\top}]Y \quad (7)$$

$$:= Y^{\top}AY \quad (8)$$

3. Additional properties of Wishart Distribution (proof of (a)-4)

$$(s^{22})^{-1} = Y^{\top} [I_r - X(X^{\top}X)^{-1}X^{\top}]Y$$

- $$\begin{aligned} &= (Y - X\beta)^{\top} [I_r - X(X^{\top}X)^{-1}X^{\top}] (Y - X\beta) \\ &= (Y - X\beta)^{\top} A (Y - X\beta) \sim \sigma_{22 \cdot 1}^2 \chi_{r-(p-1)}^2 \end{aligned}$$
- $\because (Y - X\beta) \sim N_r(0, (\sigma_{22 \cdot 1})I_r) \rightarrow (\sigma_{22 \cdot 1})^{-1/2}(Y - X\beta) \sim N_r(0, I_r)$
- $\because \text{rank}(A) = \text{tr}(A) = \text{tr}(I_r - X(X^{\top}X)^{-1}X^{\top}) = r - \text{tr}[I_{p-1}] = r - (p-1)$
- $\therefore \frac{1}{s^{22}} \cdot (\sigma_{22 \cdot 1})^{-1} = \frac{\sigma_{22}^2}{s^{22}} \sim \chi_{r-(p-1)}^2$

3. Additional properties of Wishart Distribution (sketch of (c))

- $S_{22 \cdot 1} = S_{22} - S_{21}S_{11}^{-1}S_{12} \sim W_{p_2}(r - p_1, \Sigma_{22 \cdot 1})$
- It is multivariate version of (a), where Z_{i1} is p_1 dimension, Z_{i2} is p_2 dimension.

4. Additional topics of Wishart Distribution (expectation of inverse wishart (1))

Theorem 1

$M \sim W_p(r, \Sigma)$ ($r > p$), then $E(M^{-1}) = \frac{\Sigma^{-1}}{r-p-1}$

Proof Let $\Sigma = CC^\top$, where C is nonsingular. Then $M = CBC^\top \sim W_p(r, \Sigma)$, where $B \sim W_p(r, I) \rightarrow E(M^{-1}) = (C^\top)^{-1}E(B^{-1})C^{-1}$, by symmetry, the diagonal elements of $E(B^{-1})$ are same and off-diagonal elements are same $\rightarrow E(B^{-1}) = k_1 I_p + k_2 \mathbf{1}\mathbf{1}^\top$.

4. Additional topics of Wishart Distribution (expectation of inverse wishart (2))

Proof (continued) For every orthogonal matrix Q , $QBQ^\top \sim W_p(r, I)$
 $\rightarrow E[(QBQ^\top)^{-1}] = (Q^\top)^{-1}E[B^{-1}]Q^{-1} = E(B^{-1})$. $\therefore k_2 = 0$

$$\text{Let } B = \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & B_{22} \end{bmatrix}, B^{-1} = \begin{bmatrix} b^{11} & \dots & b^{1p} \\ b^{21} & \dots & b^{2p} \\ \vdots & & \vdots \\ b^{p1} & \dots & b^{pp} \end{bmatrix}$$

$$\text{Let } Z_i = \begin{bmatrix} z_{i1} \\ Z_{i2} \end{bmatrix}, Z_i \sim N_p(0, I)$$

$$\text{Put } Z_1 = \begin{bmatrix} z_{11} \\ \vdots \\ z_{r1} \end{bmatrix}, Z_2 = \begin{bmatrix} Z_{12}^\top \\ \vdots \\ Z_{r2}^\top \end{bmatrix}$$

4. Additional topics of Wishart Distribution (expectation of inverse wishart (3))

Proof (continued)

$$B = \sum_{i=1}^r Z_i Z_i^\top \quad (9)$$

$$= \sum_{i=1}^r \begin{bmatrix} z_{i1} \\ Z_{i2} \end{bmatrix} [z_{i1}^\top, Z_{i2}^\top] \quad (10)$$

$$= \begin{bmatrix} \sum_{i=1}^r z_{i1} z_{i1}^\top & \sum_{i=1}^r z_{i1} Z_{i2}^\top \\ \sum_{i=1}^r Z_{i2} z_{i1}^\top & \sum_{i=1}^r Z_{i2} Z_{i2}^\top \end{bmatrix} \quad (11)$$

$$= \begin{bmatrix} z_1 z_1^\top & z_1 Z_2^\top \\ Z_2 z_1^\top & Z_2 Z_2^\top \end{bmatrix} \quad (12)$$

4. Additional topics of Wishart Distribution (expectation of inverse wishart (4))

Proof (continued)

$$(b^{11})^{-1} = z_1^\top z_1 - z_1^\top [Z_2(Z_2^\top Z_2)Z_2^\top] z_1 \quad (13)$$

$$= z_1^\top [I_r - Z_2(Z_2^\top Z_2)Z_2^\top] z_1 \quad (14)$$

$$= z_1^\top V z_1 \sim \chi^2_{r-(p-1)} \quad (15)$$

($\because z_1 \sim N_r(0, I)$), $V = I_r - Z_2(Z_2^\top Z_2)Z_2^\top$ is idempotent with
rank(V) = tr(V) = r-(p-1))
 $\Rightarrow b^{11} \sim [\chi^2_{r-(p-1)}]^{-1}$

4. Additional topics of Wishart Distribution (expectation of inverse wishart (5))

Proof (continued) Back to problem :

$$E(B^{-1}) = k_1 I_p, \quad E(b^{11}) = E[(\chi^2_{r-p+1})^{-1}] = \frac{1}{r-p-1} = k_1$$

$$\therefore E(M^{-1}) = (C^\top)^{-1} \left(\frac{1}{r-p-1} I_p \right) (C^{-1}) = \frac{1}{r-p-1} (CC^\top)^{-1} = \frac{1}{r-p-1} \Sigma^{-1}$$

4. Additional topics of Wishart Distribution (conditional wishart (1))

Theorem 2

Let $W \sim W_p(n, \Sigma)$ and W and Σ be partitioned as

$$W = \begin{bmatrix} W_{11} & W_{12} \\ W_{21} & W_{22} \end{bmatrix}, \Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$$

respectively, where $i_j : p_i \times p_j$ and $\Sigma_{ij} : p_i \times p_j$. Then:

- $W_{11.2} \perp\!\!\!\perp \{W_{21}, W_{22}\}$
- The conditional distribution of $W_{12} = (B_1, \dots, B_{p_2})$ is normal with mean $E(W_{12}|W_{22}) = \Sigma_{12}\Sigma_{22}^{-1}W_{22}$ and

$$\text{Cov}(B_i, B_j|W_{22}) = (W_{22})_{ij}\Sigma_{11.2},$$

where $(A)_{ij}$ denotes the (i, j) th element of A .

4. Additional topics of Wishart Distribution (conditional wishart (2))

Proof We may write W as

$$W = \sum_{i=1}^n Z_i Z_i^\top = \sum_{i=1}^n \begin{bmatrix} Z_{i1} \\ Z_{i2} \end{bmatrix} [Z_{i1}^\top, Z_{i2}^\top], \quad Z_i \sim N_p(0, I), i = 1, \dots, n$$

Consider orthogonal matrix $H = [H_1 \ H_2]$ such that $H_1 = X_2(X_2^\top X_2)^{-1/2}$. Then it can be shown that :

- $H^\top H = H H^\top = I_n, \ H_1^\top H_2 = 0.$
- $H_1^\top H_1 = I_{p_2}, \ H_1 H_1^\top = Z_2(Z_2^\top Z_2)^{-1} Z_2^\top$
- $H_2 = Z_{1|2}(Z_{1|2}^\top Z_{1|2})^{-1} Z_{1|2}^\top$, where $Z_{1|2} = (I - Z_2(Z_2^\top Z_2)^{-1} Z_2^\top) Z_1$

4. Additional topics of Wishart Distribution (conditional wishart (3))

Proof (continued)

$$\Rightarrow \text{Consider } \mathbf{X} = (\mathbf{X}_1, \dots, \mathbf{X}_n)^\top = \mathbf{H}^\top \mathbf{Z}_1 = \begin{bmatrix} \mathbf{H}_1^\top \\ \mathbf{H}_2^\top \end{bmatrix} \mathbf{Z}_1 = \begin{bmatrix} \mathbf{H}_1^\top \mathbf{Z}_1 \\ \mathbf{H}_2^\top \mathbf{Z}_1 \end{bmatrix},$$

$$\text{shortly } \mathbf{X} = \begin{bmatrix} \mathbf{X}_1^\top \\ \vdots \\ \mathbf{X}_{p_2}^\top \\ \mathbf{X}_{p_2+1}^\top \\ \vdots \\ \mathbf{X}_n^\top \end{bmatrix}, \Rightarrow \mathbf{Z}_1^\top \mathbf{H}_1 \mathbf{H}_1^\top \mathbf{Z}_1 = [\mathbf{X}_1, \dots, \mathbf{X}_{p_2}] \begin{bmatrix} \mathbf{X}_1^\top \\ \vdots \\ \mathbf{X}_{p_2}^\top \end{bmatrix}$$

4. Additional topics of Wishart Distribution (conditional wishart (4))

Proof (continued)

$$\Rightarrow W_{11 \cdot 2} = Z_1^\top Z_1 - Z_1^\top H_1 H_1^\top Z_1 = \sum_{i=p_2+1}^n X_i X_i^\top \sim W_{p_1}(n - p_2, \Sigma_{11 \cdot 2})$$

Since $W_{11 \cdot 2}$ does not depend on Z_2 , and $W_{12} = Z_1 H_1 (Z_2^\top Z_2)^{1/2}$,

hence, the conditional distributions of $W_{11 \cdot 2}$ and $W_{12}|W_{22}$ are independent, which implies $W_{11 \cdot 2} \perp\!\!\!\perp \{W_{21}, W_{22}\}$.

Now, All we need to check is : $X_i \stackrel{iid}{\sim} N_{p_1}(0, \Sigma_{11 \cdot 2})$

$\Rightarrow Z_{i1}|Z_{i2} \sim N_{p_1}(\Sigma_{12}\Sigma_{22}^{-1}Z_{i2}, \Sigma_{11 \cdot 2})$. For checking above, we need another theorem.

4. Additional topics of Wishart Distribution (conditional wishart (5))

Theorem 3

Let $X = (X_1, \dots, X_n)^\top$ be an $n \times p$ random matrix whose rows are independent normal variates with mean $E(X) = M$ and the same covariance matrix Σ . Let $Y = HX$ be the random matrix obtained by an $n \times n$ orthogonal transformation H . Then the rows of Y are independent normal variates with mean $E(Y) = HM$ and the same covariance matrix Σ .

Proof (SKIP)

4. Additional topics of Wishart Distribution (conditional wishart (6))

Proof (continued)

If $X_i \stackrel{iid}{\sim} N_{p_1}(0, \Sigma_{11.2})$, then $\sum_{i=p_2+1}^n X_i X_i^\top \sim W_{p_1}(n - p_2, \Sigma_{11.2})$.

We need to check for distribution of rows of $H_2^\top Z_1$. Since this is too difficult, instead, we will think of rows of $X = H_2^\top Z_1 | Z_2$.

\Rightarrow rows of $H_2^\top Z_1 | Z_2 : Z_{p_2+1}, \dots, Z_n \stackrel{iid}{\sim} N_{p_1}(H_2^\top E(Z_1 | Z_2), \Sigma_{11.2})$
(\because Since H_2 is orthonormal, $H_2^\top \Sigma_{11.2} H_2 = \Sigma_{11.2}$)
 $\Rightarrow H_2^\top E(Z_1 | Z_2) = H_2^\top Z_2 \Sigma_{22}^{-1} \Sigma_{21} = 0$. ($\because H_2^\top Z_2 = 0$)
 $\therefore X_i \stackrel{iid}{\sim} N_{p_1}(0, \Sigma_{11.2})$.

4. Additional topics of Wishart Distribution (conditional wishart (7))

Proof (continued)

Since $W_{12} = Z_1^\top Z_2 = Z_1^\top H_1 (Z_2^\top Z_2)^{1/2}$, $W_{11 \cdot 2} = (Z_1^\top H_2)(Z_1^\top H_2)^\top$,
And we know that $H_1^\top H_2 = 0 \rightarrow \therefore W_{11 \cdot 2} \perp\!\!\!\perp W_{12} | Z_2, W_{11 \cdot 2} \perp\!\!\!\perp W_{22}$
Also,

$$E[W_{12} | Z_2] = E[Z_1^\top | Z_2] Z_2 = [E(Z_{11} | Z_2), \dots, E(Z_{n1} | Z_2)] Z_2 \quad (16)$$

$$= \Sigma_{12} \Sigma_{22}^{-1} Z_2^\top Z_2 = \Sigma_{12} \Sigma_{22}^{-1} W_{22} \quad (17)$$

We skip the covariance part.

- Johan Lim : Lecture note of multivariate statistics
- Y. Fujikoshi, V.V. Ulyanov, R. Shimizu : Multivariate statistics - High dimensional and large sample approximations
- T.W. Anderson : An introduction to multivariate statistical analysis 3rd ed.
- R.Johnson, D.Wichern : Applied multivariate staitistical analysis 6th ed.

Thank you