



**1.2)** Dil, Emma decides to meet at the bus stop and take a bus to JP. Dil comes to the bus stop at time  $t = 0$  and waits for Emma.  $p$  179A buses arrive between 0 and 10 minutes. Emma finally comes to the bus stop at time  $t = 10$  and finds a visibly angry Dil waiting. Dil argues that the probability (say,  $P_1$ ) of at least one 179A arriving in the next 5 minutes, given that he has already seen  $p$  179A arriving in the past 10 minutes, is less than the usual probability of at least one 179A arriving in the next 5 minutes (say,  $P_2$ ). Emma feels otherwise. Dil and Emma are comparing  $P_1 =$  the conditional probability of greater than  $p$  arrivals in 0 to 15 minutes given there were  $p$  arrivals in 0 to 10 minutes  $= P\{\mathbf{x}(15) > p \mid \mathbf{x}(10) = p\}$ , with  $P_2 =$  the probability of at least 1 arrival in 10 to 15 minutes  $= P\{\mathbf{x}(15) - \mathbf{x}(10) \geq 1\}$ .

(1.2.a) Compare  $P\{\mathbf{x}(15) > p \mid \mathbf{x}(10) = p\}$  and  $P\{\mathbf{x}(15) - \mathbf{x}(10) \geq 1\}$  to see if Dil is correct ( $P_1 < P_2$ ), or if Emma is correct ( $P_1$  not less than  $P_2$ ). Show your steps.

(10 marks)

(1.2.b) Evaluate  $P_2 = P\{\mathbf{x}(15) - \mathbf{x}(10) \geq 1\}$  for  $\lambda = 0.2$  arrival/minute.

$$P\{\mathbf{x}(15) - \mathbf{x}(10) \geq 1\} = \boxed{\phantom{0.0000}}$$

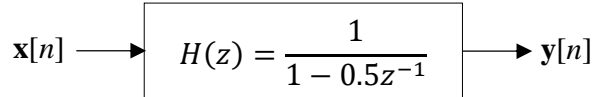
(5 marks)

(5 marks)

(2.d) Unfortunately Mr. Vestor never took EE7401 or any similar course, and failed to appreciate orthogonality. However, he understands what correlated means (recall that you already explained correlation to him in part 1.a). To convince him, show that for a given  $t_1$ , the random variables  $\mathbf{x}(t_1)$  and  $\mathbf{x}'(t_1)$  are uncorrelated.

(5 marks)

**Q3) When did you start your simulations:** Consider the simulation of a first order auto-regressive, or AR(1), process. A wide-sense stationary real white noise process  $\mathbf{x}[n]$  with autocorrelation  $R_{xx}[m] = 5\delta[m]$  is passed through the AR(1) filter with  $a = 0.5$ , such that the output is  $\mathbf{y}[n] = \mathbf{x}[n] + 0.5\mathbf{y}[n - 1]$ .

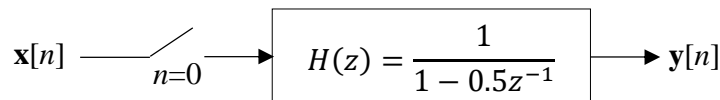


**3.1)** (3.1.a) Assume the above filter operates at all times. Then  $\mathbf{y}[n]$  is jointly wide-sense stationary. Find the cross-correlation between  $\mathbf{x}[n]$  and  $\mathbf{y}[n]$ ,  $R_{xy}[m] = E\{\mathbf{x}[n_1]\mathbf{y}[n_2]\}$ , where  $m = n_1 - n_2$ . [Hint: There are many ways to solve this problem. You may use the frequency domain or the  $z$  domain approach, finding  $S_{xy}(\omega)$  or  $\mathbf{S}_{xy}(z)$  first, and then taking the inverse Fourier or  $z$  transform. You may express  $\mathbf{y}[n]$  as a convolution of  $\mathbf{x}[n]$  with  $h[n]$  (impulse response), multiply the equation by  $\mathbf{x}[n_1]$ , and take the expectation.]

$$R_{xy}[m] = \begin{cases} \boxed{\phantom{0}} & \text{for } m > 0 \\ \boxed{\phantom{0}} & \text{for } m \leq 0 \end{cases} \quad (4 \text{ marks})$$

(3.1.b) Find the autocorrelation of the AR(1) process,  $R_{yy}[m]$ .

$$R_{yy}[m] = \boxed{\phantom{0}} \quad (2 \text{ marks})$$



**3.2)** Part (3.1) results are true if the simulation is started at  $n = -\infty$ . However, real-life simulations start at finite times. Consider a real-life simulation of the same AR(1) process as above, where the simulation starts at  $n = 0$ .  $\mathbf{x}[n]$  is still the same wide-sense stationary with the same autocorrelation. However, the AR(1) filter starts at  $n = 0$  (meaning, there was no filter before  $n = 0$ , or the output was zero). Therefore, the output becomes  $\mathbf{y}[n] = \begin{cases} 0 & n < 0 \\ \mathbf{x}[n] + 0.5\mathbf{y}[n - 1] & n \geq 0 \end{cases}$ . Note that  $\mathbf{y}[n]$  is no longer stationary. Therefore, the auto/cross-correlations involving  $\mathbf{y}[n]$  no longer depend on the time difference  $m$  but depend on both times, like  $R_{yy}[n_1, n_2]$ . As a result, the power spectrums of  $\mathbf{y}[n]$  do not exist, and the power spectrum based approaches can no longer be used to find the auto/cross-correlations. The time-domain approach may still be used.

(3.2.a) Express  $\mathbf{y}[0]$  using only the input  $\mathbf{x}[0]$ . Express  $\mathbf{y}[1]$  as a sum of only input terms of the form  $\mathbf{x}[k]$ . There should not be any past output term such as  $\mathbf{y}[n - 1]$ . Continuing as above, express  $\mathbf{y}[n_2]$  for any  $n_2 \geq 0$  as a sum of only input terms of the form  $\mathbf{x}[k]$ . There should not be any past output term such as  $\mathbf{y}[n_2 - 1]$ .

$$\mathbf{y}[n_2] = \boxed{\phantom{\sum_{k=0}^{n_2} \mathbf{x}[k]}} \quad \text{for } n_2 \geq 0 \quad \dots \text{eq.(1)}$$

(3 marks)

(3.2.b) Find the cross-correlation between  $\mathbf{x}[n]$  and  $\mathbf{y}[n]$ ,  $R_{xy}[n_1, n_2] = E\{\mathbf{x}[n_1]\mathbf{y}[n_2]\}$ , for all  $n_1, n_2 \geq 0$  by multiplying eq.(1) by  $\mathbf{x}[n_1]$ , and taking the expectation of both sides. [Hint: You should find 2 cases: write answers in 2 left boxes, and write the cases in 2 right boxes.]

$$R_{xy}[n_1, n_2] = \begin{cases} \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} & \text{for } n_1 \geq 0, n_2 \geq 0, \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} \\ \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} & \text{for } n_1 \geq 0, n_2 \geq 0, \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} \end{cases}$$

(4 marks)

(3.2.c) Extend your result of (3.2.b) to all possible  $n_1, n_2$  values. [Hint: You should find 2 cases: write answers in the 2 left boxes, and write the condition of the primary case in the right box.]

$$R_{xy}[n_1, n_2] = \begin{cases} \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} & \text{for } n_1 \geq 0, n_2 \geq 0, \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} \\ \boxed{\phantom{\sum_{k=0}^{n_1} \mathbf{x}[k] \mathbf{y}[n_2]}} & \text{otherwise} \end{cases}$$

(2 marks)

(3.2.d) Find the autocorrelation of the AR(1) process,  $R_{yy}[n_1, n_2]$  for all possible  $n_1, n_2$  values. [Hint: Replace both  $y[n_1]$  and  $y[n_2]$  in  $E\{y[n_1]y[n_2]\}$  by eq.(1) twice, then take the expectation of this double summation. You should find 3 cases (it is possible to combine 2 primary cases into a single case): write answers in the 3 left boxes, and write the conditions of 2 primary cases in the 2 right boxes.]

$$R_{yy}[n_1, n_2] = \begin{cases} \boxed{\phantom{0}} & \text{for } n_1 \geq 0, n_2 \geq 0, \boxed{\phantom{0}} \\ \boxed{\phantom{0}} & \text{for } n_1 \geq 0, n_2 \geq 0, \boxed{\phantom{0}} \\ \boxed{\phantom{0}} & \text{otherwise} \end{cases}$$

(6 marks)

(3.2.e) For all non-negative  $n_1$  and  $n_2$ , is  $R_{xy}[n_1, n_2]$  of (3.2.c) equal to  $R_{xy}[m]$  of (3.1.a)?

If not, when are they equal?

For all non-negative  $n_1$  and  $n_2$ , is  $R_{yy}[n_1, n_2]$  of (3.2.d) equal to  $R_{yy}[m]$  of (3.1.b)? If not, when are they equal?

(4 marks)

**Q4) Bandlimited process does change with time:** In lecture we have upper bounded the change in value of a bandlimited process over a small time  $\tau$ . Here we obtain a lower bound (and a new upper bound). First, two intermediate results are obtained as below.

**4.1)** Assume that the time is bounded by  $|\tau| < (\pi/\sigma)$  and that the frequency is bounded by  $|\omega| \leq \sigma$ . Then, using the fact that if  $0 < \varphi < (\pi/2)$ , then  $(2\varphi/\pi) < \sin \varphi < \varphi$ , find a lower bound and an upper bound on  $\sin^2(\omega\tau/2)$  :

$$\boxed{\phantom{0}} \leq \sin^2(\omega\tau/2) \leq \boxed{\phantom{0}} \quad (4 \text{ marks})$$

**4.2)** Let the power spectrum of  $\mathbf{x}(t)$  be  $S_{xx}(\omega)$ . If  $\mathbf{x}(t)$  is passed through a differentiator (frequency response  $H(\omega) = j\omega$ ), then you have already found  $S_{x'x'}(\omega)$ , the power spectrum of the output  $\mathbf{x}'(t)$ , in terms of  $S_{xx}(\omega)$ , in Q(2.b). Copy this  $S_{x'x'}(\omega)$  below to obtain the autocorrelation of the output  $\mathbf{x}'(t)$  as

$$R_{x'x'}(\tau) = \int_{-\infty}^{\infty} \boxed{\phantom{0}} S_{xx}(\omega) e^{j\omega\tau} d\omega/2\pi$$

Now, putting  $\tau = 0$ , the average power of the output  $\mathbf{x}'(t)$  is:

$$E\{|\mathbf{x}'(t)|^2\} = \int_{-\infty}^{\infty} \boxed{\phantom{0}} S_{xx}(\omega) d\omega/2\pi \quad (2 \text{ marks})$$

**4.3)** (4.3.a) Express the expectation of the square of the change in  $\mathbf{x}(t)$  over time  $\tau$ ,  $E\{|\mathbf{x}(t+\tau) - \mathbf{x}(t)|^2\}$ , using its autocorrelation:

$$E\{|\mathbf{x}(t+\tau) - \mathbf{x}(t)|^2\} = 2R_{xx}(\boxed{\phantom{0}}) - R_{xx}(\boxed{\phantom{0}}) - R_{xx}(\boxed{\phantom{0}}) \quad (5 \text{ marks})$$

(4.3.b) Use (4.3.a),  $R_{xx}(\tau) = \int_{-\infty}^{\infty} S_{xx}(\omega) e^{j\omega\tau} \frac{d\omega}{2\pi}$ , and  $1 - \cos\theta = 2\sin^2(\theta/2)$ , to get:

$$E\{|\mathbf{x}(t+\tau) - \mathbf{x}(t)|^2\} = \int_{-\infty}^{\infty} S_{xx}(\omega) \boxed{\phantom{0}} d\omega/2\pi \quad (4 \text{ marks})$$



(4.3.c) Let us say the integral you obtained in part (4.3.b) is  $\int_{-\infty}^{\infty} S_{xx}(\omega)g(\omega)d\omega/2\pi$  for some function  $g(\omega)$  that you wrote inside the box. Now, since  $\mathbf{x}(t)$  is bandlimited, its power spectrum  $S_{xx}(\omega) = 0$  for  $|\omega| > \sigma$ . Therefore, this integral's limits may be changed,  $E\{|\mathbf{x}(t + \tau) - \mathbf{x}(t)|^2\} = \int_{-\infty}^{\infty} S_{xx}(\omega)g(\omega)d\omega/2\pi = \int_{-\sigma}^{\sigma} S_{xx}(\omega)g(\omega)d\omega/2\pi$ . Apply the lower and upper bounds on  $\sin^2(\omega\tau/2)$  from part (4.1) to obtain the lower and upper bounds on the expectation:

$$\int_{-\sigma}^{\sigma} S_{xx}(\omega) \omega^2 \boxed{\phantom{000000}} \frac{d\omega}{2\pi} \leq E\{|\mathbf{x}(t + \tau) - \mathbf{x}(t)|^2\} \leq \int_{-\sigma}^{\sigma} S_{xx}(\omega) \omega^2 \boxed{\phantom{000000}} \frac{d\omega}{2\pi}$$

(4 marks)

(4.3.d) The difficulty is that, unlike in the lecture, we can't replace  $\omega$  by  $\sigma$  in the lower bound. Therefore, we need to evaluate  $\int_{-\sigma}^{\sigma} S_{xx}(\omega)\omega^2 d\omega/2\pi$ . This has already been done in part (4.2) using the differentiated process  $\mathbf{x}'(t)$ . Use the result of (4.2) on the lower and upper bounds of (4.3.c) to obtain the final result:

$$\boxed{\phantom{000000}} E\{|\mathbf{x}'(t)|^2\} \leq E\{|\mathbf{x}(t + \tau) - \mathbf{x}(t)|^2\} \leq \boxed{\phantom{000000}} E\{|\mathbf{x}'(t)|^2\}$$

(2 marks)

(4.3.e) Is the upper bound of (4.3.d) smaller, or larger, than the upper bound found in the class?

(4 marks)