

$$Y = (X - \mu)^2$$

$$P(Y \geq a^2) \leq \frac{E[Y]}{a^2} = \frac{\text{Var}(X)}{a^2}$$

Aus $a = k\sigma$

$$P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

(Ineq Chernoff)

Für X v.a., $a > 0$, $t > 0$

$$P(X \geq a) \leq \frac{E[e^{tX}]}{e^{ta}} \quad \forall a, t$$

Cours 14

Inequalities

$$1) E[XY] \leq \sqrt{E[X^2]E[Y^2]} : \text{Cauchy-Schwarz}$$

$$2) \begin{array}{l} \text{f. sammelv} \\ \text{f. konvex} \end{array} E[\varphi(x)] \geq \varphi(E[x]) \quad \text{Jensen}$$

$$E[\varphi(x)] \leq \varphi(E[x])$$

$$3) X > 0, a > 0$$

$$P(X > a) \leq \frac{E[X]}{a} \quad \text{Markov}$$

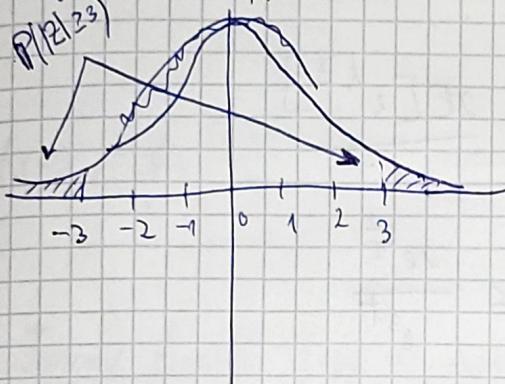
$$4) X \text{ v.a. } E[X] = \mu, \text{ Var}(X) = \sigma^2$$

$$P(|X - \mu| \geq a) \leq \frac{\text{Var}(X)}{a^2}, \forall a > 0 \quad \text{(Chebyshev)}$$

$$5) X \text{ v.a. }, a > 0, t > 0 \text{ atunii } P(X \geq a) \leq \frac{E[e^{tX}]}{e^{ta}}, \forall t > 0 \quad \text{(Chernoff)}$$

Ex:

$Z \sim N(0, 1)$. Vom zu marginum superior $P(|Z| \geq 3)$ folgendes Nachweise, Chebyshev, Chernoff



Anwendung 68 - 95 - 99,7

$$P(|Z| \leq 1) \approx 0.68$$

$$P(|Z| \leq 2) \approx 0.95$$

$$P(|Z| \leq 3) \approx 0.997$$

Aufgabe: $P(|Z| \geq 3) \approx 0.003$

a) Marksche

$$P(|Z| \geq 3) \leq \frac{E[|Z|]}{3} \leq \sqrt{\frac{2}{\pi}}$$

$$t^x - \frac{x^2}{2} = \\ = x(t - \frac{x}{2})$$

$$\begin{aligned} E[|Z|] &= \int |z| \varphi(z) dz \\ &= \int |z| \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz \\ &= 2 \int_0^\infty \frac{z}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = \sqrt{\frac{2}{\pi}} \int_0^\infty z e^{-\frac{z^2}{2}} dz = \sqrt{\frac{2}{\pi}} \left(-e^{-\frac{z^2}{2}} \right) \Big|_0^\infty = \sqrt{\frac{2}{\pi}} \\ &= 0.26 \end{aligned}$$

b) $P(|Z| \geq 3) = P(|Z-0| \geq 3) \leq \frac{\text{Var}(Z)}{9} = \frac{1}{9} = 0.11$

c) Chernoff

$$P(|Z| \geq 3) = 2P(Z \geq 3) \leq \frac{E[e^{tZ}]}{e^{3t}}, t > 0$$

$$E[e^{tZ}] = \int_{-\infty}^{+\infty} e^{tx} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2} + tx} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{(x-t)^2}{2} + t^2} dx =$$

$$= \frac{1}{\sqrt{2\pi}} \int e^{-\frac{(x-t)^2}{2} + \frac{t^2}{2}} dx = \frac{e^{\frac{t^2}{2}}}{\sqrt{2\pi}} \int e^{-\frac{(x-t)^2}{2}} dt = e^{\frac{t^2}{2}}$$

Var $\rightarrow x \sim N(t, 1)$

$$P(|Z| \geq 3) = 2P(Z \geq 3) \leq \frac{2E[e^{tZ}]}{e^{3t}}$$

$$P(|Z| \geq 3) \leq 2e^{\frac{t^2}{2} - 3t} \leq \frac{2e^{\frac{t^2}{2}}}{e^{3t}}$$

Pt. $t=3$ ansem

$$P(|Z| \geq 3) \leq 2e^{-\frac{9}{2}} = 0.02$$

GEP: $X \in [a, b]$ $E[X] = \mu$, $\text{Var}(X) = \sigma^2$

$$\begin{aligned} P(|X-\mu| \geq t) &\leq \frac{\sigma^2}{t^2} \quad t > 0 \\ &\leq \frac{(b-a)^2}{4t^2} \end{aligned}$$

mechanosent

$$\sigma^2 \leq \frac{(b-a)^2}{4}$$

$$g(\gamma) = E[(x-\gamma)^2]$$

$$\gamma = E[X]$$

$$E[(x-\gamma)^2] \geq E[(x-E[X])^2], \quad \forall \gamma \in \mathbb{R}$$

$$\begin{aligned} E[(x-\gamma)^2] &= E[(x - E[X] + E[X] - \gamma)^2] \\ &= E[(x - E[X])^2] + 2E[(x - E[X])(E[X] - \gamma)] + \\ &\quad + (E[X] - \gamma)^2 \end{aligned}$$

$\overbrace{}^n \quad \overbrace{}^v$

$$\sigma^2 = \text{Var}(X) = E[(X - E[X])^2] \leq E[(X - \bar{x})^2], \quad \forall \bar{x}$$

$$\text{Pr} \quad \bar{x} = \frac{a+b}{2} \quad E\left[\left(X - \frac{a+b}{2}\right)^2\right] = E\left[\left(X-a\right)\left(X-b\right) + \frac{(b-a)^2}{4}\right] \\ = E\left[\underbrace{\left(X-a\right)\left(X-b\right)}_{\leq 0}\right] + \frac{(b-a)^2}{4} \leq \frac{(b-a)^2}{4}$$

$$(X-\bar{x})^2 = \underbrace{A}_{\leq 0} + \frac{(b-a)^2}{4}$$

Teoreme limită. Legea Mării

Def.: Fie $(X_n)_{n \geq 1}$ un sir de v.a. și X o v.a. pe spațiu (Ω, \mathcal{F}, P)

Să spunem că sirul $(X_n)_n$ convergență la X dacă și numai dacă

$$X_n \xrightarrow{\text{a.s.}} X$$

$$\text{daca } P\left(\lim_n X_n = X\right) = 1$$

un eveniment

$$A = \left\{ \omega \in \Omega \mid \lim_n X_n(\omega) = X(\omega) \right\}$$

Def.:

Fie $(X_n)_{n \geq 1}$ un sir de v.a. și X o v.a. def. (Ω, \mathcal{F}, P)

Să spunem că sirul X_n convergență probabilă la X și numai

$$X_n \xrightarrow{P} X$$

daca $\forall \varepsilon > 0$

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \varepsilon) = 0$$

Aleg., $\forall \varepsilon > 0, \forall \delta > 0, \exists n_0 \in \mathbb{N} \text{ a. s. } n \geq n_0$

$$P(|X_n - X| \geq \varepsilon) \leq \delta \leftarrow \text{măsură de invecindere}$$

Exp: $X_m \sim U([0,1])$ indep

$$Y_m = \min \{X_1, X_2, \dots, X_m\}$$

Ast. $Y_m \xrightarrow{P} 0$

Pentru $\varepsilon > 0$

$$P(|Y_m - 0| \geq \varepsilon) \xrightarrow{m \rightarrow \infty} 0$$

$$P(|Y_m| \geq \varepsilon) = P(Y_m \geq \varepsilon) \quad (\text{pt că } Y_m \geq 0)$$

$$= P(X_1 \geq \varepsilon, X_2 \geq \varepsilon, \dots, X_m \geq \varepsilon) =$$

$$\stackrel{\text{indep}}{=} P(X_1 \geq \varepsilon) P(X_2 \geq \varepsilon) \dots P(X_m \geq \varepsilon)$$

$$= (1 - P(X_1 < \varepsilon)) (1 - P(X_2 < \varepsilon)) \dots (1 - P(X_m < \varepsilon)) =$$

$$= (1 - \varepsilon) \times (1 - \varepsilon) \dots (1 - \varepsilon) = (1 - \varepsilon)^m$$

pt. $\varepsilon \in [0,1]$

$$P(|Y_m| \geq \varepsilon) = (1 - \varepsilon)^m \quad \text{pt. } \varepsilon \in (0,1)$$

\downarrow
 $m \rightarrow \infty$
0

Def: Numim esantion de volum n din populatia Q , tota rea
 X_1, X_2, \dots, X_n indep. si identic repartizate (i.i.d) in ~~repartita~~ $P[X_i] = Q$

$$\text{Media esantionului } \bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

P.P. X_1, X_2, \dots, X_n esantion de media μ si varianta σ^2
 $(E[X_i] = \mu, \text{Var}[X_i] = \sigma^2)$

$$E[\bar{X}_n] = E\left[\frac{X_1 + \dots + X_n}{n}\right] = \frac{1}{n} (E[X_1] + E[X_2] + \dots + E[X_n]) = \mu$$

$$\text{Var}(\bar{X}_n) = \text{Var}\left(\frac{X_1 + \dots + X_n}{n}\right) = \frac{1}{n^2} (\text{Var}(X_1) + \dots + \text{Var}(X_n)) = \frac{\sigma^2}{n}$$

T (LNM) - Legem Nr. Mari (Legea)

Fie $(X_n)_n$ un sir de v.a. i.i.d cu $E[X_1] = \mu < \infty$, $\text{Var}[X_1] = \sigma^2 < \infty$

Afumci $\bar{X}_n \xrightarrow{\mathbb{P}} \mu$

Versiunea Pare

$(X_n)_n$ v.a. i.i.d $E[|X_1|] < \infty$, $E[X_1] = \mu$

$$\bar{X}_n \xrightarrow{\text{a.s.}} \mu$$

(starea)

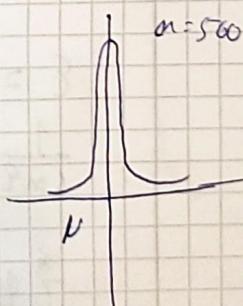
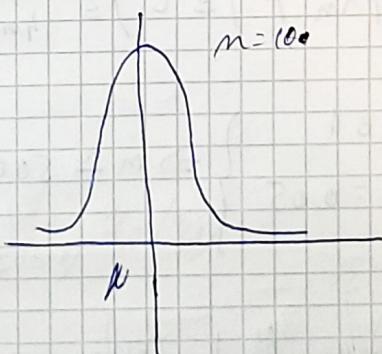
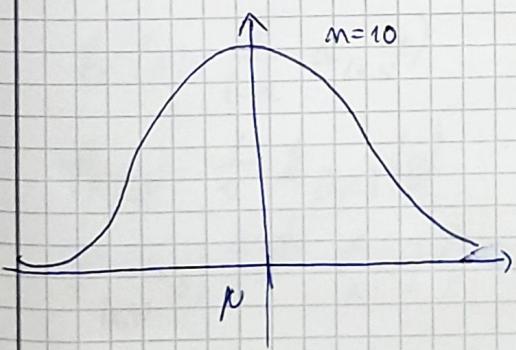
Dem

Pt. $\varepsilon > 0$

$$P(|\bar{X}_n - \mu| \geq \varepsilon) \rightarrow 0$$

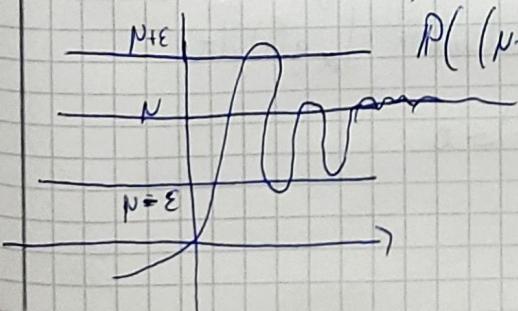
Inegalitatea Chebyshev

$$P(|\bar{X}_n - \mu| \geq \varepsilon) \leq \frac{\text{Var}(\bar{X}_n)}{\varepsilon^2} = \frac{\sigma^2}{n\varepsilon^2} \xrightarrow{n \rightarrow \infty} 0$$



Pt. un nivel de acuritate dat

$$P((\mu - \varepsilon, \mu + \varepsilon) \rightarrow \bar{X}_n) \xrightarrow{n \rightarrow \infty} 1$$



Expo: (Ω, \mathcal{F}, P) c.p., $A \in \mathcal{F}$

$$\text{Fie } X_i = \begin{cases} 1, & w \in A \\ 0, & \text{altfel} \end{cases} \quad X_i \sim B(p) \\ p = P(X_i=1) = P(A)$$

$$\bar{X}_m = \frac{X_1 + X_2 + \dots + X_m}{m} = \text{frecvență relativă de apariție } A \text{ în } m \text{ rep ale exp.}$$

$$\bar{X}_m \xrightarrow{P} E[X_1] = P(A)$$

Expo: Fie P . procentul din pop. care votăză cu A .

$$X_1, X_2, \dots, X_n \sim B(P) \text{ indep.}$$

$$P(1-P) \leq \frac{1}{9}$$

$$\bar{X}_m = \frac{X_1 + X_2 + \dots + X_m}{m} \quad P - \text{neconvenient}$$

$$P - P^2 - \frac{1}{9} \leq 0 \\ -P^2 + P - \frac{1}{9} \leq 0$$

$$P(|\bar{X}_m - P| \geq \varepsilon) \leq \frac{\text{Var}(X_i)}{\varepsilon^2} = \frac{P(1-P)}{m\varepsilon^2} \leq \frac{1}{4m\varepsilon^2}$$

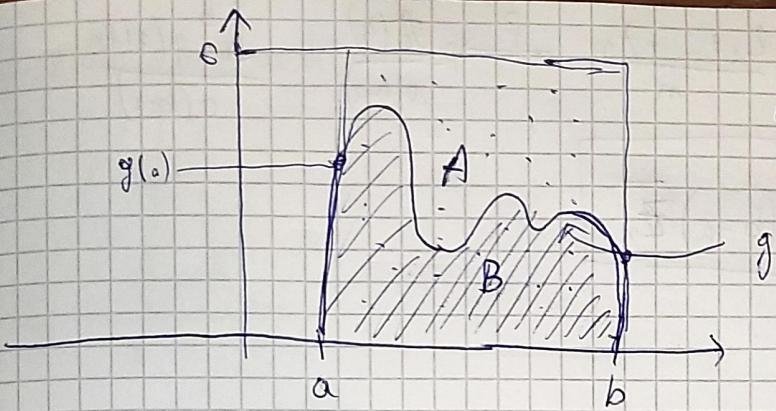
$$P(|\bar{X}_m - P| \geq \varepsilon) \leq \frac{1}{4m\varepsilon^2}$$

$$\left. \begin{array}{l} \varepsilon = 0.01 \\ \frac{1}{4m\varepsilon^2} = 0.05 \end{array} \right\} \Rightarrow m \approx 50000$$

Integrarea Monte-Carlo

P. că avem o fct g și vrem să calculăm $\int_a^b g(x) dx$

P. că pe $[a,b]$ avem $0 \leq g(x) \leq x$



$A = [a, b] \times [0, c]$ - idependent

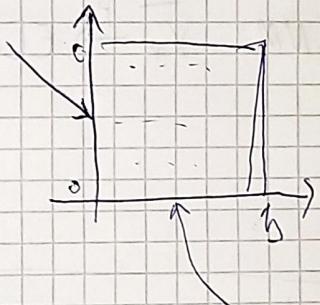
$B = \{(x, y) \mid a \leq x \leq b, 0 \leq y \leq g(x)\}$

Generum pot. unif $\propto (A)$

$(x_1, y_1), \dots, (x_n, y_n) \sim U(A)$

$(x, y) \sim U(A)$

$$f_{(x,y)}(x,y) = \begin{cases} \frac{1}{\text{vol}(A)}, & (x, y) \in A \\ 0, & \text{otherwise} \end{cases}$$



$$\text{vol}(A) = c(b-a)$$

$$X \sim U([a, b])$$

$$Y \sim U([a, c]) \quad \text{indep.}$$

$$f_X(x) = \frac{1}{b-a} \mathbb{1}_{[a,b]}(x)$$

$$f_Y(y) = \frac{1}{c-a} \mathbb{1}_{[a,c]}(y)$$

$$f_{(X,Y)}(x,y) \stackrel{\text{indep}}{=} f_X(x) f_Y(y) = \frac{1}{c(b-a)} \mathbb{1}_A(x,y)$$

$$\text{Die } z_i = \begin{cases} 1, & (x_i, y_i) \in B \\ 0, & \text{otherwise} \end{cases}, \quad z_i \sim B(p)$$

$$p = P(z_i=1) = P((x_i, y_i) \in B)$$

$$= \iint_B f_{(X,Y)}(x,y) dx dy = \frac{\text{vol}(B)}{\text{vol}(A)}$$

$$\text{Dim LNM} \bar{x}_m = \frac{\sum_1^m \bar{x}_m}{m} \xrightarrow{P} p = \frac{f(x)}{n(b-a)} = \frac{\int_a^b g(x) dx}{c(b-a)}$$

$$\boxed{\int_a^b g(x) dx \approx n(b-a) \bar{x}_m}$$

(Kz 2)

Fix $U_1, U_2, \dots, U_m \sim U[a, b]$ - - ind.

$$X_1 = g(U_1), X_2 = g(U_2), \dots, X_m = g(U_m) \quad \text{i.i.d.}$$

$$\begin{aligned} \text{LNM: } \bar{x}_m &\xrightarrow{P} E[X_1] = E[g(U_1)] \\ &= \int g(x) f_{U_1}(x) dx \\ &= \int g(x) \frac{1}{b-a} \mathbb{1}_{[a,b]}(x) dx = \frac{1}{b-a} \int_a^b g(x) dx \end{aligned}$$

$$\boxed{\int_a^b g(x) dx \approx (b-a) \cdot \frac{\sum_1^m x_i}{n}}$$

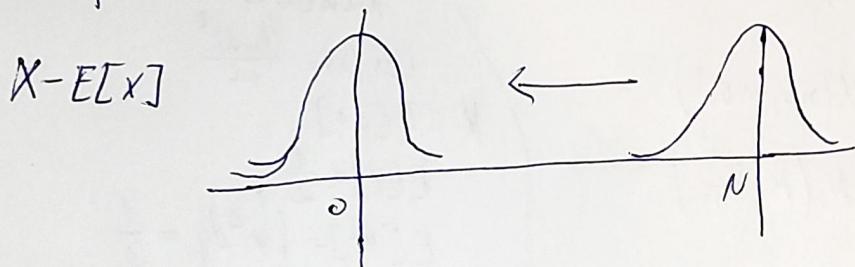
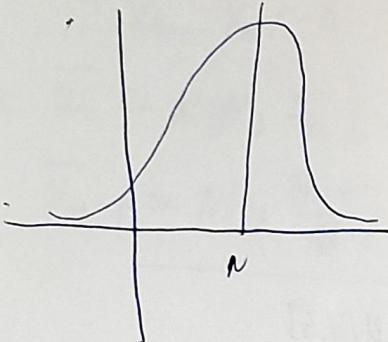
$$\boxed{\int_a^b g(x) dx \approx (b-a) \cdot \frac{g(U_1) + \dots + g(U_n)}{n}} \quad U_i \sim U[a, b]$$

The Limite Centrali (TLC)

$$\text{Dim LNM} \quad \bar{x}_m \xrightarrow{P} E[X_1]$$

Fix X_1, \dots, X_m u.a. i.i.d. $E[X_i] = \mu$, $\text{Var}(X_i) = \sigma^2$

Transf. de locatie si de scala



$$\text{Var}\left(\frac{X - E[X]}{\sqrt{\text{Var}(X)}}\right) = 1$$

$$Z = \frac{X - E[X]}{\sqrt{\text{Var}(X)}} \quad - z \text{ scor} \quad \text{var. normalizata}$$

$$Z_m = \frac{\bar{X}_m - \mu}{\sqrt{\text{Var}(\bar{X}_m)}}$$

$$= \frac{\bar{X}_m - \mu}{\sqrt{m \sigma^2}} \quad - \text{variabila de scor} \\ \text{var. normalizata}$$

$$Z_m = \sqrt{m} \left(\frac{\bar{X}_m - \mu}{\sigma} \right)$$

⑦ (Teorema Limită Centrală)

Fie $(X_n)_n$ un sir r.a. i.i.d de medie $E[\bar{X}_n] = \mu < \infty$, $\text{Var}(X_n) = \sigma^2 < \infty$. Atunci

$$\lim_{n \rightarrow \infty} P(Z_m \leq x) = \Phi(x), \forall x$$

$$\text{unde } Z_m = \sqrt{n} \frac{\bar{X}_n - \mu}{\sigma}, \quad \Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt \quad (\text{fct. de rap a normaliz} \mathcal{N}(0, 1))$$

def: X_1, X_2, \dots, X_m , $E[X_i] = \mu$, $\text{Var}(X_i) = \sigma^2$, $S_m = X_1 + \dots + X_m$

$$\begin{aligned} P(S_m \leq c) &= P\left(\frac{S_m - E[S_m]}{\sqrt{\text{Var}(S_m)}} \leq \frac{c - E[S_m]}{\sqrt{\text{Var}(S_m)}}\right) \\ &= P\left(Z_m \leq \frac{c - \mu N}{\sqrt{N\sigma^2}}\right) \xrightarrow{TLC} \underline{\Phi}\left(\frac{c - \mu N}{\sigma\sqrt{N}}\right) \end{aligned}$$

Pt. n suficient de mare

$$\begin{cases} S_m \sim N(\mu N, N\sigma^2) \\ \bar{X}_m \sim N\left(\mu, \frac{\sigma^2}{n}\right) \end{cases}$$

$X \sim U[0, 5]$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$

$X \sim U[0, 1]$

$$E[Y] = \frac{1}{2}$$

$$E[Y^2] = \int_0^1 y^2 dy = \frac{1}{3}$$

$$\text{Var}(Y) = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}$$

$$X = a + (b-a)y$$

Ex.: 100 pachete
 $U[5, 50]$

Care este proba ca greutatea totală ≥ 3000 kg?

Fie $X_1, \dots, X_{100} \sim U[5, 50]$ (greutățile pachetelor)

$S_{100} = X_1 + \dots + X_{100}$ greutatea totală

$$P(S_{100} \geq 3000) = P\left(\frac{S_{100} - E[S_{100}]}{\sqrt{\text{Var}(S_{100})}} \geq \frac{3000 - E[S_{100}]}{\sqrt{\text{Var}(S_{100})}}\right) \xrightarrow{TLC} 1 - \underline{\Phi}\left(\frac{3000 - E[S_{100}]}{\sqrt{\text{Var}(S_{100})}}\right)$$

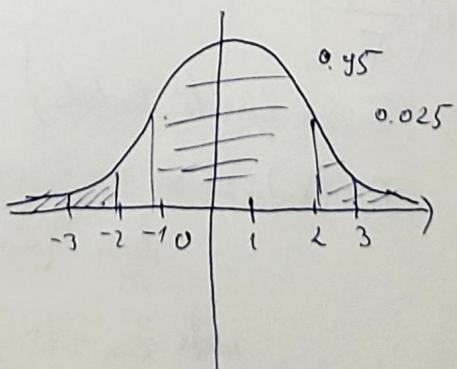
$Z_{100} \sim N(0, 1)$

$$E[S_{100}] = 100 E[X_1] = 100 \cdot \frac{5+50}{2} = 27.5 \cdot 100 = 2750$$

$$\text{Var}(S_{100}) = 100 \text{Var}(X_1) = 100 \cdot \frac{(50-5)^2}{12} = 16875$$

$$\begin{aligned} P(S_{100} \geq 3000) &\approx 1 - \underline{\Phi}\left(\frac{3000 - 2750}{\sqrt{16875}}\right) \\ &\approx 1 - \underline{\Phi}(1.92) = 0.0274 \end{aligned}$$

$$1 - \text{pnorm}(1.92)$$

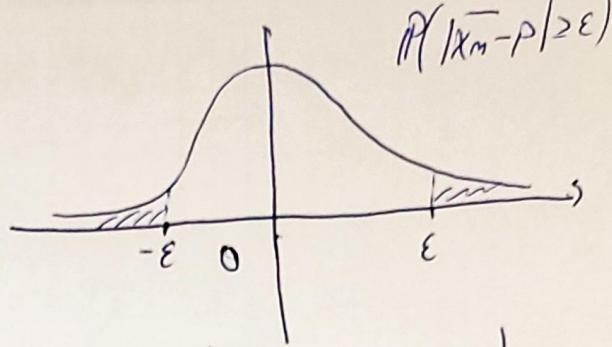


Exp: $P \in A$ in populatie

$$X_1, \dots, X_n \sim B(P)$$

$$\text{TL: } \bar{X}_n \sim N(\mu, \frac{\sigma^2}{n}) = N(P, \frac{P(1-P)}{n})$$

$$\bar{X}_n - P \sim N(0, \frac{P(1-P)}{n})$$



$$P(|\bar{X}_n - P| > \epsilon) = 2P(\bar{X}_n - P > \epsilon)$$

$$(P) \quad \bar{X}_n - P \sim N(0, \frac{P(1-P)}{n})$$

$$P(|\bar{X}_n - P| > \epsilon) \approx 2P(\bar{X}_n - P > \epsilon)$$

$$\approx 2P\left(\frac{\bar{X}_n - P}{\sqrt{\frac{P(1-P)}{n}}} \geq \frac{\epsilon}{\sqrt{\frac{P(1-P)}{n}}}\right)$$

$$\stackrel{\text{TL}}{=} 2\left(1 - \Phi\left(\frac{\epsilon}{\sqrt{\frac{P(1-P)}{n}}}\right)\right) \leq 2\left(1 - \Phi\left(2\epsilon\sqrt{n}\right)\right)$$

$$P(1-P) \leq \frac{1}{4} \Rightarrow \frac{P(1-P)}{n} \leq \frac{1}{4n} \Rightarrow \frac{1}{\sqrt{\frac{P(1-P)}{n}}} \geq \frac{1}{\sqrt{\frac{1}{4n}}} \Rightarrow$$

$$\Rightarrow \frac{\epsilon}{\sqrt{\frac{P(1-P)}{n}}} \geq 2\epsilon\sqrt{n}$$

$$\Rightarrow \Phi\left(\frac{\epsilon}{\sqrt{\frac{P(1-P)}{n}}}\right) \geq \Phi(2\epsilon\sqrt{n})$$

$$P(|\bar{X}_n - P| \geq \epsilon) \leq 2\left(1 - \Phi(2\epsilon\sqrt{n})\right)$$

$$\epsilon = 0.01$$

$$0.05$$

$$2\left(1 - \Phi(2\epsilon\sqrt{n})\right) = 0.05 \quad \epsilon = 0.01 \Rightarrow \Phi(2 \times 0.01\sqrt{n}) = 0.975$$

$$2 \times 0.01\sqrt{n} = \underbrace{\Phi^{-1}(0.975)}_{\approx 2}$$

$$\sqrt{n} \approx 100$$

$$n \approx 10000$$