

Media și momentele de ordin superior

Def: Fie $(\mathcal{R}, \mathcal{F}, P)$ un câmp de probabilitate și $X: \mathcal{R} \rightarrow \mathbb{R}$ v.a. discretă, definim media

$$\mathbb{E}[X] = \sum_x x \cdot P(X=x) = \sum_x x \cdot f(x)$$

ori de căte ori $\sum_x x \cdot f(x) < \infty$. În cazul în care seria este ∞ atunci spunem că v.a. X nu are medie.

Ex: Aruncăm cu un zar

$$X \in \{1, 2, 3, 4, 5, 6\}$$

$$P(X=x) = \frac{1}{6}, \text{ deci } \mathbb{E}[X] = \sum_x x \cdot P(X=x) = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6} = 3,5$$

$$\text{Ex: } X \sim \begin{pmatrix} x_1 & x_2 & \dots & x_n \\ p_1 & p_2 & \dots & p_n \end{pmatrix}$$

$$\mathbb{E}[X] = x_1 p_1 + x_2 p_2 + \dots + x_n p_n \Rightarrow \mathbb{E}[X] = -\frac{1}{2} + \frac{1}{4} = -\frac{1}{4}$$

Interpretare fizică



$$\begin{aligned} xM &= x_1 \cdot m_1 + x_2 \cdot m_2 \\ \bar{x} &= x_1 \cdot \frac{m_1}{M} + x_2 \cdot \frac{m_2}{M} \end{aligned}$$

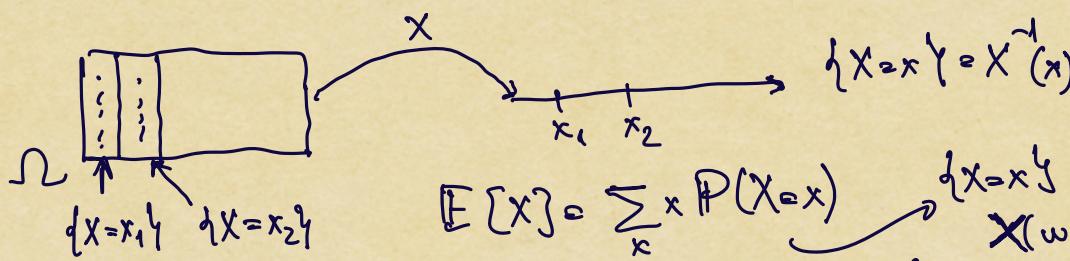
Proprietăți: 1) Dacă X este constantă \Rightarrow media este aceea constantă ($X=c \Rightarrow \mathbb{E}[X]=c$)

2) Dacă $X \geq 0$ atunci $\mathbb{E}[X] \geq 0$ (pozitivitate)

3) Dacă $X \geq Y$ atunci $\mathbb{E}[X] \geq \mathbb{E}[Y]$ (monotonie)
($X(\omega) \geq Y(\omega) \quad \forall \omega \in \mathcal{R}$)

4) Dacă X și Y v.a. discrete și $a, b \in \mathbb{R}$ atunci
 $\mathbb{E}[aX+bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$ (linearitate)

Dem: $\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$



$$\begin{aligned} \mathbb{E}[X] &= \sum_x x \cdot P(X=x) \\ &\quad \xrightarrow{\{X=x\} = \{w \in \mathcal{R} | X(w)=x\}} \\ &\quad \Rightarrow P(X=x) = \sum_w P(X(w)=x) \end{aligned}$$

Ex: Aruncăm cu o monedă de 10 ori

$X = \{nr\ de H\ din cele 10 aruncări\}$

$\Omega = \{H, T\}^{10} \quad X \sim B(10, p)$

$P(X=k) = \binom{10}{k} p^k (1-p)^{10-k}$

$\{X=k\} = \{(w_1, \dots, w_{10}) \mid w_i \in \{H, T\}\}$
 și exact
 k sunt H

Deci, $E[X+Y] = \sum_{\omega} (x+y) P(\{\omega\}) = \sum_{\omega} (x(\omega) + y(\omega)) P(\{\omega\})$

$= \sum_{\omega} x(\omega) P(\{\omega\}) + \sum_{\omega} y(\omega) P(\{\omega\})$

b) (degăsirea dintre medie și prob.)

Fie $A \in \mathcal{F}$ eveniment

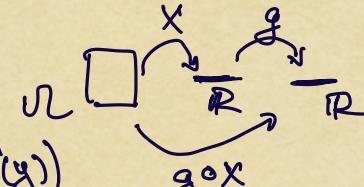
$$\mathbb{1}_A = \begin{cases} 1, & \omega \in A \\ 0, & \omega \notin A \end{cases} \quad P(\mathbb{1}_A = 1) = P(A)$$

$$\mathbb{1}_A \sim \begin{pmatrix} 0 & 1 \\ 1 - P(A) & P(A) \end{pmatrix} \quad E(\mathbb{1}_A) = P(A)$$

c) Fie X v.a discretă și $g: \mathbb{R} \rightarrow \mathbb{R}$ și $y = g(x)$. Arătați:

$$E[g(x)] = \sum_x g(x) P(X=x)$$

$$E[g] = \sum_y y P(Y=y)$$



$$P(Y=y) = P(g(x)=y) = P(X \in g^{-1}(y))$$

$$g^{-1}(y) = \{x \mid g(x)=y\} = \sum_{x \in g^{-1}(y)} P(X=x)$$

$$E[g(x)] = E[Y] = \sum_y y \sum_{x \in g^{-1}(y)} P(X=x) = \sum_x g(x) P(X=x)$$

Ex: $X \sim \begin{pmatrix} -2 & -1 & 1 & 3 \\ 1/4 & 1/8 & 1/2 & 1/8 \end{pmatrix} \quad Y = X^2$ (Pentru exercență!)

Met 1: $y \in \{1, 4, 9\}$

$$P(Y=1) = P(X=-1) + P(X=1) = 1/8 + 1/2 = 5/8$$

$$Y = X^2 \sim \begin{pmatrix} 1 & 4 & 9 \\ 5/8 & 8/8 & 1/8 \end{pmatrix}$$

$$E[X^2] = \frac{5}{8} + \frac{8}{8} + \frac{9}{8} = \frac{22}{8}$$

Met 2: $g(x_1=2)$

$$\mathbb{E}[x^2] = (-2)^2 \cdot \frac{1}{4} + (-1)^2 \cdot \frac{1}{8} + 1^2 \cdot \frac{1}{2} + 3^2 \cdot \frac{1}{8} = \frac{23}{8}$$

*) Fie X și Y 2 v.a. independente. Atunci

$$\mathbb{E}[x \cdot y] = \mathbb{E}[x] \cdot \mathbb{E}[y]$$

Dacă g, h - funcții atunci $g(X)$ și $h(Y)$ sunt independente at.

$$\mathbb{E}[g(x) \cdot h(y)] = \mathbb{E}[g(x)] \cdot \mathbb{E}[h(y)]$$

$$y \sim \begin{pmatrix} 0 & 2 \\ 1/2 & 1/2 \end{pmatrix} \quad X \perp\!\!\!\perp Y$$

$$\mathbb{E}[x^*y^3] = \mathbb{E}[x^*] \cdot \mathbb{E}[y^3]$$

Obs! În general, $\mathbb{E}[x \cdot y] \neq \mathbb{E}[x] \cdot \mathbb{E}[y]$

Def: Fie X o v.a. discretă, numărul moment de ordin $k \geq 1$, ca fiind $\mathbb{E}[X^k]$.

Se numește moment de ordin k centrat în $a \in \mathbb{R}$, $\mathbb{E}[(x-a)^k]$ și moment central de ordin k , $\mathbb{E}[(x-\mathbb{E}[x])^k]$

Def: Varianta sau dispersia v.a. X este momentul central de ordin 2 și se notează cu:

$$\text{Var}(X) = \mathbb{E}[(x-\mathbb{E}[x])^2]$$

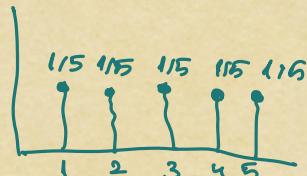
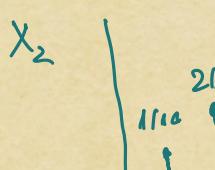
Obs!: Arată gradul de similitudine a des. făță de medie.

$$\text{Ex: } X_1 \sim \begin{pmatrix} 1 & 2 & 3 & 4 & 5 \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \end{pmatrix} \quad X_3 \sim \begin{pmatrix} 1 & 2 & 3 & 4 & 5 \\ 1/2 & 0 & 0 & 0 & 1/2 \end{pmatrix}$$

$$X_2 \sim \begin{pmatrix} 1 & 2 & 3 & 4 & 5 \\ 2/10 & 3/10 & 2/10 & 1/10 & 1/10 \end{pmatrix} \quad X_4 \sim \begin{pmatrix} 1 & 2 & 3 & 4 & 5 \\ 0 & 0 & 1/2 & 0 & 0 \end{pmatrix}$$

$$\mathbb{E}[X_1] = \mathbb{E}[X_2] \approx \mathbb{E}[X_3] = \mathbb{E}[X_4] \approx 3$$

$$\text{Var}(X_1) = \mathbb{E}[(X_1-3)^2] = \frac{(1-3)^2}{5} + \frac{(2-3)^2}{5} + \dots + \frac{(5-3)^2}{5} = \frac{10}{5} = 2$$



$$\text{Var}(X_2) = \mathbb{E}[(X_2-3)^2] = \frac{(1-3)^2}{10} + \frac{(2-3)^2 \cdot 2}{10} + \cancel{\frac{(3-3)^2}{10}} + \frac{(4-3)^2 \cdot 2}{10} + \frac{(5-3)^2}{10} = \frac{12}{10}$$



$$\text{Var}(X_3) = \frac{(1-3)^2}{2} + \frac{(5-3)^2}{2} = \frac{8}{2} = 4$$

Proprietăți ale variantei: 1) Dacă $X = c$ (const.) $\Rightarrow \text{Var}(X) = 0$.

$$2) \text{Var}(X) \geq 0$$

(translație) 3) Dacă X v.a. și $a \in \mathbb{R}$ atunci $\text{Var}(a+X) = \text{Var}(X)$

(scalare) 4) Dacă X v.a. și $b \in \mathbb{R}^*$ atunci $\text{Var}(b \cdot X) = b^2 \cdot \text{Var}(X)$

$$\begin{aligned} \text{Var}(bX) &= E\{(bX - E[bX])^2\} \\ &= E[b^2(X - E[X])^2] \end{aligned}$$

$$5) \text{Var}(a+bX) = b^2 \cdot \text{Var}(X); \quad (\forall) a, b \in \mathbb{R}$$

$$6) \text{Var}(X) = E[X^2] - E[X]^2$$

$$\begin{aligned} E[(X - E[X])^2] &= E[X^2 - 2XE[X] + E[X]^2] \\ &= E[X^2] - 2E[X]E[X] + E[X]^2 \\ &= E[X^2] - 2E[X]^2 + E[X]^2 \end{aligned}$$

7) X și Y v.a. independente

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

Def: Fie X, Y - v.a. se numește covarianta lui X și Y :

$$\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

(dacă este 0, at. X, Y sunt necoreslate)

$$\rightarrow \text{In general, } \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) - 2\text{Cov}(X, Y)$$

Def: Se numește abatere standard:

$$SD(x) = \sqrt{\text{Var}(x)}$$

σ^2 Varianta
(not.)

σ Abatere standard
(not.)

Exemplu de calcul al mediei și variantei

$$(1) X \sim B(p) \Leftrightarrow X \sim \begin{pmatrix} 0 & 1 \\ 1-p & p \end{pmatrix}$$

$$E[X] = p$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = p \cdot p^2 = p(1-p)$$

$$(2) X \sim B(n, p) \quad P(X=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$\mathbb{E}[X] = \sum_{k=0}^n k \cdot P(X=k) = \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n k \cdot \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$$

$\underbrace{\sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} (1-p)^{n-k}}$

$$\Rightarrow \mathbb{E}[X] = n \cdot p$$

$$X = \underbrace{X_1 + \dots + X_n}_{\mathbb{E}[X] = \mathbb{E}[\sum]} = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n]$$

$$\text{Var}(X) = \text{Var}(X_1 + \dots + X_n)$$

$$= \sum_{i=1}^n \text{Var}(X_i) = np(1-p)$$

(3) Hipergeometrică HG(n, N, M)

$$P(X=k) = \frac{\binom{M}{k} \binom{N-M}{n-k}}{\binom{N}{n}}$$

X_j - la extragere j arem bilă neagră $X_j = 1$
bilă albă $X_j = 0$

$$X = X_1 + \dots + X_n$$

$$P(X_j=1) = \frac{M}{N}$$

Extragere fără înlocuire

$$\mathbb{E}[X] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n] = n \cdot \frac{M}{N}$$

(4) Poisson $X \sim \text{Poisson}(\lambda)$ $P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$

$$\mathbb{E}[X] = \sum_{k=0}^{\infty} k \cdot P(X=k) = \sum_{k=0}^{\infty} k \cdot e^{-\lambda} \frac{\lambda^k}{k!}$$

$$= \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} e^{-\lambda} = \lambda$$

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

$$\mathbb{E}[X^2] = \sum_{k=1}^{\infty} k^2 e^{-\lambda} \frac{\lambda^k}{k!} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} k \frac{\lambda^{k-1}}{(k-1)!}$$

$$= \lambda e^{-\lambda} \sum_{k=0}^{\infty} (k+1) \frac{\lambda^k}{k!} =$$

$$= \lambda e^{-\lambda} \sum_{k=0}^{\infty} k \frac{\lambda^k}{k!} + \lambda e^{-\lambda} \underbrace{\sum_{k=0}^{\infty} \frac{\lambda^k}{k!}}_{= \lambda^2 + \lambda} = \lambda^2 + \lambda$$

$$\Rightarrow \text{Var}(X) = \lambda^2 + \lambda - \lambda^2 = \lambda$$

(5) $X \sim \text{Geom}(p)$

$$\begin{aligned} P(X=k) &= (1-p)^{k-1} \cdot p \\ E[X] &= \sum_{k=1}^{\infty} k \cdot p (1-p)^{k-1} = \sum_{k=1}^{\infty} kpq^{k-1} = p \cdot \sum_{k=1}^{\infty} k \cdot q^{k-1} \\ &= p \cdot \sum_{k=1}^{\infty} (q^k)' \\ &= p \cdot \left(\sum_{k=1}^{\infty} q^k \right)' \\ &= p \cdot \left(\frac{q}{1-q} \right)' = \frac{p}{(1-q)^2} \\ &= \frac{1}{p} \end{aligned}$$

$$\text{Var}(X) = \frac{1-p}{p^2}$$

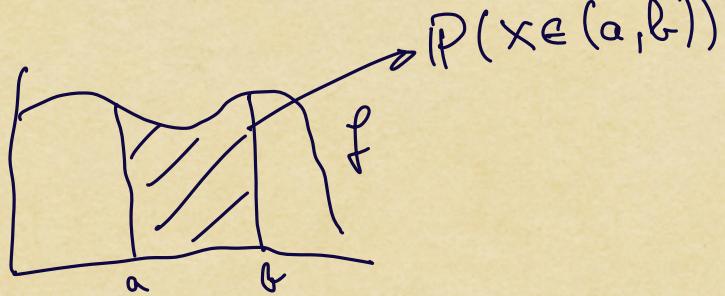
Variabile aleatoare continue

Def: (Ω, \mathcal{F}, P) nu e.p. și $X: \Omega \rightarrow \mathbb{R}$ v.o. V.A. X este continuă (absolut continuă) dacă există $f: \mathbb{R} \rightarrow \mathbb{R}_+$ cu prop.

$$P(x \in A) = \int_A f(x) dx, \quad (\forall) A \subseteq \mathbb{R}$$

Obs: Dacă $A = (a, b)$

$$P(a < X < b) = \int_a^b f(x) dx$$



f - d.m. densitate de repartitie

P: Dacă f e densitate de repartitie

$$1) f \geq 0$$

$$2) \int_{\mathbb{R}} f(x) dx = \int_{-\infty}^{\infty} f(x) dx = 1$$

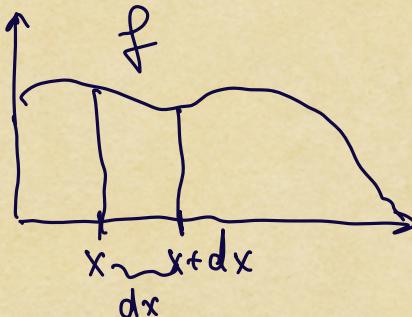
$$P(X \in \mathbb{R}) = P(\Omega) = 1$$

$$\text{Obs! } P(X = a) = \int_a^a f(x) dx = 0.$$

$$A = \{a\}$$

$$\begin{aligned} P(a < X < b) &\approx P(a \leq X \leq b) \\ &= P(a \leq X \leq b) \\ &= P(a \leq X \leq b) \end{aligned}$$

Interpretare:

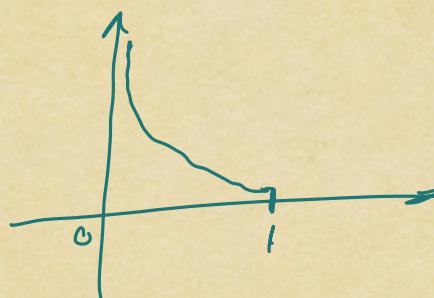


$$\begin{aligned} P(X \in (x, x+dx)) \\ = \int_x^{x+dx} f(t) dt \end{aligned}$$

Dacă dx mic $f(x)(x+dx-x) = dx \cdot f(x)$

$$\Rightarrow f(x) = \frac{P(X \in (x, x+dx))}{dx} \quad (\text{Probabilitatea pe unitatea de lungime})$$

$$\text{Ex: } f(x) = \begin{cases} \frac{1}{2\sqrt{x}}, & 0 < x < 1 \\ 0, & \text{altfel} \end{cases} \quad f(x) \geq 0$$



$$\begin{aligned} \int_{\mathbb{R}} f(x) dx &= \int_{-\infty}^{\infty} f(x) dx \\ &= \int_0^1 \frac{1}{2\sqrt{x}} dx \\ &= 1 \end{aligned}$$

\sum → \int
(discret) (continuu)

$$f(x) = P(X=x) \rightarrow 0$$

$$P(X \in A) = \sum_{x \in A} P(X=x) \rightarrow \int_A f(x) dx$$

$$\int_A f(x) dx$$