

Investigation of Non-normality in a Simple Errors-in-variables Model

STAT3799 Directed studies in statistics

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Abstract

In a classical linear regression setting, we often assume that the explanatory variable is nonrandom without any observation error, and that the errors are normally distributed. However, this may not be the case in real-life applications, where observation errors may exist, and the errors may be heavy-tailed or skewed. We use the computer simulation technique to demonstrate the impacts of non-normality in the *errors-in-variables* model. We present empirical results from simulations based on normal, Student's t and χ^2 distributions on the ordinary least squares and method of moments estimation of regression slope parameter β and residual variance σ_ϵ^2 .

1 Introduction

Consider the problem of regression through the origin with only one explanatory variable:

$$y = \beta x + \epsilon \tag{1}$$

In real life applications, usually we will first obtain pairs of observations $(\tilde{x}_i, \tilde{y}_i)$, then apply a model by performing linear regression on the data. However, we only have observations on the observed $\tilde{x}_i = x_i + u_i$ and $\tilde{y}_i = y_i + v_i$, which have some additive error compared to the true x and y that these model assumptions are based on. We further assume that the measurement errors u_i and v_i have mean zero and constant variances, and that the measurement error is uncorrelated and independent to the true values x and y . This gives rise to the *errors-in-variables model*. This imposes a different problem from the classical linear regression model, because the classical model assumes that the observed \tilde{x} is nonrandom, that we have access to the true value of explanatory variable x without any error.

Furthermore, in the classical linear regression model, we often assume that the observed dependent variable y is subject to some ϵ following $N(0, \sigma^2)$. However, the normality assumption often does not hold for real life datasets. For example, when the errors in the dataset have heavy-tails and/or have skewed shapes, then the normal assumption may not be appropriate. For instance, when dealing with datasets with heavy-tailed errors, one of the practices is to assume t -distributed errors instead of normal-distributed errors, as the light-tailedness of the normal distribution essentially implies that we assume that large errors occur with very low probability, which may not be true in datasets of poorer quality. To make the situation more complicated, in practice non-normality in errors may happen in both the residual ϵ and the observation matter u . Thus, there is a need to investigate the impacts of non-normality on the *errors-in-variables model*.

In this work, we investigate how the non-normality of errors in both the explanatory variable x and the residual of the dependent variable y affects the estimation of the regression coefficient $\hat{\beta}$ and the estimation of variance of error σ_ϵ^2 . We first perform a literature review on existing results on the *errors-in-variables model*. Then we will describe in detail the methodology, which is the computer simulation technique used to produce results. Finally, we will present the results and findings from the computer simulation experiments. The code used in this work can be found in the appendix.

2 Literature review

We mainly refer to the lecture notes written by Pischke [1] for the errors-in-variables model.

Suppose we wish to estimate the relationship $y = \beta x + \epsilon$, but we only have data on $\tilde{x} = x + u$. Also, let's further assume that $\sigma_v^2 = 0$, i.e. there is only measurement error in x .

If we substitute $\tilde{x} = x + u$ into $y = \beta x + \epsilon$, we obtain:

$$y = \beta(\tilde{x} - u) + \epsilon = \beta\tilde{x} + (\epsilon - \beta u) \quad (2)$$

As the measurement error in x becomes part of the residual error term in the model, the exogeneity assumption of the Gauss-Markov theorem is violated as $cov(u, \tilde{x}) \neq 0$. Thus, the ordinary least-squares (OLS) estimator of β may not be the *best linear unbiased estimator* (BLUE). Unlike the case of with measurement error, the OLS and MME estimators are different. In fact, We will see that the OLS estimators of β and σ_ϵ^2 are biased. In order to obtain unbiased and consistent estimates, we would have to resort to the method of moments (MME) estimators instead. Furthermore, although the OLS estimator is the same as the maximum likelihood estimator under normal-distributed errors, this may not be the case under non-normal errors.

2.1 OLS and MME for β

Suppose we use the ordinary least-squares (OLS) estimator for β :

$$\hat{\beta}_{OLS} = \frac{cov(\tilde{x}, y)}{var(\tilde{x})} = \frac{cov(x + u, \beta x + \epsilon)}{var(x + u)} \quad (3)$$

Because ϵ , u and x are independent to each other, we can obtain the limit of $\hat{\beta}_{OLS}$

$$plim \hat{\beta}_{OLS} = \frac{\beta\sigma_x^2}{\sigma_x^2 + \sigma_u^2} = \lambda\beta \quad (4)$$

where

$$\lambda \equiv \frac{\sigma_x^2}{\sigma_x^2 + \sigma_u^2} \quad (5)$$

This λ is also called the reliability ratio or signal-to-variance ratio.

Therefore, we can see that the OLS estimator $\hat{\beta}$ is biased towards zero because $0 < \lambda < 1$. The sign of the bias depends on the sign of the true β .

In order to obtain consistent estimates for β , we can use the MME estimator instead. Suppose we have some prior knowledge on the measurement errors and have obtained the value of σ_x^2 , σ_u^2 or λ . Then we can apply the appropriate

adjustment for the bias in the OLS $\hat{\beta}$ as $\sigma_{\tilde{x}}^2 = \sigma_x^2 + \sigma_u^2$, as we have can estimate $\sigma_{\tilde{x}}^2$ using $var(\tilde{x})$, which can be measured directly from the observation data.

Some MME estimators for β using the first and second moments alone are shown below:

Table 1: Various MME estimators for β

| Assumption | Method of Moments Estimator |
|-----------------------------------|---|
| σ_x^2 known | $\frac{cov(\tilde{x}, y)}{\sigma_{\tilde{x}}^2}$ |
| σ_u^2 known | $\frac{cov(\tilde{x}, y)}{var(\tilde{x}) - \sigma_u^2}$ |
| Reliability ratio λ known | $\frac{cov(\tilde{x}, y)}{\lambda var(\tilde{x})}$ |

There are practical methods available to obtain information on the measurement errors. For example, to obtain the value of σ_u^2 , we can perform repeated measurements [2]. To obtain the reliability ratio λ , we can use methods including "intraclass correlation via an internal replication study" [2].

2.2 OLS and MME for σ_ϵ^2

The usual way in OLS to estimate σ_ϵ^2 is to calculate $\hat{\sigma}_\epsilon^2$ as the sum of squares of the residuals divided by the degrees of freedom $n - 1$. To find out what happens to $\hat{\sigma}_\epsilon^2$, we can first look at what happens to the estimated residual variance first [1]:

$$\begin{aligned}
\hat{\epsilon} &= y - \hat{\beta}\tilde{x} \\
&= y - \hat{\beta}(x + u) \\
&= \epsilon - (y - \beta x) + y - \hat{\beta}x - \hat{\beta}u \\
&= \epsilon + (\beta - \hat{\beta})x - \hat{\beta}u
\end{aligned} \tag{6}$$

Thus, we can see that the true error term is obfuscated by two additional sources of error.

We can further obtain the limit of the OLS variance estimator, as we assumed earlier that ϵ , x and u are uncorrelated:

$$plim \hat{\sigma}_\epsilon^2 = \sigma_\epsilon^2 + (1 - \lambda)^2 \beta^2 \sigma_x^2 + \lambda^2 \beta^2 \sigma_u^2 \tag{7}$$

In order to obtain the MME estimator for σ_ϵ^2 , similar to the case for $\hat{\beta}_{MME}$, if we have some prior knowledge on the observation error, we can obtain consistent

estimates for σ_x^2 , σ_u , λ and β . Thus, by rearranging the expression for $\hat{\sigma}_\epsilon^2$, we have:

$$\hat{\sigma}_{\epsilon MME}^2 = \hat{\sigma}_{\epsilon OLS}^2 - (1 - \lambda)^2 \hat{\beta}_{MME}^2 \sigma_x^2 - \lambda^2 \hat{\beta}_{MME}^2 \sigma_u^2 \quad (8)$$

Based on the relation $\tilde{x}_i = x_i + u_i$, we can obtain $var(x) = var(\tilde{x}) - var(u)$. Thus, if σ_u^2 is known, and we use $var(x)$ as an estimator for σ_x^2 , we can substitute this into Equation 8, obtaining:

$$\hat{\sigma}_{\epsilon MME}^2 = \hat{\sigma}_{\epsilon OLS}^2 - (1 - \lambda)^2 \hat{\beta}_{MME}^2 (var(\tilde{x}) - \sigma_u^2) - \lambda^2 \hat{\beta}_{MME}^2 \sigma_u^2 \quad (9)$$

3 Methodology

3.1 Procedures

The computer simulation method was employed in order to investigate the impacts of non-normal errors on the OLS and MME estimates.

For all simulations, we use a common β_{truth} , n , x_{lo} , x_{hi} . β_{truth} is the underlying ground truth β of the model $y = \beta x + \epsilon$. n is the number of observations used in each simulation trial. x_{lo} and x_{hi} are the lower bound and upper bound of the x being sampled respectively.

In order to produce comparable results when using different error distributions, for each experiment, we first fix a certain σ_u^2 and σ_ϵ^2 . Afterwards, for the observation error u , we choose a distribution where the observation errors u_i are sampled from. The parameters of this distribution are chosen such that it has variance equal to σ_u^2 . Similarly, for the residual error ϵ , we choose a distribution where the residual errors ϵ_i are sampled from. The parameters of this distribution are chosen such that it has variance equal to σ_ϵ^2 .

After determining which distributions to use, we perform the observation data generation step. First, we pick n instances of explanatory variable x_i uniformly from the interval $[x_{lo}, x_{hi}]$. Afterwards, for each of the instances x_i , we compute $y_i = \beta_{truth} \cdot x_i + \epsilon_i$, where ϵ_i is drawn independently from the distribution for residual errors determined for this simulation trial. For each x_i , we also compute the observed explanatory variable $\tilde{x}_i = x_i + u_i$, where observation error u_i is added for each x_i . u_i is drawn independently from the distribution for observation errors determined for this simulation trial.

After generating the observation data, we perform estimations based on the observation data. We perform both the OLS and MME estimations for β and

σ_ϵ^2 . For the MME estimations, we use the case where prior information on the observation error σ_u^2 is known. The relevant expressions for the estimators are listed below:

$$\begin{aligned}
\hat{\beta}_{OLS} &= \frac{cov(x, \tilde{y})}{var(\tilde{x})} = \frac{cov(x + u, \beta x + \epsilon)}{var(x + u)} \\
\hat{\beta}_{MME} &= \frac{var(\tilde{x})}{var(\tilde{x}) - \sigma_u^2} \\
\hat{\sigma}_{\epsilon OLS}^2 &= var\left(y - \hat{\beta}_{OLS}\tilde{x}\right) \\
\hat{\sigma}_{\epsilon MME}^2 &= \hat{\sigma}_{\epsilon OLS}^2 - (1 - \lambda)^2 \hat{\beta}_{MME}^2 (var(\tilde{x}) - \sigma_u^2) - \lambda^2 \hat{\beta}_{MME}^2 \sigma_u^2
\end{aligned} \tag{10}$$

After computing the estimators, we compute the bias and squared error of each of these estimators when compared to the ground truth β_{truth} and σ_ϵ^2 .

For the same observation error distribution and residual error distribution, we perform the whole observation data generation and estimation procedures for multiple iterations. This is to obtain more reliable conclusions on the mean bias error (MBE) and mean squared error (MSE) of the estimators.

The computer simulation procedures are summarized in procedure 1.

3.2 Choice of distribution

In order to investigate the impacts on non-normality on the estimators, there is a need to use non-normal distributions to sample the observation error u and residual error ϵ . In order to satisfy the assumptions of the errors-in-variables model, the distributions used to generate the error terms should have mean zero. Furthermore, in our experiment procedures, we also have to derive the model parameters from a fixed σ_u^2 and σ_ϵ^2 . Thus, in order to ease calculation, the distribution chosen should have finite variance and the variance is in a closed form such that the derivation of model parameters from a fixed variance is easy to compute.

Taking the above factors into consideration, we decided to use the following distributions: The normal/Gaussian distribution, the Student's t distribution, and the χ^2 distribution re-centered at mean 0. They are chosen because they demonstrate features not seen in the normal distribution - the Student's t distribution has heavier tails than the normal distribution, and the χ^2 distribution is skewed. They pose substantial differences from the normal distribution, which has light tails and is symmetric at 0.

We will then perform the simulations using these distributions as the under-

Procedure 1: Computer simulation procedure

```
Initialize  $\beta_{truth}, n, x_{lo}, x_{hi}$ ;  
forall  $\sigma_u^2 \in \{\text{list of observation variance}\}$  do  
  forall  $\sigma_\epsilon^2 \in \{\text{list of residual variance}\}$  do  
    Initialize  $MSE_{\hat{\beta}_{OLS}} = 0, MSE_{\hat{\beta}_{MME}} = 0, MBE_{\hat{\beta}_{OLS}} = 0,$   
       $MBE_{\hat{\beta}_{MME}} = 0;$   
    Initialize  $MSE_{\hat{\sigma}_{OLS}^2} = 0, MSE_{\hat{\sigma}_{MME}^2} = 0, MBE_{\hat{\sigma}_{OLS}^2} = 0,$   
       $MBE_{\hat{\sigma}_{MME}^2} = 0;$   
    for 10000 iterations do  
      Pick  $n$  instances of explanatory variable  $x_i$  uniformly from  
        the interval  $[x_{lo}, x_{hi}]$ ;  
      Compute  $n$  observations of  $y_i = \beta_{truth} \cdot x_i + \epsilon_i$ , where  $\epsilon_i$  is  
        drawn independently from a distribution with variance  $\sigma_\epsilon^2$ ;  
      Compute  $n$  instances of observed explanatory variable  
         $\tilde{x}_i = x_i + u_i$ , where observation  $u_i$  is drawn independently  
        from a distribution with variance  $\sigma_u^2$ ;  
      Calculate  $\hat{\beta}_{OLS}, \hat{\beta}_{MME}, \hat{\sigma}_{OLS}^2, \hat{\sigma}_{MME}^2$  from the  
        observations  $\tilde{x}, y$ , and prior information on observation  
        error  $\sigma_u^2$ ;  
      Update  $MSE_{\hat{\beta}_{OLS}}, MSE_{\hat{\beta}_{MME}}, MBE_{\hat{\beta}_{OLS}}, MBE_{\hat{\beta}_{MME}},$   
         $MSE_{\hat{\sigma}_{OLS}^2}, MSE_{\hat{\sigma}_{MME}^2}, MBE_{\hat{\sigma}_{OLS}^2}, MBE_{\hat{\sigma}_{MME}^2};$   
    end  
  end  
end
```

lying sampling distributions of the observation error u and residual error ϵ , as described in the previous subsection.

3.3 Experiment parameters used

To produce the results in the subsequent section, we used the following parameters: $n = 30$, $\beta_{truth} = 2$, $x_{lo} = -20$, $x_{hi} = 20$. For the simulation procedures, we used $\sigma_u^2 \in \{0, 1.5, 2, 2.5, 3.0, 4.5, 6.0, 8.0, 10.0\}$, and $\sigma_\epsilon^2 \in \{1.5, 2, 2.5\}$. We are unable to experiment with variances from $(0, 1]$ because the variance of the Student's t distribution is only defined when $\sigma^2 > 1$.

For each combination of observation error distribution and residual error distribution, we performed 10,000 iterations to produce the MBE and MSE results.

3.4 Implementation details

The codes were implemented in Python in a Jupyter notebook. We used the sampling distribution implementations provided in the `numpy.random` package. The detailed implementation and source code can be found in the appendix.

4 Results and discussions

We split this section into several subsections.

First, we describe the results for the estimation of β . We consider both the MBE and MSE for the OLS and MME estimators of β . We compare the simulation outcomes for different distributions of residual error for when there is absence of observation error (i.e. $\sigma_u^2 = 0$). For the cases under the presence of observation error (i.e. $\sigma_u^2 > 0$), we compare the MBE and MSE for different distributions (Normal, Student's t , χ^2) while maintaining the same variance.

Afterwards, we similarly describe the results for the estimation of σ_ϵ^2 . We consider both the MBE and MSE for the OLS and MME estimators of σ^2 . We similarly compare the simulation outcomes under the cases of absence of observation error, and cases under the presence of observation error of different distributions respectively.

4.1 Estimation of β

We use the following estimators of β and apply them to the generated observation data (\tilde{x}_i, y_i) :

$$\begin{aligned}\hat{\beta}_{OLS} &= \frac{\text{cov}(\tilde{x}, y)}{\text{var}(\tilde{x})} = \frac{\text{cov}(x + u, \beta x + \epsilon)}{\text{var}(x + u)} \\ \hat{\beta}_{MME} &= \frac{\text{var}(\tilde{x})}{\text{var}(\tilde{x}) - \sigma_u^2}\end{aligned}\tag{11}$$

4.1.1 Absence of observation error

Under the absence of observation error, the OLS and MME estimators of β are the same.

Table 2 summarizes the MBE of β when $\sigma_u^2 = 0$ for different error distri-

Table 2: Estimation MBE of β when $\sigma_u^2 = 0$, under different distributions and variance of ϵ .

| | Normal | Student's t | χ^2 |
|---------------------------|-----------|---------------|-----------|
| $\sigma_\epsilon^2 = 1.5$ | -0.000260 | 0.000252 | 0.000052 |
| $\sigma_\epsilon^2 = 2.0$ | 0.000151 | 0.000114 | -0.000356 |
| $\sigma_\epsilon^2 = 2.5$ | -0.000176 | -0.000105 | 0.000112 |

butions for ϵ . There is no clear trend of increasing/decreasing pattern of the MBE for each of the distributions. This is mostly because the bias is centered at zero, leading to the signs of bias to be different between simulation trials, which cancel out each other in the calculation of MBE.

Table 3: Estimation MSE of β when $\sigma_u^2 = 0$, under different distributions and variance of ϵ .

| | Normal | Student's t | χ^2 |
|---------------------------|----------|---------------|----------|
| $\sigma_\epsilon^2 = 1.5$ | 0.000352 | 0.000351 | 0.000343 |
| $\sigma_\epsilon^2 = 2.0$ | 0.000470 | 0.000478 | 0.000458 |
| $\sigma_\epsilon^2 = 2.5$ | 0.000588 | 0.000594 | 0.000573 |

Table 3 summarizes the MSE of β when $\sigma_u^2 = 0$ for different error distributions. There is a clear trend of increasing of the MSE for each of the distributions as σ_ϵ^2 increases. We can also observe that the MSE for χ^2 distribution is the least in all 3 experiments, and the MSE for Student's t distribution is the greatest in 2 out of 3 experiments.

4.1.2 Normal-distributed observation error and normal-distributed residual error

Under the presence of observation error, the OLS and MME estimators of β are different. We would expect the MME estimator to have lower absolute bias and lower MSE because under the presence of observation error, the MME estimator is consistent but the OLS estimator is inconsistent.

Table 4 and Table 5 summarize the simulated MBE and MSE of the estimation of β respectively, under normal-distributed errors.

We first focus on the observations of the MBE and MSE of $\hat{\beta}_{OLS}$. Firstly, we observe that the MBE of $\hat{\beta}_{OLS}$ is negative. This is result is expected because $\text{plim } \hat{\beta}_{OLS} = \frac{\beta\sigma_x^2}{\sigma_x^2 + \sigma_u^2} = \lambda \cdot \beta_{truth}$ and is biased towards zero. In this simulation, $\beta_{truth} = 2 > 0$. Thus, the bias of $\hat{\beta}_{OLS}$ is expected to be negative, which

Table 4: Estimation MBE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under normal-distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | -0.019467 | 0.000711 | -0.018743 | 0.001453 | -0.018929 | 0.001258 |
| $\sigma_u^2 = 2.0$ | -0.025206 | 0.001655 | -0.023977 | 0.002920 | -0.025265 | 0.001602 |
| $\sigma_u^2 = 2.5$ | -0.030908 | 0.002607 | -0.030422 | 0.003106 | -0.030886 | 0.002635 |
| $\sigma_u^2 = 3.0$ | -0.036965 | 0.003140 | -0.036627 | 0.003527 | -0.036902 | 0.003217 |
| $\sigma_u^2 = 4.5$ | -0.055003 | 0.004770 | -0.056174 | 0.003501 | -0.056439 | 0.003171 |
| $\sigma_u^2 = 6.0$ | -0.072019 | 0.007202 | -0.071406 | 0.007854 | -0.073418 | 0.005616 |
| $\sigma_u^2 = 8.0$ | -0.096652 | 0.007750 | -0.096022 | 0.008548 | -0.096813 | 0.007639 |
| $\sigma_u^2 = 10.0$ | -0.119142 | 0.010226 | -0.121257 | 0.007779 | -0.117934 | 0.011878 |

Table 5: Estimation MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under normal-distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.002049 | 0.001761 | 0.002120 | 0.001862 | 0.002254 | 0.001992 |
| $\sigma_u^2 = 2.0$ | 0.002801 | 0.002331 | 0.002747 | 0.002338 | 0.003020 | 0.002552 |
| $\sigma_u^2 = 2.5$ | 0.003447 | 0.002739 | 0.003563 | 0.002895 | 0.003810 | 0.003124 |
| $\sigma_u^2 = 3.0$ | 0.004320 | 0.003313 | 0.004423 | 0.003452 | 0.004584 | 0.003595 |
| $\sigma_u^2 = 4.5$ | 0.007218 | 0.004996 | 0.007379 | 0.005007 | 0.007404 | 0.004989 |
| $\sigma_u^2 = 6.0$ | 0.010421 | 0.006639 | 0.010245 | 0.006529 | 0.010711 | 0.006687 |
| $\sigma_u^2 = 8.0$ | 0.015781 | 0.008820 | 0.015708 | 0.008875 | 0.016050 | 0.009091 |
| $\sigma_u^2 = 10.0$ | 0.022013 | 0.011578 | 0.022421 | 0.011359 | 0.021891 | 0.011826 |

is consistent with the simulation findings. Second, we observe that the MBE of $\hat{\beta}_{OLS}$ decreases as σ_u^2 increases. This is also consistent with the literature as $\lambda \equiv \frac{\sigma_x^2}{\sigma_x^2 + \sigma_u^2}$. When σ_u^2 increases, λ decreases, and the bias $(\lambda - 1) \cdot \beta_{truth}$ decreases. Third, we observe that the MSE of $\hat{\beta}_{OLS}$ increases as σ_u^2 increases. Intuitively, this means that the error of the estimation of β increases as the observation data is dirtier.

Now, we turn our attention to the observations of the MBE and MSE of $\hat{\beta}_{MME}$, and compare them with $\hat{\beta}_{OLS}$. Firstly, we observe that the absolute value of MBE and MSE of $\hat{\beta}_{MME}$ are both lower than that of $\hat{\beta}_{OLS}$ at the same σ_u^2 and σ_ϵ^2 . This is expected because $\hat{\beta}_{MME}$ is a consistent estimator of β while $\hat{\beta}_{OLS}$ is not. Second, we observe that the MBE of $\hat{\beta}_{MME}$ is positive and in general increases as σ_u^2 increases. Third, we observe that the MSE of $\hat{\beta}_{MME}$

increases as σ_u^2 increases. This is similar to the case of $\hat{\beta}_{OLS}$.

4.1.3 Student's t -distributed observation error and Student's t -distributed residual error

We perform the similar simulations for Student's t distributed observation error and residual error for different variances. Elementary statistical theory tells us that the variance for a Student's t distribution with $\nu > 2$ degrees of freedom has variance equal to $\frac{\nu}{\nu-2}$. Taking inverse of the equation $\sigma^2 = \frac{\nu}{\nu-2}$ gives $\nu = 2 \cdot \frac{\sigma^2}{\sigma^2-1}$. Note that the range of σ_u^2 and σ_ϵ^2 are specially chosen such that $\nu > 2$ because the Student's t distribution only has finite variance when $\nu > 2$. Thus, in the tables shown below, when we write $\sigma^2 = \delta$, it means the simulation trial was performed using the Student's t distribution with $2 \cdot \frac{\delta}{\delta-1}$ degrees of freedom.

Table 6 and Table 7 summarize the simulated MBE and MSE of the estimation of β respectively, under Student's t -distributed errors.

Table 6: Estimation MBE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under Student's t -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | -0.018770 | 0.001429 | -0.019233 | 0.000955 | -0.018866 | 0.001329 |
| $\sigma_u^2 = 2.0$ | -0.024866 | 0.002008 | -0.025077 | 0.001785 | -0.025835 | 0.001001 |
| $\sigma_u^2 = 2.5$ | -0.030725 | 0.002796 | -0.030205 | 0.003343 | -0.030965 | 0.002552 |
| $\sigma_u^2 = 3.0$ | -0.035638 | 0.004529 | -0.035464 | 0.004732 | -0.034195 | 0.006041 |
| $\sigma_u^2 = 4.5$ | -0.047532 | 0.012747 | -0.047863 | 0.012389 | -0.047730 | 0.012474 |
| $\sigma_u^2 = 6.0$ | -0.053997 | 0.026746 | -0.054018 | 0.026632 | -0.055230 | 0.025393 |
| $\sigma_u^2 = 8.0$ | -0.063497 | 0.044623 | -0.062978 | 0.045236 | -0.060738 | 0.047766 |
| $\sigma_u^2 = 10.0$ | -0.066641 | 0.070167 | -0.065806 | 0.070902 | -0.065357 | 0.071511 |

We first focus on the observations based solely on the Student's t -distributed error simulations. This will be followed by comparing the results in normal-distributed errors and Student's t -distributed errors.

Firstly, similar to the normal-distributed errors case, we observe that the MBE of $\hat{\beta}_{OLS}$ is negative and decreasing in σ_u^2 , and that the MBE of $\hat{\beta}_{MME}$ is positive and increasing in σ_u^2 . The MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ are both increasing in σ_u^2 .

When comparing $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$, we find that switching from $\hat{\beta}_{OLS}$ to $\hat{\beta}_{MME}$ still provides a decrease in absolute MBE or MSE, but the decrease is not

Table 7: Estimation MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under Student's t -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.002021 | 0.001761 | 0.002213 | 0.001939 | 0.002321 | 0.002065 |
| $\sigma_u^2 = 2.0$ | 0.002921 | 0.002473 | 0.003258 | 0.002809 | 0.003236 | 0.002744 |
| $\sigma_u^2 = 2.5$ | 0.004083 | 0.003415 | 0.004288 | 0.003668 | 0.004529 | 0.003870 |
| $\sigma_u^2 = 3.0$ | 0.005745 | 0.004929 | 0.005835 | 0.005041 | 0.005557 | 0.004851 |
| $\sigma_u^2 = 4.5$ | 0.011736 | 0.010821 | 0.011447 | 0.010442 | 0.010748 | 0.009719 |
| $\sigma_u^2 = 6.0$ | 0.013428 | 0.013042 | 0.013088 | 0.012611 | 0.014464 | 0.013965 |
| $\sigma_u^2 = 8.0$ | 0.018133 | 0.019265 | 0.017601 | 0.018818 | 0.017577 | 0.019249 |
| $\sigma_u^2 = 10.0$ | 0.020794 | 0.025946 | 0.019452 | 0.024433 | 0.019140 | 0.024228 |

that large when compared to the normal-distributed errors case. This situation is much more apparent in large values of σ_u^2 . For MBE, when $\sigma_u^2 = 10.0$, we observe that the absolute value of the MBE of $\hat{\beta}_{MME}$ is in fact greater than that of $\hat{\beta}_{OLS}$. Similarly for MSE, when $\sigma_u^2 \in \{8.0, 10.0\}$, we observe that the MSE of $\hat{\beta}_{MME}$ is greater than that of $\hat{\beta}_{OLS}$.

To compare the findings between the normal-distributed errors case and Student's t -distributed errors case, it is useful to take the cell difference between the estimation MBE in Tables 4 and 6, and estimation MSE in Tables 5 and 7. We define the difference here as (value from Student's t) – (value from normal). These differences are tabulated in Tables 8 and 9 respectively.

Table 8: Difference between Student's t error case and normal-distributed error case: Estimation MBE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | -0.000028 | -0.000000 | 0.000093 | 0.000077 | 0.000067 | 0.000073 |
| $\sigma_u^2 = 2.0$ | 0.000120 | 0.000142 | 0.000511 | 0.000471 | 0.000215 | 0.000191 |
| $\sigma_u^2 = 2.5$ | 0.000636 | 0.000676 | 0.000725 | 0.000773 | 0.000719 | 0.000746 |
| $\sigma_u^2 = 3.0$ | 0.001425 | 0.001616 | 0.001412 | 0.001589 | 0.000973 | 0.001257 |
| $\sigma_u^2 = 4.5$ | 0.004518 | 0.005825 | 0.004068 | 0.005434 | 0.003343 | 0.004731 |
| $\sigma_u^2 = 6.0$ | 0.003007 | 0.006403 | 0.002843 | 0.006083 | 0.003753 | 0.007279 |
| $\sigma_u^2 = 8.0$ | 0.002352 | 0.010444 | 0.001893 | 0.009943 | 0.001527 | 0.010159 |
| $\sigma_u^2 = 10.0$ | -0.001219 | 0.014368 | -0.002969 | 0.013074 | -0.002751 | 0.012402 |

It should be noted that when the variance of a Student's t distribution is low, the number of degrees of freedom of the distribution is high. From elementary statistical theory we know that as the number of degrees of freedom of a

Table 9: Difference between Student's t error case and normal-distributed error case: Estimation MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.000697 | 0.000718 | -0.000491 | -0.000498 | 0.000063 | 0.000072 |
| $\sigma_u^2 = 2.0$ | 0.000340 | 0.000354 | -0.001099 | -0.001135 | -0.000570 | -0.000601 |
| $\sigma_u^2 = 2.5$ | 0.000183 | 0.000189 | 0.000218 | 0.000237 | -0.000079 | -0.000083 |
| $\sigma_u^2 = 3.0$ | 0.001327 | 0.001390 | 0.001162 | 0.001205 | 0.002707 | 0.002824 |
| $\sigma_u^2 = 4.5$ | 0.007471 | 0.007977 | 0.008311 | 0.008888 | 0.008708 | 0.009303 |
| $\sigma_u^2 = 6.0$ | 0.018022 | 0.019544 | 0.017388 | 0.018778 | 0.018188 | 0.019777 |
| $\sigma_u^2 = 8.0$ | 0.033155 | 0.036873 | 0.033045 | 0.036688 | 0.036074 | 0.040126 |
| $\sigma_u^2 = 10.0$ | 0.052501 | 0.059941 | 0.055451 | 0.063122 | 0.052577 | 0.059633 |

Student's t distribution increases, it converges to the normal distribution. Thus, for some low fixed variance, we should expect similar results from a Student's t distribution and normal distribution with that variance.

We first focus on the comparisons in MBE. For $\hat{\beta}_{OLS}$, we observe that the difference is first close to zero, then becomes positive and increases to a certain point (around at $\sigma_u^2 = 4.5$), then again starts to decrease. As the bias of $\hat{\beta}_{OLS}$ is negative in both Student's t -distributed errors and normal-distributed errors, when the difference in MBE is positive, it means that $\hat{\beta}_{OLS}$ under Student's t errors is closer to zero. On the other hand, for $\hat{\beta}_{MME}$, we observe that the difference is positive and increasing throughout as σ_u^2 is increased. As the MBE for $\hat{\beta}_{MME}$ is positive in both Student's t errors and normal errors, this means that the bias for $\hat{\beta}_{MME}$ under Student's t errors is greater and further away from zero.

We then focus on the comparisons in MSE. For both $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$, we observe that the difference is mostly positive, meaning that the MSE under Student's t -distributed errors is in general greater than that under normal-distributed errors, and that this difference is in general increasing in σ_u^2 .

4.1.4 χ^2 -distributed observation error

We perform the similar simulations for χ^2 distributed observation error and residual error for different variances. From elementary statistical theory, a χ^2 distribution with $k > 0$ degrees of freedom has mean k and variance $2k$. To let the distribution follow the model assumptions to have mean zero, we subtract k from each value drawn from the distribution. Thus, when we write $\sigma^2 = \delta$, it means the simulation trial was performed using the χ^2 distribution with $2 \cdot \frac{\delta}{2}$ degrees of freedom minus $\frac{\delta}{2}$ to recenter the mean to zero.

Table 10 and Table 11 summarize the simulated MBE and MSE of the estimation of β respectively, under χ^2 -distributed errors.

Table 10: Estimation MBE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under χ^2 -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | -0.018714 | 0.001479 | -0.018709 | 0.001491 | -0.018640 | 0.001558 |
| $\sigma_u^2 = 2.0$ | -0.025420 | 0.001430 | -0.024128 | 0.002761 | -0.025012 | 0.001851 |
| $\sigma_u^2 = 2.5$ | -0.030573 | 0.002949 | -0.031135 | 0.002355 | -0.031329 | 0.002149 |
| $\sigma_u^2 = 3.0$ | -0.036298 | 0.003864 | -0.036901 | 0.003226 | -0.037024 | 0.003085 |
| $\sigma_u^2 = 4.5$ | -0.054917 | 0.004847 | -0.055502 | 0.004227 | -0.054086 | 0.005756 |
| $\sigma_u^2 = 6.0$ | -0.072913 | 0.006221 | -0.072631 | 0.006482 | -0.072691 | 0.006443 |
| $\sigma_u^2 = 8.0$ | -0.097304 | 0.007052 | -0.095563 | 0.009114 | -0.097776 | 0.006451 |
| $\sigma_u^2 = 10.0$ | -0.119601 | 0.009722 | -0.119559 | 0.009724 | -0.117377 | 0.012370 |

Table 11: Estimation MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under χ^2 -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.002214 | 0.001963 | 0.002375 | 0.002129 | 0.002378 | 0.002131 |
| $\sigma_u^2 = 2.0$ | 0.003024 | 0.002552 | 0.002964 | 0.002557 | 0.003111 | 0.002658 |
| $\sigma_u^2 = 2.5$ | 0.003698 | 0.003029 | 0.003881 | 0.003179 | 0.003957 | 0.003242 |
| $\sigma_u^2 = 3.0$ | 0.004428 | 0.003481 | 0.004697 | 0.003716 | 0.004774 | 0.003782 |
| $\sigma_u^2 = 4.5$ | 0.007516 | 0.005330 | 0.007542 | 0.005260 | 0.007606 | 0.005527 |
| $\sigma_u^2 = 6.0$ | 0.011204 | 0.007381 | 0.011033 | 0.007194 | 0.011208 | 0.007384 |
| $\sigma_u^2 = 8.0$ | 0.016869 | 0.009983 | 0.016494 | 0.009954 | 0.016978 | 0.009951 |
| $\sigma_u^2 = 10.0$ | 0.022575 | 0.012067 | 0.022936 | 0.012550 | 0.022595 | 0.012841 |

For the MBE and MSE, we have similar observations to the normal-distributed errors and Student's t -distributed errors cases, where the MBE of $\hat{\beta}_{OLS}$ is negative and decreasing in σ_u^2 , and the MBE of $\hat{\beta}_{MME}$ is positive and increasing in σ_u^2 . Similar to before, the MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ are both increasing in σ_u^2 , and the MSE for $\hat{\beta}_{MME}$ is smaller than that for $\hat{\beta}_{OLS}$.

To compare findings between the normal-distributed errors and χ^2 -distributed errors case, we again take the cell difference from the two distributions in their MBE and MSE respectively. The difference here is (value from χ^2) – (value from normal). These differences are tabulated in Tables 12 and 13 respectively.

Table 12: Difference between χ^2 -distributed error case and normal-distributed error case: Estimation MBE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.000752 | 0.000768 | 0.000034 | 0.000038 | 0.000289 | 0.000301 |
| $\sigma_u^2 = 2.0$ | -0.000214 | -0.000225 | -0.000150 | -0.000159 | 0.000253 | 0.000249 |
| $\sigma_u^2 = 2.5$ | 0.000335 | 0.000342 | -0.000712 | -0.000751 | -0.000442 | -0.000486 |
| $\sigma_u^2 = 3.0$ | 0.000667 | 0.000724 | -0.000274 | -0.000301 | -0.000122 | -0.000133 |
| $\sigma_u^2 = 4.5$ | 0.000086 | 0.000077 | 0.000672 | 0.000725 | 0.002353 | 0.002585 |
| $\sigma_u^2 = 6.0$ | -0.000895 | -0.000981 | -0.001225 | -0.001371 | 0.000728 | 0.000827 |
| $\sigma_u^2 = 8.0$ | -0.000651 | -0.000698 | 0.000460 | 0.000566 | -0.000963 | -0.001188 |
| $\sigma_u^2 = 10.0$ | -0.000459 | -0.000504 | 0.001698 | 0.001944 | 0.000557 | 0.000492 |

Table 13: Difference between χ^2 -distributed error case and normal-distributed error case: Estimation MSE of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{MME}$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------|---------------------|---------------------------|---------------------|---------------------------|---------------------|
| | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ | $\hat{\beta}_{OLS}$ | $\hat{\beta}_{MME}$ |
| $\sigma_u^2 = 1.5$ | 0.000165 | 0.000202 | 0.000255 | 0.000267 | 0.000124 | 0.000139 |
| $\sigma_u^2 = 2.0$ | 0.000223 | 0.000221 | 0.000216 | 0.000219 | 0.000090 | 0.000105 |
| $\sigma_u^2 = 2.5$ | 0.000251 | 0.000290 | 0.000318 | 0.000285 | 0.000147 | 0.000118 |
| $\sigma_u^2 = 3.0$ | 0.000108 | 0.000167 | 0.000274 | 0.000264 | 0.000190 | 0.000188 |
| $\sigma_u^2 = 4.5$ | 0.000298 | 0.000334 | 0.000164 | 0.000253 | 0.000202 | 0.000538 |
| $\sigma_u^2 = 6.0$ | 0.000783 | 0.000742 | 0.000788 | 0.000666 | 0.000498 | 0.000697 |
| $\sigma_u^2 = 8.0$ | 0.001088 | 0.001163 | 0.000785 | 0.001079 | 0.000928 | 0.000860 |
| $\sigma_u^2 = 10.0$ | 0.000562 | 0.000489 | 0.000516 | 0.001191 | 0.000705 | 0.001015 |

When comparing findings from χ^2 -distributed errors and normal-distributed errors, it should be again noted that the χ^2 distribution re-centered at mean zero converges to a normal distribution with the same variance when the number of degrees of freedom is large.

For the MBE, there seems to be no clear comparison between the normal-distributed and χ^2 -distributed cases, because the sign of the difference fluctuates in $\hat{\beta}_{OLS}$, $\hat{\beta}_{MME}$ and different σ_u^2 .

For the MSE, the difference is always positive, meaning that the MSE for χ^2 -distributed case is always higher than normal-distributed case. The difference seems to be increasing in σ_u^2 , but the trend is not that clear.

4.2 Estimation of σ_ϵ^2

We use the following estimators of σ_ϵ^2 and apply them to the generated observation data (\tilde{x}_i, y_i) :

$$\begin{aligned}\hat{\sigma}_{OLS}^2 &= \text{var} \left(y - \hat{\beta}_{OLS} \tilde{x} \right) \\ \hat{\sigma}_{MME}^2 &= \hat{\sigma}_{OLS}^2 - (1 - \lambda)^2 \hat{\beta}_{MME}^2 (\text{var}(\tilde{x}) - \sigma_u^2) - \lambda^2 \hat{\beta}_{MME}^2 \sigma_u^2\end{aligned}\tag{12}$$

4.2.1 Absence of observation error

Under the absence of observation error, the OLS and MME estimators of σ_ϵ^2 are the same. Tables 14 and 15 summarize respectively the MBE and MSE of σ_ϵ^2 when $\sigma_u^2 = 0$ for different error distributions.

Table 14: Estimation MBE of σ_ϵ^2 when $\sigma_u^2 = 0$, under different distributions and variance of ϵ .

| | Normal | Student's t | χ^2 |
|---------------------------|-----------|---------------|-----------|
| $\sigma_\epsilon^2 = 1.5$ | 0.000038 | 0.006738 | -0.009608 |
| $\sigma_\epsilon^2 = 2.0$ | 0.000135 | -0.010998 | -0.005749 |
| $\sigma_\epsilon^2 = 2.5$ | -0.001355 | 0.018754 | 0.026852 |

Table 15: Estimation MSE of σ_ϵ^2 when $\sigma_u^2 = 0$, under different distributions and variance of ϵ .

| | Normal | Student's t | χ^2 |
|---------------------------|----------|---------------|----------|
| $\sigma_\epsilon^2 = 1.5$ | 0.154338 | 0.367452 | 1.325430 |
| $\sigma_\epsilon^2 = 2.0$ | 0.278130 | 1.806585 | 1.919026 |
| $\sigma_\epsilon^2 = 2.5$ | 0.437989 | 10.128111 | 2.414450 |

For MBE, similar to the case of $\hat{\beta}$, there seems to be no clear monotonic trend of the MBE from the simulations performed under different distributions. For MSE, we can observe that the MSE for normal-errors is the least out of the distributions in concern.

4.2.2 Normal-distributed observation error

We again perform the estimation of σ_ϵ^2 under normal-distributed observation error and normal-distributed residual error, both using different variances. The simulated MBE and MSE are tabulated in Tables 16 and 17 respectively.

Table 16: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under normal-distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 5.944546 | -0.005842 | 5.959183 | 0.010743 | 5.931388 | -0.014544 |
| $\sigma_u^2 = 2.0$ | 7.886910 | -0.024392 | 7.899105 | -0.010766 | 7.826871 | -0.081642 |
| $\sigma_u^2 = 2.5$ | 9.825091 | -0.041342 | 9.837348 | -0.020427 | 9.810995 | -0.038816 |
| $\sigma_u^2 = 3.0$ | 11.752766 | -0.042576 | 11.827087 | 0.014613 | 11.813583 | 0.010360 |
| $\sigma_u^2 = 4.5$ | 17.563742 | -0.015657 | 17.409776 | -0.158051 | 17.471703 | -0.083689 |
| $\sigma_u^2 = 6.0$ | 23.104857 | -0.118361 | 23.084622 | -0.121116 | 23.031219 | -0.217114 |
| $\sigma_u^2 = 8.0$ | 30.385471 | -0.234288 | 30.443624 | -0.211370 | 30.481465 | -0.197597 |
| $\sigma_u^2 = 10.0$ | 37.582805 | -0.319186 | 37.466108 | -0.471654 | 37.645894 | -0.250725 |

Table 17: Estimation MSE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under normal-distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 39.106310 | 3.818130 | 39.828669 | 4.367114 | 40.113631 | 4.986267 |
| $\sigma_u^2 = 2.0$ | 68.251852 | 6.135680 | 69.143363 | 6.878215 | 68.728170 | 7.618354 |
| $\sigma_u^2 = 2.5$ | 105.404358 | 9.135046 | 106.313434 | 9.829482 | 106.637167 | 10.601876 |
| $\sigma_u^2 = 3.0$ | 149.786174 | 11.983789 | 153.149474 | 13.570971 | 153.738092 | 14.681478 |
| $\sigma_u^2 = 4.5$ | 334.039435 | 26.905638 | 328.920230 | 26.949451 | 333.116014 | 29.293710 |
| $\sigma_u^2 = 6.0$ | 575.168146 | 44.066269 | 576.296059 | 46.827447 | 575.827556 | 48.625474 |
| $\sigma_u^2 = 8.0$ | 993.932566 | 78.364319 | 999.708303 | 80.049208 | 1003.433290 | 81.952748 |
| $\sigma_u^2 = 10.0$ | 1517.354461 | 120.073998 | 1511.381767 | 122.354870 | 1528.536252 | 126.664946 |

For $\hat{\sigma}_{\epsilon OLS}^2$, from the literature review, we know that $\text{plim } \hat{\sigma}_{\epsilon OLS}^2 = \sigma_\epsilon^2 + (1 - \lambda)^2 \beta^2 \sigma_x^2 + \lambda^2 \beta^2 \sigma_u^2$, meaning that there is a positive bias of $\hat{\sigma}_{\epsilon OLS}^2$. Indeed, we observe that the MBE of $\hat{\sigma}_{\epsilon OLS}^2$, is positive for all simulated values of σ_u^2 and σ_ϵ^2 . The MBE is increasing in σ_u^2 and there is no clear monotonic trend of MBE with σ_ϵ^2 , which are both consistent with the expression of additional bias terms in the explicit derivation of $\text{plim } \hat{\sigma}_{\epsilon OLS}^2$.

On the other hand, for $\hat{\sigma}_{\epsilon MME}^2$, the MBE is negative and roughly decreasing with σ_u^2 . There is no clear trend with σ_ϵ^2 .

4.2.3 Student's t -distributed observation error

We perform the similar simulations and estimation of σ_ϵ^2 under the case of Student's t -distributed observation error and residual error. The MBE and MSE results are tabulated in Tables 18 and 19 respectively:

Table 18: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under Student's t -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 5.901136 | -0.051106 | 5.949806 | 0.002570 | 5.921078 | -0.029820 |
| $\sigma_u^2 = 2.0$ | 7.808126 | -0.103737 | 7.873660 | -0.038787 | 7.785658 | -0.128063 |
| $\sigma_u^2 = 2.5$ | 9.403487 | -0.473446 | 9.674954 | -0.187735 | 9.548231 | -0.327451 |
| $\sigma_u^2 = 3.0$ | 10.991609 | -0.841798 | 11.224290 | -0.613500 | 11.263660 | -0.552675 |
| $\sigma_u^2 = 4.5$ | 14.854718 | -2.893005 | 14.739074 | -3.008623 | 15.304472 | -2.413692 |
| $\sigma_u^2 = 6.0$ | 16.737684 | -6.991530 | 16.764042 | -6.956211 | 17.054863 | -6.659021 |
| $\sigma_u^2 = 8.0$ | 19.173947 | -12.709664 | 19.076132 | -12.745150 | 19.562751 | -12.287313 |
| $\sigma_u^2 = 10.0$ | 21.736085 | -18.363607 | 20.925550 | -19.274530 | 20.519368 | -19.725613 |

Table 19: Estimation MSE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under Student's t -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 42.007008 | 7.330577 | 44.265186 | 9.028256 | 57.591174 | 22.712470 |
| $\sigma_u^2 = 2.0$ | 85.073222 | 25.150626 | 90.945333 | 30.167468 | 101.248433 | 41.912878 |
| $\sigma_u^2 = 2.5$ | 151.504695 | 66.572613 | 178.829111 | 90.060470 | 216.508799 | 130.531705 |
| $\sigma_u^2 = 3.0$ | 242.209257 | 130.228570 | 296.302408 | 181.223069 | 306.113309 | 190.717377 |
| $\sigma_u^2 = 4.5$ | 674.176226 | 504.183618 | 647.724335 | 480.312491 | 853.425135 | 679.140708 |
| $\sigma_u^2 = 6.0$ | 812.257778 | 650.836042 | 853.904589 | 694.262448 | 961.547494 | 800.543447 |
| $\sigma_u^2 = 8.0$ | 1213.885175 | 1159.121657 | 1260.906736 | 1218.715769 | 1323.105685 | 1251.042547 |
| $\sigma_u^2 = 10.0$ | 1873.268174 | 2046.381778 | 1691.114328 | 1887.541448 | 1558.572354 | 1770.970898 |

We find that the observations are similar to the case of estimating β under Student's t -distributed errors: We find that the improvement in estimation accuracy when using $\hat{\sigma}_{\epsilon MME}^2$ instead of using $\hat{\sigma}_{\epsilon OLS}^2$ is not as great as compared to the normal case. Similar to the case of estimating β , we also find that when σ_u^2 increases, the improvement of using $\hat{\sigma}_{\epsilon MME}^2$ to replace $\hat{\sigma}_{\epsilon OLS}^2$ decreases. In fact, we see that when $\sigma_u^2 = 10.0$, the MSE of $\hat{\sigma}_{\epsilon MME}^2$ is greater than that of $\hat{\sigma}_{\epsilon OLS}^2$.

To compare the findings between normal-distributed errors and Student's t distributed errors, we take differences in a fashion similar to the previous sections. The differences are tabulated in Tables 20 and 21 respectively:

Table 20: Difference between Student's t -distributed error case and normal-distributed error case: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | -0.066959 | -0.069825 | 0.001343 | 0.002872 | -0.000262 | 0.000937 |
| $\sigma_u^2 = 2.0$ | -0.109399 | -0.114493 | 0.016825 | 0.012610 | -0.150998 | -0.158049 |
| $\sigma_u^2 = 2.5$ | -0.430338 | -0.443487 | -0.183308 | -0.187789 | -0.298113 | -0.310561 |
| $\sigma_u^2 = 3.0$ | -0.803268 | -0.836666 | -0.559063 | -0.595413 | -0.515155 | -0.526378 |
| $\sigma_u^2 = 4.5$ | -2.704054 | -2.880683 | -2.761439 | -2.954847 | -2.155957 | -2.311831 |
| $\sigma_u^2 = 6.0$ | -6.380179 | -6.874139 | -6.318024 | -6.798425 | -5.968561 | -6.496184 |
| $\sigma_u^2 = 8.0$ | -11.271448 | -12.496345 | -11.273698 | -12.442319 | -10.695063 | -11.949910 |
| $\sigma_u^2 = 10.0$ | -15.743113 | -17.890013 | -16.732464 | -19.050908 | -16.942578 | -19.309470 |

Table 21: Difference between Student's t -distributed error case and normal-distributed error case: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 2.567444 | 3.439570 | 4.459091 | 4.544724 | 17.715820 | 17.829044 |
| $\sigma_u^2 = 2.0$ | 16.250164 | 18.866020 | 22.518222 | 23.336357 | 30.675791 | 34.186680 |
| $\sigma_u^2 = 2.5$ | 45.666405 | 57.182168 | 71.846646 | 79.995463 | 109.022235 | 119.665395 |
| $\sigma_u^2 = 3.0$ | 90.881267 | 117.709726 | 144.389898 | 167.697770 | 153.319059 | 176.251063 |
| $\sigma_u^2 = 4.5$ | 340.965781 | 477.641574 | 315.979763 | 453.268744 | 521.592655 | 650.493605 |
| $\sigma_u^2 = 6.0$ | 236.006213 | 605.414741 | 277.237814 | 647.609167 | 385.797868 | 751.716660 |
| $\sigma_u^2 = 8.0$ | 217.295838 | 1082.203769 | 267.746928 | 1137.929792 | 334.182042 | 1169.656013 |
| $\sigma_u^2 = 10.0$ | 364.077884 | 1927.985594 | 164.195051 | 1764.231817 | 45.362629 | 1644.894215 |

For MBE, we observe that the differences for both $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ are both all negative and decreasing in σ_u^2 . As the MBE for $\hat{\sigma}_{\epsilon OLS}^2$ is positive for both normal errors and Student's t errors, this means that the MBE for $\hat{\sigma}_{\epsilon OLS}^2$ is closer to zero for Student's t errors. However, for $\hat{\sigma}_{\epsilon MME}^2$, as the MBE is negative for both types of errors, this means that $\hat{\sigma}_{\epsilon MME}^2$ has a mean bias further away from zero under Student's t errors.

For MSE, we observe that the difference is positive and increasing in σ_u^2 . The difference for $\hat{\sigma}_{\epsilon MME}^2$ is greater than that of $\hat{\sigma}_{\epsilon OLS}^2$ for the same values of σ_u^2 and σ_ϵ^2 .

4.2.4 χ^2 -distributed observation error

Lastly, we perform the simulations on estimating σ_ϵ^2 under re-centered χ^2 errors. The estimation MBE and MSE are summarized in Tables 22 and 23 respectively.

Table 22: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under χ^2 -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 5.846227 | -0.107955 | 5.852595 | -0.097985 | 5.956031 | 0.001673 |
| $\sigma_u^2 = 2.0$ | 7.916757 | 0.010683 | 7.853233 | -0.059238 | 7.983631 | 0.070941 |
| $\sigma_u^2 = 2.5$ | 9.848993 | -0.016529 | 9.828657 | -0.039154 | 9.799226 | -0.076836 |
| $\sigma_u^2 = 3.0$ | 11.705806 | -0.108001 | 11.658828 | -0.148038 | 11.752845 | -0.055206 |
| $\sigma_u^2 = 4.5$ | 17.374873 | -0.210306 | 17.373310 | -0.183002 | 17.402963 | -0.193137 |
| $\sigma_u^2 = 6.0$ | 22.814673 | -0.435660 | 22.796974 | -0.457350 | 23.039812 | -0.174432 |
| $\sigma_u^2 = 8.0$ | 30.256286 | -0.377694 | 30.314447 | -0.328371 | 30.343469 | -0.340969 |
| $\sigma_u^2 = 10.0$ | 37.120092 | -0.841162 | 37.539161 | -0.432518 | 37.117431 | -0.840077 |

Table 23: Estimation MSE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under χ^2 -distributed observation error with different variances σ_u^2 and residual error with different variances σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 56.563197 | 22.996346 | 57.309948 | 23.677762 | 60.483599 | 25.715861 |
| $\sigma_u^2 = 2.0$ | 93.680355 | 32.217787 | 91.819179 | 31.360443 | 98.074845 | 35.627879 |
| $\sigma_u^2 = 2.5$ | 137.435981 | 42.521144 | 137.176570 | 42.563518 | 136.832234 | 42.626278 |
| $\sigma_u^2 = 3.0$ | 184.303172 | 50.044663 | 183.957501 | 50.917001 | 188.524146 | 53.241392 |
| $\sigma_u^2 = 4.5$ | 377.787401 | 82.590910 | 379.565750 | 84.655555 | 381.000349 | 84.963603 |
| $\sigma_u^2 = 6.0$ | 625.109020 | 116.664481 | 623.575613 | 117.049269 | 640.975827 | 123.237757 |
| $\sigma_u^2 = 8.0$ | 1067.198766 | 175.652108 | 1074.095498 | 179.207401 | 1072.267645 | 174.251989 |
| $\sigma_u^2 = 10.0$ | 1571.092644 | 232.458616 | 1612.083101 | 239.745418 | 1581.926597 | 242.492254 |

We again observe similar properties to the normal and Student's t cases: For MBE, the MBE of $\hat{\sigma}_{\epsilon OLS}^2$ is positive and increasing in σ_u^2 , and that of $\hat{\sigma}_{\epsilon MME}^2$ is negative and decreasing in σ_u^2 . For MSE, the MSE of both $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ are both increasing in σ_u^2 , and the MSE of $\hat{\sigma}_{\epsilon MME}^2$ is smaller than that of $\hat{\sigma}_{\epsilon OLS}^2$.

The MBE and MSE results after taking difference with the normal case are shown in Tables 24 and 25 respectively.

For the difference in MBE between the χ^2 -distributed and normal-distributed

Table 24: Difference between χ^2 -distributed error case and normal-distributed error case: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | -0.121868 | -0.126673 | -0.095868 | -0.097683 | 0.034691 | 0.032430 |
| $\sigma_u^2 = 2.0$ | -0.000769 | -0.000073 | -0.003602 | -0.007841 | 0.046975 | 0.040955 |
| $\sigma_u^2 = 2.5$ | 0.015169 | 0.013430 | -0.029604 | -0.039207 | -0.047118 | -0.059946 |
| $\sigma_u^2 = 3.0$ | -0.089071 | -0.102870 | -0.124525 | -0.129951 | -0.025971 | -0.028909 |
| $\sigma_u^2 = 4.5$ | -0.183899 | -0.197984 | -0.127204 | -0.129226 | -0.057467 | -0.091276 |
| $\sigma_u^2 = 6.0$ | -0.303190 | -0.318269 | -0.285091 | -0.299563 | 0.016387 | -0.011594 |
| $\sigma_u^2 = 8.0$ | -0.189109 | -0.164376 | -0.035383 | -0.025541 | 0.085655 | -0.003566 |
| $\sigma_u^2 = 10.0$ | -0.359106 | -0.367568 | -0.118853 | -0.208896 | -0.344515 | -0.423934 |

Table 25: Difference between χ^2 -distributed error case and normal-distributed error case: Estimation MBE of $\hat{\sigma}_{\epsilon OLS}^2$ and $\hat{\sigma}_{\epsilon MME}^2$ under different σ_u^2 and σ_ϵ^2 .

| | $\sigma_\epsilon^2 = 1.5$ | | $\sigma_\epsilon^2 = 2.0$ | | $\sigma_\epsilon^2 = 2.5$ | |
|---------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ | $\hat{\sigma}_{\epsilon OLS}^2$ | $\hat{\sigma}_{\epsilon MME}^2$ |
| $\sigma_u^2 = 1.5$ | 17.123633 | 19.105339 | 17.503852 | 19.194230 | 20.608246 | 20.832436 |
| $\sigma_u^2 = 2.0$ | 24.857297 | 25.933181 | 23.392068 | 24.529332 | 27.502203 | 27.901681 |
| $\sigma_u^2 = 2.5$ | 31.597691 | 33.130699 | 30.194104 | 32.498510 | 29.345670 | 31.759968 |
| $\sigma_u^2 = 3.0$ | 32.975182 | 37.525819 | 32.044992 | 37.391702 | 35.729896 | 38.775078 |
| $\sigma_u^2 = 4.5$ | 44.576956 | 56.048866 | 47.821178 | 57.611808 | 49.167870 | 56.316500 |
| $\sigma_u^2 = 6.0$ | 48.857455 | 71.243179 | 46.908839 | 70.395988 | 65.226202 | 74.410970 |
| $\sigma_u^2 = 8.0$ | 70.609430 | 98.734220 | 80.935690 | 98.421423 | 83.344002 | 92.865456 |
| $\sigma_u^2 = 10.0$ | 61.902354 | 114.062432 | 85.163824 | 116.435787 | 68.716872 | 116.415571 |

cases, we do not see a clear monotonic trend in the difference value, and the sign fluctuates. For the difference in MSE, we observe that the difference is always positive, meaning that the MSE under χ^2 -distributed errors is greater than that under normal-distributed errors. This difference is increasing in σ_u^2 , and the difference when using $\hat{\sigma}_{\epsilon MME}^2$ is greater than that for $\hat{\sigma}_{\epsilon OLS}^2$.

5 Conclusion

We have presented some empirical results on the case of non-normality under the errors-in-variables model, and have shown how non-normality in the observation error affects the estimation of the regression parameters β and σ_ϵ^2 . We have compared the errors in different metrics when using different distributions and different estimators for $\hat{\beta}$ and $\hat{\sigma}_\epsilon^2$.

Future work can be extended upon this topic. We have only presented the empirical results arising from problem of estimation, and the problem of inference of statistical parameters has not been explored in this work. Exact inference of statistical parameters under non-normality can also be explored. The effects of non-normality on multivariate errors-in-variables models and other estimators can also be explored in the future.

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7 Appendix

The computer code used in this project can be found at <https://github.com/jevril/fyp>.

```
var_u_list = [0, 1.5, 2, 2.5, 3.0, 4.5, 6.0, 8.0, 10.0]
var_eps_list = [1.5, 2, 2.5]

def experiment(distribution_label, gen_err_u, gen_err_eps):
    print(f"{distribution_label}_distributed_epsilon")

    beta_dict = {}
    sigma_dict = {}

    for var_u in var_u_list:
        for var_eps in var_eps_list:
            bias_beta = np.array([])
            sqerr_beta = np.array([])
            bias_beta_adj = np.array([])
            sqerr_beta_adj = np.array([])

            bias_sigma = np.array([])
            sqerr_sigma = np.array([])
            bias_sigma_adj_v2 = np.array([])
            sqerr_sigma_adj_v2 = np.array([])

            for _ in range(iterations):
                x = np.linspace(-20, 20, n)
```

```

if var_u > 0:
    x_obs = x + gen_err_u(var_u, n)
else:
    x_obs = x
y = x*secret_beta + gen_err_x(var_eps, n)

beta_est = np.cov(x_obs, y)[0][1]/np.var(x_obs, ddof=1)
beta_est_adj = np.cov(x_obs, y)[0][1]/np.var(x_obs, ddof=1) * np

vx = np.var(x_obs, ddof=1) - var_u
vu = var_u
l = vx/(vx+vu);

sigma_est = np.sum((y - beta_est * x_obs) ** 2) / (n-1)
sigma_est_adj_v2 = np.sum((y - beta_est * x_obs) ** 2) / (n-1) -

bias_beta = np.append(bias_beta, (beta_est-secret_beta))
sqerr_beta = np.append(sqerr_beta, (beta_est-secret_beta)**2)
bias_beta_adj = np.append(bias_beta_adj, (beta_est_adj-secret_beta))
sqerr_beta_adj = np.append(sqerr_beta_adj, (beta_est_adj-secret_beta)**2)

bias_sigma = np.append(bias_sigma, sigma_est-var_eps)
sqerr_sigma = np.append(sqerr_sigma, (sigma_est-var_eps)**2)
bias_sigma_adj_v2 = np.append(bias_sigma_adj_v2, sigma_est_adj_v2-var_eps)
sqerr_sigma_adj_v2 = np.append(sqerr_sigma_adj_v2, (sigma_est_adj_v2-var_eps)**2)

print("var_u=%.2f, var_eps=%.2f: OLS: MBE=%%f, MSE=%%f; Sigma_MBE=%%f, Sigma_MSE=%%f")

print("var_u=%.2f, var_eps=%.2f: MME: MBE=%%f, MSE=%%f; Sigma_MBE=%%f, Sigma_MSE=%%f")

beta_dict[(var_u, var_eps)] = {}
beta_dict[(var_u, var_eps)][ 'bias_ols' ] = bias_beta.mean()
beta_dict[(var_u, var_eps)][ 'sqerr_ols' ] = sqerr_beta.mean()
beta_dict[(var_u, var_eps)][ 'bias_ols_corr' ] = bias_beta_adj.mean()
beta_dict[(var_u, var_eps)][ 'sqerr_ols_corr' ] = sqerr_beta_adj.mean()

sigma_dict[(var_u, var_eps)] = {}
sigma_dict[(var_u, var_eps)][ 'bias_ols' ] = bias_sigma.mean()
sigma_dict[(var_u, var_eps)][ 'sqerr_ols' ] = sqerr_sigma.mean()
sigma_dict[(var_u, var_eps)][ 'bias_ols_corr_v2' ] = bias_sigma_adj_v2.mean()
sigma_dict[(var_u, var_eps)][ 'sqerr_ols_corr_v2' ] = sqerr_sigma_adj_v2.mean()

return beta_dict, sigma_dict

```

References

- [1] Pischke S. Lecture Notes on Measurement Error. 2007.
- [2] Gillard J. Method of Moments Estimation in Linear Regression with Errors in both Variables. Communications in Statistics - Theory and Methods. 2014;43(15):320822.