Maximum likelihood estimation

For a Gaussian distribution

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Probabilistic approaches in machine learning

- In many applications it is useful to deal with uncertainty
- Probability theory gives a principled way to do this
- Probabilistic perspective often useful for defining and combining loss functions
- Need a way to estimate the parameters in a probabilistic model
- Maximum likelihood estimation is one of the most fundamental methods

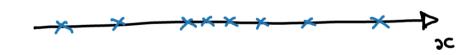
The Gaussian distribution

Maximum likelihood estimation (MLE)

Given samples $x^{(1)}, x^{(2)}, \ldots, x^{(N)}$ from a univariate Gaussian with unknown mean and variance, could we devise a way (maybe with a "loss function") to find optimal estimates of the mean $\hat{\mu}$ and variance $\hat{\sigma}^2$?

How would these estimates compare with the sample mean and variance?

We assume the samples are independent and identically distributed (IID), each a draw from the Gaussian $\mathcal{N}(x; \mu, \sigma^2)$.



MLE for univariate Gaussian

Given IID samples { x m } = each sample a draw from n (x; M,52).

Joint density:

b(x,,, ..., x,,) = N(x,, h'e,, x) x (x,, h'e,, x ... x M(x,, h'e,) Some settings of (µ,5²) will give high value on data, others low. = I W (x,, 'W'es)

Idea: Choose (M,02) that maximises this, i.e. ψ, ξ, = σιδ ωσχ <u>μ</u> ω (κ,ω); Μ, ες)

Called the likelihood of the parameters.

This approach is therefore called maximum likelihood estimation.

Terminology used for any distribution, not just Gaussians.

Instead of maximizing likelihead diractly, it is often easier to maximize log likelihead:

We like minimizing a loss, so let's minimize the negative log likelihood:

$$J(\mu,\sigma^2) = -\log \prod_{n=1}^{N} \mathcal{N}(\infty^n); \mu, \sigma^2)$$

Strategy: Set $\frac{\partial J}{\partial \mu} = 0$ and $\frac{\partial J}{\partial \sigma^2} = 0$
and solve jointly to
find $\hat{\mu}$ and $\hat{\sigma}$.

$$J(\mu_{\vec{p}}) = -\frac{N}{2} \log \left[\frac{1}{2\pi e^{2}} e^{-\frac{(x^{(n)} - \mu)^{2}}{2e^{2}}} \right]$$

$$= -\frac{N}{2} \left[-\frac{1}{2} \log (2\pi e^{2}) - \frac{(x^{(n)} - \mu)^{2}}{2e^{2}} \right]$$

$$= \frac{N}{2} \log (2\pi) + \frac{N}{2} \log^{2} + \frac{1}{2e^{2}} \sum_{n=1}^{N} (x^{(n)} - \mu)^{2}$$

$$= \frac{3}{2e^{2}} \sum_{n=1}^{N} \left[2e^{-\frac{n}{2}} \sum_{n=1}^{N} (x^{(n)} - \mu)^{2} \right]$$

$$= \frac{1}{2e^{2}} \sum_{n=1}^{N} \frac{3}{2} \left[x^{(n)} - \mu^{2} \right]$$

$$= \frac{1}{20^2} \sum_{n=1}^{\infty} (x^{(n)} - y^{(n)}) \cdot (-1)$$

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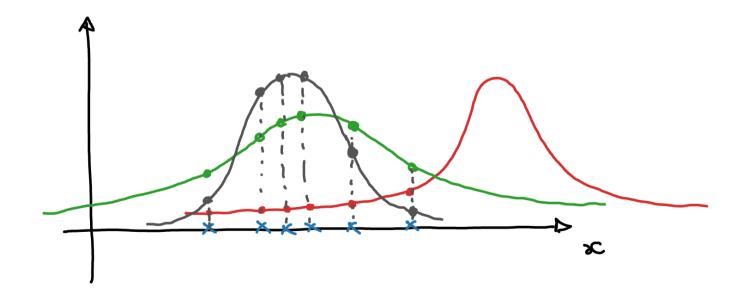
 $\frac{\partial \sigma^2}{\partial x} = 0: \left(\frac{\partial \sigma^2}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$ $\frac{\partial \sigma}{\partial x} = 0: \left(\frac{\partial \sigma}{\partial x} \right) - \frac{\partial \sigma}{\partial x} \left(\frac{\partial \sigma}{\partial x} - \frac{\partial \sigma}{\partial x} \right)$

More about the likelihood

$$\rho\left(\infty^{(1)},...,\infty^{(N)}\right) = \mathcal{N}\left(\infty^{(1)};\mu,\sigma^{2}\right) \times \mathcal{N}\left(\infty^{(2)};\mu,\sigma^{2}\right) \times ... \times \mathcal{N}\left(\infty^{(N)};\mu,\sigma^{2}\right)$$

$$= \mathcal{N}\left(\infty^{(N)};\mu,\sigma^{2}\right)$$

$$=$$



MLE for the multivariate Gaussian

Given samples $\left\{\mathbf{x}^{(n)}\right\}_{n=1}^{N}$ from a multivariate Gaussian:

$$x \in \mathbb{R}_p$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\}$$

it can be shown in a similar way that the maximum likelihood estimates are:

$$\hat{\boldsymbol{\mu}} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}^{(n)}$$

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}^{(n)} - \hat{\boldsymbol{\mu}}) (\mathbf{x}^{(n)} - \hat{\boldsymbol{\mu}})^{\top}$$

Need vector and matrix derivatives to derive this

