

14.385 Nonlinear Econometric Analysis  
(Causal) matrix completion

Maximilian Kasy

Department of Economics, MIT

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# Outline

- Setup: Filling in missing entries in a matrix  $Y$ .
  1. Recommender systems.
  2. Counterfactual outcomes in panel data.
- Recap: Singular value decompositions and principal components.
- Empirical risk minimization methods.
- Nearest neighbor methods.
- Missigness assumptions.
- Reweighting.
- The synthetic nearest neighbor algorithm.

## Takeaways for this part of class

- Typical assumption:  
 $Y$  is a sum of a low rank matrix  $A$  and idiosyncratic noise  $E$ .
- Any matrix has a singular value decomposition.  
Principal components correspond to the largest singular values.
- Popular methods for matrix completion:
  1. *Empirical risk minimization.*
  2. *Nearest neighbors.*
- Standard methods suffer from bias with non-random missingness.
- For (conditionally) missing at random data, reweighting can provide a solution.
- An algorithm for more general missingness is *Synthetic nearest neighbors*.

Setup

Standard algorithms

Missigness assumptions

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References

# Setup

- Random matrices with  $m$  rows,  $n$  columns:
  - Latent matrix  $A$ .
  - Error matrix  $E$ , where  $E[E|A] = 0$ .
  - Outcome matrix  $Y = A + E$ , thus  $E[Y|A] = A$ .
  - Observability matrix  $D$ .
  - Probability of observability:  $P = E[D]$ .
- Goal: Estimate the entries of the latent matrix  $A$  based on observations  $(Y_{ij} \cdot D_{ij}, D_{ij})$ .
- Typical loss function:

$$\frac{1}{m \cdot n} \sum_{ij} (\hat{A}_{ij} - A_{ij})^2.$$

# Interpretations

- Recommender systems:
  - Rows  $i$  index individuals, columns  $j$  index movies.
  - $Y_{ij}$  are movie ratings by individual  $i$  for movie  $j$ .
  - $D_{ij}$  are indicators for whether a movie was rated by an individual.
  - Goal: Recommend movies that would receive high ratings.
- Panel data causal inference (cf. synthetic controls):
  - Rows  $i$  index cross-sectional units, columns  $j$  index time-periods.
  - $Y_{ij}$  are potential outcomes absent treatment.
  - $D_{ij}$  are indicators for untreated units.
  - Goal: Recover missing  $Y_{ij}$  to recover causal effects.

Setup

Standard algorithms

Missigness assumptions

Synthetic nearest neighbors

References

## Reminder: Singular value decomposition

- Any real valued matrix  $A$  with  $m$  rows,  $n$  columns, rank  $k = \text{rank}(A)$  can be decomposed as

$$A = U \cdot \Sigma \cdot V = \sum_{l=1}^k \sigma_l \cdot u_l \cdot v_l'.$$

- $U$  is an  $m \times k$  matrix with orthonormal columns  $u_l$ .
- $V$  is an  $n \times k$  matrix with orthonormal columns  $v_l$ .
- $\Sigma$  is a  $k \times k$  diagonal matrix with entries  $\sigma_l$  of decreasing magnitude.
- Special case: Diagonalization of square matrices  $A$ .
- Consider the largest singular values  $\sigma_l$ :
  - Principal components:  $\sigma_l \cdot u_l$ .
  - Low-rank approximation:  $A \approx \sum_{l=1}^{\kappa} \sigma_l \cdot u_l \cdot v_l'$ , where  $\kappa < k$ .



# Empirical risk minimization (ERM) methods

- Minimize average prediction error for observed outcomes:

$$\hat{A} = \operatorname{argmin}_a \sum_{i,j} D_{ij} \cdot (a_{ij} - Y_{ij})^2 + \lambda \cdot \operatorname{Reg}(a).$$

- Here **Reg** is one of several possible regularization penalties,  $\lambda$  is a tuning parameter.
- Popular choice: Nuclear norm (or trace norm).

$$\operatorname{Reg}(a) = \operatorname{tr}(\sqrt{a' \cdot a}) = \sum_l \sigma_l(a).$$

The  $\sigma_l(a)$  are the singular values of  $a$ .

⇒ Lasso penalty for the singular values of  $\hat{A}$ .

⇒ *SoftImpute* algorithm

- Variant: Rather than penalizing  $\hat{A}$ , constrain  $\hat{A}$  to be low rank.

## Nearest neighbor methods

- Consider a specific  $i, j$  with  $D_{ij} = 0$ .
- Find a set  $\mathcal{I}$  of  $k$  rows  $i'$ , such that
  1.  $D_{i'j} \neq 0$ .
  2. Row  $i'$  is “similar” to row  $i$ .
- “Similar” often means a small distance of the vector of observed values,

$$\sum_{j'} D_{ij'} \cdot D_{i'j'} \cdot (Y_{ij'} - Y_{i'j'})^2.$$

- Impute an estimate for  $Y_{ij}$  as

$$\hat{Y}_{ij} = \frac{1}{k} \sum_{i' \in \mathcal{I}} Y_{i'j}.$$

Setup

Standard algorithms

Missigness assumptions

Synthetic nearest neighbors

References

# Missingness assumptions

## 1. **Missing completely at random** (MCAR):

$$D_{ij}|Y \sim^{iid} \text{Ber}(p).$$

## 2. **Missing at random** (MAR):

$$D_{ij}|Y, X \sim \text{Ber}(P_{ij}(X)),$$

independently across  $i, j$ ,

where  $X$  are observable controls, and  $P_{ij}(X) > 0$ .

## 3. **Missing not at random** (MNAR):

- $D$  and  $Y$  are not independent,
- $D_{ij}$  and  $D_{i'j'}$  are not independent,
- $P_{ij} = 0$  is allowed.

## Practice problem

- Suppose MCAR holds.  
Consider any empirical risk minimization (ERM) algorithm.  
What is the expectation of the objective function for such an algorithm?
- Suppose MAR holds.  
How could you modify empirical risk minimization, to avoid biases?

# Reweighting under MAR

- Suppose that  $\mathbf{P}$  takes the form

$$P_{ij} = g(X_i \cdot \beta_x + W_j \cdot \beta_w + \delta_i + \gamma_j),$$

where  $g(\cdot)$  is a link function; e.g. the logistic  $g(x) = \frac{\exp(x)}{1+\exp(x)}$ .

- We can estimate  $\mathbf{P}$  by logistic regression of  $D_{ij}$  on  $X_i$ ,  $W_j$ , and row and column fixed effects.
- Reweighted ERM:

$$\hat{\mathbf{A}} = \operatorname{argmin}_a \sum_{i,j} \frac{D_{ij}}{\hat{P}_{ij}} \cdot (a_{ij} - Y_{ij})^2 + \lambda \cdot \operatorname{Reg}(a).$$

Setup

Standard algorithms

Missigness assumptions

Synthetic nearest neighbors

References

# A restricted form of MNAR

Assumptions:

1. **Low rank factor model:**

$\text{rank}(A) = k < \min(m, n)$ , so that

$$A = \sum_{l=1}^k \sigma_l \cdot u_l \cdot v_l'.$$

2. **Selection on latent factors:**

$$E[E|U, V, D] = 0.$$

3. **Linear span inclusion:**

Any set of  $k$  rows of  $U$  has full rank.



# Identification

- Assumption 3 could be weakened, but holds generically.
- Immediate implication of these assumptions:
  - Fix a pair  $(i,j)$ .

- Let  $\mathcal{J}$  be such that  $D_{i'j} = 1$  for all  $i' \in \mathcal{J}$  and  $|\mathcal{J}| \geq k$ .

- Then there is a  $\beta$  such that

$$u_{i,.} = \sum_{i' \in \mathcal{J}} \beta_{i'} \cdot u_{i',.}$$

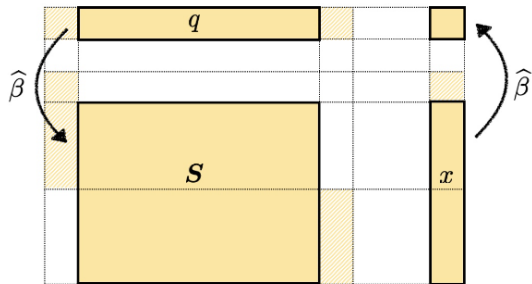
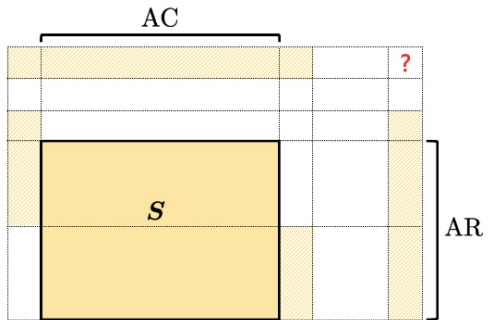
- Furthermore,

$$A_{ij} = \sum_{i' \in \mathcal{J}} \beta_{i'} \cdot E[Y_{i'j} | U, V, D].$$

- Thus:
  - Suppose we could estimate  $\beta$ .
  - Then we could impute

$$\hat{A}_{ij} = \sum_{i' \in \mathcal{J}} \beta_{i'} \cdot Y_{i'j}.$$

# Synthetic nearest neighbors



## Synthetic nearest neighbors (1)

Algorithm proposed by Agarwal et al. (2021);

1. Fix tuning parameter  $\kappa \in \mathbb{N}$  (rank of approximations).
2. Consider some  $(i, j)$  for which we want to estimate  $\mathbf{A}_{ij}$ .
3. Find a set of rows and columns  $\mathbf{AR}$  and  $\mathbf{AC}$  such that

$$D_{i'j'} = D_{i'j} = D_{ij'} = 1$$

for all  $i' \in \mathbf{AR}$  and  $j' \in \mathbf{AC}$ .

Let  $\mathbf{S}$  be the submatrix of  $\mathbf{Y}$  corresponding to rows  $\mathbf{AR}$ , columns  $\mathbf{AC}$ .

4. Find the singular value decomposition

$$\mathbf{S} = \sum_{l=1}^k \sigma_l \cdot \hat{\mathbf{u}}_l \cdot \hat{\mathbf{v}}_l'$$

## Synthetic nearest neighbors (2)

### 5. Estimate

$$\hat{\beta} = \left( \sum_{l=1}^{\kappa} \frac{1}{\sigma_l} \cdot \hat{u}_l \cdot \hat{v}_l' \right) \cdot \mathbf{A}_{i,AC}.$$

(Note we are truncating the sum at  $\kappa$ .)

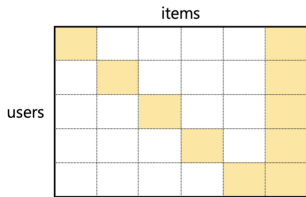
### 6. Impute

$$\hat{A}_{ij} = \sum_{i' \in \mathcal{I}} \beta_{i'} \cdot Y_{i'j}.$$

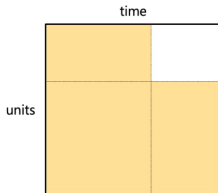
### 7. Repeat for different rows $\mathbf{AC}$ , columns $\mathbf{AR}$ , and average.

Role of columns and rows could be switched, without affecting the estimate.

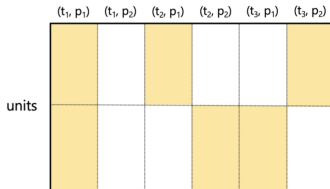
# Typical patterns of missingness



(a) Recommender systems.



(b) Panel data.



(c) Sequential decision-making.

# References

- *Athey, S., Bayati, M., Doudchenko, N., Imbens, G., and Khosravi, K. (2021). Matrix completion methods for causal panel data models. Journal of the American Statistical Association, 116(536):1716–1730*
- *Agarwal, A., Dahleh, M., Shah, D., and Shen, D. (2021). Causal matrix completion. arXiv preprint arXiv:2109.15154*