

# Multiple testing

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## Multiple Hypothesis Testing for differential expression detection

- The test statistics and hence the p-values are likely correlated due to co-regulation of the genes.
- Would like multiple testing procedures that take into account the dependence structure of the genes.
- This could be accomplished by estimating the joint null distribution of the unadjusted, unknown p-values.

## Multiple testing problem

- With thousands of genes on a microarray we're not testing one hypothesis, but many hypotheses – one for each gene
- Analysis of 20,000 genes using commonly accepted significance level  $\alpha = 0.05$  will identify 1,000 differentially expressed genes simply by chance
- If probability of making an error in one test is 0.05, probability of making at least one error in ten tests is

$$(1 - (1 - 0.05)^{10}) = 0.40126$$

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## Permutation based methods

### Permutation based adjusted p-values

- Under the  $H_0$ , the joint distribution of the test statistics can be estimated by permuting the columns of the gene expression matrix
- Permuting entire columns creates a situation in which membership to the groups being compared is independent of gene expression but preserves the dependence structure between genes

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## Permutation based methods

- Permutation algorithm for the  $b^{th}$  permutation,  $b = 1, \dots, B$ 
  1. Permute the  $n$  columns of the data matrix  $X$
  2. Compute test statistics  $t_{j,b}$  for each hypothesis (gene,  $j = 1, \dots, g$ )
- The permutation distribution of the test statistic  $T_j$  for hypothesis  $H_j$  is given by the empirical distribution of  $t_{j,1}, \dots, t_{j,B}$

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## Permutation based methods

- For two-sided alternative hypotheses, the permutation p-value for hypothesis  $H_j$  is

$$p_j^* = \frac{\sum_{b=1}^B I(|t_{j,b}| \geq |t_j|)}{B}$$

where  $I(*)$  is the indicator function, equaling 1 if the condition in parentheses is true and 0 otherwise.

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## Permutation based methods

- Permutation method permits estimation of the joint null distribution of the unadjusted unknown p-values.
- Dependency structure between the genes is preserved.
- May suffer from a granularity problem (when two groups, should have 6 arrays in each group to use permutation based method).

$n/n1!n2!$  ways of forming two groups

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## Results of Multiple hypothesis testing

Assume we are testing  $H_1, H_2, \dots, H_m$ .  $m_0$  - # of true null hypotheses

	# false null hypo.	# true null hypo.	
# non-signif.	U	T	$m - R$
# significant	V	S	R
	$m_0$	$m - m_0$	

- **U, S** - True negatives/positives \*unobservable random variable
- **V** - False positives [Type I errors] \*
- **T** - False negatives [Type II errors] \*
- **R** - All positives (# of rejected null hypotheses) Observable

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## Error rates

False Discovery rate (FDR)

$$E \left[ \frac{\text{False Discoveries}}{\text{True Discoveries}} \right]$$

Family wise error rate (FWER)

$$Pr(\text{Number of False positives} \geq 1)$$

Expected number of false positives

$$E[\text{Number of False positives}]$$

## Interpretation

Suppose 550 out of 10,000 genes are significant at  $\alpha = 0.05$

**P-value < 0.05**

- Expect  $0.05 * 10,000 = 500$  false positives

**False Discovery Rate < 0.05**

- Expect  $0.05 * 550 = 27.5$  false positives

**Family Wise Error Rate < 0.05**

- The probability of at least 1 false positive is  $\leq 0.05$

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## Multiple Hypothesis Testing: FWER

- Given  $p$  is the probability of error,  $1 - p$  is the probability of correct choice in one test
- $1 - (1 - p)^g$  is the probability of one error in  $g$  tests

## Multiple Hypothesis Testing: FWER

- Given  $p$  is the probability of error,  $1 - p$  is the probability of correct choice in one test
- $1 - (1 - p)^g$  is the probability of one error in  $g$  tests

**Sidak single step**

- Testing  $g$  null hypotheses
- Reject any  $H_i$  with  $p \leq 1 - \sqrt[g]{1 - \alpha}$
- When testing 22,283 genes for differential expression, use the following cutoff:

$$1 - \sqrt[22,283]{1 - 0.05} = 0.000002302$$

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## Multiple Hypothesis Testing: FWER

### Bonferroni procedure

- Testing  $g$  null hypothesis
- Reject any  $H_i$  with  $p_i \leq \alpha/g$
- $0.05/22,283 = 0.0000022$

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## Multiple Hypothesis Testing: FWER

### Bonferroni procedure

- Testing  $g$  null hypothesis
- Reject any  $H_i$  with  $p_i \leq \alpha/g$
- $0.05/22,283 = 0.0000022$
- Controls the FWER to be  $\leq \alpha$  and to be equal to  $\alpha$  if all hypotheses are true.
- As the number of hypotheses increases, the average power for an individual hypothesis decreases
- Very conservative; no attempt to incorporate dependence between tests

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## Multiple Hypothesis Testing: FWER

### Holm step-down procedure

1. Order the p-values and hypotheses  $P_1 \geq \dots \geq P_g$  corresponding to  $H_1, \dots, H_g$
2. Let  $i = 1$
3. If  $P_{g-i+1} > \alpha/(g - i + 1)$  then accept all remaining hypotheses  $H_{g-i+1}$  and STOP
4. If  $P_{g-i+1} \leq \alpha/(g - i + 1)$  then reject  $H_{g-i+1}$  and increment  $i$ , then return to step 3.

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## Multiple Hypothesis Testing: FWER

### Sidak step down

1. Order the p-values and hypotheses  $P_1 \geq \dots \geq P_g$  corresponding to  $H_1, \dots, H_g$
2. Let  $i = 1$
3. If  $P_{g-i+1} > 1 - \sqrt[g-i+1]{1 - \alpha}$  then accept all remaining hypotheses  $H_{g-i+1}$  and STOP
4. If  $P_{g-i+1} \leq 1 - \sqrt[g-i+1]{1 - \alpha}$  then reject  $H_{g-i+1}$  and increment  $i$ , then return to step 3.

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## Multiple Hypothesis Testing: FWER

### Hochberg step up

1. Order the p-values and hypotheses  $P_1 \geq \dots \geq P_g$  corresponding to  $H_1, \dots, H_g$
2. Let  $i = 1$
3. If  $P_i \leq \alpha/i$  then reject all remaining hypotheses  $H_i, \dots, H_g$  and STOP
4. If  $P_i > \alpha/i$  then accept  $H_i$  and increment  $i$ , then return to step 3.

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## Considerations for controlling the FWER

- Approaches that set out to control the FWER seek to control the probability of at least one false positive regardless of the number of hypotheses being tested.
- When the number of hypotheses  $N$  is very large, this may be too strict = too many missed findings.

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## Considerations for controlling the FWER

- Control over FWER is only appropriate in situations where the intent is to identify only a small number of genes that are truly different.
- Otherwise, the severe loss in power in controlling FWER is not justified.

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## False discovery rates: FDR

- It may be more appropriate to emphasize the proportion of false positives among the differentially expressed genes.
- The expectation of this proportion is the false discovery rate (FDR) (Benjamini & Hochberg, 1995)

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## False discovery rate

Benjamini and Hochberg 1995

**Definition:** FDR is the proportion of false positives among all positives

$$FDR = E \left[ \frac{V}{V+S} \right] = E \left[ \frac{V}{R} \right]$$

- Select the desired proportion  $q$ , e.g., 0.1 (10%)
- Rank the p-values  $p_1 \leq p_2 \leq \dots \leq p_m$ .
- Find the largest rank  $i$  such that  $p_i \leq \frac{i}{m} * q$
- Reject null hypotheses corresponding to  $p_1, \dots, p_i$

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## False positive vs. False discovery rates

False positive rate is **the rate at which truly null genes are called significant**

$$FPR \approx \frac{\text{false positives}}{\text{truly null}} = \frac{V}{m_0}$$

False discovery rate is **the rate at which significant genes are truly null**

$$FDR \approx \frac{\text{false positives}}{\text{called significant}} = \frac{V}{R}$$

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## False Discovery Rates

Two procedures for controlling FDR:

- Fix the acceptable FDR level  $\sigma$  a priori, then find a data-dependent threshold so that the  $FDR \geq \sigma$ . (Benjamini & Hochberg)
- Fix the threshold rule and then form an estimate of the FDR whose expectation is  $\geq$  the FDR rule over the significance region. (Storey)

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## Storey's positive FDR (pFDR)

$$BH : FDR = E \left[ \frac{V}{R} | R > 0 \right] p(R > 0)$$

$$Storey : pFDR = E \left[ \frac{V}{R} | R > 0 \right]$$

- Since  $P(R > 0)$  is  $\sim 1$  in most genomics experiments, FDR and pFDR are very similar
- Omitting  $P(R > 0)$  facilitated development of a measure of significance in terms of the FDR for each hypothesis

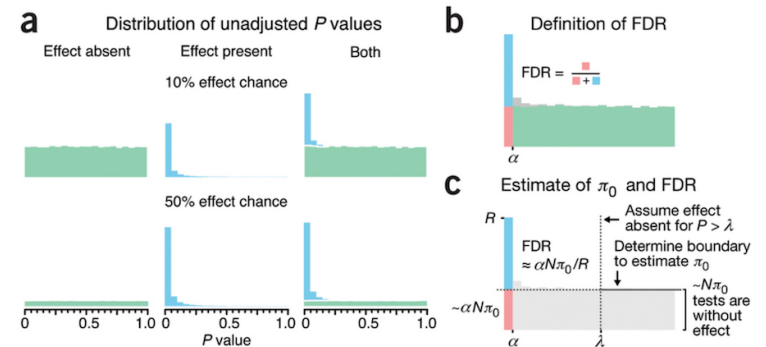
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## Q-value

- Storey & Tibshirani, "Statistical significance for genomewide studies", PNAS, 2003 <http://www.pnas.org/content/100/16/9440.full>
- Empirically derived – uses the p-value distribution
- Storey's method first estimates the fraction of comparisons for which the null is true,  $\pi_0$ , counting the number of  $P$  values larger than a cutoff  $\lambda$  (such as 0.5) relative to  $(1 - \lambda) * N$  (such as  $N/2$ ), the count expected when the distribution is uniform
- Multiply the Benjamini & Hochberg FDR by  $\pi_0$ , thus less conservative

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## Q-value



Martin Krzywinski & Naomi Altman "Points of significance: Comparing samples—part II" *Nature Methods* 2016  
<http://www.nature.com/nmeth/journal/v11/n4/full/nmeth.2900.html>

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## Q-value

- q-value is defined as the minimum FDR that can be attained when calling a "feature" significant (i.e., expected proportion of false positives incurred when calling that feature significant)
- The estimated q-value is a function of the p-value for that test and the distribution of the entire set of p-values from the family of tests being considered
- Thus, in an array study testing for differential expression, if gene X has a q-value of 0.013 it means that 1.3% of genes that show p-values at least as small as gene X are false positives

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