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8-36

To account for length-based selectivity, which is implemented only for the British Columbia data, we applied a penalty to the likelihood function as follows:

Equation 1
$$L(D | \theta) = \frac{S_{L_i} \prod_i \frac{1}{\sqrt{2\pi}\sigma_{a_i}} e^{-\frac{(L_i - \hat{L}_i)^2}{2[\sigma_{a_i}]^2}}}{S_{\hat{L}_i} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_{a_i}} e^{-\frac{(L_i - l)^2}{2[\sigma_{a_i}]^2}} dl}$$

Where L_i is the observed length at a given age a_i , \hat{L}_i is the corresponding estimate based on VBGF parameters θ , S is a logistic selectivity function with parameter L_{50} , the length at which 50% of individuals (male or female) are fully selected, set to 52.976 cm (Samuel Johnson, SFU, pers. comm.)

Equation 2
$$S_L = \frac{1}{1 + \exp(L_{50} - L)}$$

As length-based selectivity is assumed constant in both the California Current and Alaskan assessments, S_L for both estimated and observed lengths is set to 1.0 when fitting data points from those regions.

Original format from AEP:

$$L(D | \theta) = \prod_i \left\{ \frac{S_{L_i}}{\sqrt{2\pi}\sigma_{a_i}} e^{-\frac{(L_i - \hat{L}_i)^2}{2[\sigma_{a_i}]^2}} \right\} \cdot \left\{ \frac{1}{\sqrt{2\pi}\sigma_{a_i}} e^{-\frac{(L_i - l)^2}{2[\sigma_{a_i}]^2}} dl \right\}$$

ENTIRE DIST
↓
EVERY POSSIBLE LENGTH

NO DATA

CALL EXTERNALLY

A BAND OF 25 W/O SEL WINDOW DOESN'T MOVE

TO DOUBLE CHECK INTEGRAL (DENOM) SHOULD BE 1 W/O SELECT

INTEGRATE (TMB)
f. function
a. lower
b. upper
n. SUBDIVS ($2^{N-1} + 1$ EVALS)

JUST ALL AGES

$P(L | a, \hat{L}, \sigma^2, S_L)$
 $P(L | \hat{a}, \hat{L}, \sigma^2, S_L)$

USUAL

$$\frac{SEL(OBS) \cdot DNORM(OBS, EST, \sigma)}{SEL(EST) \cdot DNORM(EST, EST, \sigma)}$$

$$\int \left[\frac{SEL(EST) \cdot DNORM(EST, EST, \sigma)}{SEL(EST) \cdot DNORM(EST, EST, \sigma)} dl \right] \text{ will be 1 w/o SEL}$$

STORE (max 2)
↑ RESCALE EACH TIME

THE THIN AREA OF THIN

$$SEL(S) \cdot DNORM(S)$$

CAN MAKE LOOKUP TABLE @ EACH AGE



MANY FUNC CALLS

30 ET SOME DAY