

Lecture 7: ANCOVA, short introduction to Linear Algebra

BIO144 Data Analysis in Biology

Stefanie Muff, Owen Petchey and Erik Willems

University of Zurich

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Overview

- ▶ ANCOVA (*ANalysis of COVAriance*)
- ▶ Introduction to linear algebra

Course material covered today

- ▶ "Getting Started with R" chapter 6.3
- ▶ "Lineare regression" chapters 3.A (p. 43-45) and 3.4, 3.5 (p. 39-42)

Besherman & Petchey

Stahel

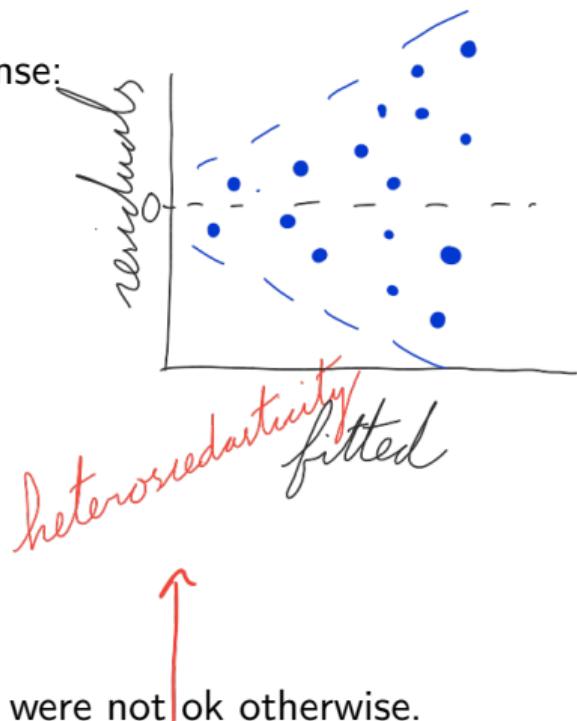
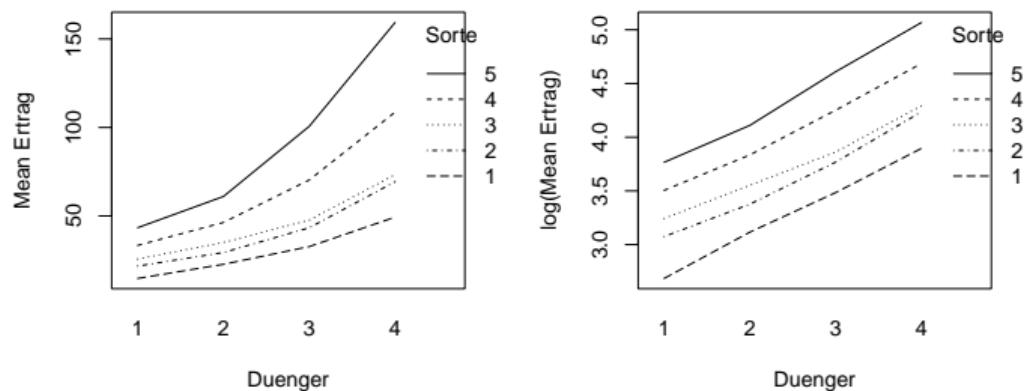
Recap of ANOVA

- ▶ ANOVA is a method to test whether the means of **two or more groups differ**
- ▶ Post-hoc tests and contrasts, including correction for p -values, to understand the differences between the groups
- ▶ Two-way ANOVA for factorial designs, interactions
- ▶ ANOVA: 'linear regression with categorical predictor(s)'
 - ▶ One categorical predictor = one-way ANOVA
 - ▶ Two categorical predictors= two-way ANOVA
 - ▶ etc.

Recap of two-way ANOVA example

The influence of four levels of fertilizer (DUENGER) on the yield (ERTRAG) of 5 crop species (SORTE) was investigated. For each DUENGER \times ERTRAG combination, 3 measurements were made.

Interaction plot with ERTRAG and $\log(\text{ERTRAG})$ as response:



Remember: We used $\log(\text{ERTRAG})$, because residual plots were not ok otherwise.

```
r.duenger2 <- lm(log(ERTRAG) ~ DUENGER*SORTE,d.duenger)
anova(r.duenger2)
```

```
## Analysis of Variance Table
##
## Response: log(ERTRAG)
##             Df  Sum Sq Mean Sq F value Pr(>F)
## DUENGER      3 11.6917  3.8972 854.0505 <2e-16 ***
## SORTE         4   8.5202  2.1300 466.7851 <2e-16 ***
## DUENGER:SORTE 12  0.0929  0.0077   1.6958 0.1045
## Residuals    40  0.1825  0.0046
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Questions:

- ▶ Number of parameters?
- ▶ Degrees of freedom (60 data points)?
- ▶ Interpretation?

↳ interaction can be dropped from model

intercept

4 levels - 1

5 levels - 1

read "bottom-up": if interaction is significant, don't interpret main effects

$\frac{59}{5}$ → intercept

$\frac{60}{6}$

```
##  
## Call:  
## lm(formula = log(ERTRAG) ~ DUENGER * SORTE, data = d.duenger)  
##  
## Residuals:  
##      Min       1Q   Median     3Q    Max  
## -0.120968 -0.045595  0.008984  0.049072  0.102175  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) 2.68505   0.03900 68.846 < 2e-16 ***  
## DUENGER2    0.43165   0.05516  7.826 1.36e-09 ***  
## DUENGER3    0.79997   0.05516 14.504 < 2e-16 ***  
## DUENGER4    1.21152   0.05516 21.966 < 2e-16 ***  
## SORTE2      0.38979   0.05516  7.067 1.51e-08 ***  
## SORTE3      0.55799   0.05516 10.117 1.38e-12 ***  
## SORTE4      0.82018   0.05516 14.870 < 2e-16 ***  
## SORTE5      1.08169   0.05516 19.612 < 2e-16 ***  
## DUENGER2:SORTE2 -0.12949   0.07800 -1.660  0.105  
## DUENGER3:SORTE2 -0.10613   0.07800 -1.361  0.181  
## DUENGER4:SORTE2 -0.04924   0.07800 -0.631  0.531  
## DUENGER2:SORTE3 -0.12180   0.07800 -1.562  0.126  
## DUENGER3:SORTE3 -0.18034   0.07800 -2.312  0.026 *  
## DUENGER4:SORTE3 -0.16061   0.07800 -2.059  0.046 *  
## DUENGER2:SORTE4 -0.10138   0.07800 -1.300  0.201  
## DUENGER3:SORTE4 -0.05311   0.07800 -0.681  0.500  
## DUENGER4:SORTE4 -0.02954   0.07800 -0.379  0.707  
## DUENGER2:SORTE5 -0.08779   0.07800 -1.125  0.267  
## DUENGER3:SORTE5  0.04370   0.07800  0.560  0.578  
## DUENGER4:SORTE5  0.09014   0.07800  1.156  0.255  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

interaction

↑ read "bottom up"

Analysis of Covariance

ANCOVA:

- ▶ An extension of ANOVA
- ▶ A method to test whether the means of two or more groups differ, **controlling for the effect of one (or more) continuous covariate(s)**
- ▶ Makes an additional assumption about the "**homogeneity of regression slopes**"
 - ▶ No interaction between the categorical and (any of the) continuous covariate(s)
 - ▶ If there is an interaction, comparing group means becomes uninformative (the model may still be biologically interesting though!)
- ▶ A **linear model** (just like regression and ANOVA)

Given a categorical covariate x_i and a continuous covariate z_i , the ANCOVA equation is:

intercept

$$y_i = \beta_0 + \beta_1 x_i^{(1)} + \dots + \beta_k x_i^{(k)} + \beta_z z_i + \epsilon_i ,$$

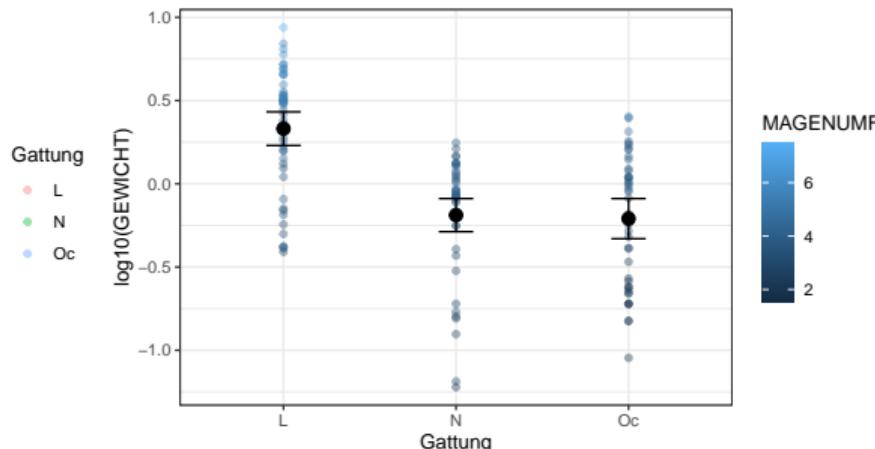
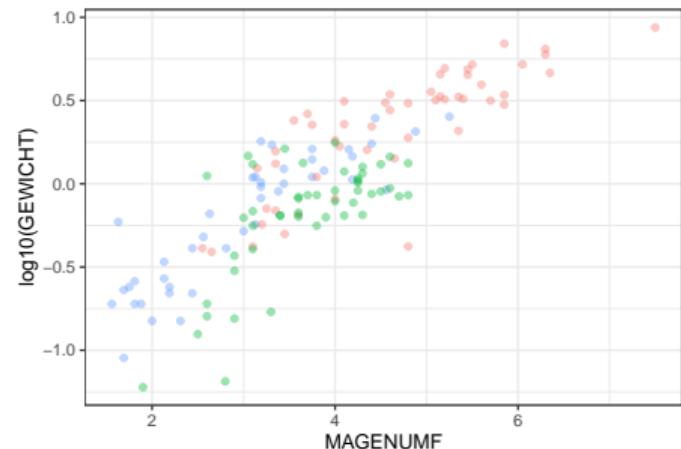
where $x_i^{(k)}$ is the k th dummy variable ($x_i^{(k)}=1$ if i th observation belongs to category k , 0 otherwise).

Note 1: Again, for reasons of identifiability, we typically set $\beta_1 = 0$

Note 2: It is easy to add the interaction between x_i with z_i , but strictly speaking such a model would no longer be an ANCOVA

Once more: the earthworms

“Gewicht” of the worm was expressed as a function of “Magenumfang” and “Gattung”



Categorical and continuous covariates were used to predict a continuous outcome →
ANCOVA?

*↳ compare species means,
controlling for stomachsize*

```
r.lm <- lm(log(GEWICHT) ~ MAGENUMF + Gattung,d.wurm)
summary(r.lm)$coef

##             Estimate Std. Error    t value   Pr(>|t|)
## (Intercept) -2.5355459 0.22147279 -11.4485663 8.617670e-22
## MAGENUMF     0.7118725 0.04528843  15.7186392 1.232126e-32
## GattungN    -0.5151344 0.11009219  -4.6791186 6.760621e-06
## Gattung0c   -0.0907298 0.12791000  -0.7093254 4.793107e-01
```

Important: The p -values for the estimates of (Intercept), GattungN and Gattung0c are not very meaningful (why?).

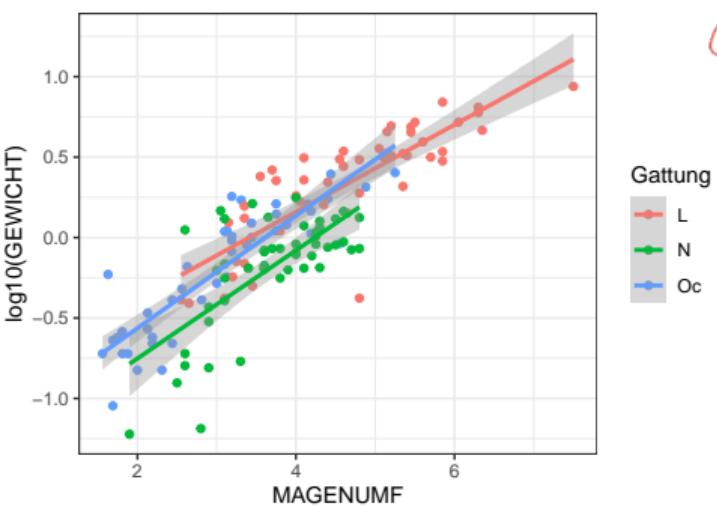
- ↳ - multiple comparisons using same data:
 - * species L vs species N
 - * species L vs species 0c
- intercept = mean of species L when Magenanzfamilie = 0

To understand if “Gattung” has an effect, **we need to carry out an F-test → ANOVA table:**

```
anova(r.lm)
```

```
## Analysis of Variance Table
##
## Response: log(GEWICHT)
##           Df  Sum Sq Mean Sq F value    Pr(>F)
## MAGENUMF     1 104.866 104.866  409.69 < 2.2e-16 ***
## Gattung      2    7.177   3.589   14.02 2.842e-06 ***
## Residuals 139   35.579   0.256
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

To check whether the assumption of *homogeneity of regression slopes* holds, we need to make sure the **interaction** between MAGENUMF and Gattung is not significant:



↙ slopes are similar
 ↓
 seems reasonable,
 but let's formally
 test!

→ We fit a new model, and again use the F -test:

```
r.lm2<- lm(log(GEWICHT) ~ MAGENUMF * Gattung,d.wurm)
anova(r.lm2)
```

```
## Analysis of Variance Table
##
## Response: log(GEWICHT)
##                               Df  Sum Sq Mean Sq F value    Pr(>F)
## MAGENUMF                  1 104.866 104.866 414.4743 < 2.2e-16 ***
## Gattung                   2   7.177   3.589  14.1835 2.521e-06 ***
## MAGENUMF:Gattung          2   0.917   0.458   1.8112   0.1673
## Residuals                 137  34.662   0.253
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

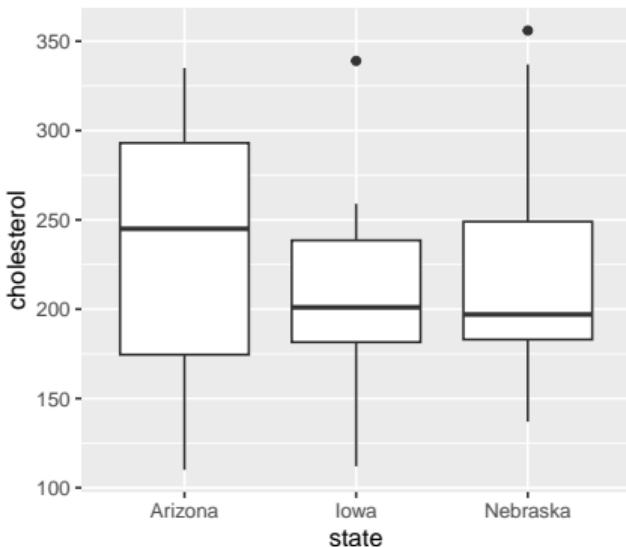


→ $p = 0.167$, the interaction is probably not relevant → ANCOVA makes sense

A new example: cholesterol levels

Example: Cholesterol levels [mg/ml] of 45 women from three US states were measured.

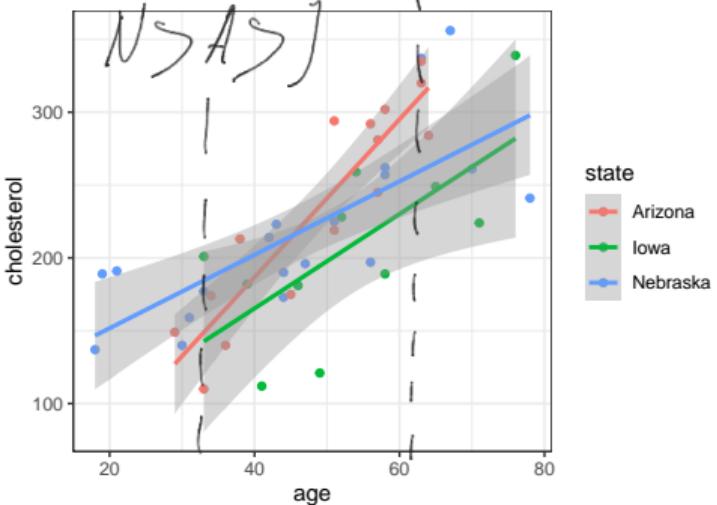
Question: Do these levels differ between the states, controlling for the age (years) of each subject?



The scatter plot already gives us a clue here.

ASNDI

→ which mean is
larger than which
other mean, depends
on the age at will
you do the comparison



→ The slopes look somewhat different, so we include state, age and the interaction between the two into our model.

→ no "homogeneity of regression slopes"

Doing the analysis:

```
r.lm <- lm(cholesterol ~ age * state, data= d.chol)
anova(r.lm)
```

Analysis of Variance Table

##

Response: cholesterol

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
## age	1	96524	96524	61.8961	1.424e-09 ***
## state	2	11474	5737	3.6789	0.03438 *
## age:state	2	12665	6332	4.0606	0.02501 *
## Residuals	39	60819	1559		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

What does this mean?

*comparing means is not informative,
it depends on the age*

Compare the results from the previous slide to the estimated coefficients:

```
r.lm <- lm(cholesterol ~ age*state, data=d.chol)
```

```
summary(r.lm)$coef
```

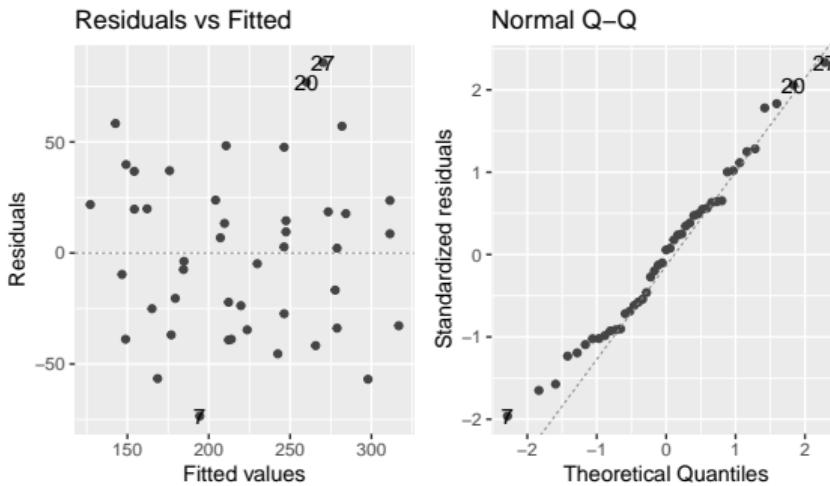
	Estimate	Std. Error	t value	Pr(> t)
## (Intercept)	-29.895169	43.7353712	-0.6835467	4.983027e-01
## age slope in Alaska	5.416908	0.8679635	6.2409400	2.396876e-07
## stateIowa	65.706383	66.7677031	0.9841043	3.311303e-01
## stateNebraska	131.192935	50.8573164	2.5796276	1.377434e-02
## age:stateIowa	-2.178763	1.2672928	-1.7192264	9.350204e-02
## age:stateNebraska	-2.896470	1.0166558	-2.8490174	6.967607e-03

mean cholesterol in Alaska at age = 0

difference in slope with Alaska : $5.41 + -2.18$
 $5.41 + -2.90$

Note: The strength of the association between cholesterol and age is less pronounced in Iowa and Nebraska than in Arizona → no ANCOVA!

As always, some model checking is necessary:



→ This seems ok.

An introduction to linear algebra

Who remembers linear algebra, perhaps from high school?

Overview

- ▶ Some basics about
 - ▶ vectors
 - ▶ matrices
 - ▶ matrix algebra
 - ▶ matrix multiplication
- ▶ Why is linear algebra useful?
- ▶ What does it have to do with data analysis and statistics?
- ▶ Linear models in matrix notation.

Motivation

Why are vectors, matrices and their algebraic rules useful?

- ▶ **Example 1:** The observations for a covariate x or the response y for all individuals $1 \leq i \leq n$ can be stored as a vector:

$$\textcolor{red}{x} = \begin{pmatrix} x_1 \\ x_2 \\ \dots \\ x_n \end{pmatrix}, \quad \textcolor{red}{y} = \begin{pmatrix} y_1 \\ y_2 \\ \dots \\ y_n \end{pmatrix}.$$

columns in spreadsheet!

- ▶ **Example 2:** Covariance matrices for multiple variables. Say we have $x^{(1)}$ and $x^{(2)}$. The **covariance matrix** is then given as:

$$\begin{pmatrix} \textcolor{red}{Var}(x^{(1)}) & \textcolor{red}{Cov}(x^{(1)}, x^{(2)}) \\ \textcolor{red}{Cov}(x^{(1)}, x^{(2)}) & \textcolor{red}{Var}(x^{(2)}) \end{pmatrix}.$$

diagonal

- **Example 3:** The **data** (e.g. of some regression model) can be stored in a **matrix**:

$$\tilde{X} = \begin{pmatrix} \text{Intercept} & \text{Var}_1 & \text{Var}_2 \\ 1 & x_1^{(1)} & x_1^{(2)} \\ 1 & x_2^{(1)} & x_2^{(2)} \\ \dots & \dots & \dots \\ 1 & x_n^{(1)} & x_n^{(2)} \end{pmatrix}.$$

"X title"

This is the so-called **design matrix** with a vector of 1's in the first column.

- **Example 4:** A linear regression model can be written compactly using **matrix multiplication**:

$$y = \tilde{X} \cdot \tilde{\beta} + e,$$

↗ Parameters

with $\tilde{\beta}$ the vector of regression coefficients and e the vector of errors

Why do we discuss this topic in our course?

- ▶ Useful for **compact notation**.
- ▶ Enables you to understand many statistical texts (books, research articles) that remain inaccessible otherwise.
- ▶ Useful for **efficient coding**, e.g. in R, which helps to increase speed and to reduce error rates.
- ▶ More advanced statistical concepts often rely on linear algebra, e.g. **Principal Component Analysis** (PCA) or **random effects** models.
- ▶ Is part of a general education (Allgemeinbildung) ;-)?

Matrices

An $n \times m$ Matrix is given as:

$$A = \begin{pmatrix} a_{11} & a_{12} & \dots & a_{1m} \\ a_{21} & a_{22} & \dots & a_{2m} \\ \vdots & \vdots & & \vdots \\ a_{n1} & a_{n2} & \dots & a_{nm} \end{pmatrix},$$

n rows n columns

with rows $1 = 1, \dots, n$ and columns $j = 1, \dots, m$.

Square matrix: $n = m$. Example:

$$\stackrel{=}{\exists} \begin{pmatrix} 1 & 2 & 3 \\ 4 & 3 & 2 \\ 6 & 1 & 9 \end{pmatrix}$$

Symmetric matrix: $a_{ij} = a_{ji}$. Example:

$$\begin{pmatrix} 1 & 2 & 3 \\ 2 & 3 & 4 \\ 3 & 4 & 5 \end{pmatrix}$$

The matrix is annotated with colored boxes and a red X. The diagonal elements (1, 3, 5) are highlighted with yellow boxes. The off-diagonal elements (2, 3, 4) are highlighted with purple boxes. A large red X is drawn across the entire matrix, indicating it is not symmetric.

The diagonal of a square matrix is given by $(a_{11}, a_{22}, \dots, a_{nn})$. Example: the diagonal of the above matrix is given as

$$(a_{11}, a_{22}, a_{33}) = (1, 3, 5)$$

Diagonal matrix: A matrix that has entries $\neq 0$ only on the diagonal. Example:

$$\begin{pmatrix} 1 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 5 \end{pmatrix}$$

Transposing a matrix: Given a matrix A . Exchange the rows by the columns and vice versa. This leads to the **transposed matrix** A^\top :

$$A = \left(\begin{array}{cccc} a_{11} & a_{12} & \dots & a_{1m} \\ a_{21} & a_{22} & \dots & a_{2m} \\ \vdots & \vdots & & \vdots \\ a_{n1} & a_{n2} & \dots & a_{nm} \end{array} \right) \Rightarrow A^\top = \left(\begin{array}{cccc} a_{11} & a_{21} & \dots & a_{n1} \\ a_{12} & a_{22} & \dots & a_{n2} \\ \vdots & \vdots & & \vdots \\ a_{1m} & a_{2m} & \dots & a_{nm} \end{array} \right)$$

Examples (note the “flip” in dimensions with non-square matrices):

$$A = \left(\begin{array}{cccc} 1 & 2 & 3 & 4 \\ 5 & 6 & 7 & 8 \\ 9 & 10 & 11 & 12 \\ 13 & 14 & 15 & 16 \end{array} \right) \Rightarrow A^\top = \left(\begin{array}{cccc} 1 & 5 & 9 & 13 \\ 2 & 6 & 10 & 14 \\ 3 & 7 & 11 & 15 \\ 4 & 8 & 12 & 16 \end{array} \right)$$

$$A = \left(\begin{array}{cccc} 1 & 2 & 3 & 4 \\ 5 & 6 & 7 & 8 \end{array} \right) \Rightarrow A^\top = \left(\begin{array}{cc} 1 & 5 \\ 2 & 6 \\ 3 & 7 \\ 4 & 8 \end{array} \right)$$

[2 x 4] Matrix

- ▶ Transposing a matrix **twice** leads to the original matrix:

$$\underbrace{(A^T)^T}_{\text{red underline}} = A .$$

- ▶ When a matrix is **symmetric**, then

$$A^T = A .$$

This is true in particular for diagonal matrices.

Vectors

A vector is nothing else than n numbers written in a column:

$$\begin{array}{c} \text{n} \times 1 \text{ Matrix} \\ \curvearrowright \end{array} \quad b = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_n \end{pmatrix}$$

Transposing a vector leads to a *row vector*:

$$\begin{array}{c} \curvearrowright \\ \text{1} \times n \text{ Matrix} \end{array} \quad \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_n \end{pmatrix}^\top = \begin{pmatrix} b_1 & b_2 & \dots & b_n \end{pmatrix}$$

Note: By definition (by default), a vector is always a column vector.

Addition and subtraction

- ▶ Adding and subtracting matrices and vectors is only possible when the objects have the **same dimensions**.
- ▶ Examples: Elementwise addition (or subtraction)

$$\begin{pmatrix} \underline{1} & \underline{2} & 3 \\ 4 & 5 & 6 \end{pmatrix} + \begin{pmatrix} \underline{3} & \underline{2} & 1 \\ 6 & 5 & 4 \end{pmatrix} = \begin{pmatrix} \underline{\frac{4}{10}} & \underline{\frac{4}{10}} & 4 \\ 10 & 10 & 10 \end{pmatrix}$$

$$\begin{pmatrix} 1 \\ \underline{4} \end{pmatrix} - \begin{pmatrix} 3 \\ \underline{9} \end{pmatrix} = \begin{pmatrix} -2 \\ \underline{-5} \end{pmatrix}$$

- ▶ But this addition is **not defined**:

$\begin{pmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{pmatrix} + \begin{pmatrix} 3 & 6 \\ 2 & 5 \\ 1 & 4 \end{pmatrix} = ?!$

$\boxed{[2 \times 3]}$ $\boxed{[3, 2]}$

Multiplication by a scalar

$\rightarrow [? \times ?] \text{ Matrix}$

Multiplication with a “number” (scalar) is simple: Multiply each element in a vector or a matrix.

Examples:

$$[2, 3] \quad [2, 3]$$

$$(3) \cdot \begin{pmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{pmatrix} = \begin{pmatrix} 3 & 6 & 9 \\ 12 & 15 & 18 \end{pmatrix}$$

$$-2 \cdot \begin{pmatrix} 1 \\ 4 \\ -2 \end{pmatrix} = \begin{pmatrix} -2 \\ -8 \\ 4 \end{pmatrix}$$

Matrix multiplication

The multiplication of two matrices A and B is **only defined if**
number of columns in A = number of rows in B .

It is easiest to explain matrix multiplication with an example:

$$\begin{pmatrix} 2 & 1 \\ -1 & 0 \\ 3 & 1 \end{pmatrix} \cdot \begin{pmatrix} 3 & 1 \\ 4 & -2 \end{pmatrix} = \begin{pmatrix} 2 \cdot 3 + 1 \cdot 4 & 2 \cdot 1 + 1 \cdot -2 \\ -1 \cdot 3 + 0 \cdot 4 & -1 \cdot 1 + 0 \cdot -2 \\ 3 \cdot 3 + 1 \cdot 4 & 3 \cdot 1 + 1 \cdot -2 \end{pmatrix}$$

$$= \begin{pmatrix} 10 & 0 \\ -3 & -1 \\ 13 & 1 \end{pmatrix}$$

► Matrix multiplication app

In general:

An $n \times m$ Matrix multiplied by an $m \times p$ Matrix = an $n \times p$ Matrix

Matrix multiplication rules I

Matrix multiplication **does not** follow the same rules as scalar multiplication!!

!

- ▶ The **commutative property** does not hold:
 - ▶ It is possible that $A \cdot B$ can be calculated, whereas $B \cdot A$ is not defined (see example on previous slide).
 ↳ wrong dimensions
 - ▶ In general, $A \cdot B \neq B \cdot A$, even if both are defined.
- ▶ It can happen that $A \cdot B = 0$ (a "zero matrix"), although both $A \neq 0$ and $B \neq 0$.
- ▶ The **associative property** holds: $\underline{A} \cdot (\underline{B} \cdot C) = (\underline{A} \cdot \underline{B}) \cdot \underline{C}$.
- ▶ The **distributive property** holds:

$$\begin{aligned} A \cdot (B + C) &= (A \cdot B) + (A \cdot C) \\ (A + B) \cdot C &= A \cdot C + B \cdot C \end{aligned}$$

Matrix multiplication rules II

- ▶ Transposing inverts the order: $(A \cdot B)^\top = \underline{B^\top} \cdot \underline{A^\top}$.
- ▶ The product $A \cdot A^\top$ is **always symmetric**.
- ▶ All these rules also hold for **vectors**, which can be interpreted as $n \times 1$ matrices:

$$a \cdot b^\top = \begin{pmatrix} a_1 b_1 & a_1 b_2 & \dots & a_1 b_m \\ a_2 b_1 & a_2 b_2 & \dots & a_2 b_m \\ \vdots & \vdots & & \vdots \\ a_n b_1 & a_n b_2 & \dots & a_n b_m \end{pmatrix}$$

Column vector $[n \times 1]$ · Row vector $[1 \times m]$ = $[n \times m]$

If a and b have the **same length**:

$$(1, 2, 3) \cdot \begin{pmatrix} 1 \\ 2 \\ 3 \end{pmatrix} = \underbrace{(1 \cdot 1) + (2 \cdot 2) + (3 \cdot 3)}_{\text{scalar}} = 14$$

$$a^\top \cdot b = \overbrace{\sum_i a_i b_i}^{\text{scalar}} = \text{scalar}$$

Short exercise

$$[n \times m] \cdot [m \times p] = [n \times p]$$



Given vectors a and b and matrix C :

$$a = \begin{pmatrix} 1 \\ -2 \\ 3 \\ 0 \end{pmatrix}, \quad b = \begin{pmatrix} -2 \\ 4 \end{pmatrix}, \quad C = \begin{pmatrix} 1 & 0 \\ -1 & 2 \end{pmatrix}$$

$$\begin{pmatrix} -2 & 4 \\ 4 & -8 \\ -6 & 12 \\ 0 & 0 \end{pmatrix}$$

! not defined!

Calculate, if defined

- ▶ $a^\top \cdot b$ $[1 \times 4] \cdot [2 \times 7] \rightarrow \begin{pmatrix} 1 \\ -2 \\ 3 \\ 0 \end{pmatrix} \cdot \begin{pmatrix} -2 & 4 \end{pmatrix} =$
- ▶ $a \cdot b^\top$ $[4 \times 7] \cdot [7 \times 2] \rightarrow$
- ▶ $C \cdot a$ \leadsto not defined!
- ▶ $C \cdot b$ $\rightarrow \begin{pmatrix} -2 \\ 10 \end{pmatrix}$

The length of a vector

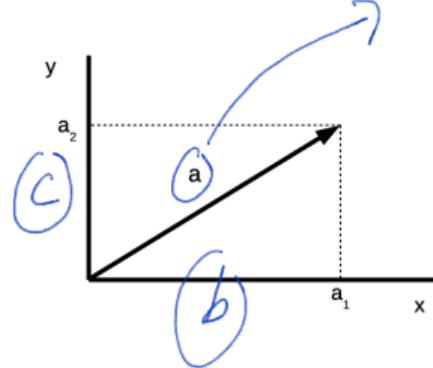
The **length of a vector** $a^\top = (a_1, a_2, \dots, a_n)$ is defined as $\|a\|$ with

$$\|a\|^2 = \underline{a^\top \cdot a} = \boxed{\sum_i a_i^2}.$$

This is basically the **Pythagoras** idea in 2, 3, ... n dimensions.

In 2 dimensions: $\|a\| = \sqrt{a_1^2 + a_2^2}$:

" $a^2 = b^2 + c^2$ "



Identity matrix (Einheitsmatrix)

The identity matrix (of dimension m) is probably the simplest matrix that exists. It has 1's on the diagonal and 0's everywhere else:

$$I = \begin{pmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & & \vdots \\ 0 & 0 & \dots & 1 \end{pmatrix}$$

Multiplication with the identity matrix leaves a $m \times n$ matrix A unchanged:

$$\underbrace{A \cdot I}_{\text{in blue}} = \underbrace{A}_{\text{in blue}}.$$

Inverse matrix

Given a square matrix A that fulfills

$$B \cdot A = I,$$

then B is called the **inverse** of A (and vice versa). One then writes

$$B = A^{-1}.$$

Note:

- ▶ In that case it also holds that $A \cdot B = I$.
- ▶ Therefore: $A = B^{-1} \Leftrightarrow B = A^{-1}$

- ▶ The inverse of A may **not exist**. If it exists, A is **regular**, otherwise **singular**,
- ▶ $(A^{-1})^{-1} = A$.
- ▶ The inverse of a matrix product is given as

$$(A \cdot B)^{-1} = B^{-1} \cdot A^{-1} .$$

- ▶ It is

$$(A^\top)^{-1} = (A^{-1})^\top .$$

Therefore one may also write $A^{-\top}$.

Linear regression in matrix notation

Linear regression with n data points can be understood as an **equation system with n equations**.

Remember the example from slide 21/22: We said that a linear regression model can be written compactly using **matrix multiplication**:

$$y = \tilde{X} \cdot \tilde{\beta} + e.$$

Let's illustrate with a model with two predictor variables $x^{(1)}$ and $x^{(2)}$:

$$y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}, \quad \tilde{X} = \begin{pmatrix} 1 & x_1^{(1)} & x_1^{(2)} \\ 1 & x_2^{(1)} & x_2^{(2)} \\ \dots & \dots & \dots \\ 1 & x_n^{(1)} & x_n^{(2)} \end{pmatrix}, \quad \tilde{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}, \quad e = \begin{pmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{pmatrix}.$$



It can be shown (see Stahel 3.4f,g) that **the least-squares estimates** $\hat{\beta}$ are calculated as:

$$\hat{\beta} = \underline{(\tilde{X}^\top \tilde{X})^{-1} \cdot \tilde{X}^\top \cdot y}$$

Does this look complicated?

Let's have a look in R...

→ It does to me, but a little practice helps !!

Doing linear algebra in R

Let us look at model $y = \tilde{X} \cdot \tilde{\beta} + e$ with coefficients:

$$\beta_0 = 10, \beta_1 = 5, \beta_2 = -2,$$

and variables:

i	$x_i^{(1)}$	$x_i^{(2)}$
1	0	4
2	1	1
3	2	0
4	3	1
5	4	4

Thus the model is given as

$$y_i = 10 + 5x_i^{(1)} - 2x_i^{(2)} + \epsilon_i, \text{ for } 1 \leq i \leq n.$$

Let's start by generating the "true" response, calculated as $\tilde{X}\tilde{\beta}$

```
x1 <- c(0,1,2,3,4)
x2 <- c(4,1,0,1,4)
Xtilde <- matrix(c(rep(1,5),x1,x2),ncol=3)
Xtilde
```

```
##      [,1] [,2] [,3]
## [1,]     1    0    4
## [2,]     1    1    1
## [3,]     1    2    0
## [4,]     1    3    1
## [5,]     1    4    4
```

```
t.beta <- c(10,5,-2)
t.y <- Xtilde%*%t.beta
t.y
```

```
##      [,1]
## [1,]     2
## [2,]    13
## [3,]    20
## [4,]    23
```

Next, we generate the vector containing the $\epsilon_i \sim N(0, \sigma^2)$ with $\sigma^2 = 1$:

```
t.e <- rnorm(5,0,1)
t.e
```

```
## [1] 0.7606833 -0.3257157 0.6830309 0.9070262 0.9342162
```

which we add to the “true” $y = \tilde{X}\tilde{\beta}$ values, to obtain the “observed” values:

```
t.Y <- t.y + t.e
t.Y
```

```
##          [,1]
## [1,] 2.760683
## [2,] 12.674284
## [3,] 20.683031
## [4,] 23.907026
## [5,] 22.934216
```

It is now possible to fit the model with lm:

```
r.lm <- lm(t.Y ~ x1 + x2)
summary(r.lm)$coef
```

	Estimate	Std. Error	t value	Pr(> t)
## (Intercept)	10.069826	0.5556231	18.12348	0.003030672
## x1	5.157981	0.1866953	27.62780	0.001307540
## x2	-1.896970	0.1577864	-12.02239	0.006847617

Alternatively, we can use formula

$$\hat{\beta} = (\tilde{X}^\top \tilde{X})^{-1} \tilde{X}^\top y$$

to find the parameter estimates:

```
solve(t(Xtilde) %*% Xtilde) %*% t(Xtilde) %*% t.Y
```

```
## [1,] 10.069826
## [2,] 5.157981
## [3,] -1.896970
```

- ▶ `solve()` calculates the **inverse** (here the inverse of $\tilde{X}^\top \tilde{X}$).
- ▶ `t()` gives the **transposed** (here of \tilde{X}^\top).

Task: Do this calculation by yourself and verify for each step that the dimensions of the matrices and the vector are indeed fitting, so that this expression is defined.

Appendix

Some R commands for matrix algebra

Reading vectors and matrices into R:

```
a <- c(1,2,3)  
a
```

```
## [1] 1 2 3
```

```
A <- matrix(c(1,2,3,4,5,6), byrow=T, nrow=2)  
B <- matrix(c(6,5,4,3,2,1), byrow=T, nrow=2)  
A
```

```
##      [,1] [,2] [,3]  
## [1,]     1     2     3  
## [2,]     4     5     6
```

```
B
```

```
##      [,1] [,2] [,3]  
## [1,]     6     5     4  
## [2,]     3     2     1
```

Adding and subtracting:

A + B

```
##      [,1] [,2] [,3]
## [1,]    7    7    7
## [2,]    7    7    7
```

A - B

```
##      [,1] [,2] [,3]
## [1,]   -5   -3   -1
## [2,]    1    3    5
```

However, be careful, R sometimes does unreasonable things:

A + a

```
##      [,1] [,2] [,3]
## [1,]    2    5    5
## [2,]    6    6    9
```

What happened here??

Matrix multiplication:

```
C <- A %*% t(B)
C
```

```
##      [,1] [,2]
## [1,]    28   10
## [2,]    73   28
A%*%a
```

```
##      [,1]
## [1,]    14
## [2,]    32
```

Matrix inversion (possible for square matrices only):

```
solve(C)
```

```
##      [,1]      [,2]
## [1,]  0.5185185 -0.1851852
## [2,] -1.3518519  0.5185185
C %*% solve(C)
```

```
##      [,1] [,2]
## [1,]    1    0
## [2,]    0    1
```

Why does `solve(A)` or `solve(B)` not work?