Genomic Analysis Combining Genotyped and Non-Genotyped Individuals

Why a Combined Analysis?

- To exploit all the available phenotypic data in GWAS and genomic prediction
 - Not just the records on genotyped individuals
 - Account for preselection of genotyped individuals
- To ensure that genomic predictions include all available information
- To avoid approximations required in multistep analyses (that lead to double-counting)

Multi-step Genomic Prediction Analysis

- Mixed model evaluation using all phenotypes and pedigree information to generate EBV and R²
- Deregression of EBV on genotyped individuals using EBV and R² of trios of every genotyped individual, its sire and its dam
- Weighted multiple regression analysis of deregressed EBV to estimate SNP effects
- Genomic prediction DGV of genotyped individuals
- Pedigree prediction of DGV for nongenotyped
- Selection Index blending of DGV & EBV for GE-EBV

Selection Index Blending Assumptions

$$varegin{bmatrix} \widehat{u} \ \widehat{u} \end{bmatrix} = egin{bmatrix} r_p^2 & r_p^2 r_m^2 \end{bmatrix} egin{bmatrix} r_p^2 \ r_p^2 r_m^2 & r_m^2 \end{bmatrix} egin{bmatrix} r_p^2 \ r_p^2 & r_m^2 \end{bmatrix} oldsymbol{\sigma}_g^2 \ r_p^2 & r_m^2 & 1 \end{bmatrix}$$

$$varigg[rac{u-\widehat{u}}{m-\widehat{m}} igg] = igg[rac{1-r_p^2}{(1-r_p^2)\left(1-r_m^2
ight)} rac{(1-r_p^2)\left(1-r_m^2
ight)}{1-r_m^2} igg]_{-1}$$

Kachman (unpublished)

Pedigree Prediction

$$\begin{bmatrix} y_n \\ y_g \end{bmatrix} = \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} u_n \\ u_g \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix}$$

with

$$var\begin{bmatrix} u_n \\ u_q \end{bmatrix} = \begin{bmatrix} A_{nn} & A_{ng} \\ A_{gn} & A_{gg} \end{bmatrix} \sigma_a^2$$

Where **A** is the numerator relationship matrix (from pedigree) with subscripts n=non-genotyped & g=genotyped

Nejati-Javaremi et al (1997)

Replace A with G

 $M = k \ columns \ of \ (0, 1, 2) \ marker \ covariates$

$$G = [MM' + (2 - M)(2 - M)']/k$$

Various other authors expanded this with various approaches to center the marker covariates to create a Genomic Relationship Matrix

Fitting G⁻¹ in the mixed model equations is known as GBLUP and gives the same estimates of genomic merit as MHG "BLUP"

Genotyped Animals

$$y_g = X_g b + Z_g u_g + e_g$$

Meuwissen, Hayes & Goddard (2001)

$$egin{aligned} with \ u_g &= M_g lpha = \sum_{j=1}^{j=\#loci} m_j lpha_j oldsymbol{\delta}_j \ lpha_j &= substitution \ effect \ oldsymbol{\delta}_j &= (0,1) \ indicator \ variable \end{aligned}$$

Bayesian Alphabet

$$\delta_j = 1, \ \sigma_{\alpha_j}^2 = (known) \sigma_{\alpha}^2 was "BLUP"$$

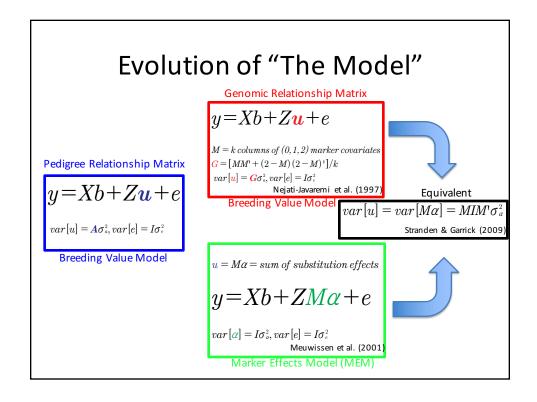
$$oldsymbol{\delta}_{\scriptscriptstyle j}=1,\;\; \sigma^{\scriptscriptstyle 2}_{\scriptscriptstyle lpha_{\scriptscriptstyle j}}=\left(\mathit{unknown}
ight)\sigma^{\scriptscriptstyle 2}_{\scriptscriptstyle lpha_{\scriptscriptstyle j}}\,\mathit{was}\,\mathit{BayesA}$$

$$\int eta_{\scriptscriptstyle j} = 0 \ with \ known \ probability = \pi \ \sigma_{\scriptscriptstyle lpha_{\scriptscriptstyle j}}^{\scriptscriptstyle 2} = (unknown) \, \sigma_{\scriptscriptstyle lpha_{\scriptscriptstyle j}}^{\scriptscriptstyle 2} \, was \, BayesB$$

Meuwissen, Hayes & Goddard (2001)

$$egin{aligned} eta_{\scriptscriptstyle j} &= 0 \ with \ (un) known \ probability = \pi \ &\sigma_{\scriptscriptstyle lpha_{\scriptscriptstyle j}}^{\scriptscriptstyle 2} &= (unknown) \ \sigma_{\scriptscriptstyle lpha}^{\scriptscriptstyle 2} \ was \ BayesC \ or \ (BayesC\pi) \end{aligned}$$

Kizilkaya et al (2010); Habier et al (2011)



What to do with the non-genotyped?

Known as Single-Step "First Attempt"

$$varegin{bmatrix} u_n \ u_g \end{bmatrix} = egin{bmatrix} A_{nn} & A_{ng} \ A_{gn} & G_{gg} \end{bmatrix} \sigma_a^2$$

Just replace that part of the numerator relationship matrix with genomic relationships

Then need a "brute-force" inversion of the var-cov matrix

Misztal et al (2009)

What to do with the non-genotyped?

Known as Single-Step "Second Attempt" (with brute force inverse)

$$H = var \begin{bmatrix} u_n \\ u_g \end{bmatrix} \sigma_a^{-2} = \begin{bmatrix} A_{nn} + A_{ng} A_{gg}^{-1} G_{gg} A_{gg}^{-1} A_{gn} & A_{ng} A_{gg}^{-1} G_{gg} \\ G_{gg} A_{gg}^{-1} A_{gn} & G_{gg} \end{bmatrix}$$
Legarra et al (2009)

Then with recognition of its simply structured inverse

$$H^{-1} = A^{-1} + \begin{bmatrix} 0 & 0 \ 0 & G_{gg}^{-1} - A_{gg}^{-1} \end{bmatrix}$$

Aguilar et al (2010)

Offering programming appeal by simply replacing A⁻¹ in MME by H⁻¹ known as Single-Step GBLUP and variants of which are widely used

What's wrong with Single-Step GBLUP?

• Its predictive ability can be improved by introducing another ad hoc constant κ whose optimal value can be found by trial and error

$$H^{-1} = A^{-1} + egin{bmatrix} 0 & 0 \ 0 & arkappa(G_{qq}^{-1} - A_{qq}^{-1}) \end{bmatrix}$$

What's wrong with Single-Step GBLUP?

- When there are less loci than genotyped individuals, G is singular
- When there are more loci than genotyped individuals, G is singular if locus covariates are centered by allele frequency

(since G=MM' and M'1=0 then G1=0)

 These problems can be overcome by adhoc regression of G towards A

What's wrong with Single-Step GBLUP?

- The var-cov matrix involves a blending of A and G requiring that they represent the same "base"
 - The base in A is the pedigree founders but the allele frequencies are not usually known in that population
- It is not clear what to use to center locus covariates in populations of mixed breeds, or populations with variable breed percentages

Issues with single-step GBLUP

- The matrix G is often singular
 - More animals than markers
 - If G is centered with observed allele frequency
- The matrix G must be "on the same base" as A

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Rather than using G_{gg}^{-1} - A_{gg}^{-1}
The model is tuned using \tau [a + b((1-c)G_{gg} + cA_{gg})]^{-1} - \omega A_{gg}^{-1} with some trial and error and \tau \leq 1; \omega \leq 1; a \leq 0.1; b \leq 1; 0.05 \leq c \leq 0.2
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· Computing effort increases with numbers genotyped

What's wrong with Single-Step GBLUP?

- It requires brute force inversion of 2 matrices whose order is the number of genotyped individuals (ie G and A_{gg})
 - The inversion effort increase rapidly with number of genotyped individuals
 - Inversion is impractical beyond say 100,000 individuals
- Ignacy now has an "APY" approximation approach for computing these inverses

What's wrong with Single-Step GBLUP?

- It is not computationally straightforward for extension to Single-Step BayesA
- It is not suitable for application of mixture models (BayesB, BayesC, BayesCπ)
 - But these models that provide variable selection are particularly appealing in fine-mapping applications such as with imputed NGS genotypes

Let's revisit the basic idea

$$\begin{bmatrix} y_n \\ y_g \end{bmatrix} = \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} u_n \\ u_g \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix}$$

$$with \ u_g = M_g \alpha \ for \ genotyped \ individuals$$

$$whereas \ u_n = \widehat{u_n} / u_g + (u_n - \widehat{u_n} / u_g) = \widehat{u_n} / u_g + \varepsilon_n$$

$$with \ \widehat{u_n} / u_g = A_{ng} A_{gg}^{-1} u_g$$

$$so \ u_n = A_{ng} A_{gg}^{-1} u_g + (u_n - A_{ng} A_{gg}^{-1} u_g)$$

Substituting these results gives

$$\begin{aligned} \begin{bmatrix} y_n \\ y_g \end{bmatrix} &= \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} u_n \\ u_g \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix} \\ &= \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix} \begin{bmatrix} A_{ng} A_{gg}^{-1} M_g \alpha \\ M_g \alpha \end{bmatrix} + \begin{bmatrix} Z_n & 0 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} \boldsymbol{\varepsilon}_n \\ 0 \end{bmatrix} + \begin{bmatrix} e_n \\ e_g \end{bmatrix} \\ &= \begin{bmatrix} X_n \\ X_g \end{bmatrix} b + \begin{bmatrix} Z_n A_{ng} A_{gg}^{-1} M_g \\ Z_g M_g \end{bmatrix} \alpha + \begin{bmatrix} Z_n \\ 0 \end{bmatrix} \boldsymbol{\varepsilon}_n + \begin{bmatrix} e_n \\ e_g \end{bmatrix}$$

Fernando et al (2014) GSE

With "Hybrid" Mixed Model Equations

$$\begin{bmatrix} X'X & X'ZM & X_n'Z_n \\ M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\ Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn}\lambda \end{bmatrix} \begin{bmatrix} b \\ \alpha \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} X'y \\ M'Z'y \\ Z_n'y_n \end{bmatrix}$$

$$where \ X = \begin{bmatrix} X_n \\ X_g \end{bmatrix}, Z = \begin{bmatrix} Z_n & 0 \\ 0 & Z_g \end{bmatrix}, M = \begin{bmatrix} M_n \\ M_g \end{bmatrix} = \begin{bmatrix} A_{ng}A_{gg}^{-1}M_g \\ M_g \end{bmatrix}, y = \begin{bmatrix} y_n \\ y_g \end{bmatrix}$$

with EBV given by $\widehat{u}_{\alpha} = M_{\alpha} \widehat{\alpha}$

 $\widehat{u_n} = M_n \widehat{\alpha} + \widehat{\varepsilon_n}$

NB Single-Step GBLUP is a special case of the above (but in this equivalent model no inversion is needed)

$$M_n = A_{ng} A_{gg}^{-1} M_g$$

If everyone is genotyped

$$\begin{bmatrix} X'X & X'ZM & X_n Z_n \\ M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\ Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn} \lambda \end{bmatrix} \begin{bmatrix} b \\ \alpha \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} X'y \\ M'Z'y \\ Z_n'y_n \end{bmatrix}$$

These are the MME that form the basis of BayesA, BayesB, BayesCetc

If no one is genotyped

$$\begin{bmatrix} X'X & X'ZM & X_n'Z_n \\ M'Z'X & M'Z'ZM + \phi & M_n'Z_n'Z_n \\ Z_n'X_n & Z_n'Z_nM_n & Z_n'Z_n + A^{nn}\lambda \end{bmatrix} \begin{bmatrix} b \\ \alpha \\ \varepsilon_n \end{bmatrix} = \begin{bmatrix} X'y \\ M'Z'y \\ Z_n'y_n \end{bmatrix}$$

These MME form the basis of traditional pedigree-based BLUP

Invariant to Covariate Centering

Genotyped

$$egin{aligned} y_g &= \mathbf{1} \mu + X_g b + Z_g M_g lpha + e_g \ &= \mathbf{1} \mu + X_g b + Z_g \mathbf{1} c^{\dagger} lpha + Z_g (M_g - 1c^{\dagger}) lpha + e_g \ define \ t &= c^{\dagger} lpha \ y_g &= \mathbf{1} (\mu + t) + X_g b + Z_g (M_g - 1c^{\dagger}) lpha + e_g \ &= \mathbf{1} \mu^* + X_g b + Z_g M_g^c lpha + e_g \end{aligned}$$

.....when all animals genotyped (BayesA, BayesB etc)

But non-genotyped NOT invariant

Non-genotyped

$$y_n = \mathbf{1}\mu + X_n b + Z_n A_{ng} A_{gg}^{-1} M_g \alpha + Z_n \varepsilon_n + e_n$$

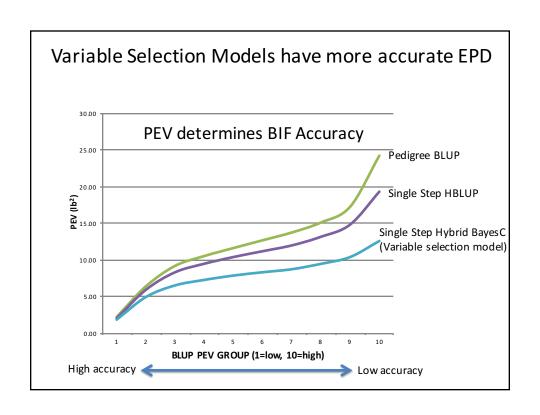
$$= \mathbf{1}\mu + X_n b + Z_n A_{ng} A_{gg}^{-1} \mathbf{1}c' \alpha + Z_n A_{ng} A_{gg}^{-1} (M_g - 1c') \alpha + Z_n \varepsilon_n + e_n$$

$$= \mathbf{1}\mu + X_n b + Z_n A_{ng} A_{gg}^{-1} \mathbf{1}t + Z_n A_{ng} A_{gg}^{-1} M_g^c \alpha + Z_n \varepsilon_n + e_n$$

So combined analysis of genotyped and non-genotype animals need to include a covariate for t if there is arbitrary centering (unless t = 0)

Computational Aspects

- It is easy to compute $A_{ng}A_{gg}^{-1}M_g$ — And this can be done in parallel
- The computing becomes easier (rather than more difficult or impossible) as more individuals are genotyped
- Readily caters for variable selection or mixture models (eg BayesB, BayesC)
- This formulation is readily extended to multibreed, maternal effects and multi-trait settings
- In an MCMC framework can provide PEV



Summary

- Genomic prediction is an immature technology
- More effort is required to extend algorithms and to develop parallel computing procedures to most efficiently implement the full range of multi-breed, multi-trait, maternal effects and other models that have been routinely applied to large-scale animal prediction in recent decades

Prediction of BVs

$$with\ EBV\ given\ by$$
 $\widehat{u_g}=M_g\,\widehat{lpha}$
 $\widehat{u_n}=M_n\,\widehat{lpha}+\widehat{oldsymbol{arepsilon}_n}$
 $or, with\ M_n=A_{ng}A_{gg}^{-1}M_g\,\widehat{lpha}+\widehat{oldsymbol{arepsilon}_n}$
 $=A_{ng}A_{gg}^{-1}M_g\,\widehat{lpha}+\widehat{oldsymbol{arepsilon}_n}$
 $=A_{ng}A_{gg}^{-1}\widehat{u_g}+\widehat{oldsymbol{arepsilon}_n}$