

# Characterising a posterior distribution

Nick Davies

Model Fitting and Inference for  
Infectious Disease Dynamics  
short course, June 2019

# Recap

Last session we saw that the **posterior distribution** of  $\theta$ , given observed data  $D$ , is

$$p(\theta|D) = \frac{p(D|\theta)p(\theta)}{p(D)}$$

$$\text{Posterior} = \frac{\text{Likelihood} \times \text{Prior}}{\text{Constant}}$$

Our aim is to characterize the **posterior**.

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$$\text{Posterior} \propto \text{Likelihood} \times \text{Prior}$$

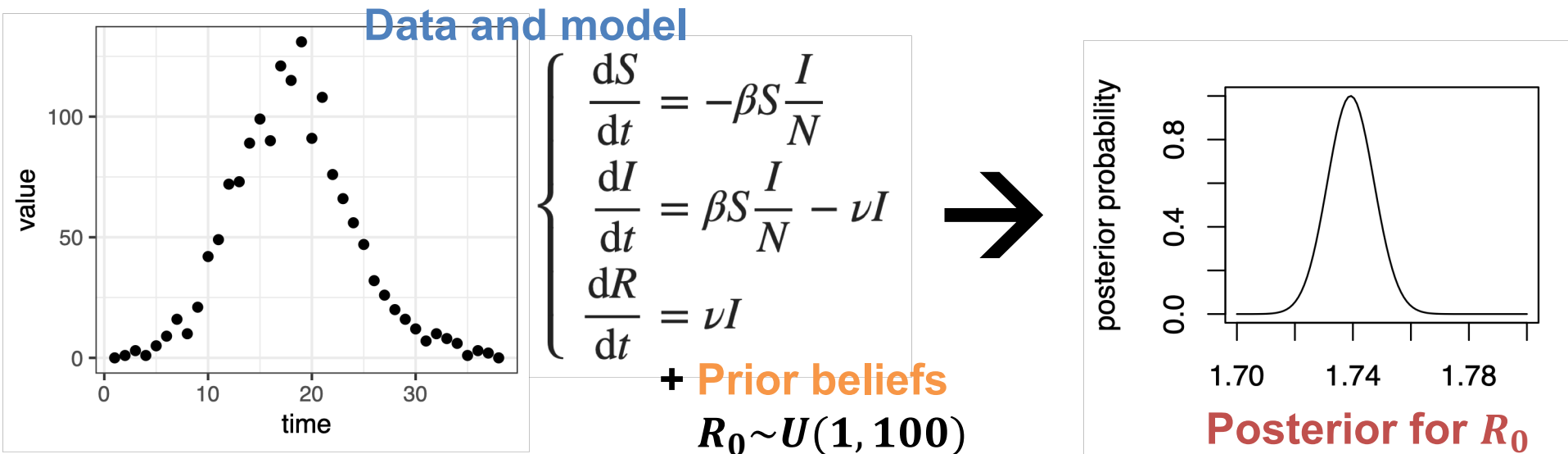
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# Recap

$$p(\theta|D) \propto p(D|\theta)p(\theta) \quad \text{Posterior} \propto \text{Likelihood} \times \text{Prior}$$

The **posterior** is a probability distribution that tells us what parameter values are credible given the data we have observed.

This allows us to answer questions like: given some case data and a model, plus some (potentially vague) prior beliefs, which values of  $R_0$  are plausible?

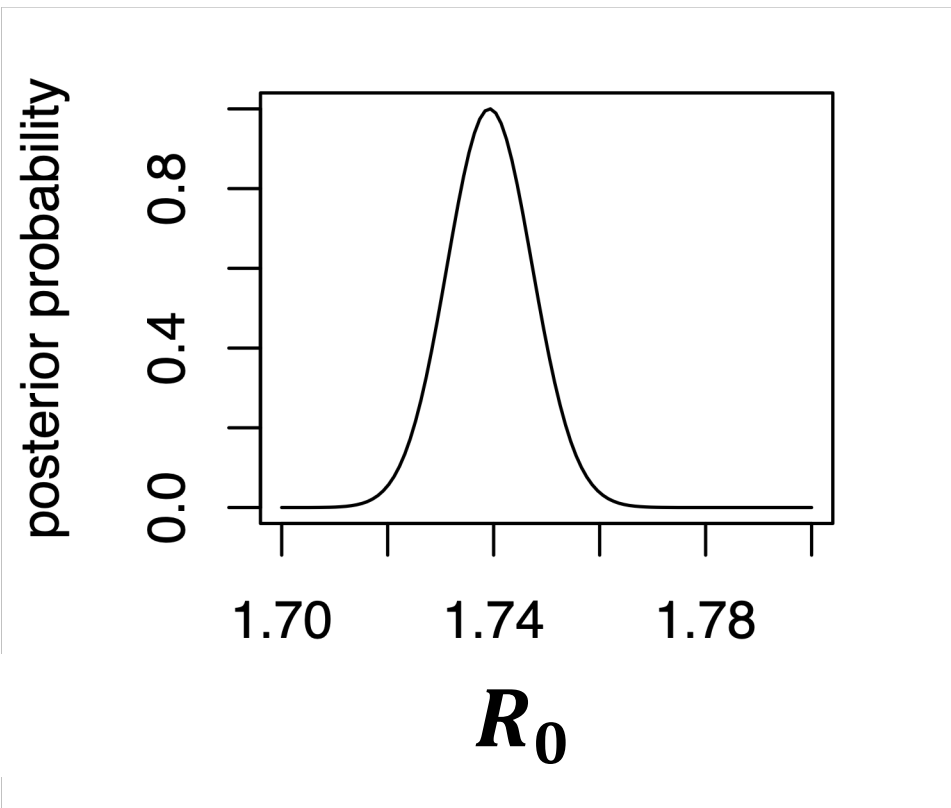


# The problem

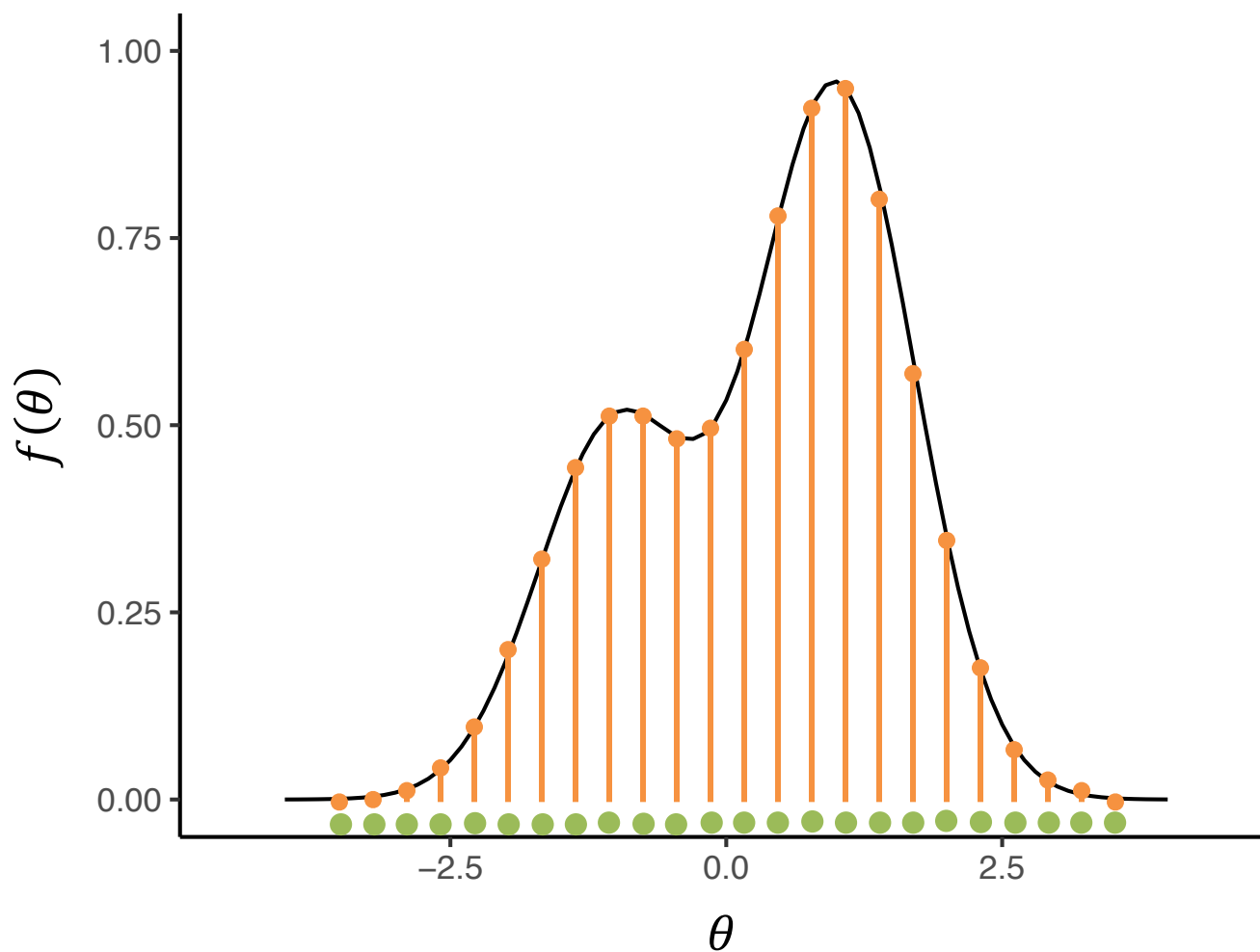
The **posterior**,  $p(\theta|D) \propto p(D|\theta)p(\theta)$  is not, generally, easy to “solve” for, generally because the **likelihood** is complicated.

So how do we characterize the posterior, *i.e.*:

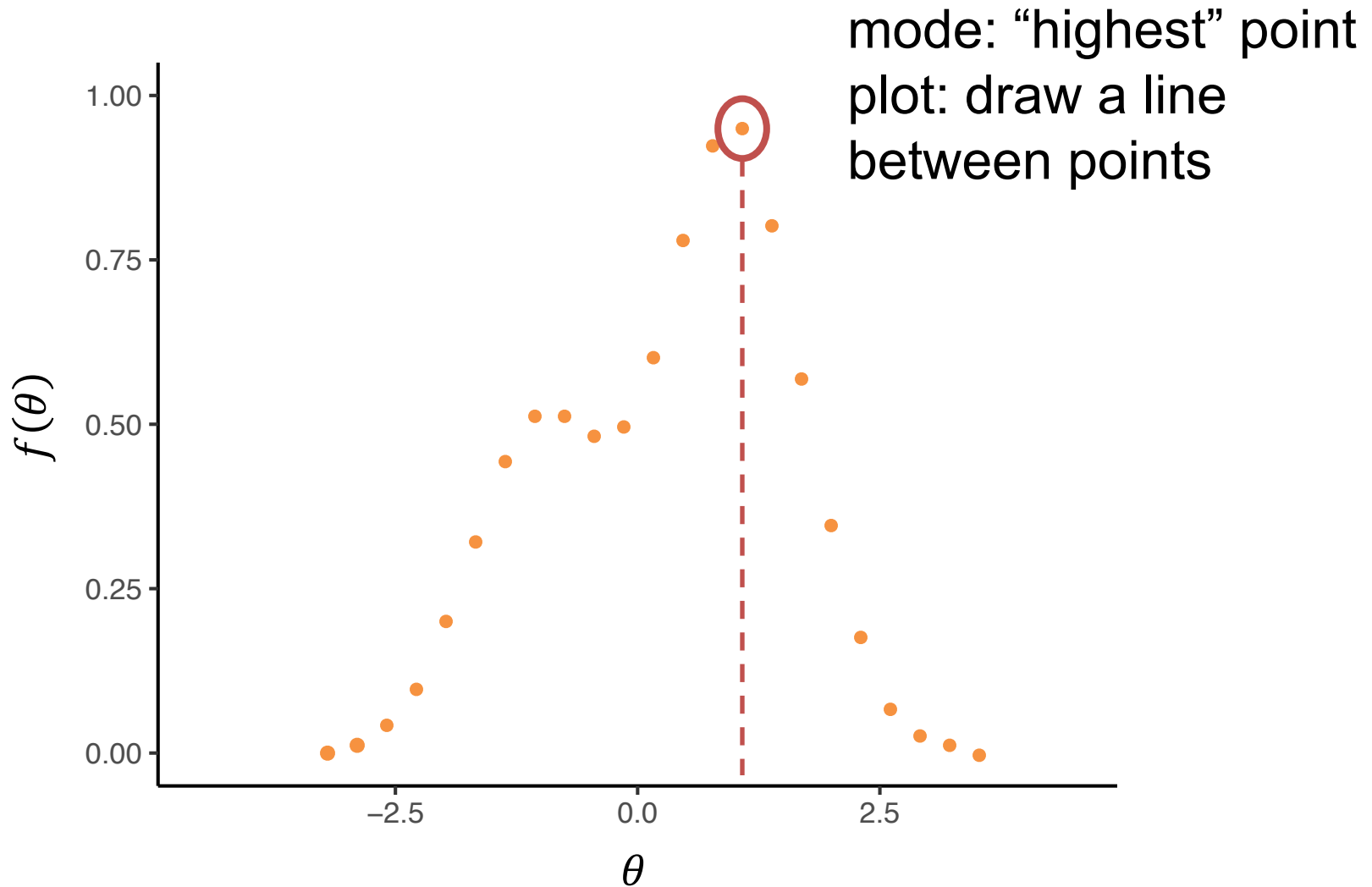
- find the mean, median, mode of  $R_0$ ?
- visualize  $R_0$  in plots?
- give “credible intervals” for  $R_0$ ?
- use fitted  $R_0$  to make predictions?



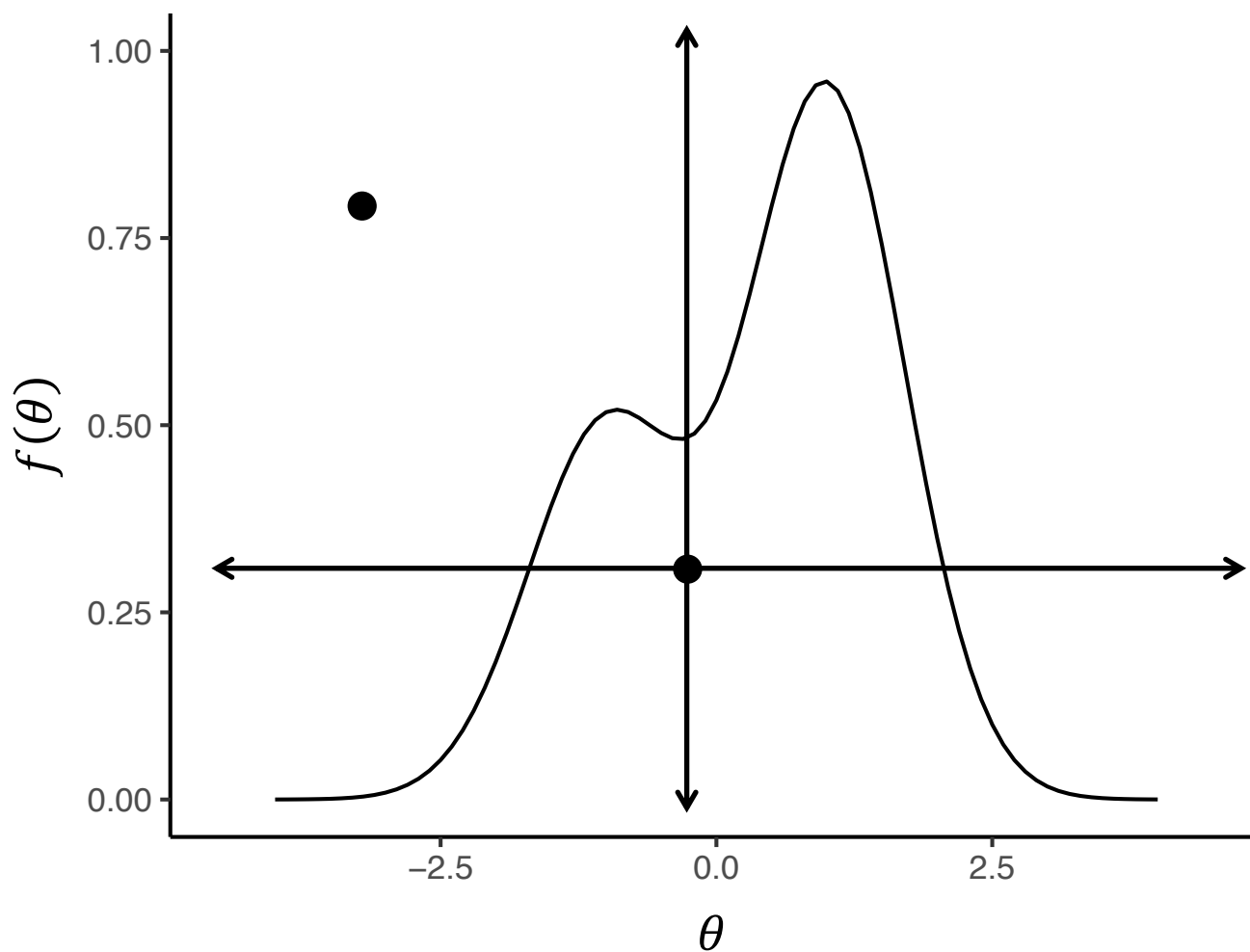
# Method 1: Grid approximation



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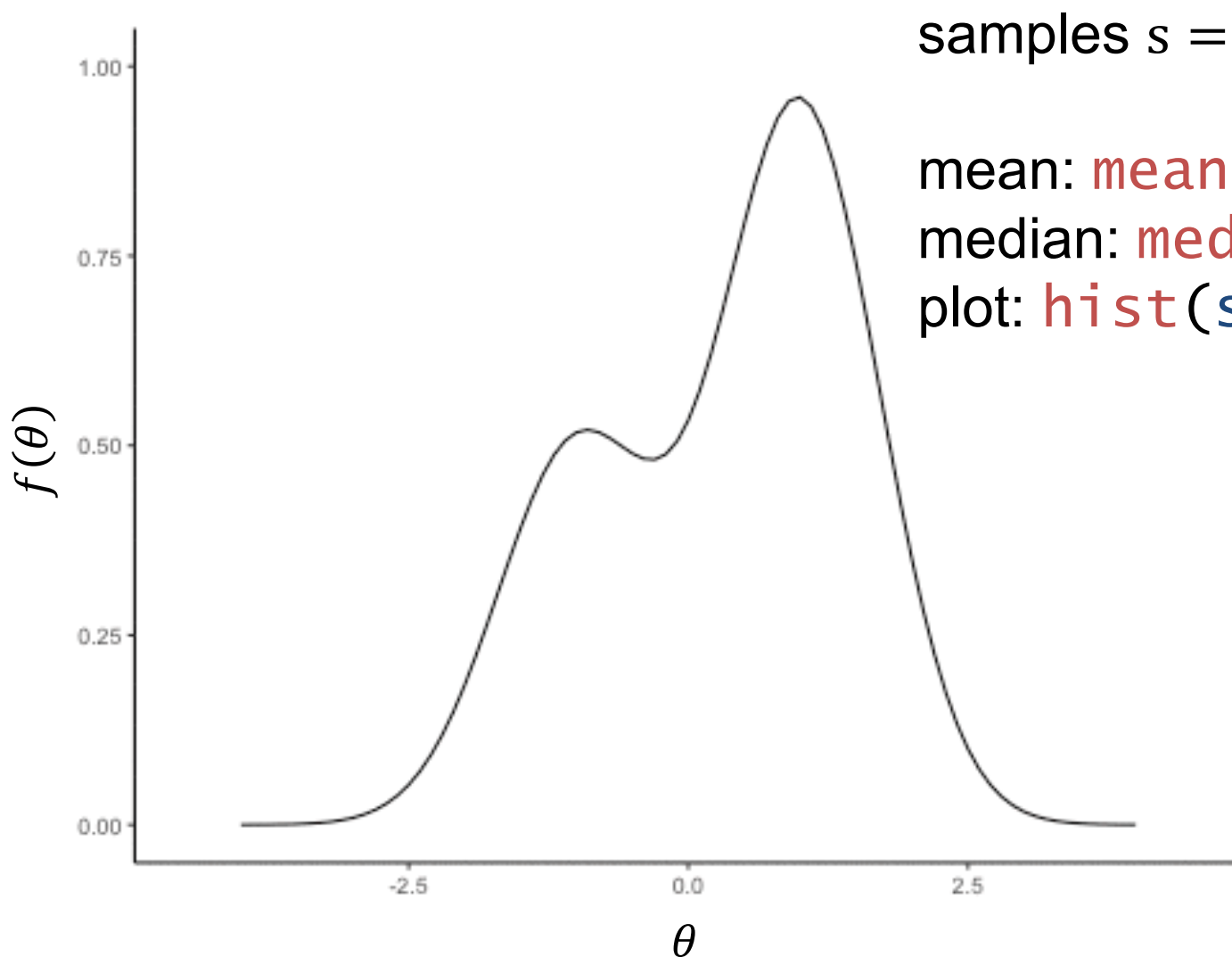


# Method 2: Rejection sampling





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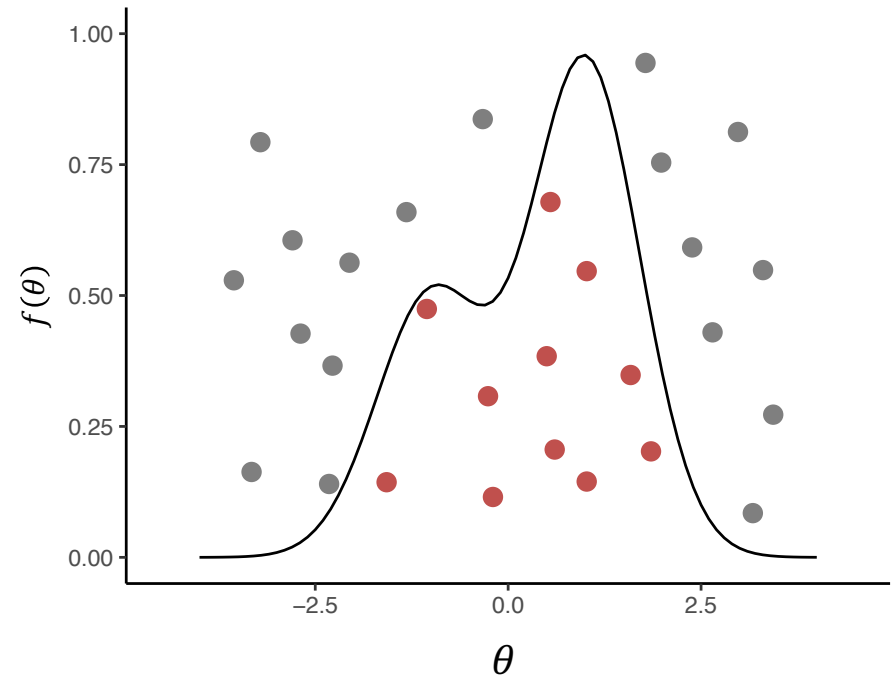
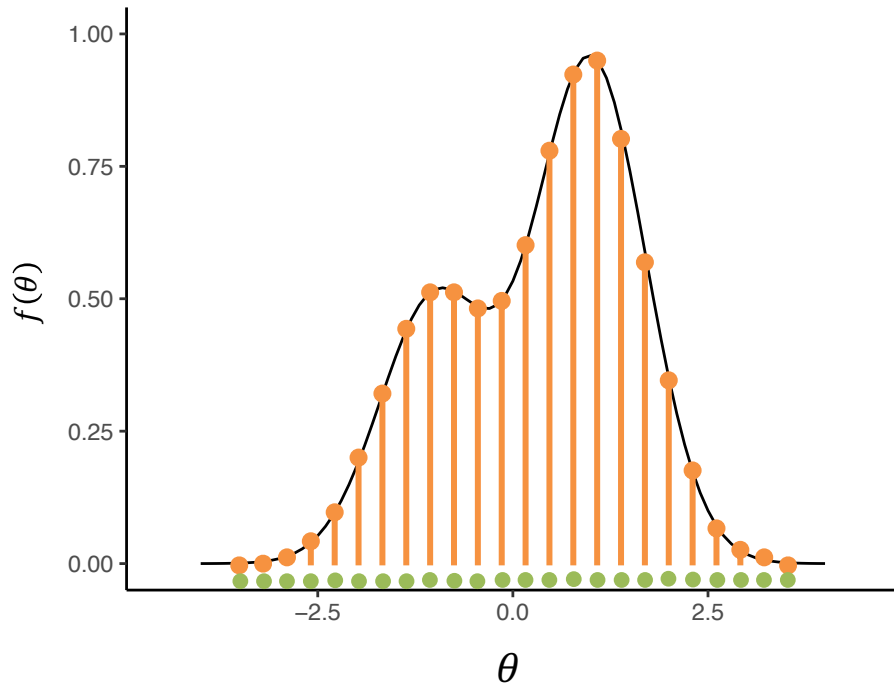
samples  $s = \{\theta_1, \theta_2, \dots\}$

mean: `mean(s)`

median: `median(s)`

plot: `hist(s)`

# Practical, part 1

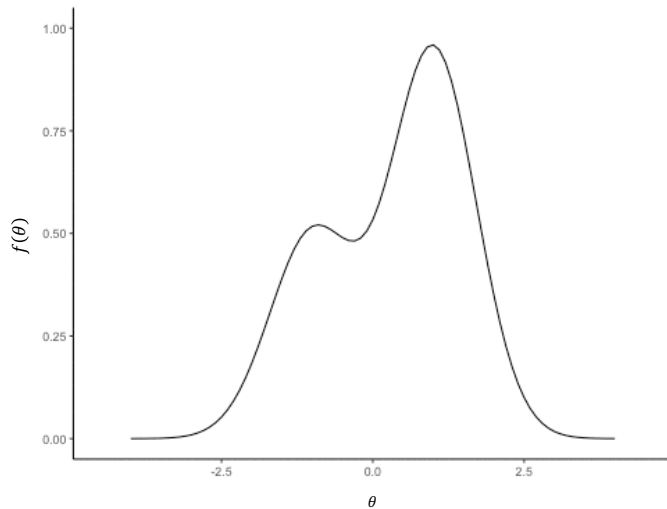


Start the practical (30–45 mins):  
“Grid approximation” and “Rejection sampling”  
in the MCMC session

# Practical 1

1. How much do the summary statistics change if you perform the sampling again?
2. If you decrease the number of attempts from 1000 to 100, would you expect the summary statistics to change more each time sampling is performed or less? What does this tell you about reliable sampling?
3. What are the advantages and disadvantages of grid approximation versus rejection sampling?

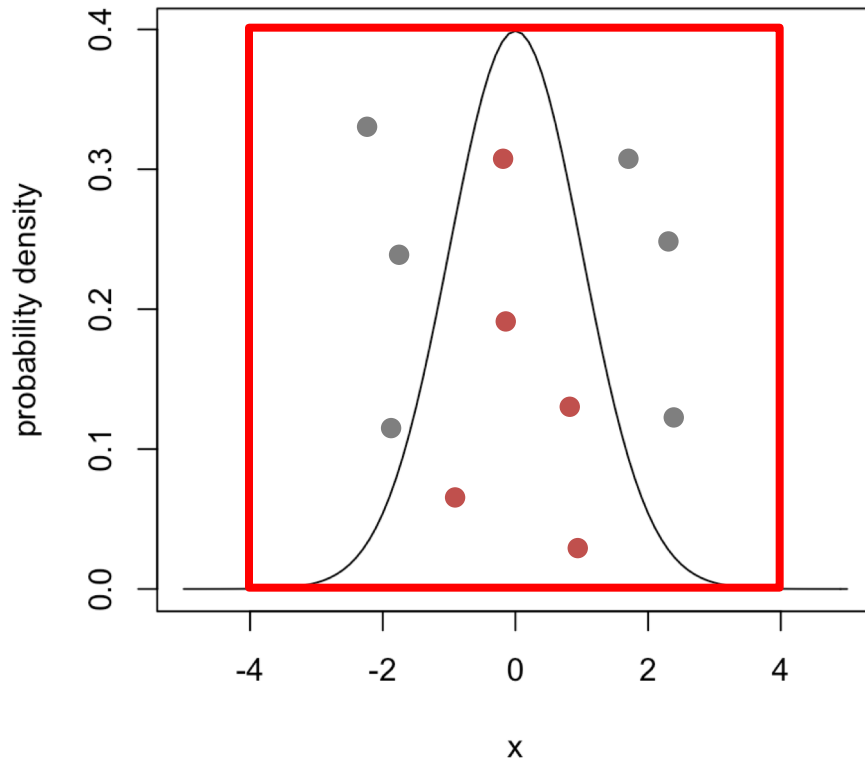
# Issues with grid / rejection methods



Need to specify the limits of the distribution

Curse of dimensionality

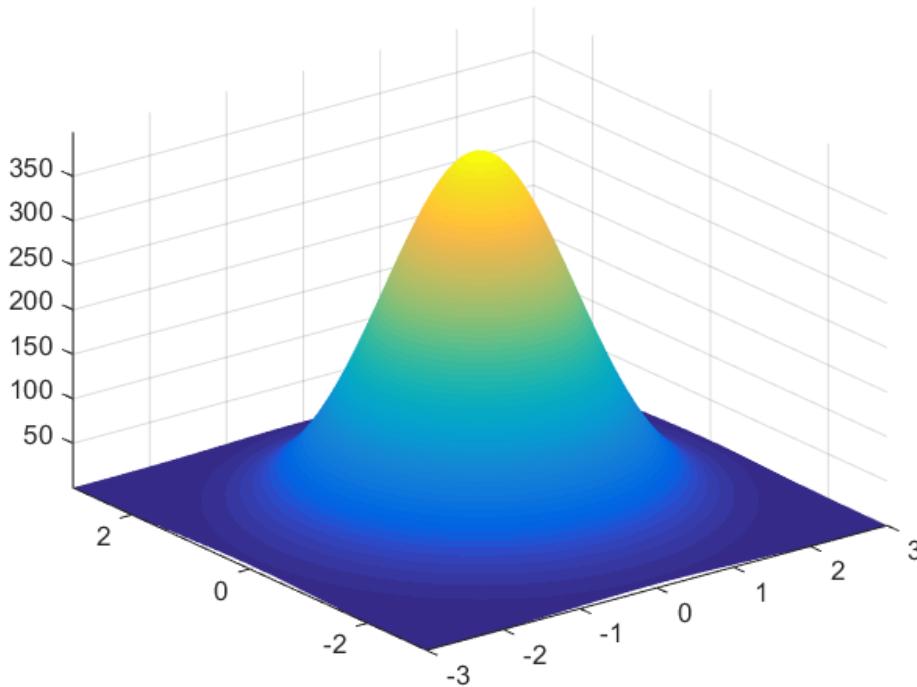
# Curse of dimensionality



Throwing darts at the standard normal distribution between  $-4 \cdot \text{SD}$  and  $4 \cdot \text{SD}$ :

Fraction 0.31 will “hit”

# Curse of dimensionality



2D normal distribution:  
 $0.31^2 = 0.098$

10D normal distribution:  
 $0.31^{10} = 0.0000091$

(to get 1000 samples,  
you need to throw 110  
million darts... and that's  
the best case scenario)

# What can we do with posterior samples?

We've been focusing on ways of *reporting* parameter estimates...

samples  $s = \{\theta_1, \theta_2, \dots\}$

`mean(s)` *# if s is vector of R0, this gives mean R0*

But we can use them to make predictions, test interventions, etc

```
for theta in s {  
  run model with R0 = theta and 0% vaccination  
  run model with R0 = theta and 50% vaccination  
  record results  
}
```

# Markov Chain Monte Carlo

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# Markov Chain Monte Carlo (MCMC)

Markov chain: stochastic sequence of states in which the next state depends only upon the current state

$$\theta_{t+1} \sim \mathbb{D}(\theta_t)$$

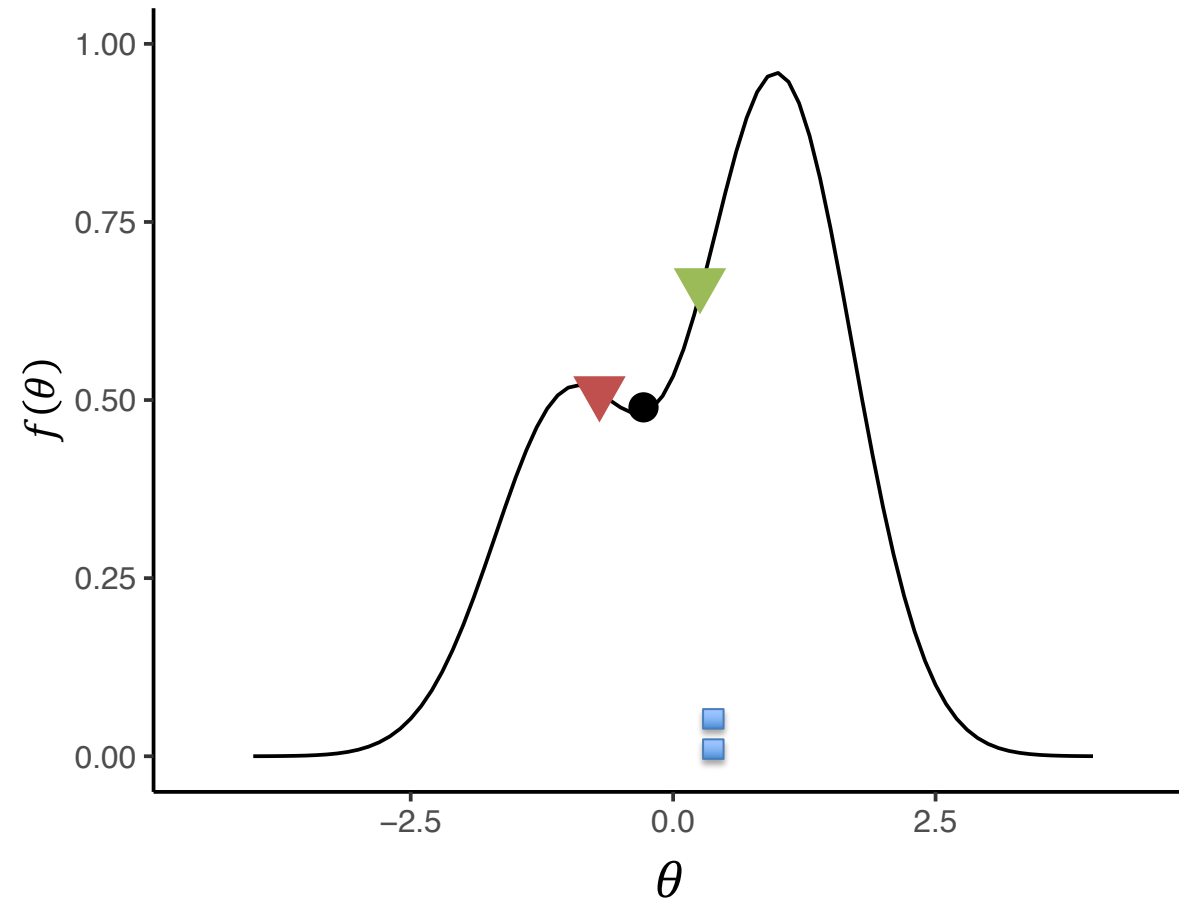
Monte Carlo: a famous casino. Also a class of algorithms in which random sampling is used to solve problems.

Metropolis-Hastings algorithm: a particular way of using MCMC to sample from a distribution

# MCMC: Outline

- What the algorithm is
- Practical: Implementing MCMC
- Why it works.

# MCMC algorithm



Choose a starting point,  $\theta = \theta_0$

*PROPOSE*

$$\theta' = \theta + \varepsilon \mid \varepsilon \sim Q$$

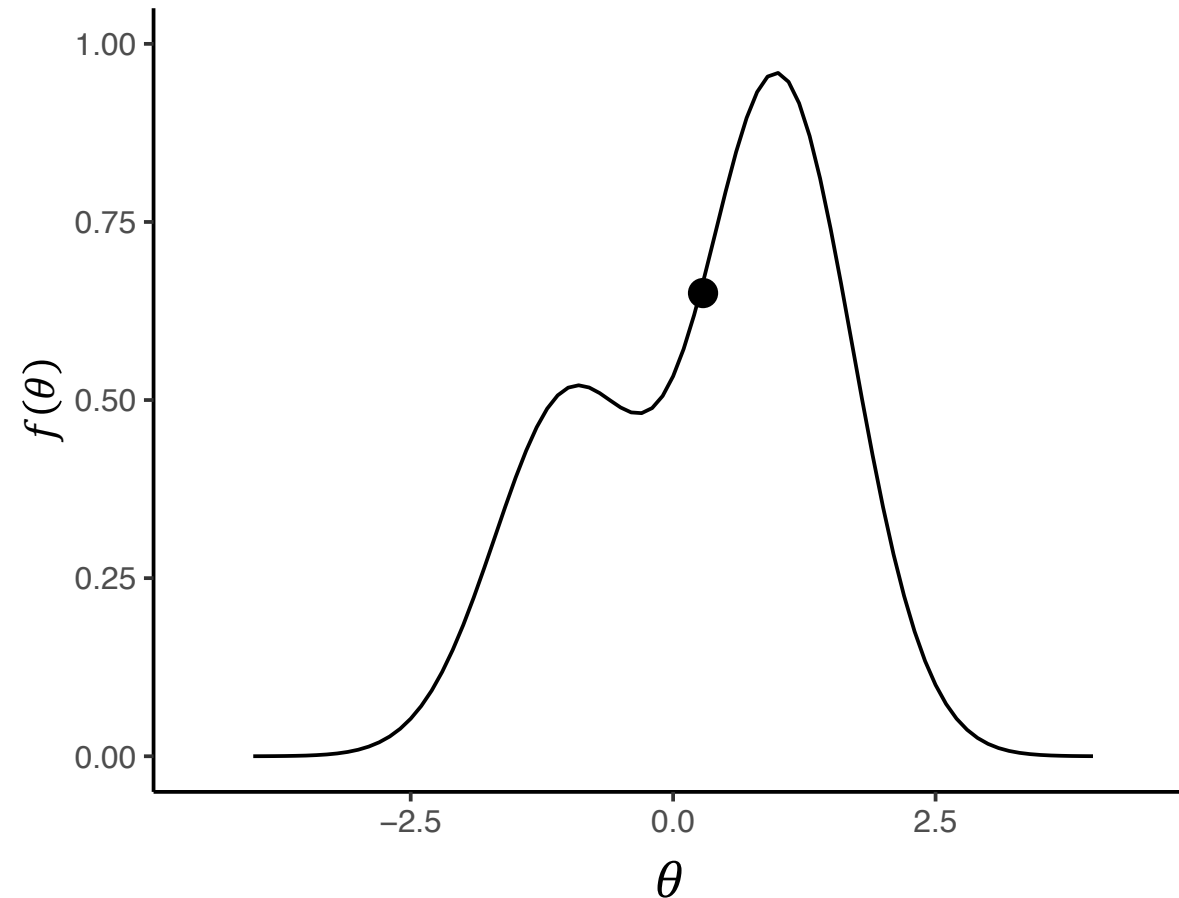
*MOVE OR STAY*

If  $f(\theta') > f(\theta)$ , definitely move.

If  $f(\theta') < f(\theta)$ , move with probability  $f(\theta')/f(\theta)$ , otherwise stay.

1. PROPOSE
2. MOVE OR STAY (“acceptance”)
3. SAVE LOCATION

# MCMC algorithm



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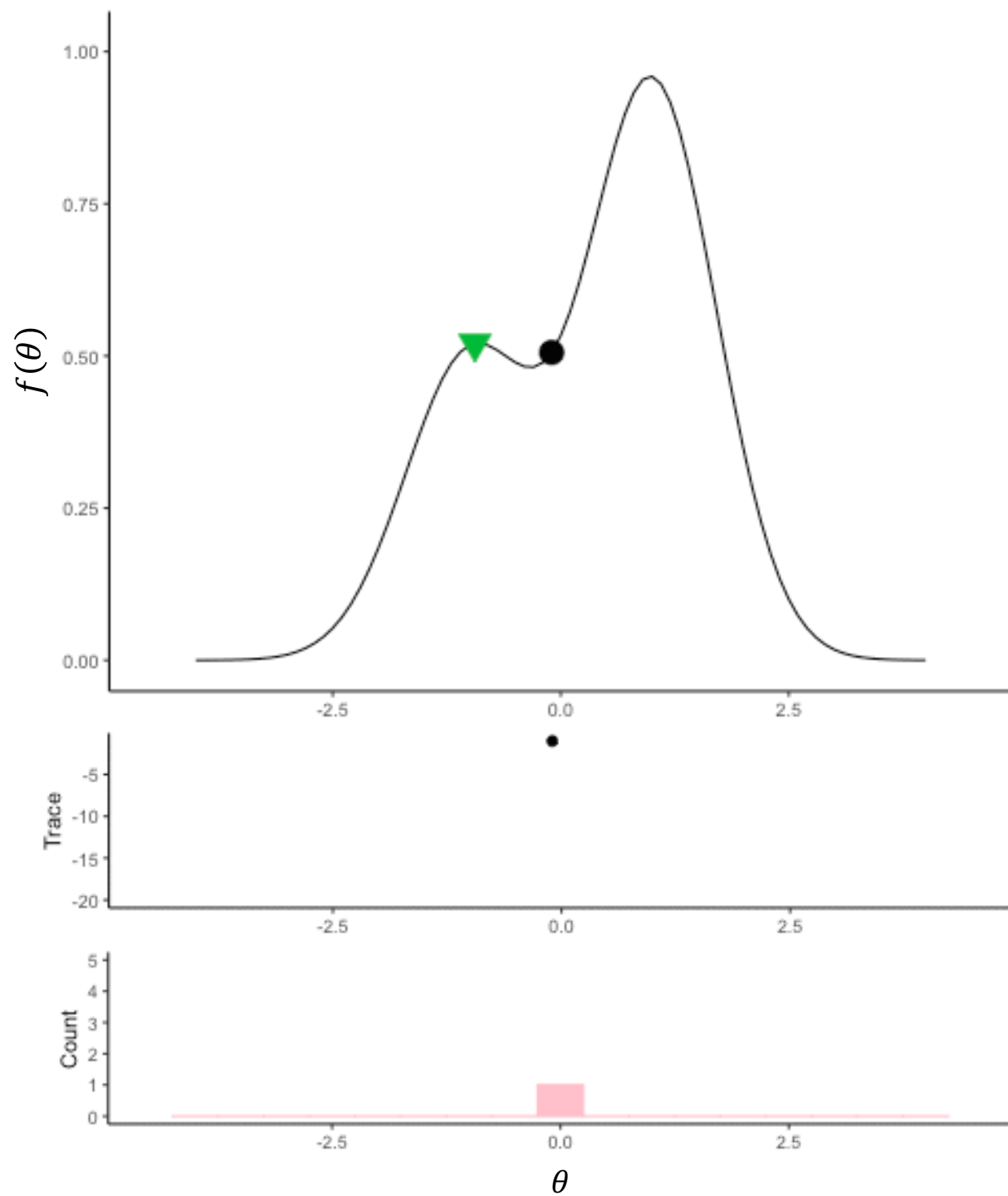
*MOVE OR STAY*

$$\theta \rightarrow \begin{cases} \theta' & \text{Pr}(a) & \text{Accept} \\ \theta & \text{Pr}(1 - a) & \text{Reject} \end{cases}$$

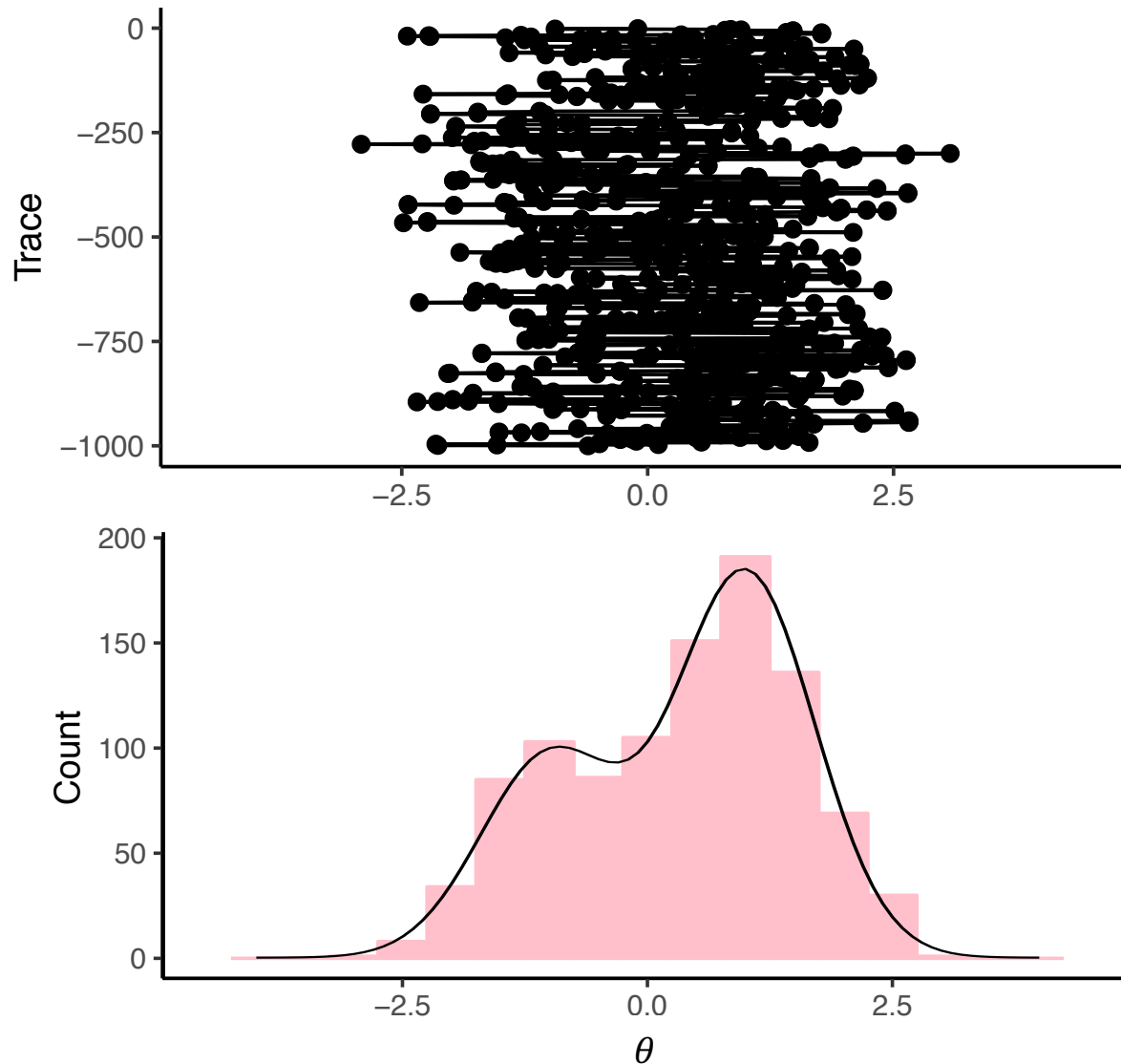
$$a = \min\left(1, \frac{f(\theta')}{f(\theta)}\right)$$

*SAVE LOCATION*

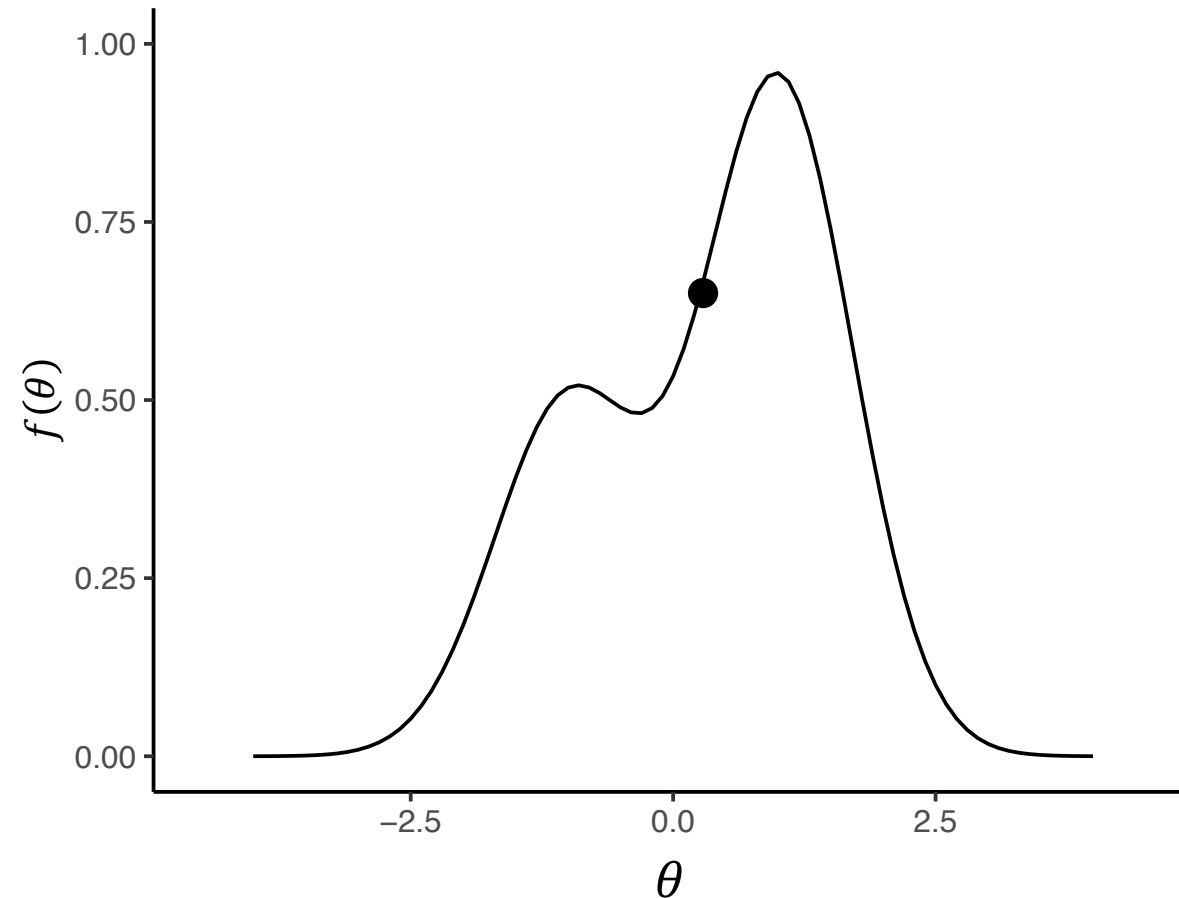
$$\theta_t \leftarrow \theta$$



# After enough iterations...



# Practical, part 2



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*SAVE LOCATION*

$$\theta_t \leftarrow \theta$$

# Why does MCMC with Metropolis-Hastings converge to the target distribution?

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THE JOURNAL OF CHEMICAL PHYSICS

VOLUME 21, NUMBER 6

JUNE, 1953

## Equation of State Calculations by Fast Computing Machines

NICHOLAS METROPOLIS, ARIANNA W. ROSENBLUTH, MARSHALL N. ROSENBLUTH, AND AUGUSTA H. TELLER,  
*Los Alamos Scientific Laboratory, Los Alamos, New Mexico*

AND

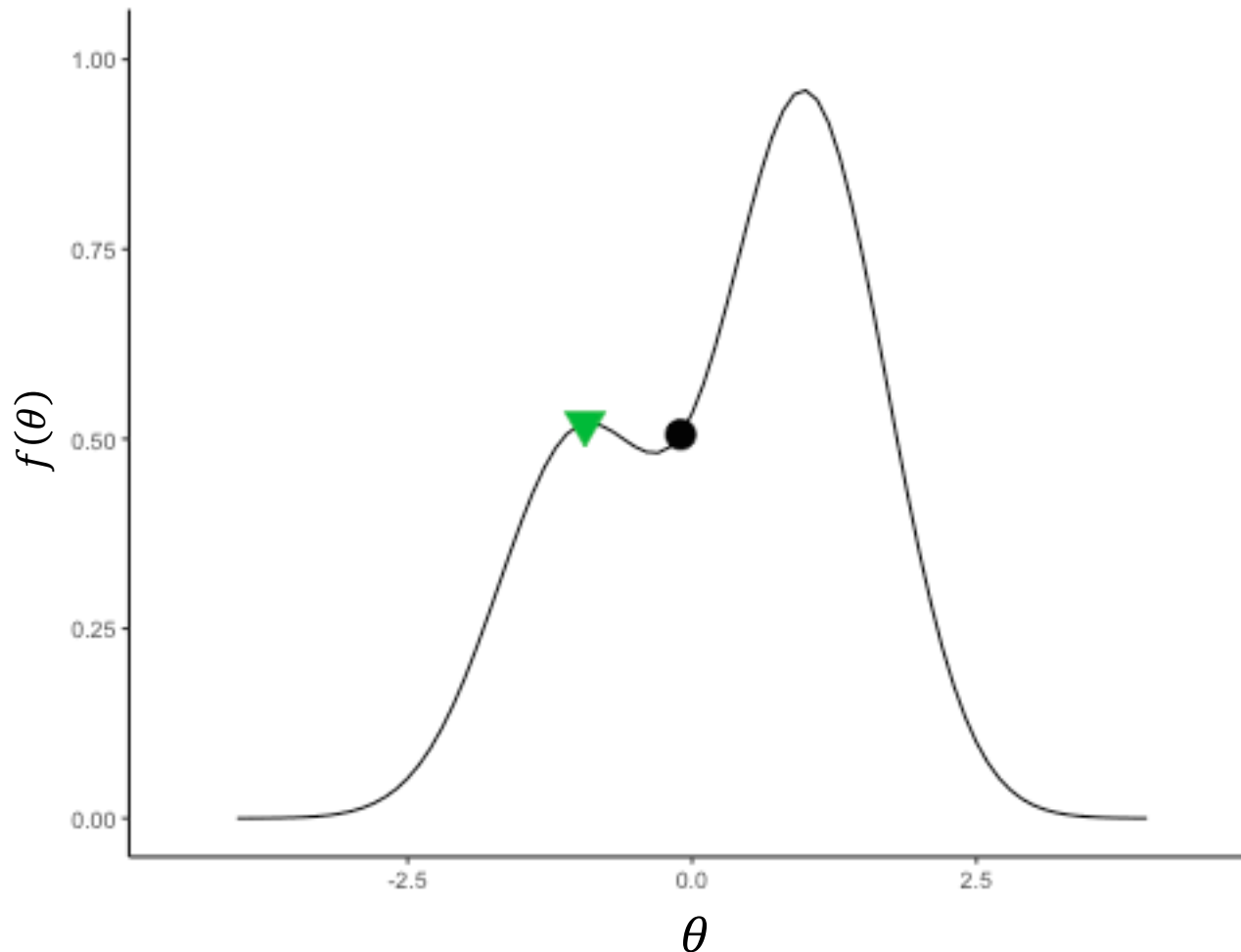
EDWARD TELLER,\* *Department of Physics, University of Chicago, Chicago, Illinois*

(Received March 6, 1953)

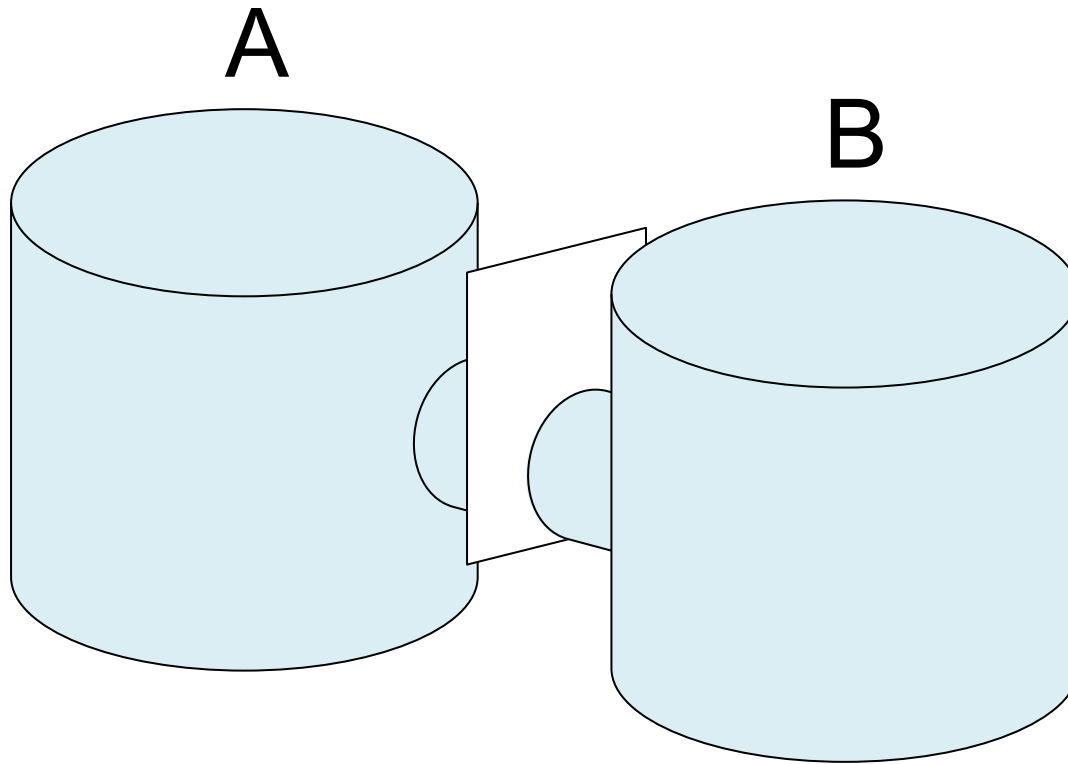
A general method, suitable for fast computing machines, for investigating such properties as equations of state for substances consisting of interacting individual molecules is described. The method consists of a modified Monte Carlo integration over configuration space. Results for the two-dimensional rigid-sphere system have been obtained on the Los Alamos MANIAC and are presented here. These results are compared to the free volume equation of state and to a four-term virial coefficient expansion.



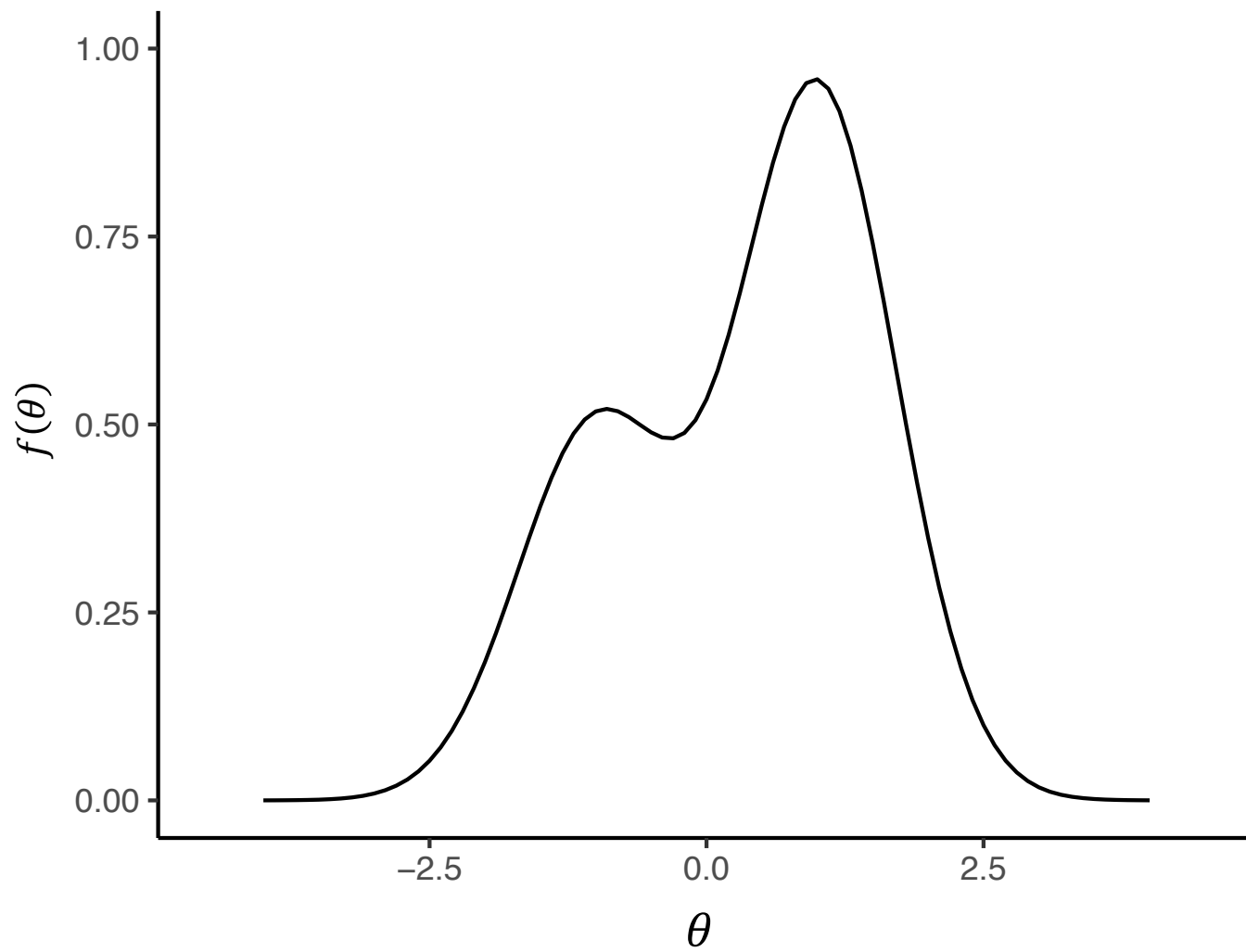
# A sloppy hill-climbing algorithm

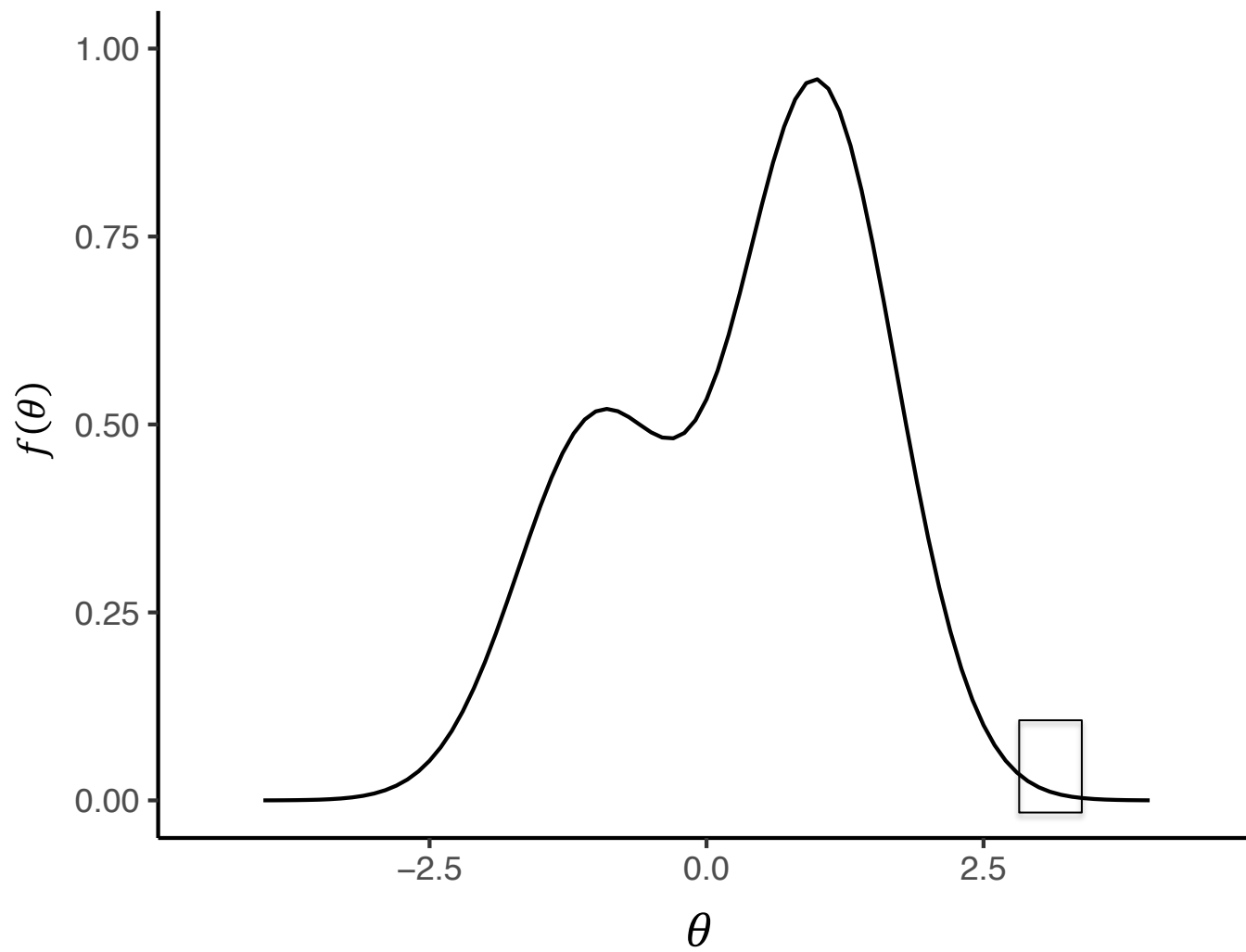


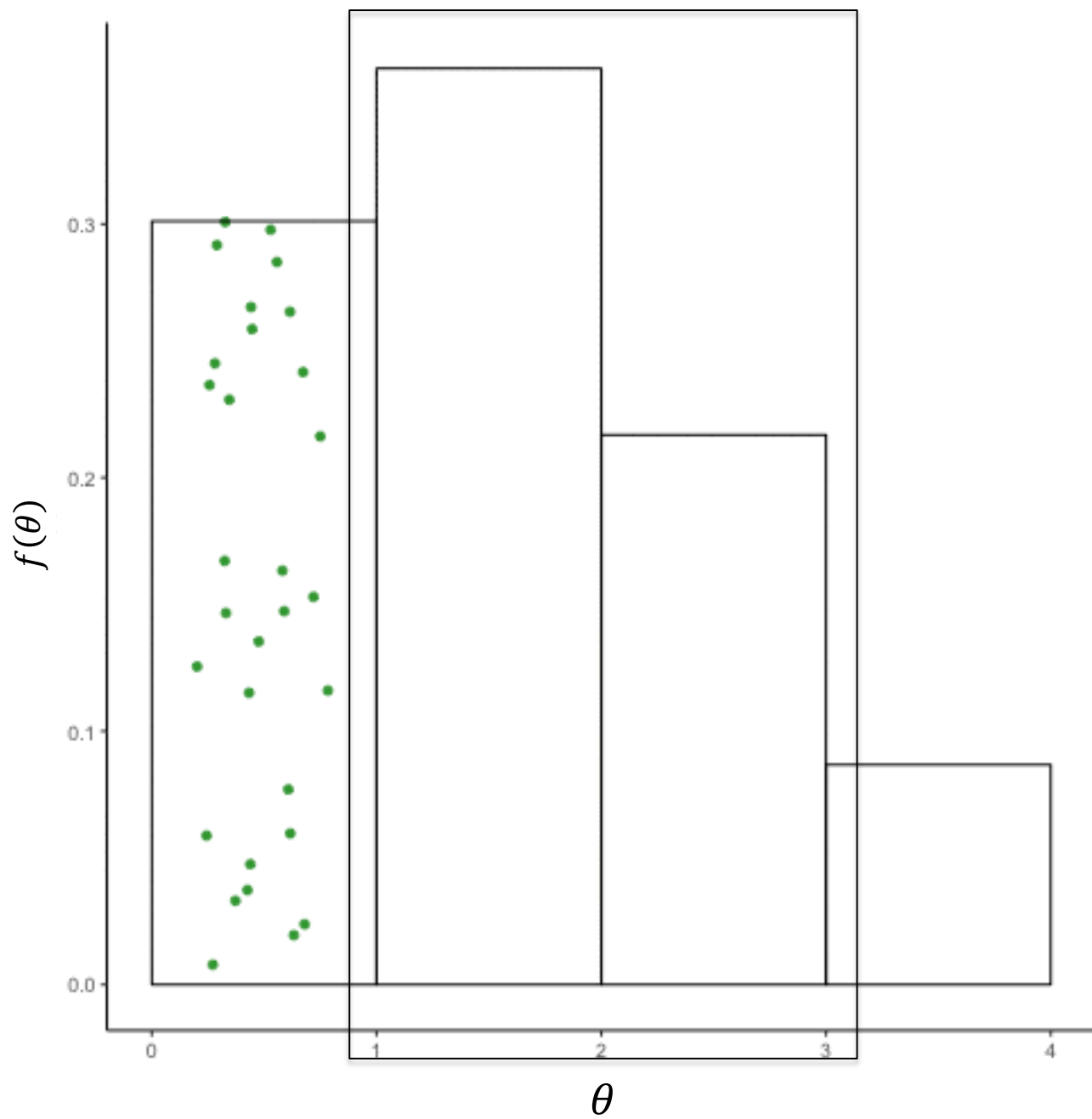
# A thought experiment...

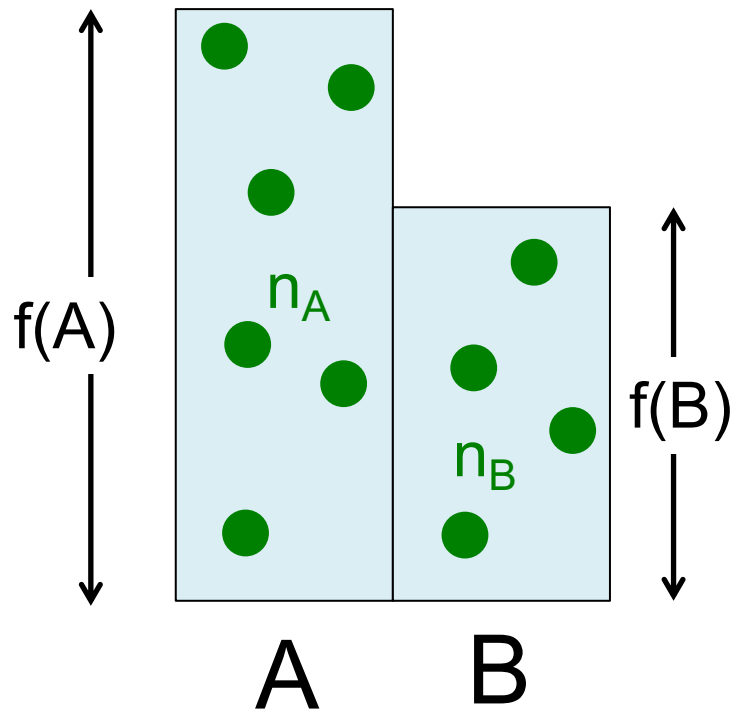


How does the dye “know” to stop flowing?









assume  $f(A) > f(B)$

$$\Pr(A \rightarrow B) = q_{A \rightarrow B} \cdot \frac{f(B)}{f(A)}$$

$$\Pr(B \rightarrow A) = q_{B \rightarrow A} \cdot 1$$

$$q_{A \rightarrow B} = q_{B \rightarrow A} = q$$

$$\Pr(A \rightarrow B) = q \frac{f(B)}{f(A)}$$

$$\Pr(B \rightarrow A) = q$$

Net movement from  $A \rightarrow B$

$$n_A \Pr(A \rightarrow B) - n_B \Pr(B \rightarrow A)$$

there will be flow  $A \rightarrow B$  if:

$$n_A \frac{f(B)}{f(A)} - n_B > 0$$

$$\frac{n_A}{n_B} > \frac{f(A)}{f(B)}$$

there will be flow  $B \rightarrow A$  if:

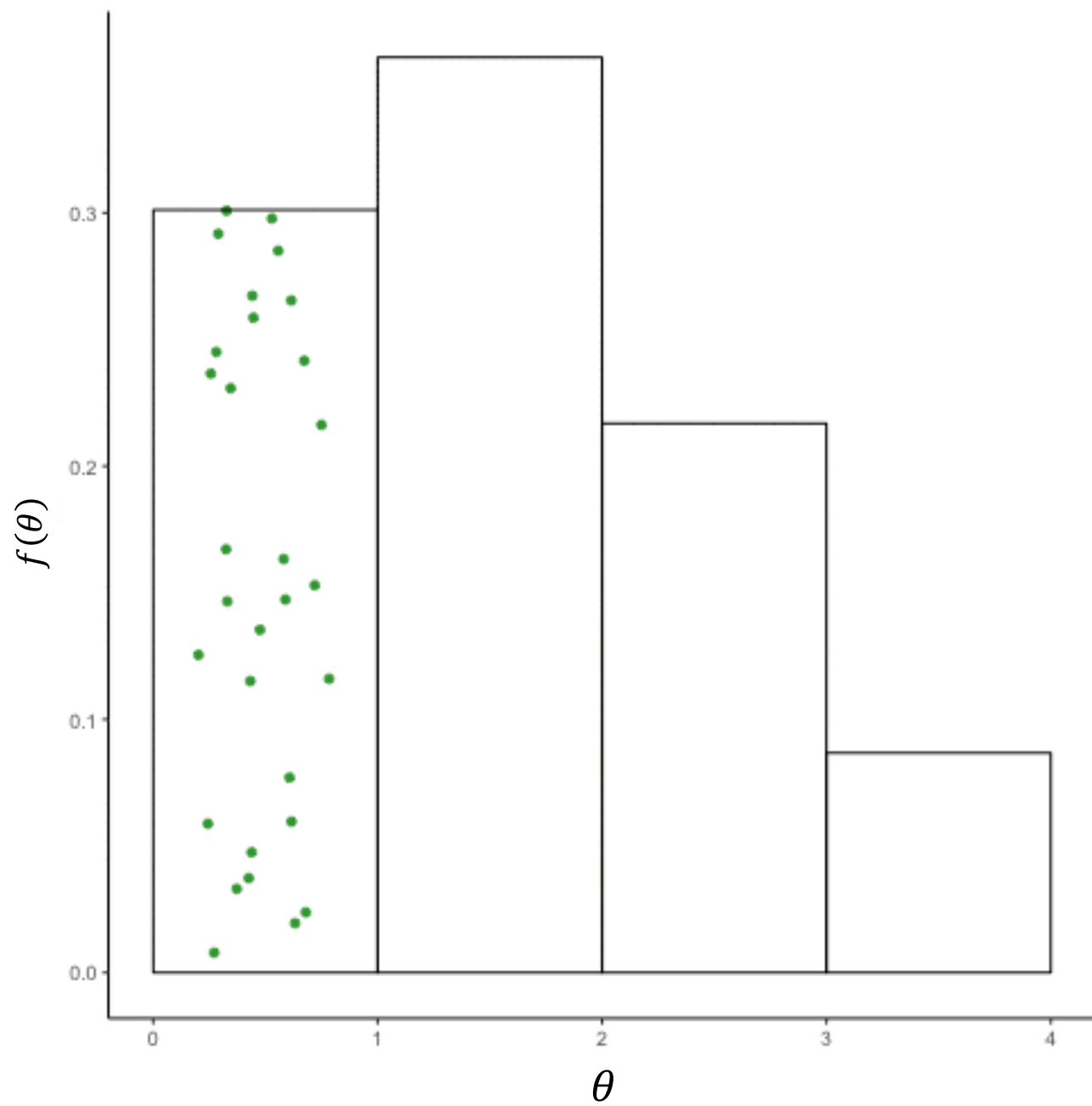
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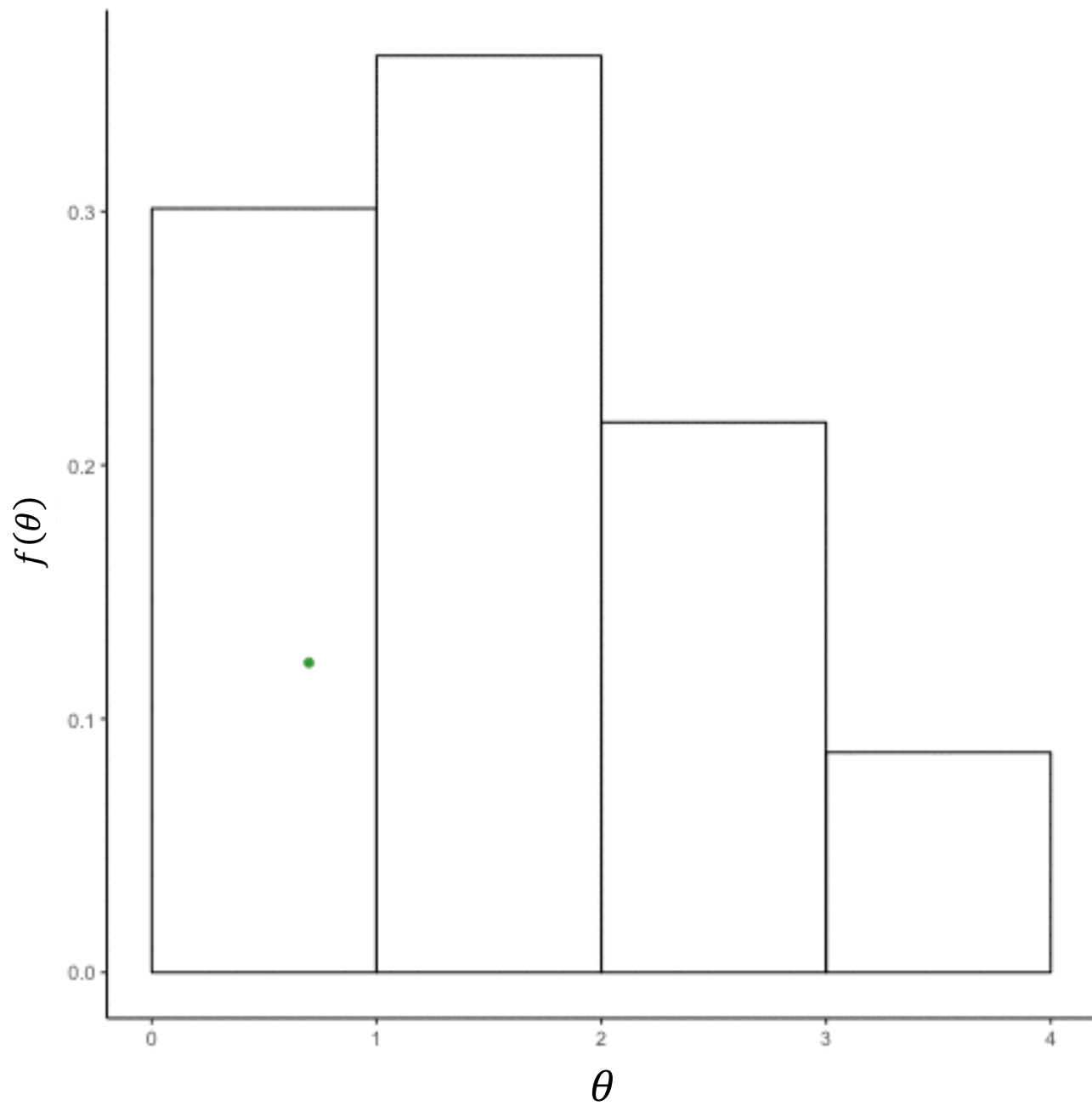
$$\frac{n_A}{n_B} < \frac{f(A)}{f(B)}$$

there will be no net flow  $A \nleftrightarrow B$  if:

$$n_A \frac{f(B)}{f(A)} - n_B = 0$$

$$\frac{n_A}{n_B} = \frac{f(A)}{f(B)}$$







# Requirements

Symmetry of proposal distribution

“detailed balance”  $q_{A \rightarrow B} = q_{B \rightarrow A}$

if  $q_{A \rightarrow B} \neq q_{B \rightarrow A}$ , use acceptance ratio

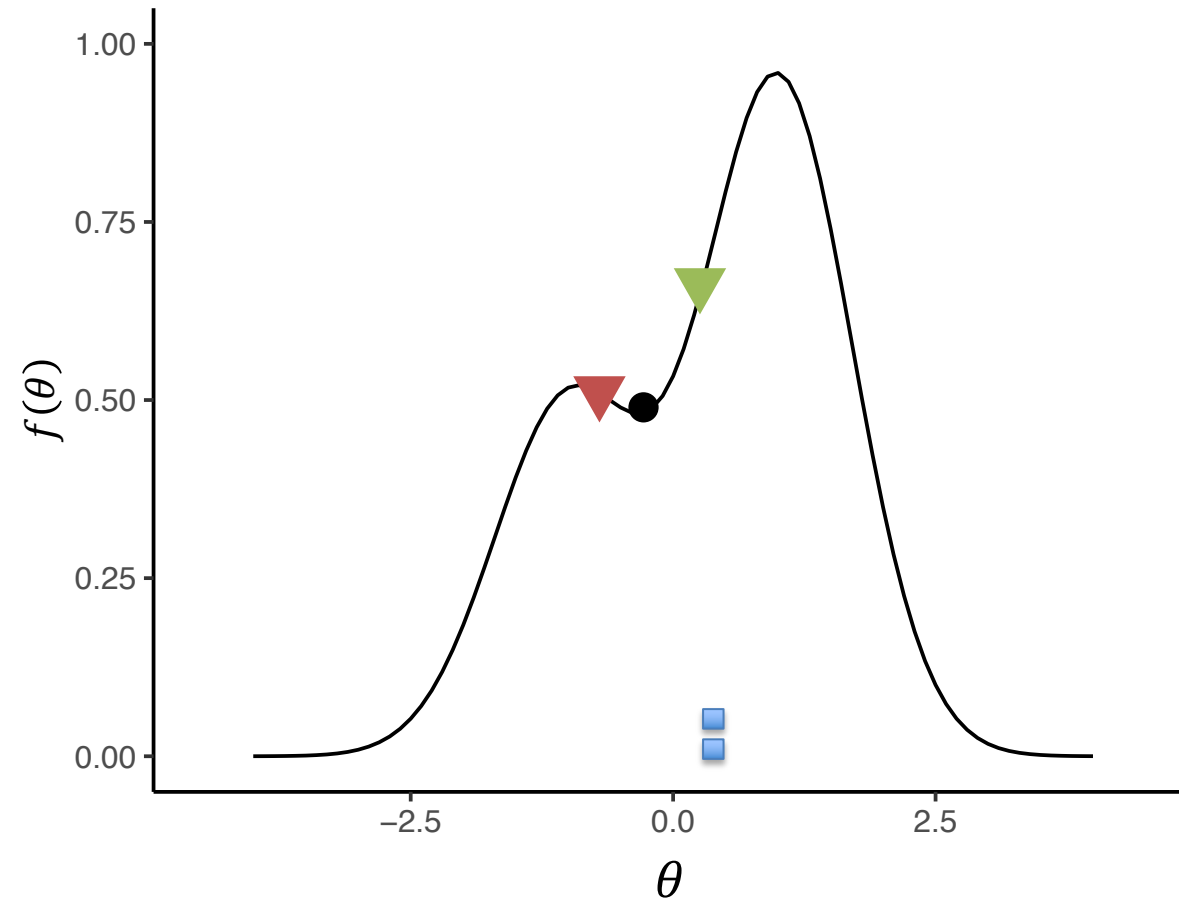
$$A = \min \left( 1, \frac{f(\theta')}{f(\theta)} \frac{q_{\theta \rightarrow \theta'}}{q_{\theta' \rightarrow \theta}} \right)$$

(Hastings 1970)

Connectedness of distribution

ensures “ergodicity”

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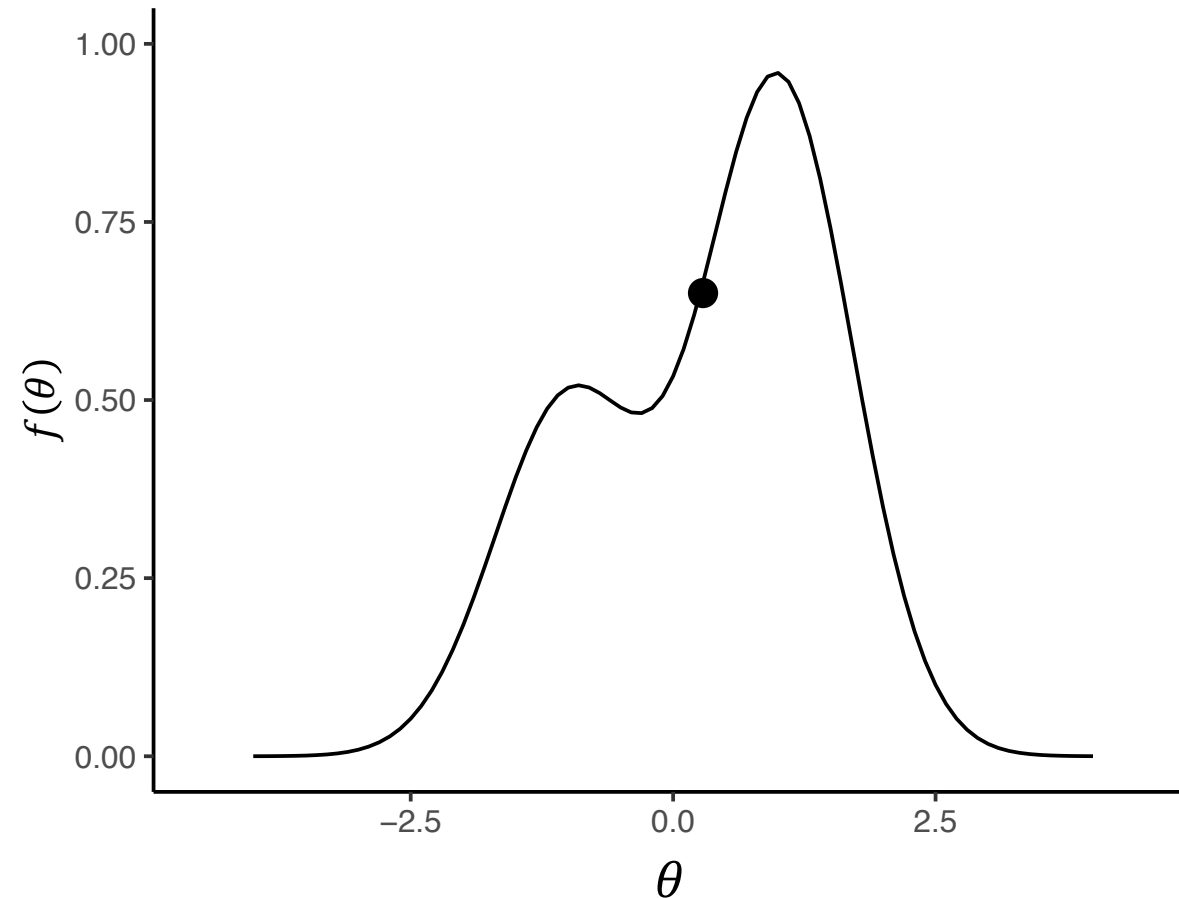
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