

Stat511(Section001): Homework #1

Due on Jan. 12, 2021 at 11:59pm

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Problem 1

Consider $X \sim \text{Uni}(-1, 1)$, $Y = X^2$. It is obvious that they are dependent. However, we can have:

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = E(X^3) - E(X)E(X^2) = 0 - 0 = 0$$

That is to say, X and Y are dependent but uncorrelated. [2]

Problem 2

We should use Scheme2. Because, for Scheme1, we obtain w_A, w_B respectively, for each w , we have standard deviation is σ . And we consider Scheme 2, we can get:

$$w_A = \frac{(w_{A+B} + w_{A-B})}{2}, w_B = \frac{(w_{A+B} - w_{A-B})}{2}$$

Since two measurement should be independent, we can get:

$$\text{Var}(w_A) = \frac{\text{Var}(w_{A+B}) + \text{Var}(w_{A-B})}{4} = \text{Var}(w_B) = \frac{\sigma^2}{2}$$

$$\text{Sd}(w_A) = \text{Sd}(w_B) = \sqrt{\text{Var}(w_A)} = \frac{\sqrt{2}}{2}\sigma$$

which is smaller than the standard deviation in Scheme1. Hence, we think Scheme2 is more precise and we prefer Scheme2.

Problem 3

Since (X, Y) is a joint normal distribution, we can know that $X - Y$ is also a normal distribution, with expectation equal to 0 and variance equal to $\text{Var}(X) + \text{Var}(Y) - 2\text{Cov}(X, Y)$. And we have:

$$\text{Cov}(X, Y) = \rho \text{Sd}(X) \text{Sd}(Y) = \rho$$

Hence, we can know $X - Y \sim N(0, 2 - 2\rho)$. Let $Z = X - Y$, Hence:

$$P(X_i > Y_i) = P(X_i - Y_i > 0) = \frac{1}{\sqrt{2\pi(2 - 2\rho)}} \int_{z=0}^{\infty} e^{-\frac{z^2}{2(2-2\rho)}} dz = \frac{1}{2}$$

Now we can see that for one i , $I(X_i > Y_i)$ is a Bernoulli distribution with $p = \frac{1}{2}$. T is sum of n Bernoulli distributio. Hence T is a binomial distribution, and $T \sim \text{Bin}(n, \frac{1}{2})$

Problem 4

First, let us compare the mean. for $\frac{T}{n-1}$, we have:

$$E\left(\frac{T}{n-1}\right) = E\left(\frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - n\bar{X}^2\right)\right) = \frac{1}{n-1} (nE(X_1^2) - nE(\bar{X}^2)) = \frac{1}{n-1} (n(\sigma^2 + \mu^2) - n(\frac{\sigma^2}{n} + \mu^2)) = \sigma^2$$

$$E\left(\frac{T}{n}\right) = E\left(\frac{n-1}{n} \frac{T}{n-1}\right) = \frac{n-1}{n} E\left(\frac{T}{n-1}\right) = \frac{n-1}{n} \sigma^2$$

We can get the conclusion that $n^{-1}T$ is a bias estimator while $(n-1)^{-1}T$ is an unbiased estimator. $(n-1)^{-1}T$ seems a better estimator under this circumstance.

Now, let us consider the variance. For this part, we will use a conclusion that $\frac{T}{\sigma^2} \sim \chi^2(n-1)$. [1]

We can know:

$$\begin{aligned} \text{Var}(\chi^2(n-1)) &= 2(n-1) \\ \text{Var}\left(\frac{T}{n-1}\right) &= \frac{1}{(n-1)^2} \text{Var}(\sigma^2 \chi^2(n-1)) = \frac{2}{n-1} \sigma^4 \\ \text{Var}\left(\frac{T}{n}\right) &= \frac{2(n-1)}{n^2} \sigma^4 \end{aligned}$$

It is easy to see that $n^{-1}T$ has smaller variance than $(n-1)^{-1}T$. Hence $n^{-1}T$ is a better estimator when we only consider variance. In fact, $n^{-1}T$ also has smaller mean square error than $(n-1)^{-1}T$.

References

- [1] William G Cochran. The distribution of quadratic forms in a normal system, with applications to the analysis of covariance. In *Mathematical Proceedings of the Cambridge Philosophical Society*, volume 30, pages 178–191. Cambridge University Press, 1934.
- [2] Clare Brown ([https://stats.stackexchange.com/users/38489/clare brown](https://stats.stackexchange.com/users/38489/clare%20brown)). Simple examples of uncorrelated but not independent x and y . Cross Validated. URL:<https://stats.stackexchange.com/q/85363> (version: 2016-03-05).