# Stat511(Section001): Homework #10

Due on Apr. 13, 2022 at 5:00pm

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#### Problem 1

(a)

We know under  $H_0$  or  $H_1$ ,  $\hat{\mu} = \bar{X}$ . And under  $H_1$ , we also know the the MLE of  $\sigma^2$  is:

$$\hat{\sigma^2} = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X})^2$$

Then we can get the log LR for this test is:

$$logL = maxL_{H_0}(x) - maxLH_1(x) = -\frac{n}{2}log(\sigma_0^2) + \frac{n}{2}log(\hat{\sigma}^2) - \frac{n}{2}\frac{\hat{\sigma}^2}{\sigma_0^2} + \frac{n}{2} = \frac{n}{2}log(\frac{\hat{\sigma}^2}{\sigma_0^2}) - \frac{n}{2}\frac{\hat{\sigma}^2}{\sigma_0^2} + \frac{n}{2}$$

We also know that, under  $H_0$ :

$$\frac{n\hat{\sigma^2}}{\sigma_0^2} \sim \chi^2(n-1)$$

Then choose  $\frac{n\hat{\sigma^2}}{\sigma_0^2} = T$ , we have:

$$logL = \frac{n}{2}log(\frac{T}{n}) - \frac{1}{2}T + \frac{n}{2}$$

We will reject  $H_0$ , when log L < k. And we can see that:

$$\frac{dlogL}{dT} = \frac{n}{2T} - \frac{1}{2}$$

Thus, when T < n, log L is a increasing function of T, and when T > n it is a decreasing function of T. Hence:

$$log L < k \rightarrow T > k_1 \ or \ T < k_2$$

Then we can know:

$$k_1 = \chi^2(n-1)_{1-\alpha/2}, k_2 = \chi^2(n-1)_{\alpha/2}$$

where  $\chi^2(n-1)_{\alpha/2}$  is the  $\frac{\alpha}{2}$  percentile of  $\chi^2(n-1)$ .

(b)

We know the  $1 - \alpha$  confidence interval for T is:

$$k_2 < T < k_1 \to k_2 < \frac{n\hat{\sigma}^2}{\sigma^2} < k_1$$

Thus we can get:

$$\frac{n\hat{\sigma^2}}{k_1} < \sigma^2 < \frac{n\hat{\sigma^2}}{k_2} \to \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{k_1} < \sigma^2 < \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{k_2}$$

 $[\frac{\sum_{i=1}^{n}(X_{i}-\bar{X})^{2}}{k_{1}}, \frac{\sum_{i=1}^{n}(X_{i}-\bar{X})^{2}}{k_{2}}] \text{ is the } 1-\alpha \text{ confidence interval of } \sigma^{2}, \text{where } k_{1}=\chi^{2}(n-1)_{1-\alpha/2}, k_{2}=\chi^{2}(n-1)_{\alpha/2}.$ 

## Problem 2

(a)

We know:

$$E(X_i) = \lambda$$

$$Var(X_i) = \lambda$$

Thus, by central limit theorem, we have:

$$\sqrt{n}(\bar{X} - \lambda) \sim N(0, \lambda) \rightarrow \frac{T - n\lambda}{\sqrt{n\lambda}} \sim N(0, 1)$$

And we can get:

$$\frac{T - n\lambda}{\sqrt{2n\lambda}} \sim N(0, \frac{1}{2})$$

Thus it is asymptotically pivotal.

(b)

From (a), we know the  $1-\alpha$  confidence level of  $\frac{T-n\lambda}{\sqrt{n\lambda}}$  is:

$$z_{\alpha/2} < \frac{T - n\lambda}{\sqrt{n\lambda}} < z_{1-\alpha/2}$$

where  $z_{\alpha/2}$  is  $\frac{\alpha}{2}$  percentile of standard normal.  $z_{1-\alpha/2}$  is  $1-\frac{\alpha}{2}$  precentile of standard normal. Then we have:

$$z_{\alpha/2} < \frac{T - n\lambda}{\sqrt{n\lambda}} < -z_{\alpha/2}$$

Solving to get:

$$\frac{2T + z_{\alpha/2}^2 - \sqrt{4Tz_{\alpha/2}^2 + z_{\alpha/2}^4}}{2n} < \lambda < \frac{2T + z_{\alpha/2}^2 + \sqrt{4Tz_{\alpha/2}^2 + z_{\alpha/2}^4}}{2n}$$

And that is the confidence interval for  $\lambda$ . (c)

By (b), we can get the 95% interval is [1.094, 2.888]

### Problem 3

It easy to know if X is the number of success in 5 independent missile trials. Then we know  $X \sim Bin(5, p)$ . Now to seek lower confidence bound. We consider  $H_0: p = p_0, H_1: p > p_0$  with  $\alpha = 0.1$ . Then under  $H_0$ , we have:

$$P(X \ge 4) = 5p_0^4(1 - p_0) + p_0^5 = 5p_0^4 - 4p_0^5 > 0.1$$

Then we can have  $p_0 = 0.416$ . And we can get 90% lower confidence bound is 0.416.

#### Problem 4

(a)

We know the likelihood function for one  $X_i$  can been written as:

$$f(X_i) = (1 - 2\theta)^{I(X_i = 0)} \theta^{I(X_i = 1) + I(X_i = -1)} = (1 - 2\theta)^{I(X_i = 0)} \theta^{1 - I(X_i = 0)}$$

Thus, take  $T = sum_{i=1}^n I(X_i = 0)$ , then we can get the likelihood function of  $X_1, ..., X_n$  is:

$$L(\theta) = \theta^{n-T} (1 - 2\theta)^T$$

$$logL(\theta) = (n-T)log(\theta) + Tlog(1-2\theta)$$

Then, we have the MLE of  $\theta$  is:

$$hat\theta = \frac{n-T}{2n}$$

Hence, under  $H_1$ , if  $\frac{n-T}{2n} > \theta_0$   $\theta_{H_1} = \frac{n-T}{2n}$ . else we get  $\theta_{H_1} = \theta_0$ . Thus we can get the Likelihood Ratio is:

$$L = \begin{cases} 1 & \theta_0 > \frac{n-T}{2n} \\ \frac{\theta_0^{n-T} (1 - 2\theta_0)^T}{(\frac{n-T}{2n})^{n-T} (\frac{T}{n})^T} & else \end{cases}$$

This is a function a T. We will reject  $H_0$  when this function is less than  $k_1$ . Then we have:

$$\frac{dlogL}{dT} = log(\frac{n-T}{T}) + log(\frac{1-2\theta_0}{\theta_0}) - log2 \geq 0$$

Thus  $L < k_1$  is equivalent to T < k. That is end of our proof.

(b)

We have the reject region  $T < k(\alpha)$ . Thus  $T \ge k(\alpha)$  is the acceptance region. Then we can get  $1 - \alpha$  lower confidence bound for  $\theta$  is just  $k(\alpha)$ .

(c)

We know under  $H_0$   $T \sim Bin(10, \frac{1}{3})$ . Then we need to find to biggest k.  $P(T < k) \le 0.05$ . And By python, we get k = 0.