Mathematical Tools for CS - Lecture 2

February 20, 2024

Today's Plan

- See some tools/examples for the analysis of probabilities and random variables.
- Threshold Phenomena

• ..

1 Copies of k_4 in G(n, p)

Let $G \sim G(n, p)$.

 $X_{n,p}=X$ Counts the nuber of copies of k_4 (clique of size 4) in G.

$$q_{n,p} = q = Pr(X_{n,p} \ge 1).$$

We will show the following threshold phenomonon:



Lemma. $\mathbb{E}(X_{n.p}) = \binom{n}{4} \cdot p^6$

Proof. For
$$S\subseteq [n]$$
 where $|S|=4$, define $X_S=\begin{cases} 1 & S \text{ induces } k_4\\ 0 & \text{otherwise} \end{cases}$. So:

$$X = \sum_{S \in \binom{[n]}{4}} X_S$$

by linearity of expectation \implies

$$\mathbb{E}(X) = \sum_{S \in \binom{[n]}{4}} \mathbb{E}(X_S) = \sum_{S \in \binom{[n]}{4}} Pr(S \text{ induces } k_4) = \sum_{S \in \binom{[n]}{4}} p^6 = \binom{n}{4} \cdot p^6$$

 $\begin{array}{ll} \textbf{Corollaries} & \Rightarrow \text{By markov inequality: } q_{n,p} \leq \frac{\mathbb{E}(X)}{1} = \binom{n}{4} \cdot p^6 = \Theta(n^4 \cdot p^6) \\ \Rightarrow \text{If } p << n^{-\frac{2}{3}} \text{ then } q_{n,p} << 1 \\ \end{array}$

Lemma.
$$\frac{Var(X)}{\mathbb{E}(X)^2} \leq \Theta(n^{-2} \cdot p^{-1} + n^{-3} \cdot p^{-3} + n^{-4} \cdot p^{-6})$$

Proof. Same decomposition as previous proof:

$$X = \sum_{S \in \binom{[n]}{4}} X_S$$

$$\begin{split} Var(X) &= \mathbb{E}(X - \mathbb{E}(X))^2 \\ &= \mathbb{E}(\sum_{S \in \binom{[n]}{4}} X_S - \mathbb{E}(\sum_{S \in \binom{[n]}{4}} X_S))^2 \\ &= \mathbb{E}(\sum_{S \in \binom{[n]}{4}} X_S - \sum_{S \in \binom{[n]}{4}} \mathbb{E}(X_S))^2 \\ &= \mathbb{E}(\sum_{S \in \binom{[n]}{4}} X_S - \sum_{S \in \binom{[n]}{4}} \mathbb{E}(X_S))^2 \\ &= \mathbb{E}(\sum_{S \in \binom{[n]}{4}} \sum_{T \in \binom{[n]}{4}} \mathbb{E}(X_S - \mathbb{E}(X_S)) (X_T - \mathbb{E}(X_T))) \\ &= \mathbb{E}\left(\sum_{S \in \binom{[n]}{4}} \sum_{T \in \binom{[n]}{4}} (X_S - \mathbb{E}(X_S)) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left((X_S - \mathbb{E}(X_S)) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T))\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_S) (X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_T - \mathbb{E}(X_T)\right) \\ &= \sum_{S,T \in \binom{[n]}{4}} \mathbb{E}\left(X_S - \mathbb{E}(X_T -$$

*: because X_S, X_T are positive:

$$\mathbb{E}\left(\left(X_{s} - \mathbb{E}(X_{S})\right)\left(X_{T} - \mathbb{E}(X_{T})\right)\right) = \mathbb{E}\left(X_{s}X_{t} - X_{S}\mathbb{E}(X_{T}) - \mathbb{E}(X_{S})X_{T} + \mathbb{E}(X_{S})\mathbb{E}(X_{T})\right)$$
$$= \mathbb{E}(X_{s}X_{t}) - \mathbb{E}(X_{S})\mathbb{E}(X_{T}) \leq \mathbb{E}(X_{s}X_{t})$$

Corollaries \Rightarrow by chevishev inequality (ex1/2) $q_{n,p} \ge 1 - \frac{Var(X)}{\mathbb{E}(X)^2}$ \Rightarrow if $p >> n^{-\frac{2}{3}}$ then $q_{n,p} = 1 - o(1)$

2 Birthday Paradox & Coupon Collector Problem

Consider the following: n bins, and k balls. We throw the balls into the bins uniformally at random.

Question (Birthday Paradox): What is the probability of the event $B_{n,k} = B$ where all balls fall into different pits.

Answer: (Threshold Phenomenon)

$$k >> \sqrt{2n} \implies Pr(B) = o(1)$$

$$k << \sqrt{2n} \implies Pr(B) = 1 - o(1)$$

Question (Coupon Collector): Assuming we have infinite balls, how many balls should we throw until (with high probability) each bin contains at least 1 balls.

Answer: (Threshold Phenomenon)

 $k = nlog(n) + \omega(n) \implies$ with high prob (1 - $o_n(1)$) all bins have a ball

 $k = nlog(n) - \omega(n) \implies$ with high prob (1 - $o_n(1)$) all bins have a ball there is an empty bin.

Definition. A random variable X is Geometric with parameter $0 \le p \le 1$ if $Pr(X = k) = (1 - P)^{k-1} \cdot p$. (Amount of tries until first success).

Fact.
$$\mathbb{E}(X) = \frac{1}{p}$$
, $Var(X) = \frac{1-p}{p^2} \le \frac{1}{p^2}$

Fact.
$$\sum_{k=1}^{n} \frac{1}{k} = ln(n) + \Theta(1)$$

Coupon Collector

 T_n - number of balls that were thrown until all bins had a ball.

Lemma.
$$\mathbb{E}(T_n) = n \cdot ln(n) + \Theta(1)$$

Proof. $T_{n,m} := \text{Number of balls until } m \text{ bins are full.}$

$$T_n = T_{n,n} = \sum_{m=1}^{n} \underbrace{(T_{n,m} - T_{n,m-1})}_{:=S_m}$$

 S_m is a geometric R.V. with parameter $\frac{n-(m-1)}{n}.$

$$\Rightarrow \mathbb{E}(T_n) = \sum_{m=1}^n \mathbb{E}(S_m)$$

$$= \sum_{m=1}^n \frac{n}{n - (m-1)}$$

$$= n \cdot \sum_{m=1}^n \frac{1}{n - (m-1)}$$

$$= n \cdot \sum_{m=1}^n \frac{1}{m}$$

$$= n \cdot (\ln(n) + \Theta(1))$$

Now for the variance:

$$Var(T_n) = Var(\sum_{m=1}^n S_m)$$

$$independence = \sum_{m=1}^n Var(S_m)$$

$$\leq \sum_{m=1}^n \frac{n^2}{(n - (m-1))^2}$$

$$= n^2 \sum_{m=1}^n \frac{1}{(n - (m-1))^2}$$

$$= n^2 \sum_{m=1}^n \frac{1}{m^2}$$

$$= \Theta(n^2)$$

Birthday Paradox

Mark $B_{n,k}$ The event that all balls fall into different bins. so:

$$Pr(B_{n,k}) = 1 \cdot \left(\frac{n-1}{n}\right) \cdot \left(\frac{n-2}{n}\right) \cdot \dots \cdot \left(\frac{n-(k-1)}{n}\right)$$

Fact. for any x, it hold that $1+x \leq e^x$ additionally, for sufficiently small x, $1+x \geq e^{x-x^2}$

so:

$$Pr(B_{n,k}) = \prod_{i=1}^{k-1} \left(1 - \frac{i}{n}\right)$$

$$\approx \prod_{i=1}^{k-1} e^{-\frac{i}{n}}$$

$$= e^{-\sum_{i=1}^{k-1} \frac{i}{n}}$$

$$= e^{-\frac{1}{n} \cdot \sum_{i=1}^{k-1} i}$$

$$= e^{-\frac{k \cdot (k-1)}{2n}}$$

But \approx is not formal. formally we know for the upper bound:

$$Pr(B_{n,k}) \le e^{-\frac{k \cdot (k-1)}{2n}}$$

Also:

$$k >> \sqrt{2n} \implies Pr(B_{n,k}) = o(1)$$

Now, assuming $k < \sqrt{2n}$:

$$Pr(B_{n,k}) \ge \prod_{i=1}^{k-1} e^{-\frac{i}{n} - (\frac{i}{n})^2}$$

$$= e^{-\sum_{i=1}^{k-1} \frac{i}{n}} \cdot e^{-\sum_{i=1}^{k-1} (\frac{i}{n})^2}$$

$$\ge e^{-\frac{k \cdot (k-1)}{2n}} \cdot e^{-\sum_{i=1}^{k} (\frac{k}{n})^2}$$

$$= e^{-\frac{k \cdot (k-1)}{2n}} \cdot \underbrace{e^{-\frac{k^3}{n^2}}}_{1+o(1)}$$