LASSO Prediction on Churn Data

```
library(data.table) # Open the data.table library for use in this script
library(ggplot2)
library(glmnet)

set.seed(2093) # Set the seed so that all of our random operations produce the same results
```

CRM

CRM stands for Customer Retail Management. This refers to when the marketer has enough data to tailor the marketing experience of each customer. Here we will explore models that can be used to predict whether a customer will churn.

Data

The data is a development sample that consists of 100,000 randomly selected customers from the data base of a a credit card company in Singapore. The random sample represents only a small percentage of the whole data base.

Customers are identified by a unique customer_id. The data also include demographic information on each customer, including whether their gender is male (genderM), their employment classification (employXX), income (hhIncome), household size (hhSize), whether they own a home (homeOwner), marital status (married). Additionally, some partial information on each customer's purchase history is included: the share of past months they have made late payments (avgLatePayment), the number of months they have been a customer of the credit card company (nMonthsCust), their average monthly bill (avgMoBill) and their credit score (creditScore).

Load the churnCC data set. We will use the fread command from the data.table package which opens text files like ".csv" much, much faster than base-R commands like read.csv.

```
churn_DT = fread("~/Dropbox/LargeRepoFiles/teaching/Teaching Data/output/Churn CC/data/churnCC.csv")
churn_DT[,.N] # .N is a special character for calling the number of observations in the data.table
churn_DT = churn_DT[runif(.N) <.01] # This creates a subset of the original data for testing/developm
churn_DT[,.N]</pre>
```

Data inspection

Summarize some key aspects of the churned variable. In particular, what is the churn incidence, and the average churn value? Note that data.table offers a special character that may be used in the j position, .N which we will use here. .N reports the number of observations in the data set. When .N is combined with a by variable, then data.table reports the number of observations for each by group.

```
# Number of observations for each value of `churned`
churn_DT[, .N, churned]
   churned
         0 7650
1:
2:
         1 2372
# Share of observations for which `churned` is zero or one.
churn_DT[, .N / churn_DT[,.N], churned]
   churned
                  V1
         0 0.7633207
1:
         1 0.2366793
# Average value of `churned` which matches the output above.-
churn_DT[, mean(churned)]
```

[1] 0.2366793

Summarize the data

summary(churn DT)

```
genderM
  customerID
                                      employSelfEmployed
                                                           employCrafts
       :10000054
                            :0.0000
                                              :0.00000
                                                                  :0.0000
Min.
                    Min.
                                                           Min.
1st Qu.:10243584
                    1st Qu.:0.0000
                                      1st Qu.:0.00000
                                                           1st Qu.:0.0000
Median :10499934
                    Median :1.0000
                                      Median :0.00000
                                                           Median :0.0000
Mean
       :10497001
                    Mean
                            :0.5019
                                      Mean
                                              :0.06735
                                                           Mean
                                                                  :0.1107
3rd Qu.:10744425
                    3rd Qu.:1.0000
                                      3rd Qu.:0.00000
                                                           3rd Qu.:0.0000
       :10999971
                            :1.0000
Max.
                    Max.
                                      Max.
                                              :1.00000
                                                           Max.
                                                                  :1.0000
employProfessional
                    employClerical
                                        employRetired
                                                           employStudent
                            :0.0000
                                               :0.00000
                                                                  :0.0000
Min.
       :0.0000
                    Min.
                                       Min.
                                                           Min.
1st Qu.:0.0000
                    1st Qu.:0.00000
                                       1st Qu.:0.00000
                                                           1st Qu.:0.00000
Median :1.0000
                    Median :0.00000
                                       Median :0.00000
                                                           Median :0.00000
Mean
       :0.6488
                    Mean
                            :0.07633
                                       Mean
                                               :0.05448
                                                           Mean
                                                                  :0.03233
3rd Qu.:1.0000
                    3rd Qu.:0.00000
                                       3rd Qu.:0.00000
                                                           3rd Qu.:0.00000
Max.
       :1.0000
                    Max.
                            :1.00000
                                       Max.
                                               :1.00000
                                                           Max.
                                                                  :1.00000
   hhIncome
                      hhSize
                                     homeOwner
                                                        married
       :
                          :1.000
                                           :0.0000
                                                             :0.0000
Min.
              0
                  Min.
                                   Min.
                                                     Min.
1st Qu.: 40000
                  1st Qu.:5.000
                                   1st Qu.:1.0000
                                                     1st Qu.:0.0000
Median: 60000
                  Median :5.000
                                   Median :1.0000
                                                     Median :1.0000
Mean
       : 60010
                          :4.561
                                   Mean
                                           :0.8639
                                                     Mean
                                                             :0.7094
                  Mean
3rd Qu.: 80000
                  3rd Qu.:5.000
                                   3rd Qu.:1.0000
                                                     3rd Qu.:1.0000
Max.
       :180000
                          :5.000
                                   Max.
                                           :1.0000
                                                             :1.0000
                  Max.
                                                     Max.
avgLatePayment
                      nMonthsCust
                                          avgMoBill
                                                          creditScore
       :3.901e-05
                             : 0.000
                                               : 10.0
                                                                :200.0
Min.
                     Min.
                                       Min.
                                                        Min.
1st Qu.:5.016e-02
                     1st Qu.: 3.000
                                       1st Qu.: 80.0
                                                         1st Qu.:440.0
Median :1.008e-01
                     Median : 4.000
                                       Median:100.0
                                                        Median :520.0
                             : 3.977
                                               :100.2
                                                                :520.2
Mean
       :1.003e-01
                     Mean
                                       Mean
                                                        Mean
3rd Qu.:1.506e-01
                     3rd Qu.: 5.000
                                       3rd Qu.:120.0
                                                         3rd Qu.:600.0
                             :14.000
                                               :240.0
                                                                :800.0
Max.
       :2.000e-01
                     Max.
                                       Max.
                                                        Max.
   churned
Min.
       :0.0000
1st Qu.:0.0000
Median :0.0000
Mean
       :0.2367
3rd Qu.:0.0000
Max.
       :1.0000
```

In this analysis we would like to predict who will churn, or the churned variable. We should take a minute to think about which of these variables we want to use to predict churned. Most of the variables appear to have the potential to predict churned, however, there are a few that should not be included as predictive variables.

First, customerID should have no relationship to whether a customer churned; these values are just a way to track customers. Additionally, even if customerID was somehow able to predict churned we would not want to use it because when we apply our model to predict churn for a different group of customers, their customerID values will be different. In this case, a model that includes customerID will not be helpful. In general, we want to restrict our predictor variables to those that will be available to us, with the same types of values.

Add additional variables that are random numbers

To help illustrate the value of LASSO regression, we are going to add several new variables that are random and do not help predict churned. (Normally, we would not do this.)

```
# This syntax is more advanced data.table and not a core part of this course
# This code creates many "garbage variables" that are just random noise.

garbNum = 300
garbVars = paste0("X", 1:garbNum)
churn_DT[, (garbVars) := 1]
churn_DT[, (garbVars) := lapply(.SD, function (x) runif(.N) ), .SDcols = garbVars]
```

Holdout sample

Because we want to make predictions and assess the quality of those predictions, we need to produce a **confusion matrix** for our predictions. To produce this matrix we need observations which include both our model's predictions and the actual outcomes. How do we generate such data? Well, we already have observations with actual outcomes, so we just need to add model predictions.

However, we have to be careful not to use the same data for estimating the model and for producing the **confusion matrix**. Using the same data leads to **overfit** a topic we will discuss further. For now, suffice it to say that we want to evaluate the predictions our model on data we have never seen before. After all, this is how we plan to use the predictions, so this is how we should evaluate their quality.

A good way to estimate the **confusion matrix** is by generating a **holdout sample**. A holdout sample is a random subset of the data we have. We will "estimate" or "train" our model using the "training data." Then, with our model estimated, we will test its predictions using the holdout data. The holdout data will be used to produce the confusion matrix.

```
churn_DT[, holdout := sample(0:1, size=.N, prob=c(0.5, 0.5), replace=TRUE)]
```

Estimate the model

For now we will use a standard logit model to make our predictions. This step is also known as *training* the model or *calibrating* the model. In later classes we will introduce other methods for making predictions. First we create a vector of the names of variables we don't want included in the regression. Placing them in noRegVar allows us to refer to these variables later.

```
# Variables that are not to be used for estimation
names(churn_DT)
```

```
[1] "customerID"
                            "genderM"
                                                   "employSelfEmployed"
                            "employProfessional"
                                                  "employClerical"
 [4] "employCrafts"
                                                   "hhIncome"
 [7] "employRetired"
                            "employStudent"
[10] "hhSize"
                            "homeOwner"
                                                   "married"
                                                   "avgMoBill"
[13] "avgLatePayment"
                            "nMonthsCust"
[16] "creditScore"
                            "churned"
                                                   "X1"
[19] "X2"
                            "X3"
                                                   "X4"
                            "X6"
[22] "X5"
                                                   "X7"
[25] "X8"
                            "X9"
                                                   "X10"
[28] "X11"
                            "X12"
                                                   "X13"
[31] "X14"
                            "X15"
                                                   "X16"
[34] "X17"
                            "X18"
                                                   "X19"
[37] "X20"
                            "X21"
                                                   "X22"
[40] "X23"
                            "X24"
                                                   "X25"
[43] "X26"
                            "X27"
                                                   "X28"
```

[46]	"X29"	"X30"	"X31"
[49]	"X32"	"X33"	"X34"
[52]	"X35"	"X36"	"X37"
[55]	"X38"	"X39"	"X40"
[58]	"X41"	"X42"	"X43"
[61]	"X44"	"X45"	"X46"
[64]	"X47"	"X48"	"X49"
[67]	"X50"	"X51"	"X52"
[70]	"X53"	"X54"	"X55"
[73]	"X56"	"X57"	"X58"
[76]	"X59"	"X60"	"X61"
[79]	"X62"	"X63"	"X64"
[82]	"X65"	"X66"	"X67"
[85]	"X68"	"X69"	"X70"
[88]	"X71"	"X72"	"X73"
[91]	"X74"	"X75"	"X76"
[94]	"X77"	"X78"	"X79"
[97]	"X80"	"X81"	"X82"
[100]	"X83"	"X84"	"X85"
[103]	"X86"	"X87"	"X88"
[106]	"X89"	"X90"	"X91"
[109]	"X92"	"X93"	"X94"
[112]	"X95"	"X96"	"X97"
[115]	"X98"	"X99"	"X100"
[118]	"X101"	"X102"	"X103"
[121]	"X104"	"X105"	"X106"
[124]	"X107"	"X108"	"X109"
[127]	"X110"	"X111"	"X112"
[130]	"X113"	"X114"	"X115"
[133]	"X116"	"X117"	"X118"
[136]	"X119"	"X120"	"X121"
[139]	"X122"	"X123"	"X124"
[142]	"X125"	"X126"	"X127"
[145]	"X128"	"X129"	"X130"
[148] [151]	"X131"	"X132" "X135"	"X133" "X136"
[154]	"X134"	"X138"	"X136"
[154]	"X137" "X140"	"X141"	"X142"
[160]	"X143"	"X144"	"X142"
[163]	"X146"	"X147"	"X148"
[166]	"X149"	"X150"	"X151"
[160]	"X152"	"X153"	"X151"
[172]	"X155"	"X156"	"X154"
[175]	"X158"	"X159"	"X160"
[178]	"X161"	"X162"	"X163"
[181]	"X164"	"X165"	"X166"
[184]	"X167"	"X168"	"X169"
[187]	"X170"	"X171"	"X172"
[190]	"X173"	"X174"	"X175"
[190]	"X176"	"X177"	"X178"
[196]	"X179"	"X180"	"X181"
[190]	"X182"	"X183"	"X184"
[202]	"X185"	"X186"	"X187"
[202]	"X188"	"X189"	"X190"
[200]			11100

```
[208] "X191"
                             "X192"
                                                   "X193"
[211] "X194"
                             "X195"
                                                   "X196"
                                                   "X199"
[214] "X197"
                             "X198"
[217] "X200"
                             "X201"
                                                   "X202"
[220] "X203"
                             "X204"
                                                   "X205"
[223] "X206"
                             "X207"
                                                   "X208"
[226] "X209"
                             "X210"
                                                   "X211"
                                                   "X214"
[229] "X212"
                             "X213"
[232] "X215"
                             "X216"
                                                   "X217"
[235] "X218"
                             "X219"
                                                   "X220"
[238] "X221"
                             "X222"
                                                   "X223"
[241] "X224"
                             "X225"
                                                   "X226"
[244] "X227"
                             "X228"
                                                   "X229"
[247] "X230"
                             "X231"
                                                   "X232"
[250] "X233"
                             "X234"
                                                   "X235"
[253] "X236"
                             "X237"
                                                   "X238"
[256] "X239"
                             "X240"
                                                   "X241"
[259] "X242"
                             "X243"
                                                   "X244"
[262] "X245"
                             "X246"
                                                   "X247"
[265] "X248"
                             "X249"
                                                   "X250"
[268] "X251"
                             "X252"
                                                   "X253"
[271] "X254"
                             "X255"
                                                   "X256"
[274] "X257"
                             "X258"
                                                   "X259"
[277] "X260"
                             "X261"
                                                   "X262"
[280] "X263"
                             "X264"
                                                   "X265"
[283] "X266"
                             "X267"
                                                   "X268"
[286] "X269"
                             "X270"
                                                   "X271"
[289] "X272"
                             "X273"
                                                   "X274"
                             "X276"
[292] "X275"
                                                   "X277"
[295] "X278"
                             "X279"
                                                   "X280"
[298] "X281"
                             "X282"
                                                   "X283"
[301] "X284"
                             "X285"
                                                   "X286"
[304] "X287"
                             "X288"
                                                   "X289"
[307] "X290"
                             "X291"
                                                   "X292"
[310] "X293"
                             "X294"
                                                   "X295"
[313] "X296"
                             "X297"
                                                   "X298"
[316] "X299"
                             "X300"
                                                   "holdout"
noRegVars = c("holdout", "customerID")
churn.fit.base = glm(churned ~ ., family="binomial",
   data=churn_DT[holdout==0, -c(..noRegVars, ..garbVars)] )
# Estimate a logit model with garbage variables
churn.fit.garb = glm(churned ~ ., family="binomial",
   data=churn_DT[holdout==0, -..noRegVars] )
```

We use the glm command to estimate a logit model with the setting family="binomial". Note that the first entry in gml is the formula for the regression we want to estimate. churned ~ . indicates that we want to predict churned using all of the other variables in the data set, represented by ..

For the data set, we entered into to glm we have used some data.table commands to produce a subset of the original data. First, we selected only observations that are *not* in the holdout sample using holdout==0. In other words, we have selected the "training" data to estimate the model. Second, we removed the variables included in noRegVars using some data.table syntax. The minus sign - indicates that we are removing

the variables. The double dots, .. indicate that we are entering a vector of variable names and want to (de-)select those variables from the data set.

Inspect the estimates

```
summary(churn.fit.garb)
Call:
glm(formula = churned ~ ., family = "binomial", data = churn_DT[holdout ==
   0, -..noRegVars])
Deviance Residuals:
   Min
              10
                   Median
                               3Q
                                        Max
-2.9476
        -0.6948 -0.4356 -0.1656
                                    3.1015
Coefficients:
                     Estimate Std. Error z value Pr(>|z|)
                                         -2.805
(Intercept)
                   -3.537e+00
                              1.261e+00
                                                0.00504 **
genderM
                   -1.752e-01
                              7.793e-02
                                         -2.249
                                                 0.02454 *
employSelfEmployed -3.130e-01 3.964e-01
                                         -0.790
                                                 0.42969
employCrafts
                   -4.044e-01 3.871e-01
                                         -1.045
                                                 0.29620
employProfessional -6.853e-01 3.734e-01
                                         -1.835
                                                 0.06644 .
employClerical
                                         -0.839
                                                 0.40170
                   -3.302e-01
                              3.937e-01
employRetired
                   -2.941e-02 4.010e-01
                                         -0.073
                                                 0.94153
employStudent
                   -5.278e-01 4.297e-01
                                         -1.228
                                                 0.21929
hhIncome
                                         -2.609
                   -4.135e-06
                              1.585e-06
                                                 0.00907 **
hhSize
                   3.907e-02 3.749e-02
                                          1.042 0.29736
homeOwner
                   1.021e-01 1.164e-01
                                          0.877 0.38051
married
                   -1.388e-01 8.572e-02 -1.620 0.10534
avgLatePayment
                   5.782e-01
                              6.689e-01
                                          0.865
                                                 0.38731
nMonthsCust
                   9.718e-03
                                          0.499
                                                 0.61811
                              1.949e-02
avgMoBill
                   3.242e-02 1.396e-03
                                         23.223
                                                 < 2e-16 ***
creditScore
                                         -1.207
                   -4.242e-04 3.516e-04
                                                 0.22761
X1
                   -2.765e-02 1.354e-01
                                         -0.204
                                                 0.83819
X2
                   -3.111e-01 1.352e-01
                                         -2.301
                                                 0.02138 *
ХЗ
                   6.729e-02 1.342e-01
                                          0.501
                                                 0.61620
Х4
                                         -1.350
                   -1.848e-01 1.369e-01
                                                 0.17692
Х5
                   -3.662e-03
                              1.342e-01
                                         -0.027
                                                 0.97823
Х6
                                          1.316 0.18834
                   1.772e-01 1.347e-01
X7
                   1.272e-02 1.354e-01
                                          0.094 0.92517
Х8
                   -1.852e-01 1.345e-01
                                         -1.377
                                                 0.16856
Х9
                   -1.396e-01 1.368e-01
                                         -1.020
                                                 0.30753
                   -5.693e-02 1.360e-01
                                         -0.419 0.67544
X10
X11
                   -2.117e-01 1.362e-01
                                         -1.555
                                                 0.11999
X12
                   1.671e-02 1.344e-01
                                          0.124
                                                 0.90105
X13
                   -3.274e-02 1.334e-01
                                         -0.245
                                                 0.80612
X14
                   -1.335e-01 1.339e-01
                                         -0.997
                                                 0.31867
X15
                   2.141e-02 1.374e-01
                                          0.156
                                                 0.87613
X16
                   -2.615e-01 1.345e-01
                                         -1.944
                                                 0.05188
                   2.513e-02 1.371e-01
X17
                                          0.183 0.85459
X18
                   2.054e-01 1.353e-01
                                          1.518 0.12905
X19
                   2.595e-01 1.354e-01
                                          1.917
                                                 0.05525 .
X20
                   1.114e-01 1.362e-01
                                          0.818
                                                 0.41337
                   -7.357e-02 1.352e-01 -0.544 0.58633
X21
```

```
X22
                    9.944e-02 1.353e-01
                                            0.735 0.46225
X23
                    3.903e-02 1.327e-01
                                            0.294
                                                    0.76870
                                            2.432
                                                    0.01502 *
X24
                    3.275e-01
                               1.347e-01
X25
                    1.543e-01
                               1.339e-01
                                            1.153
                                                    0.24906
X26
                    -1.220e-01
                                1.351e-01
                                           -0.903
                                                    0.36649
                                           -0.583
X27
                    -7.936e-02 1.362e-01
                                                    0.56015
X28
                    -1.167e-01
                               1.360e-01
                                           -0.857
                                                    0.39119
                                                    0.35747
X29
                    1.250e-01
                                1.359e-01
                                            0.920
X30
                    -8.931e-02
                                1.358e-01
                                           -0.658
                                                    0.51062
X31
                    9.830e-02
                                1.345e-01
                                            0.731
                                                    0.46490
X32
                    2.600e-02
                                1.357e-01
                                            0.192
                                                    0.84807
X33
                                           -0.323
                    -4.366e-02
                                1.351e-01
                                                    0.74663
X34
                   -1.921e-01
                               1.350e-01
                                           -1.422
                                                    0.15488
X35
                    1.584e-01
                               1.329e-01
                                            1.192
                                                    0.23308
X36
                    1.080e-01
                                            0.801
                                                    0.42304
                               1.348e-01
X37
                    5.450e-02
                                1.342e-01
                                            0.406
                                                    0.68458
X38
                    -4.427e-03
                                1.359e-01
                                           -0.033
                                                    0.97401
X39
                    -4.531e-02
                                1.356e-01
                                           -0.334
                                                    0.73826
X40
                                            1.570
                    2.142e-01
                               1.364e-01
                                                    0.11632
X41
                    9.877e-02
                                1.350e-01
                                            0.732
                                                    0.46431
X42
                    -2.809e-02 1.344e-01
                                           -0.209
                                                    0.83441
X43
                    -1.619e-01
                               1.361e-01
                                           -1.190
                                                    0.23423
X44
                    -6.595e-02
                                           -0.485
                                                    0.62754
                               1.359e-01
X45
                                            0.020
                    2.759e-03
                                1.352e-01
                                                    0.98372
                                                    0.01318 *
X46
                    3.426e-01
                               1.382e-01
                                            2.479
X47
                    -1.789e-01
                                1.337e-01
                                           -1.338
                                                    0.18084
X48
                                            0.594
                                                    0.55232
                    8.038e-02
                                1.353e-01
X49
                    2.036e-01
                                1.337e-01
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                                                    0.12783
X50
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                    4.018e-03
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                                                    0.97653
X51
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                                                    0.84992
X52
                    -9.330e-02
                                1.359e-01
                                           -0.687
                                                    0.49239
X53
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                                1.349e-01
                                           -0.091
                                                    0.92776
X54
                    2.443e-02
                               1.347e-01
                                            0.181
                                                    0.85608
X55
                               1.355e-01
                                            1.209
                    1.639e-01
                                                    0.22660
X56
                    1.704e-01
                                1.352e-01
                                            1.261
                                                    0.20734
X57
                    3.320e-02
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X58
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                               1.343e-01
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                                                    0.19577
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                    2.159e-01
                               1.335e-01
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                                                    0.10595
X60
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                                1.353e-01
                                           -0.974
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                                            0.740
                    9.889e-02 1.336e-01
X61
                                                    0.45900
X62
                                            0.170
                    2.310e-02
                               1.360e-01
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X63
                    1.947e-01
                               1.346e-01
                                            1.446
                                                    0.14817
X64
                    5.929e-02 1.359e-01
                                            0.436
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X65
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                               1.340e-01
                                           -1.328
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X66
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                    -6.873e-02
                                           -0.510
                                                    0.61026
X67
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X68
                    -1.561e-02
                                1.363e-01
                                           -0.115
                                                    0.90878
X69
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                                1.344e-01
                                           -0.682
                                                    0.49527
X70
                    3.430e-01
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                                                    0.01178 *
X71
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                                1.362e-01
                                           -1.006
                                                    0.31443
X72
                    8.402e-02
                               1.356e-01
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                                                    0.53553
X73
                    -1.168e-01 1.365e-01
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                                                    0.39215
X74
                    8.782e-02 1.350e-01
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                                                    0.51525
X75
                    6.272e-02 1.377e-01
                                            0.455
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X78
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                                            -0.626
X79
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                                1.355e-01
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X80
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                                                    0.27691
X81
                     1.443e-01
                               1.354e-01
                                             1.066
                                                    0.28656
X82
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                                1.358e-01
                                             1.550
                                                    0.12112
X83
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                                1.349e-01
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X84
                     2.701e-02
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X85
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                                1.347e-01
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X86
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X87
                                            -0.600
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                                1.380e-01
                                                    0.54831
X88
                     3.203e-01
                                1.352e-01
                                             2.369
                                                    0.01785 *
                    -6.149e-02
X89
                                1.350e-01
                                            -0.455
                                                    0.64887
X90
                    9.841e-02
                               1.351e-01
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X91
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                                1.355e-01
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                                                    0.30208
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                                1.369e-01
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X93
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                                1.327e-01
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                                                    0.17378
                                            -1.359
X94
                    -1.824e-01
                                1.342e-01
                                                    0.17417
X95
                     2.789e-01
                                1.349e-01
                                             2.068
                                                    0.03862 *
X96
                     2.724e-01
                               1.363e-01
                                             1.999
                                                    0.04563 *
X97
                    -8.091e-02
                                            -0.595
                                1.361e-01
                                                    0.55214
X98
                    -5.001e-02
                                            -0.365
                                                    0.71492
                                1.369e-01
X99
                                            -0.418
                    -5.776e-02
                                1.383e-01
                                                    0.67625
X100
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                                1.351e-01
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                                                    0.97761
X101
                     1.979e-01
                                1.342e-01
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                                                    0.14043
X102
                                             2.179
                                                    0.02932 *
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                                1.347e-01
X103
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                                1.354e-01
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                                                    0.09354
                                             0.831
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                                1.356e-01
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X112
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X113
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X114
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X115
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X116
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X119
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X120
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                                1.352e-01
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X121
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X122
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                                1.349e-01
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X124
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X127
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X129
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X132
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X133
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X134
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                                1.357e-01
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                                                    0.99969
X135
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X136
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                                                    0.00538 **
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X137
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                                1.345e-01
                                            -1.027
X138
                    -1.549e-01
                                1.364e-01
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X139
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                                1.329e-01
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                                                    0.02998 *
X140
                     2.922e-02
                                1.328e-01
                                             0.220
                                                    0.82590
                                            -1.585
X141
                    -2.116e-01
                                1.335e-01
                                                    0.11300
X142
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                     1.510e-01
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X143
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                                            -1.764
                                                    0.07766
                                            -1.905
X144
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                                1.355e-01
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X145
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                                1.371e-01
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                                                    0.53060
X146
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X147
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X148
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X150
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                                1.375e-01
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X151
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X152
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                                1.350e-01
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X153
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                                1.347e-01
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                                                    0.22630
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                                1.359e-01
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                                                    0.23914
X155
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                                1.348e-01
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                                                    0.29669
X156
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                                1.358e-01
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X157
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                                1.342e-01
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                                                    0.31468
X158
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                                1.353e-01
                                             0.153
                                                    0.87852
X159
                                             2.572
                                                    0.01010 *
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                                1.393e-01
X160
                    -8.826e-02
                                1.350e-01
                                            -0.654
                                                    0.51324
X161
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                                                    0.01032 *
X162
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                                                    0.61746
X163
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                                1.374e-01
                                             1.094
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X164
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X165
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X167
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                                1.352e-01
                                             1.324
X168
                                                    0.18556
X169
                     3.149e-01
                                1.344e-01
                                             2.344
                                                    0.01909 *
X170
                     2.024e-01
                                1.356e-01
                                             1.493
                                                    0.13540
X171
                                             0.466
                     6.335e-02
                                1.360e-01
                                                    0.64144
X172
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X173
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                                1.343e-01
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X174
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X175
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X176
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                                                    0.83687
X178
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                                1.347e-01
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X179
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                                                    0.24852
X180
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                                1.352e-01
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                                                    0.71961
X181
                    1.095e-01
                                1.346e-01
                                             0.814
                                                    0.41593
X182
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                                                    0.69010
X183
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                                             1.985
                                                   0.04717 *
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X184
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                               1.358e-01
                                            -1.773 0.07629 .
X185
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                                                     0.85810
X186
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                                1.343e-01
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                                                     0.44175
X187
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X188
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X191
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X192
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                                                     0.47860
X193
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X194
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X195
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                                 1.356e-01
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X196
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                                             2.799
                     3.749e-01
                                                     0.00513 **
                                 1.351e-01
X197
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                                                     0.27784
X198
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X199
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X200
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X201
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X206
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                                1.348e-01
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X207
                     2.970e-01
                                 1.367e-01
                                             2.173
                                                     0.02980 *
X208
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X209
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X210
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                                1.358e-01
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X213
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                                1.361e-01
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X214
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                                                     0.54664
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                                1.354e-01
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                                                     0.45094
X216
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X217
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                                                    0.43505
X218
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X219
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                                1.358e-01
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                                                     0.33819
X220
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X221
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                                                    0.90954
                                1.349e-01
X222
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X223
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X225
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X238
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                                1.348e-01
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X246
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X247
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                                1.375e-01
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X248
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                                1.344e-01
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X249
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                                1.363e-01
                                            -0.190
                                                     0.84934
X250
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                                1.335e-01
                                            -2.403
                                                     0.01626 *
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                                1.360e-01
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X252
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                                                     0.94158
X253
                    -5.352e-02
                                1.343e-01
                                            -0.398
                                                     0.69029
X254
                                1.350e-01
                     4.234e-01
                                             3.137
                                                     0.00171 **
X255
                    -2.090e-01
                                1.361e-01
                                            -1.536
                                                     0.12462
X256
                    -8.626e-02
                                1.351e-01
                                            -0.638
                                                     0.52332
X257
                     2.149e-02
                                1.328e-01
                                             0.162
                                                     0.87144
X258
                     1.929e-01
                                1.355e-01
                                             1.424
                                                    0.15452
X259
                                            -0.066
                    -8.838e-03
                                1.340e-01
                                                     0.94742
                                                    0.47495
X260
                    -9.629e-02
                                1.348e-01
                                            -0.714
                                            -0.220
X261
                    -2.989e-02
                                1.361e-01
                                                     0.82613
X262
                    -2.606e-01
                                1.356e-01
                                            -1.922
                                                     0.05461
X263
                    -2.954e-01
                                1.340e-01
                                            -2.204
                                                     0.02755 *
X264
                                             1.452
                     1.966e-01
                                1.354e-01
                                                     0.14647
X265
                     1.427e-01
                                1.363e-01
                                             1.046
                                                     0.29534
X266
                     5.816e-02
                                1.359e-01
                                             0.428
                                                     0.66858
X267
                    -2.174e-01
                                            -1.604
                                                    0.10866
                                1.355e-01
X268
                     6.107e-03
                                1.335e-01
                                             0.046
                                                    0.96352
X269
                    -6.670e-02
                                1.348e-01
                                            -0.495
                                                     0.62067
X270
                     3.654e-02
                                1.371e-01
                                             0.267
                                                     0.78974
X271
                                            -0.820
                    -1.101e-01
                                1.343e-01
                                                     0.41229
X272
                     4.710e-02
                                1.359e-01
                                             0.347
                                                     0.72889
X273
                    -4.377e-02 1.362e-01
                                            -0.321
                                                     0.74786
X274
                     3.096e-01
                                1.345e-01
                                             2.301
                                                     0.02138 *
X275
                    -1.129e-01
                                            -0.830
                                                    0.40650
                                1.361e-01
X276
                                             1.350
                     1.841e-01
                                1.363e-01
                                                     0.17692
X277
                    -3.917e-02
                                1.352e-01
                                            -0.290
                                                     0.77202
X278
                    -9.402e-02
                                1.340e-01
                                            -0.702
                                                     0.48275
X279
                                            -1.701
                    -2.305e-01
                                1.355e-01
                                                     0.08896
X280
                     6.473e-02
                                1.363e-01
                                             0.475
                                                     0.63480
X281
                     6.566e-02
                                1.350e-01
                                             0.486
                                                     0.62663
X282
                    -7.138e-02
                                1.353e-01
                                            -0.528
                                                     0.59784
X283
                                             1.454
                     1.968e-01
                                1.354e-01
                                                     0.14603
X284
                     2.173e-01
                                1.348e-01
                                             1.612
                                                    0.10687
X285
                     3.216e-02
                                1.379e-01
                                             0.233
                                                     0.81556
                                             0.187
X286
                     2.506e-02
                                1.343e-01
                                                     0.85203
X287
                    -4.410e-02
                                1.345e-01
                                            -0.328
                                                    0.74294
X288
                                1.371e-01
                                             0.862
                                                     0.38850
                     1.182e-01
X289
                    -1.698e-01
                                1.340e-01
                                            -1.267
                                                     0.20503
X290
                    -2.758e-01
                                1.356e-01
                                            -2.035
                                                     0.04189 *
X291
                    -3.469e-02 1.358e-01
                                            -0.255
                                                   0.79834
```

```
X292
                   1.148e-01 1.351e-01
                                          0.849
                                                 0.39561
                   1.044e-01 1.338e-01
                                          0.780
X293
                                                 0.43533
                                          0.157
X294
                   2.119e-02 1.352e-01
                                                 0.87549
X295
                  -4.318e-01 1.349e-01
                                         -3.202
                                                 0.00137 **
X296
                  -2.274e-02 1.332e-01
                                         -0.171
                                                 0.86444
                   7.367e-02 1.356e-01
                                          0.543 0.58683
X297
                  -1.379e-01 1.360e-01
X298
                                         -1.014 0.31049
X299
                   2.090e-01 1.354e-01
                                          1.544
                                                 0.12256
X300
                   8.687e-02 1.344e-01
                                          0.646
                                                0.51814
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 5520.5 on 5021 degrees of freedom
Residual deviance: 4431.0 on 4706 degrees of freedom
AIC: 5063
```

Number of Fisher Scoring iterations: 5

Based on the coefficient estimates and their z-values, only a few of the variables we used in the model are statistically significant. This is okay, we didn't have clear sense of which variables would be important. Additionally, because we are primarily interested in making prediction with this model, we are not especially interested in the coefficient estimates.

Make predictions in a Holdout Sample

Now that we have estimated the model, we can generate the confusion matrix. Let's create a new data.table just for this purpose:

```
churn_HO = churn_DT[holdout==1]
chHOmat = model.matrix( churned ~ ., churn_HO[, -..noRegVars])
```

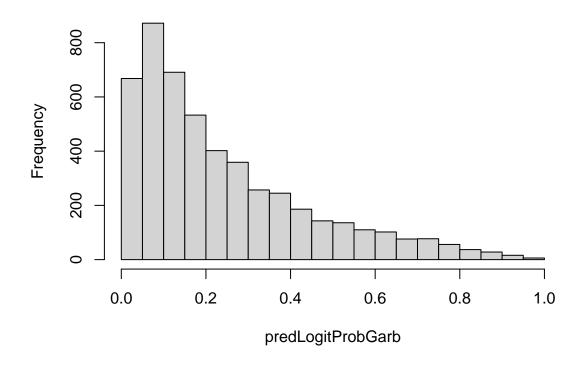
Next, let's make predictions for this data set. Recall that none of this data was used to estimate the model, which is important for the validity of the predictions. We want to test our "out-of-sample" predictions (predictions on data other than the training data) since that is how we will use the model.

```
churn_HO[, predLogitProbBase := predict(churn.fit.base, newdata=churn_HO, type = "response")]
churn_HO[, predLogitProbGarb := predict(churn.fit.garb, newdata=churn_HO, type = "response")]
```

This step used the estimated logit model to predict the probability that each customer in the holdout data would churn. Let's have a look at the distribution of these probabilities

```
churn_HO[, hist(predLogitProbGarb)]
```

Histogram of predLogitProbGarb



\$breaks

[1] 0.00 0.05 0.10 0.15 0.20 0.25 0.30 0.35 0.40 0.45 0.50 0.55 0.60 0.65 0.70 [16] 0.75 0.80 0.85 0.90 0.95 1.00

\$counts

[1] 668 872 691 533 402 359 257 245 186 143 136 110 102 76 77 56 37 28 16 [20] 6

\$density

[1] 2.672 3.488 2.764 2.132 1.608 1.436 1.028 0.980 0.744 0.572 0.544 0.440 [13] 0.408 0.304 0.308 0.224 0.148 0.112 0.064 0.024

\$mids

[1] 0.025 0.075 0.125 0.175 0.225 0.275 0.325 0.375 0.425 0.475 0.525 0.575 [13] 0.625 0.675 0.725 0.775 0.825 0.875 0.925 0.975

\$xname

[1] "predLogitProbGarb"

\$equidist

[1] TRUE

attr(,"class")
[1] "histogram"

Plot the Profit Curve

A function for calculating expected profits.

The following function, called expProfit provides a way to compute the expected profits associated with targeting different portions of the population. It should be applied to the holdout sample.

expProfit requires five input values. First is the name of the data.table that contains your holdout sample. Second is the name of the variable in the holdout sample that contains the **score** variables. In this application, this is predLogitProb. Third is the name of the variable that contains the actual outcomes that were observed. In this application, this is churned. Fourth is the value from the Cost-Benefit Matrix of correct predictions that the consumer will take an action. Fifth is the value from the Cost-Benefit Matrix

You are not responsible for understanding the coding details in this function, but I am happy to discuss them with you if interested.

```
# Function to calculate the expected profit at each score level
expProfit <- function(DT, score, actual, v11, v10) {</pre>
    # prepare copied data set for operations within the function with standardized names
    setnames(DT, c(actual, score), c("actual", "score"))
   DTloc <- copy(DT[, .(actual, score)]) # Copy a new version of the data set to work with in this fun
    setnames(DT, c("actual", "score"), c(actual, score)) # Return original variable names to original d
   DTloc[, origOrder := 1:.N] # Capture the order given for returning the output
    # Sort by score, highest to lowest.
    setorder( DTloc, -score )
   DTloc[, p11 := cumsum(actual==1) /.N] # Probability of true positives at a given score threshold
   DTloc[, p10 := cumsum(actual==0) /.N] # Probability of false positives at a given score threshold
   DTloc[, expProfit := v11*p11 + v10*p10] # Profits
    # Output expected profit estimates
    setorder(DTloc, origOrder)
   return(DTloc$expProfit)
}
```

Apply the expProfit function

```
### Base version of the model
# Apply the function to calculate expected profits for each threshold
churn_HO[, expProfitLogitBase := expProfit(
  DT = churn HO,
                          # data.table use for the calculation
  score = "predLogitProbBase", # Which variable corresponds to the score
  actual = "churned",
                       # Which variable corresponds to the actual outcomes
 v11 = 4,
                          # Value of a true positive
  v10 = -1)
                          # Value of a false positive.
### Version of the model with extra "garbage" variables
# Apply the function to calculate expected profits for each threshold
churn_HO[, expProfitLogitGarb := expProfit(
 DT = churn_H0,
                          # data.table use for the calculation
  score = "predLogitProbGarb", # Which variable corresponds to the score
 actual = "churned",  # Which variable corresponds to the actual outcomes
                          # Value of a true positive
 v11 = 4,
 v10 = -1)
                          # Value of a false positive.
```

Estimate using LASSO

We will use the glmnet function for LASSO and ridge regressions. glmnet does not take data.tables as its input, but instead requires that the X data be converted to a matrix and the y data a vector.

We will use the "cross-validated" version of glmnet, which is called cv.glmnet. It uses the cross validation method we described in class to help discover the optimal value for lambda.

```
# Convert the data.table to a matrix and vector to input to glmnet
X = model.matrix( churned ~ . , churn_DT[holdout==0, -..noRegVars] ) # Creates a matrix of the X variab
y = churn_DT[holdout==0, churned] # Creates a vector of the y variables
lasso.fit <- cv.glmnet(X, y, family="binomial", alpha=1.0) # Estimates the model for binary outcomes. a
# (Ridge can be estimated with alpha=0.)
coef(lasso.fit, s = "lambda.min") # Report the estimates for each coefficients estimated
317 x 1 sparse Matrix of class "dgCMatrix"
(Intercept)
                   -3.813829e+00
(Intercept)
genderM
                   -5.970516e-02
employSelfEmployed
employCrafts
employProfessional -2.220383e-01
employClerical
employRetired
                    9.774171e-02
employStudent
hhIncome
                   -1.234730e-06
hhSize
homeOwner
married
avgLatePayment
nMonthsCust
avgMoBill
                    2.775950e-02
creditScore
Х1
Х2
                   -4.351818e-02
ХЗ
Х4
Х5
Х6
X7
Х8
Х9
X10
X11
X12
X13
X14
X15
X16
                   -9.094375e-03
X17
X18
                    4.839712e-02
X19
X20
X21
```

```
X22
X23
X24
                    8.513802e-02
X25
X26
X27
X28
X29
X30
X31
X32
Х33
X34
X35
X36
X37
X38
X39
X40
X41
X42
X43
X44
X45
X46
                    1.109828e-01
X47
X48
X49
                    1.541288e-02
X50
X51
X52
X53
X54
X55
X56
X57
X58
X59
X60
X61
X62
X63
X64
X65
X66
X67
X68
X69
X70
                    8.799158e-02
X71
X72
X73
X74
X75
```

```
X76
X77
                   3.533740e-02
X78
X79
X80
X81
X82
                   2.576636e-03
X83
X84
X85
X86
X87
X88
                    1.033319e-01
X89
X90
X91
X92
X93
X94
                   -4.833706e-03
X95
X96
                   2.333157e-02
X97
X98
X99
X100
X101
                   1.262619e-03
X102
                   9.272279e-02
X103
X104
X105
X106
X107
X108
                   8.462068e-03
X109
X110
X111
X112
X113
X114
X115
X116
X117
X118
X119
X120
X121
X122
X123
X124
X125
X126
X127
X128
X129
```

```
X130
X131
X132
X133
X134
X135
X136
                  -1.014284e-01
X137
X138
                   -4.815335e-02
X139
X140
X141
X142
X143
X144
X145
X146
X147
X148
X149
X150
X151
X152
X153
X154
X155
X156
X157
X158
X159
                  8.493787e-02
X160
X161
                   -1.470461e-01
X162
X163
X164
X165
X166
X167
                   -7.960225e-02
X168
                   4.623966e-02
X169
X170
                   5.844784e-05
X171
X172
X173
X174
X175
X176
X177
X178
X179
X180
X181
X182
X183
                   3.040603e-02
```

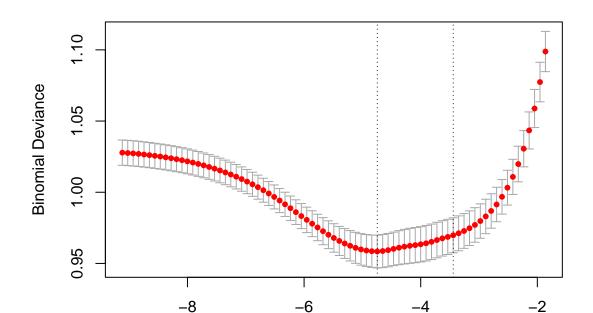
X184	-2.731483e-02
X185	
X186	
X187	•
X188	•
X189	•
X190	•
X191	-3.834900e-02
X192	•
X193	
X194	
X195	
X196	1.260179e-01
X197	
X198	
X199	•
X200	•
X201	•
X202	
X203	•
X204	•
X205	•
X206	•
X207	7.062899e-02
X208	•
X209	•
X210	•
X211	•
X212	•
X213	•
X214	•
X215	•
X216	•
X217	•
X218	•
X219	
X220	•
X221	
X222	
X223	•
X224	
X225	•
X226	•
X227	•
X228	•
X229	•
X230	
X231	
X232	•
X233	
X234	•
X235	
X236	
X237	

```
X238
                   -1.033301e-01
X239
                   -6.660676e-02
X240
                   -3.570075e-02
X241
X242
X243
X244
X245
X246
X247
                   -1.890971e-01
X248
X249
X250
                   -5.483952e-02
X251
X252
X253
X254
                   1.758450e-01
X255
                   -3.400405e-02
X256
X257
X258
X259
X260
X261
                   -6.914809e-02
X262
X263
                   -7.893707e-02
X264
X265
X266
X267
X268
X269
X270
X271
X272
X273
X274
                    5.608535e-02
X275
X276
X277
X278
X279
X280
X281
X282
X283
X284
X285
X286
X287
X288
X289
X290
                   -1.258814e-02
X291
```

```
X292
X293
X294
X295
                     -1.977191e-01
X296
X297
X298
X299
X300
```

312

plot(lasso.fit) # Plot the quality of the predictions for different lambda values.



 $Log(\lambda)$

306 294 272 224 139 44 8

2

Calculate the expected profits based on the LASSO model

```
churn_HO[, predLASSOProb := predict(lasso.fit, chHOmat, s= "lambda.min", type="response")]
# Apply the function to calculate expected profits for each threshold
churn_HO[, expProfitLASSO := expProfit(
 DT = churn_HO,
                           # data.table use for the calculation
  score = "predLASSOProb", # Which variable corresponds to the score
  actual = "churned",
                           # Which variable corresponds to the actual outcomes
                           # Value of a true positive
  v11 = 4,
  v10 = -1)]
                           # Value of a false positive.
```

Reformat the data for plotting

ggplot requires that the data be in a specific format for plotting multiple lines. In particular, the variables on the x and y-axes need to each correspond to a single variable name. In this case, x corresponds to shareTest and y corresponds to expProfit. In order to plot multiple lines on the same plot, we require a third variable that indicates how the different plots are different. In this case, I have called this third variable method and it takes values "LASSO" and "Logit."

To put the data in the required shape, I have made two new subsets of the original holdout data. The first reports the x and y values for our "Logit" estimates and the second the values for the "LASSO" estimates. These two new data.tables are then stacked one on top of the other. To accomplish this stacking I place each of the two new data.tables as different objects in the list and then call <code>rbindlist</code> which "binds" the two data.tables by row.

The resulting data.table is called profitDT and can be plot with ggplot.

```
# Stack the data from Logit and LASSO profits
chHOlogitBase = churn_HO[order(-predLogitProbBase), .(shareTest=(1:.N)/.N, expProfit = expProfitLogitBa
chHOlogitGarb = churn_HO[order(-predLogitProbGarb), .(shareTest=(1:.N)/.N, expProfit = expProfitLogitGa
chHOlasso = churn_HO[order(-predLASSOProb), .(shareTest=(1:.N)/.N, expProfit = expProfitLASSO, method="L

profitDT = rbindlist( list(chHOlogitBase,chHOlogitGarb, chHOlasso ))

# Plot the two profit curves on one plot
ggplot(data=profitDT, aes(x=shareTest, y= expProfit, color=method)) +
geom_line() + theme_bw()

0.4
```

