On the estimation of spectral distribution of integrated covariance matrices of high dimensional stochastic processes

Aleksandr Samarin

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Abstract

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We consider the estimation of integrated covariance (ICV) matrices of high dimensional diffusion processes based on high frequency observations. We start by studying the most commonly used estimator, the realized covariance (RCV) matrix. We show that in the high dimensional case when the dimension p and the observation frequency n grow in the same rate, the limiting spectral distribution (LSD) of RCV depends on the covolatility process not only through the targeting ICV, but also on how the covolatility process varies in time. We establish a Marcenko–Pastur type theorem for weighted sample covariance matrices, based on which we obtain a Marcenko– Pastur type theorem for RCV for a class C of diffusion processes. The results explicitly demonstrate how the time variability of the covolatility process affects the LSD of RCV. We further propose an alternative estimator, the time-variation adjusted realized covariance (TVARCV) matrix. We show that for processes in class C, the TVARCV possesses the desirable property that its LSD depends solely on that of the targeting ICV through the Marcenko–Pastur equation, and hence, in particular, the TVARCV can be used to recover the empirical spectral distribution of the ICV by using existing algorithms

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Part 1

Introduction

Motivation

Remark 1.1. Diffusion processes are widely used to model financial asset price processes. For example, suppose that we have multiple stocks, say, p stocks, whose price processes are denoted by $S_j(t)$ and $X_j(t) := \log S_j(t)$ for $j = 1, \ldots, p$. Let

$$\mathbf{X}(t) = \big(X_1(t), \ldots, X_p(t)\big)^T.$$

Then a widely used model for $\mathbf{X}_t$ is

$$d\mathbf{X}(t) = \boldsymbol{\mu}(t)dt + \Theta(t)d\mathbf{W}(t), \tag{1}$$

where

- $\boldsymbol{\mu}(t) = (\mu_1(t), \ldots, \mu_p(t))^T$ is a p-dimensional drift process,

- $\Theta(t) - p \times p$ **_covolatility process_**,

- $\mathbf{W}(t) - p$-dimensional standard Brownian motion.

If $\Theta(t) = \sigma(t) \cdot \mathbb{I}_{p \times p}$, we are in the case when all price processes are independent of each other and we can study them separately. In practice price on one good strongly depends on the price of the other. In this paper we consider special class of possibly dependent processes, which interrelation stays stable in time.

Definition 1.2.

(i) The **_integrated covariance matrix (ICV)_** is defined as follows

$$\Sigma_p := \int_0^1 \Theta(t) \cdot \Theta^T(t)dt.$$

For $p = 1$ ICV is called **_integrated volatility_**.

(ii) Assume that we can observe the processes $X_j(t)$ at high frequency synchronously, say at time points $\{\tau_\ell\}_{0 \leq \ell \leq n}$:

$$X_j(\tau_\ell) = \log S_j(\tau_\ell), \quad \ell = 0, 1, \ldots, n, j = 1, \ldots, p.$$

Then **_realized covariance matrix (RCV)_** is

$$\Sigma_p^{RCV} := \sum_{\ell=1}^n \Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T, \tag{2}$$

where

$$\Delta\mathbf{X}_\ell := \begin{pmatrix} \Delta X_{1\ell} \\ \vdots \\ \Delta X_{p\ell} \end{pmatrix} = \begin{pmatrix} X_1(\tau_\ell) - X_1(\tau_{\ell-1}) \\ \vdots \\ X_p(\tau_\ell) - X_p(\tau_{\ell-1}) \end{pmatrix}.$$

For $p = 1$ RCV is called **realized volatility**.

Example 1.3. Let's take $p = 3$, zero drift $\boldsymbol{\mu}(t) = (0,0,0)^T$ and

$$\Theta(t) = \begin{pmatrix} 1 & \cos(\pi t) & \sin(\pi t) \\ \cos(\pi t) & 1 & 0 \\ \sin(\pi t) & 0 & 1 \end{pmatrix}.$$

Then our process satisfies

$$d\mathbf{X}(t) = \begin{pmatrix} dW_1(t) + \cos(\pi t)dW_2(t) + \sin(\pi t)dW_3(t) \\ \cos(\pi t)dW_1(t) + dW_2(t) \\ \sin(\pi t)dW_1(t) + dW_3(t) \end{pmatrix}.$$

We have

$$\Theta(t) \cdot \Theta^T(t) = \begin{pmatrix} 2 & 2\cos(\pi t) & 2\sin(\pi t) \\ 2\cos(\pi t) & \cos^2(\pi t) + 1 & \cos(\pi t)\sin(\pi t) \\ 2\sin(\pi t) & \cos(\pi t)\sin(\pi t) & \sin^2(\pi t) + 1 \end{pmatrix}$$

and corresponding ICV matrix is

$$\Sigma_3 = \begin{pmatrix} 2 & 0 & \frac{4}{\pi} \\ 0 & \frac{3}{2} & 0 \\ \frac{4}{\pi} & 0 & \frac{3}{2} \end{pmatrix}.$$

Remark 1.4.

(i) Intuitively, one can wish zero elements of the ICV matrix to show the independence of one process from another. But in Example 1.3, where all the processes are related with each other in some sense, we still have zero values in Σ_3. It should be noted that if $\mathbf{X}(t)$ is a vector of log-price processes, then over short time period, say, one day, it is reasonable to assume that the correlation structure doesn't change.

(ii) One of the major problems in most of the models, based on the theory of stochastic calculus and applied to financial markets, is the estimation of true covariance. In our case the final goal is to find an estimator of an ICV matrix. In the next section we show how RCV matrix does the job.

Figure 1: $\mathbf{X}(t)$ from Example 1.3

One can observe that $X_1(t)$ and $X_2(t)$ have strong positive correlation in the beginning, which decreases over time and later, when t approaches 1, turns into negative one. Meanwhile $X_3(t)$ has independent fluctuations on the borders, but in the middle it repeats the behavior of $X_1(t)$.

The relation between RCV and ICV

Remark 1.5. Recall the basic notions from the theory of stochastic calculus (w.l.o.g. we assume that all the processes have a starting point at 0 a.s.):

(i) An **Itô process** $X(t)$ is defined to be an adapted stochastic process that can be expressed as the sum of an integral with respect to time and an integral with respect to Brownian motion:

$$X(t) = \int_0^t \mu(s)ds + \int_0^t \sigma(s)dW(s).$$

(ii) Let X be a continuous local martingale. Then $[X]_t$ with

$$[X](t) = X^2(t) - 2\int_0^t X(s)dX(s)$$

is called **quadratic variation** of X. Note that for deterministic processes quadratic variation is equal to 0.

(iii) The **quadratic covariation** of two continuous semimartingales X and Y is defined as follows:

$$[X, Y](t) = \frac{1}{4}([X + Y](t) - [X - Y](t)).$$

(iv) Let X and Y be Itô processes with volatilities $\sigma_X(t)$ and $\sigma_Y(t)$ with respect to the same Brownian motion, then

$$[X, Y](t) = \int_0^t \sigma_X(s)\sigma_Y(s)ds.$$

(v) Quadratic covariation is symmetric and bilinear map:

$$[X + Y, Z](t) = [X, Z](t) + [Y, Z](t),$$

$$[X, Y + Z](t) = [X, Y](t) + [X, Z](t),$$

$$[X, Y](t) = [Y, X](t), \quad [X](t) = [X, X](t).$$

Example 1.6.

(i) Let us first consider two-dimensional case with $\boldsymbol{\mu}(t) = (\mu_1(t), \mu_2(t))^T$ and

$$\Theta(t) = \begin{pmatrix} \sigma_{11}(t) & \sigma_{12}(t) \\ \sigma_{21}(t) & \sigma_{22}(t) \end{pmatrix}$$

Then $\mathbf{X}(t) = (X(t), Y(t))^T$, where

$$X(t) = \int_0^t \mu_1(s)ds + \int_0^t \sigma_{11}(s)dW_1(s) + \int_0^t \sigma_{12}(s)dW_2(s)$$

and

$$Y(t) = \int_0^t \mu_2(s)ds + \int_0^t \sigma_{21}(s)dW_1(s) + \int_0^t \sigma_{22}(s)dW_2(s).$$

Using Remark 1.5 we get

$$[X](t) = \int_0^t (\sigma_{11}^2(s) + \sigma_{12}^2(s))ds, \quad [Y](t) = \int_0^t (\sigma_{21}^2(s) + \sigma_{22}^2(s))ds \qquad (3)$$

and

$$[X, Y](t) = \int_0^t (\sigma_{11}(s)\sigma_{21}(s) + \sigma_{12}(s)\sigma_{22}(s))ds. \qquad (4)$$

Let us look at the quadratic covariation $[X, Y]$ at time 1. It is defined as follows: let $\Pi = \{\tau_0, \ldots, \tau_n\}$ be a partition of $[0, 1]$ (i.e. $0 = \tau_0 < \tau_1 < \ldots \tau_n = 1$) and set up the sampled cross variation

$$\sum_{\ell=1}^n \Delta X_\ell \Delta Y_\ell = \sum_{\ell=1}^n (X(\tau_\ell) - X(\tau_{\ell-1})) \cdot (Y(\tau_\ell) - Y(\tau_{\ell-1})). \qquad (5)$$

Now let the number of partition points n go to infinity as the length of the longest subinterval $\|\Pi\| = \max_{1 \le \ell \le n}(\tau_\ell - \tau_{\ell-1})$ goes to zero. The sum in (5) converges in probability to $[X, Y](1)$. This limit is given by the right-hand side of (4). This assertion follows from the standard theorems in the theory of stochastic calculus. Writing formally,

$$\lim_{n \to \infty} \sum_{\ell=1}^n \Delta X_\ell \Delta Y_\ell \xrightarrow{\mathbb{P}} [X, Y](1).$$

Hence,

$$\lim_{n\to\infty}\sum_{\ell=1}^{n}\Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T = \lim_{n\to\infty}\sum_{\ell=1}^{n}\begin{pmatrix}(\Delta X_\ell)^2 & \Delta X_\ell\Delta Y_\ell \\ \Delta Y_\ell\Delta X_\ell & (\Delta Y_\ell)^2\end{pmatrix} \xrightarrow{\mathbb{P}} \begin{pmatrix}[X](1) & [X,Y](1) \\ [Y,X](1) & [Y](1)\end{pmatrix}.$$

On the other hand,

$$\Theta(t)\cdot\Theta^T(t) = \begin{pmatrix}\sigma_{11}^2(t)+\sigma_{12}^2(t) & \sigma_{11}(t)\sigma_{21}(t)+\sigma_{12}(t)\sigma_{22}(t) \\ \sigma_{11}(t)\sigma_{21}(t)+\sigma_{12}(t)\sigma_{22}(t) & \sigma_{21}^2(t)+\sigma_{22}^2(t)\end{pmatrix}.$$

Taking the integral of $\Theta(t)\cdot\Theta^T(t)$ we conclude that

$$\lim_{n\to\infty}\Sigma_2^{RCV} \xrightarrow{\mathbb{P}} \Sigma_2.$$

(ii) The convergence of realized covariance to ICV can be generalized for arbitrary p. Let's set $\boldsymbol{\mu}(t)=(\mu_1(t),\ldots,\mu_p(t))^T$ and

$$\Theta(t) = \begin{pmatrix}\Theta_{11}(t) & \cdots & \Theta_{1p}(t) \\ \vdots & \ddots & \vdots \\ \Theta_{p1}(t) & \cdots & \Theta_{pp}(t)\end{pmatrix}.$$

Then we have $\mathbf{X}(t)=(X_1(t),\ldots,X_p(t))^T$ with

$$X_j(t) = \int_0^t \mu_j(s)ds + \sum_{i=1}^{p}\int_0^t \Theta_{ji}(s)dW_i(s)$$

for all $1\le j\le p$. By Remark 1.5 again

$$[X_j](t) = \int_0^t\left(\sum_{i=1}^{p}\Theta_{ji}^2(s)\right)ds$$

and

$$[X_j,X_k](t) = \int_0^t\left(\sum_{i=1}^{p}\Theta_{ji}(s)\cdot\Theta_{ki}(s)\right)ds$$

for all $1\le j,k\le p$. We also have

$$\left(\Theta(t)\cdot\Theta^T(t)\right)_{jk} = \sum_{i=1}^{p}\Theta_{ji}(t)\cdot\Theta_{ki}(t).$$

Therefore, we get

$$\lim_{n\to\infty}\sum_{\ell=1}^{n}\Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T \xrightarrow{\mathbb{P}} [\mathbf{X},\mathbf{X}](1),$$

where

$$\left([\mathbf{X},\mathbf{X}](t)\right)_{ij} = [X_i,X_j](t),$$

and finally

$$\lim_{n\to\infty}\Sigma_p^{RCV} \xrightarrow{\mathbb{P}} \Sigma_p.$$

Remark 1.7. Let ICV matrix Σ_p be deterministic. RCV matrix Σ_p^{RCV}, on the other side, is random by definition. In Example 1.6 we used quadratic covariation as an intermediate step to prove the convergence of one matrix to another. Alternatively to Remark 1.5 (iii) quadratic covariation can be defined as a limit of the sum of squared differences, so it is also random. Having almost sure degenerate distribution of quadratic covariation is one of the major properties of Itô processes and it doesn't need to hold for all martingales in general.

Part 2
Marchenko-Pastur law

Empirical spectral distribution

Remark 2.1. In Remark 1.6 we claimed the convergence of RCV matrix to ICV for any finite p. In practice, however, it might be the case that the number of processes is approximately equal to the number of observations for each process. Theoretically, the problem arises when the dimension of p has at least the same rate of growth as n. In this case the convergence is not well defined, because the RCV matrix has the unlimited size. Then, it is more convenient to study another criteria.

Definition 2.2. Let $\{\lambda_j : j = 1, \ldots, p\}$ be set of eigenvalues of matrix A, then

$$F^A(x) := \frac{\#\{j : \lambda_j \leq x\}}{p}, \quad x \in \mathbb{R},$$

is called **_empirical spectral distribution (ESD)_** of A.

Example 2.3. Take the simplest case with $\boldsymbol{\mu}(t) = 0$ and $\Theta(t) = \sigma \cdot \mathbb{I}_{p \times p}$ with some constant σ. Then $\Sigma_p = \sigma^2 \mathbb{I}_{p \times p}$ and its ESD

$$F^{\Sigma_p}(x) = 1_{[\sigma^2, \infty)}(x)$$

is a simple step function. In general case, where Σ_p have different eigenvalues, ESD is a combination of such step functions.

Remark 2.4. A naive estimator of the spectral distribution of the ICV matrix Σ_p is the spectral distribution of the RCV matrix Σ_p^{RCV}. In particular, one might wish that the ESD $F^{\Sigma_p^{RCV}}$ of Σ_p^{RCV} would approximate F^{Σ_p} well when the value of n sufficiently high, namely

$$\lim_{n \to \infty} F^{\Sigma_p^{RCV}} \to F^{\Sigma_p}.$$

From the large dimensional random matrix theory (LDRMT), we now understand quite well that in the high dimensional setting this good wish won't always come true. For example, in the simplest case when the drift process $\boldsymbol{\mu}(t)$ is 0, covolatility process $\Theta(t)$ is constant, and observation times τ_ℓ are equally spaced ($\tau_\ell = \ell/n$), we are in the setting of estimating the usual covariance matrix using the sample covariance matrix, given n i.i.d. p-dimensional observations $(\Delta \mathbf{X}_\ell)_{\ell=1,\ldots,n}$. From LDRMT, we know that if p/n converges to a non-zero number and the ESD F^{Σ_p} of the true covariance matrix converges, then the ESD $F^{\Sigma_p^{RCV}}$ of the sample covariance matrix also converges (see e.g. Marchenko and Pastur (1967), Yin(1968), Silverstein and Bai (1995) and Silverstein(1995)). However, the relationship between the limiting spectral distribution (LSD) of Σ_p^{RCV} in this case and the LSD of Σ_p still can be described, namely by a Marchenko-Pastur equation through Stieltjes transforms.

Definition 2.5. In mathematics, the ***Stieltjes transformation*** $m_\mu(z)$ of a measure μ on $\mathbb{R}$ is the function of the complex variable $z \in \mathbb{C}_+ := \{z \in \mathbb{C} : \Im(z) > 0\}$ defined by the formula

$$m_\mu(z) = \int_{t \in \mathbb{R}} \frac{1}{t - z} d\mu(t).$$

The inversion formula is given by

$$\mu([a, b]) + \frac{\mu(\{a\}) + \mu(\{b\})}{2} = \frac{1}{\pi} \lim_{\eta \to 0^+} \int_a^b \Im(m_\mu(z + i\eta)) dz. \tag{6}$$

Proposition 2.6.

(i) for $p = 1, 2, \ldots$ and for $1 \le \ell \le n$, $\mathbf{Z}_\ell^{(p)} = \left(Z_{j\ell}^{(p)}\right)_{1 \le j \le p}$ is a p-dimensional vector with $Z_{j\ell}^{(p)}$ i.i.d. with mean 0 and variance 1;

(ii) $n = n(p)$ with $y_n := p/n \to y > 0$ as $p \to \infty$;

(iii) Σ_p is a (possibly random) nonnegative definite $p \times p$ matrix such that its ESD F^{Σ_p} converges a.s. in distribution to a probability distribution H on $[0, \infty)$ as $p \to \infty$;

(iv) Σ_p and $\mathbf{Z}_\ell^{(p)}$ are independent.

Let $\Sigma_p^{1/2}$ be the (nonnegative) square root matrix of Σ_p and

$$S_p := \frac{1}{n} \sum_{\ell=1}^n \Sigma_p^{1/2} \mathbf{Z}_\ell^{(p)} (\mathbf{Z}_\ell^{(p)})^T \Sigma_p^{1/2}.$$

Then a.s. the ESD of S_p converges in distribution to a probability distribution F, which is determined by H in that its Stieltjes transform

$$m_F(z) := \int_{\lambda \in \mathbb{R}} \frac{1}{\lambda - z} dF(\lambda), \quad z \in \mathbb{C}_+$$

solves the equation

$$m_F(z) = \int_{\tau \in \mathbb{R}} \frac{1}{\tau(1 - y(1 + zm_F(z))) - z} dH(\tau). \tag{7}$$

Proof. See Theorem 1.1 of Silverstein (1995). □

Remark 2.7.

(i) It is easy to see that for ESD F its Stieltjes transform can be found by formula

$$m_F(z) = \frac{1}{p} \sum_{j=1}^p \frac{1}{\lambda_j - z} = \frac{1}{p} \operatorname{tr}(A - z\mathbb{I}_{p \times p})^{-1}. \tag{8}$$

(ii) Schematically Proposition 2.6 can be depicted as follows

$$(F^{\Sigma_p^{RCV}}, F^{\Sigma_p}) \xrightarrow[p,n\to\infty]{} (F, H).$$

Our goal is to find F^{Σ_p} for sufficiently large p. First, we approximate F with $F^{\Sigma_p^{RCV}}$ and calculate $m_F(z)$, using, for example, (8). Then we get H, via Marchenko-Pastur equation (7), and finally approximate F^{Σ_p} with H.

(iii) Note that if $y \to 0$ limiting distribution function F of S_p matches limiting distribution function H of Σ_p. This is the case, when n grows faster than p and as we stated before Σ_p^{RCV} converges to Σ_p.

(iv) In the special case when $\Sigma_p = \sigma^2 \mathbb{I}_{p\times p}$, where $\mathbb{I}_{p\times p}$ is the $p \times p$ identity matrix, the LSD F has an analytical expression, which can be derived from the following proposition.

Proposition 2.8. Suppose that $\mathbf{Z}_\ell^{(p)}$'s are as in the Proposition 2.6, and $\Sigma_p = \sigma^2 \mathbb{I}_{p\times p}$ for some $\sigma^2 > 0$. Then the LSD F has density

$$f(x) = \left(1 - \frac{1}{y}\right)_+ \delta_0(x) + \frac{1}{2\pi\sigma^2 xy}\sqrt{(b-x)(x-a)}\mathbf{1}_{[a,b]}(x),$$

where $\delta_0(x)$ is a Dirac delta function and

$$a = \sigma^2(1 - \sqrt{y})^2 \quad \text{and} \quad b = \sigma^2(1 + \sqrt{y})^2.$$

The LSD F in this proposition is called the Marchenko-Pastur law with ratio index y and scale index σ^2, and will be denoted by $\mathcal{MP}(y, \sigma^2)$.

Proof. See, e.g., Theorem 2.5 in Bai (1999). $\square$

Remark 2.9.

(i) We have for $\Sigma_p = \sigma^2 \mathbb{I}_{p\times p}$ the density of limiting spectral distribution

$$dH(x) = \delta_{\sigma^2}(x).$$

Define $m_{\mathcal{MP}}(z)$ as the Stieltjes transform of the Marchenko-Pastur distribution. Then it satisfies the self-consistent equation

$$
\begin{aligned}
m_{\mathcal{MP}}(z) &= \int_{\tau \in \mathbb{R}} \frac{1}{\tau(1 - y - zm_{\mathcal{MP}}(z)) - z}\delta_{\sigma^2}(x)dx \\
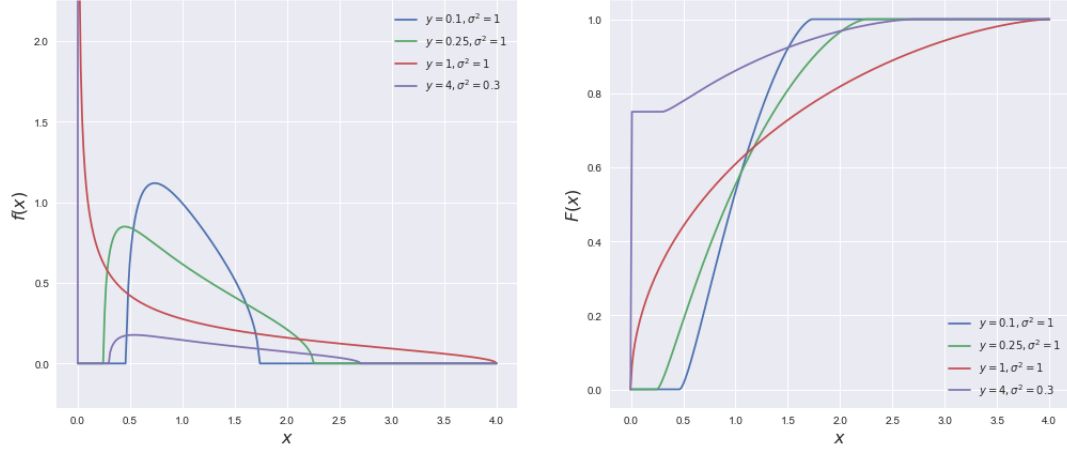&= \frac{1}{\sigma^2(1 - y - zm_{\mathcal{MP}}(z)) - z}.
\end{aligned}
$$

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Figure 2: Marchenko-Pastur law

The left panel shows graphs of the density function $f(x)$ for different parameters y and σ^2. On the right there are graphs of corresponding cumulative distribution functions $F(x)$. Notice that as y becomes larger the eigenvalues of Σ_p spread out instead of concentrating near σ^2 and the length of the density's support increases.

(ii) Note again that for $y \to 0$ the density in Proposition 2.8 is

$$f(x) = \delta_{\sigma^2}(x),$$

which coincides with the density in Example 2.3 for finite value of p and with density $dH(x)$, because in this case Σ_p^{RCV} converges to Σ_p.

(iii) One can show that

$$m_{\mathcal{MP}}(z) = \frac{\sigma^2(1-y) - z + i\sqrt{(b-z)(z-a)}}{2yz}.$$

Therefore, one can find $m_{\mathcal{MP}}(z)$ through (8) and derive σ^2 from it easily.

Proof. See, e.g., Lemma 2.18 in Ben Adlam $\qquad \square$

Non-constant volatility case

Remark 2.10. In practice, the covolatility process is typically not constant. Furthermore, stock volatility might not only vary in time, but also be random and depend on the price process itself. In general, for any time-varying covolatility process $\Theta(t)$, we associate it with a constant covolatility process given by the square root of the ICV matrix

$$\Theta^0(t) = \sqrt{\int_0^1 \Theta(s) \cdot \Theta^T(s) ds} \quad \forall t \in [0,1].$$

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Let $\mathbf{X}^0(t)$ be such that

$$d\mathbf{X}^0(t) = \Theta^0(t)d\mathbf{W}(t).$$

Note that $\mathbf{X}(t)$ and $\mathbf{X}^0(t)$ share the same ICV matrix at time 1:

$$\Sigma_p^0 := \int_0^1 \Theta^0(t) \cdot \Theta^{0T}(t)dt = \int_0^1 dt \int_0^1 \Big(\Theta(s) \cdot \Theta^T(s)\Big)ds = \Sigma_p.$$

Based on $\mathbf{X}^0(t)$, we have an associated RCV matrix

$$\Sigma_p^{RCV^0} = \sum_{\ell=1}^n \Delta\mathbf{X}_\ell^0(\Delta\mathbf{X}_\ell^0)^T.$$

Since Σ_p^{RCV} and $\Sigma_p^{RCV^0}$ are based on the same estimation method and share the same targeting ICV matrix, it is desirable that their ESDs have similar properties. In particular, based on the results in LDRMT and the discussion about constant covolatility case, we have the following property for $\Sigma_p^{RCV^0}$: if the ESD F^{Σ_p} converges, then so does $F^{\Sigma_p^{RCV^0}}$. Furthermore, their limits are related to each other via the Marchenko-Pastur equation (7). The question is: does this property hold for ESD of Σ_p^{RCV}? Later we will see that even in the most ideal case, when the covolatility process has the form $\Theta(t) = \gamma(t)\mathbb{I}_{p \times p}$ for some deterministic scalar function $\gamma(t)$, such convergence results may *not* hold for Σ_p^{RCV}. Moreover, the limit of $F^{\Sigma_p^{RCV}}$ (when it exists) changes according to how the covolatility process evolves over time.

Definition 2.11. Let A, B be two symmetric $p \times p$ matrices. Then we write

$$A \succeq B$$

to denote that $A - B$ is a positive semidefinite matrix. Similarly, we say that $A \preceq B$ if $A - B$ is a negative semidefinite. The partial ordering defined by $\succeq$ is called Löwner ordering.

Lemma 2.12 (Weyl's Monotonicity Theorem)**.** *Suppose A and B are symmetric, $p \times p$ matrices. Let $\lambda_i(A)$ be the i-th largest eigenvalue of A. If $A \preceq B$, then $\lambda_i(A) \leq \lambda_i(B)$ for all i, or, equivalently*

$$F^B(x) \leq F^A(x) \quad \forall x \geq 0.$$

Proof. Corollary 4.3.3 in Horn and Johnson (1990). $\qquad\square$

Proposition 2.13. Suppose that for any p, $\mathbf{X}(t)$ is a p-dimensional process satisfying

$$d\mathbf{X}(t) = \gamma(t)d\mathbf{W}(t), \quad t \in [0,1], \tag{9}$$

where $\gamma(t) > 0$ is a nonrandom (scalar) càdlàg process. Let $\sigma^2 = \int_0^1 \gamma^2(t)dt$ and so that the ICV matrix $\Sigma_p = \sigma^2\mathbb{I}_{p \times p}$. Assume further that the observation times τ_ℓ are equally spaced, that is, $\tau_\ell = \ell/n$, and that the RCV matrix Σ_p^{RCV} is defined by (2). Then so

long as $\gamma(t)$ is not constant on $[0, 1)$, for any $\varepsilon > 0$, there exists $y_c = y_c(\gamma, \varepsilon) > 0$ such that if $\lim p/n = y \geq y_c$,

$$\limsup F^{\Sigma_p^{RCV}}(b(y) + \sigma^2 \varepsilon) < 1 \quad \text{a.s.} \tag{10}$$

In particular, $F^{\Sigma_p^{RCV}}$ doesn't converge to the Marchenko-Pastur law $\mathcal{MP}(y, \sigma^2)$.

Proof. By assumption if $\gamma(t)$ is non-contant, there exists $\delta > 0$ and an interval $[c, d] \subseteq [0, 1]$ such that

$$\gamma(t) \geq \sigma(1 + \delta) \quad \forall t \in [c, d].$$

Therefore, if $\left[\frac{\ell-1}{n}, \frac{\ell}{n}\right] \subseteq [c, d]$ for some $1 \leq \ell \leq n$, then

$$\Delta \mathbf{X}_\ell (\Delta \mathbf{X}_\ell)^T \stackrel{d}{=} \int_{(\ell-1)/n}^{\ell/n} \gamma^2(t) dt \cdot \mathbf{Z}_\ell (\mathbf{Z}_\ell)^T \succeq \frac{(1+\delta)^2}{n} \sigma^2 \mathbf{Z}_\ell (\mathbf{Z}_\ell)^T,$$

where $\mathbf{Z}_\ell = (Z_{1\ell}, \dots, Z_{p\ell})^T$ consists of independent standard normals. Hence, if we let $J_n = \left\{\ell : \left[\frac{\ell-1}{n}, \frac{\ell}{n}\right] \subseteq [c, d]\right\}$ and

$$\Gamma_p = \sum_{\ell \in J_n} \Delta \mathbf{X}_\ell (\Delta \mathbf{X}_\ell)^T, \quad \Lambda_p = \frac{\sigma^2}{n(d-c)} \sum_{\ell \in J_n} \mathbf{Z}_\ell (\mathbf{Z}_\ell)^T,$$

then for any $x \geq 0$, by Weyl's Monotonicity Theorem,

$$F^{\Sigma_p^{RCV}}(x) \leq F^{\Gamma_p}(x) \leq F^{\Lambda_p}\left(\frac{x}{(1+\delta)^2(d-c)}\right).$$

Now note that $\#J_n \sim (d-c)n$, hence if $p/n \to y$, by Proposition 2.6, F^{Λ_p} will converge a.s. to the Marchenko-Pastur law with ratio index $y' = \frac{y}{d-c}$ and scale index σ^2. By the formula of $b(\cdot)$ in Marchenko-Pastur density

$$(1+\delta)^2(d-c)b(y') = (1+\delta)\sigma^2 \cdot (1+\delta)(d-c)(1 + 2\sqrt{y'} + y')$$
$$= (1+\delta)\sigma^2 \cdot (1+\delta)(d-c+2\sqrt{(d-c)y}+y)$$
$$:= (1+\delta)\sigma^2 \cdot g(y).$$

Note that the $g(y)$ has a linear growth in y with coefficient $1 + \delta$. Hence, for any $\varepsilon > 0$, there exists $y_c > 0$, such that for all $y \geq y_c$

$$g(y) \geq (1 + \sqrt{y})^2 + \varepsilon,$$

that is,

$$(1+\delta)^2(d-c)b(y') \geq (1+\delta)\sigma^2 \cdot ((1+\sqrt{y})^2 + \varepsilon) = (1+\delta)(b(y) + \sigma^2 \varepsilon)$$

or, equivalently,

$$\frac{b(y) + \sigma^2 \varepsilon}{(1+\delta)^2(d-c)} \leq \frac{b(y')}{1+\delta}.$$

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Therefore, when the above inequality holds,

$$\limsup F^{\Sigma_p^{RCV}}(b(y) + \sigma^2 \varepsilon) \le \limsup F^{\Lambda_p}\left(\frac{b(y')}{1+\delta}\right) < 1.$$

$\square$

Remark 2.14. In the proof of Proposition 2.13 we saw that more volatility process vary in time, the bigger value of δ can be found, and hence more the maximum value of domain of $F^{\Sigma_p^{RCV}}$ goes away from $b(y)$. Therefore ESD of Σ_p^{RCV} cannot be considered as consistent estimator anymore. In the next part we will study different estimator, which is based on the whole behavior of volatility process.

Part 3
Limit theorems for non-constant covolatility processes

Limiting distribution for realized covariance

Remark 3.1. In this part we will state two major theorems. The first one is a generalization of Proposition 2.6 for limiting spectral distribution of RCV matrix in the case of non-constant covolatility. The second theorem claims the convergence of ESD of alternative estimator to Marchenko-Pastur law. Both theorems require some additional assumptions.

Definition 3.2. Suppose that for all p $\mathbf{X}(t) = \mathbf{X}^{(p)}(t)$ is a p-dimensional process satisfying (1), and $\Theta(t)$ is càdlàg. We say that $\mathbf{X}(t)$ belongs to **class** $\mathcal{C}$ if, almost surely, there exist $\gamma(t) : [0,1] \mapsto \mathbb{R}$ and Λ a $p \times p$ matrix satisfying $\mathrm{tr}(\Lambda\Lambda^T) = p$ such that

$$\Theta(t) = \gamma(t)\Lambda. \tag{11}$$

Observe that if (11) holds, then the ICV matrix $\Sigma_p = \int_0^1 \gamma^2(t)dt \cdot \Lambda\Lambda^T$. The special case when $\Lambda = \mathbb{I}_{p \times p}$ will be studied in simulations.

Example 3.3. Suppose that $X_j(t)$ satisfy

$$dX_j(t) = \mu_j(t)dt + \sigma_j(t)dW_j(t), \quad j = 1, \ldots, p,$$

where $\mu_j(t), \sigma_j(t) : [0,1] \to \mathbb{R}$ are the drift and volatility processes for stock j, and $W_j(t)$'s are (one-dimensional) standard Brownian motions. If the following conditions hold:

- the correlation matrix process of $(W_j(t))$

$$R(t) := \left(\frac{[W_j, W_k](t)}{t} \right)_{1 \leq j,k \leq p} =: (r_{jk})_{1 \leq j,k \leq p}$$

  is constant in $t \in [0,1]$.

- $r^{(jk)} \neq 0$ for all $1 \leq j, k \leq p$; and

- the correlation matrix process of $X_j(t)$

$$\left(\frac{\int_0^t \sigma_j(s)\sigma_k(s)d[W_j, W_k](s)}{\sqrt{\int_0^t \sigma_j^2(s)ds \cdot \int_0^t \sigma_k^2(s)ds}} \right)_{1 \leq j,k \leq p} =: (\rho_{jk})_{1 \leq j,k \leq p}$$

  is constant in $t \in [0,1]$;

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then $\mathbf{X}(t)$ belongs to class $\mathcal{C}$.

Proof. For any $t \in [0,1]$:

$$\rho_{jk} = \frac{r_{jk} \int_0^t \sigma_j(s)\sigma_k(s)ds}{\sqrt{\int_0^t \sigma_j^2(s)ds \cdot \int_0^t \sigma_k^2(s)ds}},$$

therefore

$$\frac{\rho_{jk}}{r_{jk}} = \frac{\int_0^t \sigma_j(s)\sigma_k(s)ds/t}{\sqrt{\int_0^t \sigma_j^2(s)ds/t \cdot \int_0^t \sigma_k^2(s)ds/t}}.$$

Letting $t \downarrow 0$, using l'Hôpital's rule and noting that $\sigma_j(t)$ are càdlàg, we observe that

$$\frac{\rho_{jk}}{r_{jk}} = \frac{\sigma_j(0)\sigma_k(0)}{\sqrt{\sigma_j^2(0) \cdot \sigma_k^2(0)}} = \pm 1.$$

Hence, for all $t \in [0,1]$:

$$\left| \int_0^t \sigma_j(s)\sigma_k(s)ds \right| = \sqrt{\int_0^t \sigma_j^2(s)ds \cdot \int_0^t \sigma_k^2(s)ds}.$$

By Cauchy-Schwartz inequality, this holds only if $\sigma_j(s)$ and $\sigma_k(s)$ are proportional to each other. Therefore, almost surely, there exists a scalar process $\gamma(t) : [0,1] \to \mathbb{R}$ and a p-dimensional vector $(\sigma_1, \ldots, \sigma_p)^T$, such that

$$(\sigma_1(t), \ldots, \sigma_p(t))^T = \gamma(t) \cdot (\sigma_1, \ldots, \sigma_p)^T.$$

Now we show that $\mathbf{X}(t)$ belongs to class $\mathcal{C}$. In fact one can always find a p-dimensional standard Brownian motion $\widetilde{\mathbf{W}}(t)$, such that

$$\mathbf{W}(t) = R^{1/2}\widetilde{\mathbf{W}}(t),$$

where $\mathbf{W} = (W_1(t), \ldots, W_p(t))^T$ and R is a correlation matrix of $\mathbf{W}(t)$, which is constant for all $t \in [0,1]$ by assumption. Hence, writing $\mu(t) = (\mu_1(t), \ldots, \mu_p(t))^T$, we have

$$d\mathbf{X}(t) = \mu(t)dt + \mathrm{diag}(\sigma_1(t), \ldots, \sigma_p(t))d\mathbf{W}(t)$$
$$= \mu(t)dt + \gamma(t) \cdot \mathrm{diag}(\sigma_1, \ldots, \sigma_p)R^{1/2}d\widetilde{\mathbf{W}}(t).$$

$\square$

Remark 3.4.

(i) In particular, if all Brownian motions in Example 3.3 are independent of each other, then $R(t)$ is an identity matrix $\mathbb{I}_{p \times p}$. Then correlation matrix process of $X_j(t)$ is a diagonal matrix and we need only to require ρ_{jj} to be constant in t.

(ii) The process in Example 1.3 doesn't belong to the class $\mathcal{C}$, because its correlation structure varies in time.

(iii) Observe that if a diffusion process $\mathbf{X}(t)$ belongs to class $\mathcal{C}$, the drift process $\boldsymbol{\mu}(t) = 0$, and τ_ℓ's and $\gamma(t)$ are independent of $\mathbf{W}(t)$, then

$$\Delta \mathbf{X}_\ell = \int_{\tau_{\ell-1}}^{\tau_\ell} \gamma(t) \Lambda d\mathbf{W}(t) \stackrel{d}{=} \sqrt{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma^2(t) dt} \cdot \overline{\Sigma}^{1/2} \cdot \mathbf{Z}_\ell,$$

where $\overline{\Sigma}^{1/2}$ is the nonnegative square root matrix of

$$\overline{\Sigma} := \Lambda\Lambda^T$$

and $\mathbf{Z}_\ell = (Z_{1\ell}, \ldots, Z_{p\ell})^T$ consists of independent standard normals. Therefore, the RCV matrix

$$\Sigma_p^{RCV} = \sum_{\ell=1}^n \Delta \mathbf{X}_\ell (\Delta \mathbf{X}_\ell)^T \stackrel{d}{=} \sum_{\ell=1}^n w_\ell^n \cdot \overline{\Sigma}^{1/2} \mathbf{Z}_\ell (\mathbf{Z}_\ell)^T \overline{\Sigma}^{1/2}$$

with

$$w_\ell^n = \int_{\tau_{\ell-1}}^{\tau_\ell} \gamma^2(t) dt.$$

This is similar to the S_p in Proposition 2.6, except that here the weights w_ℓ^n may vary in ℓ, while in Proposition 2.6 the weights are constantly $1/n$.

Theorem 3.5. *Assume that all the conditions in Proposition 2.6 are satisfied. Furthermore,*

(i) $Z_{j\ell}^{(p)}$ have finite moments of all orders;

(ii) H has a finite second moment;

(iii) the weights w_ℓ^n, $1 \le \ell \le n$, $n = 1, 2, \ldots$, are all positive, and there exists $\kappa < \infty$ such that the rescaled weights (nw_ℓ^n) satisfy

$$\max_n \max_{\ell=1,\ldots,n} (nw_\ell^n) \le \kappa;$$

moreover, almost surely, there exists a càdlàg function $w(s) : [0,1] \to \mathbb{R}_+$, such that

$$\lim_n \sum_{1 \le \ell \le n} \int_{(\ell-1)/n}^{\ell/n} |nw_\ell^n - w(s)| ds = 0;$$

(iv) there exists a sequence $\eta_p = o(p)$ and a sequence of index sets $\mathcal{I}_p$ satisfying $\mathcal{I}_p \subset \{1, \ldots, p\}$ and $\#\mathcal{I}_p \le \eta_p$ such that for all n and all ℓ, w_ℓ^n may depend on $\mathbf{Z}_\ell^{(p)}$ but only on $\{Z_{j\ell}^{(p)} : j \in \mathcal{I}_p\}$;

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(v) there exist $C < \infty$ and $\delta < 1/6$ such that for all p, $\|\Sigma_p\| \le Cp^{\delta}$ a.s.

Define $S_p = \sum_{\ell=1}^{n} w_{\ell}^n \cdot \Sigma_p^{1/2} \mathbf{Z}_{\ell}^{(p)} (\mathbf{Z}_{\ell}^{(p)})^T \Sigma_p^{1/2}$. Then, almost surely, the ESD of S_p converges in distribution to a probability distribution F^w, which is determined by H and $w(s)$ in that its Stiltjes trasform $m_{F^w}(z)$ is given by

$$m_{F^w}(z) = -\frac{1}{z} \int_{\tau \in \mathbb{R}} \frac{1}{\tau M(z) + 1} dH(\tau),$$

where $M(z)$, together with another function $\tilde{m}(z)$, uniquely solve the following equation in $\mathbb{C}_+ \times \mathbb{C}_+$:

$$\begin{cases} M(z) = -\frac{1}{z} \int_{\tau \in \mathbb{R}} \frac{w_s}{1 + y\tilde{m}(z)w_s} ds, \\ \tilde{m}(z) = -\frac{1}{z} \int_{\tau \in \mathbb{R}} \frac{\tau}{\tau M(z) + 1} dH(\tau). \end{cases}$$

Remark 3.6. If $w_{\ell}^n = 1/n$, then $w(s) = 1$, and Theorem 3.5 reduces to Proposition 2.6. Moreover, if $w(s)$ is not constant, that is, $w(s) \ne \int_0^1 w(t)dt$ on $[0,1]$, then except in the trivial case when H is a delta measure at 0, the LSD $F^w \ne F$, where F is the LSD in Proposition 2.6 determined by $H(\cdot / \int_0^1 w(t)dt)$. (ADD MORE EXPLANATION FROM SUPPLEMENTARY ARTICLE)

Time-variance adjusted realized covariance

Remark 3.7. In order to avoid confusion, we should specify the designations. For any vector $A = (A_1, \ldots, A_n)^T$, $\|A\|_2$ denotes its L^2-norm, that is

$$\|A\|_2 = \sqrt{\sum_{\ell=1}^{n} A_{\ell}^2}.$$

Note that

$$\|A\|_2^2 = A^T A.$$

Let $\mathbf{A}$ be $m \times n$ matrix and $\mathbf{A}^* = \overline{\mathbf{A}^T}$ its conjugate transpose. Let also $\lambda_{\max}(\cdot)$ be the operator, taking the largest eigenvalue. Then

$$\|\mathbf{A}\| = \sqrt{\lambda_{\max}(\mathbf{A}^* \mathbf{A})}$$

denote a spectral norm of $\mathbf{A}$. The value $\lambda_{\max}(\mathbf{A}^* \mathbf{A})$ is also called singular value of the matrix $\mathbf{A}$.

Definition 3.8. Suppose that a diffusion process $\mathbf{X}(t)$ belongs to class $\mathcal{C}$. We define *time-variation adjusted realized covariance (TVARCV) matrix* as follows:

$$\widehat{\Sigma}_p := \frac{\operatorname{tr}\left(\Sigma_p^{RCV}\right)}{n} \sum_{\ell=1}^{n} \frac{\Delta \mathbf{X}_{\ell}(\Delta \mathbf{X}_{\ell})^T}{\|\Delta \mathbf{X}_{\ell}\|_2^2} = \frac{\operatorname{tr}\left(\Sigma_p^{RCV}\right)}{p} \widetilde{\Sigma}_p, \tag{12}$$

where

$$\widetilde{\Sigma}_p := \frac{p}{n} \sum_{\ell=1}^{n} \frac{\Delta \mathbf{X}_{\ell}(\Delta \mathbf{X}_{\ell})^T}{\|\Delta \mathbf{X}_{\ell}\|_2^2}. \tag{13}$$

Remark 3.9. Let us explain $\widetilde{\Sigma}_p$. Consider the simplest case when $\mu(t) = 0$, $\gamma(t)$ is deterministic, $\Lambda_t = \mathbb{I}_{p \times p}$, and $\tau_\ell = \ell/n$, $\ell = 0, 1, \ldots, n$. In this case,

$$\Delta\mathbf{X}_\ell = \sqrt{\int_{(\ell-1)/n}^{\ell/n} \gamma^2(t)dt} \cdot \frac{\mathbf{Z}_\ell}{\sqrt{n}},$$

where $\mathbf{Z}_\ell = (Z_{1\ell}, \ldots, Z_{p\ell})^T$ is a vector of i.i.d. standard normal random variables. Hence,

$$\frac{\Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T}{\|\Delta\mathbf{X}_\ell\|_2^2} = \frac{\mathbf{Z}_\ell\mathbf{Z}_\ell^T}{\|\mathbf{Z}_\ell\|_2^2}.$$

However, as $p \to \infty$, $\|\mathbf{Z}_\ell\|_2^2 \sim p$, hence

$$\widetilde{\Sigma}_p \sim \frac{\sum_{\ell=1}^n \mathbf{Z}_\ell\mathbf{Z}_\ell^T}{n},$$

the latter being the usual sample covariance matrix. We will show that, first, $\mathrm{tr}(\Sigma_p^{RCV}) \sim \mathrm{tr}(\Sigma_p)$; and second, if $\mathbf{X}(t)$ belongs to class $\mathcal{C}$ and satisfies certain additional assumptions, then the LSD of $\widetilde{\Sigma}_p$ is related to that of $\breve{\Sigma}_p$ via the Marchenko-Pastur equation, where

$$\breve{\Sigma}_p = \frac{p}{\mathrm{tr}(\Sigma_p)}\Sigma_p = \Lambda\Lambda^T.$$

Hence, the LSD of $\widehat{\Sigma}_p$ is also related to that of Σ_p via the same Marchenko-Pastur equation. We will now state our assumptions for the major theorem in this paper.

Assumptions 3.10.

(i) there exists $C_0 < \infty$ such that for all p and all $j = 1, \ldots, p$, $|\mu_j^{(p)}(t)| \leq C_0$ for all $t \in [0, 1]$ a.s.;

(ii) there exist constants $C_1 < \infty$, $0 \leq \delta_1 < 1/2$, a sequence $\eta_p < C_1 p^{\delta_1}$ and a sequence of index sets $\mathcal{I}_p$ satisfying $\mathcal{I}_p \subset \{1, \ldots p\}$ and $\#\mathcal{I}_p \leq \eta_p$ such that $\gamma^{(p)}(t)$ may depend on $\mathbf{W}^{(p)}(t)$ but only on $\{W_j^{(p)}(t) : j \in \mathcal{I}_p\}$; moreover, there exists $C_2 < \infty$ such that for all p, $|\gamma^{(p)}(t)| \in (1/C_2, C_2)$ for all $t \in [0, 1]$ a.s.;

(iii) there exists $C_3 < \infty$ such that for all p and for all j, the individual volatilities

$$\sigma_j(t) = \gamma^{(p)}(t)\sqrt{\sum_{k=1}^p (\Lambda_{jk}^{(p)})^2} \in (1/C_3, C_3)$$

for all $t \in [0, 1]$ a.s.;

(iv)

$$\lim_{p \to \infty} \frac{\mathrm{tr}(\Sigma_p)}{p} = \lim_{p \to \infty} \int_0^1 (\gamma^{(p)}(t))^2 dt := \theta > 0 \quad \text{a.s.};$$

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(v) almost surely, as $p \to \infty$, the ESD F^{Σ_p} converges to a probability distribution H on $[0, \infty)$;

(vi) there exist $C_5 < \infty$ and $0 \le \delta_2 < 1/2$ such that for all p, $\|\Sigma_p\| \le C_5 p^{\delta_2}$ a.s.;

(vii) the δ_1 in (ii) and δ_2 in (vi) satisfy that $\delta_1 + \delta_2 < 1/2$;

(viii) $p/n \to y \in (0, \infty)$ as $p \to \infty$; and

(ix) there exists $C_4 < \infty$ such that for all n,

$$\max_{1 \le \ell \le n} n \cdot (\tau_\ell - \tau_{\ell-1}) \le C_4 \quad \text{a.s.}$$

moreover, τ_ℓ's are independent of $\mathbf{X}(t)$.

Remark 3.11. Before stating and proving theorem for TVARCV we provide some useful lemmas.

Lemma 3.12 (Burkholder-Davis-Gundy inequality). *For any real number $p \ge 1$ there exist positive constants c_p and C_p such that, for all local martingales $X(t)$ with $X(0) = 0$, its maximum processes $M(t) = \sup_{t \le s} |X(s)|$ and stopping times τ, the following inequalities hold*

$$c_p \mathbb{E}[[X]^{p/2}(\tau)] \le \mathbb{E}[M^p(\tau)] \le C_p \mathbb{E}[[X]^{p/2}(\tau)].$$

Proof. LOOK IN THE BOOKS □

Definition 3.13. Let F and G be two cumulative distribution functions. Then the Lévy distance between them is

$$L(F, G) := \inf\{\varepsilon > 0 : F(x - \varepsilon) - \varepsilon \le G(x) \le F(x + \varepsilon) + \varepsilon \quad \forall x \in \mathbb{R}\}$$

Lemma 3.14. *Suppose that for each p, $\mathbf{v}_\ell^{(p)} = (v_{1\ell}^{(p)}, \ldots, v_{p\ell}^{(p)})^T$ and $\mathbf{w}_\ell^{(p)} = (w_{1\ell}^{(p)}, \ldots, w_{p\ell}^{(p)})^T$, $\ell = 1, \ldots, n$, are all p-dimensional vectors. Define*

$$\widetilde{S}_n = \sum_{\ell=1}^n (\mathbf{v}_\ell^{(p)} + \mathbf{w}_\ell^{(p)}) \cdot (\mathbf{v}_\ell^{(p)} + \mathbf{w}_\ell^{(p)})^T$$

and

$$S_n = \sum_{\ell=1}^n \mathbf{w}_\ell^{(p)} (\mathbf{w}_\ell^{(p)})^T.$$

If the following conditions are satisfied:

(i) $\lim_{p \to \infty} p/n = y > 0$;

(ii) *there exists a sequence $\varepsilon_p = o(1/\sqrt{p})$ such that for all p and all ℓ, all the entries of $\mathbf{v}_\ell^{(p)}$ are bounded by ε_p in absolute value;*

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(iii) $\limsup_{p\to\infty} \operatorname{tr}(S_n)/p < \infty$ *almost surely.*

Then $L(F^{\widetilde{S}_n}, F^{S_n}) \to 0$ *almost surely.*

Proof. For notational ease we shall omit the superscript p and write $\mathbf{v}_\ell = \mathbf{v}_\ell^{(p)}$, $\mathbf{w}_\ell = \mathbf{w}_\ell^{(p)}$, $v_{j\ell} = v_{j\ell}^{(p)}$, etc. For each p set

$$A = A_p = (\mathbf{v}_1 + \mathbf{w}_1, \ldots, \mathbf{v}_n + \mathbf{w}_1)$$

and

$$B = B_p = (\mathbf{w}_1, \ldots, \mathbf{w}_n).$$

Then

$$\widetilde{S}_n = AA^T \quad \text{and} \quad S_n = BB^T.$$

By Lemma 2.7 of Bai (1999)

$$L^4(F^{\widetilde{S}_n}, F^{S_n}) \le \frac{2}{p^2} \operatorname{tr}((A - B)(A - B)^T) \cdot \operatorname{tr}(\widetilde{S}_n + S_n).$$

By condition (ii),

$$\operatorname{tr}((A - B)(A - B)^T) = \sum_{\ell=1}^{n} \sum_{j=1}^{p} v_{j\ell}^2 \le np \cdot \varepsilon_p^2.$$

Moreover,

$$\operatorname{tr}(\widetilde{S}_n) = \sum_{\ell=1}^{n} \sum_{j=1}^{p} (v_{j\ell} + w_{j\ell})^2$$
$$\le 2 \sum_{\ell=1}^{n} \sum_{j=1}^{p} v_{j\ell}^2 + 2 \sum_{\ell=1}^{n} \sum_{j=1}^{p} w_{j\ell}^2$$
$$\le 2np \cdot \varepsilon_p^2 + 2 \operatorname{tr}(S_n).$$

Therefore,

$$L^4(F^{\widetilde{S}_n}, F^{S_n}) \le 4n^2 \varepsilon_p^4 + 6n\varepsilon_p^2 \frac{\operatorname{tr}(S_n)}{p}.$$

The conclusion follows since $\varepsilon_p = o(1/\sqrt{p})$ and hence $o(1/\sqrt{n})$. $\qquad\square$

Lemma 3.15. *Under Assumptions 3.10(iv), namely, suppose that*

$$\lim_{p\to\infty} \frac{\operatorname{tr}(\Sigma_p)}{p} = \theta \quad a.s.,$$

then

$$\lim_{p\to\infty} \frac{\operatorname{tr}(\Sigma_p^{RCV})}{p} = \theta \quad a.s.$$

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Proof. The full proof can be found in supplementary article OF THE PAPER .... We will provide here the main idea: it suffices to show that

$$\frac{\text{tr}(\Sigma_p^{RCV}) - \text{tr}(\Sigma_p)}{p} \xrightarrow[p \to \infty]{} 0 \quad \text{a.s.}$$

Let us first consider one-dimensional process

$$dX(t) = \mu(t)dt + \sigma(t)dW(t).$$

Under slight assumptions it can be shown that there exist constants c, C_1 and C_2, such that for all $0 \le x \le c\sqrt{n}$

$$\mathbb{P}\left(\sqrt{n}\left|[X, X](1) - \int_0^1 \sigma^2(t)dt\right| \ge x\right) \le C_1 \exp\{-C_2 x^2\}.$$

Therefore, generalizing it for p-dimensional processes, we have for every $1 \le j \le p$

$$\mathbb{P}(\sqrt{n}|(\Sigma_p^{RCV})_{jj} - (\Sigma_p)_{jj}| > x) \le C_1 \exp\{-C_2 x^2\}.$$

Hence, for any $0 < \varepsilon < c$ and for all p,

$$\mathbb{P}\left(\left|\frac{\text{tr}(\Sigma_p^{RCV}) - \text{tr}(\Sigma_p)}{p}\right| \ge \varepsilon\right) \le \sum_{j=1}^p \mathbb{P}(\sqrt{n}|(\Sigma_p^{RCV})_{jj} - (\Sigma_p)_{jj}| \ge \varepsilon)$$

$$\le p \cdot C_1 \exp\{-C_2 \varepsilon^2 \cdot n\}.$$

The almost sure convergence then follows from the Borel-Cantelli lemma[1]. $\qquad\square$

Corollary 3.16. *The main results of the Lemma 3.15 are the convergence of $\breve{\Sigma}_p$*

$$\lim_{p \to \infty} \breve{\Sigma}_p \to \frac{1}{\theta}\Sigma_p$$

and of TVARCV estimator

$$\lim_{p \to \infty} \widehat{\Sigma}_p \to \theta\widetilde{\Sigma}_p.$$

Remark 3.17. Using Lemma 3.15 we state that ESD of $\breve{\Sigma}_p$ converges almost surely:

$$\lim_{p,n \to \infty} F^{\breve{\Sigma}_p} \to \breve{H} \quad \text{a.s.,}$$

where $\breve{H}$ is defined by

$$\breve{H}(x) = H(\theta x) \quad \forall x \ge 0.$$

---

[1]Let $X_1, X_2, \ldots$ be a sequence of random variables in some probability space and $\varepsilon > 0$ be arbitrary. Then

$$\sum_{n=1}^{\infty} \mathbb{P}(|X_n| \ge \varepsilon) < \infty \implies X_n \xrightarrow{a.s.} 0.$$

It is obvious, since

$$F^{\breve{\Sigma}_p} = F^{\Sigma_p}(\text{tr}(\Sigma_p)/p \cdot x) \quad \forall x \geq 0.$$

Now it is left to prove the following:

$$\lim_{p,n \to \infty} F^{\widetilde{\Sigma}_p} \to \widetilde{F} \quad \text{a.s.,}$$

where the LSD $\widetilde{F}$ of $\widetilde{\Sigma}_p$ is determined by $\breve{H}$ in that its Stieltjes transform $m_{\widetilde{F}}(z)$ satisfies the equation

$$m_{\widetilde{F}}(z) = \int_{\tau \in \mathbb{R}} \frac{1}{\tau(1 - y(1 + zm_{\widetilde{F}}(z))) - z} d\breve{H}(\tau).$$

Then the convergence of $\widehat{\Sigma}_p$ will follow from the Corollary 3.16.

Theorem 3.18. *Suppose that for all p, $\mathbf{X}(t) = \mathbf{X}^{(p)}(t)$ is a p-dimensional process in class $\mathcal{C}$ for some drift process $\boldsymbol{\mu}^{(p)}(t) = (\mu_1^{(p)}(t), \ldots, \mu_p^{(p)}(t))^T$, covolatility process $\Theta^{(p)}(t) = \gamma^{(p)}(t)\Lambda^{(p)}$ and a p-dimensional Brownian motion $\mathbf{W}_t^{(p)}$, which satisfy Assumptions 3.10 (i)-(vii) above. Suppose also that p and n satisfy Assumptions 3.10 (viii), and the observation times satisfy Assumptions 3.10 (ix). Let $\hat{\Sigma}_p$ be as in (12). Then, as $p \to \infty$, $F^{\hat{\Sigma}_p}$ converges a.s. to a probability distribution F, which is determined by H through Stieltjes transforms via the same Marchenko-Pastur equation as in Proposition 2.6.*

Proof. We only need to show the convergence of $F^{\widetilde{\Sigma}_p}$. For notational ease, we shall omit the superscript p in the arguments below: thus we write $\boldsymbol{\mu}(t)$ instead of $\boldsymbol{\mu}^{(p)}(t)$, $\gamma(t)$ instead of $\gamma^{(p)}(t)$, Λ instead of $\Lambda^{(p)}$, etc. First, note that

$$\Delta\mathbf{X}_\ell = \int_{\tau_{\ell-1}}^{\tau_\ell} \boldsymbol{\mu}(t)dt + \Lambda \cdot \int_{\tau_{\ell-1}}^{\tau_\ell} \gamma(t)d\mathbf{W}(t) := \sqrt{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma^2(t)dt} \cdot (\mathbf{v}_\ell + \Lambda \cdot \mathbf{Z}_\ell),$$

where

$$\mathbf{v}_\ell = (v_{1\ell}, \ldots, v_{p\ell})^T = \frac{\int_{\tau_{\ell-1}}^{\tau_\ell} \boldsymbol{\mu}(t)dt}{\sqrt{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma^2(t)dt}}$$

and

$$\mathbf{Z}_\ell = (Z_{1\ell}, \ldots, Z_{p\ell})^T = \frac{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma(t)d\mathbf{W}(t)}{\sqrt{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma^2(t)dt}}$$

By performing an orthogonal transformation if necessary, w.l.o.g., we may assume that the index set $\mathcal{I}_p \subset \{1, \ldots, \eta_p\}$. Then by Assumptions 3.10 (ii) and (ix), for $j > \eta_p$, $Z_{j\ell}$ are i.i.d. $\sim \mathcal{N}(0, 1)$. Write

$$\mathbf{U}_\ell = (Z_{1\ell}, \ldots Z_{\eta_p\ell})^T$$

and

$$\mathbf{D}_\ell = (Z_{\eta_p+1\ell}, \ldots Z_{p\ell})^T.$$

With the above notation, $\widetilde{\Sigma}_p$ can be rewritten as

$$\widetilde{\Sigma}_p = y_n \sum_{\ell=1}^n \frac{\Delta\mathbf{X}_\ell (\Delta\mathbf{X}_\ell)^T}{\|\Delta\mathbf{X}_\ell\|_2^2} = y_n \sum_{\ell=1}^n \frac{(\mathbf{v}_\ell + \Lambda\mathbf{Z}_\ell)(\mathbf{v}_\ell + \Lambda\mathbf{Z}_\ell)^T}{\|\mathbf{v}_\ell + \Lambda\mathbf{Z}_\ell\|_2^2} \qquad (14)$$

By Assumptions 3.10 (i), (ii) and (ix), there exists $C > 0$, such that $|v_{j\ell}| \leq C/\sqrt{n}$ for all j and ℓ, hence $\|\mathbf{v}_\ell\|_2$'s are uniformly bounded. We will show that

$$\max_{\ell=1,\ldots,n} \left| \frac{\|\Lambda\mathbf{Z}_\ell\|_2^2}{p} - 1 \right| = \max_{\ell=1,\ldots,n} \left| \frac{\mathbf{Z}_\ell^T \breve{\Sigma}_p \mathbf{Z}_\ell}{p} - 1 \right| \to 0 \quad \text{a.s.,} \qquad (15)$$

which clearly implies that

$$\max_{\ell=1,\ldots,n} \left| \frac{\|\mathbf{v}_\ell + \Lambda\mathbf{Z}_\ell\|_2^2}{p} - 1 \right| \to 0 \quad \text{a.s.} \qquad (16)$$

To prove (15) let us write

$$\breve{\Sigma}_p = \begin{pmatrix} A & B \\ B^T & C \end{pmatrix},$$

where A, B and C are $\eta_p \times \eta_p$, $\eta_p \times (n - \eta_p)$ and $(n - \eta_p) \times (n - \eta_p)$ matrices, respectively. Then

$$\mathbf{Z}_\ell^T \breve{\Sigma}_p \mathbf{Z}_\ell = \mathbf{U}_\ell^T A \mathbf{U}_\ell + 2\mathbf{D}_\ell^T B^T \mathbf{U}_\ell + \mathbf{D}_\ell^T C \mathbf{D}_\ell.$$

By a well-known fact about the spectral norm,

$$\|A\| \leq \|\breve{\Sigma}_p\|, \quad \|B\| \leq \|\breve{\Sigma}_p\| \quad \text{and} \quad \|C\| \leq \|\breve{\Sigma}_p\|.$$

In particular, by Assumptions 3.10 (ii), (vi) and (vii),

$$0 \leq \text{tr}(A) \leq \eta_p \cdot \|\breve{\Sigma}_p\| \leq C p^{\delta_1 + \delta_2} = o(p),$$

hence

$$\frac{\text{tr}(C)}{p} = \frac{\text{tr}(\breve{\Sigma}_p) - \text{tr}(A)}{p} \to 1.$$

Now using the fact that $\mathbf{D}_\ell$ consists of i.i.d. standard normals and by the same proof as that for (READ THE PROOF AND WRITE IT HERE!!!) we get

$$\max_{\ell=1,\ldots,n} \left| \frac{\mathbf{D}_\ell^T C \mathbf{D}_\ell}{p} - 1 \right| \to 0 \quad \text{a.s.}$$

To complete the proof of (15), it then suffices to show that

$$\max_{\ell=1,\ldots,n} \frac{|\mathbf{U}_\ell^T A \mathbf{U}_\ell|}{p} \to 0 \quad \text{and} \quad \max_{\ell=1,\ldots,n} \frac{|\mathbf{D}_\ell^T B \mathbf{U}_\ell|}{p} \to 0 \quad \text{a.s.}$$

We shall only prove the first convergence, the second one can be proved similarly. We have

$$|\mathbf{U}_\ell A \mathbf{U}_\ell| \le \|A\| \cdot \|\mathbf{U}_\ell\|_2^2 \le C_5 p^{\delta_2} \cdot \|\mathbf{U}_\ell\|_2^2 \tag{17}$$

Observe that for all $1 \le i \le \eta_p$, by Assumptions 3.10 (ii),

$$|Z_\ell^{(i)}|^2 = \frac{|\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma_t d\mathbf{W}_t|^2}{\int_{\tau_{\ell-1}}^{\tau_\ell} \gamma_t^2 dt} \le \frac{C_2^2}{\tau_\ell - \tau_{\ell-1}} \cdot \left| \int_{\tau_{\ell-1}}^{\tau_\ell} \gamma_t d\mathbf{W}_t \right|^2.$$

By Lemma 3.12 , we then get that for any $k \in \mathbb{N}$, there exists $\lambda_k > 0$ such that

$$\mathbb{E}[|Z_\ell^{(i)}|^{2k}] \le \lambda_k C_2^{4k}. \tag{18}$$

Now we are ready to show that $\max_{\ell=1,\dots,n} \frac{|\mathbf{U}_\ell^T A \mathbf{U}_\ell|}{p} \to 0$. In fact, for any $\varepsilon > 0$, for any $k \in \mathbb{N}$, by Markov's inequality[2], (17), Hölder's inequality and (18),

$$
\begin{aligned}
\mathbb{P}(\max_{\ell=1,\dots,n} |\mathbf{U}_\ell^T A \mathbf{U}_\ell| \ge p\varepsilon) &\le \sum_{\ell=1}^n \mathbb{P}(|\mathbf{U}_\ell^T A \mathbf{U}_\ell| \ge p\varepsilon) \\
&\le \sum_{\ell=1}^n \frac{\mathbb{E}[|\mathbf{U}_\ell^T A \mathbf{U}_\ell|^k]}{p^k \varepsilon^k} \\
&\le \sum_{\ell=1}^n \frac{C_5^k p^{k\delta_2} \cdot \eta_p \cdot \lambda_k C_2^{4k} \cdot \eta_p^{k-1}}{p^k \varepsilon^k} \\
&\le C p^{1+k\delta_2+k\delta_1-k}.
\end{aligned}
$$

By Assumptions 3.10 (vii), $\delta_1 + \delta_2 < 1/2$, hence by choosing k to be large enough, the right hand side will be summable in p:

$$\sum_{p=1}^\infty \mathbb{P}\big(\max_{\ell=1,\dots,n} |\mathbf{U}_\ell^T A \mathbf{U}_\ell| \ge p\varepsilon\big) < \infty.$$

Hence by Borel–Cantelli lemma, almost surely, $\max_{\ell=1,\dots,n} \frac{|\mathbf{U}_\ell^T A \mathbf{U}_\ell|}{p} \to 0$. We now get back to $\widetilde{\Sigma}_p$. By (16), for any $\varepsilon > 0$, almost surely, for all n sufficiently large, for all $\ell = 1, \dots, n$,

$$p(1-\varepsilon) \le \|\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell\|_2^2 \le p(1+\varepsilon).$$

Hence, almost surely, for all n sufficiently large,

$$\frac{1}{1+\varepsilon} \widetilde{S}_p \preceq \widetilde{\Sigma}_p = y_n \sum_{\ell=1}^n \frac{(\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell)(\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell)^T}{\|\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell\|_2^2} \preceq \frac{1}{1-\varepsilon} \widetilde{S}_p,$$

---

[2]Markov's inequality states that for any non-negative random variable X, $k \in \mathbb{N}$ and $n \in \mathbb{R}_+$ the following inequality holds:

$$\mathbb{P}(X \ge n) \le \frac{\mathbb{E}[X^k]}{n^k}.$$

where

$$\widetilde{S}_p = \frac{1}{n} \sum_{\ell=1}^{n} (\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell)(\mathbf{v}_\ell + \Lambda \mathbf{Z}_\ell)^T.$$

Hence, by Weyl's Monotonicity theorem,

$$F^{\widetilde{S}_p}((1+\varepsilon)x) \geq F^{\widetilde{\Sigma}_p}(x) \geq F^{\widetilde{S}_p}((1-\varepsilon)x) \quad \forall x \geq 0. \tag{19}$$

Next, by Lemma 3.14 $\widetilde{S}_p$ has the same LSD as

$$S_p := \frac{1}{n} \sum_{\ell=1}^{n} \Lambda \mathbf{Z}_\ell (\mathbf{Z}_\ell)^T \Lambda^T.$$

Moreover, by using the same trick as in the beginning of the proof of THEOREM 1 (WHICH TRICK???), F^{S_p} has the same limit as $F^{S'_p}$, where

$$S'_p := \frac{1}{n} \sum_{\ell=1}^{n} \Lambda \widetilde{\mathbf{Z}}_\ell (\widetilde{\mathbf{Z}}_\ell)^T \Lambda^T$$

and $\widetilde{\mathbf{Z}}_\ell$ consists of i.i.d. standard normals. For $F^{S'_p}$, it follows easily from Proposition 2.6 that it converges to $\widetilde{F}$, because $\check{\Sigma}_p = \Lambda \Lambda^T$. Moreover, by BY SOME THEOREMS IN THIS F** BOOKS, $\widetilde{F}$ is differentiable and in particular continuous at all $x > 0$. It follows from (19) that $F^{\widetilde{\Sigma}_p}$ must also converge to $\widetilde{F}$. $\qquad\square$

Part 4
Stochastic volatility case

Remark 4.1. It is useful to consider stochastic volatility case.

Definition 4.2.

(i) The stochastic process $Y(t)$ is called **Cox-Ingersoll-Ross (CIR) process**, if it is determined by the stochastic differential equation (SDE):

$$dY(t) = \beta(\alpha - Y(t))dt + \xi\sqrt{Y(t)}d\overline{W}(t), \qquad (20)$$

where α, β and ξ are positive constants, and $\overline{W}(t)$ is a standard Brownian motion. Conditional on parameters, time t and initial value $Y(0)$, scaled CIR process has a non-central chi-squared distribution with k degrees of freedom and non-centrality parameter λ:

$$cY(t) \sim \chi'^2_k(\lambda),$$

where

$$k = \frac{4\alpha\beta}{\xi^2}, \ \ \lambda = Y(0)ce^{-\beta t} \text{ and } c = \frac{4\beta}{\xi^2(1 - e^{-\beta t})}.$$

CIR process never goes below 0. Moreover, it is a.s. positive if the condition

$$2\alpha\beta \geq \xi^2$$

is met. This property makes CIR process a perfect tool for volatility (or interest rate) modeling.

(ii) The standard **Heston model** assumes that one-dimensional log-price process X_t is determined by the following SDE:

$$dX(t) = \mu dt + \sqrt{Y(t)}dW(t), \qquad (21)$$

where $Y(t)$ satisfies SDE (20) and $W(t)$ is another standard Brownian motion, such that

$$[W, \overline{W}](t) = \int_0^t \rho(s)ds \qquad (22)$$

holds with a (stochastic) process $\rho(t)$, taking values strictly between -1 and 1. Sometimes equality (22) is informally written as

$$dW(t) \cdot d\overline{W}(t) = \rho(t)dt.$$

Example 4.3. Let's assume

$$\gamma(t) = \sqrt{Y(t)},$$

where Y_t is a CIR process, determined by (20). Then it is easy to see that

$$\sigma^2 = \beta\alpha - \beta \int_0^1 \gamma^2(t)dt + \xi \int_0^1 \gamma(t)d\overline{W}(t).$$

Therefore, as long as $\xi > 0$, the value of σ^2 is random and for different realizations of γ_t we have different LSD. Moreover, let

$$\overline{W}(t) = \frac{1}{\sqrt{\eta_p}} \sum_{j=1}^{\eta_p} W_j(t),$$

where η_p is determined by Assumption 3.10 (ii). The process $\overline{W}(t)$ is still the standard Brownian motion, because it is a.s. continuous, has independent increments and

$$\overline{W}(t) - \overline{W}(s) = \frac{1}{\sqrt{\eta_p}} \sum_{j=1}^{\eta_p} (W_j(t) - W_j(s)) \sim \mathcal{N}(0, t-s) \quad \forall t < s.$$

We also take $\eta_p = p$, that means that $\overline{W}(t)$ is a sample average of $\mathbf{W}(t)$ and depends on every its component $W_j(t)$. Simulations show that even in this case, when Assumptions 3.10 (ii) are not satisfied, Theorem 3.18 still holds and F^{TVARCV} converges to Marchenko-Pastur law. Intuitively it can be explained in the following way: For any $0 \leq i \leq p$ we have

$$dW_i(t) \cdot d\overline{W}(t) = dW_i(t) \cdot \frac{1}{\sqrt{\eta_p}} \sum_{j=1}^{\eta_p} dW_j(t) = \frac{dt}{\sqrt{\eta_p}} \mathbf{1}_{\{i \leq \eta_p\}},$$

or equivalently

$$[W_i, \overline{W}](t) = \frac{t}{\sqrt{\eta_p}} \mathbf{1}_{\{i \leq \eta_p\}} \leq \frac{1}{\eta_p}.$$

Hence, with growing η_p the dependence of $\overline{W}_t$ on ... becomes smaller and for $\eta_p = p \to \infty$ correlation between them goes to 0.
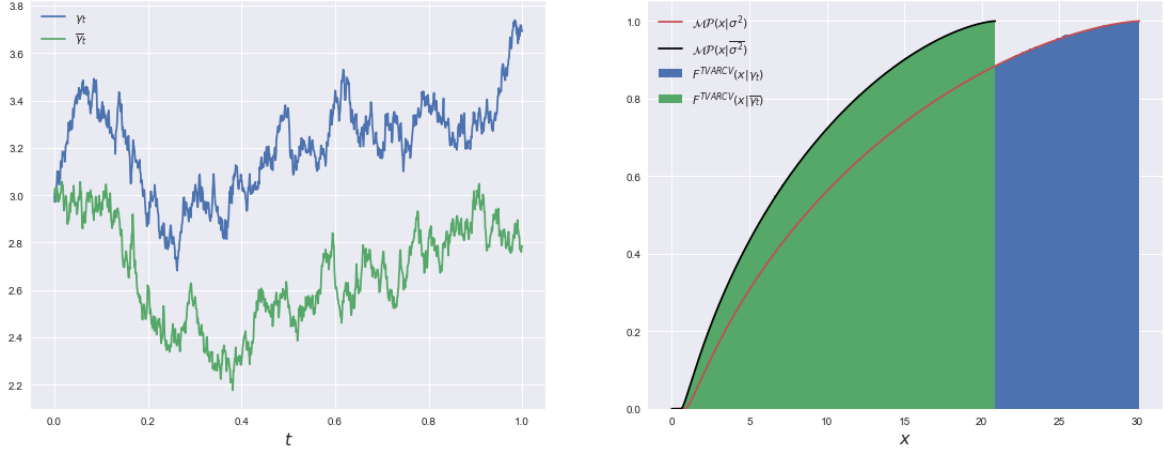
Figure 3: $\eta_p = p, y = 0.5, \alpha = 5, \beta = 1, \gamma_0 = 3, \xi = 2$

Left panel shows two different realizations of $\gamma(t)$. On the right panel blue and green histograms are ESD of $\widehat{\Sigma}_p$ for corresponding volatility processes, red and black curve – Marchenko-Pastur law for different realizations of σ^2. We see that even with the same parameters of volatility model, we have the different scale parameter in asymptotic spectral distribution for different realizations.

Part 5

Jump-diffusion processes

Remark 5.1. (CHECK TEXT HERE!) Before we discussed pure diffusion processes, namely that these processes are continuous almost sure. It is quite useful to consider so-called jump-diffusion processes... How should we change our estimator then? ALSO MENTION ONE- AND FINITE-DIMENSIONAL PROCESSES, THEN GO FOR P->INF ALSO WE ONLY CONSIDER EQUIDISTANT TIMES AND ONLY SIGMA EQUAL TO IDENTITY, WHICH MEANS INDPENENCE AND MU = 0

Example 5.2.

(i) Simplest model: let $U \sim \mathcal{U}[0,1]$ be standard uniformly distributed random variable and

$$d\mathbf{X}(t) = \gamma(t)d\mathbf{W}(t) + \Delta_U(t)dt, \qquad (23)$$

where $\Delta_U(t) = (\delta_U(t), \ldots, \delta_U(t))^T$ is a vector of Dirac delta functions shifted to U. In other words, all the components of $\mathbf{X}(t)$ make a jump of size 1 at the same random time $t = U$. Let us denote by $\mathbf{X^c}(t)$ the continuous part of $\mathbf{X}(t)$:

$$d\mathbf{X^c}(t) = \gamma(t)d\mathbf{W}(t).$$

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Moreover, let ℓ' denote the index of interval, where U is located:

$$U \in \left[\frac{\ell'-1}{n}, \frac{\ell'}{n}\right).$$

Then

$$\Sigma_p^{RCV} = \sum_{\ell=1, \ell \neq \ell'}^{n} \Delta\mathbf{X}_\ell (\Delta\mathbf{X}_\ell)^T + \Delta\mathbf{X}_{\ell'}(\Delta\mathbf{X}_{\ell'})^T$$

$$= \sum_{\ell=1, \ell \neq \ell'}^{n} \Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T + (\Delta\mathbf{X^c}_{\ell'} + \mathbb{U}_{p\times 1})(\Delta\mathbf{X^c}_{\ell'} + \mathbb{U}_{p\times 1})^T$$

$$= \sum_{\ell=1}^{n} \Delta\mathbf{X^c}_\ell(\Delta\mathbf{X^c}_\ell)^T + \Delta\mathbf{X^c}_{\ell'} \cdot \mathbb{U}_{1\times p} + (\Delta\mathbf{X^c}_{\ell'} \cdot \mathbb{U}_{1\times p})^T + \mathbb{U}_{p\times p},$$

where $\mathbb{U}_{k\times j}$ stands for $k \times j$ matrix filled with ones. Note that the elements of $\Delta\mathbf{X^c}_{\ell'} \cdot \mathbb{U}_{1\times p}$ are i.i.d. $\sim \mathcal{N}\left(0, \Gamma_{\ell'}/n\right)$ with $\Gamma_{\ell'} = \int_{(\ell'-1)/n}^{\ell'/n} \gamma^2(t)dt$. Therefore, if we set $n \to \infty$, $\text{tr}(\Sigma_p^{RCV})$ will converge to the same value as before plus p. Also,

$$\|\Delta\mathbf{X}_{\ell'}\|_2^2 = \|\Delta\mathbf{X^c}_{\ell'} + \mathbb{U}_{p\times 1}\|_2^2 = \sum_{j=1}^{p}(\Delta X_{\ell'}^c)^2 + 2\sum_{j=1}^{p}\Delta X_{\ell'}^c + p.$$

We have asymptotically:

$$\sum_{j=1}^{p}(\Delta X_{\ell'}^c)^2 \sim \frac{p}{n}\Gamma_{\ell'}$$

as a sum of p i.i.d. normally distributed random variables with variance $\Gamma_{\ell'}/n$ and

$$\sum_{j=1}^{p}\Delta X_{\ell'}^c \sim \mathcal{N}\left(0, \frac{p}{n}\Gamma_{\ell'}\right).$$

As long as p/n converges to some finite constant, $\Gamma_{\ell'}$ is bounded and p converges to infinity, we have

$$\frac{\Delta\mathbf{X}_{\ell'}(\Delta\mathbf{X}_{\ell'})^T}{\|\Delta\mathbf{X}_{\ell'}\|_2^2} \to 0.$$

Hence, if we know that all components jumped at the same time and the size of the jump is equal to 1, new estimator will be

$$\widehat{\Sigma}_p = \frac{\text{tr}(\Sigma_p^{RCV}) - p}{n}\sum_{\ell=1}^{n}\frac{\Delta\mathbf{X}_\ell(\Delta\mathbf{X}_\ell)^T}{\|\Delta\mathbf{X}_\ell\|_2^2}.$$

(ii) Now let $U_1, \ldots, U_p$ be i.i.d. $\sim \mathcal{U}[0,1]$, $\mathbf{U} = (U_1, \ldots, U_p)^T$ and

$$d\mathbf{X}(t) = \gamma(t)d\mathbf{W}(t) + \Delta_\mathbf{U}(t)dt$$

with $\Delta_\mathbf{U}(t) = (\delta_{U_1}(t), \ldots, \delta_{U_p}(t))^T$.

Remark 5.3. In Example 5.2 we assumed that in the interval $[0, 1]$ each component of the process $\mathbf{X}(t)$ had exactly one jump. In practice, jump-diffusion processes are often modeled using Poisson processes, where one cannot be sure how many jumps can happen in given interval.

Definition 5.4.

(i) To construct a Poisson process, we begin with a sequence of arrival times $\tau_1^N, \tau_2^N, \dots$, where

$$\tau_{j+1}^N - \tau_j^N \sim \text{Exp}(\lambda) \quad \forall j \geq 1.$$

This setting is similar to the one in Example 6.3, except that now τ_j^N stands for the time of j-th jump. ***The Poisson process*** $N(t)$ counts the number of jumps that occur at or before time t. More precisely,

$$N(t) = \sum_{j \geq 1} \mathbf{1}_{\{\tau_j^N \leq t\}}.$$

Because the expected time between jumps is $1/\lambda$, the jumps are arriving at an average rate of λ per unit time. We say the Poisson process $N(t)$ has intensity λ.

(ii) Let $N(t)$ be a Poisson process with intensity λ, and let $Y_1, Y_2, \dots$ be a sequence of i.i.d. random variables with mean

$$\beta = \mathbb{E}[Y_1].$$

We assume the random variables $Y_1, Y_2, \dots$ are also independent of $N(t)$. We define the ***compound Poisson process***

$$Q(t) = \sum_{i=1}^{N(t)} Y_i, \quad t \geq 0.$$

The jumps in $Q(t)$ occur at the same times as the jumps in $N(t)$, but whereas the jumps in $N(t)$ are always of size 1, the j-th jump in $Q(t)$ is of random size Y_j.

(iii) A process $X(t)$ of the form

$$X(t) = \int_0^t \mu(s)ds + \int_0^t \sigma(s)dW(s) + J(t),$$

where $\mu(s), \sigma(s)$ are adapted processes, $W(s)$ is a standard Brownian motion and $J(t)$ is an adapted right-continuous pure jump process with $J(0) = 0$, is called a ***jump-diffusion process***.

Proposition 5.5. Let $X_1(t) = \int_0^t \mu_1(s)ds + \int_0^t \sigma_1(s)dW(s) + J_1(t)$ be a jump-diffusion process. Then

$$[X_1, X_1](t) = \int_0^t \sigma_1^2(s)ds + \sum_{0 < s \leq t} (\Delta J_1(s))^2.$$

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Let also $X_2(t) = \int_0^t \mu_2(s)ds + \int_0^t \sigma_2(s)dW(s) + J_2(t)$, then

$$[X_1, X_2](t) = \int_0^t \sigma_1(s)\sigma_2(s)ds + \sum_{0<s\leq t} \Delta J_1(s)\Delta J_2(s).$$

Proof. See Theorem 11.4.7 in Shreve $\qquad\qquad\square$

Remark 5.6. Let us again look at the case with fixed p. IF WE HAVE U1, THEN WE HAVE

$$\Sigma_p^{RCV} \xrightarrow{\mathbb{P}} \Sigma_p + \mathbb{U}_{p\times p}$$

OR IN THE CASE OF U2:

$$\Sigma_p^{RCV} \xrightarrow{\mathbb{P}} \Sigma_p + \mathbb{I}_{p\times p}$$
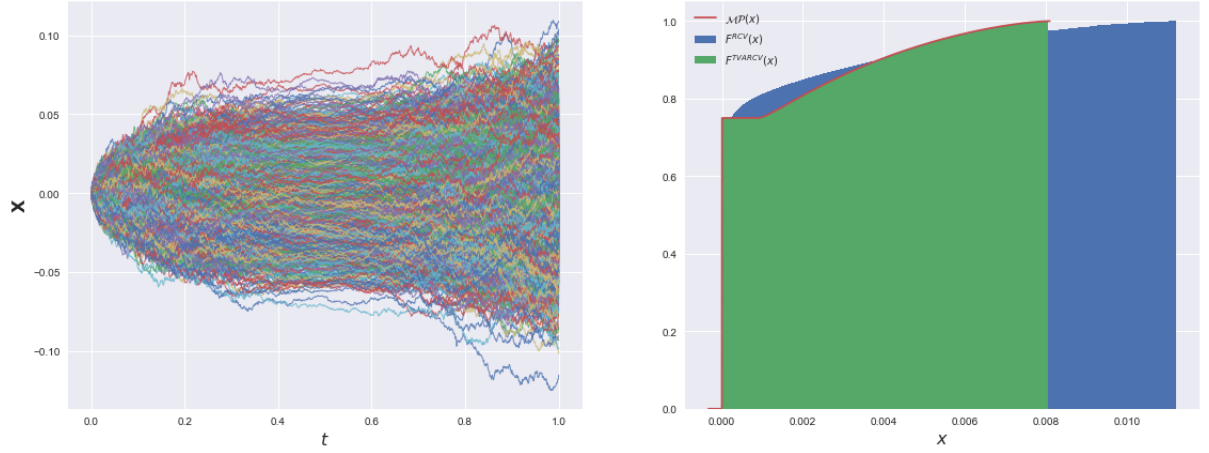
(CHECK IT IN SIMULATIONS) BUT WHAT IF P IS INFINITE?

Figure 4: $\mathbf{X}(t)$ from Example 6.2, $n = 1000$, $p = 4000$, $\sigma^2 = 0.0009$

In this setting $y = p/n = 4$. One the left panel one sees four thousands independent stochastic processes. On the right blue histogram shows ESD of Σ_p^{RCV}, green histogram – ESD of $\widehat{\Sigma}_p$ and red curve – Marchenko-Pastur law. Note point mass equal to 0.75 at the origin.

Part 6
Simulation studies

Non-constant volatility

Remark 6.1. BLA BLA BLA Here we do simulations

Example 6.2. Let's take an example from the paper...

$$\gamma(t) = \sqrt{0.0009 + 0.0008\cos(2\pi t)}.$$

Then

$$\sigma^2 = 0.0009 + \int_0^1 0.0008\cos(2\pi t)dt = 0.0009.$$

The simulations show that TVARCV estimator converges asymptotically to Marchenko-Pastur law, while RCV converges to ???

Non-equidistant observation times

Example 6.3. Let us set time points $\{\tau_\ell\}_{0 \le \ell \le n}$, where $\tau_0 = 0$ a.s. and interarrival times are independent and obey the exponential distribution with rate λ, that is

$$\mathbb{P}(\tau_\ell - \tau_{\ell-1} > x) = e^{-\lambda x}, \quad 1 \le \ell \le n.$$
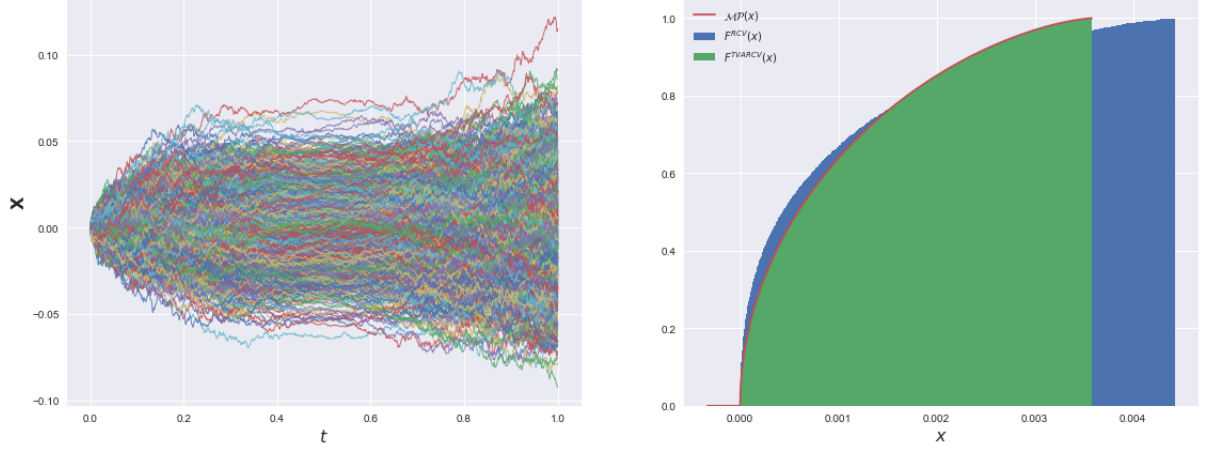
Figure 5: $\mathbf{X}(t)$ from Example 6.2, $n = 1000$, $p = 1000$

In this setting $y = p/n = 1$. Note, that the domain of F^{RCV} is larger than the domain of Marchenko-Pastur law, as it was proved in Proposition 2.13.

Then by Strong Law of Large Numbers

$$\frac{\tau_n}{n} = \frac{1}{n}\sum_{\ell=1}^{n}(\tau_\ell - \tau_{\ell-1}) \xrightarrow{a.s.} \mathbb{E}[\tau_1] = \lambda.$$

Define

$$\mathcal{T} = \{\tau_\ell : \tau_\ell \in [0,1]\},$$

then by setting $\lambda = \frac{1}{n}$, we get $\#\mathcal{T} \sim n$. Let's take volatility γ_t as in Example 6.2. (CONSIDER WHEN TAU DEPENDS ON X, SO THE LAST ASSUMPTION WON'T BE SATISFIED)

INTERESTING CASE

Remark 6.4. For TVARCV to converge to Marchenko-Pastur law, it is sufficient, but not necessary, that $\mathbf{X}(t)$ belongs to the class $\mathcal{C}$. Morevover, we can show that it might be the case when even have convergence for RCV.

Example 6.5. Let as usual $\boldsymbol{\mu}(t) = 0$ and $\Theta(t)$ diagonal matrix, satisfying the following:

$$\left(\Theta(t)\right)_{jj} = \cos(\pi t)\mathbf{1}_{\{j \leq p/2\}} + \sin(\pi t)\mathbf{1}_{\{j > p/2\}}, \quad 1 \leq j \leq p.$$

In other words, first $p/2$ elements of the diagonal of $\Theta(t)$ are $\cos(\pi t)$ and the others are
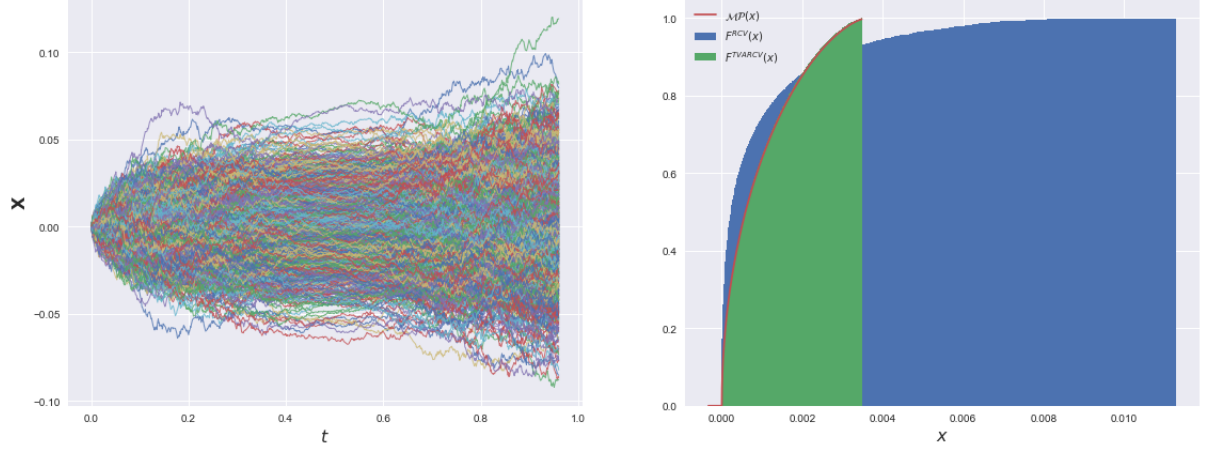
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Figure 6: $\mathbf{X}(t)$ from Example 6.3, $n \simeq 1000$, $p = 1000$

We see that if observation times are not equidistant, the ESD of RCV changes drastically, even if visually there is no difference in realizations of $\mathbf{X}(t)$ and volatility process $\gamma(t)$ stays the same.

$\sin(\pi t)$. Obviously, $\mathbf{X}(t)$ doesn't belong to class $\mathcal{C}$. The ICV matrix is in this case

$$\Sigma_p = \begin{pmatrix} \int_0^1 \cos(\pi t)^2 dt & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \int_0^1 \sin(\pi t)^2 dt \end{pmatrix} = \frac{1}{2}\mathbb{I}_{p \times p}.$$

Example 6.6. Let us change slightly the structure of $\Theta(t)$:

$$\left(\Theta(t)\right)_{jj} = \cos(\pi t)\mathbf{1}_{\{j \leq p/4\}} + \sin(\pi t)\mathbf{1}_{\{j > p/4\}}, \quad 1 \leq j \leq p.$$

The ICV matrix Σ_p stays the same. But we see that both estimators RCV and TVARCV are not converging to Marchenko-Pastur law anymore.
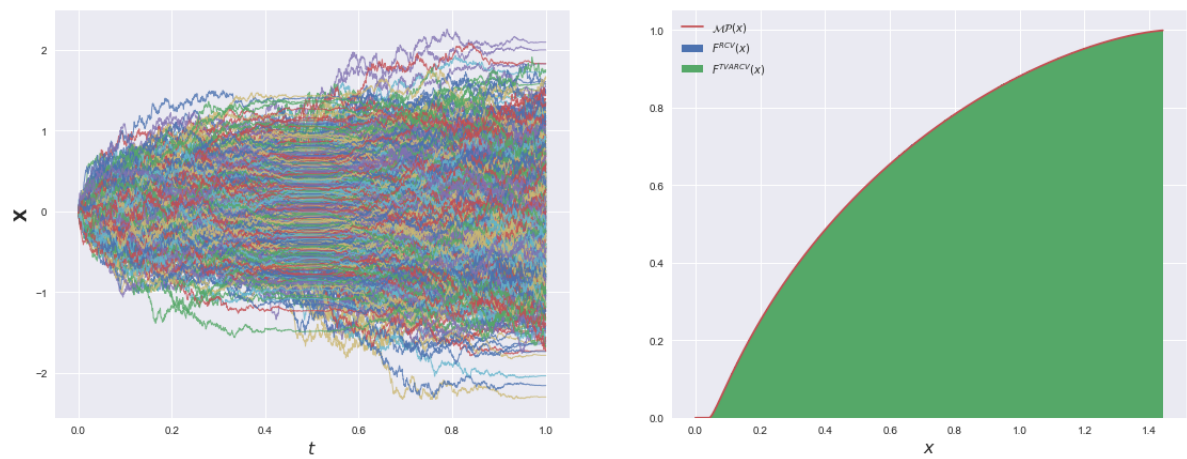
Figure 7: $\mathbf{X}(t)$ from Example 6.5, $n = 3000$, $p = 1000$

Blue and green histograms coincide...

Part 7
Conclusion

Appendices

Cumulative distribution for Marchenko-Pastur law

Python code

```python
import seaborn as sns
import matplotlib.pyplot as plt
import matplotlib
import pandas as pd
import scipy as sc
import numpy as np


# Marchenko-Pastur probability density function
def MPPdf(x, y, sigmaSq):
    if x == 0:
        if y > 1:
            return np.inf
        return 0
    sqrtY = np.sqrt(y)
    a = sigmaSq * (1 - sqrtY) * (1 - sqrtY)
    b = sigmaSq * (1 + sqrtY) * (1 + sqrtY)
    if x < a or x > b:
        return 0
    pdf = np.sqrt((b - x) * (x - a))
    pdf /= (y * x)
    pdf /= (2 * np.pi * sigmaSq)
    return pdf


# Marchenko-Pastur cumulative distribution function
def MPCdf(x, y, sigmaSq):
    # Marchenko-Pastur cdf for y > 1
    def MPCdfLargeRatio(x, y):
        y1 = 1 - x + y
        temp = np.sqrt(4 * y - y1 * y1)
        y1 /= temp
        y1 = (1 + y) * np.arctan(y1)
        y2 = x * (1 + y)
        y2 -= (y - 1) * (y - 1)
        y2 /= (temp * (y - 1))
        y2 = (y - 1) * np.arctan(y2)
        cdf = np.pi - temp + y1 + y2
        cdf /= (2 * np.pi * y)
        return 1 - cdf
```

```python
    # Marchenko-Pastur cdf for y <= 1
    def MPCdfSmallRatio(x, y):
        y1 = 1 - x + y
        temp = np.sqrt(4 * y - y1 * y1)
        y1 /= temp
        y1 = (1 + y) * np.arctan(y1)
        if y == 1:
            y2 = 0
        else:
            y2 = x * (1 + y)
            y2 -= (y - 1) * (y - 1)
            y2 /= (temp * (1 - y))
            y2 = (y - 1) * np.arctan(y2)
        cdf = np.pi * y + temp - y1 + y2
        cdf /= (2 * np.pi * y)
        return cdf

    x0 = x / sigmaSq
    sqrtY = np.sqrt(y)
    a = (1 - sqrtY) * (1 - sqrtY)
    b = (1 + sqrtY) * (1 + sqrtY)
    if x0 >= b:
        return 1
    if x0 < 0:
        return 0
    if y > 1:
        if x0 <= a:
            return 1 - 1 / y
        return MPCdfLargeRatio(x0, y)
    if x0 <= a:
        return 0
    return MPCdfSmallRatio(x0, y)


# get sorted eigenvalues
def get_sorted_eigvals(Sigma):
    return np.sort(np.linalg.eigvals(Sigma))


# get empirical spectral distribution
def ESD(x, sorted_eig_values):
    numerator = len(sorted_eig_values[sorted_eig_values <= x])
    denominator = len(sorted_eig_values)
```

```python
        return numerator / denominator


# TVARCV
def TVARCV(dX, Sigma_RCV):
    (p, n) = dX.shape
    Sigma = np.zeros((p, p))
    minnorm = p
    maxnorm = 0
    for l in range(n):
        dX_l = np.matrix(dX[:, l]).transpose()
        norm = np.linalg.norm(dX_l, ord=2)
        Sigma += dX_l.dot(np.transpose(dX_l)) / (norm * norm)
    Sigma *= np.trace(Sigma_RCV) / n
    return Sigma


# RCV
def RCV(dX):
    (p, n) = dX.shape
    Sigma = np.zeros((p, p))
    for l in range(n):
        dX_l = np.matrix(dX[:, l]).transpose()
        Sigma += dX_l.dot(np.transpose(dX_l))
    return Sigma


# function to plot stochastic processes in Figure 1
def plot_exmp_1():
    n = 10000
    time = np.linspace(0, 1, n)
    mu = np.array([0, 0, 0])
    X = np.zeros((n, 3))
    dt = time[1] - time[0]
    sqrt_dt = np.sqrt(dt)
    dX = np.random.normal(0, sqrt_dt, 3)
    X[1] = X[0] + dX
    for i in range(1, n - 1):
        dW = np.random.normal(0, sqrt_dt, 3)
        Theta = np.identity(3)
        Theta[0][1] = Theta[1][0] = np.cos(np.pi * time[i])
        Theta[0][2] = Theta[2][0] = np.sin(np.pi * time[i])
        dX_l = np.dot(Theta, dW)
```

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        dX_l += mu * dt
        X[i + 1] = X[i] + dX_l
        dX = np.c_[dX, dX_l]

    plt.plot(time, X)
    plt.legend(["$X_1(t)$", "$X_2(t)$", "$X_3(t)$"], fontsize = 16)
    plt.xlabel('$t$', fontsize = 16)
    plt.ylabel('$\mathbf{X}$', fontsize = 16)
    plt.show()


# function to plot graphs of Marchenko-Pastur law (Figure 2)
def plot_MP_law
    y = [0.1, 0.25, 1, 4]
    sigma_sq = [1, 1, 1, 0.3]
    m = 300
    xpdf = np.linspace(1/m, 4, m)
    xcdf = np.linspace(-1/m, 4, m)
    cdf = np.zeros(m)
    pdf = np.zeros(m)

    plt.subplot(1, 2, 1)
    for ratio, vol in zip(y, sigma_sq):
        for i in range(m):
            pdf[i] = MPPdf(xpdf[i], ratio, vol)
        plt.plot(xpdf, pdf)
    plt.ylim(ymax=2.25)
    # plot infinity line for y > 1
    plt.plot((0, 0), (0, 3), color='C3')
    plt.legend(["$y=0.1, \sigma^2=1$", "$y=0.25, \sigma^2=1$", \
        "$y=1, \sigma^2=1$", "$y=4, \sigma^2=0.3$"])
    plt.xlabel('$x$', fontsize = 16)
    plt.ylabel('$f(x)$', fontsize = 16)

    plt.subplot(1, 2, 2)
    for ratio, vol in zip(y, sigma_sq):
        for i in range(m):
            cdf[i] = MPCdf(xcdf[i], ratio, vol)
        plt.plot(xcdf, cdf)
    plt.legend(["$y=0.1, \sigma^2=1$", "$y=0.25, \sigma^2=1$", \
        "$y=1, \sigma^2=1$", "$y=4, \sigma^2=0.3$"])
    plt.xlabel('$x$', fontsize = 16)
    plt.ylabel('$F(x)$', fontsize = 16)
```

```
plt.show()
```