Proof: Asymptotic Covariance Reduction in Linear Regression

Objective

To prove that in standard linear regression with fixed observation noise variance σ^2 , the covariance of the maximum likelihood estimator $\hat{\theta}_{\rm ML}$ decreases as the number of samples $D \to \infty$. Specifically,

$$\operatorname{Cov}(\hat{\theta}_{\operatorname{ML}}) \approx \frac{\sigma^2}{D} Q^{-1}, \quad \text{where } Q = \mathbb{E}[\phi(\zeta)\phi(\zeta)^{\top}]$$

1. Setup

We consider a standard linear regression model:

$$\gamma^{(d)} = \phi(\zeta^{(d)})^{\top} \theta^* + \varepsilon^{(d)}, \quad \varepsilon^{(d)} \sim \mathcal{N}(0, \sigma^2), \quad d = 1, \dots, D$$

Define the design matrix:

$$\Phi = \begin{bmatrix} \phi(\zeta^{(1)})^{\top} \\ \phi(\zeta^{(2)})^{\top} \\ \vdots \\ \phi(\zeta^{(D)})^{\top} \end{bmatrix} \in \mathbb{R}^{D \times B}, \quad \Gamma = \begin{bmatrix} \gamma^{(1)} \\ \gamma^{(2)} \\ \vdots \\ \gamma^{(D)} \end{bmatrix} = \Phi \theta^* + \varepsilon$$

where $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_D)$.

2. Maximum Likelihood Estimator

The MLE (least-squares) solution is given by:

$$\hat{\theta}_{\mathrm{ML}} = (\Phi^{\top}\Phi)^{-1}\Phi^{\top}\Gamma$$

Substitute $\Gamma = \Phi \theta^* + \varepsilon$:

$$\hat{\theta}_{\mathrm{ML}} = \theta^* + (\Phi^{\top} \Phi)^{-1} \Phi^{\top} \varepsilon$$

3. Covariance of $\hat{\theta}_{ML}$

Taking covariance on both sides:

$$\operatorname{Cov}(\hat{\theta}_{\mathrm{ML}}) = \operatorname{Cov}\left((\Phi^{\top}\Phi)^{-1}\Phi^{\top}\varepsilon\right)$$

Using the linear transformation rule:

$$Cov(A\varepsilon) = A \cdot Cov(\varepsilon) \cdot A^{\top}$$

Let $A = (\Phi^{\top}\Phi)^{-1}\Phi^{\top}$, and $Cov(\varepsilon) = \sigma^2 I_D$, then:

$$\operatorname{Cov}(\hat{\theta}_{\operatorname{ML}}) = \sigma^2(\Phi^{\top}\Phi)^{-1}\Phi^{\top}\Phi(\Phi^{\top}\Phi)^{-1} = \sigma^2(\Phi^{\top}\Phi)^{-1}$$

4. Asymptotic Behavior as $D \to \infty$

Assume the samples $\zeta^{(d)}$ are i.i.d., and define:

$$Q := \mathbb{E}[\phi(\zeta)\phi(\zeta)^{\top}] \in \mathbb{R}^{B \times B}$$

Then by the Law of Large Numbers:

$$\frac{1}{D}\Phi^{\top}\Phi = \frac{1}{D}\sum_{d=1}^{D}\phi(\zeta^{(d)})\phi(\zeta^{(d)})^{\top} \xrightarrow{a.s.} Q$$

So for large D:

$$\boldsymbol{\Phi}^{\top}\boldsymbol{\Phi} \approx \boldsymbol{D}\boldsymbol{Q} \quad \Rightarrow \quad (\boldsymbol{\Phi}^{\top}\boldsymbol{\Phi})^{-1} \approx \frac{1}{D}\boldsymbol{Q}^{-1}$$

Hence:

$$\operatorname{Cov}(\hat{\theta}_{\mathrm{ML}}) = \sigma^{2}(\Phi^{\top}\Phi)^{-1} \approx \frac{\sigma^{2}}{D}Q^{-1}$$

5. Conclusion

As $D \to \infty$, the covariance of the parameter estimate $\hat{\theta}_{\rm ML}$ decreases proportionally to 1/D, and we have:

$$\hat{\theta}_{\mathrm{ML}} \xrightarrow{p} \theta^*$$
 (consistency)

Thus, with fixed noise variance σ^2 , increasing the dataset size makes the estimate more concentrated and accurate.