Machine Learning Exercise Sheet 10

Dimensionality Reduction & Matrix Factorization

In-class Exercises

Problem 1: In this exercise, we use proof by induction to show that the linear projection onto an M-dimensional subspace that maximizes the variance of the projected data is defined by the M eigenvectors of the data covariance matrix S, given by

$$S = \frac{1}{N} \sum_{n=1}^{N} (\boldsymbol{x}_n - \bar{\boldsymbol{x}}) (\boldsymbol{x}_n - \bar{\boldsymbol{x}})^T \qquad \bar{\boldsymbol{x}} = \frac{1}{N} \sum_{n=1}^{N} \boldsymbol{x}_n$$

corresponding to the M largest eigenvalues. In Section 12.1 in Bishop this result was proven for the case of M=1. Now suppose the result holds for some general value of M and show that it consequently holds for dimensionality M+1. To do this, first set the derivative of the variance of the projected data with respect to a vector u_{M+1} defining the new direction in data space equal to zero. This should be done subject to the constraints that u_{M+1} be orthogonal to the existing vectors $u_1, \ldots u_M$, and also that it be normalized to unit length. Use Lagrange multipliers to enforce these constraints. Then make use of the orthonormality properties of the vectors u_1, \ldots, u_M to show that the new vector u_{M+1} is an eigenvector of S. Finally, show that the variance is maximized if the eigenvector is chosen to be the one corresponding to eigenvector λ_{M+1} where the eigenvalues have been ordered in decreasing value.

Suppose that the result holds for projection spaces of dimensionality M. The M+1 dimensional principal subspace will be defined by the M principal eigenvectors u_1, \ldots, u_M together with an additional direction vector u_{M+1} whose value we wish to determine. We must constrain u_{M+1} such that it cannot be linearly related to u_1, \ldots, u_M (otherwise it will lie in the M-dimensional projection space instead of defining an M+1 independent direction). This can easily be achieved by requiring that u_{M+1} be orthogonal to u_1, \ldots, u_M , and these constraints can be enforced using Lagrange multipliers η_1, \ldots, η_M .

Following the argument given in section 12.1.1 for u_1 we see that the <u>variance in the direction</u> u_{M+1} is given by $u_{M+1}^T S u_{M+1}$. We now maximize this using a Lagrange multiplier λ_{M+1} to enforce the normalization constraint $u_{M+1}^T u_{M+1} = 1$. Thus we seek a maximum of the function:

$$m{u}_{M+1}^T m{S} m{u}_{M+1} + \lambda_{M+1} (1 - m{u}_{M+1}^T m{u}_{M+1}) + \sum_{i=1}^M \eta_i m{u}_{M+1}^T m{u}_i$$

with respect to u_{M+1} . The stationary points occur when

$$0 = 2Su_{M+1} - 2\lambda_{M+1}u_{M+1} + \sum_{i=1}^{M} \eta_{i}u_{i}$$

Left multiplying with u_j^T , and using the orthogonality constraints, we see that $\eta_j = 0$ for j = 1, ..., M. We therefore obtain

$$Su_{M+1} = \lambda_{M+1}u_{M+1}$$

and so u_{M+1} must be an eigenvector of S with eigenvalue λ_{M+1} . The variance in the direction u_{M+1} is given by $u_{M+1}^T S u_{M+1} = \lambda_{M+1}$ and so is maximized by choosing u_{M+1} to be the eigenvector having the largest eigenvalue amongst those not previously selected. Thus the result holds also for projection spaces of dimensionality M+1, which completes the inductive step. Since we have already shown this result explicitly for M=1 if follows that the result must hold for any $M \ll D$.

Problem 2: Proof that minimizing the error is equivalent to maximizing the variance.

See Bishop chapter 12.1.2.

Homework

PCA

Problem 3: Let the matrix $X \in \mathbb{R}^{N \times D}$ represent N data points of dimension D = 10 (samples stored as rows). We applied PCA to X. By using the K = 5 top principal components, we transformed/projected X into $\tilde{X} \in \mathbb{R}^{N \times K}$. We computed that \tilde{X} preserves 70% of the variance of the original data X.

Suppose now we apply PCA on the following matrices:

- a) $Y_1 = XS$ where $S = \lambda I$, with $\lambda \in \mathbb{R}$ and $I \in \mathbb{R}^{D \times D}$ is the identity matrix
- b) $Y_2 = XR$ where $R \in \mathbb{R}^{D \times D}$ and $RR^T = I$
- c) $Y_3 = XP$ where $P = \text{diag}(+5, -5, \dots, +5, -5)$ is a $D \times D$ diagonal matrix
- d) $Y_4 = XQ$ where Q = diag(1, 2, 3, ..., D 1, D) is a $D \times D$ diagonal matrix
- e) $Y_5 = X + \mathbf{1}_N \boldsymbol{\mu}^T$ where $\boldsymbol{\mu} \in \mathbb{R}^D$ and $\mathbf{1}_N$ is an N-dimensional column vector of all ones
- f) $Y_6 = XA$ where $A \in \mathbb{R}^{D \times D}$ and rank(A) = 5

and obtain the projected data $\tilde{\mathbf{Y}}_1, \dots \tilde{\mathbf{Y}}_6 \in \mathbb{R}^{N \times K}$ using the principal components corresponding to the top K = 5 largest eigenvalues of the respective \mathbf{Y}_i .

What fraction of variance of each Y_i will be preserved by each respective \tilde{Y}_i ? Justify your answer.

The answer "cannot tell without additional information" is also valid if you provide a justification.

- a) 70%. All eigenvalues are scaled by the same amount λ^2 , so the fraction doesn't change.
- b) 70%. **R** is a rotation/reflection/permutation matrix. The direction of the eigenvectors of the covariance matrix is changed, but the eigenvalues stay the same.
- c) 70%. This is just combination of (a) and (b). All data points are scaled by 5 (i.e. eigenvalues of X^TX are all scaled by 25), and some dimensions are reflected around origin, but the fraction of variance explained by the first K components stays the same.
- d) We <u>cannot tell</u> without additional information. since each column (i.e. each dimension) is scaled by a different amount.
- e) 70%. All data points are shifted by μ . But since we center the data as the first step of PCA, shifting has no effect.
- f) 100%. Since $\operatorname{rank}(\mathbf{A}) = 5$, $\operatorname{rank}(\mathbf{Y}_6) \leq 5$ as well. This means that the <u>data lies in a</u> ≤ 5 dimensional subspace, and the <u>first 5 principal components</u> cap<u>ture all the variance</u>.

Problem 4: You are given N = 4 data points: $\{x_i\}_{i=1}^4, x_i \in \mathbb{R}^3$, represented with the matrix $X \in \mathbb{R}^{4 \times 3}$.

$$\boldsymbol{X} = \begin{bmatrix} 4 & 3 & 2 \\ 2 & 1 & -2 \\ 4 & -1 & 2 \\ -2 & 1 & 2 \end{bmatrix}$$

Hint: In this task the results of all (final and intermediate) computations happen to be integers.

a Perform principal component analysis (PCA) of the data X, i.e. find the principal components and their associated variances in the transformed coordinate system. Show your work.

First we center the data. The mean is $\bar{x} = [2, 1, 1]$, thus we have

$$m{X}_c = m{X} - ar{m{x}} = egin{bmatrix} 2 & 2 & 1 \ 0 & 0 & -3 \ 2 & -2 & 1 \ -4 & 0 & 1 \ \end{bmatrix}$$

Then we compute the covariance matrix.

$$m{\Sigma}_{X_c} = rac{1}{N} m{X}_c^T m{X}_c = egin{bmatrix} 6 & 0 & 0 \ 0 & 2 & 0 \ 0 & 0 & 3 \end{bmatrix}$$

Since Σ_{X_c} it is already in a diagonal form we can conclude that $\Lambda = \Sigma_{X_c}$ and $\Gamma = I_3$, and that it holds $\Sigma_{X_c} = \Gamma \Lambda \Gamma^T$. The principal components are the canonical basis vectors.

Project the data to two dimensions, i.e. write down the transformed data matrix $Y \in \mathbb{R}^{4\times 2}$ using the top-2 principal components you computed in (a). What fraction of variance of X is preserved by Y?

The projection matrix is:

$$oldsymbol{\Gamma}_{trunc} = egin{bmatrix} 1 & 0 \ 0 & 0 \ 0 & 1 \end{bmatrix}$$

since we pick the <u>first and the third principal vector</u> corresponding to the two <u>largest eigenvalues</u>. Thus, we have

$$oldsymbol{Y} = oldsymbol{X} oldsymbol{\Gamma}_{trunc} = egin{bmatrix} 2 & 1 \ 0 & -3 \ 2 & 1 \ -4 & 1 \end{bmatrix}$$

We preserve $\frac{6+3}{6+2+3} = \frac{9}{11}$ of the variance.

c Let $x_5 \in \mathbb{R}^3$ be a new data point. Specify the vector x_5 such that performing PCA on the data including the new data point $\{x_i\}_{i=1}^5$ leads to exactly the same principal components as in (a).

Let $x_5 = \bar{x}$, i.e. the <u>new data point</u> equals the <u>mean before including</u> x_5 to the dataset. Therefore, the new mean including x_5 is equal to the old mean. We have:

$$m{X}_c = m{X} - ar{m{x}} = egin{bmatrix} 2 & 2 & 1 \ 0 & 0 & -3 \ 2 & -2 & 1 \ -4 & 0 & 1 \ 0 & 0 & 0 \end{bmatrix}$$

which leads to the same Σ_{X_c} as in (a) up to a difference in the multiplicative constant. In (a) we had $\frac{1}{4}X_c^TX_c$ and here we have $\frac{1}{5}X_c^TX_c$. While this difference leads to different eigenvalues, the eigenvectors and thus the principal components stay the same.

SVD

Problem 5: Use the SVD shown below. Suppose a new user Leslie assigns rating 3 to Alien and

	Matrix	Alien	Star Wars	Casablanca	Titanic
Joe	1	1	1	0	0
Jim	3	3	3	0	0
John	4	4	4	0	0
Jack	5	5	5	0	0
Jill	0	0	0	4	4
Jenny	0	0	0	5	5
Jane	0	0	0	2	2

Figure 11.6: Ratings of movies by users

$$\begin{bmatrix} 1 & 1 & 1 & 0 & 0 \\ 3 & 3 & 3 & 0 & 0 \\ 4 & 4 & 4 & 0 & 0 \\ 5 & 5 & 5 & 0 & 0 \\ 0 & 0 & 0 & 4 & 4 \\ 0 & 0 & 0 & 5 & 5 \\ 0 & 0 & 0 & 2 & 2 \end{bmatrix} = \begin{bmatrix} .14 & 0 \\ .42 & 0 \\ .56 & 0 \\ .70 & 0 \\ 0 & .60 \\ 0 & .75 \\ 0 & .30 \end{bmatrix} \begin{bmatrix} 12.4 & 0 \\ 0 & 9.5 \end{bmatrix} \begin{bmatrix} .58 & .58 & .58 & 0 & 0 \\ 0 & 0 & 0 & .71 & .71 \end{bmatrix}$$

$$M \qquad \qquad U \qquad \qquad \Sigma \qquad \qquad V^{T}$$

rating 4 to Titanic, giving us a representation of Leslie in the 'original space' of [0, 3, 0, 0, 4]. Find the representation of Leslie in concept space. What does that representation predict about how well Leslie would like the other movies appearing in our example data?

The projection is given by $P = M \cdot V$, thus the representation of Leslie in concept space is given by $[0,3,0,0,4] \cdot V = [1.74,2.84]$. It seems that Leslie has a higher preference for "classic" movies (the score is 2.84) such as "Titanic" and "Casablanca" compared to the "sci-fi" movies (the score is 1.74). Thus, since she already saw "Titanic", "Casablanca" would be a reasonable recommendation.

In general, if $\hat{\boldsymbol{U}}$, $\hat{\boldsymbol{\Sigma}}$, $\hat{\boldsymbol{V}}^T$ are the full singular values/vectors of \boldsymbol{M} (obtained by performing full SVD on \boldsymbol{M}) and \boldsymbol{U} , $\boldsymbol{\Sigma}$, \boldsymbol{V}^T are the respective truncated versions (i.e. by taking only the top K singular values/vectors) it holds that the projected data \boldsymbol{P} can be obtained in two alternative and equivalent ways: $\boldsymbol{P} = \boldsymbol{U} \cdot \boldsymbol{\Sigma}$ or $\boldsymbol{P} = \boldsymbol{M} \cdot \boldsymbol{V}$. We usually prefer the second way since we only need to compute the top k singular vectors.

Problem 6: You want to perform linear regression on a data set with features $X \in \mathbb{R}^{N \times D}$ and targets $y \in \mathbb{R}^N$. Assume that you have already computed the SVD of the feature matrix $X = U \Sigma V^T$. Additionally, assume that X has full rank.

Show how we can compute the optimal linear regression weights \boldsymbol{w}^{\star} in $\mathcal{O}(ND)$ operations by using the result of the SVD.

Hint: Matrix operations have the following asymptotic complexity

- Matrix multiplication AB for arbitrary $A \in \mathbb{R}^{P \times Q}$ and $B \in \mathbb{R}^{Q \times R}$ takes $\mathcal{O}(PQR)$
- Matrix multiplication AD for an arbitrary $A \in \mathbb{R}^{P \times Q}$ and a diagonal $D \in \mathbb{R}^{Q \times Q}$ takes $\mathcal{O}(PQ)$
- Matrix inversion C^{-1} for an arbitrary matrix $C \in \mathbb{R}^{M \times M}$ takes $\mathcal{O}(M^3)$
- Matrix inversion D^{-1} for a diagonal matrix $D \in \mathbb{R}^{M \times M}$ takes $\mathcal{O}(M)$

$$\begin{aligned} \boldsymbol{w}^* &= (\boldsymbol{X}^T \boldsymbol{X})^{-1} \boldsymbol{X}^T \boldsymbol{y} \\ &= ((\boldsymbol{U} \boldsymbol{\Sigma} \boldsymbol{V}^T)^T (\boldsymbol{U} \boldsymbol{\Sigma} \boldsymbol{V}^T))^{-1} (\boldsymbol{U} \boldsymbol{\Sigma} \boldsymbol{V}^T)^T \boldsymbol{y} \\ &= (\boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{U} \boldsymbol{\Sigma} \boldsymbol{V}^T)^{-1} \boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{y} \\ &= (\boldsymbol{V} \boldsymbol{\Sigma}^2 \boldsymbol{V}^T)^{-1} \boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{y} \\ &= (\boldsymbol{V}^T)^{-1} (\boldsymbol{\Sigma}^2)^{-1} \boldsymbol{V}^{-1} \boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{y} \\ &= \boldsymbol{V} \boldsymbol{\Sigma}^{-2} \boldsymbol{V}^T \boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{y} \\ &= \boldsymbol{V} \boldsymbol{\Sigma}^{-2} \boldsymbol{V}^T \boldsymbol{V} \boldsymbol{\Sigma} \boldsymbol{U}^T \boldsymbol{y} \\ &= \boldsymbol{V} \boldsymbol{\Sigma}^{-1} \boldsymbol{U}^T \boldsymbol{y} \end{aligned}$$

Multiplication $\boldsymbol{a} = \boldsymbol{U}^T \boldsymbol{y}$ takes $\mathcal{O}(N \cdot D \cdot 1)$

Multiplication $\boldsymbol{b} = \boldsymbol{\Sigma}^{-1} \boldsymbol{a}$ takes $\mathcal{O}(D)$

Multiplication $\boldsymbol{w} = \boldsymbol{V}\boldsymbol{b}$ takes $\mathcal{O}(D^2)$

In total, $\mathcal{O}(ND + D + D^2) = \mathcal{O}(ND)$ if N > D.

Coding

Problem 7: Download the notebook exercise_10_notebook.ipynb from Moodle. Fill in the missing code and run the notebook. Convert the evaluated notebook to pdf and add it to the printout of your homework.

The solution notebook is uploaded on Moodle.