

1.2 Least Squares

$$\hat{\mathcal{R}}(\theta) = \frac{1}{n} \sum_{i=1}^{n} \left(y_i - \langle \alpha_i, \theta \rangle \right)^2 = \frac{1}{n} ||y - X\theta||_2^2$$

- Suppose) 1) $X \in \mathbb{R}^{n \times d}$ matrix
 - 2) full column rank → 멸들이 선병됩니 →
- 2. 입대면 어떻게 구하는데?

1. Linear Regression 3H7+ Z2H6H4?

3. 이게 왜 貂/ 나쁜 해일까?

X^TX invertible possible If not) ≠singular matrix

= 역행결 없음 → 刊世皇十→胡器.

> 그래프가 잘반난다 위에 있는 &

f(y)

Proposition 1: If X has full rank, $\hat{\theta}_{is} = (X^T X)^{-1} X^T Y$

1.2.1 Convex Optimization based on proof of Prop.1

PH? 다른모델(logistic Reg,...) 에도 학장가능

→ Prop 1 을 Convex Optimization 관점에서 광정해보자.

Proposition 2. Optimality Condition (1/21/21/272)

· Definition 1: What 3 坚 convex?

时间唱片就教介于话还是 聖新思 Convex 計好站.

 $f(y) \ge f(x) + \langle y - x, \nabla f(x) \rangle$

 $\langle y-x, \nabla f(x) \rangle + f(x)$

function f is convex & differentiable, 2 x* is global minimizer Consider a point x^* obeying $\nabla f(x^*) = 0$

→ 今日) f(y) ≥ f(x*) + < y-x, ▽f(x*)> =0

つ f(y) ≥ f(x*) □

· Least Square on Tighter.

- function $f(\theta) = \frac{1}{2} || x\theta y ||_{2}^{2}$
- · f(0) is convex, diff-ble
- Gradient) $n \hat{R}(\theta) = \langle X\theta y, X\theta y \rangle$ $= \langle \theta, \chi^T \chi \theta \rangle - 2 \langle \theta, \chi^T y \rangle + \langle y, y \rangle$

$\nabla n \hat{k}(\theta) = 2 X^T X \theta - 2 X^T Y$ $\nabla \hat{R}(\hat{\theta}_{ls}) = 0$ $\Rightarrow 0 = \cancel{2} X^{\mathsf{T}} X \hat{\theta}_{LS} - \cancel{2} X^{\mathsf{T}} Y$

- $\Rightarrow \quad X^{\mathsf{T}} X \ \hat{\theta}_{\mathsf{LS}} = X^{\mathsf{T}} Y$ $\Rightarrow \qquad \hat{\theta}_{\mathsf{LS}} = (X^{\mathsf{T}} X)^{-1} X^{\mathsf{T}} Y \square$

* 2智 理? "gradient = 0 "인 789 >> 84! 2日 2时 7时:

· Linear Algebraic Proof (716+6579) 865 91 9101

- → Y를 XO로 가장 가깝게 근사하는 Projection (직교 투명)
 - · XO는 X의 column vector 위에 존재
 - · 위는 you now note that XO 를 책 나는

- · X ERnxd. has full column rank,
- · U E R^{n×d}, 弘 罗 버터 Column orthogonal
- ∑ ∈ Rdxd, tHythe singular value on, o2,..., od >0

$$\hat{\theta}_{LS} = (X^T X)^{-1} X^T y$$

• $X^T X = (U \Sigma V^T)^T (U \Sigma V^T) = (V \Sigma^T U^T) (U \Sigma V^T)$

 $= V \Sigma^{T} \Sigma V^{T} = V \Sigma^{2} V^{T}$

 \Rightarrow $(X^T X)^{-1} = (V \sum^2 V^T)^{-1} = V \sum^{-2} V^T$

• $X^T y = (U \Sigma V^T)^T y = V \Sigma^T U^T$

 $\Rightarrow \hat{\theta}_{LS} = V \Sigma^{-2} V^{\mathsf{T}} . V \Sigma^{\mathsf{T}} U^{\mathsf{T}} y = V \Sigma^{-4} U^{\mathsf{T}} y$

• $Y\hat{\theta}_{LS} = U\Sigma V^T \cdot V\Sigma^{-1}U^Ty = UU^Ty$

XÔLS = UUTY 9 elo) UUTE X9 colum space on chtole projectiou watrik

 $X = U \Sigma V^{\tau}$

Ridge Regression			
(dea) LROIH ÔLS 가 라하게 변하게 가하기 의하기 의 기계			기를 해졌더니 우등이 세를 좀 더 게임하네 안녕적인 추정을 사내고나
		•	
2.1 Ridge Regression Estimate			λ>0: reg. parameter
· θ̂LS = argmin y-Xθ (2 (713)			(3°C G 26L Mg)
· 이수 큰 IIBII는 메슥은 불어내 해. > 11811를 건트들해			White $\hat{\theta}_{\text{Higgs}} = \operatorname{argmin} \ y - X\theta \ _2^2 + \lambda \ \theta\ _2^2$
		→ 110(1 on 24th IH	JEI O TOTAL
· 어전히 convex, diff-able > 3 closed-form fy.			Attiny error regularization penalty
$\hat{\theta}_{\text{ridge}} = (X^T X + \lambda I)^{-1} X^T Y$			
			$X\hat{\theta}_{rigde} = X(X^TX + \lambda I)^{-1}X^Ty$
SVD기반생덕) Shrinkage 관점.			$= U \Sigma V^{T} \left(V \underline{\Sigma}^{T} \underline{U}^{T} \underline{U} \underline{\Sigma} V^{T} + \lambda V V^{T} \right)^{-1} V \Sigma U^{T} Y$
· X毫 SVD, X= UΣVT.			$= U\Sigma V^{T} \left(V\Sigma^{2} V^{T} + \lambda V V^{T} \right)^{-1} V\Sigma U^{T} Y$
$\hat{\nabla} \hat{\theta} = \sum_{i=1}^{d} \frac{\sigma_{i}^{2}}{\sigma_{i}^{2} + \lambda} \langle y, u_{i} \rangle u_{i}$			$= U \Sigma V^{T} \left(V (\Sigma^2 + \lambda \mathbf{I}) V^{T} \right)^{-1} V \Sigma U^{T} Y$
			$= U\Sigma V^{T} V (\Sigma^2 + \lambda \mathbf{I})^{-1} V^{T} V \Sigma U^{T} Y$
shrink factor ∈ (0.1)			$= U\Sigma(\Sigma^2 + \lambda D)^{-1}\Sigma U^{T} Y$
> 해灯 각 양병 Ui oil 대해 Shrink factor를 급한다.			
平, data 7+ \$\$\$ (3号 (3号 (5)) MH는 bias 是 Bol 全时			diagonal matrix with $\sigma_i^2 + \lambda$ $\Rightarrow (\Sigma^2 + \lambda I)^{-1} = \text{diagonal matrix with } \frac{1}{\sigma_{i+\lambda}^2}$
→ 40년에 덜 만남 (더 경육하나가)			
> 1 2 20 (of John (b)			• $\Sigma = \text{diagond matrix with } \tau_i$
2.2 Page Halange Tall 18			$\Rightarrow \Sigma(\Sigma^2 + \lambda I)^{-1}\Sigma = diag\left(\frac{\sigma_1^2}{\sigma_1^2 + \lambda}, \dots, \frac{\sigma_1^2}{\sigma_1^2 + \lambda}\right)$
2.2 Bias - Variance - Trade off			$\Rightarrow uu^{T}y = \sum u_i \langle u_i, y \rangle = \sum \langle y, u_i \rangle u_i$
· Ridge Reg. 쿠팅likol bias를 일부고입 > Variance 쿨링.			To the state of th
factor	Least Squares	Ridge Requession	. \=0: \text{9Ht LS (Bias 0 . Var 1)}
Bias	(ow (o)	high	권적 2 선택방법 · 2→∞: β ridge → 0 (국단적 단당씨)
Variance	大	/J\	· LE Closs - validation Als
ranance 떠날 Lyr	·	-	TA CIOSS - NIMI WALLON 1.1.2
- I S LXF	overfitting possible	CALLED MAN V	

Optimization: Maximum Likelihood Estimate

- A method of estimating the parameters of a statistical model given observations by finding the parameter values that maximize the likelihood of making the observations given the parameters

 $\Theta_{ML} = \partial \Theta_{0}^{MC} \circ P_{model}(y|X,\theta)$
- Approach

$$\theta_{ML} = \underset{\theta}{arg \max} \prod_{i=1}^{n} p_{model}(y_i|x_i,\theta) \qquad \text{ Some rated by same probability distribution}$$

Podata (y|X): true underlying distribution (observation)

• We can replace the product by applying the
$$\frac{\log \operatorname{rithmic}}{\log \operatorname{c}(ab)} = \log_{\operatorname{c}}(a) + \log_{\operatorname{c}}(b)$$
:

$$\theta_{ML} = \arg\max_{\theta} \sum_{i=1}^{n} \log \frac{\operatorname{p}_{model}(y_{i}|\boldsymbol{x}_{i}, \theta)}{\operatorname{p}_{model}(y_{i}|\boldsymbol{x}_{i}, \theta)} \qquad \text{if } \frac{\operatorname{vi}(\boldsymbol{x}_{i}, \theta, \sigma^{\star})}{\operatorname{equation}} = \frac{\operatorname{vi}(\theta + \operatorname{vi}(\theta, \sigma^{\star}))}{\operatorname{equation}} = \frac{\operatorname{vi}(\theta + \operatorname{vi}(\theta, \sigma^{\star}))}{\operatorname{e$$

• Assuming Gaussian distribution, we get the same result for the optimization as for Linear least squares

$$\boldsymbol{\theta} = (\boldsymbol{X^TX})^{-1}\boldsymbol{X^Ty}$$

+
$$P(y:|x_i,\theta) = (2\pi\sigma^2)^{\frac{4}{3}} \cdot e^{-\frac{\lambda}{2\sigma^2}(y_i-x_i,\theta)^2}$$

Assuming
$$y_i = \mathcal{N}(x;\theta,\sigma^2) = x_i\theta + \mathcal{N}(0,\sigma^2)$$

of) Gaupian:
$$p(y_i) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x_i}{2\sigma^2}(y_i - \mu)^2} \quad y_i \sim \mathcal{N}(\mu,\sigma^2)$$

original optimization problem

$$\theta_{ML} = \underset{\theta}{\text{arg max}} \sum_{i=0}^{N} \underset{\theta}{\text{log pmodel}} (y_i | x_i, \theta)$$

$$\underset{\theta}{\text{log}} \left[(2\pi\sigma^2)^{-\frac{1}{2}} e^{-\frac{\lambda}{2\sigma^2}(y_i - x_i; \theta)^2} \right]$$

$$\sum_{i=1}^{n} {\binom{o}{g}} \left[(2\pi\sigma^{2})^{-\frac{1}{2}} \cdot e^{-\frac{\lambda}{2\sigma^{2}} (y_{i} - x_{i} \theta)^{2}} \right]$$

$$= \sum_{i=1}^{n} -\frac{1}{2} \log(2\pi\sigma^{2}) + \sum_{i=1}^{n} \left(-\frac{1}{2\sigma^{2}}\right) (\gamma_{i} - \kappa_{i}\theta)^{2}$$

$$= -\frac{N}{2} \left(\cos \left(2\pi \sigma^2 \right) - \frac{1}{2\sigma^2} \left(y_i - X\theta \right)^T (y - X\theta) \right)$$

$$\left(\frac{\partial J(\theta)}{\partial x_i} \stackrel{?}{\to} 0 \right) \Rightarrow 0 = \left(\frac{X^T X^{-1} X^T}{2\sigma^2} \right)$$

$$\oint \frac{\partial \theta}{\partial x} = 0 \quad \Rightarrow \quad \theta = (X_{\perp} X)_{-1} X_{\perp} A$$

अर्धि अर्थित ति तिक्षित्र देशा अर्थि हेशा है है ।

→ व्यम् अन्ति भूतिकान प्रमाह त्या क्रिका र

- @ Least squares @ Maximum Likelihood
- (7/6/65-2-2)
- · 정답 8는 로스지만
- · 0를 写에서 阿勢맛 ŷ;= X; 0이 (3m) 맛 y; 다

उत्पर सहिन्त्र रूप.

े हमें सिंधान ये प्रधा गरे से मा best

* (nterpretation*

११(y)न धन्म माह्य भगटियां गरिट छे रूप

2 Maximum Likelihood

· data = 整 model out Uffe TOP. y;~ N(0x; 52)

· "이전 연를 사명하는 다면 사건(yi)이 나올 학률 "이 가상 높은 결 갖자

$$\hat{\theta}_{ME} = argmax \prod_{i=1}^{n} \mathcal{N}(y_i | \theta x_i, \sigma^2)$$

$$P(A(B) = \frac{P(A\cap B)}{P(B)}$$
 ⇒ B가 일이낮을 Œ내 그것이너 A가 일이날 학률

* Interpretation *

지금 막 본 결과들이 실제로 나를 맞춤이 기참 높은 0를 찾자. = 기참 2절등한 0를 찾자

· Why? Y=ax+b

LS가 식민적인데 "라 그 용나는 사망하는지"의 근거 부족 MLE는 "data가 라 이렇게 나타는가 "를 한국되고 연명가능

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3.3 Bias-Trade-Off Ridge Requession Of PH 일반하 告이 器시 유하적으로 얼아버고다.
                                                           어느 Phr 을 Bias2+ Variance + Noise 3 분배해서 어떻게 Variance를 클러는게 분배비자
  · 문제 설명 - Prediction Risk of ô
 →M32 H를 X ∈ Rd on that only by tour
                                                                                              (예술 2123)
     · Model: y=h(x)+z, Z~N(0,02)
                                                                                                 Prediction Risk R(\hat{h}) = E_{x,y} [(\hat{h}(x) - y)^2]
     · Dataset: D= { (y1, x1), -- (yn, xn)}
                                                                                                  + Dostaset
                                                                                                        randour
                                                                                                                             \Rightarrow E_{\mathbf{p}}[R(\hat{\mathbf{h}}_{\mathbf{p}})] = E_{\mathbf{x},\mathbf{y},\mathbf{p}}[(\hat{\mathbf{h}}_{\mathbf{p}}(\mathbf{x}) - \mathbf{y})^{2}]
     • 阿乌尔: \hat{y} = \hat{h}(x) = \langle \hat{x}, \hat{\theta}_{nidae} \rangle
                                                                                                     (強強)
     • 日間: Y= <x,0*>+2, そ~り(0,0)
  · Goal: Exy.D & Bias Variance, Noise ? Hot
                                                                           y= h(x)+2
  · 44)
   E_{\mathcal{D}}[R(\hat{h}_{\mathcal{D}})] = E_{\mathcal{X},\mathcal{Y},\mathcal{D}}[(\hat{h}_{\mathcal{D}}(\mathcal{X}) - \mathcal{Y})^2] \stackrel{*}{=} E_{\mathcal{D}}[(\hat{h}_{\mathcal{D}}(\mathcal{X}) - h(\mathcal{X}) - \mathcal{Z})^2]
                                                                             = \mathbb{E}_{\mathbb{P}} \left[ \left( \widehat{h}_{\mathbb{P}}(\mathbb{X}) - h(\mathbb{X}) \right)^{2} \right] + 2 \mathbb{E} \left[ \left( \widehat{h}_{\mathbb{P}}(\mathbb{X}) - h(\mathbb{X}) \right) \mathcal{Z} \right] + \mathbb{E} \left[ \mathcal{Z}^{2} \right]
                                                                             = \mathbb{E}_{\mathcal{D}} \left[ \left( \hat{h_{\mathcal{D}}}(X) - h(X) \right)^{2} \right] + \mathbb{E} \left[ \mathbb{E}^{2} \right]
                                                                                                                                                  Z has zero mean.
    \mathsf{E}_{\mathsf{D}}\Big[\big(\hat{\mathsf{h}}_{\mathsf{D}}(\mathsf{x}) - \mathsf{h}(\mathsf{x})\big)^{2}\Big] = \mathsf{E}_{\mathsf{D}}\Big[\big(\hat{\mathsf{h}}_{\mathsf{D}}(\mathsf{2}) - \mathsf{E}\big[\hat{\mathsf{h}}_{\mathsf{D}}(\mathsf{x})\big] + \mathsf{E}\big[\hat{\mathsf{h}}_{\mathsf{D}}(\mathsf{x})\big] + \mathsf{h}(\mathsf{x})\big)^{2}\Big]
                                              = \mathbb{E}_{\mathcal{D}} \left[ \left( h_{\mathcal{D}}(\mathbf{x}) - \mathbb{E} \left[ \hat{h}_{\mathcal{D}}(\mathbf{x}) \right]^{2} \right] - 2 \left( \mathbb{E}_{\mathcal{D}} \left[ \hat{h}_{\mathcal{D}}(\mathbf{x}) \right] - \mathbb{E}_{\mathcal{D}} \left[ h(\mathbf{x}) \right] \right) \left( \mathbb{E} \left[ \hat{h}_{\mathcal{D}}(\mathbf{x}) - h(\mathbf{x}) \right) + \mathbb{E} \left[ \left( \mathbb{E} \left[ \hat{h}_{\mathcal{D}}(\mathbf{x}) \right] - h(\mathbf{x}) \right)^{2} \right] 
                                                                                                           = D, :: 기약값 # 53년
                                               =\mathbb{E}_{\mathcal{D}}\Big[\big(h_{\mathsf{D}}(2)-\mathbb{E}\big[\hat{h_{\mathsf{D}}}(\mathcal{X})\big]^2\Big]+\mathbb{E}\Big[\big(\mathbb{E}\big[\hat{h_{\mathsf{D}}}(\mathcal{X})\big]-h(\mathcal{X})\big)^2\Big]
   Thus we have,
    E_D[R(\hat{h})] = E_D[(\hat{h}_D(x) - h(x))^2] + E[z^2]
                           = E \left[ \left( E \left[ h_{D}(x) \right] - h(x) \right)^{2} \right] + E_{D} \left[ \left( h_{D}(2) - E \left[ h_{D}(x) \right] \right)^{2} \right] + E \left[ Z^{2} \right]
                                              Bias<sup>2</sup>
                                                                                                Variance
                                                                                                                                          Noise
                               = 写起了图2+ 25xx+站个h(x) 站Dataset 电影则 四是
                                                                                                                                           明朝外的
                                  एव भय
                                                                                         어느의 환원
米型强烈0的111
     · Bias7+3ct > 모델이 단문하片 가까나 파턴 吴 四年十 (underfitting)
     · Variance 7+ 五十 > GIPIET 快口 强 HPIPE 어떻이 크게 WHA (overfitting)
    · Noise는 叫起 午 部
 米대 강화하다? 오덴 선택& Hyperparameter 튜닝의 근거.
                             Next Chapter
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