

# Tutorial 2: Maximum-Likelihood and Bayesian Parameter Estimation

Rui Zhao

[rzhao@ee.cuhk.edu.hk](mailto:rzhao@ee.cuhk.edu.hk)

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- 1 ML estimate:  $\hat{\Sigma}$  of multivariate Gaussian
- 2 ML estimate bias:  $\hat{\sigma}^2$  of univariate Gaussian
- 3 Bayesian estimate: Univariate Gaussian
  - brief review
  - posteriori  $p(\mu|\mathcal{D})$
  - conditional probability density  $p(x|w_i, \mathcal{D}_i)$

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# 1. ML estimate: $\hat{\Sigma}$ of multivariate Gaussian

## Multivariate Gaussian Case: unknown $\mu$ and $\Sigma$

$$\hat{\mu} = \frac{1}{n} \sum_{k=1}^n \mathbf{x}_k \quad (1)$$

$$\hat{\Sigma} = \frac{1}{n} \sum_{k=1}^n (\mathbf{x}_k - \hat{\mu})(\mathbf{x}_k - \hat{\mu})^t \quad (2)$$

- $\hat{\mu}$  is the sample mean.
- $\hat{\Sigma}$  is the arithmetic average of the  $n$  matrices  $(\mathbf{x}_k - \hat{\mu})(\mathbf{x}_k - \hat{\mu})^t$ .

# 1. ML estimate: $\hat{\Sigma}$ of multivariate Gaussian

## Multivariate normal density

$$p(\mathbf{x} \mid \mu, \Sigma) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[ -\frac{1}{2} (\mathbf{x} - \mu)^t \Sigma^{-1} (\mathbf{x} - \mu) \right] \quad (3)$$

Draw  $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n$  independently from  $p(\mathbf{x} \mid \mu, \Sigma)$ , and the joint density (likelihood) is:

$$p(\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n \mid \mu, \Sigma) = \quad (4)$$

$$\frac{1}{(2\pi)^{nd/2} |\Sigma|^{n/2}} \exp \left[ -\frac{1}{2} \sum_{k=1}^n (\mathbf{x}_k - \mu)^t \Sigma^{-1} (\mathbf{x}_k - \mu) \right] \quad (5)$$

Log-likelihood  $l(\mu, \Sigma)$  is

$$l(\mu, \Sigma) = -\frac{nd}{2} \ln(2\pi) - \frac{n}{2} \ln |\Sigma| - \frac{1}{2} \sum_{k=1}^n (\mathbf{x}_k - \mu)^t \Sigma^{-1} (\mathbf{x}_k - \mu) \quad (6)$$

# 1. ML estimate: $\hat{\Sigma}$ of multivariate Gaussian

## Multivariate normal density

Let  $\mathbf{A} = \Sigma^{-1}$

$$l(\mu, \Sigma) = -\frac{nd}{2} \ln(2\pi) - \frac{n}{2} \ln |\Sigma| - \frac{1}{2} \sum_{k=1}^n (\mathbf{x}_k - \mu)^t \Sigma^{-1} (\mathbf{x}_k - \mu) \quad (7)$$

$$l(\mu, \Sigma) = -\frac{nd}{2} \ln(2\pi) + \frac{n}{2} \ln \mathbf{A} - \frac{1}{2} \sum_{k=1}^n (\mathbf{x}_k - \mu)^t \mathbf{A} (\mathbf{x}_k - \mu) \quad (8)$$

$$\frac{\partial l(\mu, \Sigma)}{\partial \mathbf{A}} = \frac{n}{2} \mathbf{A}^{-1} - \frac{1}{2} \sum_{k=1}^n (\mathbf{x}_k - \mu)(\mathbf{x}_k - \mu)^t = 0 \quad (9)$$

Replace  $\mathbf{A}$  by  $\Sigma^{-1}$

$$\hat{\Sigma} = \frac{1}{n} \sum_{k=1}^n (\mathbf{x}_k - \hat{\mu})(\mathbf{x}_k - \hat{\mu})^t, \quad (\hat{\mu} = \frac{1}{n} \sum_{k=1}^n \mathbf{x}_k) \quad (10)$$

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# ML estimate bias: $\hat{\sigma}^2$ of univariate Gaussian

## Univariate Gaussian Case

ML estimator  $\hat{\sigma}^2$  is biased.

$$E[\hat{\sigma}^2] = E\left[\frac{1}{n} \sum_{k=1}^n (x_k - \hat{\mu})^2\right] = \frac{n-1}{n} \sigma^2 \quad (11)$$

An elementary unbiased estimator for  $\sigma^2$  is given by  $\frac{1}{n-1} \sum_{k=1}^n (x_k - \hat{\mu})^2$ .

# ML estimate bias: $\hat{\sigma}^2$ of univariate Gaussian

## Univariate Gaussian Case

$$E\left[\frac{1}{n} \sum_{k=1}^n (x_k - \hat{\mu})^2\right] = E\left[\frac{1}{n} \sum_{k=1}^n \left(x_k - \frac{1}{n} \sum_{j=1}^n x_j\right)^2\right] \quad (12)$$

$$= E\left[\frac{1}{n} \sum_{k=1}^n \left(x_k^2 - \frac{2}{n} x_k \sum_{j=1}^n x_j + \frac{1}{n^2} \left(\sum_{j=1}^n x_j\right)^2\right)\right] \quad (13)$$

$$= E\left[\frac{1}{n} \left(\sum_{k=1}^n x_k^2 - \frac{2}{n} \left(\sum_{k=1}^n x_k\right)^2 + \frac{n}{n^2} \left(\sum_{k=1}^n x_k\right)^2\right)\right] \quad (14)$$

$$= E\left[\frac{1}{n} \left(\sum_{k=1}^n x_k^2 - \frac{1}{n} \left(\sum_{k=1}^n x_k\right)^2\right)\right] \quad (15)$$

$$= \frac{1}{n} E\left[\sum_{k=1}^n x_k^2\right] - \frac{1}{n^2} E\left[\left(\sum_{k=1}^n x_k\right)^2\right] \quad (16)$$

$$= E[x^2] - \frac{1}{n^2} E\left[\sum_{k=1}^n x_k^2 + \sum_{i \neq j} x_i x_j\right] \quad (17)$$

$$= E[x^2] - \frac{1}{n^2} E\left[\sum_{k=1}^n x_k^2\right] - \frac{1}{n^2} E\left[\sum_{i \neq j} x_i x_j\right] \quad (18)$$

$$= E[x^2] - \frac{1}{n} E[x^2] - \frac{n^2 - n}{n^2} E[x_i x_j] = \frac{n-1}{n} \left(E[x^2] - (E[x])^2\right) \quad (19)$$

$$= \frac{n-1}{n} \sigma^2 \quad (20)$$

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