

# RMT

## LECTURE 5

### HYPOTHESIS TESTING

2023/2024



Denis Marinšek

# HYPOTHESIS TESTING

# HYPOTHESIS TESTING

1. Research Question: Is there a difference in the share of online banking users between younger and older generations?
  2. Research Hypothesis: A higher percentage of younger generations use online banking compared to older generations.
  3. Collect the data, find the appropriate statistical test, and see if the data supports your hypothesis.
- Hypothesis can be tested with parametric or with nonparametric statistical tests. What is the difference?
  - When to use non-parametrical tests:
    - If assumptions of parametrical tests are violated (for example assumption about normal distribution) or if there are serious problems of outliers.
    - If variables are measured on ordinal measurement scale.

# HYPOTHESIS TESTING

Parametrical test		Alternative non-parametrical test
Hypothesis about the population arithmetic mean	<i>t</i> -test Effect size: CohenD – Sawilowsky, 2009	Wilcoxon Signed Rank Test (at least interval variables) – Hypothesis about the population median Effect size: Biserial Correlation – Funder, 2019
Hypothesis about the difference between two population arithmetic means (dependent or paired samples)	Paired <i>t</i> -test Effect size: CohenD – Sawilowsky, 2009	Sign Test (ordinal variables) – Special version of Binomial test  Wilcoxon Signed Rank Test (at least interval variables) Effect size: Biserial Correlation – Funder, 2019
Hypothesis about the difference between two population arithmetic means (independent samples)	Independent <i>t</i> -test with Welch correction Effect size: CohenD – Sawilowsky, 2009	Wilcoxon Rank Sum Test, sometimes called Mann-Whitney U-test (at least ordinal variables) Effect size: Biserial Correlation – Funder, 2019
Hypothesis about the equality of three or more population arithmetic means for dependent samples	One-way repeated measures ANOVA, rANOVA Effect size: $\eta^2$ - Cohen, 1992	Friedman ANOVA (at least ordinal variables) Effect size: Kendall W - Landis, 1977

# HYPOTHESIS TESTING

Parametrical test		Alternative non-parametrical test
Hypothesis about the equality of three or more population arithmetic means for independent samples	<i>One-way ANOVA</i> Effect size: $\eta^2$ - Cohen, 1992  <i>Welch Heteroscedastic F-test</i> – robustness check	<i>Kruskal-Wallis Rank Sum Test</i> (at least ordinal variables) Effect size: $\eta^2$ - Field, 2013
Hypothesis about the population proportion	$\chi^2$ -test (Test of proportion)	Binomial test
Hypothesis about association between two categorical variables	$\chi^2$ -test Effect size: CramerV statistics (Funder, 2019) or Odds ratio (Chen et al., 2010)	Fisher's Exact Probability Test
Hypothesis about population distribution for given probabilities	$\chi^2$ -test	
Hypothesis about the equality in two population proportions	$\chi^2$ -test (Test of proportion)	

# HYPOTHESIS TESTING

Sawilowsky (2009) ("sawilowsky2009")

- $d < 0.1$  - Tiny
- $0.1 \leq d < 0.2$  - Very small
- $0.2 \leq d < 0.5$  - Small
- $0.5 \leq d < 0.8$  - Medium
- $0.8 \leq d < 1.2$  - Large
- $1.2 \leq d < 2$  - Very large

Funder & Ozer (2019) ("funder2019"; default)

- $r < 0.05$  - Tiny
- $0.05 \leq r < 0.1$  - Very small
- $0.1 \leq r < 0.2$  - Small
- $0.2 \leq r < 0.3$  - Medium
- $0.3 \leq r < 0.4$  - Large
- $r \geq 0.4$  - Very large

Landis & Koch (1977) ("landis1977"; default)

- $0.00 \leq w < 0.20$  - Slight agreement
- $0.20 \leq w < 0.40$  - Fair agreement
- $0.40 \leq w < 0.60$  - Moderate agreement
- $0.60 \leq w < 0.80$  - Substantial agreement
- $w \geq 0.80$  - Almost perfect agreement

Cohen (1992) ("cohen1992") applicable to one-way anova, or to *partial eta* / omega / epsilon squared in multi-way anova.

- $ES < 0.02$  - Very small
- $0.02 \leq ES < 0.13$  - Small
- $0.13 \leq ES < 0.26$  - Medium
- $ES \geq 0.26$  - Large

Chen et al. (2010) ("chen2010"; default)

- $OR < 1.68$  - Very small
- $1.68 \leq OR < 3.47$  - Small
- $3.47 \leq OR < 6.71$  - Medium
- $**OR \geq 6.71 **$  - Large

# HYPOTHESIS ABOUT THE POPULATION ARITHMETIC MEAN

# HYPOTHESIS ABOUT THE POPULATION MEDIAN

# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

Testing the population arithmetic mean with the chosen parameter.

$$H_0: \mu = \mu_0$$

$$H_1: \mu \neq \mu_0$$

$$t = \frac{\bar{y} - \mu_0}{se(\bar{y})}$$

$$se(\bar{y}) = \frac{s_y}{\sqrt{n}} \quad df = n - 1$$

Conditions and assumptions:

- Variable is numeric.
- Normality - variable on the population is normally distributed.
- No outliers.

# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

## Example: Body mass

A researcher examined the effects of online schooling on the body weight of ninth graders. Based on data from the 2018/2019 school year, the average weight was 59.5 kg and the median was 58.3 kg. At the beginning of the 2021/2022 school year, the researchers randomly selected 50 ninth graders and measured their weight. What conclusion can the researcher draw?

```
mydata <- read.table("./Body mass.csv", header=TRUE, sep=";", dec=",")  
head(mydata)
```

```
##   ID Mass  
## 1  1 62.1  
## 2  2 64.5  
## 3  3 56.5  
## 4  4 53.4  
## 5  5 61.3  
## 6  6 62.2
```

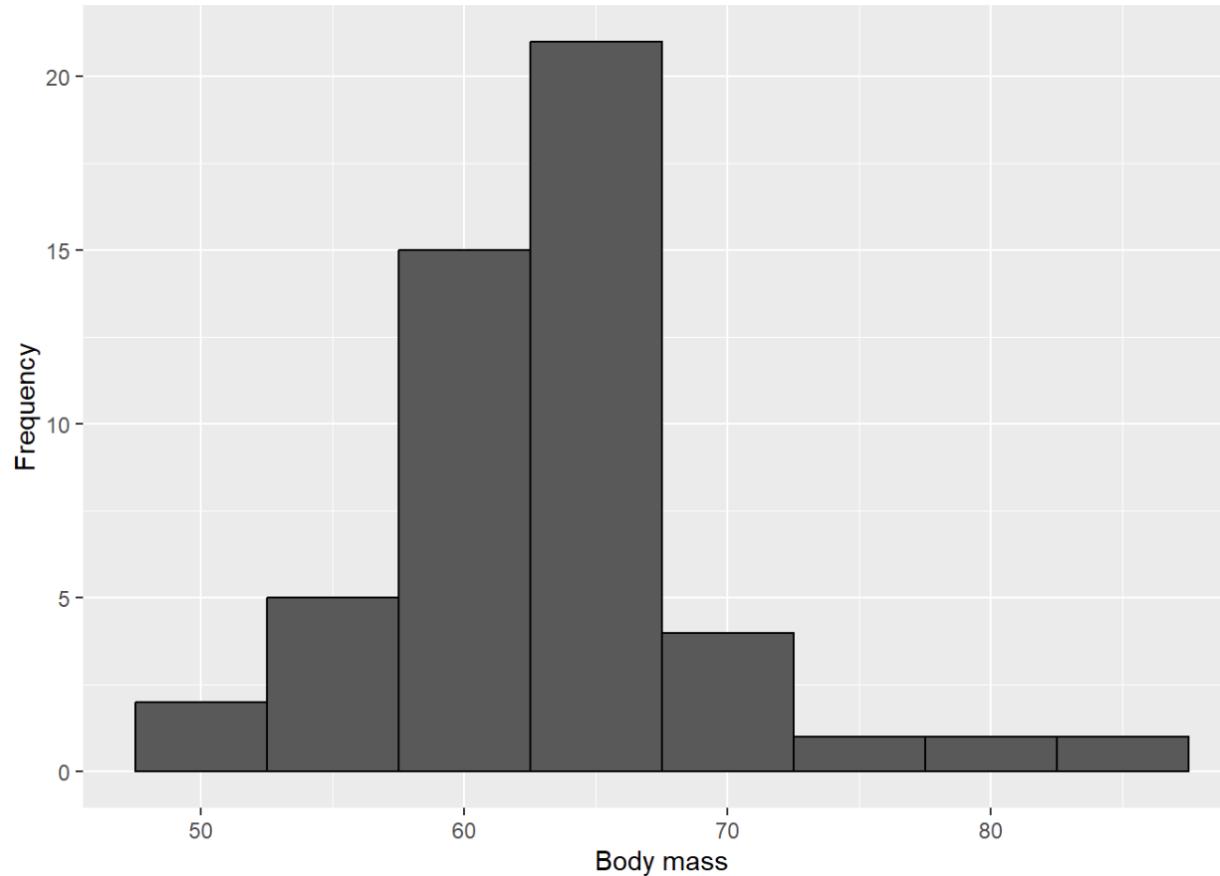
Description of data:

- ID: Identifier.
- Mass: Body mass in kg.

# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

## Example: Body mass

```
library(ggplot2)
ggplot(mydata, aes(x = Mass)) +
  geom_histogram(binwidth = 5, colour = "black") +
  ylab("Frequency") +
  xlab("Body mass")
```



# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

## Example: Body mass

```
mean(mydata$Mass)
```

```
## [1] 62.876
```

```
sd(mydata$Mass)
```

```
## [1] 6.011403
```

```
nrow(mydata)
```

```
## [1] 50
```

# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

## Example: Body mass

```
ybar = mean(mydata$Mass); sd = sd(mydata$Mass); n = nrow(mydata)

se = sd/sqrt(n)

ybar_lower_t = ybar + qt(0.025, df=n-1)*se
ybar_upper_t = ybar + qt(0.975, df=n-1)*se
```

95%-confidence interval:  $61.168 < \mu < 64.584$

```
t.test(mydata$Mass,
       mu = 59.5,
       alternative = "two.sided")
```

```
##
## One Sample t-test
##
## data: mydata$Mass
## t = 3.9711, df = 49, p-value = 0.000234
## alternative hypothesis: true mean is not equal to 59.5
## 95 percent confidence interval:
## 61.16758 64.58442
## sample estimates:
## mean of x
## 62.876
```

# HYPOTHESIS ABOUT THE ARITHMETIC MEAN

## Example: Body mass

```
#install.packages("effectsize")
library(effectsize)

effectsize::cohens_d(mydata$Mass, mu = 59.5)

## Cohen's d |      95% CI
## -----
## 0.56    | [0.26, 0.86]
##
## - Deviation from a difference of 59.5.

interpret_cohens_d(0.56, rules = "sawilowsky2009")

## [1] "medium"
## (Rules: sawilowsky2009)
```

$$d = \frac{\bar{y} - \mu_0}{s_y} = \frac{62.876 - 59.5}{6.011} = 0.56$$

Sawilowsky (2009) ("sawilowsky2009")

- **d < 0.1** - Tiny
- **0.1 <= d < 0.2** - Very small
- **0.2 <= d < 0.5** - Small
- **0.5 <= d < 0.8** - Medium
- **0.8 <= d < 1.2** - Large
- **1.2 <= d < 2** - Very large

Conclusions:

Based on the sample data, we found that the average weight of ninth graders at the beginning of the 2021/2022 school year was 62.88 kg and has increased compared to the 2018/2019 school year ( $p < 0.001$ ,  $d = 0.56$  – medium sized effect).

# HYPOTHESIS ABOUT THE MEDIAN

Testing the population median with the chosen parameter.

$$H_0: Me = Me_0$$

$$H_1: Me \neq Me_0$$

$$z = \frac{T^+ - \bar{T}}{SE_T}$$

$$\bar{T} = \frac{n(n + 1)}{4}$$

$$SE_T = \sqrt{\frac{n(n + 1)(2n + 1)}{24}}$$

Requirement:

- Variable must be numeric.

# HYPOTHESIS ABOUT THE MEDIAN

## Example: Body mass

```
median(mydata$Mass)
```

```
## [1] 62.8
```

```
wilcox.test(mydata$Mass,  
            mu = 58.3,  
            correct = FALSE)
```

```
##  
## Wilcoxon signed rank test  
##  
## data: mydata$Mass  
## V = 1116, p-value = 3.847e-06  
## alternative hypothesis: true location is not equal to 58.3
```

# HYPOTHESIS ABOUT THE MEDIAN

## Example: Body mass

```
effectsize(wilcox.test(mydata$Mass,
  mu = 58.3,
  correct = FALSE))
```

```
## r (rank biserial) |      95% CI
## -----
## 0.75            | [0.58, 0.86]
##
## - Deviation from a difference of 58.3.
```

```
interpret_rank_biserial(0.75, rules = "funder2019")
```

```
## [1] "very large"
## (Rules: funder2019)
```

Funder & Ozer (2019) ("funder2019"; default)

- **r < 0.05** - Tiny
- **0.05 <= r < 0.1** - Very small
- **0.1 <= r < 0.2** - Small
- **0.2 <= r < 0.3** - Medium
- **0.3 <= r < 0.4** - Large
- **r >= 0.4** - Very large

Conclusions:

Based on the sample data, we found that the median weight of ninth graders at the beginning of the 2021/2022 school year was 62.8 kg and has increased compared to the 2018/2019 school year ( $p < 0.001$ ,  $r = 0.75$  – very large effect).

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO POPULATION ARITHMETIC MEANS

SIGN TEST

WILCOXON SIGNED RANK TEST

WILCOXON RANK SUM TEST

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## *t*-test

*t*-test for repeated measures (Paired samples *t*-test):

Pairs of data belong to the same unit, measured at different conditions.

*t*-test for independent measures (Independent samples *t*-test):

Data belong to two different groups of units (from two different populations). Each unit is measured once.

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Assumptions

Independent Samples  $t$ -test and Paired Samples  $t$ -test are parametric tests:

- Variable is numeric.
- Normality:
  - Paired Samples  $t$ -test: Differences on the population are normally distributed.
  - Independent Samples  $t$ -test: The distribution of the variable is normal in both populations.
- For independent Samples  $t$ -test:
  - The data must come from two independent populations,
  - Variable has the same variance in both populations – since this assumption is often violated, we apply Welch correction.

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples $t$ -test

$$H_0: \mu_1 = \mu_2 \quad \text{or} \quad \mu_d = 0$$

$$H_1: \mu_1 \neq \mu_2 \quad \text{or} \quad \mu_d \neq 0$$

$$t = \frac{\bar{d} - \mu_d}{se(\bar{d})}$$

$$d_i = y_{i1} - y_{i2} \quad \bar{d} = \frac{1}{n} \sum_{i=1}^n d_i; \quad se(\bar{d}) = \frac{s_d}{\sqrt{n}};$$

$$s_d = \sqrt{s_d^2}$$

$$s_d^2 = \frac{1}{n-1} \sum_{i=1}^n (d_i - \bar{d})^2$$

$$df = n - 1$$

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Independent samples $t$ -test ( $\sigma_1^2 \neq \sigma_2^2$ )

$$H_0: \mu_1 = \mu_2 \quad \text{or} \quad \mu_1 - \mu_2 = 0$$

$$H_1: \mu_1 \neq \mu_2 \quad \text{or} \quad \mu_1 - \mu_2 \neq 0$$

$$t = \frac{\bar{y}_1 - \bar{y}_2}{se(\bar{y}_1 - \bar{y}_2)}$$

$$se(\bar{y}_1 - \bar{y}_2) = \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

$$df = \frac{\left( \frac{s_1^2}{n_1} + \frac{s_2^2}{n_2} \right)^2}{\left( \frac{s_1^2}{n_1} \right)^2 \frac{1}{n_1 - 1} + \left( \frac{s_2^2}{n_2} \right)^2 \frac{1}{n_2 - 1}}$$

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples $t$ -test – Example 1: Literacy

In elementary school, it was decided to test students' literacy skills. 92 sixth grade students were randomly selected, and each of them took a simple test on reading comprehension (maximum 100 points) and writing knowledge (maximum 100 points). Can we say that the average scores obtained in reading comprehension and written knowledge are different? If so, where did the students perform better?

```
mydata <- read.table("./Literacy.csv", header=TRUE, sep=";", dec=",")  
head(mydata)
```

```
##   ID Reading Writing  
## 1  1    81.50   81.44  
## 2  2    85.25   73.27  
## 3  3    86.88   84.24  
## 4  4    88.68   73.16  
## 5  5    77.30   80.45  
## 6  6    87.53   70.79
```

Description:

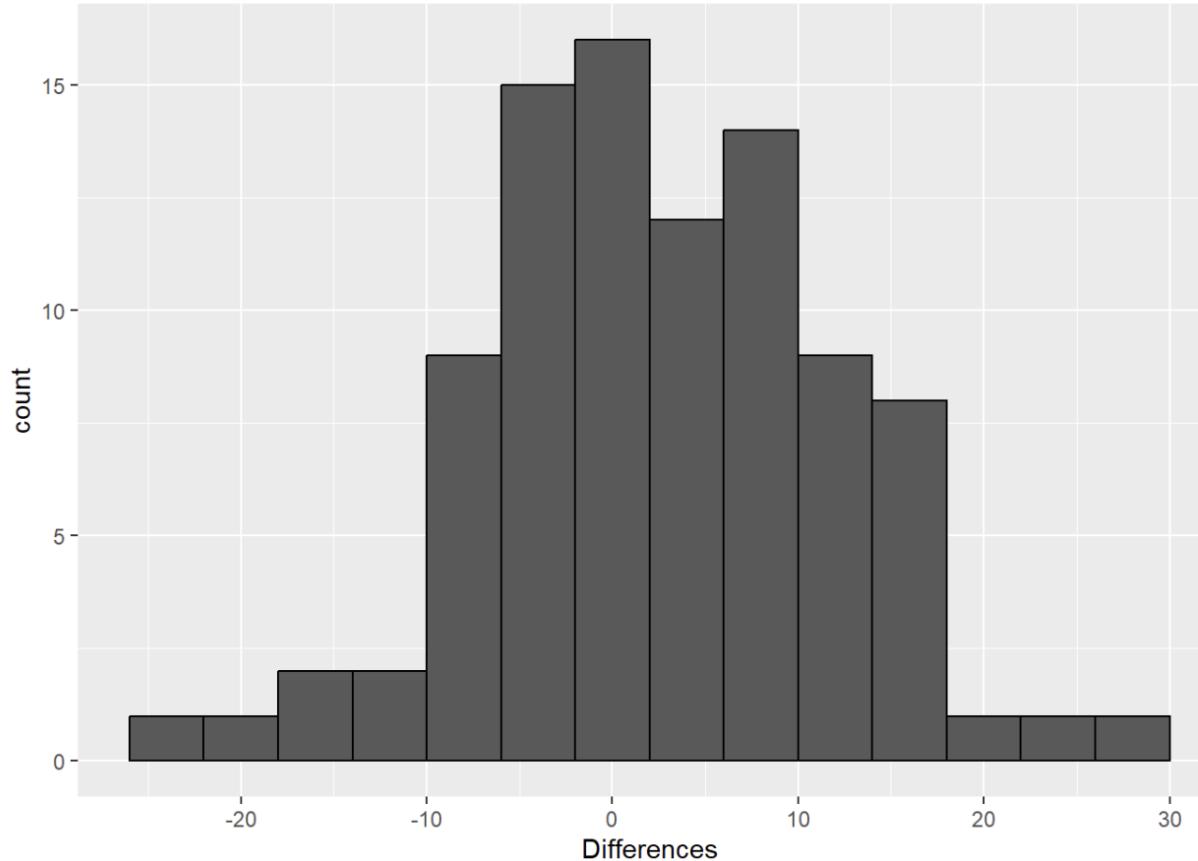
- ID: Student ID
- Reading: Result at reading
- Writing: Result at writing

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples $t$ -test – Example: Literacy

```
mydata$Difference <- mydata$Reading - mydata$Writing

library(ggplot2)
ggplot(mydata, aes(x = Difference)) +
  geom_histogram(binwidth = 4, color = "black") +
  xlab("Differences")
```



# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples *t*-test – Example: Literacy

```
shapiro.test(mydata$Difference)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: mydata$Difference  
## W = 0.99245, p-value = 0.8842
```

```
library(psych)
```

```
describe(mydata[ , -1])
```

```
##          vars n  mean   sd median trimmed   mad    min    max range skew kurtosis    se  
## Reading     1 92 80.99 7.94  81.34   81.20 7.67  55.11 100.00 44.89 -0.43     0.90 0.83  
## Writing     2 92 78.47 5.63  78.75   78.43 6.31  64.06  92.30 28.24 -0.03     -0.48 0.59  
## Difference  3 92  2.51 9.37   2.05    2.50 9.10 -25.51  28.33 53.84 -0.03     0.26 0.98
```

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples *t*-test – Example: Literacy

```
t.test(mydata$Reading, mydata$Writing,  
       paired = TRUE,  
       alternative = "two.sided")
```

```
##  
## Paired t-test  
##  
## data: mydata$Reading and mydata$Writing  
## t = 2.5711, df = 91, p-value = 0.01176  
## alternative hypothesis: true mean difference is not equal to 0  
## 95 percent confidence interval:  
##  0.5714427 4.4539920  
## sample estimates:  
## mean difference  
##          2.512717
```

```
library(effectsize)
```

```
cohens_d(mydata$Difference)
```

```
## Cohen's d |      95% CI  
## -----  
## 0.27     | [0.06, 0.48]
```

```
interpret_coherens_d(0.27, rules = "sawilowsky2009")
```

```
## [1] "small"  
## (Rules: sawilowsky2009)
```

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Paired samples $t$ -test – Example: Literacy

Conclusions:

Based on sample data we found the difference between average points received at reading ( $\bar{y} = 80.99$ ) and average points received at writing ( $\bar{y} = 78.47$ ). The difference is statistically significant at  $p = 0.012$  (effect size is small,  $d = 0.27$ ).

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 2: Grades

We have received data from 8 elementary school students about the result of the test in mathematics and English they wrote last week. Can we claim that there is a difference in the knowledge of elementary school students in the two subjects mentioned?

ID	Math		English		Sign	Difference	Abs(Difference)	Rank		
1	Excl	90	Good	70	Plus	20	20	5	Positive signs	2
2	Good	72	Very good	83	Minus	-11	11	4	Negative signs	5
3	Very good	85	Excellent	95	Minus	-10	10	3		
4	Fair	60	Very good	82	Minus	-22	22	6		
5	Very good	70	Good	68	Plus	2	2	1,5	Sum of positive signs	8
6	Very good	74	Very good	72	/	2	2	1,5	Sum of negative signs	28
7	Fair	62	Excelent	95	Minus	-33	33	8		
8	Fair	57	Very good	88	Minus	-31	31	7		

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 2: Grades

```
mydata <- read.table("./Grades.csv", header=TRUE, sep=";", dec=",")  
print(mydata)
```

```
##   lD      Math Math_points  English English_points  
## 1  1      Excl        90     Good        70  
## 2  2      Good        72  VeryGood        83  
## 3  3 VeryGood        85      Excl        95  
## 4  4      Fair        60  VeryGood        82  
## 5  5 VeryGood        70     Good        68  
## 6  6 VeryGood        74  VeryGood        72  
## 7  7      Fair        62      Excl        95  
## 8  8      Fair        57  VeryGood        88
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 2: Grades

```
library(dplyr)
```

```
mydata <- mydata %>%
  mutate(Math_code = dplyr::recode(mydata$Math,
    "Excl" = 5, "VeryGood" = 4, "Good" = 3, "Fair" = 2))
```

```
mydata <- mydata %>%
  mutate(English_code = dplyr::recode(mydata$English,
    "Excl" = 5, "VeryGood" = 4, "Good" = 3, "Fair" = 2))
```

```
print(mydata)
```

```
##   lD      Math Math_points  English English_points Math_code English_code
## 1  1      Excl        90     Good         70          5            3
## 2  2      Good        72  VeryGood        83          3            4
## 3  3  VeryGood        85      Excl        95          4            5
## 4  4      Fair        60  VeryGood        82          2            4
## 5  5  VeryGood        70     Good        68          4            3
## 6  6  VeryGood        74  VeryGood        72          4            4
## 7  7      Fair        62      Excl        95          2            5
## 8  8      Fair        57  VeryGood        88          2            4
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 2: Grades

### Sign Test

```
mydata$Difference <- mydata$Math_code - mydata$English_code

table(mydata$Difference)
```

```
##
## -3 -2 -1  0  1  2
##  1  2  2  1  1  1
```

```
binom.test(x = 5,
            n = 7,
            p = 0.5,
            alternative = "two.sided")
```

```
##
## Exact binomial test
##
## data: 5 and 7
## number of successes = 5, number of trials = 7, p-value = 0.4531
## alternative hypothesis: true probability of success is not equal to 0.5
## 95 percent confidence interval:
## 0.2904209 0.9633074
## sample estimates:
## probability of success
## 0.7142857
```

$$\binom{n}{k} p^k (1-p)^{n-k}$$

$$\binom{7}{5} 0,5^5 (1-0,5)^{7-5} = 21 \cdot \frac{1}{128} = 0,1641$$

$$\binom{7}{6} 0,5^6 (1-0,5)^{7-6} = 7 \cdot \frac{1}{128} = 0,0547$$

$$\binom{7}{7} 0,5^7 (1-0,5)^{7-7} = 1 \cdot \frac{1}{128} = 0,0078$$

$$\binom{7}{2} 0,5^3 (1-0,5)^{7-2} = 21 \cdot \frac{1}{128} = 0,1641$$

$$\binom{7}{1} 0,5^2 (1-0,5)^{7-1} = 7 \cdot \frac{1}{128} = 0,0547$$

$$\binom{7}{0} 0,5^1 (1-0,5)^{7-0} = 1 \cdot \frac{1}{128} = 0,0078$$

$$\sum = 0,4531$$

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 2: Grades

### Wilcoxon Signed Rank Test

```
wilcox.test(mydata$Math_points, mydata$English_points,  
           paired = TRUE,  
           correct = FALSE,  
           exact = FALSE,  
           alternative = "two.sided")
```

```
##  
## Wilcoxon signed rank test  
##  
## data: mydata$Math_points and mydata$English_points  
## V = 8, p-value = 0.1609  
## alternative hypothesis: true location shift is not equal to 0
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 3: Party

10 participants attend the party two weekends in a row. On the first weekend they consume alcohol and cannabis, on the second weekend they consume only alcohol. The day after the party, each of them rates how severe the headache is on a scale of 1 to 7 (higher score - higher intensity). Check if there are differences in the intensity of the headache after the first and second weekend.

```
mydata <- read.table("./Party.csv", header=TRUE, sep=";", dec=",")  
print(mydata)
```

```
##      ID Cannabis Alcohol  
## 1     1        3       3  
## 2     2        7       3  
## 3     3        3       4  
## 4     4        4       3  
## 5     5        5       3  
## 6     6        4       2  
## 7     7        5       3  
## 8     8        3       2  
## 9     9        2       1  
## 10   10       6       2
```

Description:

- ID: Person ID
- Cannabis: The headaches (1 - weak to 7 - strong)
- Alcohol: The headaches (1 - weak to 7 - strong)

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

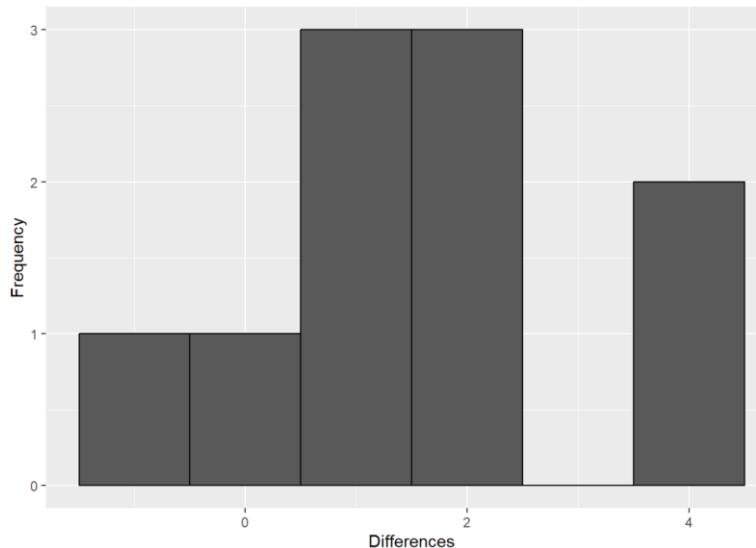
## Example 3: Party

```
mydata$Difference <- mydata$Cannabis - mydata$Alcohol

describe(mydata[, -1])
```

```
##          vars   n  mean    sd median trimmed   mad min  max range skew kurtosis    se
## Cannabis     1 10  4.2 1.55     4.0    4.12 1.48     2    7      5  0.35    -1.22  0.49
## Alcohol      2 10  2.6 0.84     3.0    2.62 0.74     1    4      3 -0.28    -0.84  0.27
## Difference   3 10  1.6 1.58     1.5    1.62 0.74    -1    4      5  0.14    -1.08  0.50
```

```
library(ggplot2)
ggplot(mydata, aes(x=Difference)) +
  geom_histogram(binwidth = 1, color = "black") +
  xlab("Differences") +
  ylab("Frequency")
```



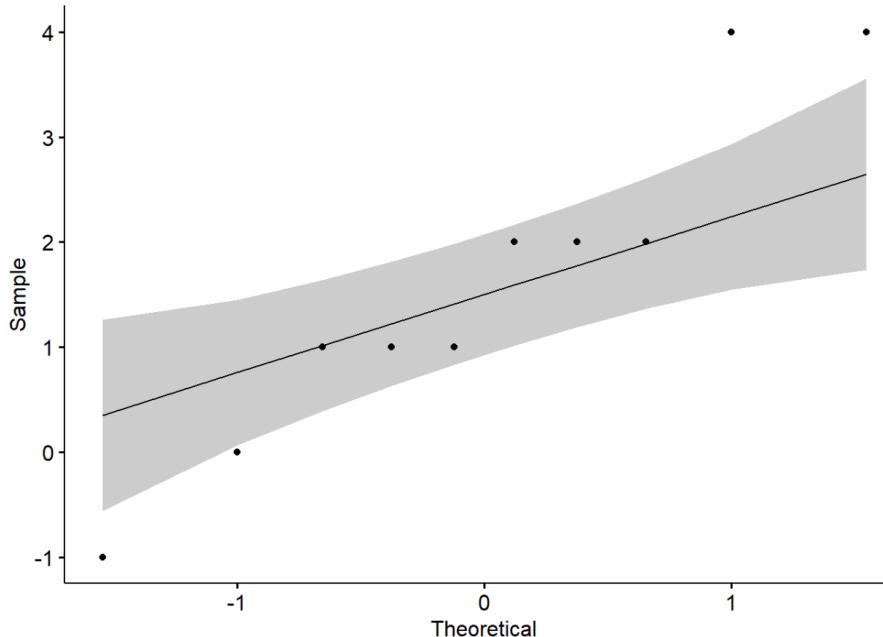
# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 3: Party

```
shapiro.test(mydata$Difference)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: mydata$Difference  
## W = 0.92897, p-value = 0.4378
```

```
library(ggpubr)  
ggqqplot(mydata$Difference)
```



# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 3: Party

```
wilcox.test(mydata$Cannabis, mydata$Alcohol,
            paired = TRUE,
            correct = FALSE,
            exact = FALSE,
            alternative = "two.sided")
```

```
##  
## Wilcoxon signed rank test  
##  
## data: mydata$Cannabis and mydata$Alcohol  
## V = 42.5, p-value = 0.01634  
## alternative hypothesis: true location shift is not equal to 0
```

```
library(effectsize)
```

```
effectsize(wilcox.test(mydata$Cannabis, mydata$Alcohol,
                      paired = TRUE,
                      correct = FALSE,
                      exact = FALSE,
                      alternative = "two.sided"))
```

```
## r (rank biserial) |      95% CI
## -----
## 0.89 | [0.62, 0.97]
```

Funder & Ozer (2019) ("funder2019"; default)

- **$r < 0.05$**  - Tiny
- **$0.05 \leq r < 0.1$**  - Very small
- **$0.1 \leq r < 0.2$**  - Small
- **$0.2 \leq r < 0.3$**  - Medium
- **$0.3 \leq r < 0.4$**  - Large
- **$r \geq 0.4$**  - Very large

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 3: Party

Conclusions:

Using the sample data, we find that the intensity of the headache with cannabis and alcohol differs from the intensity of the headache with alcohol alone ( $p = 0.017$ ). More severe headaches are expected after the consumption of alcohol and cannabis together (effect size  $r = 0.89$ ).

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Independent samples $t$ -test – Example 4: Math

The math professor was interested in whether there is a difference in exam scores between students who exercise at least three times a week and those who do not. Over the past few years, she has collected data on first exam scores, with students also indicating whether they are actively exercising.

```
mydata <- read.table("./Math.csv", header=TRUE, sep=";", dec=",")  
head(mydata)
```

```
##   ID Active  Math  
## 1  1      1 93.78  
## 2  2      0 86.66  
## 3  3      1 85.69  
## 4  4      0 85.57  
## 5  5      0 85.32  
## 6  6      0 85.19
```

Description:

- ID: ID of a student
- Active: Is student regularly exercising?
- Math: Result of math exam

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Independent samples $t$ -test – Example 4: Math

```
mydata$ActiveF <- factor(mydata$Active,  
                          levels = c(0, 1),  
                          labels = c("No", "Yes"))
```

```
library(psych)
```

```
describeBy(mydata$Math, mydata$ActiveF)
```

```
##  
## Descriptive statistics by group  
## group: No  
##   vars   n  mean   sd median trimmed   mad   min   max range skew kurtosis    se  
## X1     1 242 65.58 8.58   65.78   65.72 8.04 40.38 86.66 46.28 -0.14      0.06 0.55  
## -----  
## group: Yes  
##   vars   n  mean   sd median trimmed   mad   min   max range skew kurtosis    se  
## X1     1 180 65.32 8.12   65.12   65.37 6.99 35.32 93.78 58.46 -0.06      1.07 0.61
```

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

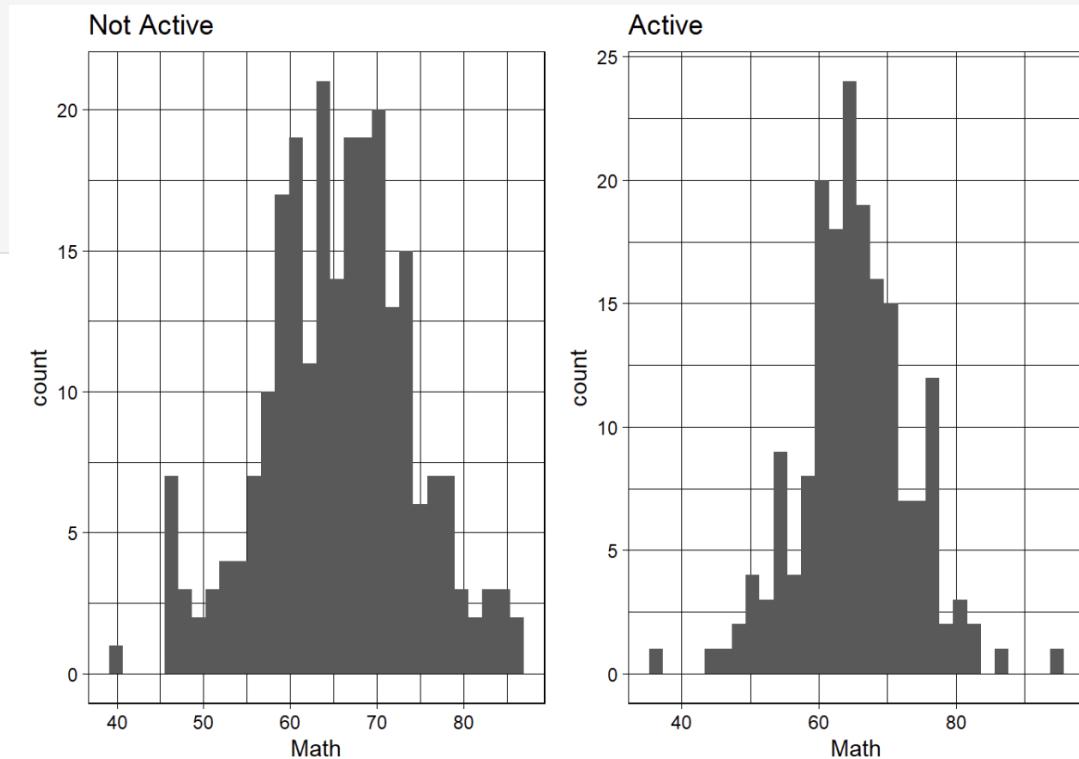
## Independent samples $t$ -test – Example 4: Math

```
library(ggplot2)

Active_no <- ggplot(mydata[mydata$ActiveF == "No", ], aes(x = Math)) +
  theme_linedraw() +
  geom_histogram() +
  ggtitle("Not Active")

Active_yes <- ggplot(mydata[mydata$ActiveF == "Yes", ], aes(x = Math)) +
  theme_linedraw() +
  geom_histogram() +
  ggtitle("Active")

library(ggpubr)
ggarrange(Active_no, Active_yes,
          ncol = 2, nrow = 1)
```



# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Independent samples $t$ -test – Example 4: Math

```
t.test(mydata$Math ~ mydata$ActiveF,  
       paired = FALSE,  
       var.equal = FALSE,  
       alternative = "two.sided")
```

```
##  
## Welch Two Sample t-test  
##  
## data: mydata$Math by mydata$ActiveF  
## t = 0.31548, df = 396.45, p-value = 0.7526  
## alternative hypothesis: true difference in means between group No and group Yes is not equal to 0  
## 95 percent confidence interval:  
## -1.351309 1.867895  
## sample estimates:  
## mean in group No mean in group Yes  
## 65.57818 65.31989
```

# HYPOTHESIS ABOUT THE DIFFERENCE BETWEEN TWO ARITHMETIC MEANS

## Independent samples $t$ -test – Example 4: Math

```
library(effectsize)
```

```
effectsize::cohens_d(mydata$Math ~ mydata$ActiveF,  
                      pooled_sd = FALSE)
```

```
## Cohen's d |      95% CI  
## -----  
## 0.03     | [-0.16, 0.22]  
##  
## - Estimated using un-pooled SD.
```

```
interpret_cohens_d(0.03, rules = "sawilowsky2009")
```

```
## [1] "tiny"  
## (Rules: sawilowsky2009)
```

Conclusions:

Using the sample data, the professor was unable to demonstrate that the average score in mathematics differed between athletically active and non-athletically active students (the effect size is tiny,  $d = 0.03$ ).

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

You wonder if there are differences between men and women in the amount of time they spend at the shopping mall. You measured the time a person spends in a single visit to a large shopping mall in Vienna for 17 men and 20 women. What is your conclusion?

```
mydata <- read.table("./Shopping.csv", header=TRUE, sep=";", dec=",")  
head(mydata)
```

```
##   ID Gender Time  
## 1  1      1   22  
## 2  2      1  167  
## 3  3      1  160  
## 4  4      1  205  
## 5  5      1  279  
## 6  6      1   15
```

Description:

- ID: Identifactor.
- Gender: 0 = Male, 1 = Female,
- Time: Time shopping in minutes

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

```
mydata$GenderF <- factor(mydata$Gender,  
                          levels = c(0, 1),  
                          labels = c("Male", "Female"))
```

```
library(psych)  
describeBy(mydata$Time, g = mydata$GenderF)
```

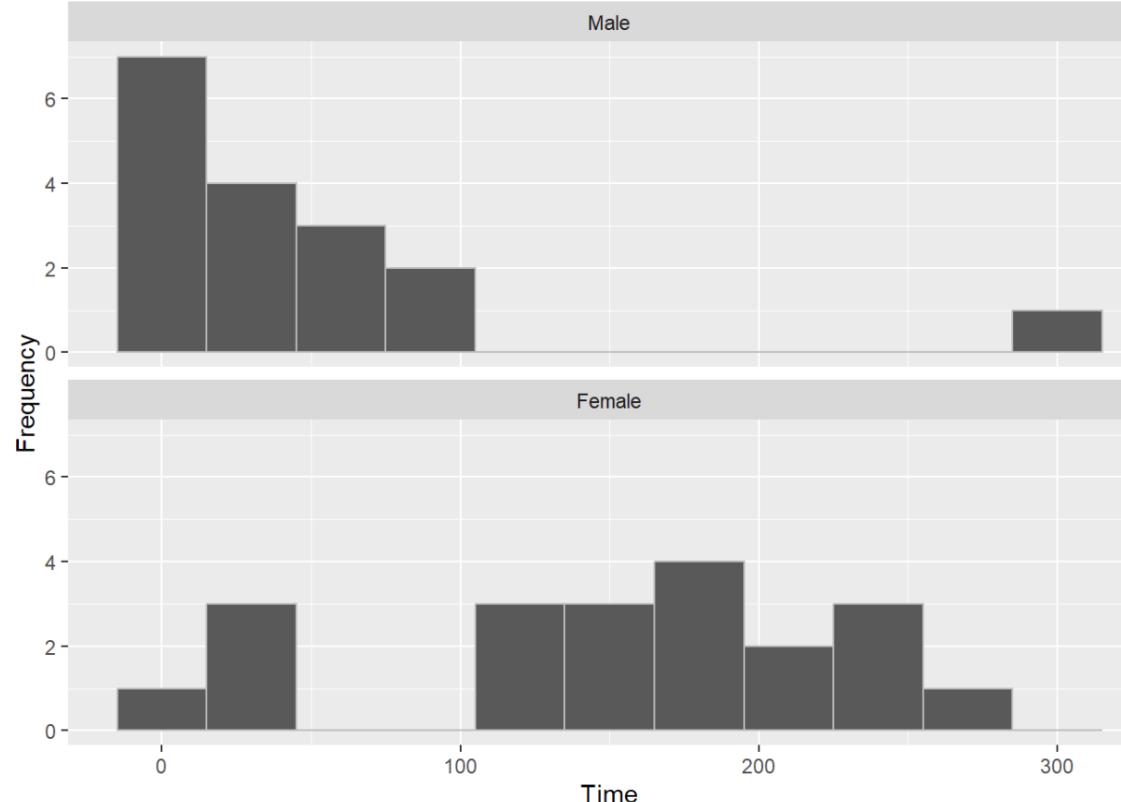
```
##  
## Descriptive statistics by group  
## group: Male  
##   vars n  mean    sd median trimmed   mad min max range skew kurtosis    se  
## X1    1 17 52.82 70.85     22      39 20.76    8 305   297 2.59      6.5 17.18  
## -----  
## group: Female  
##   vars n  mean    sd median trimmed   mad min max range skew kurtosis    se  
## X1    1 20 151.35 80.67   163.5  153.81 66.72   15 279   264 -0.39      -1 18.04
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

```
library(ggplot2)
```

```
ggplot(mydata, aes(x = Time)) +  
  geom_histogram(binwidth = 30, colour="gray") +  
  facet_wrap(~GenderF, ncol = 1) +  
  ylab("Frequency")
```



# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

```
library(rstatix)
```

```
mydata %>%
  group_by(GenderF) %>%
  shapiro_test(Time)
```

```
## # A tibble: 2 × 4
##   GenderF variable statistic      p
##   <fct>    <chr>     <dbl>    <dbl>
## 1 Male     Time      0.602 0.0000109
## 2 Female   Time      0.927 0.138
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

```
wilcox.test(mydata$Time ~ mydata$GenderF,  
            paired = FALSE,  
            correct = FALSE,  
            exact = FALSE,  
            alternative = "two.sided")
```

```
##  
## Wilcoxon rank sum test  
##  
## data: mydata$Time by mydata$GenderF  
## W = 52.5, p-value = 0.0003355  
## alternative hypothesis: true location shift is not equal to 0
```

# HYPOTHESIS ABOUT THE DIFFERENCE IN TWO DISTRIBUTION LOCATIONS

## Example 5: Shopping

```
library(effectsize)
```

```
effectsize(wilcox.test(mydata$Time ~ mydata$GenderF,
                      paired = FALSE,
                      correct = FALSE,
                      exact = FALSE,
                      alternative = "two.sided"))
```

```
## r (rank biserial) |      95% CI
## -----
## -0.69 | [-0.84, -0.44]
```

```
interpret_rank_biserial(0.69)
```

```
## [1] "very large"
## (Rules: funder2019)
```

Funder & Ozer (2019) ("funder2019"; default)

- **r < 0.05** - Tiny
- **0.05 <= r < 0.1** - Very small
- **0.1 <= r < 0.2** - Small
- **0.2 <= r < 0.3** - Medium
- **0.3 <= r < 0.4** - Large
- **r >= 0.4** - Very large

Conclusions:

Based on the sample data, we find that men and women differ in the amount of time spent at the mall ( $p < 0.001$ ) - women spend more time, the difference in distribution is very large ( $r = 0.69$ )

# REPEATED MEASURES ANALYSIS OF VARIANCE, rANOVA

## FRIEDMAN ANALYSIS OF VARIANCE

# REPEATED MEASURES ANOVA - rANOVA

Comparison of three or more arithmetic means for dependent samples (extension of paired samples  $t$ -test), rANOVA.

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

$H_1$ : At least one  $\mu_j$  is different

- The same unit (e.g., person) is measured over time (at least 3 times) and the differences in the averages are compared.
- Using three or more comparable variables (e.g., agreement measured on a Likert scale of 1-5), belonging to the same unit (e.g., person), and checking to see if there are differences in averages.

# REPEATED MEASURES ANOVA - rANOVA

Conditions and assumptions:

- Analyzed variables are numeric.
- Normality: The variables in the population are normally distributed (this can be checked with the Shapiro-Wilk test).
- Sphericity: the variances of the differences between the examined variables are equal (tested with the Mauchly test for sphericity).
- No significant outliers.

# FRIEDMAN ANOVA

Comparison of three or more locations for dependent samples (extension of Wilcoxon Signed Rank Test).

$H_0$ : All distribution locations of variables are the same

$H_1$ : At least one distribution location of variable is different

Comparison of three or more distribution locations of variables belonging to the same units:

- The same unit is measured over time (at least 3 times) and differences in the location of variable distributions are compared using rank analysis.
- Using three or more comparable variables (e.g., agreement measured on a Likert scale of 1-5) belonging to the same unit and checking by rank analysis whether there are differences in the location of the variable distributions.

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

In his master's thesis, the student investigated the level of trust in different social systems in Slovenia. He conducted an online survey using the 1KA program and received 1,318 responses. He identified the following as possible social systems, on which respondents indicated how much they trust them on a scale between 0 and 10 (the higher the value, the greater the trust):

- Trust in parliament (trstprl)
- Trust in the rule of law (trstlgl)
- Trust in the police (trstpplc)
- Trust in politics (trstplt)

Are there differences in trust in these social systems?

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
mydata <- read.table("./Trust.csv", header=TRUE, sep=";", dec=",")  
head(mydata)
```

```
##   ID trstprl trstlg1 trstplc trstplt  
## 1  1      3      0      7      0  
## 2  2      2      3      3      2  
## 3  3      0      4     10      0  
## 4  4      0      0      0      0  
## 5  5      1      6      7      3  
## 6  6      4     88      7      4
```

Description:

- ID: Person ID
- trstprl: Trust in the parliament (0-10)
- trstlg1: Trust in the legal system (0-10)
- trstplc: Trust in the police (0-10)
- trstplt: Trust in the politics (0-10)

77, 88, 99 missing values

```
set.seed(1) #Setting initial point of sampling.  
mydata <- mydata[sample(nrow(mydata), 50), ] #Random sample of 50 units.
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(dplyr)
library(nanrar)

#Missing values are changed to NA.
mydata <- mydata %>%
  replace_with_na(replace = list(trstprl = c(77, 88, 99),
                                 trstlgl = c(77, 88, 99),
                                 trstplc = c(77, 88, 99),
                                 trstplt = c(77, 88, 99)))

tail(mydata, 10) #Showing last 10 units.
```

```
##      ID trstprl trstlgl trstplc trstplt
## 810   810      4      0      4      0
## 526   526     NA      0      9      0
## 642   642      2      4      6      3
## 1069  1069     3      2      7      6
## 22    22       4      3      7      3
## 193   193      6      8      9      3
## 499   499      5      4      7      5
## 1128  1128     5      3      7      5
## 983   983      3      2      4      1
## 843   843      5      4      6      5
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(tidyverse)
```

```
mydata <- drop_na(mydata) #Removing units with missing values.
```

```
tail(mydata, 10)
```

```
##      ID trstprl trstlgl trstplc trstplt
## 40  435      0      0      0      0
## 41  810      4      0      4      0
## 42  642      2      4      6      3
## 43 1069      3      2      7      6
## 44   22      4      3      7      3
## 45  193      6      8      9      3
## 46  499      5      4      7      5
## 47 1128      5      3      7      5
## 48  983      3      2      4      1
## 49  843      5      4      6      5
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(rstatix)
```

```
library(tidyverse)
```

```
#Wide format of data is changed to Long format.  
mydata_long <- mydata %>%  
  pivot_longer(  
    cols = c("trstprl", "trstlgl", "trstplc", "trstplt"),  
    names_to = "System",  
    values_to = "Trust") %>%  
  convert_as_factor(System)  
  
mydata_long <- as.data.frame(mydata_long)  
  
tail(mydata_long, 10)
```

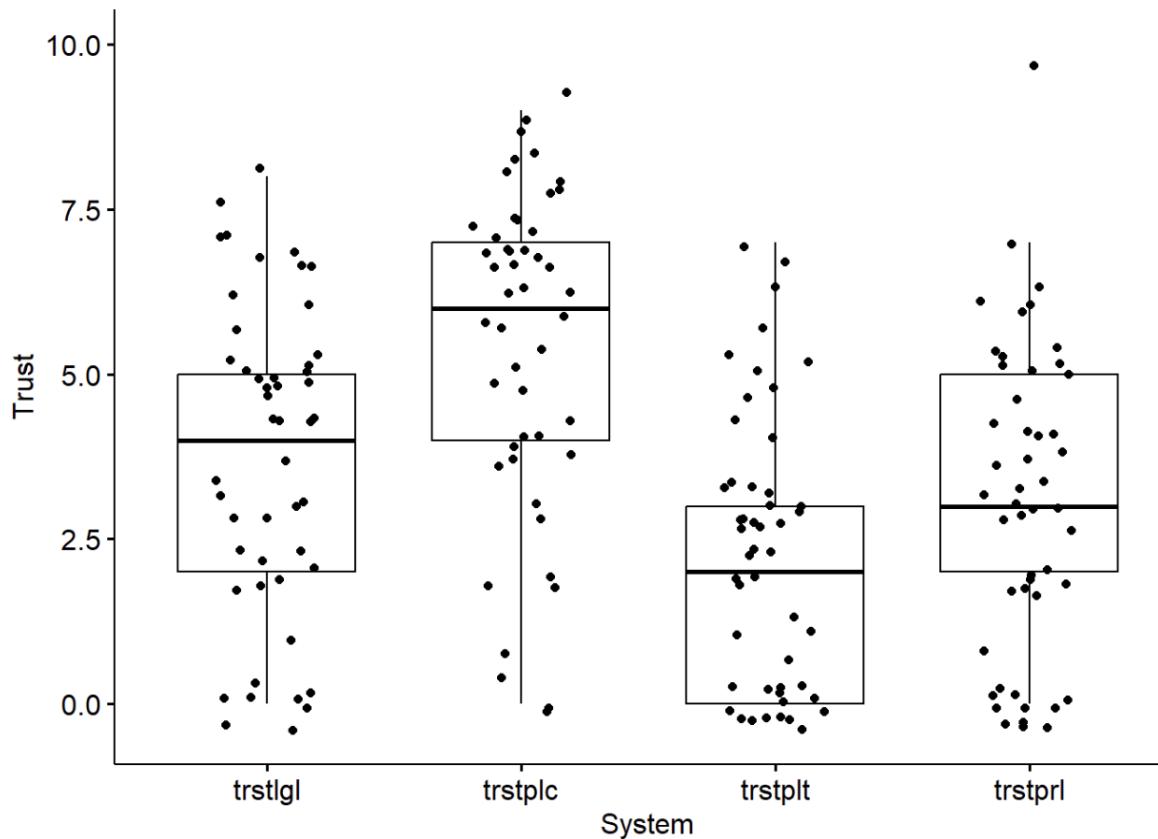
```
##      ID System Trust  
## 187 1128 trstplc    7  
## 188 1128 trstplt    5  
## 189  983 trstprl    3  
## 190  983 trstlgl    2  
## 191  983 trstplc    4  
## 192  983 trstplt    1  
## 193  843 trstprl    5  
## 194  843 trstlgl    4  
## 195  843 trstplc    6  
## 196  843 trstplt    5
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(ggpubr)

#Boxplot for each variable.
ggboxplot(mydata_long,
  x = "System",
  y = "Trust",
  add = "jitter")
```



# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(tidyverse)
library(ggpubr)
library(rstatix)
```

#*Finding outliers.*

```
mydata_long %>%
  group_by(System) %>%
  identify_outliers(Trust)
```

```
## # A tibble: 1 × 5
##   System     ID Trust is.outlier is.extreme
##   <fct>    <int> <int>    <lgl>      <lgl>
## 1 trstprl    248     10  TRUE      FALSE
```

#*Removing outliers.*

```
mydata_long <- mydata_long %>%
  filter(!ID == 248)
```

```
mydata_long %>%
  group_by(System) %>%
  identify_outliers(Trust)
```

```
## [1] System     ID         Trust      is.outlier is.extreme
## <0 rows> (or 0-length row.names)
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(rstatix)

#Checking normality with Shapiro-Wilk test
mydata_long %>%
  group_by(System) %>%
  shapiro_test(Trust)
```

```
## # A tibble: 4 × 4
##   System  variable statistic      p
##   <fct>   <chr>     <dbl>    <dbl>
## 1 trstlgl Trust     0.935 0.0104
## 2 trstplc Trust     0.914 0.00187
## 3 trstplt Trust     0.891 0.000317
## 4 trstprl Trust     0.919 0.00269
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(rstatix)

#Descriptive statistics for each variable.
mydata_long %>%
  group_by(System) %>%
  get_summary_stats(Trust, type = "common")
```

```
## # A tibble: 4 × 11
##   System variable     n   min   max median   iqr   mean    sd    se    ci
##   <fct>   <fct> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1 trstlgl Trust     48     0     8     4     3     3.71  2.40  0.347  0.698
## 2 trstp1c Trust     48     0     9     6     3     5.46  2.43  0.351  0.706
## 3 trstp1t Trust     48     0     7     2     3     2.23  1.98  0.286  0.575
## 4 trstp1r Trust     48     0     7     3     3.25  2.94  2.07  0.298  0.6
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(rstatix)

#rANOVA
ANOVA_results <- anova_test(dv = Trust, #Dependent variable
                             wid = ID, #Subject identifier
                             within = System, #Within-subject factor variable
                             data = mydata_long)

ANOVA_results #Summary of results.
```

```
## ANOVA Table (type III tests)
##
## $ANOVA
##   Effect DFn DFD      F      p p<.05    ges
## 1 System    3 141 39.534 1.33e-18     * 0.229
##
## $`Mauchly's Test for Sphericity`
##   Effect      W      p p<.05
## 1 System 0.747 0.021     *
##
## $`Sphericity Corrections`
##   Effect    GGe      DF[GG]  p[GG] p[GG]<.05    HFe      DF[HF]  p[HF] p[HF]<.05
## 1 System 0.874 2.62, 123.3 1.45e-16     * 0.931 2.79, 131.25 1.76e-17     *
```

```
get_anova_table(ANOVA_results, correction = "auto")
```

```
## ANOVA Table (type III tests)
##
##   Effect DFn DFD      F      p p<.05    ges
## 1 System 2.62 123.3 39.534 1.45e-16     * 0.229
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(effectsize)
```

```
interpret_eta_squared(0.229, rules = "cohen1992")
```

```
## [1] "medium"  
## (Rules: cohen1992)
```

Cohen (1992) ("cohen1992") applicable to one-way anova, or to *partial* eta / omega / epsilon squared in multi-way anova.

- **ES < 0.02** - Very small
- **0.02 <= ES < 0.13** - Small
- **0.13 <= ES < 0.26** - Medium
- **ES >= 0.26** - Large

# REPEATED MEASURES ANOVA – rANOVA

## Post-Hoc tests

Post-Hoc tests compare all combinations of two variables.

Because the probability of a Type I error increases when the statistical tests are repeated on the same sample of data, the criterion for rejecting null hypothesis is more stringent. The most commonly used method is the Bonferroni correction, which increases the reported  $p$ -value according to the number of tests performed.

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

```
library(rstatix)

#Comparing all pairs of variables
pwc <- mydata_long %>%
  pairwise_t_test(Trust ~ System,
                  paired = TRUE,
                  p.adjust.method = "bonferroni")

pwc

## # A tibble: 6 × 10
##   .y.   group1  group2     n1     n2 statistic    df      p    p.adj p.adj.signif
## * <chr> <chr>   <chr> <int> <int>     <dbl> <dbl>    <dbl> <dbl> <chr>
## 1 Trust trstlgl trstplc    48     48     -5.09    47 6.27e- 6 3.76e- 5 ****
## 2 Trust trstlgl trstplt    48     48      4.44    47 5.36e- 5 3.22e- 4 ***
## 3 Trust trstlgl trstprl    48     48      2.50    47 1.6 e- 2 9.6 e- 2 ns
## 4 Trust trstplc trstplt    48     48     10.1     47 2   e-13 1.2 e-12 ****
## 5 Trust trstplc trstprl    48     48      7.54    47 1.24e- 9 7.44e- 9 ****
## 6 Trust trstplt trstprl    48     48     -3.24    47 2   e- 3 1.3 e- 2 *
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

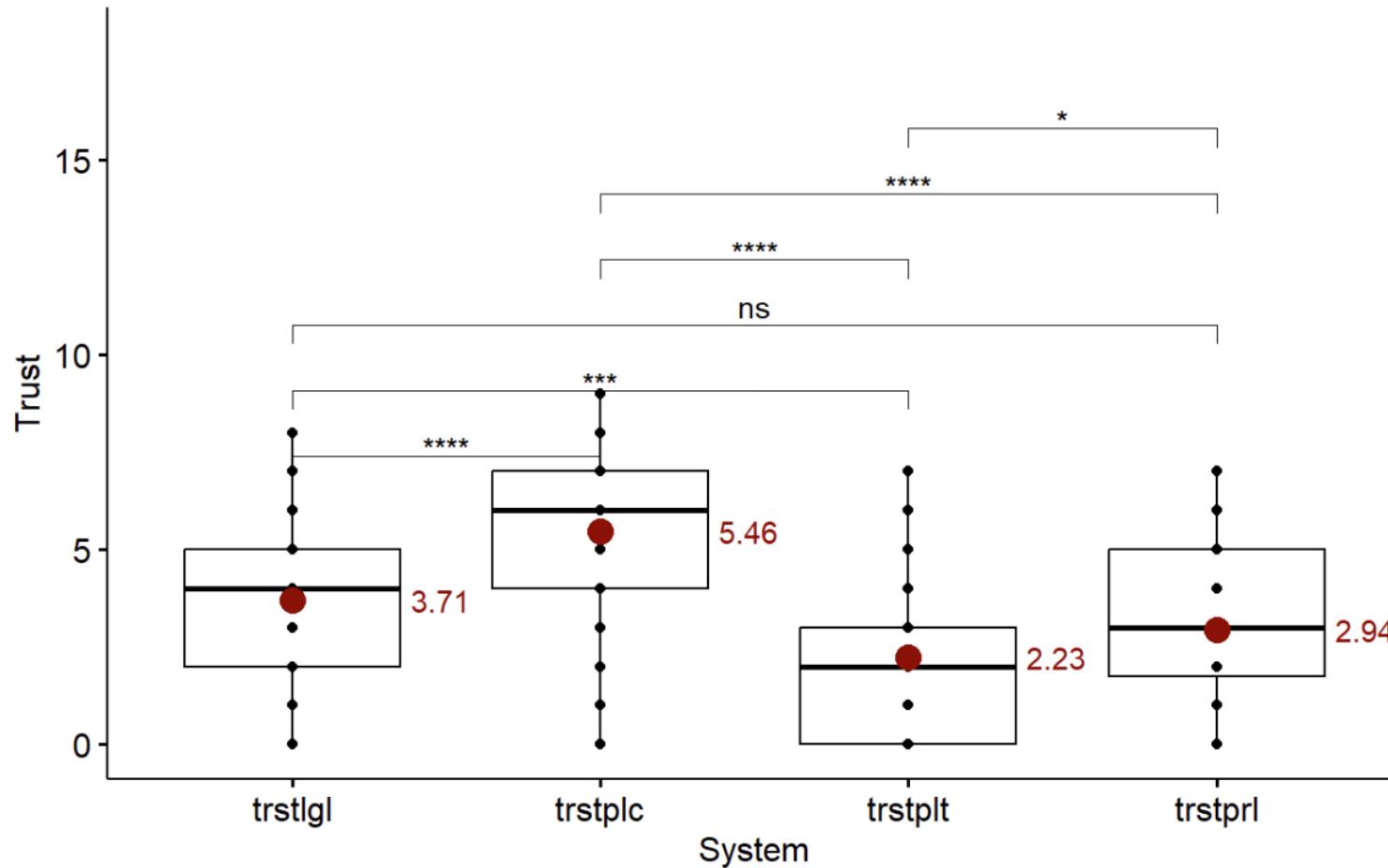
```
library(rstatix)
pwc <- pwc %>%
  add_y_position(fun = "median", step.increase = 0.35)

library(ggpubr)
ggboxplot(mydata_long, x = "System", y = "Trust", add = "point", ylim=c(0, 18)) +
  stat_pvalue_manual(pwc, hide.ns = FALSE) +
  stat_summary(fun = mean, geom = "point", shape = 16, size = 4,
               aes(group = System), color = "darkred",
               position = position_dodge(width = 0.8)) +
  stat_summary(fun = mean, colour = "darkred",
               position = position_dodge(width = 0.8),
               geom = "text", vjust = 0.5, hjust = -2,
               aes(label = round(after_stat(y), digits = 2), group = System)) +
  labs(subtitle = get_test_label(ANOVA_results, detailed = TRUE),
       caption = get_pwc_label(pwc))
```

# REPEATED MEASURES ANOVA – rANOVA

## Example: Trust

Anova,  $F(2.62, 123.3) = 39.53, p = <0.0001, \eta_g^2 = 0.23$



pwc: T test; p.adjust: Bonferroni

# FRIEDMAN ANOVA

## Example: Trust

```
library(rstatix)

#Friedman ANOVA.
FriedmanANOVA <- friedman_test(Trust ~ System | ID,
                                data = mydata_long)

FriedmanANOVA #Summary of results.
```

```
## # A tibble: 1 × 6
##   .y.      n statistic    df      p method
## * <chr> <int>     <dbl> <dbl> <dbl> <chr>
## 1 Trust     48      68.2     3 1.02e-14 Friedman test
```

```
library(effectsize)
effectsize::kendalls_w(Trust ~ System | ID,
                        data = mydata_long)
```

```
## Warning: 36 block(s) contain ties, some containing only 1 unique ranking.
```

```
## Kendall's W |      95% CI
## -----
## 0.47      | [0.37, 1.00]
##
## - One-sided CIs: upper bound fixed at [1.00].
```

```
interpret_kendalls_w(0.47)
```

```
## [1] "moderate agreement"
## (Rules: landis1977)
```

Landis & Koch (1977) ("landis1977"; default)

- **0.00 <= w < 0.20** - Slight agreement
- **0.20 <= w < 0.40** - Fair agreement
- **0.40 <= w < 0.60** - Moderate agreement
- **0.60 <= w < 0.80** - Substantial agreement
- **w >= 0.80** - Almost perfect agreement

# FRIEDMAN ANOVA

## Example: Trust

```
library(rstatix)

#Wilcoxon signed rank tests - comparing all possible pairs.
paires_nonpar <- wilcox_test(Trust ~ System,
                             paired = TRUE,
                             p.adjust.method = "bonferroni",
                             data = mydata_long)

paires_nonpar
```

```
## # A tibble: 6 × 9
##   .y.   group1  group2     n1     n2 statistic          p      p.adj p.adj.signif
## * <chr> <chr>   <chr> <int> <int>    <dbl>      <dbl>      <dbl> <chr>
## 1 Trust trstlgl trstplc    48     48     109  0.0000279  0.000167    ***
## 2 Trust trstlgl trstplt    48     48     548.  0.00013    0.00078    ***
## 3 Trust trstlgl trstprtl  48     48     464.  0.014     0.082     ns
## 4 Trust trstplc trstplt    48     48     942.  0.0000000133 0.0000000798 ****
## 5 Trust trstplc trstprtl  48     48     908.  0.000000138  0.000000828 ****
## 6 Trust trstplt trstprtl  48     48     56.5  0.004     0.023     *
```

# FRIEDMAN ANOVA

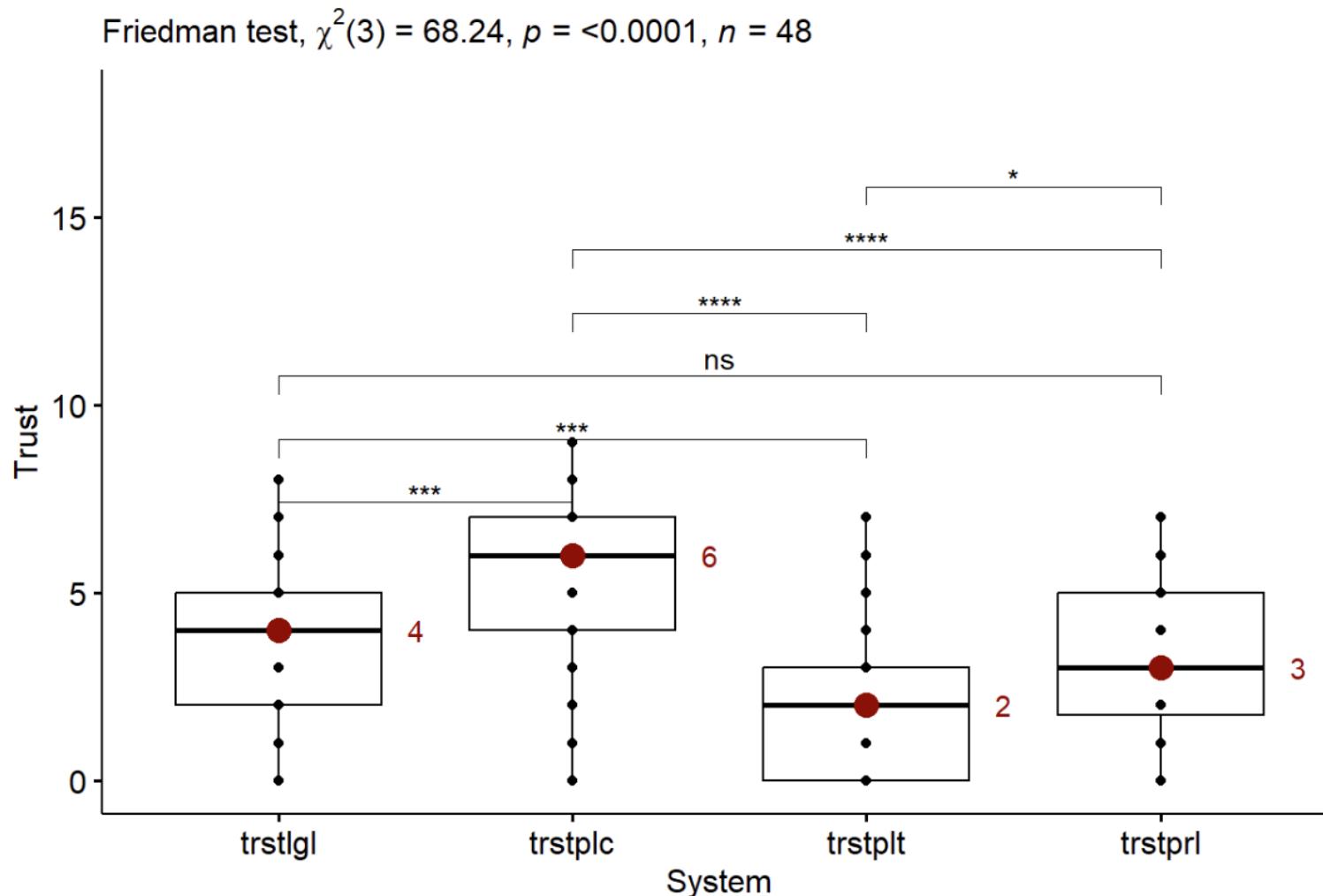
## Example: Trust

```
library(rstatix)
comparisons <- paires_nonpar %>%
    add_y_position(fun = "median", step.increase = 0.35)

library(ggpubr)
ggboxplot(mydata_long, x = "System", y = "Trust", add = "point", ylim=c(0, 18)) +
    stat_pvalue_manual(comparisons, hide.ns = FALSE) +
    stat_summary(fun = median, geom = "point", shape = 16, size = 4,
                 aes(group = System), color = "darkred",
                 position = position_dodge(width = 0.8)) +
    stat_summary(fun = median, colour = "darkred",
                 position = position_dodge(width = 0.8),
                 geom = "text", vjust = 0.5, hjust = -8,
                 aes(label = round(after_stat(y), digits = 2), group = System)) +
    labs(subtitle = get_test_label(FriedmanANOVA, detailed = TRUE),
         caption = get_pwc_label(comparisons))
```

# FRIEDMAN ANOVA

## Example: Trust



pwc: Wilcoxon test; p.adjust: Bonferroni

# ONE-WAY ANALYSIS OF VARIANCE (ANOVA)

## KRUSKAL-WALLIS RANK SUM TEST

# ONE-WAY ANALYSIS OF VARIANCE - ANOVA

Comparison of three or more arithmetic means for independent samples (extension of independent samples  $t$ -test).

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

$H_1$ : At least one  $\mu_j$  is different from others

Different groups of units are measured (at the same time) and the differences in the group averages are compared.

# ONE-WAY ANALYSIS OF VARIANCE - ANOVA

Conditions and assumptions:

- Analyzed variable is numeric.
- Variable in the population is normally distributed within each group.
- Homoscedasticity: the variance of analyzed variable is the same within all groups.
- No significant outliers.

# KRUSKAL-WALLIS RANK SUM TEST

Comparison of three or more distribution locations of variables for independent samples (extension of the Wilcoxon Rank Sum Test).

$H_0$ : All distribution locations of variables are the same

$H_1$ : At least one distribution location of variable is different

Comparison of three or more distribution locations of the same variable, belonging to different groups of units.

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

Puppy therapy.

90 people is randomly assigned to one of three groups:

- Control group – no puppy (28 people)
- 15 minutes of puppy therapy (32 people)
- 30 minutes of puppy therapy (30 people)

The analyzed variable is a measure of happiness between 1 (as unhappy as I can imagine) and 10 (as happy as I can imagine). Are puppy therapy and happiness related?

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
mydata <- read.table("./Puppies.csv", header=TRUE, sep=";", dec=".")  
head(mydata)
```

```
##   ID Dose Happiness  
## 1  1    1        3  
## 2  2    1        3  
## 3  3    1        4  
## 4  4    1        1  
## 5  5    1        3  
## 6  6    1        6
```

Description:

- Person: ID
- Dose: 1: Control group, 2: 15 minutes exposed to Puppy, 3: 30 minutes exposed to Puppy
- Happiness: How happy is a person between 1 and 10.

```
mydata$DoseFactor <- factor(mydata$Dose,  
                           levels = c(1, 2, 3),  
                           labels = c("Control", "15 minutes", "30 minutes"))
```

```
library(psych)
```

```
describe(mydata$Happiness)
```

```
##     vars  n mean   sd median trimmed  mad min max range skew kurtosis    se  
## X1     1 90 5.33 2.12      5  5.33 2.97    1 10      9 0.05    -0.87 0.22
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
describeBy(x = mydata$Happiness, group = mydata$DoseFactor)
```

```
##  
## Descriptive statistics by group  
## group: Control  
##   vars n mean sd median trimmed mad min max range skew kurtosis se  
## X1    1 28  3.5 1.32     3.5    3.5 0.74    1   6     5     0    -0.67  0.25  
## -----  
## group: 15 minutes  
##   vars n mean sd median trimmed mad min max range skew kurtosis se  
## X1    1 32  5.06 1.72      5      5 1.48    2   9     7     0.24    -0.54  0.3  
## -----  
## group: 30 minutes  
##   vars n mean sd median trimmed mad min max range skew kurtosis se  
## X1    1 30  7.33 1.24     7     7.33 1.48    5  10     5     0    -0.73  0.23
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
library(car)
leveneTest(mydata$Happiness, group = mydata$DoseFactor)
```

```
## Levene's Test for Homogeneity of Variance (center = median)
##          Df F value Pr(>F)
## group    2  1.0404 0.3577
##          87
```

```
library(dplyr)
```

```
library(rstatix)
```

```
mydata %>%
  group_by(DoseFactor) %>%
  shapiro_test(Happiness)
```

```
## # A tibble: 3 × 4
##   DoseFactor variable  statistic     p
##   <fct>      <chr>       <dbl> <dbl>
## 1 Control    Happiness  0.946  0.161
## 2 15 minutes Happiness  0.964  0.356
## 3 30 minutes Happiness  0.943  0.108
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
ANOVA_results <- aov(Happiness ~ DoseFactor,
                      data = mydata)
```

```
summary(ANOVA_results)
```

```
##           Df Sum Sq Mean Sq F value    Pr(>F)
## DoseFactor   2  216.5  108.23    51.3 1.92e-15 ***
## Residuals   87  183.5     2.11
## ---
## Signif. codes:  0 '****' 0.001 '***' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
library(effectsize)
```

```
eta_squared(ANOVA_results)
```

```
## # Effect Size for ANOVA
##
## Parameter | Eta2 |      95% CI
## -----
## DoseFactor | 0.54 | [0.42, 1.00]
##
## - One-sided CIs: upper bound fixed at [1.00].
```

```
interpret_eta_squared(0.54, rules = "cohen1992")
```

```
## [1] "large"
## (Rules: cohen1992)
```

$$\eta^2 = \frac{k_k}{k_y} = \frac{216.5}{216.5 + 183.5} = 0.54$$

Cohen (1992) ("cohen1992") applicable to one-way anova, or to *partial* eta / omega / epsilon squared in multi-way anova.

- **ES < 0.02** - Very small
- **0.02 <= ES < 0.13** - Small
- **0.13 <= ES < 0.26** - Medium
- **ES >= 0.26** - Large

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

If the assumption of homogeneity of variances is violated, we use Welch's  $F$ -test, which is less affected by the conditions of heteroskedasticity.

```
library(onewaytests)
welch.test(Happiness ~ DoseFactor,
           data = mydata)

## 
##   Welch's Heteroscedastic F Test (alpha = 0.05)
## -----
##   data : Happiness and DoseFactor
##
##   statistic : 65.16615
##   num df    : 2
##   denom df   : 57.50688
##   p.value    : 1.655095e-15
##
##   Result     : Difference is statistically significant.
## -----
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
pairwise.t.test(x = mydata$Happiness, g = mydata$DoseFactor,  
                 p.adj = "none")
```

```
##  
##  Pairwise comparisons using t tests with pooled SD  
##  
## data: mydata$Happiness and mydata$DoseFactor  
##  
##          Control 15 minutes  
## 15 minutes 7.5e-05 -  
## 30 minutes 3.3e-16 2.3e-08  
##  
## P value adjustment method: none
```

```
pairwise.t.test(x = mydata$Happiness, g = mydata$DoseFactor,  
                 p.adj = "bonf")
```

```
##  
##  Pairwise comparisons using t tests with pooled SD  
##  
## data: mydata$Happiness and mydata$DoseFactor  
##  
##          Control 15 minutes  
## 15 minutes 0.00023 -  
## 30 minutes 9.9e-16 6.8e-08  
##  
## P value adjustment method: bonferroni
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

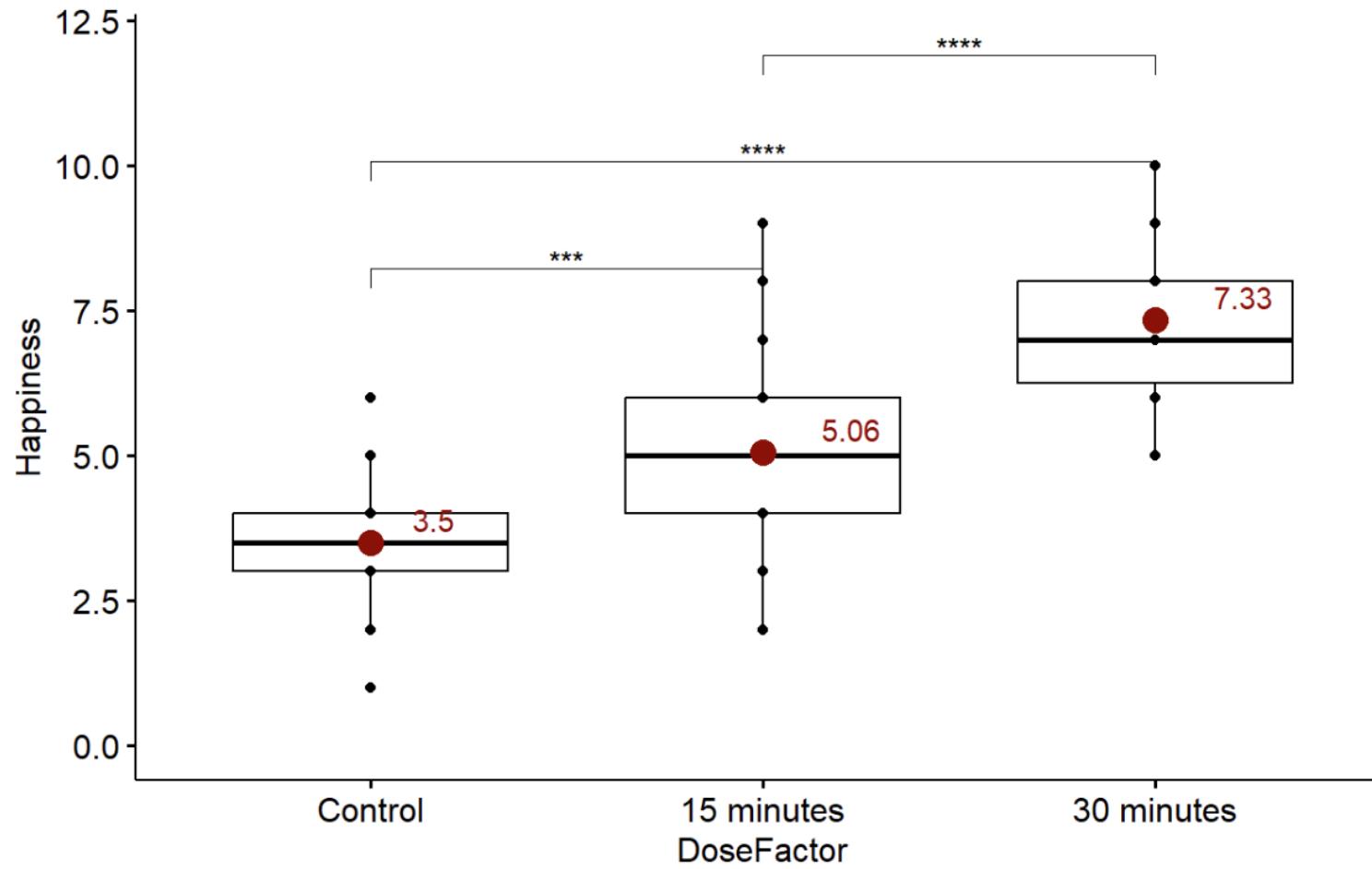
## Example: Puppies

```
pwc <- mydata %>%  
  pairwise_t_test(Happiness ~ DoseFactor,  
                  paired = FALSE,  
                  p.adjust.method = "bonferroni")  
  
ANOVA_results <- anova_test(Happiness ~ DoseFactor,  
                             data = mydata)  
  
library(rstatix)  
pwc <- pwc %>%  
  add_y_position(fun = "median", step.increase = 0.35)  
  
library(ggpubr)  
ggboxplot(mydata, x = "DoseFactor", y = "Happiness", add = "point", ylim=c(0, 12)) +  
  stat_pvalue_manual(pwc, hide.ns = FALSE) +  
  stat_summary(fun = mean, geom = "point", shape = 16, size = 4,  
               aes(group = DoseFactor), color = "darkred",  
               position = position_dodge(width = 0.8)) +  
  stat_summary(fun = mean, colour = "darkred",  
               position = position_dodge(width = 0.8),  
               geom = "text", vjust = -0.5, hjust = -1,  
               aes(label = round(after_stat(y), digits = 2), group = DoseFactor)) +  
  labs(subtitle = get_test_label(ANOVA_results, detailed = TRUE),  
       caption = get_pwc_label(pwc))
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

Anova,  $F(2,87) = 51.3$ ,  $p = <0.0001$ ,  $\eta_g^2 = 0.54$



pwc: **T test**; p.adjust: **Bonferroni**

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

Conclusions:

We found a statistically significant relationship between puppy therapy and average happiness level ( $F = 51.3$ ,  $p < 0.001$ ), the effect size is large ( $\eta^2 = 0.54$ ). Post hoc tests revealed differences for each pair of groups ( $p < 0.001$ ).

# KRUSKAL-WALLIS RANK SUM TEST

## Example: Puppies

```
kruskal.test(Happiness ~ DoseFactor,  
             data = mydata)
```

```
##  
## Kruskal-Wallis rank sum test  
##  
## data: Happiness by DoseFactor  
## Kruskal-Wallis chi-squared = 49.612, df = 2, p-value = 1.686e-11
```

```
kruskal_effsize(Happiness ~ DoseFactor,  
                 data = mydata)
```

```
## # A tibble: 1 × 5  
##   .y.          n effsize method  magnitude  
## * <chr>     <int>    <dbl> <chr>    <ord>  
## 1 Happiness    90    0.547 eta2[H] large
```

# KRUSKAL-WALLIS RANK SUM TEST

## Example: Puppies

```
library(rstatix)

groups_nonpar <- wilcox_test(Happiness ~ DoseFactor,
                               paired = FALSE,
                               p.adjust.method = "bonferroni",
                               data = mydata)
```

```
groups_nonpar
```

```
## # A tibble: 3 × 9
##   .y.    group1    group2     n1     n2 statistic      p   p.adj p.adj.signif
## * <chr>   <chr>    <chr>    <int>   <int>     <dbl>    <dbl>   <dbl>   <chr>
## 1 Happiness Control 15 minutes    28     32      218 5.48e- 4 2   e- 3   **
## 2 Happiness Control 30 minutes    28     30       14 2.02e-10 6.06e-10 ****
## 3 Happiness 15 minutes 30 minutes  32     30      142 1.49e- 6 4.47e- 6 ****
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

```
pwc <- mydata %>%
  wilcox_test(Happiness ~ DoseFactor,
               paired = FALSE,
               p.adjust.method = "bonferroni")

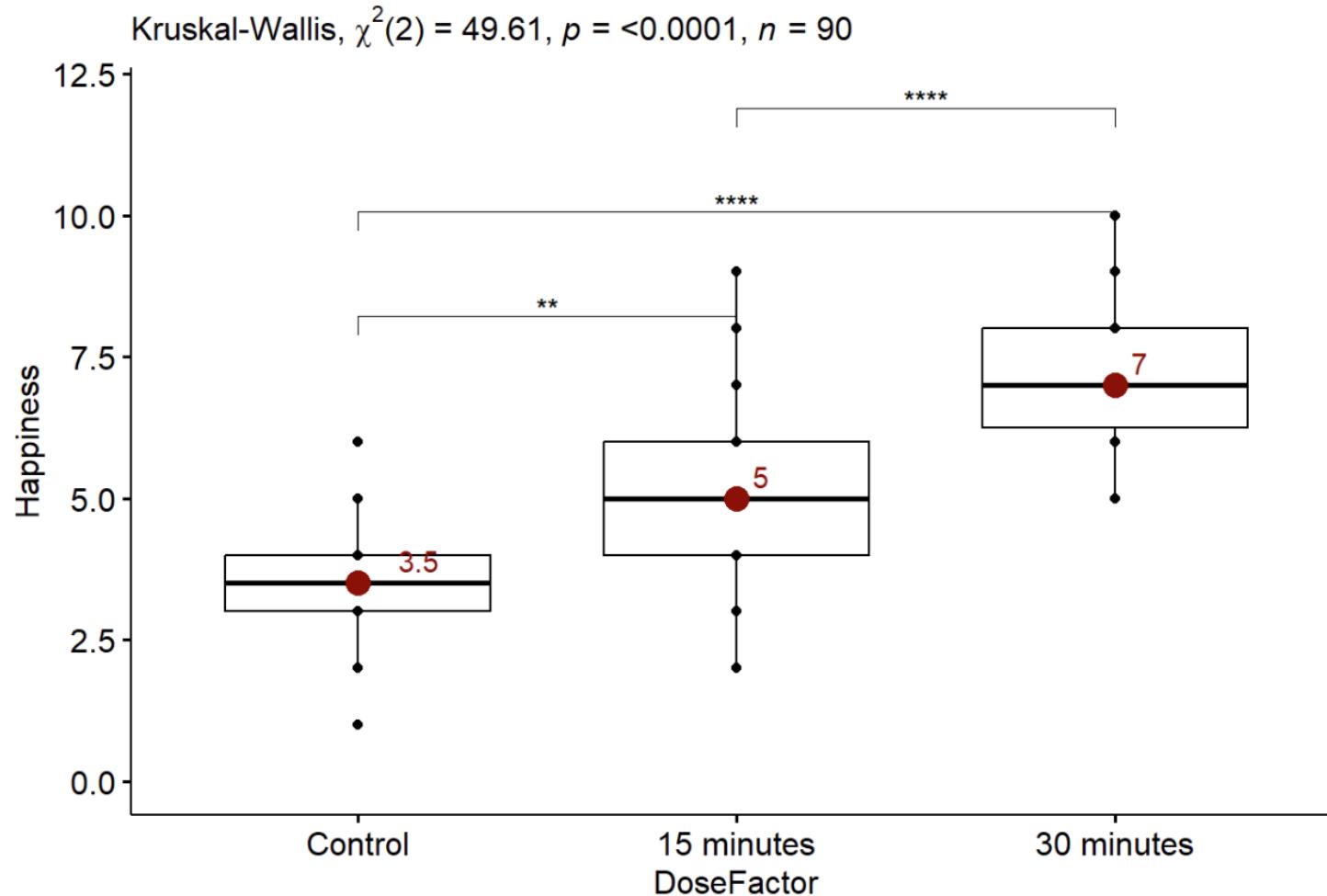
Kruskal_results <- kruskal_test(Happiness ~ DoseFactor,
                                 data = mydata)

library(rstatix)
pwc <- pwc %>%
  add_y_position(fun = "median", step.increase = 0.35)

library(ggpubr)
ggboxplot(mydata, x = "DoseFactor", y = "Happiness", add = "point", ylim=c(0, 12)) +
  stat_pvalue_manual(pwc, hide.ns = FALSE) +
  stat_summary(fun = median, geom = "point", shape = 16, size = 4,
               aes(group = DoseFactor), color = "darkred",
               position = position_dodge(width = 0.8)) +
  stat_summary(fun = median, colour = "darkred",
               position = position_dodge(width = 0.8),
               geom = "text", vjust = -0.5, hjust = -1,
               aes(label = round(after_stat(y), digits = 2), group = DoseFactor)) +
  labs(subtitle = get_test_label(Kruskal_results, detailed = TRUE),
       caption = get_pwc_label(pwc))
```

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies



pwc: Wilcoxon test; p.adjust: Bonferroni

# ONE-WAY ANALYSIS OF VARIANCE – ANOVA

## Example: Puppies

Conclusions:

We found that the distribution of happiness levels differs for at least one of the puppy doses ( $\chi^2 = 49.6$ ,  $p < 0.001$ ), the effect size was high ( $\eta^2 = 0.55$ ). Post-Hoc tests revealed differences for each pair of groups ( $p < 0.01$ ).

# TEST OF POPULATION PROPORTION

## BINOMIAL TEST

## TEST OF EQUALITY IN TWO POPULATION PROPORTIONS

# TEST OF POPULATION PROPORTION

Testing a population proportion with the chosen parameter.

$$H_0: \pi = \pi_0$$

$$H_1: \pi \neq \pi_0$$

$$\chi^2 = \left( \frac{p - \pi_0}{SE(p)} \right)^2 \quad p = \frac{n_a}{n} \quad SE(p) = \sqrt{\frac{\pi_0(1 - \pi_0)}{n}} \quad m = 1$$

Assumption:  $n\pi_0 > 5$  and  $n(1 - \pi_0) > 5$

# TEST OF POPULATION PROPORTION

## Example 1

The sample of 1,870 children born at the maternity hospital over the last year included 965 boys and 905 girls. Can you conclude that there are more boys than girls in the newborn population?

```
prop.test(x = 965,
          n = 1870,
          p = 0.5,
          correct = FALSE,
          alternative = "greater")
```

```
##  
## 1-sample proportions test without continuity correction  
##  
## data: 965 out of 1870, null probability 0.5  
## X-squared = 1.9251, df = 1, p-value = 0.08265  
## alternative hypothesis: true p is greater than 0.5  
## 95 percent confidence interval:  
## 0.4970246 1.0000000  
## sample estimates:  
## p  
## 0.5160428
```

# BINOMIAL TEST

## Example 2

The inspector visited 33 businesses in a particular food industry and found serious violations in 9 of them. Can she conclude from this that the proportion of undertakings in the industry in question is more than 0.10, which is the proportion of infringements at the last inspection five years ago?

```
binom.test(x = 9,
            n = 33,
            p = 0.10,
            alternative = "greater")
```

```
##  
##  Exact binomial test  
##  
## data: 9 and 33  
## number of successes = 9, number of trials = 33, p-value = 0.004134  
## alternative hypothesis: true probability of success is greater than 0.1  
## 95 percent confidence interval:  
## 0.1502883 1.0000000  
## sample estimates:  
## probability of success  
##                 0.2727273
```

# TEST OF EQUALITY IN TWO POPULATION PROPORTIONS

Testing the equality in two population proportions.

$$H_0: \pi_1 = \pi_2$$

$$H_1: \pi_1 \neq \pi_2$$

# TEST OF EQUALITY IN TWO POPULATION PROPORTIONS

## Example 3

Researcher surveyed ( $n = 323$ ) the general population about their use of online banking. Of 234 people under 40, 195 use online banking, while of the 89 people over 40, 45 use online banking. Is there a difference in the proportion of online banking users between the younger and older populations?

```
prop.test(c(195, 45), c(234, 89),
          correct = TRUE,
          alternative = "two.sided")

##
## 2-sample test for equality of proportions with continuity correction
##
## data: c(195, 45) out of c(234, 89)
## X-squared = 34.571, df = 1, p-value = 4.11e-09
## alternative hypothesis: two.sided
## 95 percent confidence interval:
##  0.2056395 0.4497912
## sample estimates:
##   prop 1   prop 2
## 0.8333333 0.5056180
```