Simple MCMC under SIR

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We describe how to set-up and run a simple Metropolis-Hastings-based Markov chain Monte Carlo (MCMC) sampler under the susceptible-infected-removed (SIR) model.

```
library(MultiBD)
```

This example uses the Eyam data that consist the population counts of susceptible, infected and removed individuals across several time points.

```
data(Eyam)
Eyam
```

```
time
           S
## 1
     0.0 254 7
                  0
    0.5 235 14
## 3 1.0 201 22
     1.5 153 29
                79
     2.0 121 20 120
## 6 2.5 110 8 143
## 7 3.0 97
              8 156
     4.0
         83
              0 178
```

The log likelihood function is the sum of the log of the transition probabilities between two consecutive observations. Note that, we will use $(\log \alpha, \log \beta)$ as parameters instead of (α, β) . The rows and columns of the transition probability matrix returned by $dbd_prob()$ correspond to possible values of S (from a to a0) and I (from 0 to B) respectively.

```
loglik_sir <- function(param, data) {</pre>
  alpha <- exp(param[1]) # Rates must be non-negative</pre>
  beta <- exp(param[2])
  # Set-up SIR model
  drates1 <- function(a, b) { 0 }</pre>
  brates2 <- function(a, b) { 0 }</pre>
  drates2 <- function(a, b) { alpha * b</pre>
  trans12 <- function(a, b) { beta * a * b }</pre>
  sum(sapply(1:(nrow(data) - 1), # Sum across all time steps k
             function(k) {
                log(
                  dbd_prob( # Compute the transition probability matrix
                    t = data$time[k + 1] - data$time[k], # Time increment
                    a0 = data\$S[k], b0 = data\$I[k],
                                                             # From: S(t_k), I(t_k)
                    drates1, brates2, drates2, trans12,
                    a = data\$S[k + 1], B = data\$S[k] + data\$I[k] - data\$S[k + 1],
                    computeMode = 4, nblocks = 80
                                                             # Compute using 4 threads
```

Here, we choose Normal(0, 100^2) as the prior for both $\log \alpha$ and $\log \beta$.

```
logprior <- function(param) {
  log_alpha <- param[1]
  log_beta <- param[2]

dnorm(log_alpha, mean = 0, sd = 100, log = TRUE) +
   dnorm(log_beta, mean = 0, sd = 100, log = TRUE)
}</pre>
```

We will use the random walk Metropolis algorithm implemented in the function MCMCmetrop1R() (MCMCpack package) to explore the posterior distribution. So, we first need to install the package and its dependencies.

```
source("http://bioconductor.org/biocLite.R")
biocLite("graph")
biocLite("Rgraphviz")
install.packages("MCMCpack", repos = 'http://cran.us.r-project.org')
library(MCMCpack)
```

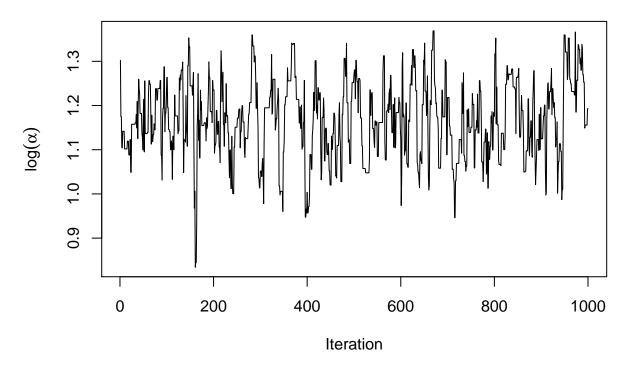
The starting point of our Markov chain is the estimated value of (α, β) from Raggett (1982).

```
alpha0 <- 3.39
beta0 <- 0.0212
```

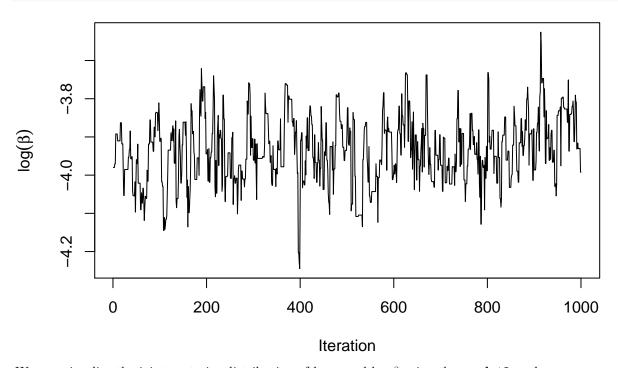
We discard the first 200 iterations and keep the next 1000 iterations of the chain.

The trace plots of both $\log \alpha$ and $\log \beta$ look good.

```
plot(as.vector(post_sample[,1]), type = "l", xlab = "Iteration", ylab = expression(log(alpha)))
```



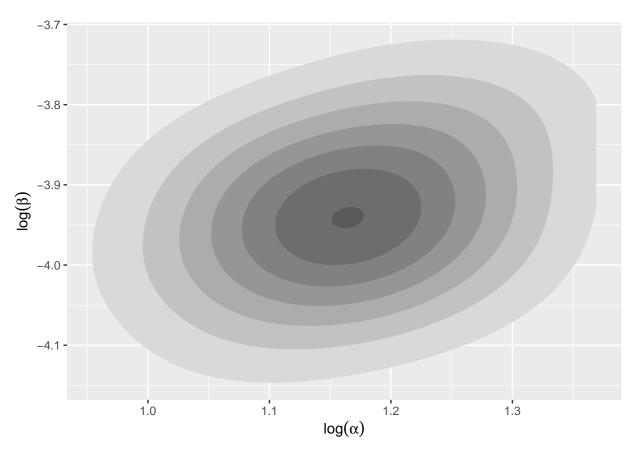
plot(as.vector(post_sample[,2]), type = "l", xlab = "Iteration", ylab = expression(log(beta)))



We can visualize the joint posterior distribution of $\log \alpha$ and $\log \beta$ using the **ggplot2** package.

```
library(ggplot2)
x = as.vector(post_sample[,1])
y = as.vector(post_sample[,2])
df <- data.frame(x, y)
ggplot(df,aes(x = x,y = y)) +
   stat_density2d(aes(fill = ..level..), geom = "polygon", h = 0.26) +</pre>
```

```
scale_fill_gradient(low = "grey85", high = "grey35", guide = FALSE) +
xlab(expression(log(alpha))) +
ylab(expression(log(beta)))
```



We can also construct the 95% Bayesian credible intervals for α and β .

```
quantile(exp(post_sample[,1]), probs = c(0.025,0.975))
```

```
## 2.5% 97.5%
## 2.721921 3.809780
```

```
quantile(exp(post_sample[,2]), probs = c(0.025,0.975))
```

```
## 2.5% 97.5%
## 0.01649866 0.02327176
```