

# Shrinkage and regression to the mean

André Meichtry

October 28, 2021

Shrinkage of results can be seen  
to be a necessary fact of life.

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(Stephen Senn)

## 1. Shrinkage is a fact of life

Assume true diastolic blood pressure,  $\tau$ , at baseline is measured with error  $\epsilon$  so that

$$X = \tau + \epsilon \quad (1)$$

is the *observed* blood pressure. Let the *true* mean difference between patients be  $\Delta$  and the *observed* mean difference  $D$ , then the expectation of  $D$  is

$$E(D \mid \Delta = \delta) = \delta, \quad (2)$$

that is, we have “classical” *unbiasedness*, since  $\beta_{D|\Delta} = \frac{\text{Cov}(\Delta, D)}{\sigma_\Delta^2} = 1$ .

However, the contrary is not true. We have for the expectation of  $\Delta$ , given an observed difference  $d$ ,

$$|E(\Delta \mid D = d)| < |d|, \quad (3)$$

since  $\beta_{\Delta|D} = \frac{\text{Cov}(\Delta, D)}{\sigma_D^2} < 1$ , and we have *regression to the mean*<sup>1</sup>.

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**In a Bayesian approach,** shrinking is *natural* and we have *inverse unbiasedness*. Most bayesians are rather unconcerned about unbiasedness (at least in the formal sampling-theory sense above) of their estimates. For example, Gelman et al (1995) write: “From a Bayesian perspective, the principle of unbiasedness is reasonable in the limit of large samples, but otherwise it is potentially misleading. Unbiasedness as conventionally understood is not a necessary property of good inferences”. Assume without loss of generality  $E(\Delta) = 0$  and  $\hat{\Delta}$  an unbiased estimate of a given effect  $\Delta$  and  $\hat{\Delta}_{shrunk}$  a shrunk estimate. Although  $\hat{\Delta}_{shrunk}$  is not unbiased in the classic forward sense,  $E(\hat{\Delta}_{shrunk} \mid \Delta = \delta) \neq \delta$  it is unbiased in the Bayesian backward sense:  $E(\Delta \mid \hat{\Delta}_{shrunk} = \hat{\delta}_{shrunk}) = \hat{\delta}_{shrunk}$ .

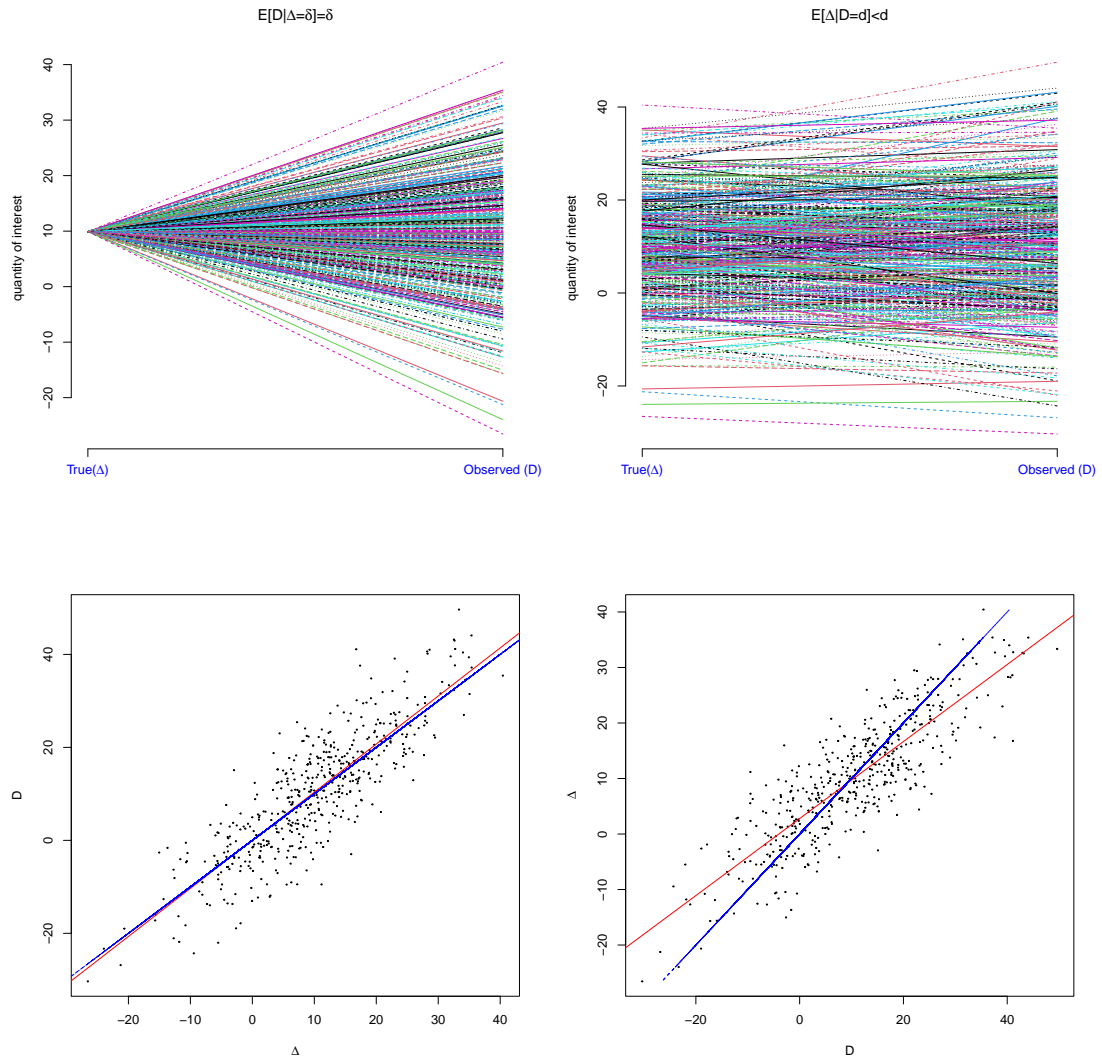


Figure 1: Shrinkage as a fact of life

**Reliability as upper bound** The maximal possible correlation between  $\Delta$  and  $D$  is  $\sqrt{rel_D}$ .

## 2. Placebo as a statistical phenomenon

Placebo effects can often be interpreted as a purely statistical – not a psychological – phenomenon.

**Assuming no true change.** We simulate correlated pre-post diastolic blood pressure data assuming *no change* from baseline to follow-up: simulations from parameters:  $\rho_{BL,FU} = 0.76, \mu_{BL} = \mu_{FU} = 90, \sigma_{BL} = \sigma_{FU} = 8$ . Then let us look at the *subgroup* of “hypertensive at baseline” only. We have regression to the mean, since  $\beta_{FU|BL} = \frac{\sigma_{FU,BL}}{\sigma_{BL}^2} = r \leq 1$ .

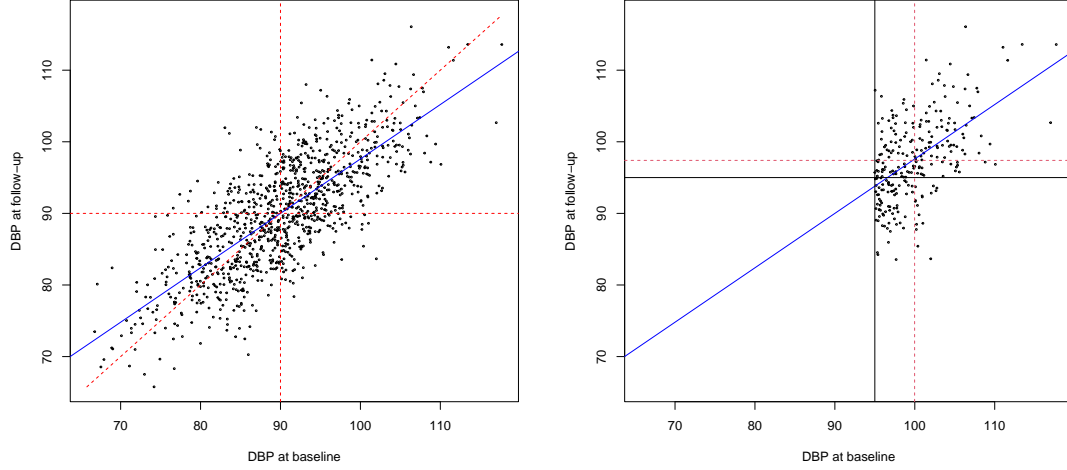
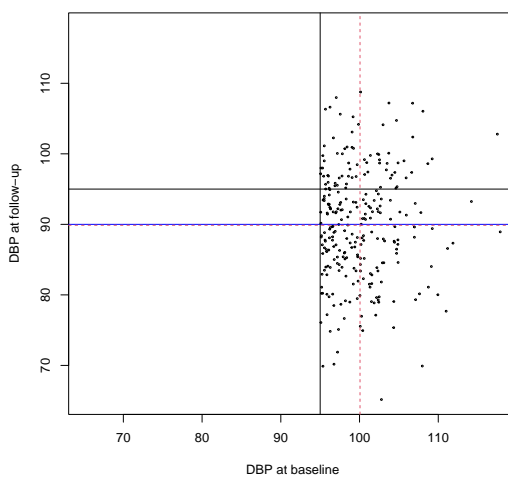
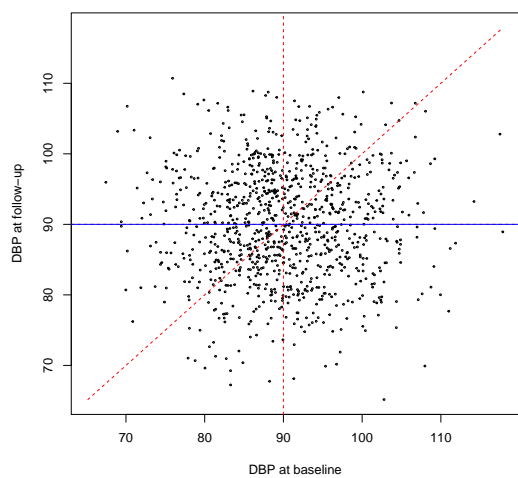
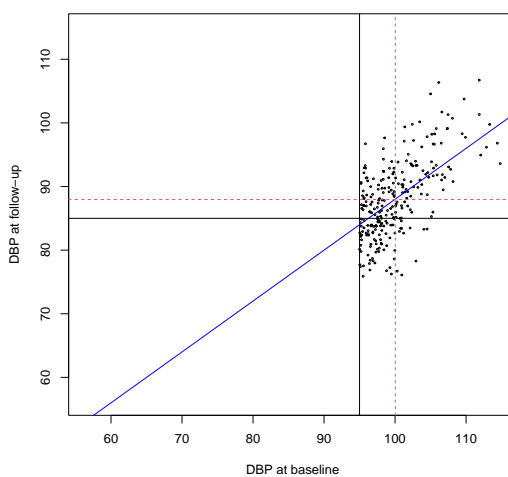
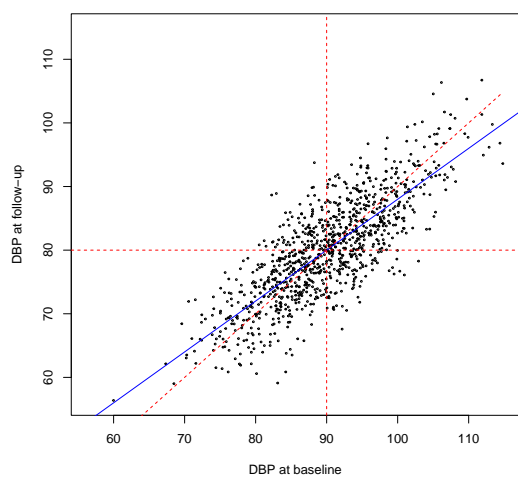


Figure 2: Simulation of diastolic blood pressure data. Simulations from parameters:  $\rho_{BL,FU} = 0.76, \mu_{BL} = \mu_{FU} = 90, \sigma_{BL} = \sigma_{FU} = 8$ . Left panel: Baseline versus Follow-up for diastolic blood pressure: no change in the mean. Right panel: Baseline versus Follow-up for “hypertensive at baseline” only. We observe an apparent change due to regression to the mean (Solid line: Regression of follow-up on baseline-measure (that is, by fixing baseline)). Dashed lines: mean values and equality lines.

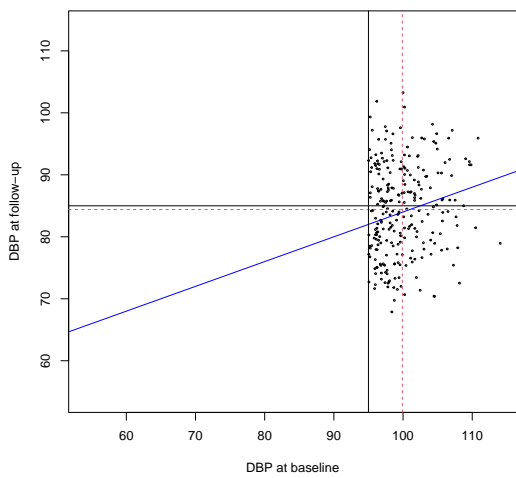
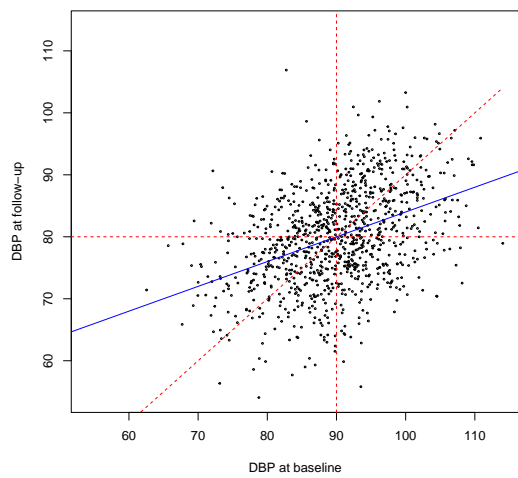
**Extreme case:  $\rho=0$**



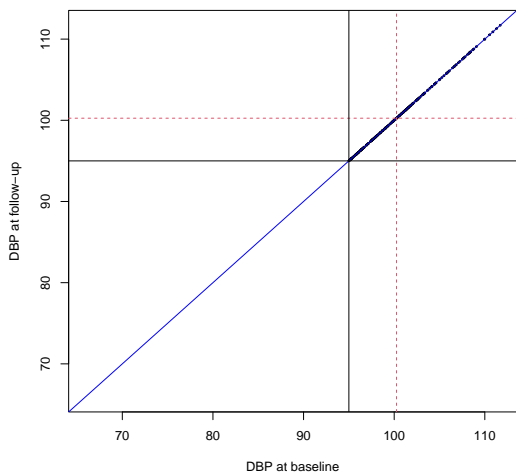
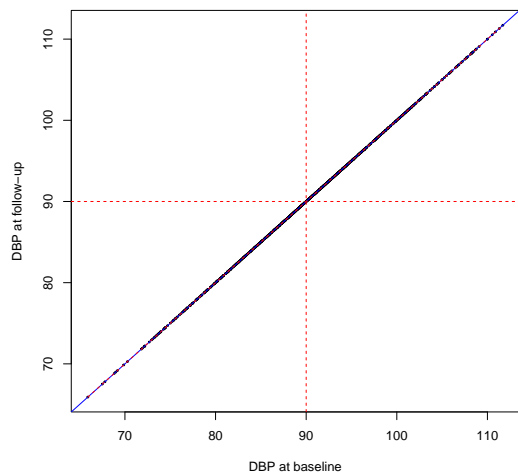
**Including a true change of -10 and  $\rho=.8$**



Including a true change of -10 and  $\rho=.4$



$\rho=1$  (Perfect reliability)



## A. Maths

The conditional expected value of  $Y$ , given that  $X$  is  $t$  standard deviations above its mean (and that includes the case where it is below its mean, when  $t < 0$ ), is  $\rho$  standard deviations above the mean of  $Y$ .

Since  $|\rho| \leq 1$ ,  $Y$  is no farther from the mean than  $X$  is, as measured in the number of standard deviations. Hence, if  $0 \leq \rho < 1$ , then  $(X, Y)$  shows regression toward the mean.

$$\boxed{\frac{E(Y | X) - E(Y)}{\sigma_Y} = \rho \frac{X - E(X)}{\sigma_X}}, \quad (4)$$

that is,  $z_{Y|X} = \rho z_x$ , leading to  $z_{Y|X} - z_X = (\rho - 1)z_x$ . The amount of RTM is

$$\boxed{|z_{Y|X} - z_x| = (1 - \rho)z_x}. \quad (5)$$

The estimated fraction of RTM is given by  $1 - \rho$ , the fraction of variance that is due to within-subject variability. This quantity represents *unreliability*.