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## pypomp: Inference for partially observed Markov process models in Python with JAX DRAFT IN PROGRESS

Aaron J. Abkemeier\*
University of Michigan

Jun Chen\*
Duke University

**Kevin Tan** University of Pennsylvania Jesse Wheeler Idaho State University

Bo Yang Columbia University Kunyang He University of Michigan

Jonathan Terhorst University of Michigan Aaron A. King

Edward L. Ionides University of Michigan \*These authors contributed equally

#### Abstract

—!!!—an abstract is required—!!!—

Keywords: —!!!—at least one keyword is required—!!!—.

#### 1. Introduction

[Topic] Partially Observable Markov Process (POMP) models, also known as state-space models or hidden Markov models, provide a flexible and mechanistic framework for modeling

time-series dynamic systems, particularly suited for scenarios where latent states are only partially observable. Characterized by transition densities and measurement densities of Markov processes, this framework bridges complex underlying dynamics with limited information in real-world data. Consequently, POMP models find extensive application in epidemiology (Mietchen, Clancey, McMichael, and Lofgren 2024; Fox, Lachmann, Tec, Pasco, Woody, Du, Wang, Ingle, Javan, Dahan, Gaither, Escott, Adler, Johnston, Scott, and Meyers 2022; Wen, Yin, Li, Peng, and He 2024), ecology (Auger-Méthé, Newman, Cole, Empacher, Gryba, King, Leos-Barajas, Mills Flemming, Nielsen, Petris, and Thomas 2021; Marino, Peacor, Bunnell, Vanderploeg, Pothoven, Elgin, Bence, Jiao, and Ionides 2019; Blackwood, Streicker, Altizer, and Rohani 2013), finance (Bretó 2014), and other domains.

[Existing package discussion] The rich POMP package ecosystem built in R has provided a solid, standardized, and extensible framework for modeling time series data using nonlinear, stochastic partially observed mechanism dynamic models. The R pomp package has become a well-established tool for fitting POMP models using a general and abstract representation, that supports multiple inference techniques. Its extension packages panelPomp, spatpomp, and phyloPomp further enhance its capabilities for panel, spatio-temporal and phylodynamic data analysis respectively

[computational challenges + potential limitations] While conceptually powerful, statistical inference for POMP models using the above R packages poses substantial computational challenges. From a methodological perspective, likelihood-based inference for POMP models typically relies on perturbations within iterated filtering (adding ref related to iterative filtering) algorithms. While many of them are stable and effective in locating a neighborhood containing the likelihood maximum, they exhibit numerical inefficiency for obtaining a precise identification of the maximum value. Particularly, when the latent states are high-dimensional or when repeated model evaluations are required, the fitting process could be computationally prohibitive by the constraints.

On the other hand, these POMP models have demonstrated strong potential to be sped up considerably. Many of the processes are embarassingly parallel, such as simulating the state process for each of thousands of particles, running the particle filter multiple times for the same parameter set, and running iterated filtering from multiple starting parameter sets, especially when estimating a profile likelihood to construct a confidence interval as per (Ionides, Breto, Park, Smith, and King 2017). Graphics Processing Units (GPUs) are well-suited for such operations. However, the existing family of POMP packages (pomp, panelPomp, and spatPomp (King, Nguyen, and Ionides 2016, Bretó, Wheeler, King, and Ionides (2025), Asfaw, Park, King, and Ionides (2024))) only runs on CPUs.

[introducing AD/JIT/GPU hardware + pypomp] As the demand grows for scalable and parallelizable inference algorithms, there is an increasing need for an accelerated framework for POMP modeling framework. pypomp(Abkemeier, Chen, Ionides, Wheeler, and Tan 2024) addresses this by tightly integrating automatic differentiation (AD), just-in-time (JIT) compilation, and GPU/TPU acceleration through JAX (Bradbury, Frostig, Hawkins, Johnson, Leary, Dougal Maclaurin, Necula, Paszke, VanderPlas, Skye Wanderman-Milne, and Zhang 2018), a high-performance numerical computing library that supports hardware acceleration and vectorization. AD techniques enable efficient and accurate derivative computation by systematically applying the chain rule to fundamental operations. Recent works have extended AD for gradient estimation using the particle filtering, yielding a class of meth-

ods termed automatic differetiation particle filters (ADPF). Building on these advancements, pypomp provides AD-enabled gradient estimation for inference for POMP models within a plug-and-play framework (Ionides, Breto, and King 2006), where users specify dynamic models solely through simulators of latent state trajectories, rather than evaluating the transition density of the latent Markov process. JAX's JIT compilation further accelerates repeated inference and evaluation by converting Python functions into machine code at runtime. Finally, JAX's GPU/TPU support enables efficient scaling of inference methods. Together, these features make pypomp more than a port of the R package pomp. Instead, it established an independent, modernized platform for fast and flexible POMP modeling in Python.

#### [structure]

The remainder of this paper is organized as follows. Section 2 discusses the Motivation for pypomp design using specific examples. Section 3 demonstrates the mathematical notation for POMP models and their related implementation in pypomp. Section 4 introduces the embedded methodologies. Section 5 presents data analysis workflows and benchmarking results. Section 6 concludes with a discussion of future directions.

[add discussion of panelpomp into introduction]

## 2. Motivation for pypomp

[This section is for some extra detailed numeric cost estimates and dataset descriptions to illustrate motivation based on Aaron's draft. It is a bit redundant now.]

#### 2.1. Real-world computational bottleneck

Computational speed is a major bottleneck in the practical application of iterated filtering methods to POMP models. In Korevaar, Metcalf, and Grenfell (2020) 's dataset, fitting and evaluating likelihoods of POMP models for 180 units required 8 days on 36 CPU cores (two 3.0 GHz Intel Xeon Gold 6154 CPUs). Scaling this up to the full dataset of 1422 units would require almost eight times as much effort, equivalent to running 36 cores for two months or 288 cores for 8 days. This is not only time consuming, but also incurs substantial computational costs, highlighting the urgent need for more efficient inference software for large-scale POMP analyses Importantly, this cost only accounts for one round of iterated filtering. In practice, to further refine the likelihood estimates, multple rounds are required, which would increase the computational burden significantly. This motivates the development of accelerated, scalable tools to make large-scale POMP inference feasible.

## 2.2. Opportunities for speeding up the POMP models

Many of the processes involved in fitting POMP models are embarrassingly parallel. Examples include simulating the state process for each of thousands of particles, running the particle filter repeatedly under the same parameter set, and executing iterated filtering from multiple starting parameter sets. Such parallelism is especially advantageous when estimating a profile likelihood to construct confidence intervals (Ionides et al. 2017). Harnessing parallel computing resources can therefore dramatically reduce computation time and make large-scale inference feasible.

Graphics Processing Units (GPUs) are well-suited for embarrassingly parallel operations, but the existing family of POMP packages (**pomp**, **panelPomp**, and **spatPomp** (King *et al.* 2016, Bretó *et al.* (2025), Asfaw *et al.* (2024))) are limited to CPU computation. None provide support for GPU acceleration or automatic differentiation. These two technologies are key to enabling scalable and efficient inference for modern POMP applications.

#### 2.3. Our solution: pypomp

To address this computational bottleneck, we are creating pypomp(Abkemeier et al. 2024), a python implementation of the R package pomp. It draws inspiration from pomp, but further implements new methods incoporating automatic differentiation techniques by forking the source code used in Tan (Tan, Hooker, and Ionides 2024), as well as leverages JAX's just-in-time(JIT) compilation and GPU core parallelization (Bradbury et al. 2018), allowing practitioners to run filtering methods significantly faster and cheaper. For example, in an SPX comparison model, we show that, compared to pomp with 36 CPU cores, pypomp can run at least 7 times faster and can finish the job at 5% of the price using 1 GPU and 1 CPU core (5120 CUDA cores on a NVIDIA Tesla V100 and one core from a 2.4 GHz Intel Xeon Gold 6148 CPU).

In addition, pypomp is gradually including functionality from panelPomp and spatPomp, offering a unified Python interface for entire POMP methodologies across multiple R packages. It also takes advantage of JAX's implementation of automatic differentiation (AD), which can be used in conjunction with the differentiable measurement off-parameter with discount factor  $\alpha$  (MOP- $\alpha$ ) particle filter to improve local optimization of the likelihood surface (Tan et al. 2024).

#### 2.4. Summary of key features

Table 1 summarizes the main differences between pypomp and pomp, highlighting the new capabilities of pypomp.

Feature	pypomp	pomp
Backend and Acceleration	JAX (GPU/CPU, JIT,vmap,	R and C Snippets (CPU
	<pre>jax.grad, jax.Hessian)</pre>	only)
Automatic Differentiation	Yes (gradient/Hessian via	No
and gradient-based inference	AD supported)	
Particle Filtering Methods	Yes (PF, MOP- $\alpha$ , IF2,	Yes (PF, IF2, pMCMC, etc.)
	IFAD)	
Plug-and-Play Property	Yes	Yes
Object Design	In-place updates on current	Returns new objects
	objects, stored in the object	
	attribute results_history	

Table 1: Feature comparison between pypomp and pomp in R ecosystem.

## 3. POMP Models in pypomp

This section introduces the structure of POMP models and its implementation in pypomp, including both mathematical setup and the package implementation.

#### 3.1. Model setup

A partially observed Markov process (POMP) model has two main components: (i) a latent Markov process that evolves over time and (ii) an observation process that links the latent states. Together, these jointly specify the mechanistic model for the observed time series, providing a framework for modeling dynamic systems where measurements are noisy. Formally, suppose we observe the process at discrete time n = 1, ..., N, with initial time 0. Let  $Y_{1:N}$  denote the observations and  $X_{0:N}$  denote the corresponding latent (unobserved) Markov process at the corresponding time. A POMP model is defined by three building components, each corresponding to user-supplied Python functions in pypomp:

- 1. initial density:  $f_{X_0}(x_0;\theta)$ , implemented via rinit, generates draws of the initial states  $X_0$ .
- 2. transition density:  $f_{X_n|X_{n-1}}(x_n\mid x_{n-1};\theta)$ , implemented via rproc, propagates the evolution of latent states.
- 3. measurement density:  $f_{Y_n|X_n}(y_n \mid x_n; \theta)$ , implemented via dmeas, evaluates or simulates observations conditional on the latent states.

Here, we hold an conditional independence assumption that, given  $X_n$ , the observation  $Y_n$  is independent of all other variables. Then, the joint density of  $(X_{0:N}, Y_{1:N})$  can be expressed as the product of the initial distribution, the transition densities, and the measurement densities:

$$f_{X_{0:N},Y_{1:N}}(x_{0:N},y_{1:N};\theta) = f_{X_0}(x_0;\theta) \prod_{n=1}^N f_{X_n\mid X_{n-1}}(x_n\mid x_{n-1};\theta) \prod_{n=1}^N f_{Y_n\mid X_n}(y_n\mid x_n;\theta)$$

The marginal likelihood of the observations is  $\mathcal{L}(\theta) = f_{Y_{1:N}}(y_{1:N};\theta) = \int f_{X_{0:N},Y_{1:N}}(x_{0:N},y_{1:N};\theta)\,dx_{0:N}$ . In practice, this integral is intractable for most nonlinear or non-Gaussian POMP models, motivating the use of simulation-based inference methods such as particle filtering.

#### 3.2. Implementations of POMP models in pypomp

#### Object-oriented interface

A POMP model in pypomp is represented as an object of class Pomp, which encapsulates the model components: the initial state distribution, process model, and measurement model. This object-oriented interface allows users to specify by passing components to the constructor, including observations, model parameters, model mechanics such as simulators and the measurement density, covariates, and times. After the components are passed into the constructor, the constructor automatically generates additional internal elements, such as extended observations and covariates required for interpolation

Table 2 summarizes the main arguments to the Pomp constructor and their correspondence to mathematical objects. RInit, RProc and RMeas are the subconstructors, where they specify the stochastic mechanisms that define a Pomp model, and their objects are passed as arguments when initializing a main Pomp class.

Constructor/Sub-			Description / Mathematical
constructor	Argument	Type	representation
Pomp	rinit	RInit	simulate initial states $X_0 \sim f_{X_0}(x_0; \theta)$
	rproc	RProc	simulate state transitions
			$X_n \sim f_{X_n \mid X_{n-1}}(x_n \mid x_{n-1}; \theta)$
	rmeas	RMeas	simulate observations
			$Y_n \sim f_{Y_n \mid X_n}(y_n \mid x_n; \theta)$
	dmeas	DMeas	evaluate measurement density
			$f_{Y_n X_n}(y_n \mid x_n; \theta)$
	ys	pandas.DataFrame	observations $y_{1:N}^*$ with times $t_{1:N}$
	covars	pandas.DataFrame	covariates $z_{1:N}^*$ with times $s_{1:N}$
	theta	list or dict	parameters $\theta$
RInit	t0	float	initial time point $t_0$ for simulation
RProc	step_type	str	method of process evolution:
			"fixedstep" or "euler"
	nstep	int	number of steps if
			step_type="fixedstep"
	dt	float	time step if step_type="euler"
	accumvars	tuple	indices of state variables to be
			accumulated
$\mathbf{RMeas}$	ydim	int	observation dimension $\dim(Y)$

Table 2: Main arguments to the Pomp class and related constructor objects.

We demonstrate here how to create a Pomp object. Specifically, we show how to create the linear Gaussian model included in the package as LG(). We begin by importing necessary packages and defining helper functions for handling the parameters. Because pypomp will run our defined model components within JAX JIT-compiled code, it is necessary to write the components to be JAX-compliant. Naturally, the JAX package has many useful functions for this purpose. We also generate a pseudorandom number generation (PRNG) key to be used with JAX random number generators. All stochastic simulations in pypomp are controlled via JAX PRNG keys, ensuring full reproducibility when using the same seed.

```
import pypomp as pp
import pandas as pd
import jax
import jax.numpy as jnp
from functools import partial
def get_thetas(theta):
```

```
theta = jnp.asarray(theta)
A = theta[0:4].reshape((2, 2))
C = theta[4:8].reshape((2, 2))
Q = theta[8:12].reshape((2, 2))
R = theta[12:16].reshape((2, 2))
return A, C, Q, R

def transform_thetas(A, C, Q, R):
    return jnp.concatenate([A.ravel(), C.ravel(), Q.ravel(), R.ravel()])
# create PRNG key correctly
key = jax.random.PRNGKey(1)
```

#### Model Components

We refer to model components describing initialization, transfer, or measurement processes as model mechanisms, including rinit, rproc, dmeas, and rmeas. Users must define these processes as Python functions. Specifically, we require users to provide function code to the object constructor, which verifies that all necessary function arguments are included and in the correct order. This requirement stems from pypomp's internal mechanism: it vectorizes component functions using jax.vmap() to efficiently run thousands of particles. Since jax.vmap() maps functions to input arrays by position rather than keyword, users must strictly adhere to parameter order. While all expected parameters must be included, the function does not need to utilize all of them.

Illustrated in Table 2,pypomp also includes object constructors for components describing the model mechanics: RInit, RProc, DMeas, and RMeas. Some constructors also require additional arguments, such as t0 for RInit. Notably, RProc takes step\_type, dt, and nstep arguments. step\_type determines how RProc should be run at intermediate steps between two observation times. If we want to model the state process as evolving in continuous time, setting step\_type="euler"uses an Euler approximation, running rproc at intermediate steps based on the time step size, dt. The number of steps taken is given by the number of times dt divides the difference between two observation times, rounded up, and is consequently dynamic. Otherwise, if we instead want a fixed number of steps for each observation time interval, we can use step\_type="fixedstep", in which case rproc will run at nstep intermediate steps equally spaced between two observation times, starting from the first observation time. Consequently, setting step\_type="fixedstep" and nstep=1 only runs rproc at the observation times. Here is an example of defining the object constructors for components under the linear gaussian model.In practice, at least one of dmeas or rmeas must be provided, while the construction of RInit and RProc are always required.

```
import pypomp as pp

@partial(pp.RInit, t0=0.0)
def rinit(theta_, key, covars=None, t0=None):
```

```
A, C, Q, R = get_thetas(theta_)
return jax.random.multivariate_normal(key=key, mean=jnp.array([0.0, 0.0]), cov=Q)

@partial(pp.RProc, step_type="fixedstep", nstep=1)
def rproc(X_, theta_, key, covars=None, t=None, dt=None):
    A, C, Q, R = get_thetas(theta_)
    return jax.random.multivariate_normal(key=key, mean=A @ X_, cov=Q)

@pp.DMeas
def dmeas(Y_, X_, theta_, covars=None, t=None):
    A, C, Q, R = get_thetas(theta_)
    # return logpdf of Y given X (mean = C @ X_, cov = R)
    return jax.scipy.stats.multivariate_normal.logpdf(Y_, mean=C @ X_, cov=R)

@partial(pp.RMeas, ydim=2)
def rmeas(X_, theta_, key, covars=None, t=None):
    A, C, Q, R = get_thetas(theta_)
    return jax.random.multivariate_normal(key=key, mean=C @ X_, cov=R)
```

#### **Parameters**

The Pomp constructor also requires model parameters. These can be provided either as a dictionary or as a list of dictionaries. Each item in a dictionary should include the parameter name as the key and the parameter value as the dictionary value. If the parameter sets are provided as a list of dictionaries, methods such as pfilter() run on each set of parameters. Here, we use Pomp.sample\_params() to sample sets of parameters from uniform distributions with bounds passed as a dictionary of length-2 tuples. Pomp.sample\_params() returns a ready-to-use list of dictionaries with the sampled parameters. Internally, parameters, even are multi-dimensional, are stored as flat dictionaries to facilitate JAX transformations and compilation.

```
theta = {
    "A11": jnp.cos(0.2), "A12": -jnp.sin(0.2),
    "A21": jnp.sin(0.2), "A22": jnp.cos(0.2),
    "C11": 1.0, "C12": 0.0, "C21": 0.0, "C22": 1.0,
    "Q11": 0.01, "Q12": 1e-6, "Q21": 1e-6, "Q22": 0.01,
    "R11": 0.1, "R12": 0.01, "R21": 0.01, "R22": 0.1,
}
param_bounds = {k: (v * 0.9, v * 1.1) for k, v in theta.items()}
n = 5
key = jax.random.PRNGKey(1)
key, subkey = jax.random.split(key)
theta_list = pp.Pomp.sample_params(param_bounds, n, subkey)
```

#### 3.3. Covariates

Scientifically, POMP models often involve external time-varying inputs, referred to as covariates, which can influence either the latent process or the measurement model. Examples include seasonality, interventions, or environmental drivers in ecological applications. In pypomp, covariates are supplied as a pandas.DataFrame indexed by time. The time at which the covariates were observed should be specified in the ctime argument. Importantly, the covariate time points may differ from the observation times, necessitating interpolation. Given the observation times, covariate times, and the step type specified in RProc, the model automatically aligns and interpolates observations and covariates to ensure consistency with the simulation of the latent and observation processes. The linear gaussian model doesn't involve any covariates, and an example using covariates is given in the Data Analysis Section.

#### POMP Object Construction

We do not have real data in this LG example, so we generate our own. To make this example cleaner, we here use the function LG() to construct the completed linear Gaussian model object and then generate the data using simulate(). Observation times are provided to the Pomp constructor via the pandas.DataFrame row index. If covariates were provided, the times at which the covariates were observed would also be provided by the pandas.DataFrame row index.

```
import jax, jax.numpy as jnp
import pandas as pd
import pypomp as pp
T = 100
# ensure `key` exists; if not, uncomment the next line
# key = jax.random.PRNGKey(1)
key, subkey = jax.random.split(key)
sims = pp.LG(T=T).simulate(key=subkey)
ys = pd.DataFrame(
    sims[0]["Y_sims"].squeeze(),
    index=range(1, T + 1),
    columns=["Y1", "Y2"],
)
LG_obj = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    ys=ys,
    theta=theta_list,
    covars=None,
)
```

print("LG\_obj created; ys.shape =", ys.shape)

```
LG_obj created; ys.shape = (100, 2)
Each argument to Pomp is accessible from the object as an attribute.
print(LG_obj.rinit) # access POMP model components
print(LG_obj.rproc)
print(LG_obj.dmeas)
print(LG_obj.rmeas)
print(LG_obj.theta)
                      # access parameters
print(LG_obj.ys.head()) # access observations
<pypomp.model_struct.RInit object at 0x12a9a67e0>
<pypomp.model struct.RProc object at 0x12aef9a90>
<pypomp.model_struct.DMeas object at 0x12aefa330>
<pypomp.model_struct.RMeas object at 0x12af103b0>
[{'A11': 0.9211453199386597, 'A12': -0.17880238592624664, 'A21': 0.21495014429092407, 'A22'
         Y1
1 0.243581 -0.125824
2 0.094966 0.149981
3 0.614238 0.213606
4 0.955786 0.462541
5 0.242421 0.490216
```

#### Premade models:

Beyond the linear gaussian model, pypomp includes several ready-to-use model constructors that serve both as examples and as tested templates for custom model development:

- 1. LG() a simple linear-Gaussian model with 2-dimensional latent and observed states; useful to validate API usage and diagnostics.
- 2. spx() the S&P500 log-return model from Sun et al. (Sun 2024).
- 3. dacca() the cholera transmission model from King et al. (King, Ionides, Pascual, and Bouma 2008).
- 4. UKMeasles.Pomp() the measles district model from He et al. (He, Ionides, and King 2010), wired to the Korevaar et al. dataset (Korevaar et al. 2020). Panel and spatial variants (PanelPOMP/SpatPOMP style) are planned.

These examples show correct component wiring (rinit, rproc, dmeas, rmeas), recommended step\_type/dt usage, and typical diagnostics. If a user model errors or runs slowly, compare its components to the matching premade model to find mistakes and performance opportunities. Meanwhile, these premade models can also replicate well-know case studies in the R pomp ecosystem, allowing direct comparison and validation.

## JAX Numerical Backend and Interface Design

A key design choice pypomp is it relys heavily on the JAX numerical backend. Unlike the R package pomp, where users typically provide POMP model components in C Snippets for acceleration, pypomp requires model components to be written as JAX-compatible Python functions. These functions are then compiled and vectorized by JAX tools such as jit and vmap. This design leads to several important interface features:

- Strict argument requirements for compilation and vectorization: JAX's jit compiler transforms the user-supplied component functions (rinit, rproc, dmeas, rmeas) into efficient machine code, while vmap efficiently run them over thousands of particles via vectorization of arguments. To ensure the compatibility with JAX's compilation and vectorization system, each component function must follow the expected input types and order, otherwise compilation would fail.
- PRNG random key policy: To ensure the reproducibility of randomness in POMP models under pypomp, the public API accept an optional jax.random.PRNGkay, which is explicitly passed through constructors and methods. Keys are internally split when it is needed. Unlike the R setting, where randomness can be controlled globally or by seed chunks, in JAX, random keys only be explicitly passed through functions
- Consistent shapes and sizes handling: model parameters, even multidimensional, are stored as flattened dictionaries. Consequently, JAX can uniformly process parameters, thereby maintaining consistency in particle propagation.

Later section will demonstrate how the JAX-based design supports further inference methods. [more introduction to JAX?]

Panel POMP class

## 4. POMP Methods in pypomp

In this section, we describe the core inference methods currently implemented in pypomp, including:

- Paricle Filter (Sequential Monte Carlo, written in pfilter()): A standard sequential Monte Carlo algorithm for likelihood evaluation and state estimation, forming the basis for most inference methods in POMP models.
- Differentiated Measurement-off-policy Particle Filter  $(MOP(\alpha), written in mop())$ : A recently proposed SMC method (Tan et al. 2024) that evaluates the likelihood at one parameter value while obtaining resampling decisions from another, adjusting via discounted off-parameter measurement weights.
- Iterated Filtering (IF2, mif()): A classical IF2 algorithm (Ionides, Nguyen, Atchadé, Stoev, and King 2015) for likelihood-based parameter inference that maximizes the likelihood via particle filtering.

• Iterated Filtering with Automatic Differentiation (train()): A recently proposed AD-based algorithm (Tan et al. 2024) that incorporates  $MOP(\alpha)$ , the differentiable particle filter, to enable efficient gradient-based parameter inference for maximum likelihood estimation.

A key feature of the above POMP inference methods lies in the **plug-and-play property** (Ionides *et al.* 2006), meaning that inference algorithms can be implemented without requiring explicit evaluation of the transition density of the latent process. Instead, it suffices for the user to provide a simulator of the latent process (rproc), initial state distribution (rinit), and observation measurement model (dmeas, rmeas). This property enables POMP methods to be widely applied to complex mechanistic models where transition densities are intractable.

In pypomp, the plug-and-play design is fully preserved: users only need to provide component functions compliant with JAX requirements, which can be directly plugged in inference methods likepfilter(), mop(), mif(), and train(). The package combines the generality of plug-and-play modeling with the efficiency of JAX compilation and vectorization.

Unlike the R family of POMP packages, some Pomp class methods including pfilter(), mif() and train() yield results by modifying the object in place instead of returning new objects. All of results are stored a list under LG\_obj.results\_history, which is an attibute under Pomp class object LG\_obj. Each element in the list correponds to one method call. Each element includes results such as the log-likelihood and parameter estimates when applicable as well as the inputs used for the function call, so it is easy to keep track of how the results were calculated. If multiple parameter sets are supplied in a list as an argument, the method evaluates at each set and the results for each are stored. Meanwhile, each following method in pypomp that takes a key as an argument stores an unused child key under LG\_obj.fresh\_key that later method calls can use by default when a key argument is not given. This design prevents repeated use of the same key across a sequence of method calls and ensures proper randomness.

```
[-75.0522 , -75.3827 , -73.30542 , -71.4641 , -75.1468 ]], dtype=float32,
Dimensions without coordinates: theta, replicate, 'theta': [{'A11': 0.9211453199386597, 'A
array([[[
                    nan, 9.21145320e-01, -1.78802371e-01,
         2.14950144e-01, 8.89322698e-01, 1.02344239e+00,
         0.00000000e+00, 0.00000000e+00, 9.79293942e-01,
         1.04345558e-02, 9.33975798e-07, 9.76017759e-07,
         1.05938269e-02, 9.44515765e-02, 9.41129215e-03,
         9.19735525e-03, 9.43083838e-02],
        [-0.00000000e+00, 1.07045341e+00, -2.18700305e-01,
         2.06873283e-01, 9.45706606e-01, 8.82167339e-01,
         -3.03104311e-01, -1.46514282e-01, 5.65519333e-01,
         1.62466958e-01, 1.77885965e-01, -6.56164214e-02,
         3.40645671e-01, 1.02493912e-01, 3.30706313e-02,
        -6.84480742e-02, 1.20453514e-01],
        [-1.24636581e+02, 8.16644430e-01, -5.28744869e-02,
         5.82232058e-01, 6.72007203e-01, 9.57372367e-01,
        -3.92328411e-01, -2.65173197e-01, 4.19366837e-01,
         2.80792415e-01, 2.61916459e-01, 6.33998886e-02,
         3.07527870e-01, 7.97012299e-02, 6.75280914e-02,
        -6.22278303e-02, 1.40667498e-01]],
       nan, 9.74913478e-01, -1.78802371e-01,
        -2.03950927e-02, 1.04405761e-01]],
       nan, 9.70987082e-01, -1.78802371e-01,
         2.18500718e-01, 9.31446433e-01, 1.07719481e+00,
         0.0000000e+00, 0.0000000e+00, 9.35545087e-01,
         1.07558817e-02, 1.06478592e-06, 9.44516046e-07,
         1.02092260e-02, 1.01782054e-01, 9.41056665e-03,
         9.20325704e-03, 9.58210155e-02],
        [-0.00000000e+00, 8.51299584e-01, -2.22326946e-02,
         2.04024658e-01, 1.01018143e+00, 7.58982956e-01,
         1.84635725e-02, 2.20715329e-01, 9.56478238e-01,
         2.77179211e-01, -1.28381521e-01, -7.14511871e-02,
         1.35017648e-01, 6.87062517e-02, 1.59603179e-01,
        -1.26104221e-01, 1.22075811e-01],
        [-1.17136009e+02, 7.31945157e-01, -1.86973646e-01,
         1.58347413e-01, 9.49631929e-01, 5.89739382e-01,
         2.68075258e-01, 2.40069270e-01, 1.08920443e+00,
         2.85785407e-01, -8.81665424e-02, -2.44462746e-03,
         9.77937952e-02, 6.82703108e-02, 8.02902132e-03,
        -2.59932708e-02, 1.85035437e-01]]])
Coordinates:
  * replicate (replicate) int64 40B 0 1 2 3 4
  * iteration (iteration) int64 24B 0 1 2
  * variable (variable) <U6 408B 'logLik' 'A11' 'A12' ... 'R12' 'R21' 'R22', 'theta': [{
```

## 4.1. Particle Filter (pfilter)

[outline: 1. purpose/role 2. implementation details in pypomp 3. outputs/results 4. remarks/highlights]

The particle filter algorithm (sequential Monte Carlo, SMC) (Arulampalam, Maskell, Gordon, and Clapp (2002), King et al. (2016)) is the core classical likelihood evaluation tool in POMP model inference. It uses Monte Carlo techniques to sequentially estimate the integrals in the prediction and filtering recursions, and produce an unbiased Monte Carlo estimate of the likelihood of the observed data under a parameter set. The likelihood estimate is essential for other parameter inference algorithms such as iterative filtering. In pypomp, the basic particle filter is implemented under pfilter(). The algorithm is summarized as Algorithm 1,

Algorithm 1 Sequential Monte Carlo (SMC, or particle filter) in LG\_obj.pfilter(J=J, reps=reps, key=key), where LG\_obj is a class Pomp object with components rinit, rproc, dmeas, rmeas, ys, and theta

**Require** Simulator for process<sub>n</sub>(·|x<sub>n</sub>;  $\theta$ ); Evaluator for  $f_{Y_n|X_n}(y_n \mid x_n; \theta)$ ; Simulator for  $f_{X_0}(x_0; \theta)$ ; Parameter  $\theta$ ; Data  $y_{1:N}^*$ ; Number of particles J.

**Initialize** particles:  $X_{0,j}^{F,\theta} \sim f_{X_0}(\cdot;\theta)$  for j=1:J.

For n = 1, ..., N:

Simulate prediction particles:  $X_{n,j}^{P,\theta} \sim \operatorname{process}_n(\cdot|x_n;\theta)$  for j=1:J.

Evaluate weights and measurement density:  $w(n,j) = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P,\theta}; \theta)$  for j = 1:J.

Normalize weights 
$$\tilde{w}(n,j) = \frac{w(n,j)}{\sum_{m=1}^{J} w(n,m)}$$
 for  $j=1:J$ . Select resample indices  $k_{1:J}$  with  $P(k_j=m)=\tilde{w}(n,m)$ .

Obtain resampled particles  $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$  for j = 1:J.

Compute conditional log likelihood:

$$\hat{\ell}_{n|1:n-1} = \log \left(\frac{1}{J} \sum_{m=1}^J w(n,m)\right).$$

**Return** Log likelihood estimate  $\hat{\ell}(\theta) = \sum_{n=1}^{N} \hat{\ell}_{n|1:n-1}$ ; filter samples  $X_{n,1:J}^{F,\theta}$  for n=1:N. Complexity:  $\mathcal{O}(J)$ 

In pypomp, the pfilter() functions is internally run in a compiled JAX routine \_pfilter\_internal is pfilter\_internal() but wrapped up into a class method. JIT-compiled with static arguments, in which the replicates and parameter sets are vectorized via jax.vmap. Both techniques enable execution on CPU/GPU and ensure the computational efficiency. In implementation, the weights and likelihood are stored in a log scale, preventing underflow. There are optional arguments (CLL, ESS, filter\_mean, prediction mean) control wheter additional diagnostic arrays are allocated and updated. The algorithm works for flexible and extendable models through the user-specified Pomp class components (rinit, rproc, dmeas, reams), number of particles J as well as other parameters. Each run returns a dict type element updated inside the LG\_obj. result\_history attribute, containing the log-likelihoods, algorithm parameters used, as well as model

diagonostic elements at each time included if their respective boolean flags are set to True. For example, suppose we run

```
LG_{obj_2} = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    ys=ys,
    theta=theta_list,
    covars=None,
)
LG_{obj_2.pfilter}(J = 1000,
                  reps = 10,
                  key = subkey)
LG_{obj_2.pfilter}(J = 1000,
                  reps = 10,
                  key = subkey,
                  CLL = True,
                  ESS = True,
                  filter_mean = True,
                  prediction_mean = True)
```

where J is the number of particles used and reps is the number of particle filtering replicates to run for each parameter set provided in the Pomp object or as an optional argument to pfilter(). Because LG\_obj2.result\_history begins as an empty list here when the model is constructed, the results are appended at LG\_obj\_2.results\_history[0] and LG\_obj\_2.results\_history[1] respectively. Both of these two dictionaries contain with the following items:

- method: The method that was run. In this case, pfilter.
- logLiks: An xarray.DataArray of particle filter log-likelihood estimates. There is one list for each parameter set given and each list has reps log-likelihood estimates. dimensions = (theta, replicate)
- theta: The list of parameter sets used.
- J: The number of particles used.
- thresh: Threshold value for resampling used, which is 0 by default.

Meanwhile, LG\_obj\_2.results\_history[1] also contains the following items that are not contained in LG\_obj\_2.results\_history[0]:

• CLL: An xarray.DataArray containing the conditional loglikelihood over observed time steps (theta, replicate, time)

- ESS: An xarray.DataArray containing the effective sample size over observed time steps (theta, replicate, time)
- filter\_mean: An xarray.DataArray containing the filtered mean over observed time steps (theta, replicate, time, state)
- prediction\_mean: An xarray.DataArray containing the prediction mean over observed time steps. (theta, replicate, time, state)

A key design is that all runs across multiple parameter sets and replications are stored in one results\_history list under the corresponding Pomp object. This avoids object duplication, and increases reproducibility by facilitating pickling Pomp objects obtained from many different search strategies, loading them into a new Python session, and then comparing the results to see how they vary with different algorithmic parameters.

#### 4.2. Differentiated Measurement Off-Parameter (DMOP) particle filter

([TO BE REWRITTEN, FOLLOWING ISSUE 2])

[Motivation for DMOP] Classical particle filters for POMP models produce unbiased Monte Carlo estimates, but are non-differentiable in model parameters  $\theta$ , due to the discrete random sampling of particles. This prevents the wide use of AD and gradient-based optimization in POMP dynamical systems. The differentiable measurement off-parameter particle filter (DMOP- $\alpha$ , Tan et al. (2024)) provides a solution: it treats a differentiable likelihood estimator produced by a variant of the standard particle filter as an objective and apply AD to obtain  $\nabla_{\theta} \hat{\ell}(\theta)$  for likelihood-based optimal parameter search. The derivative estimate from DMOP- $\alpha$  is essential for further stochastic gradient inference methods, such as iterated filtering with automatic differentiation (IFAD), where it employs stochastic gradient descent using DMOP- $\alpha$ . A tunning dicount factor  $\alpha \in [0,1]$  controls the bias-variance trade-off of the gradient estimator: larger  $\alpha$  may result in lower bias and higher variance, while smaller  $\alpha$  can reduce the variance with the cost of increasing bias. In implementation, DMOP- $\alpha$  does not require heavy diagnostics or multiple replicates. It only needs a stable ,differentiable of the (log-) likelihood to feed into jax.grad and jax.value\_and\_grad. The pseudocode is shown in 2

Algorithm 2 DMOP-α in pypomp: jax.value\_and\_grad(LG\_obj.mop(J=J, reps=reps, key=key, alpha=alpha)). LG\_obj is a Pomp object with components rinit, rproc, dmeas, ys, theta.

For n = 1, ..., N:

Simulate prediction particles,  $X_{n,j}^{P,\theta} \sim \operatorname{process}_n(\;\cdot\;|X_{n-1,j}^{F,\theta};\theta)$  for j=1:J. Accumulate discounted prediction weights,  $w_{n,j}^{P,\theta}=\left(w_{n-1,j}^{F,\theta}\right)^{\alpha}$ .

Evaluate measurement density,  $g_{n,j}^{\theta} = f_{Y_n|X_n}(y_n^*|X_{n,j}^{P,\theta};\theta)$  for j=1:J. Compute conditional likelihood in version B:  $L_n^{B,\theta,\alpha} = \sum_{j=1}^J g_{n,j}^{\theta} w_{n,j}^{P,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}$ .

Select resampling indices  $k_{1:J}$  with  $P(k_j = m) \propto g_{n,m}^{\theta}$ .

Obtain resampled particles  $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$ .

Calculate corrected weights  $w_{n,j}^{F,\theta} = w_{n,k_j}^{P,\theta} g_{n,k_j}^{\theta} / \text{stop\_gradient}(g_{n,k_j}^{\theta})$ .

Compute considtional likelihood in version A:  $L_n^{A,\theta,\alpha} = L_n^{\phi} \cdot \sum_{j=1}^J w_{n,j}^{F,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}$ .

Forward Pass: AD constructs a computation graph to evaluate the likelihood estimate  $\hat{L}^A(\theta) = \prod_{n=1}^N L_n^{A,\theta,\alpha}$  or  $\hat{L}^B(\theta) = \prod_{n=1}^N L_n^{B,\theta,\alpha}$ .

Backward Pass: AD propagates gradients through the computation graph and evaluates

gradient estimate  $\hat{\mathcal{D}}^{A}(\theta)$  or  $\hat{\hat{\mathcal{D}}}^{\bar{B}}(\theta)$ 

**Return:** likelihood estimate  $\hat{L}^{A}(\theta)$  or  $\hat{L}^{B}(\theta)$ , and gradient estimate  $\hat{\mathcal{D}}^{A}(\theta)$  or  $\hat{\mathcal{D}}^{B}(\theta)$ .

The key building block of the DMOP- $\alpha$  is the MOP- $\alpha$  algorithm, a variant of the standard particle filter designed to make the discontinuous Monte Carlo resampling process independent of the target parameter  $\theta$ . It evaluates the likelihood at the evaluation parameter  $\theta$  but draws resampling indices using weights computed at another baseline behavior parameter  $\phi$ . This renders the likelihood estimates differentiable w.r.t.  $\theta$ , further enabling AD.

In DMOP- $\alpha$ , the same idea is implemented equivalently by setting  $\phi \equiv \mathtt{stop\_gradient}(\theta)$ in JAX. This prevent its input from being differentiated. The filtering process therefore runs  $\theta = \phi$ , assuming the evaluation parameter and the bahavior parameters have identical values. This avoids the need to explicitly introduce new parameters while preserving the differentiable property. Consequently, the likelihood evaluation of DMOP is identical to that of the standard particle filter, and it does not require implement a fully separate MOP- $\alpha$  procedure. This design simplifies the algorithm and still ensures the valid likelihood and gradient estimator, while maintaining differentiability with respect to  $\theta$ . That is, DMOP computes the derivative of the likelihood estimate obtained by MOP when  $\theta = \phi$ . The pseudocode of MOP is shown in 3.

#### **Algorithm 3** MOP( $\alpha$ ), Measurement off-policy sequential Monte Carlo

**Require** Simulator for  $\operatorname{process}_n(\cdot|x_n;\theta)$ ; Evaluator for  $f_{Y_n|X_n}(y_n \mid x_n;\theta)$ ; Simulator for  $f_{X_0}(x_0;\theta)$ ; Evaluation Parameter  $\theta$ ; Behavior Parameter  $\phi$ ; Data  $y_{1:N}^*$ ; Number of particles J.

```
\begin{split} &J_{X_0}(x_0;\theta); \text{ Evaluation Parameter } \theta; \text{ Behavior Parameter } \phi; \text{ Data } y_{1:N}; \text{ Number of particles } J. \\ &\textbf{Initialize } \text{filter particles } X_{0,j}^{F,\theta} \sim f_{X_0}(\cdot;\theta) \text{ and } w_{0,j}^{F,\theta} = 1 \text{ for } j = 1 \cdot J. \\ &\textbf{For } n = 1, ..., N \cdot \\ &\textbf{Simulate prediction particles: } X_{n,j}^{P,\theta} \sim f_{X_n|X_{n-1}}(\cdot \mid X_{n-1,j}^{F,\theta};\theta) \text{ for } j = 1 \cdot J. \\ &\textbf{Accumulate discounted Prediction weight: } w_{n,j}^{P,\theta} = \left(w_{n-1,j}^{F,\theta}\right)^{\alpha} \text{ for } j = 1 \cdot J. \\ &\textbf{Evaluate measurement density: } g_{n,j}^{\theta} = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P,\theta};\theta) \text{ for } j = 1 \cdot J. \\ &\textbf{Compute conditional likelihood under } \theta \text{ in version } B \cdot L_n^{B,\theta,\alpha} = \sum_{j=1}^J g_{n,j}^{\theta} w_{n,j}^{P,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}. \\ &\textbf{Compute conditional likelihood under } \phi \cdot L_n^{\theta} = \frac{1}{J} \sum_{m=1}^J g_{n,m}^{\phi}. \\ &\textbf{Select resample indices } k_{1:J} \text{ with } P(k_j = m) \propto \tilde{g}_{n,m}^{\theta}. \\ &\textbf{Obtain resampled particles } X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta} \text{ for } j = 1 \cdot J. \\ &\textbf{Calculate resampled corrected weights:} w_{n,j}^{F,\theta} = w_{n,j}^{P,\theta} \times \frac{g_{n,j}^{\theta}}{g_{n,j}^{\theta}} \text{ for } j = 1 \cdot J. \\ &\textbf{Compute conditional likelihood under } \theta \text{ in version } A \cdot L_n^{A,\theta,\alpha} = L_n^{\phi} \cdot \sum_{j=1}^J w_{n,j}^{F,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}. \\ &\textbf{Return: Likelihood estimate: } \hat{L}(\theta) = \prod_{n=1}^N L_n^{A,\theta,\alpha} \text{ or } \hat{L}(\theta) = \prod_{n=1}^N L_n^{B,\theta,\alpha}. \end{aligned}
```

In pypomp, DMOP- $\alpha$  is implemented by wrapping the JAX-compiled mop() function inside JAX AD. It is a Pomp class method calling an internal \_mop\_internal(), which uses static

arguments and jax.lax.fori\_loop for efficiency, and stores weights and likelihoods in log scale to prevent underflow. System sampling is involved for the resampling process. Users

can directly apply jax.grad and jax.value and grad to mop() to obtain gradients.

In practice, the mop() provides an auto-differentiable alternative to pfilter(), as a lightweight version. Unlike pfilter(), which provides the choice of multiple replicates and diagnostic summaries, mop() in pypomp focuses purely on producing differentiable likelihood estimates. This makes the algorithm suitable for AD and stochastic gradient descent algorithms in accelerated iterated filtering algorithms by setting the function as input of jax.value\_and\_grad() and jax.grad(). Beyond a well-constructed Pomp object, the method requires only the number of particles J, the evaluation parameter theta, a PRNG key, and the cooling factor  $\alpha$ . Results are stored in the results\_history attribute as a list of jax.Array objects, one for each parameter set, representing the estimated log-likelihood of the observations. For example, we run

```
LG_obj_3 = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    ys=ys,
    theta=theta_list,
    covars=None,
)
```

```
LG_obj_3.mop(J = 1000,

key = subkey,

alpha = 0.97)
```

```
[Array(-76.55094, dtype=float32, weak_type=True), Array(-71.051605, dtype=float32, weak_type=True), Array(-75.82673, dtype=float32, weak_type=True), Array(-71.13629, dtype=float32, weak_type=True), Array(-72.63167, dtype=float32, weak_type=True)]
```

Generally, the result\_history only updates:

• log-likelihood(s): a list of jax. Array containing the estimated log-likelihood(s) of the observed data given the model parameters. Always a list, even if only one theta is provided.

This design highlights mop() as a lightweight but essential building block for accelarated gradient-based inference methods within pypomp. By combining jax.grad and jax.value\_and\_grad around mop(), the package effectively implements the DMOP algorithm, enabling stochastic-gradient updates for likelihood-based inference on POMP models.

#### 4.3. Iterated Filtering

Iterated filtering (IF2) provides a Monte Carlo approach to estimate parameters by combining particle filtering with sequential random-walk perturbations and cooling. Compared with a plain particle filter or MOP method, IF2 not only evaluates the likelihoods but also systematically updates parameters across iterations. Its primary role is to estimate the parameters that maximizes the likelihood, given the likelihood is intractable but can be estimated under the particle filtering methods. The pseudocode for IF2 algorithm is as follows:

#### Algorithm 4 Iterated Filtering (IF2)

```
\textbf{Require:} \ \ \text{Starting parameter} \ \ \theta_0; \ \ \text{simulator for} \ \ f_{X_0}(x_0;\theta); \ \ \text{simulator for} \ \ f_{X_n|X_{n-1}}(x_n \ \ | \ \ )
      (x_{n-1};\theta); evaluator for f_{Y_n\mid X_n}(y_n\mid x_n;\theta); data y_{1:N}^*; labels I\subset\{1,\dots,p\} for IVPs; fixed lag
      L; number of particles J; number of iterations M; cooling rate a, 0 < a < 1; perturbation
      scales \sigma_{1:p}; initial scale multiplier C > 0.
  1: for m = 1 to M do
            Initialize parameters: \Theta_{0,j,i}^P \sim \text{Normal}([\theta_{m-1}]_i, (Ca^m\sigma_i)^2) for i \in 1:p, j \in 1:J.
  2:
            Initialize states: simulate X_{0,j}^F \sim f_{X_0}(\cdot;\Theta_{0,j}^P) for j=1:J.
 3:
            Initialize filter mean: \bar{\theta}_0 = \theta_{m-1}
  4:
            Define [V]_i = (C^2 + 1)a^{2m}\sigma_i^2.
  5:
            for n = 1 to N do
  6:
                  Perturb parameters: \Theta_{n,j,i}^P \sim \text{Normal}([\Theta_{n-1,j}^F]_i, (a^m \sigma_i)^2) for i \notin I, j = 1:J.
  7:
                  Simulate prediction particles: X_{n,j}^P \sim f_{X_n|X_{n-1}}(\cdot \mid X_{n-1,j}^F; \Theta_{n,j}^P) for j=1:J.
  8:
                 Evaluate weights: w(n,j) = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^P; \Theta_{n,j}^P) for j=1:J.

Normalize weights: \tilde{w}(n,j) = \frac{w(n,j)}{\sum_{u=1}^J w(n,u)}.

Resample indices k_{1:J} with \Pr[k_u=j] = \tilde{w}(n,j).

Resample particles: X_{n,j}^F = X_{n,k_j}^P and \Theta_{n,j}^F = \Theta_{n,k_j}^P for j=1:J.
  9:
10:
11:
12:
                  Filter mean: [\bar{\theta}_n]_i = \sum_{j=1}^J \tilde{w}(n,j) [\Theta_{n,j}^P]_i for i \notin I.
13:
                  Prediction variance: [V_{n+1}]_i = (a^m \sigma_i)^2 + \sum_{j=1}^J \tilde{w}(n,j) \left( [\Theta_{n,j}^P]_i - [\bar{\theta}_n]_i \right)^2 for i \notin I.
14:
            end for
15:
            Update non-IVPs: [\theta_m]_i = [\theta_{m-1}]_i + [V]_i \sum_{n=1}^N ([\bar{\theta}_n]_i - [\theta_{m-1}]_i) for i \notin I.
16:
            Update IVPs: [\theta_m]_i = \frac{1}{J} \sum_{i=1}^J [\Theta_{L,i}^F]_i for i \in I.
17:
18: end for
```

Under pypomp, the IF2 is implemented under mif(). Arguments sigmas and sigmas\_init control the random walk standard deviation after and during time  $t_0$  respectively. Another argument a, the cooling rate, reduces the perturbation magnitudes across iterations. The purturbations controlled by sigmas, sigmas\_init, and a enable the POMP model to explore the parameter space and progressively converge to the optimal parameter estimates. Specifically, each call also updates and stores a child key into LG\_obj.fresh\_key, so subfequent method calls can safely omit an explicit assignment on key argument.

Different from previous methods for evaluating the likelihoods, Pomp methods such as mif() can update attributes directly other than resuts\_history. The LG\_obj.mif() replaces LG\_obj.theta with the parameter estimate from the end of the last iteration. The traces of log-likelihoods and parameter estimates are stored in LG\_obj.results\_history as a xr.DataArray object. The xr.DataArray has dimentions (replicate, iteration, variable), where the variable dimension includes the log-likelihood estimations and all POMP model parameters.

key, subkey = jax.random.split(key)

**Ensure:** Monte Carlo maximum likelihood estimate  $\theta_M$ .

```
\#LG_{obj_4} = pp.Pomp(
     rinit=rinit,
 #
     rproc=rproc,
 #
     dmeas=dmeas,
 #
     rmeas=rmeas,
 #
     ys=ys,
 #
     theta=theta_list,
 #
     covars=None,
#)
#LG_obj_4.mif(
  #sigmas=0.02,
  #sigmas_init=0.1,
  #J=1000,
  \#M=100,
  #a=0.5, # cooling rate
  #key=subkey
#)
#LG_obj_4.pfilter(J=1000, reps=36)
#LG_obj_4.mif(
    #sigmas=0.005,
    #sigmas_init=0.0025,
    #J=1000,
    #M=100,
    #a=0.5,
#)
#LG_obj_4.pfilter(J=1000, reps=36)
```

The pypomp implementation of IF algorithm in mif() inherits the structure of R's pomp package while adding JAX compilation, providing a substantial performance improvements over R implementation.

#### 4.4. Iterated Filtering with Automatic Differentiation

Iterated filtering with automatic differentiation (IFAD) (Tan et al. 2024) is a hybrid method that combines the complementary advanrages of iterated filtering (IF2) and a new proposed MOP- $\alpha$  gradient method. Here by using the differentiable particle filter, the MOP- $\alpha$  gradient method can obtain gradient and Hessian estimates of the log-likelihood of a POMP model. These estimates are involved in an optimization method to perform a gradient-based search in the parameter space. While IF2 can quickly approach the neighborhood of the global maximum, it struggles to achieve exact local convergence to the global maximum. It can take dozens of iterations to partially close the log-likelihood gap between a point in the neighborhood and the actual global maximum. In contrast, the new MOP- $\alpha$  gradient method can conduct a fast and strong local search, but shows weaker ability at conducting a global exploration (Tan et al. 2024). IFAD addresses this tradeoff by first running IF2 to approach the neighborhoold of the global maximum, and then swithcing to the gradient-based method to further refine the parameter estimates. NOTE:Global Maximum or Local Maximum?

## Algorithm 5 IFAD: Iterated Filtering with Automatic Differentiation

**Require:** Number of particles J, timesteps N, IF2 cooling schedule  $\eta_m$ , MOP- $\alpha$  discounting parameter  $\alpha$ , initial parameter  $\theta_0$ , iteration index m=0.

- 1: Run IF2 until initial "convergence" under cooling schedule  $\eta_m$ , or for a fixed number of iterations, to obtain  $\{\Theta_j, j=1,\ldots,J\}$ .
- 2: Set  $\theta_m := \frac{1}{J} \sum_{j=1}^{J} \Theta_j$ . 3: while procedure not converged do
- Run Algorithm 3 (MOP- $\alpha$  filter) to obtain  $\hat{\ell}(\theta_m)$ .
- Obtain gradient and Hessian: 5:

$$g(\theta_m) = \nabla_{\theta_m} \big( - \hat{\ell}(\theta_m) \big), \quad H(\theta_m) \quad \text{s.t. } \lambda_{\min}(H(\theta_m)) > c.$$

6: Update parameter:

$$\theta_{m+1} := \theta_m - \eta_m \, H(\theta_m)^{-1} g(\theta_m).$$

- Set m := m + 1.
- 8: end while

**Ensure:** Return  $\hat{\theta} := \theta_m$ .

In pypomp, the MOP- $\alpha$  gradent local search is implemented under train(), which utilizes JAX's automatic differentiation features with MOP- $\alpha$  to compute the gradient and Hessian of the log-likelihood. While IFAD is not implemented as an independent function, users can always run mif() and train() to execute the two phases sequentially. For example:

```
key, subkey = jax.random.split(key)
\#LG_{obj_5} = pp.Pomp(
     rinit=rinit,
     rproc=rproc,
     dmeas=dmeas,
     rmeas=rmeas,
     ys=ys,
     theta=theta_list,
#
     covars=None,
#)
#LG_obj_5.mif(
    sigmas=0.02,
    sigmas_init=0.1,
# J=1000,
    M=40.
    a=0.5,
    key=subkey
#LG_obj_5.train(
     J=1000,
#
     itns=40,
     eta=0.0025, # learning rate
```

```
# optimizer="Newton"
#)
```

Here, mif() runs IF2 for 40 iterations, followed by train() runs the Newton's method with MOP- $\alpha$  for 40 iterations.

The implemented function train() serves as a gradient-based optimizer for exploring the parameter space. Each iteration follows the standard optimization working flow: starting from the current parameter estimate, it computes gradient (and Hessian) of the log-likelihood via jax.grad() and jax.Hessian(), then utilizes the choosen internal optimizer to update the direction, obtains the learning rate, and finally updates the parameters as well as evaluates the new likelihood under the updated parameters. train() supports several common gradient-based optimizers for exploring the parameter space, including Newton's method, weighted Newton's method, Broyden-Fletcher-Goldfarb-Shanno (BFGS) algorithm, and gradient descent method. In addition, A linear search method (Wills and Schön 2021) is available under train for determining the learning rate (step size) under stochastic Quasi-Newton methods.

In train(), there is a key argument n\_monitors, controlling the number of log-likelihood values are evaluated during each training step. When setting n\_monitors=0, the algorithm totally skip the evaluating process and conduct the gradient-only optimization. It is the fastest choice, but doesn't support the linear search method for learning rate determination, as it utilizes the estimated log-likelihood as the current objective function value. n\_monitors=1 is a lightest monitoring form, where users can get a one log-likelihood evaluation per training step. Larger n\_monitors reduces the log-likelihood estimate variance in particle filtering using Monte Carlo average at additional computational cost. There is another useful Boolean flag, scale, which controls whether the search direction is normalized to unit length. A normalized search direction can provide a more conservative update when the gradients or Hessians are unstable.

Similarly to mif(), train() updates the result\_history with the trace of log-likelihoods and parameter estimates as xr.DataArray object, and replaces LG\_obj.theta with the final parameter estimates from the last iteration for each replicate, organized as distionaries keyed by parameter names. The xr.DataArray has dimentions (replicate, iteration, variable). The variable dimension includes the log-likelihood estimations and all POMP model parameters.

The core highlight of IFAD in algorithmic implementation lies in its internal automatic differentiation and just-in-time compilation powered by JAX. Via JAX, gradients and Hessian matrices for the MOP- $\alpha$  log-likelihood function can be directly obtained without manual derivation, ensuring correctness, efficiency and operability. This also enables flexible use of advanced gradient-based optimizers in each training step, such as Newton's method, quasi-Newton methods, BFGS, gradient descent and stochastic line search. Additionally, JAX's GPU compilation capabilities enable these optimization processes to execute efficiently, making IFAD not only theoretically feasible but also capable of effectively handling large complex POMP models in practical algorithmic implementations.

Data Analysis with pypomp / Tutorials and data analysis examples

This section demonstrates: - Log-likelihood profiling - GPU benchmarking - Conditional log-likelihood residuals

Discussion

#### References

- Abkemeier A, Chen J, Ionides E, Wheeler J, Tan K (2024). "Pypomp."
- Arulampalam M, Maskell S, Gordon N, Clapp T (2002). "A Tutorial on Particle Filters for Online Nonlinear/Non-Gaussian Bayesian Tracking." *IEEE Transactions on Signal Processing*, **50**(2), 174–188. ISSN 1053587X. doi:10.1109/78.978374.
- Asfaw K, Park J, King AA, Ionides EL (2024). "spatPomp: An R Package for Spatiotemporal Partially Observed Markov Process Models." *Journal of Open Source Software*, **9**(104), 7008. ISSN 2475-9066. doi:10.21105/joss.07008.
- Auger-Méthé M, Newman K, Cole D, Empacher F, Gryba R, King AA, Leos-Barajas V, Mills Flemming J, Nielsen A, Petris G, Thomas L (2021). "A guide to state-space modeling of ecological time series." *Ecological Monographs*, **91**(4), e01470. doi:10.1002/ecm.1470. URL https://doi.org/10.1002/ecm.1470.
- Blackwood JC, Streicker DG, Altizer S, Rohani P (2013). "Resolving the roles of immunity, pathogenesis, and immigration for rabies persistence in vampire bats." *Proceedings of the National Academy of Sciences of the United States of America*, **110**(51), 20837–20842. doi:10.1073/pnas.1308817110. URL https://doi.org/10.1073/pnas.1308817110.
- Bradbury J, Frostig R, Hawkins P, Johnson MJ, Leary C, Dougal Maclaurin, Necula G, Paszke A, VanderPlas a, Skye Wanderman-Milne, Zhang Q (2018). "JAX: Composable Transformations of Python+NumPy Programs."
- Bretó C (2014). "On idiosyncratic stochasticity of financial leverage effects." Statistics & Probability Letters, 91, 20–26. doi:10.1016/j.spl.2014.04.003. URL https://doi.org/10.1016/j.spl.2014.04.003.
- Bretó C, Wheeler J, King AA, Ionides EL (2025). "panelPomp: Analysis of Panel Data via Partially Observed Markov Processes in R." doi:10.48550/arXiv.2410.07934. 2410.07934.
- Fox SJ, Lachmann M, Tec M, Pasco R, Woody S, Du Z, Wang X, Ingle TA, Javan E, Dahan M, Gaither K, Escott ME, Adler SI, Johnston SC, Scott JG, Meyers LA (2022). "Real-time pandemic surveillance using hospital admissions and mobility data." *Proceedings of the National Academy of Sciences*, **119**(7), e2111870119. doi:10.1073/pnas.2111870119. URL https://doi.org/10.1073/pnas.2111870119.
- He D, Ionides EL, King AA (2010). "Plug-and-Play Inference for Disease Dynamics: Measles in Large and Small Populations as a Case Study." *Journal of The Royal Society Interface*, 7(43), 271–283. ISSN 1742-5689, 1742-5662. doi:10.1098/rsif.2009.0151.
- Ionides EL, Breto C, King AA (2006). "Inference for Nonlinear Dynamical Systems." Proceedings of the National Academy of Sciences, 103(49), 18438–18443. ISSN 0027-8424, 1091-6490. doi:10.1073/pnas.0603181103.

- Ionides EL, Breto C, Park J, Smith RA, King AA (2017). "Monte Carlo Profile Confidence Intervals for Dynamic Systems." *Journal of The Royal Society Interface*, **14**(132), 20170126. ISSN 1742-5669, 1742-5662. doi:10.1098/rsif.2017.0126.
- Ionides EL, Nguyen D, Atchadé Y, Stoev S, King AA (2015). "Inference for Dynamic and Latent Variable Models via Iterated, Perturbed Bayes Maps." *Proceedings of the National Academy of Sciences*, **112**(3), 719–724. ISSN 0027-8424, 1091-6490. doi:10.1073/pnas. 1410597112.
- King AA, Ionides EL, Pascual M, Bouma MJ (2008). "Inapparent Infections and Cholera Dynamics." *Nature*, **454**(7206), 877–880. ISSN 1476-4687. doi:10.1038/nature07084.
- King AA, Nguyen D, Ionides EL (2016). "Statistical Inference for Partially Observed Markov Processes via the R Package **Pomp**." Journal of Statistical Software, **69**(12). ISSN 1548-7660. doi:10.18637/jss.v069.i12.
- Korevaar H, Metcalf CJ, Grenfell BT (2020). "Structure, Space and Size: Competing Drivers of Variation in Urban and Rural Measles Transmission." *Journal of The Royal Society Interface*, 17(168), 20200010. doi:10.1098/rsif.2020.0010.
- Marino JAJ, Peacor SD, Bunnell DB, Vanderploeg HA, Pothoven SA, Elgin AK, Bence JR, Jiao J, Ionides EL (2019). "Evaluating consumptive and nonconsumptive predator effects on prey density using field time-series data." *Ecology*, **100**(3), e02583. doi:10.1002/ecy.2583. URL https://doi.org/10.1002/ecy.2583.
- Mietchen MS, Clancey E, McMichael C, Lofgren ET (2024). "Estimating SARS-CoV-2 transmission parameters between coinciding outbreaks in a university population and the surrounding community." doi:10.1101/2024.01.10.24301116. URL https://doi.org/10.1101/2024.01.10.24301116.
- Sun W (2024). "Model Based Inference of Stochastic Volatility via Iterated Filtering."
- Tan K, Hooker G, Ionides EL (2024). "Accelerated Inference for Partially Observed Markov Processes Using Automatic Differentiation." doi:10.48550/arXiv.2407.03085. 2407.03085.
- Wen L, Yin Y, Li Q, Peng Z, He D (2024). "Modeling the co-circulation of influenza and COVID-19 in Hong Kong, China." *Advances in Continuous and Discrete Models*, **2024**(1), 1–9. doi:10.1186/s13662-024-03830-7. Article 32, URL https://doi.org/10.1186/s13662-024-03830-7.
- Wills AG, Schön TB (2021). "Stochastic quasi-Newton with line-search regularisation." Automatica, 127, 109503. ISSN 0005-1098. doi:10.1016/j.automatica.2021.109503.

#### Affiliation:

Aaron J. Abkemeier\*
Department of Statistics
E-mail: aaronabk@umich.edu

E man. dar ondensami

 $\mathrm{Jun}\ \mathrm{Chen}^*$ 

Department of Statistics E-mail: chenjunc@umich.edu

Kevin Tan

Department of Statistics

E-mail: kevtan@wharton.upenn.edu

Jesse Wheeler

Department of Mathematics E-mail: jessewheeler@isu.edu

Bo Yang

Department of Marketing E-mail: ybb@umich.edu

Kunyang He

Department of Statistics

E-mail: kunyanghe@umich.edu

Jonathan Terhorst Department of Statistics E-mail: jonth@umich.edu

Aaron A. King

Edward L. Ionides
Department of Statistics
E-mail: ionides@umich.edu

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<sup>\*</sup>These authors contributed equally