pypomp: Inference for partially observed Markov process models in Python with JAX DRAFT IN PROGRESS

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Introduction

[Topic] Partially Observable Markov Process (POMP) models, also known as state-space models or hidden Markov models, provide a flexible and mechanistic framework for modeling time-series dynamic systems, particularly suited for scenarios where latent states are only partially observable. Characterized by transition densities and measurement densities of Markov processes, this framework bridges complex underlying dynamics with limited information in real-world data. Consequently, POMP models find extensive application in epidemiology (Mietchen et al. 2024; Fox et al. 2022; Wen et al. 2024), ecology (Auger-Méthé et al. 2021; Marino et al. 2019; Blackwood et al. 2013), finance (Bretó 2014), and other domains.

[Existing package discussion] The rich POMP package ecosystem built in R has provided a solid, standardized, and extensible framework for modeling time series data using nonlinear, stochastic partially observed mechanism dynamic models. The R pomp package has become a well-established tool for fitting POMP models using a general and abstract representation, that supports multiple inference techniques. Its extension packages panelPomp, spatpomp, and phyloPomp further enhance its capabilities for panel, spatio-temporal and phylodynamic data analysis respectively

[computational challenges + potential limitations] While conceptually powerful, statistical inference for POMP models using the above R packages poses substantial computational challenges. From a methodological perspective, likelihood-based inference for POMP models typically relies on perturbations within iterated filtering (adding ref related to iterative filtering) algorithms. While many of them are stable and effective in locating a neighborhood containing the likelihood maximum, they exhibit numerical inefficiency for obtaining a precise identification of the maximum value. Particularly, when the latent states are high-dimensional or

when repeated model evaluations are required, the fitting process could be computationally prohibitive by the constraints.

On the other hand, these POMP models have demonstrated strong potential to be sped up considerably. Many of the processes are embarassingly parallel, such as simulating the state process for each of thousands of particles, running the particle filter multiple times for the same parameter set, and running iterated filtering from multiple starting parameter sets, especially when estimating a profile likelihood to construct a confidence interval as per (E. L. Ionides et al. 2017). Graphics Processing Units (GPUs) are well-suited for such operations. However, the existing family of POMP packages (**pomp**, **panelPomp**, and **spatPomp** Asfaw et al. (2024)) only runs on CPUs.

[introducing AD/GPU/JAX + pypomp] As the demand grows for scalable and parallelizable inference algorithms, there is a increasing need for a accelerated POMP modeling framework. Automatic differentiation (AD) is a technique that enables efficient and accurate computation of numerical differentiation by systematically applying the chain rule to fundamental operations within computer programs. While several general-purpose AD libraries are available, we directly integrate AD technique into the inference of generic POMP models, particularly the particle filter. This leads to a novel class of algorithms, which are termed automatic differentiation particle filters (ADPF) for POMP modls. The ADPF and differentiable iterated filtering interfaces enable the gradient-based optimization, effectively resolving the previous numerous inefficiencies in the traditional algorithms. Our approach maintains the plug-and-play property (E. L. Ionides, Breto, and King 2006), allowing users to specify dynamic models solely through simulators that generate latent state trajectories between arbitrary time points. Furthermore, these methods are implemented in JAX(Bradbury et al. 2018), a high-performance numerical computing library that supports hardware acceleration (GPU) and vectorization. JAX's just-in-time (JIT) compilation further accelerates inference. With the combination of ADPF methods, JAX implementation, and GPU hardware supports, instead of merely a port from the R package pomp, pypomp(Abkemeier et al. 2024) establishes a modern platform for POMP modeling.

[structure] The remainder of this paper is organized as follows. Section 2 discusses the Motivation for pypomp design using specific examples. Section 3 demonstrates the mathematical notation for POMP models and their related implementation in pypomp. Section 4 introduces the embedded methodologies. Section 5 presents data analysis workflows and benchmarking results. Section 6 concludes with a discussion of future directions.

NOTE key points for introduction section: add discussion of panelpomp

Motivation for pypomp

NOTE: This section is for some extra detailed numeric cost estimates and dataset descriptions to illustrate motivation based on Aaron's draft. It is a bit redundant

now.

Real-world computational bottleneck

Computational speed is a major bottleneck in the practical application of iterated filtering methods to POMP models. In Korevaar, Metcalf, and Grenfell (2020) 's dataset, fitting and evaluating likelihoods of POMP models for 180 units required 8 days on 36 CPU cores (two 3.0 GHz Intel Xeon Gold 6154 CPUs). Scaling this up to the full dataset of 1422 units would require almost eight times as much effort, equivalent to running 36 cores for two months or 288 cores for 8 days. This is not only time consuming, but also incurs substantial computational costs, highlighting the urgent need for more efficient inference software for large-scale POMP analyses Importantly, this cost only accounts for one round of iterated filtering. In practice, to further refine the likelihood estimates, multiple rounds are required, which would increase the computational burden significantly. This motivates the development of accelerated, scalable tools to make large-scale POMP inference feasible.

Opportunities for speeding up the POMP models

Many of the processes involved in fitting POMP models are embarrassingly parallel. Examples include simulating the state process for each of thousands of particles, running the particle filter repeatedly under the same parameter set, and executing iterated filtering from multiple starting parameter sets. Such parallelism is especially advantageous when estimating a profile likelihood to construct confidence intervals (E. L. Ionides et al. 2017). Harnessing parallel computing resources can therefore dramatically reduce computation time and make large-scale inference feasible.

Graphics Processing Units (GPUs) are well-suited for embarrassingly parallel operations, but the existing family of POMP packages (**pomp**, **panelPomp**, and **spatPomp** Asfaw et al. (2024)) are limited to CPU computation. None provide support for GPU acceleration or automatic differentiation. These two technologies are key to enabling scalable and efficient inference for modern POMP applications.

Our solution: pypomp

To address this computational bottleneck, we are creating pypomp(Abkemeier et al. 2024), a python implementation of the R package pomp. It draws inspiration from pomp, but further implements new methods incoporating automatic differentiation techniques by forking the source code used in Tan (Tan, Hooker, and Ionides 2024), as well as leverages JAX's just-in-time(JIT) compilation and GPU core parallelization (Bradbury et al. 2018), allowing practitioners to run filtering methods significantly faster and cheaper. For example, in an SPX comparison model, we show that, compared to pomp with 36 CPU cores, pypomp can run at least 7 times faster

and can finish the job at 5% of the price using 1 GPU and 1 CPU core (5120 CUDA cores on a NVIDIA Tesla V100 and one core from a 2.4 GHz Intel Xeon Gold 6148 CPU).

In addition, pypomp is gradually including functionality from panelPomp and spatPomp, offering a unified Python interface for entire POMP methodologies across multiple R packages. It also takes advantage of JAX's implementation of automatic differentiation (AD), which can be used in conjunction with the differentiable measurement off-parameter with discount factor α (MOP- α) particle filter to improve local optimization of the likelihood surface (Tan, Hooker, and Ionides 2024).

Summary of key features

Table 1 summarizes the main differences between pypomp and pomp, highlighting the new capabilities of pypomp.

Table 1: Feature comparison between pypomp and pomp in R ecosystem.

Feature	pypomp	pomp
Backend and Acceleration	JAX (GPU/CPU, JIT, vmap, jax.grad, jax.Hessian)	R and C Snippets (CPU only)
Automatic Differentiation and gradient-based inference	Yes (gradient/Hessian via AD supported)	No
Particle Filtering Methods	Yes (PF, MOP- α , IF2, IFAD)	Yes (PF, IF2, pMCMC, etc.)
Plug-and-Play Property	Yes	Yes
Object Design	In-place updates on current objects, stored in the object attribute results_history	Returns new objects

POMP Models in pypomp

This section introduces the structure of POMP models and its implementation in pypomp, including both mathematical setup and the package implementation.

Model setup

A partially observed Markov process (POMP) model has two main components: (i) a latent Markov process that evolves over time and (ii) an observation process that links the latent states. Together, these jointly specify the mechanistic model for the observed time series, providing a framework for modeling dynamic systems where measurements are noisy. Formally,

suppose $t_1 < t_2 < ... < t_N$ be a collection of times at which measurements are available, and let t_0 be some time prior to t_1 at which the model is initialized. Let $\{Y_t\}_{t=1}^N$ denote the observations at time $t_1, ..., t_N$, and $\{X_t\}_{t=1}^N$ denote the postulated latent (unobserved) Markov process at the corresponding time. A POMP model is specified by three building components:

- 1. initial density: $f_{X_0}(x_0;\theta)$ describes the initial distribution of latent state X_0 ;
- 2. transition density: $f_{X_t \mid X_{t-1}}(x_t \mid x_{t-1}; \theta)$ characterizes the latent Markov process evolution:
- 3. measurement density: $f_{Y_t|X_t}(y_t \mid x_t; \theta)$ links the observations and the latent states.

The joint density of $(X_{0:N}, Y_{1:N})$ can be expressed as the product of the initial distribution, the transition densities, and the measurement densities:

The joint density of latent states $(X_{0:N})$ and observation $Y_{1:N}$ can be expressed as the product of initial density, transition density and the measurement density:

$$f_{X_{0:N},Y_{1:N}}(x_{0:N},y_{1:N};\theta) = f_{X_0}(x_0;\theta) \prod_{t=1}^N f_{X_t \mid X_{t-1}}(x_t \mid x_{t-1};\theta) \prod_{t=1}^N f_{Y_t \mid X_t}(y_t \mid x_t;\theta)$$

The marginal likelihood of the observations is $\mathcal{L}(\theta) = f_{Y_{1:N}}(y_{1:N}; \theta) = \int f_{X_{0:N}, Y_{1:N}}(x_{0:N}, y_{1:N}; \theta) \, dx_{0:N}$. In practice, this integral is intractable for most nonlinear or non-Gaussian POMP models, motivating the use of simulation-based inference methods such as particle filtering.

In our software, these model components are specified by user-provided functions (rinit, rproc, dmeas), and the package provides various implementations of likelihood evaluation and parameter inference.

Implementations of POMP models in pypomp

Object-oriented interface

A POMP model in pypomp is represented as an object of class Pomp, which encapsulates the model components: the initial state distribution, process model, and measurement model. This object-oriented interface allows users to specify by passing components to the constructor, including observations, model parameters, model mechanics such as simulators and the measurement density, covariates, and times. After the components are passed into the constructor, the constructor automatically generates additional internal elements, such as extended observations and covariates required for interpolation

Table 2 summarizes the main arguments to the Pomp constructor and their correspondence to mathematical objects

Table 2: Main arguments to the Pomp class and related constructor objects.

			Description / Mathematical
Constructor	Argument	Type	representation
Pomp	rinit	RInit	simulate initial states $X_0 \sim f_{X_0}(x_0; \theta)$
	rproc	RProc	simulate state transitions
			$X_n \sim f_{X_n \mid X_{n-1}}(x_n \mid x_{n-1}; \theta)$
	rmeas	RMeas	simulate observations
			$Y_n \sim f_{Y_n \mid X_n}(y_n \mid x_n; \theta)$
	dmeas	DMeas	evaluate measurement density
			$f_{Y_n X_n}(y_n \mid x_n; \theta)$
	ys	pandas.DataFrame	observations $y_{1:N}^*$ with times $t_{1:N}$
	covars	pandas.DataFrame	covariates $z_{1:N}^*$ with times $s_{1:N}$
	theta	list or dict	parameters θ
RInit	t0	float	initial time point t_0 for simulation
\mathbf{RProc}	step_type	str	method of process evolution:
			"fixedstep" or "euler"
	nstep	int	number of steps if
			step_type="fixedstep"
	dt	float	time step if step_type="euler"
	accumvars	tuple	indices of state variables to be
			accumulated
\mathbf{RMeas}	ydim	int	observation dimension $\dim(Y)$

We demonstrate here how to create a Pomp object. Specifically, we show how to create the linear Gaussian model included in the package as LG(). We begin by importing necessary packages and defining helper functions for handling the parameters. Because pypomp will run our defined model components within JAX JIT-compiled code, it is necessary to write the components to be JAX-compliant. Naturally, the JAX package has many useful functions for this purpose. We also generate a pseudorandom number generation (PRNG) key to be used with JAX random number generators. All stochastic simulations in pypomp are controlled via JAX PRNG keys, ensuring full reproducibility when using the same seed.

```
import pypomp as pp
import pandas as pd
import jax
import jax.numpy as jnp
from functools import partial

def get_thetas(theta):
    theta = jnp.asarray(theta)
```

```
A = theta[0:4].reshape((2, 2))
C = theta[4:8].reshape((2, 2))
Q = theta[8:12].reshape((2, 2))
R = theta[12:16].reshape((2, 2))
return A, C, Q, R

def transform_thetas(A, C, Q, R):
    return jnp.concatenate([A.ravel(), C.ravel(), Q.ravel(), R.ravel()])

# create PRNG key correctly
key = jax.random.PRNGKey(1)
```

Model Components

We refer to model components describing initialization, transfer, or measurement processes as model mechanisms, including rinit, rproc, dmeas, and rmeas. Users must define these processes as Python functions. Specifically, we require users to provide function code to the object constructor, which verifies that all necessary function arguments are included and in the correct order. This requirement stems from pypomp's internal mechanism: it vectorizes component functions using jax.vmap() to efficiently run thousands of particles. Since jax.vmap() maps functions to input arrays by position rather than keyword, users must strictly adhere to parameter order. While all expected parameters must be included, the function does not need to utilize all of them.

Illustrated in Table 2,pypomp also includes object constructors for components describing the model mechanics: RInit, RProc, DMeas, and RMeas. Some constructors also require additional arguments, such as t0 for RInit. Notably, RProc takes step_type, dt, and nstep arguments. step_type determines how RProc should be run at intermediate steps between two observation times. If we want to model the state process as evolving in continuous time, setting step_type="euler"uses an Euler approximation, running rproc at intermediate steps based on the time step size, dt. The number of steps taken is given by the number of times dt divides the difference between two observation times, rounded up, and is consequently dynamic. Otherwise, if we instead want a fixed number of steps for each observation time interval, we can use step_type="fixedstep", in which case rproc will run at nstep intermediate steps equally spaced between two observation times, starting from the first observation time. Consequently, setting step_type="fixedstep" and nstep=1 only runs rproc at the observation times. Here is an example of defining the object constructors for components under the linear gaussian model. In practice, at least one of dmeas or rmeas must be provided, while the construction of RInit and RProc are always required.

```
import pypomp as pp
@partial(pp.RInit, t0=0.0)
def rinit(theta_, key, covars=None, t0=None):
    A, C, Q, R = get_thetas(theta_)
    return jax.random.multivariate_normal(key=key, mean=jnp.array([0.0, 0.0]), cov=Q)
@partial(pp.RProc, step_type="fixedstep", nstep=1)
def rproc(X_, theta_, key, covars=None, t=None, dt=None):
    A, C, Q, R = get_thetas(theta_)
   return jax.random.multivariate_normal(key=key, mean=A @ X_, cov=Q)
@pp.DMeas
def dmeas(Y_, X_, theta_, covars=None, t=None):
    A, C, Q, R = get_thetas(theta_)
    # return logpdf of Y given X (mean = C @ X_, cov = R)
    return jax.scipy.stats.multivariate_normal.logpdf(Y_, mean=C @ X_, cov=R)
Opartial(pp.RMeas, ydim=2)
def rmeas(X_, theta_, key, covars=None, t=None):
    A, C, Q, R = get_thetas(theta_)
   return jax.random.multivariate_normal(key=key, mean=C @ X_, cov=R)
```

Parameters

The Pomp constructor also requires model parameters. These can be provided either as a dictionary or as a list of dictionaries. Each item in a dictionary should include the parameter name as the key and the parameter value as the dictionary value. If the parameter sets are provided as a list of dictionaries, methods such as pfilter() run on each set of parameters. Here, we use Pomp.sample_params() to sample sets of parameters from uniform distributions with bounds passed as a dictionary of length-2 tuples. Pomp.sample_params() returns a ready-to-use list of dictionaries with the sampled parameters. Internally, parameters, even are multi-dimensional, are stored as flat dictionaries to facilitate JAX transformations and compilation.

```
theta = {
    "A11": jnp.cos(0.2), "A12": -jnp.sin(0.2),
    "A21": jnp.sin(0.2), "A22": jnp.cos(0.2),
    "C11": 1.0, "C12": 0.0, "C21": 0.0, "C22": 1.0,
    "Q11": 0.01, "Q12": 1e-6, "Q21": 1e-6, "Q22": 0.01,
    "R11": 0.1, "R12": 0.01, "R21": 0.01, "R22": 0.1,
```

```
param_bounds = {k: (v * 0.9, v * 1.1) for k, v in theta.items()}
n = 5
key = jax.random.PRNGKey(1)
key, subkey = jax.random.split(key)
theta_list = pp.Pomp.sample_params(param_bounds, n, subkey)
```

Covariates

Scientifically, POMP models often involve external time-varying inputs, referred to as covariates, which can influence either the latent process or the measurement model. Examples include seasonality, interventions, or environmental drivers in ecological applications. In pypomp, covariates are supplied as a pandas.DataFrame indexed by time. The time at which the covariates were observed should be specified in the ctime argument. Importantly, the covariate time points may differ from the observation times, necessitating interpolation. Given the observation times, covariate times, and the step type specified in RProc, the model automatically aligns and interpolates observations and covariates to ensure consistency with the simulation of the latent and observation processes. The linear gaussian model doesn't involve any covariates, and an example using covariates is given in the Data Analysis Section.

POMP Object Construction

We do not have real data in this LG example, so we generate our own. To make this example cleaner, we here use the function LG() to construct the completed linear Gaussian model object and then generate the data using simulate(). Observation times are provided to the Pomp constructor via the pandas.DataFrame row index. If covariates were provided, the times at which the covariates were observed would also be provided by the pandas.DataFrame row index.

```
import jax, jax.numpy as jnp
import pandas as pd
import pypomp as pp

T = 100
# ensure `key` exists; if not, uncomment the next line
# key = jax.random.PRNGKey(1)

key, subkey = jax.random.split(key)
sims = pp.LG(T=T).simulate(key=subkey)
```

```
ys = pd.DataFrame(
    sims[0]["Y_sims"].squeeze(),
    index=range(1, T + 1),
    columns=["Y1", "Y2"],
)

LG_obj = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    ys=ys,
    theta=theta_list,
    covars=None,
)

print("LG_obj created; ys.shape =", ys.shape)
```

```
LG_obj created; ys.shape = (100, 2)
```

Each argument to Pomp is accessible from the object as an attribute.

```
print(LG_obj.rinit) # access POMP model components
print(LG_obj.rproc)
print(LG_obj.dmeas)
print(LG_obj.rmeas)
print(LG_obj.theta) # access parameters
print(LG_obj.ys.head()) # access observations
```

4 0.955786 0.462541

Premade models:

Beyond the linear gaussian model, pypomp includes several ready-to-use model constructors that serve both as examples and as tested templates for custom model development:

- 1. LG() a simple linear-Gaussian model with 2-dimensional latent and observed states; useful to validate API usage and diagnostics.
- 2. spx() the S&P500 log-return model from Sun et al. (Sun 2024).
- 3. dacca() the cholera transmission model from King et al. (King et al. 2008).
- 4. UKMeasles.Pomp() the measles district model from He et al. (He, Ionides, and King 2010), wired to the Korevaar et al. dataset (Korevaar, Metcalf, and Grenfell 2020). Panel and spatial variants (PanelPOMP/SpatPOMP style) are planned.

These examples show correct component wiring (rinit, rproc, dmeas, rmeas), recommended step_type/dt usage, and typical diagnostics. If a user model errors or runs slowly, compare its components to the matching premade model to find mistakes and performance opportunities. Meanwhile, these premade models can also replicate well-know case studies in the R pomp ecosystem, allowing direct comparison and validation.

JAX Numerical Backend and Interface Design

A key design choice pypomp is it relys heavily on the JAX numerical backend. Unlike the R package pomp, where users typically provide POMP model components in C Snippets for acceleration, pypomp requires model components to be written as JAX-compatible Python functions. These functions are then compiled and vectorized by JAX tools such as jit and vmap. This design leads to several important interface features:

- Strict argument requirements for compilation and vectorization: JAX's jit compiler transforms the user-supplied component functions (rinit, rproc, dmeas, rmeas) into efficient machine code, while vmap efficiently run them over thousands of particles via vectorization of arguments. To ensure the compatibility with JAX's compilation and vectorization system, each component function must follow the expected input types and order, otherwise compilation would fail.
- PRNG random key policy: To ensure the reproducibility of randomness in POMP models under pypomp, the public API accept an optional jax.random.PRNGkay, which is explicitly passed through constructors and methods. Keys are internally split when it is needed. Unlike the R setting, where randomness can be controlled globally or by seed chunks, in JAX, random keys only be explicitly passed through functions
- Consistent shapes and sizes handling: model parameters, even multidimensional, are stored as flattened dictionaries. Consequently, JAX can uniformly process parameters, thereby maintaining consistency in particle propagation.

Later section will demonstrate how the JAX-based design supports further inference methods.

[Question: more introductions on JAX?]

Panel POMP class

POMP Methods in pypomp

In this section, we describe the core inference methods currently implemented in pypomp, including:

- Paricle Filter (Sequential Monte Carlo, written in pfilter()): A standard sequential Monte Carlo algorithm for likelihood evaluation and state estimation, forming the basis for most inference methods in POMP models.
- Measurement-off-policy Particle Filter (MOP(α), written in mop()): A recently proposed SMC method (Tan, Hooker, and Ionides 2024) that evaluates the likelihood at one parameter value while obtaining resampling decisions from another, adjusting via discounted off-parameter measurement weights.
- Iterated Filtering (IF2, mif()): A classical IF2 algorithm (Edward L. Ionides et al. 2015) for likelihood-based parameter inference that maximizes the likelihood via particle filtering.
- Iterated Filtering with Automatic Differentiation (train()): A recently proposed AD-based algorithm (Tan, Hooker, and Ionides 2024) that incorporates $MOP(\alpha)$, the differentiable particle filter, to enable efficient gradient-based parameter inference for maximum likelihood estimation.

A key feature of the above POMP inference methods lies in the **plug-and-play property** (E. L. Ionides, Breto, and King 2006), meaning that inference algorithms can be implemented without requiring explicit evaluation of the transition density of the latent process. Instead, it suffices for the user to provide a simulator of the latent process (**rproc**), initial state distribution (**rinit**), and observation measurement model (**dmeas**, **rmeas**). This property enables POMP methods to be widely applied to complex mechanistic models where transition densities are intractable.

In pypomp, the plug-and-play design is fully preserved: users only need to provide component functions compliant with JAX requirements, which can be directly plugged in inference methods likepfilter(), mop(), mif(), and train(). The package combines the generality of plug-and-play modeling with the efficiency of JAX compilation and vectorization.

Unlike the R family of POMP packages, some Pomp class methods including pfilter(), mif() and train() yield results by modifying the object in place instead of returning new objects. All of results are stored a list under LG_obj.results_history, which is an attibute under

Pomp class object LG_obj. Each element in the list correponds to one method call. Each element includes results such as the log-likelihood and parameter estimates when applicable as well as the inputs used for the function call, so it is easy to keep track of how the results were calculated. If multiple parameter sets are supplied in a list as an argument, the method evaluates at each set and the results for each are stored.

```
[{'method': 'pfilter', 'logLiks': <xarray.DataArray (theta: 5, replicate: 5)> Size: 100B
Array([[-78.89917 , -76.7121 , -76.20436 , -77.44517 , -79.325195],
       [-74.994446, -70.438416, -70.757324, -70.6057, -72.60899],
       [-78.496506, -79.72346 , -78.77383 , -79.19662 , -76.05818 ],
       [-73.46929, -73.42637, -72.895485, -71.27409, -70.808205],
       [-75.0522 , -75.3827 , -73.30542 , -71.4641 , -75.1468 ]],
                                                                          dtype=float32, we
Dimensions without coordinates: theta, replicate, 'theta': [{'A11': 0.9211453199386597, 'A12
array([[[
                          9.21145320e-01, -1.78802371e-01,
         2.14950144e-01, 8.89322698e-01, 1.02344239e+00,
         0.0000000e+00, 0.0000000e+00, 9.79293942e-01,
         1.04345558e-02, 9.33975798e-07, 9.76017759e-07,
         1.05938269e-02, 9.44515765e-02, 9.41129215e-03,
         9.19735525e-03, 9.43083838e-02],
        [-0.00000000e+00, 1.07045341e+00, -2.18700305e-01,
         2.06873283e-01, 9.45706606e-01, 8.82167339e-01,
        -3.03104311e-01, -1.46514282e-01, 5.65519333e-01,
         1.62466958e-01, 1.77885965e-01, -6.56164214e-02,
         3.40645671e-01, 1.02493912e-01, 3.30706313e-02,
        -6.84480742e-02, 1.20453514e-01],
        [-1.24636581e+02, 8.16644430e-01, -5.28744869e-02,
         5.82232058e-01, 6.72007203e-01, 9.57372367e-01,
        -3.92328411e-01, -2.65173197e-01, 4.19366837e-01,
         2.80792415e-01, 2.61916459e-01, 6.33998886e-02,
         3.07527870e-01, 7.97012299e-02, 6.75280914e-02,
```

```
1.40667498e-01]],
         -6.22278303e-02,
       [[
                           9.74913478e-01, -1.78802371e-01,
                    nan,
. . .
         -2.03950927e-02,
                           1.04405761e-01]],
       9.70987082e-01, -1.78802371e-01,
                    nan,
          2.18500718e-01,
                           9.31446433e-01, 1.07719481e+00,
          0.00000000e+00,
                           0.0000000e+00, 9.35545087e-01,
          1.07558817e-02,
                          1.06478592e-06, 9.44516046e-07,
          1.02092260e-02,
                           1.01782054e-01, 9.41056665e-03,
          9.20325704e-03,
                           9.58210155e-02],
                          8.51299584e-01, -2.22326946e-02,
        [-0.0000000e+00,
          2.04024658e-01,
                           1.01018143e+00, 7.58982956e-01,
          1.84635725e-02,
                           2.20715329e-01,
                                           9.56478238e-01,
          2.77179211e-01, -1.28381521e-01, -7.14511871e-02,
          1.35017648e-01, 6.87062517e-02, 1.59603179e-01,
         -1.26104221e-01, 1.22075811e-01],
        [-1.17136009e+02, 7.31945157e-01, -1.86973646e-01,
          1.58347413e-01, 9.49631929e-01, 5.89739382e-01,
          2.68075258e-01, 2.40069270e-01, 1.08920443e+00,
          2.85785407e-01, -8.81665424e-02, -2.44462746e-03,
          9.77937952e-02, 6.82703108e-02, 8.02902132e-03,
         -2.59932708e-02,
                          1.85035437e-01]])
Coordinates:
               (replicate) int64 40B 0 1 2 3 4
  * replicate
               (iteration) int64 24B 0 1 2
  * iteration
               (variable) <U6 408B 'logLik' 'A11' 'A12' ... 'R12' 'R21' 'R22', 'theta': [{'A
  * variable
```

Particle Filter (pfilter)

NOTE: 1. purpose/role 2. implementation details in pypomp 3. outputs/results 4. remarks/highlights

The particle filter algorithm (sequential Monte Carlo, SMC) (Arulampalam et al. (2002), King, Nguyen, and Ionides (2016)) is the core classical likelihood evaluation tool in POMP model inference. It uses Monte Carlo techniques to sequentially estimate the integrals in the prediction and filtering recursions, and produce an unbiased Monte Carlo estimate of the likelihood of the observed data under a parameter set. The likelihood estimate is essential for other parameter inference algorithms such as iterative filtering. In pypomp, the basic particle filter is implemented under pfilter(). The algorithm is summarized as Algerithm 1,

1 Sequential Monte Carlo (SMC, or particle filter) LG_obj.pfilter(J=J, reps=reps, key=key), where LG_obj is a class Pomp object with definitions for rinit, rproc, dmeas, rmeas, ys, and theta

Require: Simulator for $f_{X_n \mid X_{n-1}}(x_n \mid x_{n-1}; \theta)$; Evaluator for $f_{Y_n \mid X_n}(y_n \mid x_n; \theta)$; Simulator for $f_{X_0}(x_0; \theta)$; Parameter θ ; Data $y_{1:N}^*$; Number of particles J.

- 1: Initialize filter particles: simulate $X_{0,j}^F \sim f_{X_0}(\cdot;\theta)$ for j=1:J.
- 2: **for** n = 1 to N **do**
- Simulate prediction particles: $X_{n,j}^P \sim f_{X_n \mid X_{n-1}}(\cdot \mid X_{n-1,j}^F; \theta)$ for j=1:J. 3:
- Evaluate weights: $w(n,j) = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P^{n-1}}; \theta)$ for j = 1:J. 4:
- Normalize weights: $\tilde{w}(n,j) = \frac{w(n,j)}{\sum_{m=1}^{J} w(n,m)}$. Resample indices $k_{1:J}$ with $\Pr[k_j = m] = \tilde{w}(n,m)$. 5:
- 6:
- Set $X_{n,j}^{\bar{F}} = X_{n,k_{\bar{s}}}^{P}$ for j = 1:J. 7:
- Compute conditional log likelihood: 8:

$$\hat{\ell}_{n|1:n-1} = \log \left(\frac{1}{J} \sum_{m=1}^{J} w(n,m) \right).$$

9: end for

Ensure: Log likelihood estimate $\hat{\ell}(\theta) = \sum_{n=1}^{N} \hat{\ell}_{n|1:n-1}$; filter samples $X_{n,1:J}^{F}$ for n=1:N. 10: Complexity: $\mathcal{O}(J)$

In pypomp, the pfilter() functions is internally run in a compiled JAX routine _pfilter_internal is pfilter internal() but wrapped up into a class method. JIT-compiled with static arguments, in which the replicates and parameter sets are vectorized via jax.vmap. Both techniques enable execution on CPU/GPU and ensure the computational efficiency. In implementation, the weights and likelihood are stored in a log scale, preventing underflow. There are optional arguments (CLL, ESS, filter_mean, prediction_mean) control wheter additional diagnostic arrays are allocated and updated. The algorithm works for flexible and extendable models through the user-specified Pomp class components (rinit, rproc, dmeas, reams), number of particles J as well as other parameters. Each run returns a dict type element updated inside the LG_obj. result_history attribute, containing the log-likelihoods, algorithm parameters used, as well as model diagonostic elements at each time included if their respective boolean flags are set to True. For example, suppose we run

```
LG_{obj_2} = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    ys=ys,
```

where J is the number of particles used and reps is the number of particle filtering replicates to run for each parameter set provided in the Pomp object or as an optional argument to pfilter(). Because LG_obj2.result_history begins as an empty list here when the model is constructed, the results are appended at LG_obj_2.results_history[0] and LG_obj_2.results_history[1] respectively. Both of these two dictionaries contain with the following items:

- method: The method that was run. In this case, pfilter.
- logLiks: An xarray.DataArray of particle filter log-likelihood estimates. There is one list for each parameter set given and each list has reps log-likelihood estimates. dimensions = (theta, replicate)
- theta: The list of parameter sets used.
- J: The number of particles used.
- thresh: Threshold value for resampling used, which is 0 by default.

Meanwhile, LG_obj_2.results_history[1] also contains the following items that are not contained in LG_obj_2.results_history[0]:

- CLL: An xarray.DataArray containing the conditional loglikelihood over observed time steps (theta, replicate, time)
- ESS: An xarray.DataArray containing the effective sample size over observed time steps (theta, replicate, time)
- filter_mean: An xarray.DataArray containing the filtered mean over observed time steps (theta, replicate, time, state)
- prediction_mean: An xarray.DataArray containing the prediction mean over observed time steps. (theta, replicate, time, state)

A key design is that all runs across multiple parameter sets and replications are stored in one results_history list under the corresponding Pomp object. This avoids object duplication, and increases reproducibility by facilitating pickling Pomp objects obtained from many different search strategies, loading them into a new Python session, and then comparing the results to see how they vary with different algorithmic parameters.

Measurement Off-Parameter (MOP)

The measurement off-parameter- paricle filter (MOP- α) algorithm (Tan, Hooker, and Ionides 2024).

It is a variant of the standard vanilla particle filter aiming to produce differentiable likelihood estimate with respect to the model parameters. More spcifically, it evaluates the likelihood at the target evaluation parameter θ and resamples the particles with the indices generated by the vanilla particle filter at another baseline behavior parameter ϕ . This makes the Monte Carlo resampling process independent of the target evaluation parameter θ , enabling the differentiability of the algorithm, and the further utilization of automatic differentiation and gradient descent. There is a discount factor $\alpha \in [0,1]$, which controls the bias-variance tradeoff of the gradient estimator. The likelihood estimate is essential for other parameter inference algorithms such as iterated filtering with automatic differentiation (IFAD), where it employs stochastic gradient descient to MOP. In pypomp, the MOP is implemented under mop(). The pseudocode is shown in 2.

Algorithm 2 MOP(α), Measurement off-policy sequential Monte Carlo

Require: Simulator for $f_{X_n|X_{n-1}}(x_n\mid x_{n-1};\theta)$; Evaluator for $f_{Y_n|X_n}(y_n\mid x_n;\theta)$; Simulator for $f_{X_0}(x_0;\theta)$; Evaluation Parameter θ ; Behavior Parameter ϕ ; Data $y_{1:N}^*$; Number of particles

- 1: Pass: Set $\theta = \phi$, yielding $X_{n,j}^{P,\phi}, X_{n,j}^{F,\phi}, g_{n,j}^{\phi}$. 2: Initialize filter particles: simulate $X_{0,j}^{F,\theta} \sim f_{X_0}(\cdot;\theta)$ and $w_{0,j}^{F,\theta} = 1$ for j = 1:J.
- 3: **for** n = 1 to N **do**
- 4:
- Simulate prediction particles: $X_{n,j}^{P,\theta} \sim f_{X_n|X_{n-1}}(\cdot \mid X_{n-1,j}^{F,\theta}; \theta)$ for j=1:J. Accumulate discounted Prediction weight: $w_{n,j}^{P,\theta} = \left(w_{n-1,j}^{F,\theta}\right)^{\alpha}$ for j=1:J. 5:
- Evaluate measurement density: $g_{n,j}^{\theta} = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P,\theta}; \theta)$ for j = 1:J. Conditional likelihood under θ in version B: 6:
- 7:

$$L_{n}^{B,\theta,\alpha} = \frac{\sum_{j=1}^{J} g_{n,j}^{\theta} w_{n,j}^{P,\theta}}{\sum_{j=1}^{J} w_{n,j}^{P,\theta}}.$$

Conditional likelihood under ϕ : 8:

$$L_n^{\phi} = \frac{1}{J} \sum_{m=1}^J g_{n,m}^{\phi}.$$

- Normalize weights: $\tilde{g}_{n,j}^{\phi} = \frac{g_{n,j}^{\phi}}{I^{\phi}}$ for j = 1:J. 9:
- 10:
- Resample indices $k_{1:J}$ with $\mathbb{P}[k_j = m] = \tilde{g}_{n,m}^{\phi}$. Resample particles: $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$ for j = 1:J. 11:
- Calculate resampled corrected weights: 12:

$$w_{n,j}^{F,\theta} = w_{n,j}^{P,\theta} \times \frac{g_{n,j}^{\theta}}{g_{n,j}^{\phi}}$$
 for $j = 1:J$.

Conditional likelihood under θ in version A: 13:

$$L_n^{A,\theta,\alpha} = L_n^{\phi} \cdot \frac{\sum_{j=1}^J w_{n,j}^{F,\theta}}{\sum_{j=1}^J w_{n,j}^{P,\theta}}.$$

- 14: **end for**
- 15: Likelihood estimate: $\hat{L}(\theta) = \prod_{n=1}^{N} L_n^{A,\theta,\alpha}$ or $\hat{L}(\theta) = \prod_{n=1}^{N} L_n^{B,\theta,\alpha}$.

In pypomp, the mop() function is internally run in a JAX-compiled function _mop_internal() but wrapped up into a class method. Similarly, it also uses static arguments and static array shapes supported by jax.lax.fori loop to enhance efficiency. The weights and likelihood are also stored in a log scale to prevent underflow. Technically, instead of setting two distinct parameters θ and ϕ , the algorithm internally uses the **stop_gradient** function in **JAX** to run one particle filter with $\theta = \phi$, and set the particles at ϕ to be copies of θ to avoid propagating gradients on these particles. This design simplifies the algorithm and still ensures the valid likelihood and gradient estimator, while maintaining differentiability with respect to θ .

In practice, the mop() provides an auto-differentiable alternative to pfilter(), as a lightweight version. Different from pfilter(), mop() in pypomp focuses purely on producing differentiable likelihood estimates, avoiding extra multiple replicates and diagnostics. This could further support the automatic differentiation and stochastic gradient descent algorithms in accelerated iterated filtering algorithms. It only requires the number of particles J, the evaluation parameter theta, a PRNG key, and the cooling factor α . The results are stored in the results_history attribute as a list of jax.Array objects, one for each parameter set, representing the estimated log-likelihood of the observations. For example, we run

```
[Array(-76.55094, dtype=float32, weak_type=True), Array(-71.051605, dtype=float32, weak_type=True), Array(-75.82673, dtype=float32, weak_type=True), Array(-71.13629, dtype=float32, weak_type=True), Array(-72.63167, dtype=float32, weak_type=True)]
```

Generally, the result_history only updates:

• list of jax.Array: The estimated log-likelihood(s) of the observed data given the model parameters. Always a list, even if only one theta is provided.

This design highlights mop() as a lightweight but essential building block for accelarated gradient-based inference methods within pypomp.

Iterated Filtering

Algorithm 3 Iterated Filtering (IF2)

```
Require: Starting parameter \theta_0; simulator for f_{X_0}(x_0;\theta); simulator for f_{X_n\mid X_{n-1}}(x_n\mid x_{n-1};\theta); evaluator for f_{Y_n\mid X_n}(y_n\mid x_n;\theta); data y_{1:N}^*; labels I\subset\{1,\ldots,p\} for IVPs; fixed lag L; number of particles J; number of iterations M; cooling rate a,\ 0< a<1; perturbation scales \sigma_{1:p}; initial scale multiplier C>0.
```

```
1: for m = 1 to M do
               \text{Initialize parameters: } \Theta^P_{0,j,i} \sim \text{Normal}\big([\theta_{m-1}]_i, \; (Ca^m\sigma_i)^2\big) \text{ for } i \in 1:p, \; j \in 1:J.
  2:
              Initialize states: simulate X_{0,j}^F \sim f_{X_0}(\cdot;\Theta_{0,j}^P) for j=1:J.
 3:
              Initialize filter mean: \bar{\theta}_0 = \theta_{m-1}.
  4:
              Define [V]_i = (C^2 + 1)a^{2m}\sigma_i^2
  5:
               for n = 1 to N do
 6:
                     Perturb parameters: \Theta^P_{n,j,i} \sim \text{Normal}\big([\Theta^F_{n-1,j}]_i, \ (a^m\sigma_i)^2\big) for i \notin I, \ j=1:J.
Simulate prediction particles: X^P_{n,j} \sim f_{X_n|X_{n-1}}(\cdot \mid X^F_{n-1,j};\Theta^P_{n,j}) for j=1:J.
Evaluate weights: w(n,j) = f_{Y_n|X_n}(y^*_n \mid X^P_{n,j};\Theta^P_{n,j}) for j=1:J.
  7:
  8:
 9:
                     Normalize weights: \tilde{w}(n,j) = \frac{w(n,j)}{\sum_{u=1}^{J} w(n,u)}.
Resample indices k_{1:J} with \Pr[k_u=j] = \tilde{w}(n,j).
Resample particles: X_{n,j}^F = X_{n,k_j}^P and \Theta_{n,j}^F = \Theta_{n,k_j}^P for j=1:J.
10:
11:
12:
                      Filter mean: [\bar{\theta}_n]_i = \sum_{j=1}^J \tilde{w}(n,j) [\Theta_{n,j}^P]_i for i \notin I.
13:
                      Prediction variance: [V_{n+1}]_i = (a^m \sigma_i)^2 + \sum_{j=1}^J \tilde{w}(n,j) \left( [\Theta_{n,j}^P]_i - [\bar{\theta}_n]_i \right)^2 for i \notin I.
14:
               end for
15:
               Update non-IVPs: [\theta_m]_i = [\underline{\theta}_{m-1}]_i + [V]_i \sum_{n=1}^N ([\bar{\theta}_n]_i - [\theta_{m-1}]_i) for i \notin I.
16:
               Update IVPs: [\theta_m]_i = \frac{1}{J} \sum_{j=1}^J [\Theta_{L,j}^F]_i for i \in I.
17:
18: end for
```

Ensure: Monte Carlo maximum likelihood estimate θ_M .

Iterated Firltering with Automatic Differentiation

Algorithm 4 IFAD: Iterated Filtering with Automatic Differentiation

Require: Number of particles J, timesteps N, IF2 cooling schedule η_m , MOP- α discounting parameter α , initial parameter θ_0 , iteration index m=0.

- 1: Run IF2 until initial "convergence" under cooling schedule η_m , or for a fixed number of iterations, to obtain $\{\Theta_i, j=1,\ldots,J\}$.
- 2: Set $\theta_m:=\frac{1}{J}\sum_{j=1}^J\Theta_j$. 3: while procedure not converged do
- Run Algorithm 2 (MOP- α filter) to obtain $\hat{\ell}(\theta_m)$. 4:
- Obtain gradient and Hessian: 5:

$$g(\theta_m) = \nabla_{\theta_m} \big(- \hat{\ell}(\theta_m) \big), \quad H(\theta_m) \quad \text{s.t. } \lambda_{\min}(H(\theta_m)) > c.$$

Update parameter: 6:

$$\theta_{m+1} := \theta_m - \eta_m \, H(\theta_m)^{-1} g(\theta_m).$$

- Set m := m + 1. 7:
- 8: end while

Ensure: Return $\hat{\theta} := \theta_m$.

Data Analysis with pypomp / Tutorials and data analysis examples

This section demonstrates: - Log-likelihood profiling - GPU benchmarking - Conditional loglikelihood residuals

Discussion

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