pypomp: Inference for partially observed Markov process models in Python with JAX DRAFT IN PROGRESS

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Introduction

[Topic]

Partially Observable Markov Process (POMP) models, also known as state-space models or hidden Markov models, provide a flexible and mechanistic framework for modeling time-series dynamic systems, particularly suited for scenarios where latent states are only partially observable. Characterized by transition densities and measurement densities of Markov processes, this framework bridges complex underlying dynamics with limited information in real-world data. Consequently, POMP models find extensive application in epidemiology (Mietchen et al. 2024; Fox et al. 2022; Wen et al. 2024), ecology (Auger-Méthé et al. 2021; Marino et al. 2019; Blackwood et al. 2013), finance (Bretó 2014), and other domains.

[Existing package discussion]

The POMP package ecosystem in R has provided a solid, standardized, and extensible framework for modeling time series data using nonlinear, stochastic partially observed mechanism dynamic models. The R pomp package has become a well-established tool for fitting POMP models using a general and abstract representation, that supports multiple inference techniques. Its extension packages panelPomp, spatpomp, and phyloPomp further enhance its capabilities for panel, spatio-temporal, and phylodynamic data analysis, respectively.

[computational challenges + potential limitations]

While conceptually powerful, statistical inference for POMP models using the above R packages poses substantial computational challenges. From a methodological perspective, likelihood-based inference for POMP models typically relies on perturbations within iterated filtering (adding ref related to iterative filtering) algorithms. While many of them are stable and

effective in locating a neighborhood containing the likelihood maximum, they exhibit numerical inefficiency for obtaining a precise identification of the maximum value. Particularly, when the latent states are high-dimensional or when repeated model evaluations are required, the fitting process could be computationally prohibitive by the constraints.

As it turns out, POMP methods can be sped up considerably. Many of the processes involved in POMP methods are embarassingly parallel, such as simulating the state process for each of thousands of particles, running the particle filter multiple times for the same parameter set, and running iterated filtering from multiple starting parameter sets, especially when estimating a profile likelihood to construct a confidence interval as per (E. L. Ionides et al. 2017). Graphics Processing Units (GPUs) are well-suited for such operations. However, the existing family of POMP packages (pomp, panelPomp, and spatPomp Asfaw et al. (2024)) only runs on CPUs.

[introducing AD/JIT/GPU hardware + pypomp]

As the demand grows for scalable and parallelizable inference algorithms, there is an increasing need for an accelerated framework for POMP modeling framework. The Python package pypomp (Abkemeier et al. 2024) addresses this by integrating automatic differentiation (AD), just-in-time (JIT) compilation, and GPU acceleration through JAX (Bradbury et al. 2018), a high-performance numerical computing library that supports hardware acceleration and vectorization. AD techniques enable efficient and accurate derivative computation by systematically applying the chain rule to fundamental operations. Recent works have extended AD for gradient estimation using particle filtering, yielding a class of methods termed automatic differentiation particle filters (ADPF). Building on these advancements, pypomp provides AD-enabled gradient estimation for inference for POMP models within a plug-and-play framework (E. L. Ionides, Breto, and King 2006), where users specify dynamic models solely through simulators of latent state trajectories, rather than evaluating the transition density of the latent Markov process. JAX's JIT compilation further accelerates repeated inference and evaluation by converting Python functions into machine code at runtime. Finally, JAX's GPU support enables efficient scaling of inference methods. Together, these features make pypomp more than a port of the R package pomp. Instead, it establishes an independent, modernized platform for fast and flexible POMP modeling in Python.

[structure]

The remainder of this paper is organized as follows. Section 2 discusses the motivation for pypomp design using specific examples. Section 3 demonstrates the mathematical notation for POMP models and their related implementation in pypomp. Section 4 introduces the embedded methodologies. Section 5 presents data analysis workflows and benchmarking results. Section 6 concludes with a discussion of future directions.

[add discussion of panelpomp into introduction]

Motivation for pypomp

[This section is for some extra detailed numeric cost estimates and dataset descriptions to illustrate motivation based on Aaron's draft. It is a bit redundant now.]

Real-world computational bottleneck

Computational speed is a major bottleneck in the practical application of iterated filtering methods to POMP models. In Korevaar, Metcalf, and Grenfell (2020) 's dataset, fitting and evaluating likelihoods of POMP models for 180 units required 8 days on 36 CPU cores (two 3.0 GHz Intel Xeon Gold 6154 CPUs). Scaling this up to the full dataset of 1422 units would require almost eight times as much effort, equivalent to running 36 cores for two months or 288 cores for 8 days. This is not only time consuming, but also incurs substantial computational costs, highlighting the urgent need for more efficient inference software for large-scale POMP analyses Importantly, this cost only accounts for one round of iterated filtering. In practice, to further refine the likelihood estimates, multple rounds are required, which would increase the computational burden significantly. This motivates the development of accelerated, scalable tools to make large-scale POMP inference feasible.

Opportunities for speeding up the POMP models

Many of the processes involved in fitting POMP models are embarrassingly parallel. Examples include simulating the state process for each of thousands of particles, running the particle filter repeatedly under the same parameter set, and executing iterated filtering from multiple starting parameter sets. Such parallelism is especially advantageous when estimating a profile likelihood to construct confidence intervals (E. L. Ionides et al. 2017). Harnessing parallel computing resources can therefore dramatically reduce computation time and make large-scale inference feasible.

Graphics Processing Units (GPUs) are well-suited for embarrassingly parallel operations, but the existing family of POMP packages (**pomp**, **panelPomp**, and **spatPomp** Asfaw et al. (2024)) are limited to CPU computation. None provide support for GPU acceleration or automatic differentiation. These two technologies are key to enabling scalable and efficient inference for modern POMP applications.

Our solution: pypomp

To address this computational bottleneck, we are creating pypomp(Abkemeier et al. 2024), a python implementation of the R package pomp. It draws inspiration from pomp, but further implements new methods incoporating automatic differentiation techniques by forking the source code used in Tan (Tan, Hooker, and Ionides 2024), as well as leverages JAX's just-in-time(JIT)

compilation and GPU core parallelization (Bradbury et al. 2018), allowing practitioners run filtering methods significantly faster and cheaper. For example, in an SPX comparison model, we show that, compared to pomp with 36 CPU cores, pypomp can run at least 7 times faster and can finish the job at 5% of the price using 1 GPU and 1 CPU core (5120 CUDA cores on a NVIDIA Tesla V100 and one core from a 2.4 GHz Intel Xeon Gold 6148 CPU).

In addition, pypomp is gradually including functionality from panelPomp and spatPomp, offering a unified Python interface for entire POMP methodologies across multiple R packages. It also takes advantage of JAX's implementation of automatic differentiation (AD), which can be used in conjunction with the differentiable measurement off-parameter with discount factor α (MOP- α) particle filter to improve local optimization of the likelihood surface (Tan, Hooker, and Ionides 2024).

Summary of key features

Table 1 summarizes the main differences between pypomp and pomp, highlighting the new capabilities of pypomp.

Table 1: Feature comparison	between pypomp and	${ m l}$ pomp ${ m in}$ ${ m l}$	R ecosystem.
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Feature	pypomp	pomp
Backend and Acceleration	JAX (GPU/CPU, JIT,vmap,	R and C Snippets (CPU
	<pre>jax.grad, jax.Hessian)</pre>	only)
Automatic Differentiation	Yes (gradient/Hessian via	No
and gradient-based inference	AD supported)	
Particle Filtering Methods	Yes (PF, MOP- α , IF2,	Yes (PF, IF2, pMCMC, etc.)
	IFAD)	
Plug-and-Play Property	Yes	Yes
Object Design	In-place updates on current	Returns new objects
	objects, stored in the object	-
	attribute results_history	

POMP Models in pypomp

This section introduces the structure of POMP models and its implementation in pypomp, including both mathematical setup and the package implementation.

Model setup

A partially observed Markov process (POMP) model has two main components: (i) a latent Markov process that evolves over time and (ii) an observation process that links the

latent states. Together, these jointly specify the mechanistic model for the observed time series, providing a framework for modeling dynamic systems where measurements are noisy. Formally, suppose we observe the process at discrete time $n=1,\ldots,N$, with initial time 0. Let $Y_{1:N}$ denote the observations and $X_{0:N}$ denote the corresponding latent (unobserved) Markov process at the corresponding time. A POMP model is defined by three building components, each corresponding to user-supplied Python functions in pypomp:

- 1. initial density: $f_{X_0}(x_0; \theta)$, implemented via rinit, generates draws of the initial states X_0 .
- 2. transition density: $f_{X_n|X_{n-1}}(x_n\mid x_{n-1};\theta)$, implemented via rproc, propagates the evolution of latent states.
- 3. measurement density: $f_{Y_n|X_n}(y_n \mid x_n; \theta)$, implemented via dmeas, evaluates or simulates observations conditional on the latent states.

Here, we hold an conditional independence assumption that, given X_n , the observation Y_n is independent of all other variables. Then, the joint density of $(X_{0:N}, Y_{1:N})$ can be expressed as the product of the initial distribution, the transition densities, and the measurement densities:

$$f_{X_{0:N},Y_{1:N}}(x_{0:N},y_{1:N};\theta) = f_{X_0}(x_0;\theta) \prod_{n=1}^N f_{X_n\mid X_{n-1}}(x_n\mid x_{n-1};\theta) \prod_{n=1}^N f_{Y_n\mid X_n}(y_n\mid x_n;\theta)$$

The marginal likelihood of the observations is $\mathcal{L}(\theta) = f_{Y_{1:N}}(y_{1:N};\theta) = \int f_{X_{0:N},Y_{1:N}}(x_{0:N},y_{1:N};\theta)\,dx_{0:N}.$ In practice, this integral is intractable for most nonlinear or non-Gaussian POMP models, motivating the use of simulation-based inference methods such as particle filtering.

Implementations of POMP models in pypomp

Object-oriented interface

A POMP model in pypomp is represented as an object of class Pomp, which encapsulates the model components: the initial state distribution, process model, and measurement model. This object-oriented interface allows users to specify by passing components to the constructor, including observations, model parameters, model mechanics such as simulators and the measurement density, covariates, and times. After the components are passed into the constructor, the constructor automatically generates additional internal elements, such as extended observations and covariates required for interpolation

Table 2 summarizes the main arguments to the Pomp constructor and their correspondence to mathematical objects. RInit, RProc and RMeas are the subconstructors, where they specify the stochastic mechanisms that define a Pomp model, and their objects are passed as arguments when initializing a main Pomp class.

Table 2: Main arguments to the Pomp class and related constructor objects.

Constructor/Sub-			Description / Mathematical
constructor	Argument	Type	representation
Pomp	rinit	RInit	simulate initial states $X_0 \sim f_{X_0}(x_0; \theta)$
	rproc	RProc	simulate state transitions
			$X_n \sim f_{X_n \mid X_{n-1}}(x_n \mid x_{n-1}; \theta)$
	rmeas	RMeas	simulate observations
			$Y_n \sim f_{Y_n \mid X_n}(y_n \mid x_n; \theta)$
	dmeas	DMeas	evaluate measurement density
			$f_{Y_n X_n}(y_n \mid x_n; \theta)$
	ys	pandas.DataFrame	observations $y_{1:N}^*$ with times $t_{1:N}$
	covars	pandas.DataFrame	covariates $z_{1:N}^*$ with times $s_{1:N}$
	theta	list or dict	parameters θ
RInit	t0	float	initial time point t_0 for simulation
RProc	step_type	str	method of process evolution:
			"fixedstep" or "euler"
	nstep	int	number of steps if
			step_type="fixedstep"
	dt	float	time step if step_type="euler"
	accumvars	tuple	indices of state variables to be
			accumulated
RMeas	ydim	int	observation dimension $\dim(Y)$

We demonstrate here how to create a Pomp object. Specifically, we show how to create the linear Gaussian model included in the package as LG(). We begin by importing necessary packages and defining helper functions for handling the parameters. Because pypomp will run our defined model components within JAX JIT-compiled code, it is necessary to write the components to be JAX-compliant. Naturally, the JAX package has many useful functions for this purpose. We also generate a pseudorandom number generation (PRNG) key to be used with JAX random number generators. All stochastic simulations in pypomp are controlled via JAX PRNG keys, ensuring full reproducibility when using the same seed.

```
import pypomp as pp
import pandas as pd
import jax
import jax.numpy as jnp
from functools import partial

def get_thetas(theta):
```

```
A = jnp.array(
      [theta["A11"], theta["A12"], theta["A21"], theta["A22"]]
    ).reshape(2, 2)
    C = jnp.array(
      [theta["C11"], theta["C12"], theta["C21"], theta["C22"]]
    ).reshape(2, 2)
    Q = jnp.array(
      [theta["Q11"], theta["Q12"], theta["Q21"], theta["Q22"]]
    ).reshape(2, 2)
    R = jnp.array(
      [theta["R11"], theta["R12"], theta["R21"], theta["R22"]]
    ).reshape(2, 2)
    return A, C, Q, R
def transform_thetas(A, C, Q, R):
    return jnp.concatenate([A.flatten(), C.flatten(), Q.flatten(), R.flatten()])
rw_sd = pp.RWSigma(
  sigmas={
    "A11": 0.02, "A12": 0.02,
    "A21": 0.02, "A22": 0.02,
    "C11": 0.02, "C12": 0.02,
    "C21": 0.02, "C22": 0.02,
    "Q11": 0.02, "Q12": 0.02,
    "Q21": 0.02, "Q22": 0.02,
    "R11": 0.02, "R12": 0.02,
    "R21": 0.02, "R22": 0.02,
  },
  init_names=[],
)
key = jax.random.key(1)
```

Model Components

We refer to model components describing initialization, transfer, or measurement processes as the model mechanics, including rinit, rproc, dmeas, and rmeas. Users must define these mechanics as Python functions. Specifically, we require users to provide function code to the object constructor, which verifies that all necessary function arguments are included and in the

correct order. This requirement stems from pypomp's internal mechanism: it vectorizes component functions using jax.vmap() to efficiently run thousands of particles. Since jax.vmap() maps functions to input arrays by position rather than keyword, users must strictly adhere to parameter order. While all expected arguments must be included, the function does not need to utilize all of them. We include here an example defining the model mechanics for the linear gaussian model. In practice, at least one of dmeas or rmeas must be provided, while definitions for rinit and rproc are always required.

```
def rinit(theta , key, covars=None, t0=None):
    """Initial state process simulator for the linear Gaussian model"""
    A, C, Q, R = get_thetas(theta_)
    result = jax.random.multivariate_normal(key=key, mean=jnp.array([0, 0]), cov=Q)
    return {"X1": result[0], "X2": result[1]}
def rproc(X_, theta_, key, covars=None, t=None, dt=None):
    """Process simulator for the linear Gaussian model"""
    A, C, Q, R = get_thetas(theta_)
    X_array = jnp.array([X_["X1"], X_["X2"]])
    result = jax.random.multivariate normal(key=key, mean=A @ X array, cov=Q)
    return {"X1": result[0], "X2": result[1]}
def dmeas(Y_, X_, theta_, covars=None, t=None):
    """Measurement model distribution for the linear Gaussian model"""
    A, C, Q, R = get_thetas(theta_)
    X_{array} = jnp.array([X_["X1"], X_["X2"]])
    return jax.scipy.stats.multivariate_normal.logpdf(Y_, X_array, R)
def rmeas(X_, theta_, key, covars=None, t=None):
    """Measurement simulator for the linear Gaussian model"""
    A, C, Q, R = get_thetas(theta_)
    X_array = jnp.array([X_["X1"], X_["X2"]])
    return jax.random.multivariate normal(key=key, mean=C @ X array, cov=R)
```

Parameters

The Pomp constructor also requires model parameters. These can be provided either as a dictionary or as a list of dictionaries. Each item in a dictionary should include the parameter name as the key and the parameter value as the dictionary value.

If the parameter sets are provided as a list of dictionaries, methods such as pfilter()

run on each set of parameters. Here, we use Pomp.sample_params() to sample sets of parameters from uniform distributions with bounds passed as a dictionary of length-2 tuples. Pomp.sample_params() returns a ready-to-use list of dictionaries with the sampled parameters.

```
theta = {
    "A11": jnp.cos(0.2), "A12": -jnp.sin(0.2),
    "A21": jnp.sin(0.2), "A22": jnp.cos(0.2),
    "C11": 1.0,
                          "C12": 0.0,
    "C21": 0.0,
                           "C22": 1.0,
    "Q11": 0.01,
                           "Q12": 1e-6,
    "Q21": 1e-6,
                           "Q22": 0.01,
    "R11": 0.1,
                           "R12": 0.01,
    "R21": 0.01,
                           "R22": 0.1,
param bounds = \{k: (v * 0.9, v * 1.1) \text{ for } k, v \text{ in theta.items()}\}
n = 5
key = jax.random.key(1)
key, subkey = jax.random.split(key)
theta list = pp.Pomp.sample params(param bounds, n, subkey)
```

Covariates

POMP models often involve external time-varying inputs, referred to as covariates, which can influence either the latent process or the measurement model. Examples include seasonality, interventions, or environmental drivers in ecological applications. In pypomp, covariates are supplied as a pandas.DataFrame indexed by time. The covariate time points may differ from the observation times, so the Pomp object automatically interpolates covariate values at the observation times. The linear gaussian model does not involve any covariates, and an example using covariates is given in the Data Analysis Section.

POMP Object Construction

We do not have real data in this LG example, so we generate our own. To make this example cleaner, we use the function LG() to construct the completed linear Gaussian model object and then generate the data using simulate(). Observation times are provided to the Pomp constructor via the pandas.DataFrame row index. If covariates were provided, the times at which the covariates were observed would also be provided by the pandas.DataFrame row index.

```
T = 100
# ensure `key` exists; if not, uncomment the next line
# key = jax.random.PRNGKey(1)
key, subkey = jax.random.split(key)
_, ys = pp.LG(T=T).simulate(key=subkey)
ys = ys.rename(columns={"obs_0": "Y1", "obs_1": "Y2"})
ys = ys[["time", "Y1", "Y2"]]
ys.set_index("time", inplace=True)
LG_obj = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    t0=0.0,
    nstep=1,
    ydim=2,
    ys=ys,
    theta=theta_list,
    statenames=["X1", "X2"],
)
print("LG_obj created; ys.shape =", ys.shape)
```

```
LG_obj created; ys.shape = (100, 2)
```

Each argument to Pomp is accessible from the object as an attribute.

```
print(LG_obj.rinit) # access POMP model components
print(LG_obj.rproc)
print(LG_obj.dmeas)
print(LG_obj.rmeas)
print(LG_obj.theta) # access parameters
print(LG_obj.ys.head()) # access observations
```

```
<pypomp.model_struct.RInit object at 0x11b186d80>
<pypomp.model_struct.RProc object at 0x1182383b0>
<pypomp.model_struct.DMeas object at 0x11b086ba0>
<pypomp.model_struct.RMeas object at 0x11ac5a570>
[{'A11': 0.963512122631073, 'A12': -0.17880238592624664, 'A21': 0.20370502769947052, 'A22':
```

	Y1	Y2
time		
1.0	-0.116254	0.725715
2.0	-0.354118	0.231588
3.0	0.078726	-0.153802
4.0	0.147694	-0.299703
5.0	0.655032	0.523176

Premade models:

Beyond the linear gaussian model, pypomp includes several ready-to-use model constructors that serve as examples for what users can do:

- 1. LG() a simple linear-Gaussian model with 2-dimensional latent and observed states; useful to validate API usage and diagnostics.
- 2. spx() the S&P500 log-return model from Sun et al. (Sun 2024).
- 3. dacca() the cholera transmission model from King et al. (King et al. 2008).
- 4. UKMeasles.Pomp() the measles district model from He et al. (He, Ionides, and King 2010), extended to the Korevaar et al. dataset (Korevaar, Metcalf, and Grenfell 2020). Panel and spatial variants (PanelPOMP/SpatPOMP style) are planned.

These examples demonstrate correct model specification. If a user's model throws an error or runs slowly, they may compare its components to a premade model to make improvements.

JAX Numerical Backend and Interface Design

A key design choice in pypomp is its usage of JAX as a numerical backend. Unlike the R package pomp, where users typically provide POMP model components in C Snippets for acceleration, pypomp requires model components to be written as JAX-compatible Python functions. These functions are then compiled and vectorized by JAX tools such as jit and vmap. This design leads to several important interface features:

[AA: How necessary is it to mention this restriction?]

• Strict argument requirements for compilation and vectorization: JAX's jit compiler transforms the user-supplied model mechanics functions into efficient machine code, while vmap efficiently run them over thousands of particles via vectorization of arguments. To ensure the compatibility with JAX's compilation and vectorization system, each component function must follow the expected input types and order, otherwise compilation would fail.

[AA: Is this necessary?]

• PRNG random key policy: To ensure reproducibility of random methods in pypomp, public API methods accept an optional jax.random.key, which is explicitly passed through constructors and methods. Unlike the R setting, where randomness can be controlled globally or by seed chunks, in JAX, random keys only be explicitly passed through functions

[AA: Is this necessary? In general, I think the internal details in this section aren't very important. We should go over internal details that lead to some advantage over other packages or which constrain what we can do.]

• Consistent shapes and sizes handling: Model parameters, even multidimensional, are stored as flattened dictionaries. Consequently, JAX can uniformly process parameters, thereby maintaining consistency in particle propagation.

Later section will demonstrate how the JAX-based design supports further inference methods.

[more introduction to JAX?]

Panel POMP class

POMP Methods in pypomp

In this section, we describe the core inference methods currently implemented in pypomp, including:

- Paricle Filter (Sequential Monte Carlo, written in pfilter()): A standard sequential Monte Carlo algorithm for likelihood evaluation and state estimation, forming the basis for most inference methods in POMP models.
- Differentiated Measurement-off-policy Particle Filter (DMOP(α), written in mop()): A recently proposed SMC method (Tan, Hooker, and Ionides 2024) that evaluates the likelihood at one parameter value while obtaining resampling decisions from another, adjusting via discounted off-parameter measurement weights.
- Iterated Filtering (IF2, mif()): A classical IF2 algorithm (Edward L. Ionides et al. 2015) for likelihood-based parameter inference that maximizes the likelihood via particle filtering.
- Iterated Filtering with Automatic Differentiation (train()): A recently proposed AD-based algorithm (Tan, Hooker, and Ionides 2024) that leverages $\mathrm{DMOP}(\alpha)$ perform maximum likelihood estimation with gradient descent methods.

A key feature of the above POMP inference methods lies in the **plug-and-play property** (E. L. Ionides, Breto, and King 2006), meaning that inference algorithms can be implemented without requiring explicit evaluation of the transition density of the latent process. Instead, it

suffices for the user to provide a the latent process simulator (rproc), the initial state simulator (rinit), and the observation density (dmeas). This property enables POMP methods to be widely applied to complex mechanistic models where transition densities are intractable.

In pypomp, the plug-and-play design is fully preserved: users only need to provide component functions compliant with JAX requirements, which can be directly plugged in inference methods likepfilter(), mop(), mif(), and train(). The package combines the generality of plug-and-play modeling with the efficiency of JAX compilation and vectorization.

Unlike the R family of POMP packages, some Pomp class methods including pfilter(), mif() and train() yield results by modifying the object in place instead of returning new objects. All of the results are stored a list under LG_obj.results_history, which is an attibute under Pomp class object LG_obj. Each element in the list correponds to one method call. Each element includes results such as the log-likelihood and parameter estimates when applicable as well as the inputs used for the function call, so it is easy to keep track of how the results were calculated. If multiple parameter sets are supplied in a list as an argument, the method evaluates at each set and the results for each are stored. Meanwhile, each following method in pypomp that takes a key as an argument stores an unused child key under LG_obj.fresh_key that later method calls can use by default when a key argument is not given. This design prevents repeated use of the same key across a sequence of method calls and ensures proper randomness.

```
LG_obj.pfilter(J=100, reps=5, key=subkey)
LG_obj.mif(rw_sd=rw_sd, M=2, a=0.5, J=100, key=subkey)
print(LG_obj.results_history)
```

```
[{'method': 'pfilter', 'logLiks': <xarray.DataArray (theta: 5, replicate: 5)> Size: 100B
Array([[ -98.343124, -98.823456, -104.67427 , -99.98291 , -100.560265],
                                              -91.484856, -89.26564],
       [ -89.46374 , -90.23834 , -92.57293 ,
       [ -94.22773 , -93.686676, -91.52622 ,
                                              -94.595406, -90.94031],
       [-106.548386, -106.57879, -112.516525, -105.55487, -103.90391],
       [ -89.703125, -89.2629 , -86.12627 , -90.139595, -90.03055 ]],
                                                                              dtype=float3
Dimensions without coordinates: theta, replicate, 'theta': [{'A11': 0.963512122631073, 'A12'
[2705940334 2639757084], 'execution_time': 1.336474895477295}, {'method': 'mif', 'traces': <
array([[[
                    nan, 9.63512242e-01, -1.78802431e-01,
          2.03704998e-01, 1.06974924e+00, 9.99111295e-01,
         0.0000000e+00, 0.0000000e+00, 1.05488253e+00,
          1.04723452e-02, 9.36353899e-07, 1.03818627e-06,
          1.08780684e-02, 1.02496698e-01, 1.07683847e-02,
          9.18221567e-03, 9.63022783e-02],
        [-1.39011429e+02, 8.68815184e-01, -2.80453652e-01,
          2.47058019e-01, 8.01628411e-01, 1.02020454e+00,
```

```
1.47310942e-01, -1.80884093e-01, 7.90077806e-01,
          1.53452471e-01, -6.92871958e-02, 1.97661936e-01,
          1.40399724e-01, 8.93150866e-02, -8.61471519e-02,
          3.39505449e-02, 1.25918418e-01],
        [-1.36756348e+02, 8.44347954e-01, -1.31342560e-01,
          4.17445600e-01, 6.97917819e-01, 8.59905064e-01,
          1.40382484e-01, -2.28263423e-01, 7.12622881e-01,
         8.68392512e-02, -3.04817408e-01, 2.71064401e-01,
         1.47589087e-01, 9.27633867e-02, -8.00307095e-03,
        -8.86493027e-02, 1.66606173e-01]],
       [[
                    nan, 9.66489673e-01, -1.78802431e-01,
. . .
         4.57636297e-01, 1.55186519e-01]],
       nan, 9.77908611e-01, -1.78802431e-01,
          1.95180431e-01, 9.28941786e-01, 1.05515659e+00,
          0.00000000e+00, 0.00000000e+00, 9.39434111e-01,
          1.03248833e-02, 1.08449956e-06, 1.02173408e-06,
          1.00193741e-02, 1.08066812e-01, 1.02444272e-02,
          1.04886629e-02, 9.63648930e-02],
        [-1.41777817e+02, 7.39164770e-01, -1.31324336e-01,
         9.75261852e-02, 7.19573200e-01, 1.14779556e+00,
        -7.76170706e-03, 7.32917413e-02, 1.23638988e+00,
          1.27380058e-01, 2.68890321e-01, -1.15902096e-01,
          2.53278881e-01, 1.99454665e-01, -9.21754837e-02,
          1.62946373e-01, 1.66802466e-01],
        [-1.49710815e+02, 6.46324933e-01, -5.29466346e-02,
          1.62493661e-01, 5.10818541e-01, 8.47000837e-01,
          1.90395802e-01, 2.71731377e-01, 1.18122673e+00,
          1.03760451e-01, 3.05648297e-01, -3.77437323e-01,
          1.42512158e-01, 1.17045403e-01, -1.69405669e-01,
         7.85537064e-02, 1.90754697e-01]]])
Coordinates:
  * replicate (replicate) int64 40B 0 1 2 3 4
  * iteration (iteration) int64 24B 0 1 2
              (variable) <U6 408B 'logLik' 'A11' 'A12' ... 'R12' 'R21' 'R22', 'theta': [{'A
[2705940334 2639757084], 'execution_time': 1.6023688316345215}]
```

Particle Filter (pfilter)

[outline: 1. purpose/role 2. implementation details in pypomp 3. outputs/results 4. remarks/highlights]

The particle filter algorithm (sequential Monte Carlo, SMC) (Arulampalam et al. (2002), King, Nguyen, and Ionides (2016)) is the core classical likelihood evaluation tool in POMP model inference. It uses Monte Carlo techniques to sequentially estimate the integrals in the prediction and filtering recursions, and produce an unbiased Monte Carlo estimate of the likelihood of the observed data under a parameter set. The likelihood estimate is essential for other parameter inference algorithms such as iterative filtering. In pypomp, the basic particle filter is implemented under pfilter(). The algorithm is summarized as Algorithm 1,

Algorithm 1 Sequential Monte Carlo (SMC, or particle filter) in textttpypomp: LG_obj.pfilter(J=J, reps=reps, key=key), where LG_obj is a class Pomp object with components rinit, rproc, dmeas, rmeas, ys, and theta

Require Simulator for process_n(·|x_n; θ); Evaluator for $f_{Y_n|X_n}(y_n \mid x_n; \theta)$; Simulator for $f_{X_0}(x_0; \theta)$; Parameter θ ; Data $y_{1:N}^*$; Number of particles J.

Initialize particles: $X_{0,j}^{F,\theta} \sim f_{X_0}(\cdot;\theta)$ for j=1:J.

For n = 1, ..., N:

Simulate prediction particles: $X_{n,j}^{P,\theta} \sim \operatorname{process}_n(\cdot|x_n;\theta)$ for j=1:J.

Evaluate weights and measurement density: $w(n,j) = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P,\theta}; \theta)$ for j = 1:J.

Normalize weights $\tilde{w}(n,j) = \frac{w(n,j)}{\sum_{m=1}^{J} w(n,m)}$ for j=1:J.

Select resample indices $k_{1:J}$ with $P(k_j = m) = \tilde{w}(n, m)$.

Obtain resampled particles $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$ for j = 1:J.

Compute conditional log likelihood:

$$\hat{\ell}_{n|1:n-1} = \log \left(\frac{1}{J} \sum_{m=1}^{J} w(n,m) \right).$$

Return Log likelihood estimate $\hat{\ell}(\theta) = \sum_{n=1}^{N} \hat{\ell}_{n|1:n-1}$; filter samples $X_{n,1:J}^{F,\theta}$ for n=1:N. Complexity: $\mathcal{O}(J)$

In pypomp, the pfilter() functions is internally run in a compiled JAX routine pfilter_internal() but wrapped up into a class method. _pfilter_internal is JIT-compiled with static arguments, in which the replicates and parameter sets are vectorized via jax.vmap. Both techniques enable execution on CPU/GPU and ensure the computational efficiency. In implementation, the weights and likelihood are stored in a log scale, preventing underflow. There are optional arguments (CLL, ESS, filter_mean, prediction_mean) control wheter additional diagnostic arrays are allocated and updated. The algorithm works for

flexible and extendable models through the user-specified Pomp class components (rinit, rproc, dmeas, reams), number of particles J as well as other parameters. Each run returns a dict type element updated inside the LG_obj. result_history attribute, containing the log-likelihoods, algorithm parameters used, as well as model diagonostic elements at each time included if their respective boolean flags are set to True. For example, suppose we run

```
LG_{obj_2} = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    t0=0.0,
    nstep=1,
    ydim=2,
    ys=ys,
    theta=theta list,
    statenames=["X1", "X2"],
)
LG_obj_2.pfilter(J=1000, reps=10, key=subkey)
LG_obj_2.pfilter(
    J=1000,
    reps=10,
    key=subkey,
    CLL=True,
    ESS=True,
    filter_mean=True,
    prediction_mean=True,
)
```

where J is the number of particles used and reps is the number of particle filtering replicates to run for each parameter set provided in the Pomp object or as an optional argument to pfilter(). Because LG_obj2.result_history begins as an empty list here when the model is constructed, the results are appended at LG_obj_2.results_history[0] and LG_obj_2.results_history[1] respectively. Both of these two dictionaries contain the following items:

- method: The method that was run. In this case, pfilter.
- logLiks: An xarray.DataArray of particle filter log-likelihood estimates. There is one list for each parameter set given and each list has reps log-likelihood estimates. dimensions = (theta, replicate)
- theta: The list of parameter sets used.

- J: The number of particles used.
- thresh: Threshold value for resampling used, which is 0 by default.

Meanwhile, $LG_obj_2.results_history[1]$ also contains the following items that are not contained in $LG_obj_2.results_history[0]$:

- CLL: An xarray.DataArray containing the conditional loglikelihood over observed time steps (theta, replicate, time)
- ESS: An xarray.DataArray containing the effective sample size over observed time steps (theta, replicate, time)
- filter_mean: An xarray.DataArray containing the filtered mean over observed time steps (theta, replicate, time, state)
- prediction_mean: An xarray.DataArray containing the prediction mean over observed time steps. (theta, replicate, time, state)

A key design is that all runs across multiple parameter sets and replications are stored in one results_history list under the corresponding Pomp object. This avoids object duplication, and increases reproducibility by facilitating pickling Pomp objects obtained from many different search strategies, loading them into a new Python session, and then comparing the results to see how they vary with different algorithmic parameters.

Differentiated Measurement Off-Parameter (DMOP) particle filter

([TO BE REWRITTEN, FOLLOWING ISSUE 2])

[Motivation for DMOP] Classical particle filters for POMP models produce unbiased Monte Carlo estimates, but are non-differentiable in model parameters θ , due to the discrete random sampling of particles. This prevents the wide use of AD and gradient-based optimization in POMP dynamical systems. The differentiable measurement off-parameter particle filter (DMOP- α , Tan, Hooker, and Ionides (2024)) provides a solution: it treats a differentiable likelihood estimator produced by a variant of the standard particle filter as an objective and apply AD to obtain $\nabla_{\theta}\hat{\ell}(\theta)$ for likelihood-based optimal parameter search. The derivative estimate from DMOP- α is essential for further stochastic gradient inference methods, such as iterated filtering with automatic differentiation (IFAD), where it employs stochastic gradient descent using DMOP- α . A tunning dicount factor $\alpha \in [0,1]$ controls the bias-variance trade-off of the gradient estimator: larger α may result in lower bias and higher variance, while smaller α can reduce the variance with the cost of increasing bias. In implementation, DMOP- α does not require heavy diagnostics or multiple replicates. It only needs a stable, differentiable of the (log-) likelihood to feed into jax.grad and jax.value_and_grad. The pseudocode is shown in 2

Algorithm 2 DMOP- α in pypomp: jax.value_and_grad(LG_obj.mop(J=J, reps=reps, key=key, alpha=alpha)). LG_obj is a Pomp object with components rinit, rproc, dmeas, ys, theta.

 $\begin{array}{l} \textbf{Require Simulator process}_n(\cdot|x_n;\theta); \ \text{Evaluator for} \ f_{Y_n|X_n}(y_n^*|x_n,\theta); \ \text{Simulator for} \ f_{X_0}(\cdot;\theta); \\ \text{Evaluation parameter} \ \theta; \ \text{Data} \ y_{1:N}^*; \ \text{Number of particles } J; \ \text{Discount factor} \ \alpha. \\ \textbf{Initialize particles} \ X_{0,j}^{F,\theta} \sim f_{X_0}\left(\cdot;\,\theta\right), \ \text{weights} \ w_{0,j}^{F,\theta} = 1 \ \text{for} \ j=1:J. \\ \end{array}$

For n = 1, ..., N:

Simulate prediction particles, $X_{n,j}^{P,\theta} \sim \operatorname{process}_n(\;\cdot\;|X_{n-1,j}^{F,\theta};\theta)$ for j=1:J. Accumulate discounted prediction weights, $w_{n,j}^{P,\theta}=(w_{n-1,j}^{F,\theta})^{\alpha}$.

Evaluate measurement density, $g_{n,j}^{\theta} = f_{Y_n|X_n}(y_n^*|X_{n,j}^{P,\theta};\theta)$ for j=1:J. Compute conditional likelihood in version B: $L_n^{B,\theta,\alpha} = \sum_{j=1}^J g_{n,j}^{\theta} \, w_{n,j}^{P,\theta} \, \Big/ \sum_{j=1}^J w_{n,j}^{P,\theta}$.

Select resampling indices $k_{1:J}$ with $P(k_j=m) \propto g_{n,m}^{\theta}$. Obtain resampled particles $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$.

Calculate corrected weights $w_{n,j}^{F,\theta} = w_{n,k_j}^{P,\theta}$ g_{n,k_j}^{θ} stop_gradient(g_{n,k_j}^{θ}).

Compute considtional likelihood in version A: $L_n^{A,\theta,\alpha} = L_n^{\phi} \cdot \sum_{j=1}^J w_{n,j}^{F,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}$.

Forward Pass: AD constructs a computation graph to evaluate the likelihood estimate $\hat{L}^A(\theta) = \prod_{n=1}^N L_n^{A,\theta,\alpha}$ or $\hat{L}^B(\theta) = \prod_{n=1}^N L_n^{B,\theta,\alpha}$.

Backward Pass: AD propagates gradients through the computation graph and evaluates

gradient estimate $\mathcal{D}^A(\theta)$ or $\mathcal{D}^B(\theta)$

Return: likelihood estimate $\hat{L}^A(\theta)$ or $\hat{L}^B(\theta)$, and gradient estimate $\hat{\mathcal{D}}^A(\theta)$ or $\hat{\mathcal{D}}^B(\theta)$.

The key building block of the DMOP- α is the MOP- α algorithm, a variant of the standard particle filter designed to make the discontinuous Monte Carlo resampling process independent of the target parameter θ . It evaluates the likelihood at the evaluation parameter θ but draws resampling indices using weights computed at another baseline behavior parameter ϕ . This renders the likelihood estimates differentiable w.r.t. θ , further enabling AD.

In DMOP- α , the same idea is implemented equivalently by setting $\phi \equiv \text{stop_gradient}(\theta)$ in JAX. This prevent its input from being differentiated. The filtering process therefore runs $\theta = \phi$, assuming the evaluation parameter and the bahavior parameters have identical values. This avoids the need to explicitly introduce new parameters while preserving the differentiable property. Consequently, the likelihood evaluation of DMOP is identical to that of the standard particle filter, and it does not require implement a fully separate MOP- α procedure. This design simplifies the algorithm and still ensures the valid likelihood and gradient estimator, while maintaining differentiability with respect to θ . That is, DMOP computes the derivative of the likelihood estimate obtained by MOP when $\theta = \phi$. The pseudocode of MOP is shown in 3.

Algorithm 3 MOP(α), Measurement off-policy sequential Monte Carlo

Require Simulator for $\operatorname{process}_n(\cdot|x_n;\theta)$; Evaluator for $f_{Y_n|X_n}(y_n\mid x_n;\theta)$; Simulator for $f_{X_0}(x_0;\theta)$; Evaluation Parameter θ ; Behavior Parameter ϕ ; Data $y_{1:N}^*$; Number of particles J. **Initialize** filter particles $X_{0,j}^{F,\theta} \sim f_{X_0}(\cdot;\theta)$ and $w_{0,j}^{F,\theta} = 1$ for j = 1:J. For n = 1, ..., N:

Simulate prediction particles: $X_{n,j}^{P,\theta} \sim f_{X_n|X_{n-1}}(\cdot \mid X_{n-1,j}^{F,\theta}; \theta)$ for j=1:J. Accumulate discounted Prediction weight: $w_{n,j}^{P,\theta} = \left(w_{n-1,j}^{F,\theta}\right)^{\alpha}$ for j=1:J.

Evaluate measurement density: $g_{n,j}^{\theta} = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^{P,\theta}; \theta)$ for j=1:J. Compute conditional likelihood under θ in version B: $L_n^{B,\theta,\alpha} = \sum_{j=1}^J g_{n,j}^{\theta} w_{n,j}^{P,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}$.

Compute conditional likelihood under $\phi:L_n^{\phi} = \frac{1}{J}\sum_{m=1}^J g_{n,m}^{\phi}$.

Select resample indices $k_{1:J}$ with $P(k_j = m) \propto \tilde{g}_{n,m}^{\phi}$. Obtain resampled particles $X_{n,j}^{F,\theta} = X_{n,k_j}^{P,\theta}$ for j = 1:J.

Calculate resampled corrected weights: $w_{n,j}^{F,\theta} = w_{n,j}^{P,\theta} \times \frac{g_{n,j}^{\theta}}{g_{n,j}^{\phi}}$ for j=1:J.

Compute conditional likelihood under θ in version A: $L_n^{A,\theta,\alpha} = L_n^{\phi} \cdot \sum_{j=1}^J w_{n,j}^{F,\theta} / \sum_{j=1}^J w_{n,j}^{P,\theta}$.

Return: Likelihood estimate: $\hat{L}(\theta) = \prod_{n=1}^N L_n^{A,\theta,\alpha}$ or $\hat{L}(\theta) = \prod_{n=1}^N L_n^{B,\theta,\alpha}$.

In pypomp, DMOP- α is implemented by wrapping the JAX-compiled mop() function inside JAX AD. It is a Pomp class method calling an internal _mop_internal(), which uses static arguments and jax.lax.fori_loop for efficiency, and stores weights and likelihoods in log scale to prevent underflow. Systematic sampling is involved for the resampling process. Users can directly apply jax.grad and jax.value_and_grad to mop() to obtain gradients.

In practice, the mop() provides an auto-differentiable alternative to pfilter(), as a lightweight version. Unlike pfilter(), which provides the choice of multiple replicates and diagnostic summaries, mop() in pypomp focuses purely on producing differentiable likelihood estimates. This makes the algorithm suitable for AD and stochastic gradient descent algorithms in accelerated iterated filtering algorithms by setting the function as input of jax.value_and_grad() and jax.grad(). Beyond a well-constructed Pomp object, the method requires only the number of particles J, the evaluation parameter theta, a PRNG key, and the cooling factor α . Results are stored in the results_history attribute as a list of jax. Array objects, one for each parameter set, representing the estimated log-likelihood of the observations. For example, we run

```
LG_{obj_3} = pp.Pomp(
    rinit=rinit,
    rproc=rproc,
    dmeas=dmeas,
    rmeas=rmeas,
    t0=0.0,
```

```
nstep=1,
  ydim=2,
  ys=ys,
  theta=theta_list,
  statenames=["X1", "X2"],
)

LG_obj_3.mop(J=1000, key=subkey, alpha=0.97)
```

```
[Array(-98.85261, dtype=float32),
Array(-89.553665, dtype=float32),
Array(-91.47145, dtype=float32),
Array(-102.464615, dtype=float32),
Array(-89.382, dtype=float32)]
```

Generally, the result_history only updates:

• log-likelihood(s): a list of jax.Array containing the estimated log-likelihood(s) of the observed data given the model parameters. Always a list, even if only one theta is provided.

This design highlights mop() as a lightweight but essential building block for accelarated gradient-based inference methods within pypomp. By combining jax.grad and jax.value_and_grad around mop(), the package effectively implements the DMOP algorithm, enabling stochastic-gradient updates for likelihood-based inference on POMP models.

Iterated Filtering

Iterated filtering (IF2) provides a Monte Carlo approach to estimate parameters by combining particle filtering with sequential random-walk perturbations and cooling. Compared with a plain particle filter or MOP method, IF2 not only evaluates the likelihoods but also systematically updates parameters across iterations. Its primary role is to estimate the parameters that maximizes the likelihood, given the likelihood is intractable but can be estimated under the particle filtering methods. The pseudocode for IF2 algorithm is as follows:

Algorithm 4 Iterated Filtering (IF2)

```
Require: Starting parameter \theta_0; simulator for f_{X_0}(x_0;\theta); simulator for f_{X_n\mid X_{n-1}}(x_n\mid x_{n-1};\theta); evaluator for f_{Y_n\mid X_n}(y_n\mid x_n;\theta); data y_{1:N}^*; labels I\subset\{1,\dots,p\} for IVPs; fixed lag L;
        number of particles J; number of iterations M; cooling rate a, 0 < a < 1; perturbation
        scales \sigma_{1:n}; initial scale multiplier C > 0.
   1: for m = 1 to M do
               Initialize parameters: \Theta_{0,j,i}^P \sim \text{Normal}([\theta_{m-1}]_i, (Ca^m\sigma_i)^2) for i \in 1:p, j \in 1:J.
   2:
               Initialize states: simulate X_{0,j}^F \sim f_{X_0}(\cdot; \Theta_{0,j}^P) for j = 1:J.
  3:
               Initialize filter mean: \bar{\theta}_0 = \theta_{m-1}.
Define [V]_i = (C^2 + 1)a^{2m}\sigma_i^2.
   4:
  5:
               for n = 1 to N do
  6:
                      Perturb parameters: \Theta_{n,j,i}^P \sim \text{Normal}([\Theta_{n-1,j}^F]_i, (a^m \sigma_i)^2) for i \notin I, j = 1:J.
   7:
                     Simulate prediction particles: X_{n,j}^P \sim f_{X_n|X_{n-1}}(\cdot \mid X_{n-1,j}^F; \Theta_{n,j}^P) for j=1:J. Evaluate weights: w(n,j) = f_{Y_n|X_n}(y_n^* \mid X_{n,j}^P; \Theta_{n,j}^P) for j=1:J. Normalize weights: \tilde{w}(n,j) = \frac{w(n,j)}{\sum_{u=1}^J w(n,u)}. Resample indices k_{1:J} with \Pr[k_u=j] = \tilde{w}(n,j). Resample particles: X_{n,j}^F = X_{n,k_j}^P and \Theta_{n,j}^F = \Theta_{n,k_j}^P for j=1:J.
  8:
  9:
 10:
11:
 12:
                      Filter mean: [\bar{\theta}_n]_i = \sum_{j=1}^J \tilde{w}(n,j) [\Theta_{n,j}^P]_i for i \notin I.
 13:
                      Prediction variance: [V_{n+1}]_i = (a^m \sigma_i)^2 + \sum_{i=1}^J \tilde{w}(n,j) \left( [\Theta_{n,j}^P]_i - [\bar{\theta}_n]_i \right)^2 for i \notin I.
14:
 15:
               Update non-IVPs: [\theta_m]_i = [\theta_{m-1}]_i + [V]_i \sum_{n=1}^N ([\bar{\theta}_n]_i - [\theta_{m-1}]_i) for i \notin I.
 16:
               Update IVPs: [\theta_m]_i = \frac{1}{I} \sum_{i=1}^J [\Theta_{L,i}^F]_i for i \in I.
17:
18: end for
```

Under pypomp, the IF2 is implemented under mif(). Arguments sigmas and sigmas_init control the random walk standard deviation after and during time t_0 respectively. Another argument a, the cooling rate, reduces the perturbation magnitudes across iterations. The purturbations controlled by sigmas, sigmas_init, and a enable the POMP model to explore the parameter space and progressively converge to the optimal parameter estimates. Specifically, each call also updates and stores a child key into $LG_obj.fresh_key$, so subfequent method calls can safely omit an explicit assignment on key argument.

Ensure: Monte Carlo maximum likelihood estimate θ_M .

Different from previous methods for evaluating the likelihoods, Pomp methods such as mif() can update attributes directly other than resuts_history. The LG_obj.mif() replaces LG_obj.theta with the parameter estimate from the end of the last iteration. The traces of log-likelihoods and parameter estimates are stored in LG_obj.results_history as a xr.DataArray object. The xr.DataArray has dimentions (replicate, iteration, variable), where the variable dimension includes the log-likelihood estimations and all

POMP model parameters.

:

```
key, subkey = jax.random.split(key)
\#LG_{obj_4} = pp.Pomp(
     rinit=rinit,
 #
     rproc=rproc,
 #
     dmeas=dmeas,
    rmeas=rmeas,
 #
     ys=ys,
 #
     theta=theta_list,
 #
     covars=None,
#)
#LG_obj_4.mif(
  #sigmas=0.02,
  #sigmas_init=0.1,
  #J=1000,
  #M=100,
  #a=0.5, # cooling rate
  #key=subkey
#)
#LG_obj_4.pfilter(J=1000, reps=36)
#LG obj 4.mif(
    #sigmas=0.005,
    #sigmas_init=0.0025,
    #J=1000,
    #M=100,
    #a=0.5,
#)
#LG_obj_4.pfilter(J=1000, reps=36)
```

The pypomp implementation of IF algorithm in mif() inherits the structure of R's pomp package while adding JAX compilation, providing a substantial performance improvements over R implementation.

Iterated Filtering with Automatic Differentiation

Iterated filtering with automatic differentiation (IFAD) (Tan, Hooker, and Ionides 2024) is a hybrid method that combines the complementary advantages of iterated filtering (IF2) and a new proposed MOP- α gradient method. Here by using the differentiable particle filter, the

 $MOP-\alpha$ gradient method can obtain gradient and Hessian estimates of the log-likelihood of a POMP model. These estimates are involved in an optimization method to perform a gradientbased search in the parameter space. While IF2 can quickly approach the neighborhood of the global maximum, it struggles to achieve exact local convergence to the global maximum. It can take dozens of iterations to partially close the log-likelihood gap between a point in the neighborhood and the actual global maximum. In contrast, the new MOP- α gradient method can conduct a fast and strong local search, but shows weaker ability at conducting a global exploration (Tan, Hooker, and Ionides 2024). IFAD addresses this tradeoff by first running IF2 to approach the neighborhoold of the global maximum, and then swithcing to the gradient-based method to further refine the parameter estimates. NOTE:Global Maximum or Local Maximum?

Algorithm 5 IFAD: Iterated Filtering with Automatic Differentiation

Require: Number of particles J, timesteps N, IF2 cooling schedule η_m , MOP- α discounting parameter α , initial parameter θ_0 , iteration index m=0.

- 1: Run IF2 until initial "convergence" under cooling schedule η_m , or for a fixed number of iterations, to obtain $\{\Theta_j, j=1,\ldots,J\}$.
- 2: Set $\theta_m := \frac{1}{J} \sum_{j=1}^J \Theta_j$. 3: **while** procedure not converged **do**
- Run Algorithm 3 (MOP- α filter) to obtain $\hat{\ell}(\theta_m)$. 4:
- Obtain gradient and Hessian: 5:

$$g(\theta_m) = \nabla_{\theta_m} \big(- \hat{\ell}(\theta_m) \big), \quad H(\theta_m) \quad \text{s.t. } \lambda_{\min}(H(\theta_m)) > c.$$

Update parameter: 6:

$$\theta_{m+1} := \theta_m - \eta_m \, H(\theta_m)^{-1} g(\theta_m).$$

- Set m := m + 1.
- 8: end while

Ensure: Return $\hat{\theta} := \theta_m$.

In pypomp, the MOP- α gradent local search is implemented under train(), which utilizes JAX's automatic differentiation features with MOP- α to compute the gradient and Hessian of the log-likelihood. While IFAD is not implemented as an independent function, users can always run mif() and train() to execute the two phases sequentially. For example:

```
key, subkey = jax.random.split(key)
\#LG_{obj_5} = pp.Pomp(
     rinit=rinit,
     rproc=rproc,
#
     dmeas=dmeas,
#
#
     rmeas=rmeas,
```

```
#
     ys=ys,
#
     theta=theta_list,
#
     covars=None,
#)
#LG_obj_5.mif(
#
    sigmas=0.02,
#
    sigmas_init=0.1,
# J=1000,
    M=40,
#
#
    a=0.5.
#
    key=subkey
#)
#LG_obj_5.train(
     J=1000,
#
     itns=40,
#
     eta=0.0025, # learning rate
#
     optimizer="Newton"
#)
```

Here, mif() runs IF2 for 40 iterations, followed by train() runs the Newton's method with MOP- α for 40 iterations.

The implemented function train() serves as a gradient-based optimizer for exploring the parameter space. Each iteration follows the standard optimization working flow: starting from the current parameter estimate, it computes gradient (and Hessian) of the log-likelihood via jax.grad() and jax.Hessian(), then utilizes the choosen internal optimizer to update the direction, obtains the learning rate, and finally updates the parameters as well as evaluates the new likelihood under the updated parameters. train() supports several common gradient-based optimizers for exploring the parameter space, including Newton's method, weighted Newton's method, Broyden-Fletcher-Goldfarb-Shanno (BFGS) algorithm, and gradient descent method. In addition, A linear search method (Wills and Schön 2021) is available under train for determining the learning rate (step size) under stochastic Quasi-Newton methods.

In train(), there is a key argument n_monitors, controlling the number of log-likelihood values are evaluated during each training step. When setting n_monitors=0, the algorithm totally skip the evaluating process and conduct the gradient-only optimization. It is the fastest choice, but doesn't support the linear search method for learning rate determination, as it utilizes the estimated log-likelihood as the current objective function value. n_monitors=1 is a lightest monitoring form, where users can get a one log-likelihood evaluation per training step. Larger n_monitors reduces the log-likelihood estimate variance in particle filtering using Monte Carlo average at additional computational cost. There is another useful Boolean flag, scale, which controls whether the search direction is normalized to unit length. A normalized

search direction can provide a more conservative update when the gradients or Hessians are unstable.

Similarly to mif(), train() updates the result_history with the trace of log-likelihoods and parameter estimates as xr.DataArray object, and replaces LG_obj.theta with the final parameter estimates from the last iteration for each replicate, organized as distionaries keyed by parameter names. The xr.DataArray has dimentions (replicate, iteration, variable). The variable dimension includes the log-likelihood estimations and all POMP model parameters.

The core highlight of IFAD in algorithmic implementation lies in its internal automatic differentiation and just-in-time compilation powered by JAX. Via JAX, gradients and Hessian matrices for the MOP- α log-likelihood function can be directly obtained without manual derivation, ensuring correctness, efficiency and operability. This also enables flexible use of advanced gradient-based optimizers in each training step, such as Newton's method, quasi-Newton methods, BFGS, gradient descent and stochastic line search. Additionally, JAX's GPU compilation capabilities enable these optimization processes to execute efficiently, making IFAD not only theoretically feasible but also capable of effectively handling large complex POMP models in practical algorithmic implementations.

Data Analysis with pypomp / Tutorials and data analysis examples

This section demonstrates: - Log-likelihood profiling - GPU benchmarking - Conditional log-likelihood residuals

Discussion

Abkemeier, Aaron, Jun Chen, Edward Ionides, Jesse Wheeler, and Kevin Tan. 2024. "Pypomp."

Arulampalam, M. S., S. Maskell, N. Gordon, and T. Clapp. 2002. "A Tutorial on Particle Filters for Online Nonlinear/Non-Gaussian Bayesian Tracking." *IEEE Transactions on Signal Processing* 50 (2): 174–88. https://doi.org/10.1109/78.978374.

Asfaw, Kidus, Joonha Park, Aaron A. King, and Edward L. Ionides. 2024. "spatPomp: An R Package for Spatiotemporal Partially Observed Markov Process Models." *Journal of Open Source Software* 9 (104): 7008. https://doi.org/10.21105/joss.07008.

Auger-Méthé, Marie, Ken Newman, Diana Cole, Fanny Empacher, Rowenna Gryba, Aaron A. King, Vianey Leos-Barajas, et al. 2021. "A Guide to State-Space Modeling of Ecological Time Series." *Ecological Monographs* 91 (4): e01470. https://doi.org/10.1002/ecm.1470.

- Blackwood, J. C., D. G. Streicker, S. Altizer, and P. Rohani. 2013. "Resolving the Roles of Immunity, Pathogenesis, and Immigration for Rabies Persistence in Vampire Bats." *Proceedings of the National Academy of Sciences of the United States of America* 110 (51): 20837–42. https://doi.org/10.1073/pnas.1308817110.
- Bradbury, James, Roy Frostig, Peter Hawkins, Matthew James Johnson, Chris Leary, Dougal Maclaurin, George Necula, et al. 2018. "JAX: Composable Transformations of Python+NumPy Programs."
- Bretó, Carles. 2014. "On Idiosyncratic Stochasticity of Financial Leverage Effects." Statistics & Probability Letters 91: 20–26. https://doi.org/10.1016/j.spl.2014.04.003.
- Bretó, Carles, Jesse Wheeler, Aaron A. King, and Edward L. Ionides. 2025. "panelPomp: Analysis of Panel Data via Partially Observed Markov Processes in R." arXiv. https://doi.org/10.48550/arXiv.2410.07934.
- Fox, S. J., M. Lachmann, M. Tec, R. Pasco, S. Woody, Z. Du, X. Wang, et al. 2022. "Real-Time Pandemic Surveillance Using Hospital Admissions and Mobility Data." *Proceedings of the National Academy of Sciences* 119 (7): e2111870119. https://doi.org/10.1073/pnas.2111870119.
- He, Daihai, Edward L. Ionides, and Aaron A. King. 2010. "Plug-and-Play Inference for Disease Dynamics: Measles in Large and Small Populations as a Case Study." *Journal of The Royal Society Interface* 7 (43): 271–83. https://doi.org/10.1098/rsif.2009.0151.
- Ionides, E. L., C. Breto, and A. A. King. 2006. "Inference for Nonlinear Dynamical Systems." *Proceedings of the National Academy of Sciences* 103 (49): 18438–43. https://doi.org/10.1073/pnas.0603181103.
- Ionides, E. L., C. Breto, J. Park, R. A. Smith, and A. A. King. 2017. "Monte Carlo Profile Confidence Intervals for Dynamic Systems." *Journal of The Royal Society Interface* 14 (132): 20170126. https://doi.org/10.1098/rsif.2017.0126.
- Ionides, Edward L., Dao Nguyen, Yves Atchadé, Stilian Stoev, and Aaron A. King. 2015. "Inference for Dynamic and Latent Variable Models via Iterated, Perturbed Bayes Maps." *Proceedings of the National Academy of Sciences* 112 (3): 719–24. https://doi.org/10.1073/pnas.1410597112.
- King, Aaron A., Edward L. Ionides, Mercedes Pascual, and Menno J. Bouma. 2008. "Inapparent Infections and Cholera Dynamics." *Nature* 454 (7206): 877–80. https://doi.org/10.1038/nature07084.
- King, Aaron A., Dao Nguyen, and Edward L. Ionides. 2016. "Statistical Inference for Partially Observed Markov Processes via the *R* Package **Pomp**." *Journal of Statistical Software* 69 (12). https://doi.org/10.18637/jss.v069.i12.
- Korevaar, Hannah, C. Jessica Metcalf, and Bryan T. Grenfell. 2020. "Structure, Space and Size: Competing Drivers of Variation in Urban and Rural Measles Transmission." *Journal of The Royal Society Interface* 17 (168): 20200010. https://doi.org/10.1098/rsif.2020.0010.
- Marino, J. A. Jr., S. D. Peacor, D. B. Bunnell, H. A. Vanderploeg, S. A. Pothoven, A. K. Elgin, J. R. Bence, J. Jiao, and E. L. Ionides. 2019. "Evaluating Consumptive and Nonconsumptive Predator Effects on Prey Density Using Field Time-Series Data." *Ecology* 100 (3): e02583. https://doi.org/10.1002/ecy.2583.
- Mietchen, Matthew S., Erin Clancey, Corrin McMichael, and Eric T. Lofgren. 2024. "Estimat-

- ing SARS-CoV-2 Transmission Parameters Between Coinciding Outbreaks in a University Population and the Surrounding Community." https://doi.org/10.1101/2024.01.10.24301
- Sun, Weizhe. 2024. "Model Based Inference of Stochastic Volatility via Iterated Filtering," April.
- Tan, Kevin, Giles Hooker, and Edward L. Ionides. 2024. "Accelerated Inference for Partially Observed Markov Processes Using Automatic Differentiation." arXiv. https://doi.org/10.48550/arXiv.2407.03085.
- Wen, L., Y. Yin, Q. Li, Z. Peng, and D. He. 2024. "Modeling the Co-Circulation of Influenza and COVID-19 in Hong Kong, China." *Advances in Continuous and Discrete Models* 2024 (1): 1–9. https://doi.org/10.1186/s13662-024-03830-7.
- Wills, Adrian G., and Thomas B. Schön. 2021. "Stochastic Quasi-Newton with Line-Search Regularisation." *Automatica* 127: 109503. https://doi.org/10.1016/j.automatica.2021.109503.