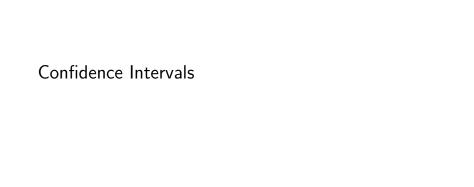
BIFX 553 - Confidence Intervals

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Setup



Confidence Interals

Two results will help us understand the derivation of confidence intervals:

- Central Limit Theorem
- ► Law of Large Numbers

Central Limit Theorem

The mean, \bar{x} , of n independent, identically distributed random variables, X, with well defined expected value, $E(X) = \mu$, and variance, $Var(X) = \sigma$, will be approximately normally distributed when n is sufficiently large:

$$\bar{x} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$

Central Limit Theorem: Simulation

Load the R function found at http://tinyurl.com/zenq9q3.

```
# Normal distribution, sample size of 10
clt.test(rnorm, 10)
# Chi-squared distribution sample size of 5
clt.test(rchisq, 5, df = 3)
# Bimodal mixture of Normals, sample size of 10
rbimodal <- function(n)
 m \leftarrow rbinom(n, 1, 0.5) \%
       as.logical()
  return(ifelse(m, rnorm(n),
                   rnorm(n, 5, 2)))
clt.test(rbimodal, 10)
```

Law of Large Numbers

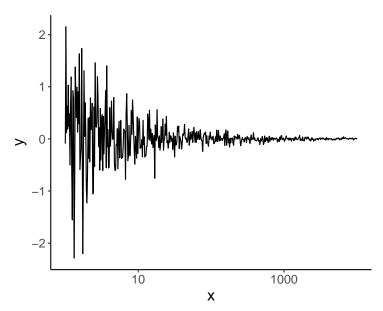
Given our sample mean, \bar{x} , the Law of Large Numbers states that \bar{x} will converge to the true population mean as the sample size increases, assuming the sample, X, are independent, identically distributed random variables.

$$\bar{x} \xrightarrow{n \to \infty} \mu$$

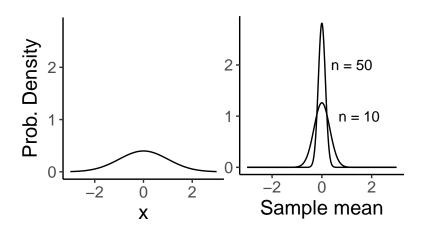
Law of Large Numbers: Simulation

```
set.seed(293874)
lln \leftarrow data_frame(x = 10^seq(from = 0, to = 4,
                                length = 500),
                    y = {map(x, rnorm) \%}
                         map(mean) %>%
                         unlist()})
g \leftarrow ggplot(lln, aes(x, y)) +
     geom line() +
     scale x log10()
```

Law of Large Numbers: Simulation

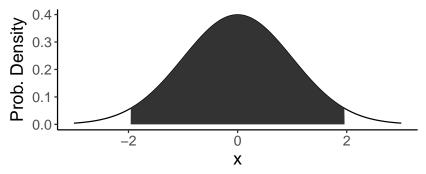


The distribution of \bar{x}



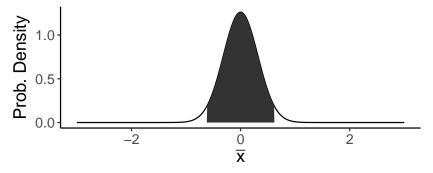
95% Confidence Region

95% of the samples of x we collect will fall in $\mu \pm 1.95\sigma$. This forms the basis of our confidence interval. Side note: the area under this curve over the range $(-\infty,\infty)$, and every probability distribution, is 1.



95% Confidence Interval construction

When the sample size is 10, the distribution of \bar{x} looks like this. 95/% of the time we will expect \bar{x} to fall in the region $\left(\mu \pm \frac{1.95\sigma}{\sqrt{n}}\right)$. From this, we infer that we are 95% confident that the true mean lies within the interval $\left(\bar{x} \pm \frac{1.95*sd}{\sqrt{n}}\right)$.



95% CI Example 1

Lets say that we are studying a population, and we have a sample of 100 blood systolic preasure measurements. The mean is 123 and the standard deviation is 12. What is our confidence interval?

$$ar{x} = 123, sd(x) = 12, n = 100$$

$$95\% \ \mathsf{CI}(\mu) = 123 \pm \frac{1.95 * 12}{\sqrt{100}} = (120.66, 125.34)$$

95% CI with gmodels

Let's simulate a similar data set in R and use the gmodels package to calculate the CI.

```
set.seed(29874)
rnorm(100, 123, 12) %>%
  ci()
```

```
## Estimate CI lower CI upper Std. Error
## 124.023961 121.599383 126.448540 1.221932
```

95% CI of Im object

A more practical use of ci() can be applied to the homework from a few weeks ago.

```
load('.../Data/06_NonLinearVariables.RData')
lm(y ~ x1*x2, data = dat1) %>%
    ci()
```

```
## Estimate CI lower CI upper Std. Error p-value

## (Intercept) 0.08622145 -0.1120740 0.2845169 0.10054832 3.922090e-01

## x1 -0.83619322 -1.0211558 -0.6512306 0.09378771 3.391344e-16

## x2 1.85230919 1.5638783 2.1407401 0.14625265 2.995820e-27

## x1:x2 0.65742039 0.3376731 0.9771677 0.16213205 7.234576e-05
```

95% CI of Im object

Now it is your turn. What is the 95% CI for the x1 variable in the second dataset from a few weeks ago? The model is provided here for your convenience. What about the 90% CI?

```
lm(log(y) ~ x1, data = dat2) %>%
ci(confidence = 0.9)
```

```
## Estimate CI lower CI upper Std. Error
## (Intercept) 0.01326985 -0.105471 0.1320107 0.07185157 8
## x1 1.98735186 1.864883 2.1098211 0.07410763 7
```

Now, suppose we want to know the 95% CI of the expected number of nodes with detectable cancer in a woman with the following measurements:

- ▶ size = 23
- ▶ grade = 2
- ▶ pgr = 32.5
- hormon = "no tamoxifen"

As you may recall, the model we chose for this last week was

```
model <- lm(lnodes ~ size + grade + lpgr + hormon, data = gbsg)
coef(model)</pre>
```

```
## (Intercept) size grade
## 0.43767500 0.01948081 0.13101418
## lpgr hormonno tamoxifen
## -0.02429904 -0.08218227
```

ci() will give us the confidence intervals for each of the betas, but won't get us very far with a specific prediction. We can get the prediction using predict(), but what is the standard error of the prediction?

```
## 1
## 2.7
```

We can use the estimable() function to give us Cl's.

```
estimable(model, cm = c(1, 23, 2, signif(log(32.5), 1), 1), conf.int = 0.95)[c(1,6,7)]
```

```
## Estimate Lower.CI Upper.CI
## (1 23 2 3 1) 0.9926826 0.9028564 1.082509
```

We can also use estimable() to explore more complicated questions. For example, we could ask, is there a difference in the expected number of expected nodes between a woman with a grade 2, 27 mm tumor and a woman with a grade 3, 20 mm tumor?

1)
$$\log E(nodes|size = 27, grade = 2) = \beta_0 + 27\beta_1 + 2\beta_2 + lpgr\beta_3 + hormon\beta_4$$

2) $\log E(nodes|size = 27, grade = 2) = \beta_0 + 20\beta_1 + 3\beta_2 + lpgr\beta_3 + hormon\beta_4$

eqn1 - eqn2 =
$$\beta_0' + 27\beta_1 + 2\beta_2 + Ipgr\beta_3' + hormon\beta_4'$$

- $(\beta_0' + 20\beta_1 + 3\beta_2 + Ipgr\beta_3' + hormon\beta_4')$
= $7\beta_1 - \beta_2$

```
estimable(model, cm = c(size = 7, grade = -1), conf.int = 0.95)
```

```
## Estimate Std. Error t value DF Pr(>|t|) Lower.CI
## (0 7 -1 0 0) 0.005351483 0.06595384 0.08113982 681 0.9353546 -0.1241458
## Upper.CI
## (0 7 -1 0 0) 0.1348488
```