# Homework 3

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## 1 Prove Chernoff Bound (2)(3)

Let  $X_1, X_2, ..., X_n$  be independent Poisson trials such that, for  $1 \leq i \leq n, Pr[X_i = 1] = p_i$ , where  $0 < p_i < 1$ . Then, for  $X = \sum_{i=1}^n X_i, \mu = E[X] = \sum_{i=1}^n p_i$ , and any  $0 < \delta < 1$ ,

$$Pr[X<(1-\delta)\mu]<[rac{e^{-\delta}}{(1-\delta)^{(1-\delta)}}]^{\mu} \ Pr[|X-\mu|>\delta\mu]<2e^{-rac{\delta^2}{3}\mu}$$

#### Prove(2)

For any positive real  $\lambda$ , and applying Markov inequality, we have

$$\begin{split} Pr[X < (1-\delta)\mu] = & Pr[-X > -(1-\delta)\mu] \\ = & Pr[e^{-\lambda X} > e^{-\lambda(1-\delta)\mu}] \\ < & \frac{E[e^{-\lambda X}]}{e^{-\lambda(1-\delta)\mu}} \end{split}$$

Since  $X_i$  are independent, the variable  $e^{-\lambda X_i}$  are also independent.

$$egin{aligned} E[e^{-\lambda X}] \ &= \prod_{1 \leq i \leq n} E[e^{-\lambda X_i}] \ &= \prod_{1 \leq i \leq n} [p_i e^{-\lambda} + (1-p_i)] \ &= \prod_{1 \leq i \leq n} [p_i (e^{-\lambda} - 1) + 1] \ &\leq \prod_{1 \leq i \leq n} [e^{p_i (e^{-\lambda} - 1)}] \ &= e^{\mu(e^{-\lambda} - 1)} \end{aligned}$$

so we have

$$Pr[X < (1-\delta)\mu] < rac{e^{\mu(e^{-\lambda}-1)}}{e^{-\lambda(1-\delta)\mu}}$$

let  $\lambda = -\ln(1-\delta)$ , which is positive, to obtain

$$Pr[X<(1-\delta)\mu]<rac{e^{-\mu\delta}}{(1-\delta)^{(1-\delta)\mu}}$$

This yields the desired result.  $\Box$ 

#### Prove(3)

Using McLaurin expansion for  $\ln(1-\delta)$  and  $\ln(1+\delta)$  to obtain

$$(1-\delta)^{1-\delta} > e^{-\delta + rac{\delta^2}{2}} \ (1+\delta)^{1+\delta} > e^{\delta + rac{\delta^2}{2} - rac{\delta^3}{6}}$$

Combine these two inequalities with chernoff bound (1) and (2)

$$egin{split} & Pr[|X-\mu| > \delta \mu] \ = & Pr[X > (1+\delta)\mu] + Pr[X < (1-\delta)\mu] \ < & e^{-rac{\delta^2}{2}\mu} + e^{-rac{\delta^2}{2}\mu + rac{\delta^3}{6}\mu} \ < & 2e^{-rac{\delta^2}{3}\mu} \end{split}$$

### 2.bounds for Binomial coefficient

for integer x range from 1 to m

$$(\frac{m}{x})^x \le {m \choose x} \le (\frac{em}{x})^x$$

Prove:

$$\binom{m}{x} = \frac{m}{x} \frac{m-1}{x-1} ... \frac{m-x+2}{2} \frac{m-x+1}{1} \ge (\frac{m}{x})^x$$

The other side

$$\binom{m}{x} = \frac{m}{x} \frac{m-1}{x-1} \dots \frac{m-x+2}{2} \frac{m-x+1}{1} \le \frac{m^x}{x!}$$

Applying stirling formula  $x!=\sqrt{2\pi x}(\frac{x}{e})^x$ , since  $m(m-1)...(m-x+1)\leq \frac{m^x}{\sqrt{2\pi x}}$ , we can omit the factor  $\sqrt{2\pi x}$  and obtain

$$\binom{m}{x} \le (\frac{em}{x})^x$$

$$3.Cor(Y_i, Y_j) < 0$$

Prove:

$$Cor(Y_i, Y_i) = E(Y_iY_i) - E(Y_i)E(Y_i)$$

Assume the probability that  $Y_i=1$  is p, that is  $E(Y_i)=E(Y_j)=p$ ,

$$E(Y_iY_j) = E(Y_i|Y_j)E(Y_j) = E(Y_i|Y_j)p$$

We can focus on  $E(Y_i|Y_j)$ , which means the expectation of  $Y_i$  with the condition that more than cx balls fall into another bin.

$$E(Y_i|Y_j) < \sum_{i=cr}^{m-cx} \binom{m-cx}{i} (\frac{1}{n})^i (\frac{n-1}{n})^{m-cx-i} < \sum_{i=cr}^{m} \binom{m}{i} (\frac{1}{n})^i (\frac{n-1}{n})^{m-i} = E(Y_j)$$

It yields the desired result that  $Cor(Y_i,Y_j)<0$