

Homework 3

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1 Prove Chernoff Bound (2)(3)

Let X_1, X_2, \dots, X_n be independent Poisson trials such that, for $1 \leq i \leq n$, $Pr[X_i = 1] = p_i$, where $0 < p_i < 1$. Then, for $X = \sum_{i=1}^n X_i$, $\mu = E[X] = \sum_{i=1}^n p_i$, and any $0 < \delta < 1$,

$$Pr[X < (1 - \delta)\mu] < \left[\frac{e^{-\delta}}{(1 - \delta)^{(1 - \delta)}} \right]^\mu$$
$$Pr[|X - \mu| > \delta\mu] < 2e^{-\frac{\delta^2}{3}\mu}$$

Prove(2)

For any positive real λ , and applying Markov inequality, we have

$$\begin{aligned} Pr[X < (1 - \delta)\mu] &= Pr[-X > -(1 - \delta)\mu] \\ &= Pr[e^{-\lambda X} > e^{-\lambda(1 - \delta)\mu}] \\ &< \frac{E[e^{-\lambda X}]}{e^{-\lambda(1 - \delta)\mu}} \end{aligned}$$

Since X_i are independent, the variable $e^{-\lambda X_i}$ are also independent.

$$\begin{aligned} &E[e^{-\lambda X}] \\ &= \prod_{1 \leq i \leq n} E[e^{-\lambda X_i}] \\ &= \prod_{1 \leq i \leq n} [p_i e^{-\lambda} + (1 - p_i)] \\ &= \prod_{1 \leq i \leq n} [p_i(e^{-\lambda} - 1) + 1] \\ &\leq \prod_{1 \leq i \leq n} [e^{p_i(e^{-\lambda} - 1)}] \\ &= e^{\mu(e^{-\lambda} - 1)} \end{aligned}$$

so we have

$$Pr[X < (1 - \delta)\mu] < \frac{e^{\mu(e^{-\lambda}-1)}}{e^{-\lambda(1-\delta)\mu}}$$

let $\lambda = -\ln(1 - \delta)$, which is positive, to obtain

$$Pr[X < (1 - \delta)\mu] < \frac{e^{-\mu\delta}}{(1 - \delta)^{(1-\delta)\mu}}$$

This yields the desired result. \square

Prove(3)

Using McLaurin expansion for $\ln(1 - \delta)$ and $\ln(1 + \delta)$ to obtain

$$\begin{aligned}(1 - \delta)^{1-\delta} &> e^{-\delta + \frac{\delta^2}{2}} \\ (1 + \delta)^{1+\delta} &> e^{\delta + \frac{\delta^2}{2} - \frac{\delta^3}{6}}\end{aligned}$$

Combine these two inequalities with chernoff bound (1) and (2)

$$\begin{aligned}&Pr[|X - \mu| > \delta\mu] \\ &= Pr[X > (1 + \delta)\mu] + Pr[X < (1 - \delta)\mu] \\ &< e^{-\frac{\delta^2}{2}\mu} + e^{-\frac{\delta^2}{2}\mu + \frac{\delta^3}{6}\mu} \\ &< 2e^{-\frac{\delta^2}{3}\mu}\end{aligned}$$

2.bounds for Binomial coefficient

for integer x range from 1 to m

$$\left(\frac{m}{x}\right)^x \leq \binom{m}{x} \leq \left(\frac{em}{x}\right)^x$$

Prove:

$$\binom{m}{x} = \frac{m}{x} \frac{m-1}{x-1} \dots \frac{m-x+2}{2} \frac{m-x+1}{1} \geq \left(\frac{m}{x}\right)^x$$

The other side

$$\binom{m}{x} = \frac{m}{x} \frac{m-1}{x-1} \dots \frac{m-x+2}{2} \frac{m-x+1}{1} \leq \frac{m^x}{x!}$$

Applying stirling formula $x! = \sqrt{2\pi x} \left(\frac{x}{e}\right)^x$ to obtain

$$\binom{m}{x} \leq \frac{1}{\sqrt{2\pi x}} \left(\frac{em}{x}\right)^x < \left(\frac{em}{x}\right)^x$$

$$\mathbf{3.} Cor(Y_i, Y_j) < 0$$

Prove:

$$Cor(Y_i, Y_j) = E(Y_i Y_j) - E(Y_i)E(Y_j)$$

Assume the probability that $Y_i = 1$ is p , that is $E(Y_i) = E(Y_j) = p$,

$$E(Y_i Y_j) = E(Y_i | Y_j) E(Y_j) = E(Y_i | Y_j) p$$

We can focus on $E(Y_i | Y_j)$, which means the expectation of Y_i with the condition that more than cx balls fall into another bin.

$$E(Y_i | Y_j) < \sum_{i=cx}^{m-cx} \binom{m-cx}{i} \left(\frac{1}{n}\right)^i \left(\frac{n-1}{n}\right)^{m-cx-i} < \sum_{i=cx}^m \binom{m}{i} \left(\frac{1}{n}\right)^i \left(\frac{n-1}{n}\right)^{m-i} = E(Y_j)$$

It yields the desired result that $Cor(Y_i, Y_j) < 0$