Master project

Halvard

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```
#Libraries
library(ggplot2) #plotting
## Warning: package 'ggplot2' was built under R version 4.3.3
library(tidyr) #for pivot_longer?
## Warning: package 'tidyr' was built under R version 4.3.3
library(ggpubr)
library(INLA) #posterior inference
## Warning: package 'INLA' was built under R version 4.3.3
## Loading required package: Matrix
## Attaching package: 'Matrix'
## The following objects are masked from 'package:tidyr':
##
##
       expand, pack, unpack
## Loading required package: sp
## Warning: package 'sp' was built under R version 4.3.3
## This is INLA_24.05.01-1 built 2024-05-01 18:49:50 UTC.
## - See www.r-inla.org/contact-us for how to get help.
## - List available models/likelihoods/etc with inla.list.models()
## - Use inla.doc(<NAME>) to access documentation
library(MASS) #for ginv
```

Some theory

Random walk 1

The easiest way to simulate a random walk is through the assumption of independent normal distributed increments. We know that '

$$x_{t+1}|x_1,...,x_t,\sigma = x_{t+1}|x_t,\sigma \sim N(x_t,\sigma^2)$$

Thus, all we need to do is sample from the standard normal, scale them by σ and add them sequentially. Standard to assert that $x_0 = 0$.

Defining a basic plotting function that will come in handy

Now, lets define functionality for the random walk.

```
RW1 <- function(sigma, N){
  # sigma^2 is the variance for the transitions and N is the number of points
  x \leftarrow rep(0, N)
  z <- sigma * rnorm(N-1)
  for(j in 2:N){
    x[j] \leftarrow x[j-1] + z[j-1]
  return(x)
}
#Also making a normalized RW1, ie. it sums to zero
Norm_RW1 <- function(sigma, N){</pre>
  # sigma^2 is the variance for the transitions and N is the number of points
  x \leftarrow rep(0, N)
  z <- sigma * rnorm(N-1)
  for(j in 2:N){
    x[j] \leftarrow x[j-1] + z[j-1]
  return(x - mean(x)) #makes the mean zero
#Parameters for simulation of RW1
n <- 10
N <- 100
sigma <- 1
```

```
df <- data.frame(matrix(NA, nrow = N, ncol = n))
set.seed(0)
for (i in 1:n){
    df[, i] <- RW1(sigma, N)
}
RW1_plot <- plot_realizations(df, "RW1 with N=100", legend = FALSE)</pre>
```

Random walk 2

$$x_{t} - 2x_{t+1} + x_{t+2} \sim N(0, \sigma^{2})$$

$$x_{t+2} \sim N(2x_{t+1} - x_{t}, \sigma^{2})$$

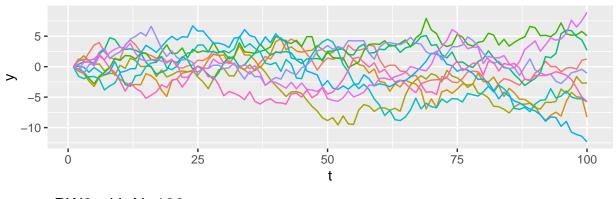
$$x_{t+2} = 2x_{t+1} - x_{t} + \epsilon_{t}, \quad \epsilon_{t} \sim N(0, \sigma^{2})$$

```
RW2 <- function(sigma, N){
  # sigma^2 is the variance for the transitions and N is the number of points
 x \leftarrow rep(0, N)
 z <- sigma * rnorm(N-1)
 for(j in 3:N){
    x[j] \leftarrow 2*x[j-1] - x[j-2] + z[j-1]
  return(x)
}
#Parameters for simulation of RW2
n <- 10
N <- 100
sigma <- 1
df2 <- data.frame(matrix(NA, nrow = N, ncol = n))</pre>
set.seed(0)
for (i in 1:n){
 df2[, i] <- RW2(sigma, N)
# Plot all lines using ggplot
RW2_plot <- plot_realizations(df2, "RW2 with N=100", legend = FALSE)
```

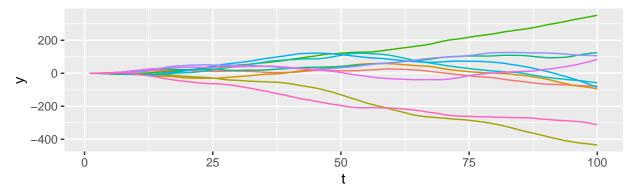
Visualize the plots.

```
RW_figure <- ggarrange(RW1_plot, RW2_plot, ncol = 1)
RW_figure</pre>
```





RW2 with N=100



Simulation study - Gaussian data with or without an offset

We want to conduct a small simulation study to see if the adaptive models improve the standard models in situations with shocks. First we will start with assessing the performance on non-shocked data.

Simulation of non-shocked Gaussian data

We will simulate data with a latent temporal structured random effect as a RW1, denoted \mathbf{x} . The total Bayesian hierarchical model can be described as

$$y_t | \eta_t \sim N(\eta_t, \sigma_t^2)$$

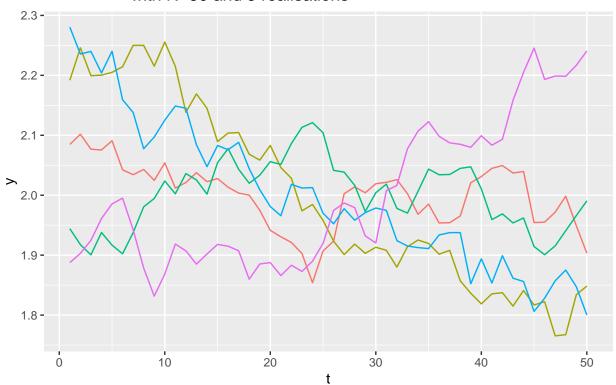
 $\eta_t = \mu + x_t.$

For the moment we assume constant σ_t for all timepoints, and choose some fixed σ_r for the random walk. First, lets make the general functions.

```
#function to simulate a realization y
sim_non_shocked_gaussian_data <- function(N, mu, sigma_obs, sigma_rw){
    #N timepoints, mean mu, and standard deviations observations and the RW1
    eta <- mu + Norm_RW1(sigma_rw, N)
    y <- sapply(eta, function(r) rnorm(1, mean = r, sd = sigma_obs))
    return(y)
}</pre>
```

```
sim_non_shocked_gaussian_dataframe <- function(N, mu, sigma_obs, sigma_rw, n, seed = 50){</pre>
  set.seed(seed)
  df <- data.frame(matrix(NA, nrow = N, ncol = n))</pre>
  for(i in 1:n){
    df[, i] <- sim_non_shocked_gaussian_data(N, mu, sigma_obs, sigma_rw)</pre>
  return(df)
}
#The dataframe for all non-shocked Gaussian data
N <- 50
n <- 50 #was 100
sigma_obs <- 0.001
sigma_rw <- 0.03
mu <- 2
NSG_dataframe <- sim_non_shocked_gaussian_dataframe(N, mu, sigma_obs, sigma_rw, n)
NSG_dataframe$t <- 1:N #needed for random effects later
NSG_dataframe$us <- 1:N #needed for random effects later
#Visualizing some simulated data
plot_realizations(NSG_dataframe[, 1:5], "Simulated non-shocked Gaussian data
                  with N=50 and 5 realisations", legend = FALSE)
```

Simulated non-shocked Gaussian data with N=50 and 5 realisations



Brief testing with INLA on some simulated data

We will fit the simple model in INLA with a latent layer as

$$\eta_t = \mu + x_t$$

where x_t is a RW1 with some precision τ with a default prior. We use a Gaussian likelihood in the observation layer, again with a default prior for the precision. Same for μ .

```
#Data preperation
NSG_data <- data.frame(matrix(c(NSG_dataframe[, 1], 1:N, 1:N), nrow = N, ncol = 3))
colnames(NSG_data) <- c("y", "time", "us") #makes the colnames match the formula

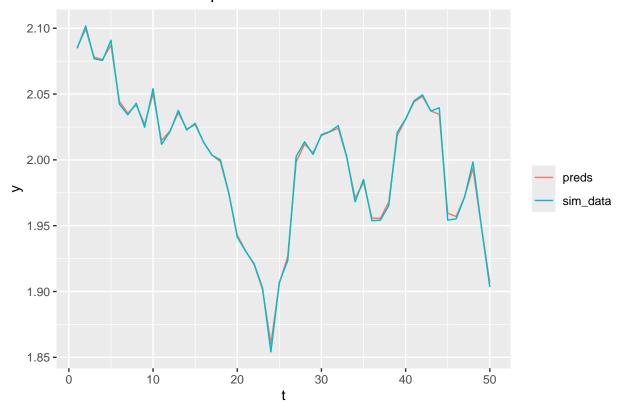
#The INLA model
formula <- y ~ f(time, model = "rw1") + f(us, model = "iid") #intercept is included automatically
res <- inla(formula, family = "gaussian", data = NSG_data)

## Warning in .recacheSubclasses(def@className, def, env): undefined subclass
## "ndiMatrix" of class "replValueSp"; definition not updated

#For plotting the data and the predicted values
plot_df <- data.frame(matrix(c(NSG_dataframe[, 1], res$summary.fitted.values$mean), ncol = 2))
colnames(plot_df) <- c("sim_data", "preds") #for legends in the plot

plot_realizations(plot_df, "Simulated data and predicted values for non-shocked data")</pre>
```

Simulated data and predicted values for non-shocked data

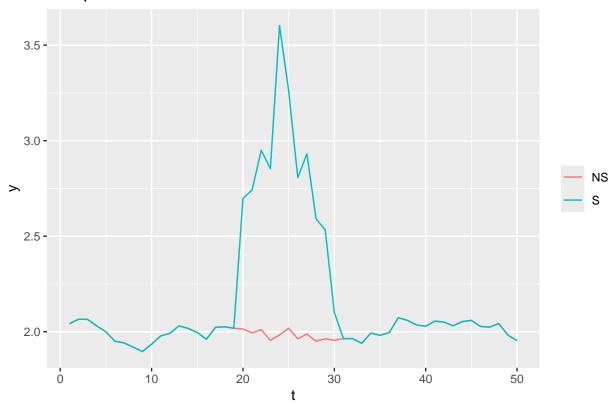


We see that the model predictions align well with the data it is fit on.

Simulation of shocked Gaussian data

We now want to simulate Gaussian data where we know that there are shocks on certain timepoints. This could be modeled by adding or subtracting a slightly randomized value from the chosen points. Lets say we want a shock from t = 20 to t = 30, which could be done by adding a $s_t \stackrel{iid}{\sim} N(0.7, 0.3)$ for instance.

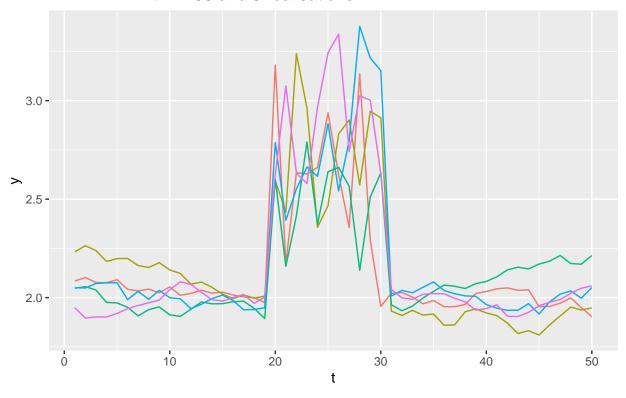
Comparison of shocked and non-shocked simulated data



Lets make some general functions for simulating shocked Gaussian data.

```
#function to simulate a shocked realization y
sim_shocked_gaussian_data <- function(N, mu, sigma_obs, sigma_rw, t_start = 20,</pre>
                              t_end = 30, mu_offset = 0.7, sigma_offset = 0.3){
  #N timepoints, mean mu, and standard deviations for observations and the RW1
  #t_start and t_end bound the offset area with specified mean and sd
  eta <- mu + Norm_RW1(sigma_rw, N)
 y <- sapply(eta, function(r) rnorm(1, mean = r, sd = sigma_obs))
  offset <- c(rep(0, t_start - 1),rnorm(t_end - t_start + 1, mu_offset,
                                         sigma_offset), rep(0, N - t_end))
 y_offset <- y + offset</pre>
 return(y_offset)
sim_shocked_gaussian_dataframe <- function(N, mu, sigma_obs, sigma_rw,</pre>
 t_start = 20, t_end = 30, mu_offset = 0.7, sigma_offset = 0.3, n, seed = 50){
 set.seed(seed)
 df <- data.frame(matrix(NA, nrow = N, ncol = n))</pre>
 for(i in 1:n){
    df[, i] <- sim_shocked_gaussian_data(N, mu, sigma_obs, sigma_rw, t_start,</pre>
                                          t_end, mu_offset, sigma_offset)
 }
 return(df)
}
#Shocked Gaussian dataframe
SG_dataframe <- sim_shocked_gaussian_dataframe(N, mu, sigma_obs, sigma_rw, n=n)
SG_dataframe$t <- 1:N #needed for random effects later
SG_dataframe$us <- 1:N #needed for random effects later
\#Visualizing\ some\ simulated\ data
plot_realizations(SG_dataframe[, 1:5], "Simulated shocked Gaussian data
                  with N=50 and 5 realisations", legend = FALSE)
```

Simulated shocked Gaussian data with N=50 and 5 realisations



Implementing the adaptive random walk in INLA

As this is somewhat complicated and new to me, I will start by a slightly easier example, namely the RW1. Can then also compare it to the already defined RW1 in INLA to ensure it works as intended. First some basic theory on defining random effects in INLA from https://becarioprecario.bitbucket.io/inla-gitbook/chnewmodels.html .

Defining new latent random effects in INLA

New latent effects must be specified as GMRFs. So we need μ , Q and θ and its log-priors and initial values. Also need a graph, which I think can just be Q as well, and some log-normalizing constant. As θ is parametrized as $\theta_1 = \log(\tau)$ and $\theta_2 = \operatorname{logit}(\rho)$ where τ is the precision and ρ is the spatial dependence, which I think we define to be 1. The general structure of the inla-regeneric is shown below:

```
inla.rgeneric.somemodel = function(
   cmd = c("graph", "Q", "mu", "initial", "log.norm.const","log.prior", "quit"),
   theta = NULL)
{
    # for reference and potential storage for objects to
    # cache, this is the environment of this function
    # which holds arguments passed as `...` in
    # `inla.rgeneric.define()`.
    envir = parent.env(environment())
    graph = function(){ <to be completed> }
```

```
Q = function() { <to be completed> }
  mu = function() { <to be completed> }
  log.norm.const = function() { <to be completed> }
  log.prior = function() { <to be completed> }
  initial = function() { <to be completed> }
  quit = function() { <to be completed> }
  # sometimes this is useful, as argument 'graph' and 'quit'
  # will pass theta=numeric(0) (or NULL in R-3.6...) as
  # the values of theta are NOT
  # required for defining the graph. however, this statement
  # will ensure that theta is always defined.
  if (!length(theta)) theta = initial()
  val = do.call(match.arg(cmd), args = list())
  return (val)
#if W is a needed argument
somemodel.model <- inla.rgeneric.define(inla.rgeneric.somemodel, W = W)</pre>
```

Implementing RW1 in INLA

In a RW1 we only have one hyperparameter, namely τ . So we get $\theta = \log(\tau)$ and the precision matrix is defined previously. Lets first define a function for the geometric variance, defined in the overleaf document, to scale the precision matrix. Use the function ginv from the library MASS to calculate the generalized inverse.

```
geometric_variance <- function(R) {
    #Input: R is a square structure matrix, often sparse
    N <- dim(R)[1]
    GV <- exp(1 / N * sum(log(diag(ginv(R)))))
    return(GV)
}</pre>
```

Then, lets define the inla.rgeneric function with all its necessary subfunctions.

```
inla.rgeneric.RW1.model = function(
   cmd = c("graph", "Q", "mu", "initial", "log.norm.const","log.prior", "quit"),
   theta = NULL)
{
   #Input:
   #N is the number of timepoints
   #R_star is the scaled structure matrix

   envir = parent.env(environment())

   interpret_theta <- function() { return(list(tau = exp(theta[1L])))}

   graph <- function() {return(Q())}

   Q <- function() {</pre>
```

```
p <- interpret_theta()</pre>
  Q <- p$tau * R_star
  return(inla.as.sparse(Q))
mu <- function() {return(numeric(0))}</pre>
initial <- function() {return(4)} #default for precisions: initial = 4
log.norm.const <- function() {return(numeric(0))} #Inla computes it</pre>
log.prior <- function() {#default: shape = 1, rate = 0.00005 for tau</pre>
  p <- interpret_theta()</pre>
  prior <- dgamma(p$tau, shape = 1, rate = 0.00005, log = TRUE) + theta[1L]
  return(prior)
}
quit <- function() {return(invisible())}</pre>
#to ensure theta is defined
if (!length(theta)) theta = initial()
vals <- do.call(match.arg(cmd), args = list())</pre>
return(vals)
```

We need to pass the function above the number of timepoints N and the scaled structure matrix R^* . The scaled structure matrix for a RW1 is defined by the function below, followed by some testing.

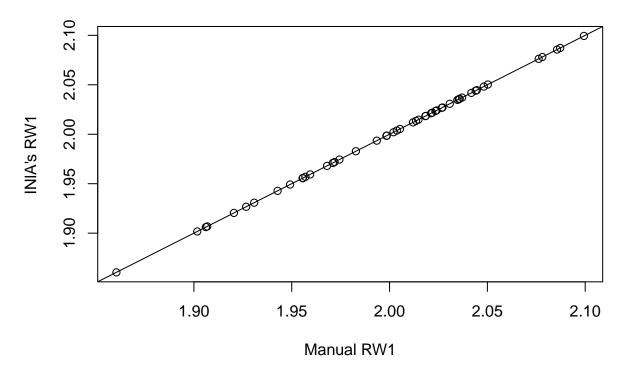
```
Q <- function(N) {
    # Input: N timepoints
    R <- toeplitz(c(2, -1, rep(0, N - 2)))# 2 on diag and -1 on firstdiags
    R[1, 1] <- R[N, N] <- 1 # 1 for first and last diag element
    gv <- geometric_variance(R)
    R_star <- gv * R
    return(R_star) #returns the scaled structure matrix for a RW1
}

N <- 50 #is defined further up as well
R_star <- Q(N)
RW1_model <- inla.rgeneric.define(inla.rgeneric.RW1.model, N = N, R_star = R_star)</pre>
```

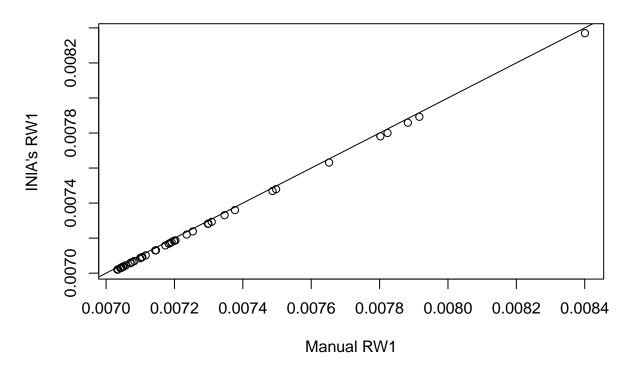
The RW1_model above is now a custom latent effect which can be included in INLA formulas to define models. However, it is only defined for N=50, and we need to define separate ones for other N.Now, lets check if it works as intended on some data from earlier. Then we need to define the INLA formula where we add a constraint to enforce a sum to zero constraint as the model includes an intercept.

```
#The standard RW1 model from INLA
formula_I <- y ~ f(time, model = "rw1") + f(us, model = "iid")</pre>
res_I <- inla(formula_I, family = "gaussian", data = NSG_data)</pre>
summary(res M)
## Time used:
       Pre = 0.837, Running = 5.03, Post = 0.206, Total = 6.07
## Fixed effects:
##
                       sd 0.025quant 0.5quant 0.975quant mode kld
               mean
## (Intercept)
                  2 0.002
                               1.997
                                            2
                                                    2.003
##
## Random effects:
##
   Name
              Model
##
      time RGeneric2
##
      us IID model
##
## Model hyperparameters:
                                                           sd 0.025quant 0.5quant
                                               mean
                                                                 4848.70 24023.35
## Precision for the Gaussian observations 32978.49 3.10e+04
## Theta1 for time
                                               5.25 2.46e-01
                                                                    4.76
                                                                             5.25
## Precision for us
                                            32957.86 3.10e+04
                                                                 4832.88 24008.10
                                           0.975quant
                                                           mode
## Precision for the Gaussian observations
                                              1.15e+05 12522.76
## Theta1 for time
                                             5.73e+00
                                                           5.26
## Precision for us
                                             1.15e+05 12496.63
## Marginal log-Likelihood: 97.52
## is computed
## Posterior summaries for the linear predictor and the fitted values are computed
## (Posterior marginals needs also 'control.compute=list(return.marginals.predictor=TRUE)')
summary(res_I)
## Time used:
##
       Pre = 0.534, Running = 0.745, Post = 0.397, Total = 1.68
## Fixed effects:
                       sd 0.025quant 0.5quant 0.975quant mode kld
               mean
                  2 0.002
## (Intercept)
                               1.997
                                            2
                                                    2.003
##
## Random effects:
##
    Name
              Model
##
      time RW1 model
##
     us IID model
##
## Model hyperparameters:
                                                           sd 0.025quant 0.5quant
                                               mean
## Precision for the Gaussian observations 33015.23 30933.78
                                                                 4872.66 24092.50
## Precision for time
                                                                 883.37 1438.74
                                            1480.29
                                                       366.52
## Precision for us
                                            33005.58 30938.40
                                                                 4868.91 24080.41
                                           0.975quant
## Precision for the Gaussian observations 114959.20 12588.04
## Precision for time
                                              2316.96 1361.02
```

Comparison of the mean



Comparison of the standard deviation



From the summaries it is clear they are very similar. However, not exactly equal, as for instance the mean of the Gaussian precision. This is supported by the plots of the means and standard deviations, which show that they are almost the same, but clearly not identical. We also note that the run-time for the manual model is much longer than the one implemented by INLA, so I will use INLA's predefined model from here on out for better computational times.

Implementing the adaptive RW1

We now make the previous model more flexible by allowing for two different precisions in the random walk. The new precision is used for transitions involving conflict years which we define. So, we start by computing the scaled structure matrices R1 star and R2 star which we need as inputs for the adaptive random walk.

```
Scaled_structure_matrices_for_ARW1 <- function(N, conflict_years) {
    #Input:
    #N timepoints
    #conflict_years is a list with the conflict years

R1 <- matrix(0, nrow = N, ncol = N) #should be N = 50, non-conflict
R2 <- matrix(0, nrow = N, ncol = N) #should be N = 50, conflict
for( i in 1:(N - 1)){
    if(i %in% conflict_years | (i + 1) %in% conflict_years) {
        R2[c(i, i+1), c(i, i+1)] <- R2[c(i, i+1), c(i, i+1)] + c(1, -1, -1, 1)
    }
    else {
        R1[c(i, i+1), c(i, i+1)] <- R1[c(i, i+1), c(i, i+1)] + c(1, -1, -1, 1)
    }
}</pre>
```

```
gv <- geometric_variance(R1 + R2) #scaling
 return(list(R1 = R1*gv, R2 = R2*gv))
}
#testing
R_star_list <- Scaled_structure_matrices_for_ARW1(7, c(3, 4, 5))</pre>
R star list$R1
##
                     [,2] [,3] [,4] [,5]
                                                      [,7]
            [,1]
                                             [,6]
## [1,] 1.039708 -1.039708
                            0
                                 0
                                     0 0.000000 0.000000
## [2,] -1.039708 1.039708
                                 0
                                     0 0.000000 0.000000
                            0
## [3,] 0.000000 0.000000 0
                                    0 0.000000 0.000000
## [4,] 0.000000 0.000000 0 0
                                    0 0.000000 0.000000
## [5,] 0.000000 0.000000
                            0 0
                                    0 0.000000 0.000000
                                    0 1.039708 -1.039708
## [6,] 0.000000 0.000000
                            0 0
## [7.] 0.000000 0.000000
                                    0 -1.039708 1.039708
R <- R_star_list$R1 + R_star_list$R2</pre>
##
            [,1]
                     [,2]
                               [,3]
                                        [,4]
                                                 [,5]
                                                           [,6]
                                                                    [,7]
## [2,] -1.039708 2.079416 -1.039708 0.000000 0.000000 0.000000 0.000000
## [3,] 0.000000 -1.039708 2.079416 -1.039708 0.000000 0.000000 0.000000
## [4,] 0.000000 0.000000 -1.039708 2.079416 -1.039708 0.000000 0.000000
## [5,] 0.000000 0.000000 0.000000 -1.039708 2.079416 -1.039708 0.000000
       0.000000 0.000000 0.000000 0.000000 -1.039708 2.079416 -1.039708
## [6,]
## [7,] 0.000000 0.000000 0.000000 0.000000 -1.039708 1.039708
Now, lets define the adaptive RW1.
inla.rgeneric.AdaptiveRW1.model = function(
 cmd = c("graph", "Q", "mu", "initial", "log.norm.const","log.prior", "quit"),
 theta = NULL)
 #Input:
 #N is the number of timepoints
 #R_star_list contains R1_star and R2_star, the scaled structure matrices
 envir = parent.env(environment())
 interpret_theta <- function() { return(list(tau1 = exp(theta[1L]),</pre>
                                           tau2 = exp(theta[2L])))}
 graph <- function() {return(Q())}</pre>
 Q <- function() {</pre>
   p <- interpret_theta()</pre>
   Q <- R_star_list$R1 * p$tau1 + R_star_list$R2 * p$tau2
   return(inla.as.sparse(Q)) #sparse representation
```

}

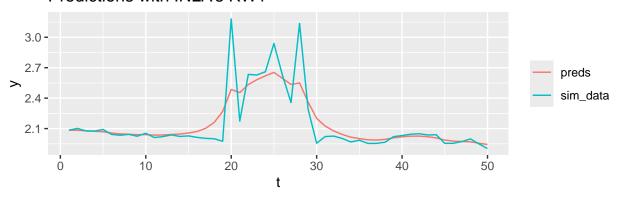
```
mu <- function() {return(numeric(0))}</pre>
  initial <- function() {return(c(4, 4))}#Default initial for precisions
  log.norm.const <- function() {return(numeric(0))}</pre>
  log.prior <- function() {#default: shape = 1, rate = 0.00005</pre>
    p <- interpret theta()</pre>
    prior <- dgamma(p$tau1, shape = 1, rate = 0.005, log = TRUE) + theta[1L]+
              dgamma(p$tau2, shape = 1, rate = 0.005, log = TRUE) + theta[2L]
    return(prior)
  } #the theta terms come from the transformation of variables to the log scale
  quit <- function() {return(invisible())}</pre>
  #to ensure theta is defined
  if (!length(theta)) theta = initial()
 vals <- do.call(match.arg(cmd), args = list())</pre>
 return(vals)
#Computing the scaled R's and defining the ARW1 model
N \leftarrow 50 #is defined further up as well
conf years <- 20:30 #as in the generated data further up
R_star_list <- Scaled_structure_matrices_for_ARW1(N, conf_years)</pre>
ARW1_model <- inla.rgeneric.define(inla.rgeneric.AdaptiveRW1.model,
                                    N = N, R_star_list = R_star_list)
```

Lets do some testing for shocked data.

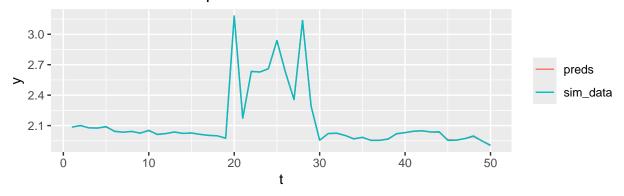
```
#The INLA formula for a latent model with intercept and a RW1
formula_I <- y ~ f(time, model = "rw1") + f(us, model = "iid")</pre>
#The INLA formula for an adaptive RW1
formula_ARW1 <- y ~ f(time, model = ARW1_model,</pre>
                extraconstr = list(A = matrix(1, nrow = 1, ncol = N), e = 0)) +
                f(us, model = "iid")
figure_list <- list()
for( i in 1:5) {
  test_data <- data.frame(matrix(c(SG_dataframe[, i], 1:N, 1:N), nrow = N, ncol = 3))
  colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula</pre>
 res_I <- inla(formula_I, family = "gaussian", data = test_data)</pre>
  plot_df_I <- data.frame(matrix(c(SG_dataframe[, i],</pre>
                                     res_I$summary.fitted.values$mean), ncol = 2))
  colnames(plot_df_I) <- c("sim_data", "preds")</pre>
  plot_I <- plot_realizations(plot_df_I, "Predictions with INLA's RW1")</pre>
  #The adaptive RW1
  res_ARW1 <- inla(formula_ARW1, family = "gaussian", data = test_data)
```

[[1]]

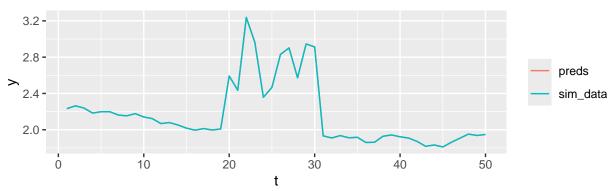
Predictions with INLA's RW1



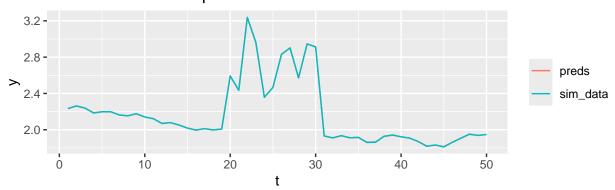
Predictions with adaptive RW1



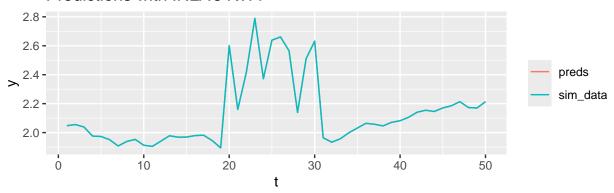
[[2]]



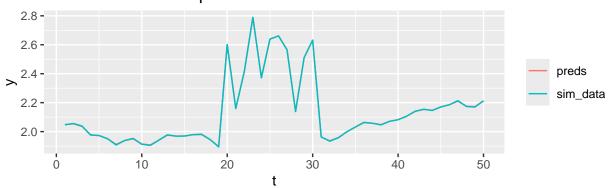
Predictions with adaptive RW1



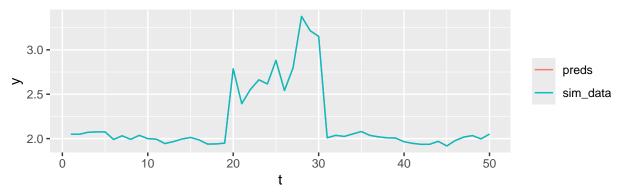
[[3]]



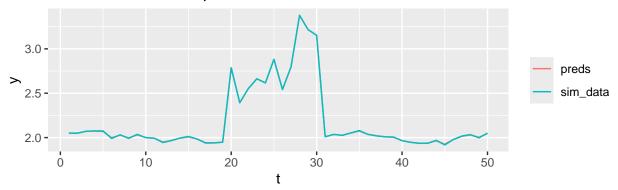
Predictions with adaptive RW1



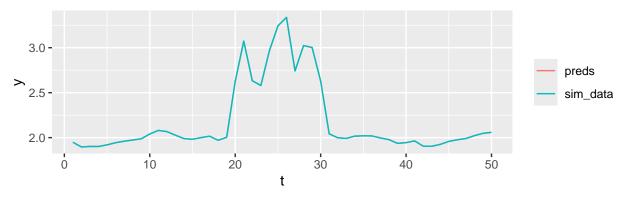
[[4]]



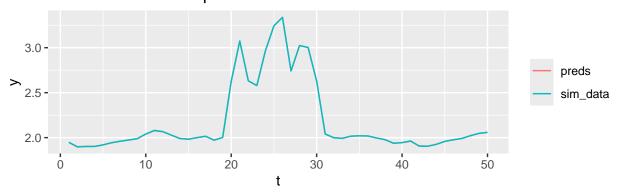
Predictions with adaptive RW1



[[5]]



Predictions with adaptive RW1



We observe that the standard RW1 struggles a lot with the first and third realization, and the rest of the plots seem nice.

Model evaluation

Model criteria

We will evaluate the models with root mean square error (RMSE) and average proper logarithmic scoring (LS).

Root mean square eroor A common model criteria is the RMSE. We define a function to calculate this below. The lower RMSE the better.

```
RMSE <- function(data, preds){
  return(sqrt( sum((data - preds)**2) / length(data))) #definition of RMSE
}</pre>
```

Average proper logarithmic scoring rule From a paper by Gneiting and Raftery (2007).

$$LS(p, \omega) = log p(\omega)$$

where p is the predicted distribution of a point, which we get from INLA, and ω is the observed value, which is the specific datapoint. We then take the average of the score for all the data points. The higher average proper LS the better.

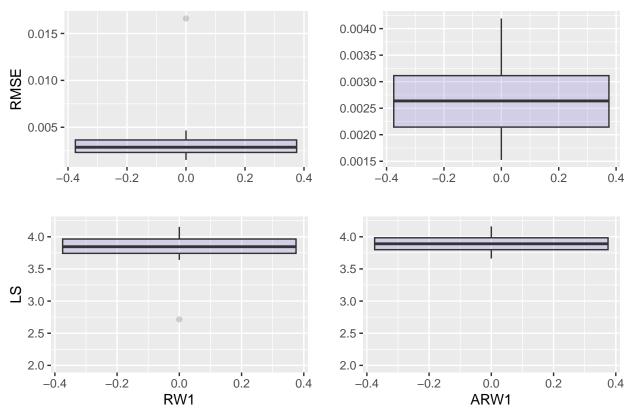
```
average_proper_LS <- function(res, data){
    #res is an inla object from calling a model on data, a vector of datapoints
    mean <- res$summary.fitted.values$mean
    sd <- res$summary.fitted.values$sd
    p <- dnorm(data, mean = mean, sd = sd)
    return(mean(log(p)))
}</pre>
```

Model evaluation for non-shocked data

Lets do some rigorous testing for the entire dataframes form earlier and evaluate them by the chosen model criterias, RMSE and LS. First lets compare the models for non-shocked Gaussian data.

```
Model eval NSG <- data.frame(matrix(NA, nrow = n, ncol = 4))
colnames(Model_eval_NSG) <- c("RMSE_RW1", "LS_RW1", "RMSE_ARW1", "LS_ARW1")</pre>
formula_RW1 <- y ~ f(time, model = "rw1") + f(us, model = "iid")</pre>
for(i in 1:n){#iterate over each simulated realization
  test_data <- NSG_dataframe[, c(i, n + 1, n + 2)] #qets the i-th realization and time and us
  colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula</pre>
  res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)</pre>
  LS_RW1 <- average_proper_LS(res_RW1, NSG_dataframe[, i])
  RMSE_RW1 <- RMSE(NSG_dataframe[, i], res_RW1$summary.fitted.values$mean )
  res_ARW1 <- inla(formula_ARW1, family = "gaussian", data = test_data)
  LS_ARW1 <- average_proper_LS(res_ARW1, NSG_dataframe[, i])
  RMSE_ARW1 <- RMSE(NSG_dataframe[, i], res_ARW1$summary.fitted.values$mean )
 Model eval NSG[i, ] <- c(RMSE RW1, LS RW1, RMSE ARW1, LS ARW1)
}
myboxplot <- function(data, Y, xlabel, ylabel, ymin = 0, ymax = 5) {</pre>
 df <- data.frame(d = data[,1])</pre>
  BPlot <- ggplot(df, aes(y = d)) +</pre>
    geom_boxplot(fill = "slateblue", alpha = 0.2) +
    labs( x = xlabel, y = ylabel)
  if(Y){}
    BPlot <- BPlot + ylim(ymin, ymax)}</pre>
  return(BPlot)
}
#Plotting
plot_RMSE_RW1_NS <- myboxplot(Model_eval_NSG["RMSE_RW1"], 0, "", "RMSE")</pre>
plot_LS_RW1_NS <- myboxplot(Model_eval_NSG["LS_RW1"], 1, "RW1", "LS", 2, 4.2)
plot_RMSE_ARW1_NS <- myboxplot(Model_eval_NSG["RMSE_ARW1"], 0, "", "")</pre>
plot LS ARW1 NS <- myboxplot(Model eval NSG["LS ARW1"], 1, "ARW1", "", 2, 4.2)
plot_eval_NS <- ggarrange(plot_RMSE_RW1_NS, plot_RMSE_ARW1_NS,</pre>
                           plot_LS_RW1_NS, plot_LS_ARW1_NS, ncol = 2, nrow = 2)
annotate_figure(plot_eval_NS, top = text_grob
                ("Model evaluation for non-shocked Gaussian data"))
```

Model evaluation for non-shocked Gaussian data



#can add more parameters like size and color...

From the box-plots it seems like the RW1 generally obtains a smaller RMSE while the ARW1 gets a slightly larger LS. In other words, the RW1 is best according to RMSE and ARW1 is best according to LS. However, they are rather close in both cases.

Model evaluation for shocked data

```
Model_eval_SG <- data.frame(matrix(NA, nrow = n, ncol = 4))
colnames(Model_eval_SG) <- c("RMSE_RW1", "LS_RW1", "RMSE_ARW1", "LS_ARW1")

for(i in 1:n){#iterate over each simulated realization
   test_data <- SG_dataframe[, c(i, n + 1, n + 2)] #gets the i-th realization and time and us
   colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula

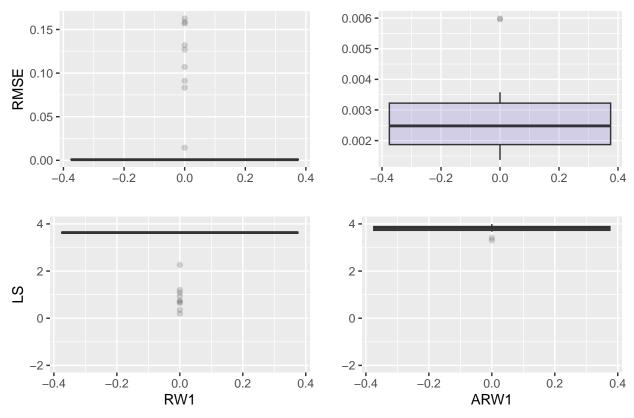
res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)
   LS_RW1 <- average_proper_LS(res_RW1, SG_dataframe[, i])
   RMSE_RW1 <- RMSE(SG_dataframe[, i], res_RW1$summary.fitted.values$mean )

res_ARW1 <- inla(formula_ARW1, family = "gaussian", data = test_data)
   LS_ARW1 <- average_proper_LS(res_ARW1, SG_dataframe[, i])
   RMSE_ARW1 <- RMSE(SG_dataframe[, i], res_ARW1$summary.fitted.values$mean )

Model_eval_SG[i, ] <- c(RMSE_RW1, LS_RW1, RMSE_ARW1, LS_ARW1)</pre>
```

```
}
#check behavior for an outlier realization
\#test\_data \leftarrow SG\_dataframe[, c(72, n + 1)] \#gets the i-th realization and time
\#colnames(test\_data) \leftarrow c("y", "time")
#res_ARW1 <- inla(formula_ARW1, family = "gaussian", data = test_data)</pre>
#plot_realizations(data.frame(res_ARW1$summary.fitted.values$mean, #SG_dataframe[, 72]), "")
#summary(res_ARW1)
#res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)</pre>
#plot_realizations(data.frame(res_RW1$summary.fitted.values$mean, SG_dataframe[, 72]), "")
#Plotting
plot_RMSE_RW1_S <- myboxplot(Model_eval_SG["RMSE_RW1"], 0, "", "RMSE")</pre>
plot_LS_RW1_S <- myboxplot(Model_eval_SG["LS_RW1"], 1, "RW1", "LS", -2, 4)</pre>
plot_RMSE_ARW1_S <- myboxplot(Model_eval_SG["RMSE_ARW1"], 0, "", "")</pre>
plot_LS_ARW1_S <- myboxplot(Model_eval_SG["LS_ARW1"], 1, "ARW1", "", -2, 4)
plot_eval_S <- ggarrange(plot_RMSE_RW1_S, plot_RMSE_ARW1_S,</pre>
                           plot_LS_RW1_S, plot_LS_ARW1_S, ncol = 2, nrow = 2)
## Warning: Removed 9 rows containing non-finite outside the scale range
## ('stat_boxplot()').
annotate_figure(plot_eval_S, top = text_grob
                 ("Model evaluation for shocked Gaussian data"))
```

Model evaluation for shocked Gaussian data



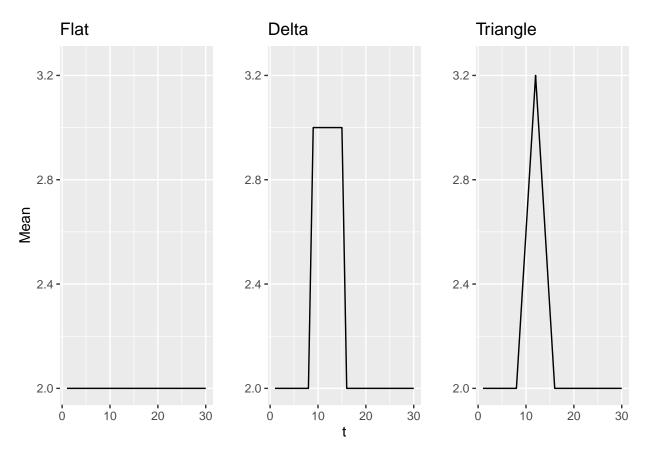
#can add more parameters like size and color...

For the shocked data we see that ARW1 clearly outperforms the RW1 according to both RMSE and LS. Nevermind, not sure about LS, the plots do not have the same axis.

Recreating the simulations study by Wakefield and Aleshin-Guendel

Their simulation study compares a standard RW1 with an adaptive RW1 with an extra precision parameter for some timepoints where we expect a shock. They look at three different means in the latent layer, namely flat, delta and triangle. The latent layer is $\eta_i = \mu_i + b_i$ where $b_i \sim N(0, \tau_i^{-1})$ for some precision τ_i . The observation layer is then $y_i | \eta_i \sim N(\eta_i, V)$ for some variance V. They choose V to be 1/75, 1/150 or 1/300. τ_i is either 20 for all points or 10 for the conflict points, which are 9-15 out of the N=30 timepoints. They then combine and evaluate the model in all these cases. Lets start my creating and plotting the means in the latent layers.

```
#parameters N \leftarrow 30 #number of timepoints n \leftarrow 20 #number of simulations in each case conf_years \leftarrow c(9, 10, 11, 12, 13, 14, 15) #the conflict years #defining the means mean_flat \leftarrow rep(2, N)
```



The next step is to make functions for simulating the data, which is done below.

```
set.seed(seed)
df <- data.frame(matrix(0, nrow = length(mu), ncol = n))
for (i in 1:n) {
   df[, i] <- sim_data_Wakefield(mu, V, tau1 = tau1, tau2 = tau2)
}
df$t <- 1:length(mu)
df$us <- 1:length(mu)
return(df)
}</pre>
```

Now we simulate the data for the eighteen different cases with n=20 realizations of each case.

```
#for constant tau
# flat mean with varying V
data_w_flat_75_NS <- sim_dataframe_Wakefield(n, mean_flat, 1/75)
data_w_flat_150_NS <- sim_dataframe_Wakefield(n, mean_flat, 1/150)
data_w_flat_300_NS <- sim_dataframe_Wakefield(n, mean_flat, 1/300)
#delta mean with varying V
data_w_delta_75_NS <- sim_dataframe_Wakefield(n, mean_delta, 1/75)
data w delta 150 NS <- sim dataframe Wakefield(n, mean delta, 1/150)
data w delta 300 NS <- sim dataframe Wakefield(n, mean delta, 1/300)
#triangle mean with varying V
data_w_triangle_75_NS <- sim_dataframe_Wakefield(n, mean_triangle, 1/75)
data_w_triangle_150_NS <- sim_dataframe_Wakefield(n, mean_triangle, 1/150)
data w triangle 300 NS <- sim dataframe Wakefield(n, mean triangle, 1/300)
#for varying tau
# flat mean with varying V
data_w_flat_75_S <- sim_dataframe_Wakefield(n, mean_flat, 1/75, tau2 = 10)
data_w_flat_150_S <- sim_dataframe_Wakefield(n, mean_flat, 1/150, tau2 = 10)
data_w_flat_300_S <- sim_dataframe_Wakefield(n, mean_flat, 1/300, tau2 = 10)
#delta mean with varying V
data_w_delta_75_S <- sim_dataframe_Wakefield(n, mean_delta, 1/75, tau2 = 10)
data_w_delta_150_S <- sim_dataframe_Wakefield(n, mean_delta, 1/150, tau2 = 10)
data w delta 300 S <- sim dataframe Wakefield(n, mean delta, 1/300, tau2 = 10)
#triangle mean with varying V
data_w_triangle_75_S <- sim_dataframe_Wakefield(n, mean_triangle, 1/75, tau2 = 10)
data_w_triangle_150_S <- sim_dataframe_Wakefield(n, mean_triangle, 1/150, tau2 = 10)
data_w_triangle_300_S <- sim_dataframe_Wakefield(n, mean_triangle, 1/300, tau2 = 10)
```

As we now use N=30 timepoints, we also have to define the ARW1 model with the corresponding scaled structure matrix.

```
#Parameters also defined further up
N <- 30
conf_years <- c(9, 10, 11, 12, 13, 14, 15)
#R_star_list_W <- Scaled_structure_matrices_for_ARW1(N, conf_years)</pre>
```

Model evaluation

```
myboxplot2 <- function(df, mean, V){</pre>
 colnames(df) <- c("RW1", "ARW1")</pre>
  # Reshape data to long format
 df long <- df %>%
    pivot_longer(cols = everything(), names_to = "Category", values_to = "Value")
  # Create the boxplot
  eval_plot <- ggplot(df_long, aes(x = Category, y = Value, fill = Category)) +
    geom_boxplot() +
    ggtitle(paste("Mu: ", mean, ", V: ", V)) + xlab("") + ylab("") +
    theme(legend.position = "none") +
    scale_fill_manual(values = c("RW1" = "skyblue", "ARW1" = "orange"))
 return(eval_plot)
}
#myboxplot2 <- function(data, mean, V) {</pre>
# df \leftarrow data.frame(d = data[,1])
# BPlot \leftarrow ggplot(df, aes(y = d)) +
       geom_boxplot(fill = "slateblue", alpha = 0.2) +
       qqtitle(paste("Mu: ", mean, ", V: ", V))+
       ylab("") + xlab("") +
   # theme(plot.title = element_text(size = 16))
  #return(BPlot)
#}
#testing the boxplot2
#myboxplot2(data w delta 150 NS["X1"], "delta", "1/150")
mod_eval_W <- function(df, mean = "", V = "") {</pre>
 n \leftarrow dim(df)[2] - 2
  eval_df <- data.frame(matrix(NA, nrow = n, ncol = 4))</pre>
  colnames(eval_df) <- c("RMSE_RW1", "LS_RW1", "RMSE_ARW1", "LS_ARW1")</pre>
  for(i in 1:n){#iterate over each simulated realization
    test_data <- df[, c(i, n + 1, n + 2)] #gets the i-th realization + time
    colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula</pre>
    res RW1 <- inla(formula RW1, family = "gaussian", data = test data)
    LS_RW1 <- average_proper_LS(res_RW1, df[, i])
    RMSE_RW1 <- RMSE(df[, i], res_RW1$summary.fitted.values$mean )</pre>
    res_ARW1 <- inla(formula_ARW1_W, family = "gaussian", data = test_data)
    LS_ARW1 <- average_proper_LS(res_ARW1, df[, i])
```

```
RMSE_ARW1 <- RMSE(df[, i], res_ARW1$summary.fitted.values$mean )

eval_df[i, ] <- c(RMSE_RW1, LS_RW1, RMSE_ARW1, LS_ARW1)
}

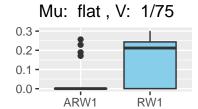
plot_RMSE <- myboxplot2(eval_df[, c("RMSE_RW1", "RMSE_ARW1")], mean, V)

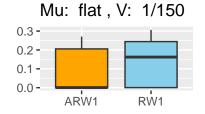
plot_LS <- myboxplot2(eval_df[, c("LS_RW1", "LS_ARW1")], mean, V)

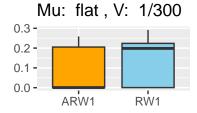
return(list(RMSE = plot_RMSE, LS = plot_LS))
}</pre>
```

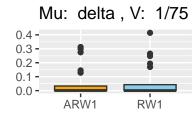
```
#Non-shocked data, ie. constant tau
f_75_NS <- mod_eval_W(data_w_flat_75_NS, "flat", "1/75")
f_150_NS <- mod_eval_W(data_w_flat_150_NS, "flat", "1/150")</pre>
f_300_NS <- mod_eval_W(data_w_flat_300_NS, "flat", "1/300")
d_75_NS <- mod_eval_W(data_w_delta_75_NS, "delta", "1/75")</pre>
d_150_NS <- mod_eval_W(data_w_delta_150_NS, "delta", "1/150")</pre>
d_300_NS <- mod_eval_W(data_w_delta_300_NS, "delta", "1/300")</pre>
t_75_NS <- mod_eval_W(data_w_triangle_75_NS, "triangle", "1/75")
t 150 NS <- mod eval W(data w triangle 150 NS, "triangle", "1/150")
t_300_NS <- mod_eval_W(data_w_triangle_300_NS, "triangle", "1/300")
#making the 3x3 boxplots
plot_eval_NS_W_RMSE <- ggarrange(f_75_NS$RMSE, f_150_NS$RMSE, f_300_NS$RMSE,
                                  d_75_NS$RMSE, d_150_NS$RMSE, d_300_NS$RMSE,
                                  t_75_NS$RMSE, t_150_NS$RMSE, t_300_NS$RMSE,
                                  ncol = 3, nrow = 3)
annotate_figure(plot_eval_NS_W_RMSE,
                top = text_grob("RMSE for non-shocked Wakefield data"))
```

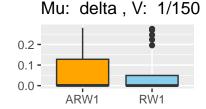
RMSE for non-shocked Wakefield data

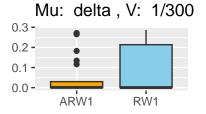


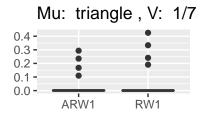


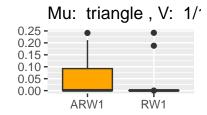


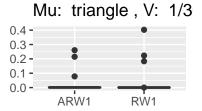




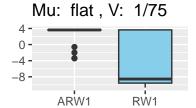


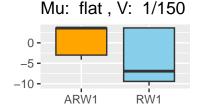


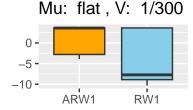




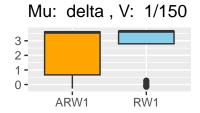
LS for non-shocked Wakefield data

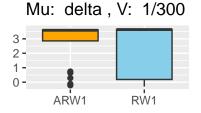


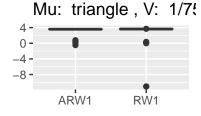


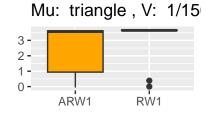


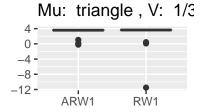
Mu: delta , V: 1/75





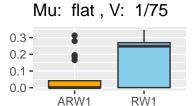


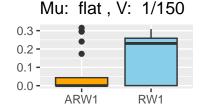


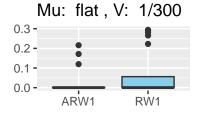


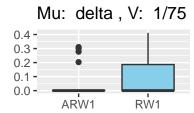
```
#Shocked data, ie. not constant tau
f_75_S <- mod_eval_W(data_w_flat_75_S, "flat", "1/75")</pre>
f_150_S <- mod_eval_W(data_w_flat_150_S, "flat", "1/150")
f_300_S <- mod_eval_W(data_w_flat_300_S, "flat", "1/300")
d 75 S <- mod eval W(data w delta 75 S, "delta", "1/75")
d_150_S <- mod_eval_W(data_w_delta_150_S, "delta", "1/150")</pre>
d_300_S <- mod_eval_W(data_w_delta_300_S, "delta", "1/300")</pre>
t_75_S <- mod_eval_W(data_w_triangle_75_S, "triangle", "1/75")
t 150 S <- mod eval W(data w triangle 150 S, "triangle", "1/150")
t_300_S <- mod_eval_W(data_w_triangle_300_S, "triangle", "1/300")
#Making the 3x3 boxplots
plot_eval_S_W_RMSE <- ggarrange(f_75_S$RMSE, f_150_S$RMSE, f_300_S$RMSE,
                                 d_75_S$RMSE, d_150_S$RMSE, d_300_S$RMSE,
                                 t_75_S$RMSE, t_150_S$RMSE, t_300_S$RMSE,
                                 ncol = 3, nrow = 3)
annotate_figure(plot_eval_S_W_RMSE,
                top = text_grob("RMSE for shocked Wakefield data"))
```

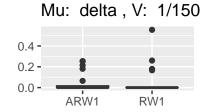
RMSE for shocked Wakefield data

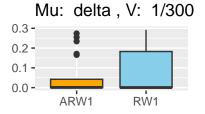


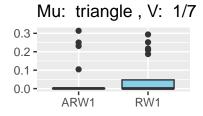


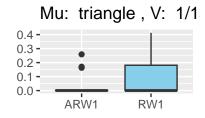


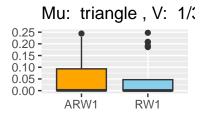




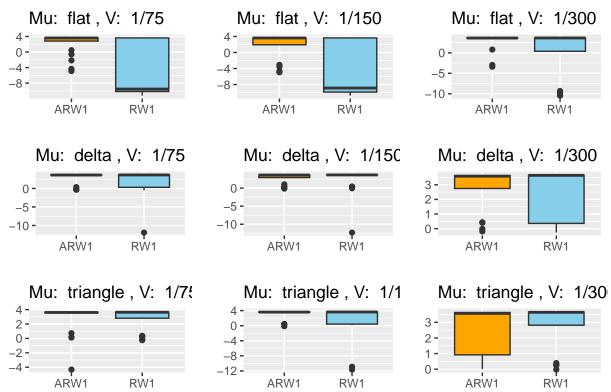








LS for shocked Wakefield data



Worth noting that their model also includes an iid normal latent term, so an unstructured random effect, in both the RW1 and the ARW1.

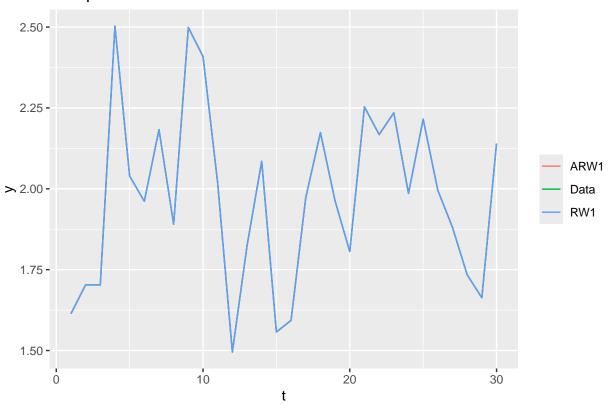
```
#some testing
df <- data_w_flat_300_S
i <- 1

test_data <- df[, c(i, n + 1, n + 2)] #gets the i-th realization + time
colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula

res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data) #control.family = list(hyper=list(p) res_ARW1 <- inla(formula_ARW1_W, family = "gaussian", data = test_data)

plotting_df <- data.frame(matrix(c(res_RW1$summary.fitted.values$mean, res_ARW1$summary.fitted.values$m colnames(plotting_df) <- c("RW1", "ARW1", "Data")
plot_realizations(plotting_df, "Comparison for Wakefield with flat mean")</pre>
```

Comparison for Wakefield with flat mean



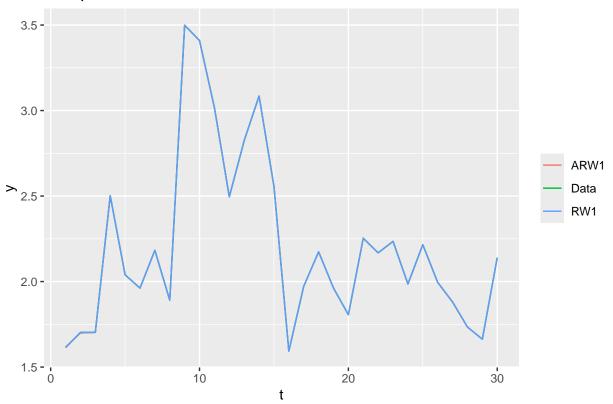
```
df <- data_w_delta_300_S
i <- 1

test_data <- df[, c(i, n + 1, n + 2)] #gets the i-th realization + time
colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula

res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)
res_ARW1 <- inla(formula_ARW1_W, family = "gaussian", data = test_data)

plotting_df <- data.frame(matrix(c(res_RW1$summary.fitted.values$mean, res_ARW1$summary.fitted.values$m
colnames(plotting_df) <- c("RW1", "ARW1", "Data")
plot_realizations(plotting_df, "Comparison for Wakefield with delta mean")</pre>
```

Comparison for Wakefield with delta mean



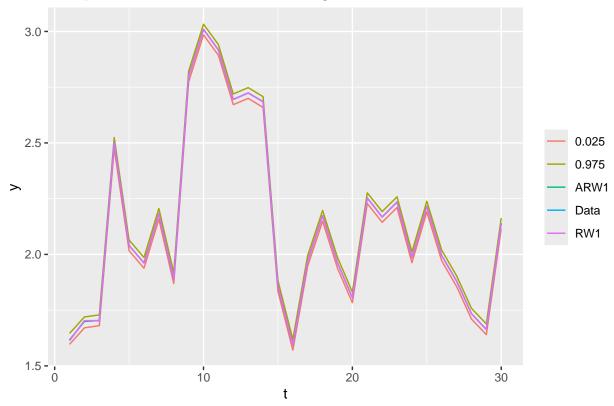
```
df <- data_w_triangle_300_S
i <- 1

test_data <- df[, c(i, n + 1, n + 2)] #gets the i-th realization + time
colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula

res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)
res_ARW1 <- inla(formula_ARW1_W, family = "gaussian", data = test_data)

plotting_df <- data.frame(matrix(c(res_RW1$summary.fitted.values$mean, res_ARW1$summary.fitted.values$"
colnames(plotting_df) <- c("RW1", "0.025", "0.975", "ARW1", "Data")
plot_realizations(plotting_df, "Comparison for Wakefield with triangle mean")</pre>
```





Seems like the ARW1 doesn't work as intended, also seems like RW1 behaves weirdly, maybe I need the unstructured random effects for them to work well on this data.

Harmonious mean

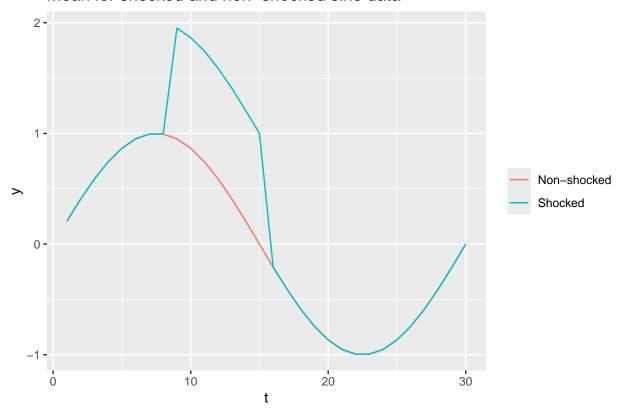
Another similar situation is when the mean is a harmonious function, say some sine function, and we again impose some offset for a subset of the time values.

```
#parameters
N <- 30
n <- 30
V <- 1/1000 #observational variance
tau <- 100 #latent precision in the iid noise
o <- 1 # the offset for points 9-15

t <- 1:N
mean_sine <- sin(pi*t/15)
mean_sine_offset <- mean_sine
mean_sine_offset[9:15] <- mean_sine_offset[9:15] + o

df_mean_sine <- data.frame(matrix(c(mean_sine, mean_sine_offset), ncol = 2))
colnames(df_mean_sine) <- c("Non-shocked", "Shocked")
plot_realizations(df_mean_sine, "Mean for shocked and non-shocked sine data")</pre>
```

Mean for shocked and non-shocked sine data



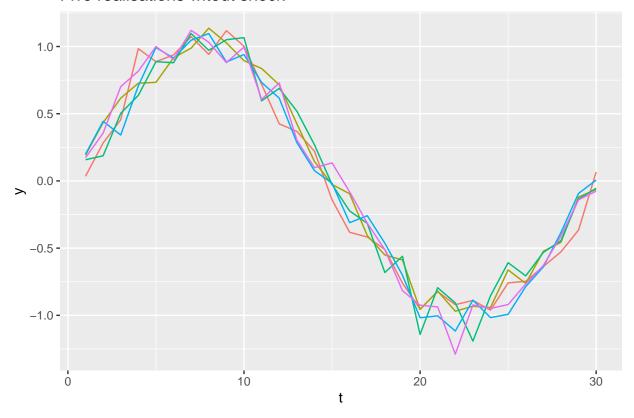
```
sim_dataframe_sine <- function(n, mu, V, tau, o, seed = 64) {
    set.seed(seed)
    df <- data.frame(matrix(0, nrow = length(mu), ncol = n))
    for (i in 1:n) {
        df[, i] <- sim_data_Wakefield(mu, V, tau1 = tau, tau2 = tau) #the function from wakefield above, st
        df[9:15, i] <- df[9:15, i] + o #adding the offset for 9-15
    }
    df$t <- 1:length(mu)
    df$us <- 1:length(mu)
    return(df)
}

df_sine_S <- sim_dataframe_sine(n, mean_sine, V, tau, o)

df_sine_NS <- sim_dataframe_sine(n, mean_sine, V, tau, o)

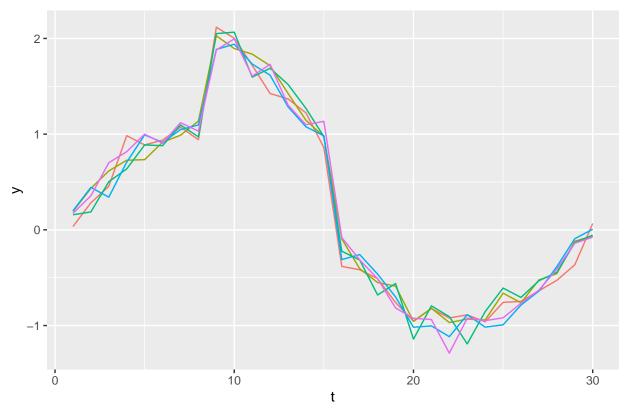
plot_realizations(df_sine_NS[, 1:5], "Five realisations witout shock",, legend = FALSE)</pre>
```

Five realisations witout shock



plot_realizations(df_sine_S[, 1:5], "Five realisations with shock at 9-15", legend =FALSE)

Five realisations with shock at 9-15

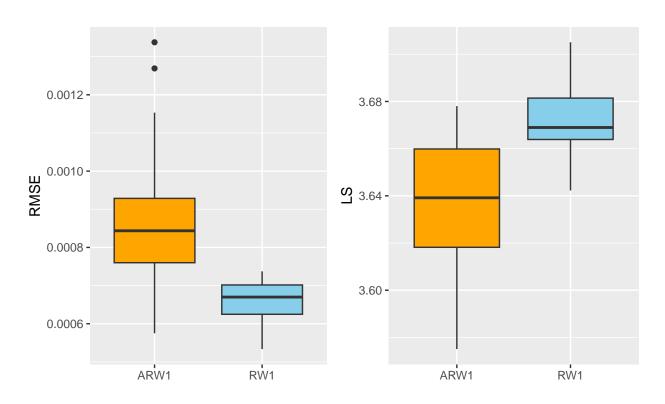


Now we have the data, lets perform the model evaluation with RMSE and LS.

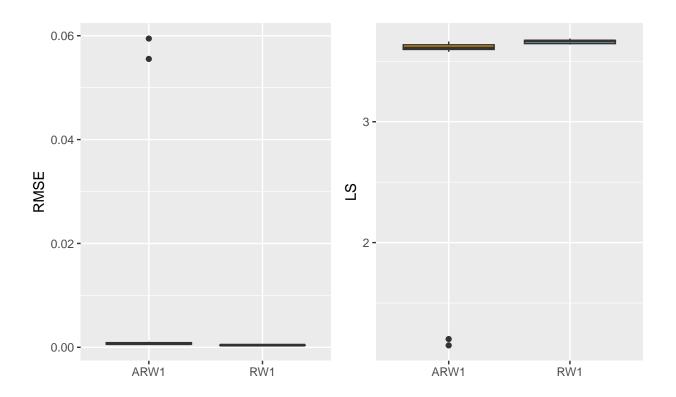
```
#model evaluation
myboxplot3 <- function(df, title, criteria){</pre>
  # Reshape data to long format
  df_long <- df %>%
    pivot_longer(cols = everything(), names_to = "Category", values_to = "Value")
  # Create the boxplot
  eval_plot <- ggplot(df_long, aes(x = Category, y = Value, fill = Category)) +
    geom_boxplot() +
    labs(title = title,x = "", y = criteria) +
    theme(legend.position = "none") +
    scale_fill_manual(values = c("RW1" = "skyblue", "ARW1" = "orange"))
  return(eval_plot)
}
mod_eval_sine <- function(df){</pre>
  n \leftarrow dim(df)[2] - 2
  eval_df <- data.frame(matrix(NA, nrow = n, ncol = 4))</pre>
  colnames(eval_df) <- c("RMSE_RW1", "LS_RW1", "RMSE_ARW1", "LS_ARW1")</pre>
  for(i in 1:n){#iterate over each simulated realization
    test_data <- df[, c(i, n + 1, n + 2)] #gets the i-th realization + time
    colnames(test_data) <- c("y", "time", "us") #makes the colnames match the formula</pre>
```

```
res_RW1 <- inla(formula_RW1, family = "gaussian", data = test_data)</pre>
    LS_RW1 <- average_proper_LS(res_RW1, df[, i])
    RMSE_RW1 <- RMSE(df[, i], res_RW1$summary.fitted.values$mean )</pre>
    res_ARW1 <- inla(formula_ARW1_W, family = "gaussian", data = test_data)</pre>
    LS_ARW1 <- average_proper_LS(res_ARW1, df[, i])
    RMSE_ARW1 <- RMSE(df[, i], res_ARW1$summary.fitted.values$mean )</pre>
    eval_df[i, ] <- c(RMSE_RW1, LS_RW1, RMSE_ARW1, LS_ARW1)</pre>
  df_RMSE <- eval_df[, c("RMSE_RW1", "RMSE_ARW1")]</pre>
  colnames(df_RMSE) <- c("RW1", "ARW1")</pre>
  df_LS <- eval_df[, c("LS_RW1", "LS_ARW1")]</pre>
  colnames(df_LS) <- c("RW1", "ARW1")</pre>
  plot_RMSE <- myboxplot3(df_RMSE, "", "RMSE")</pre>
  plot_LS <- myboxplot3(df_LS, "", "LS")</pre>
  fig <- ggarrange(plot_RMSE, plot_LS, ncol = 2)</pre>
  return(fig)
}
plot_sine_NS <- mod_eval_sine(df_sine_NS)</pre>
annotate_figure(plot_sine_NS,
                 top = text_grob("LS and RMSE for non-shocked sine data"))
```

LS and RMSE for non-shocked sine data



LS and RMSE for shocked sine data



check: maybe weird axis on line 604, definitely weird legend labels, also at 609 Check yaxis for boxplots for the model eval of Gaussian data

inla.doc("rw1") - all info om RW1, for eksempel prior og initial values inla.scale.model - optional argument to scale or not scale a latent effect NB: INLA klikker hvis man bruker funksjoner inne i den, for eksmepel geometric variance, ellers må de defineres det inne potensielt?

To do: Write about INLA - not started Use set.seed() to make results reproducible - continuous need for this dont need to set the seed for INLA, should be deterministic.

Briefly describe the methods used in the overleaf, both for data and models.

Sim-studie: Har undersøkt en RW1 med Gaussian offset for noen datapunkter. Reprodusere Wakefield: 3 ulike trends og tre ulike varianser, pluss forskjell i conf og non-conf tau eller ikke, så 18 scenarior og sjekke LS og RMSE. Andrea foreslo også harmonisk med noise pluss offset.

Må gjøres: implementer med unstructured random effect, altså bare en iid gaussian, både i RW1 og i ARW1, kan kanskje bare legge til en f(iid) i hver formel simulere data og figurer med harmonisk funksjon, med og uten offset. Present the box plots with 2 in every box, not the difference. Også ha samme y-akse på hver linje når det er tre ved siden av hverandre.

Code for ARW1 from github Alekshin Guendel

```
#The rgeneric inla function for user defined random effects, i e. a ARW1
inla.rgeneric.bym2.model = function(
    cmd = c("graph", "Q", "mu", "initial", "log.norm.const", "log.prior",
            "quit"),
   theta = NULL)
    # Assume we are passed the following inputs
    # n: the number of subdivisions (i.e. time points or regions)
    \# Q_star: the scaled structure matrix for the structured component
    # gamma_tilde: the inverse eigenvalues of Q_star
    # U_prec, alpha_prec: U and alpha parameters for precision PC prior
    # U_phi, alpha_phi: U and alpha parameters for mixing parameter PC prior
    # for reference and potential storage for objects to
    # cache, this is the environment of this function
    # which holds arguments passed as `...` in
    # `inla.rgeneric.define()`.
   envir = parent.env(environment())
   interpret.theta = function() {
        return(list(prec = exp(theta[1L]),
                    phi = 1 / (1 + exp(-theta[2L]))))
   }
   graph = function(){ return (Q()) }
    Q = function() {
        p = interpret.theta()
        D = (1 / (1 - p\$phi)) * diag(n)
        QQ = rbind(cbind(p$prec * D, -sqrt(p$phi * p$prec) * D),
                   cbind(-sqrt(p$phi * p$prec) * D, Q_star + p$phi * D))
       return (inla.as.sparse(QQ))
   }
   mu = function() { return(numeric(0)) }
   log.norm.const = function() { return (numeric(0)) }
   log.prior = function() {
       p = interpret.theta()
        # Construct prior for the precision parameter
        lambda_prec = -log(alpha_prec) / U_prec
        prec_prior = log(lambda_prec) - log(2) - theta[1L] / 2 -
            lambda_prec / sqrt(p$prec)
        # Construct prior for the mixing parameter phi
        dU = sqrt(U_phi * sum(gamma_tilde - 1) -
                      sum(log(1 + U_phi * (gamma_tilde - 1))))
        lambda_phi = -log(1 - alpha_phi) / dU
        d2 = p$phi * sum(gamma_tilde - 1) -
```

```
sum(log(1 + p$phi * (gamma_tilde - 1)))
        phi_prior = log(lambda_phi) + 2 * theta[2L] -
            2 * log1p(exp(theta[2L])) - log(2) - log(d2) / 2 -
            lambda_phi * sqrt(d2) +
            log(abs(sum( (gamma_tilde - 1) ^ 2 /
                              (1 + exp(theta[2L]) * gamma_tilde))))
        return(prec_prior + phi_prior)
    }
    initial = function() { return(c(4, 0)) }
    quit = function() { return (invisible()) }
    if (!length(theta)) theta = initial()
    val = do.call(match.arg(cmd), args = list())
    return (val)
}
# ----- The following is my work trying to understand the code -----
?do.call
?match.arg
#I believe the match.arg(cmd) are all the values we are interested in
# which were defined as functions with no arguments further up
# or by arguments passed globally to the function as specified at the top,
\# so, val = c("graph", "Q", "mu", "initial", "log.norm.const", "log.prior", "quit")
#I assume these quantities are needed later?
?inla.rgeneric.bym2.model
#might need these, dont know
library(SUMMER)
library(readstata13)
library(dplyr)
vignette()
#### Get smoothed direct estimates w/ time fixed effect, adaptive bym2 ####
# Create structure matrices
conflict_years <- 1993:1999 - 1984</pre>
conflict_years_long <- rep(0, num_years)</pre>
conflict_years_long[conflict_years] <- 1</pre>
R_conflict <- matrix(0, num_years, num_years)</pre>
R_nonconflict <- matrix(0, num_years, num_years)</pre>
for(i in 1:num_years){
  if(i == 1){
    if(conflict_years_long[i] || conflict_years_long[i + 1]){
      R_conflict[i, i] <- 1</pre>
    }
    else{
      R_nonconflict[i, i] <- 1</pre>
```

```
else if(i == num_years){
    if(conflict_years_long[i] || conflict_years_long[i - 1]){
      R_conflict[i, i] <- 1</pre>
    }
    else{
      R_nonconflict[i, i] <- 1</pre>
    }
  }
  else{
    if(conflict_years_long[i] | (conflict_years_long[i - 1] &&
                                    conflict_years_long[i + 1])){
      R_conflict[i, i] <- 2</pre>
    else if(conflict_years_long[i - 1] || conflict_years_long[i + 1]){
      R_conflict[i, i] <- 1</pre>
      R_nonconflict[i, i] <- 1</pre>
    }
    else{
      R_nonconflict[i, i] <- 2</pre>
  }
  for(j in 1:num_years){
    if(abs(i - j) == 1){
      if(conflict_years_long[i] || conflict_years_long[j]){
        R_conflict[i, j] <- -1</pre>
      else{
        R_nonconflict[i, j] <- -1</pre>
      }
    }
 }
R_1 \leftarrow R_nonconflict
R_2 \leftarrow R_{conflict}
scaled_Q <- INLA:::inla.scale.model.bym.internal(R_1 + R_2,</pre>
                                                      adjust.for.com.comp = TRUE) $Q
gv \leftarrow scaled_Q[1, 1] / (R_1 + R_2)[1, 1]
R_1_star <- gv * R_1
R_2_star <- gv * R_2
vals <- (1:num_years)[-num_years]</pre>
R_1_star_hat <- R_1_star[vals, vals]</pre>
R_2_star_hat <- R_2_star[vals, vals]</pre>
eps <- eigen(solve(R_1_star_hat + R_2_star_hat) %*% R_2_star_hat)$values
gamma_tilde <- c(1 / eigen(R_1_star + R_2_star)$values[1:(num_years - 1)], 0)</pre>
save(R_1_star, R_2_star, eps, gamma_tilde,
     file = "../Data/generated_data/structure_matrices.RData")
# Specify PC prior hyperparameters
pc.u.theta <- 0.75
```

```
pc.alpha.theta <- 0.75
# Fit model
adaptive bym2 model <-
  INLA::inla.rgeneric.define(model = inla.rgeneric.adaptive.bym2.model,
                              n = num_years, R_1_star = R_1_star,
                              R_2_star = R_2_star, gamma_tilde = gamma_tilde,
                              eps = eps, U_prec = pc.u, alpha_prec = pc.alpha,
                              U_phi = pc.u.phi, alpha_phi = pc.alpha.phi,
                              U_theta = pc.u.theta,
                              alpha_theta = pc.alpha.theta)
constr <- list(A = matrix(c(rep(0, num_years), rep(1, num_years)),</pre>
                           nrow = 1, ncol = 2 * num_years), e = 0)
mod <- logit.est ~ time +
 f(region.struct, model = adaptive_bym2_model,
    diagonal = 1e-06, extraconstr = constr, n = 2 * num_years) +
  f(survey.id, model = "iid", hyper = hyperpc1)
options <- list(dic = TRUE, mlik = TRUE, cpo = TRUE,
                openmp.strategy = "default", return.marginals.predictor = TRUE)
control.inla <- list(strategy = "adaptive", int.strategy = "auto")</pre>
fit_adaptive_linear <-
  INLA::inla(mod, family = "gaussian", control.compute = options,
             data = dat,
             control.predictor = list(compute = TRUE),
             control.family =
               list(hyper = list(prec = list(initial = log(1),
                                                fixed = TRUE))),
             scale = dat$logit.prec,
             control.inla = control.inla, verbose = FALSE)
out_adaptive_linear <- dat %>% select(region, years) %>%
  mutate(median = NA, lower = NA, upper = NA, logit.median = NA,
         logit.lower = NA, logit.upper = NA)
for (i in 1:nrow(dat)) {
  tmp.logit <-</pre>
    INLA::inla.rmarginal(1e+05,
                          fit_adaptive_linear$marginals.fitted.values[[i]])
  tmp <- expit(tmp.logit)</pre>
  out_adaptive_linear$median[i] <- median(tmp)</pre>
  out_adaptive_linear$lower[i] <- quantile(tmp, probs = 0.05)</pre>
  out_adaptive_linear$upper[i] <- quantile(tmp, probs = 0.95)</pre>
  out_adaptive_linear$logit.median[i] <- median(tmp.logit)</pre>
  out_adaptive_linear$logit.lower[i] <- quantile(tmp.logit, probs = 0.05)</pre>
  out_adaptive_linear$logit.upper[i] <- quantile(tmp.logit, probs = 0.95)</pre>
out_adaptive_bym2 <- out_adaptive_linear %>%
 filter(is.na(years))
smoothed_direct_adaptive_bym2$fit <- fit_adaptive_linear</pre>
smoothed_direct_adaptive_bym2$model <- mod</pre>
out_combined <- rbind(cbind(out_bym2, prior = "bym2"),</pre>
```

```
cbind(out_adaptive_bym2, prior = "adaptive bym2"))

# ----- The following is my work trying to understand the code ------
#first part is computing the structure and precision matrices, and scaled
# A lot of parameters and priors etc, then we define the ARW1 and include
# it in the formula for the inla call. Also a sum to zero constraint maybe?
# After the model call I am not sure whats happening.
```

Simulation of non-shocked Poisson data

We will simulate data with a latent temporal structured random effect as a RW1. The total Bayesian hierarchical model can be described as

$$Y|\lambda \sim Poisson(E\lambda)\log\lambda_t = \mu + x_t$$

Where $\mathbf{x} \sim RW1(\tau)$ where we fix E, μ and τ for the simulations. When fitting models we will need to assign priors to them.

```
#Simulating non-shocked data
#Parameters
E = 100
mu = 1
sigma = 0.2
T = 100 #Number of time points
#A single simulation
x <- RW1(sigma, T)
rates <- E * exp(mu + x) #rates is E*lambda or E*exp(mu + x)
y <- sapply(rates, function(r) rpois(1, r)) #samples from the Poisson
plot(1:100, x)
sim_non_shocked_data <- function(E, mu, sigma, T){</pre>
 x <- RW1(sigma, T)
 rates <- E * exp(mu + x) #rates is E*lambda or E*exp(mu + x)
  y <- sapply(rates, function(r) rpois(1, r)) #samples from the Poisson
  return(y) #The observed data
sim_non_shocked_dataframe <- function(E, mu, sigma, T, n, seed = 44){</pre>
  set.seed(seed)
  df <- data.frame(matrix(NA, nrow = T, ncol = n))</pre>
  for(i in 1:n){df[, i] <- sim_non_shocked_data(E, mu, sigma, T) }</pre>
  return(df)
}
test <- sim_non_shocked_dataframe(100, 4, 0.001, 100, 5)
test$t <- 1:nrow(test) # Create a time index from 1 to n
# Reshape the dataframe to long format
```

```
test_long <- test %>% pivot_longer(cols = starts_with("X"), names_to = "variable", values_to = "value")

# Plot all lines using ggplot
ggplot(test_long, aes(x = t, y = value, color = variable)) +
    geom_line() +
    labs(title = "Simulated non-shocked data with N=100 and 5 realisations",
        x = "Time", y = "y") + theme(legend.position = "none")
```

The code above seems to work fine, but should possibly tune some of the parameters, or maybe the plot of the ys should look insane.

Now the next step is to fit some models with INLA, we will start with the direct smoothed model. ## Fitting the model with INLA

```
#need to input the data as a dataframe with a y column and time column
formula <- y ~ f(time, model = "rw1") #intercept is included automatically

test_data <- data.frame(matrix(c(test[, 2], 1:100, rep(E, 100)), nrow = 100, ncol = 3)) # should make t
colnames(test_data) <- c("y", "time", "E")

res <- inla(formula, E = E, family = "poisson", data = test_data)
plot(res)
summary(res)
?inla</pre>
```

Seems like it works quite well, the posterior distributions align rather well with the true parameters for mean and precision.