Moments of the truncated bivariate Normal distribution

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Abstract

Miscellaneous notes about the computation of the moments of a truncated bivariate normal distribution.

1 Introduction

We consider a standardized normal bivariate distribution of the form:

$$\mathcal{N}\left(\mathbf{x};\rho\right) = \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{H\left(\mathbf{x}\right)}{2\left(1-\rho^2\right)}\right), \quad H\left(\mathbf{x}\right) = x_1^2 + x_2^2 - 2\rho x_1 x_2$$

where $-1 < \rho < 1$ and $\mathbf{x} = (x_1, x_2)$. The "standardization" means that $\langle x_i \rangle_{\mathcal{N}} = 0$ and $\langle x_i^2 \rangle_{\mathcal{N}} = 1$ for i = 1, 2, while $\rho = \langle x_1 x_2 \rangle_{\mathcal{N}}$. Any normal bivariate distribution can be brought into this form by a translation and scaling of the variables x_1, x_2 . Note that $\mathcal{H}(\mathbf{x})$ is concave everywhere because the Hessian

$$\nabla^2 H = \left[\begin{array}{cc} 2 & -2\rho \\ -2\rho & 2 \end{array} \right]$$

has the positive eigenvalues $2(1 \pm \rho)$.

The levelsets of $\mathcal{N}\left(\mathbf{x};\rho\right)$ are the ellipses $H\left(\mathbf{x}\right)=c$ for some constant $c\geq0$. The axes of this ellipse are the lines $x_1\pm x_2=0$ with semi-lenghts $\sqrt{c/\left(1\pm\rho\right)}$, respectively (Fig. 1). Let $\mathfrak{f}=\sqrt{|\rho|}\,c/\left(1-\rho^2\right)$. The focci lie at a distance $\sqrt{2}\mathfrak{f}$ from the origin. If $\rho<0$ the focci are at the coordinates $x_1=-x_2=\pm\mathfrak{f}$, and if $\rho>0$ the focci are at the coordinates $x_1=x_2=\pm\mathfrak{f}$. The ellipse $H\left(\mathbf{x}\right)=c$ is the locus of points whose distances to the focci add up to $2\sqrt{c/\left(1-|\rho|\right)}$.

Define a rectangular domain \mathcal{D} , with vertices at $\mathbf{a} = (a_1, a_2)$ and $\mathbf{b} = (b_1, b_2)$,

$$\mathcal{D}\left(\mathbf{a},\mathbf{b}\right)=\left\{ \mathbf{x}\in\mathbb{R}^{2}:\mathbf{a}\leq\mathbf{x}\leq\mathbf{b}\right\}$$

where the vector inequalities are component-wise and we assume that $\mathbf{a} \leq \mathbf{b}$. The bivariate density truncated to \mathcal{D} is given by:

$$\mathcal{P}(\mathbf{x}; \mathbf{a}, \mathbf{b}, \rho) = \frac{1}{Z(\mathbf{a}, \mathbf{b}, \rho)} \begin{cases} \mathcal{N}(\mathbf{x}; \rho) & \mathbf{a} \leq \mathbf{x} \leq \mathbf{b} \\ 0 & \text{otherwise} \end{cases}$$

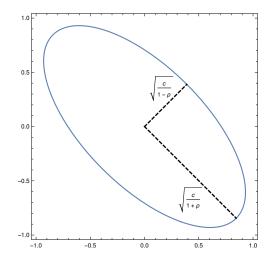


Figure 1: The level sets $H\left(\mathbf{x}\right)=c$ are ellipses. The curve shown has c=0.5, $\rho=-0.65.$

where Z is a normalization constant, given by:

$$Z\left(\mathbf{a}, \mathbf{b}, \rho\right) = \int_{a_1}^{b_1} \mathrm{d}x_1 \int_{a_2}^{b_2} \mathrm{d}x_2 \, \mathcal{N}\left(\mathbf{x}; \rho\right)$$

In this note we want to compute the moments:

$$\langle x_i \rangle_{\mathcal{P}} = \int_{a_1}^{b_1} \mathrm{d}x_1 \int_{a_2}^{b_2} \mathrm{d}x_2 \mathcal{P}\left(\mathbf{x}; \mathbf{a}, \mathbf{b}, \rho\right) x_i, \quad i = 1, 2$$

$$\langle x_i^2 \rangle_{\mathcal{P}} = \int_{a_1}^{b_1} \mathrm{d}x_1 \int_{a_2}^{b_2} \mathrm{d}x_2 \mathcal{P}\left(\mathbf{x}; \mathbf{a}, \mathbf{b}, \rho\right) x_i^2, \quad i = 1, 2$$

$$\langle x_1 x_2 \rangle_{\mathcal{P}} = \int_{a_1}^{b_1} \mathrm{d}x_1 \int_{a_2}^{b_2} \mathrm{d}x_2 \mathcal{P}\left(\mathbf{x}; \mathbf{a}, \mathbf{b}, \rho\right) x_1 x_2$$

2 General parameters

Consider now the general truncated Gaussian with a positive semi-definite covariance matrix Σ and means μ . The change of variables:

$$\xi_i = \frac{x_i - \mu_i}{\sqrt{\Sigma_{ii}}}$$

leads to the previously considered distribution, with non-truncated moments given by:

$$\langle \xi_i \rangle_{\mathcal{N}} = 0, \quad \langle \xi_i^2 \rangle_{\mathcal{N}} = 1, \quad i = 1, 2$$

$$\rho = \langle \xi_1 \xi_2 \rangle_{\mathcal{N}} = \frac{\Sigma_{12}}{\sqrt{\Sigma_{11} \Sigma_{22}}}$$

If the original variables x_i are bounded by $a_i \leq x_i \leq b_i$ then the ξ_i are bounded by $\alpha_i \leq \xi_i \leq \beta_i$, where

$$\alpha_i = \frac{a_i - \mu_i}{\sqrt{\Sigma_{ii}}}, \qquad \beta_i = \frac{b_i - \mu_i}{\sqrt{\Sigma_{ii}}}$$

Using the methods described below, we can compute the truncated moments of ξ_i . Then:

$$\begin{split} \langle x_1^m x_2^n \rangle_{\mathcal{P}} &= \left\langle \left(\sqrt{\Sigma_{11}} \xi_1 + \mu_1 \right)^m \left(\sqrt{\Sigma_{11}} \xi_2 + \mu_2 \right)^n \right\rangle_{\mathcal{P}} \\ &= \sum_{i=0}^m \sum_{j=0}^n \binom{m}{i} \binom{n}{j} \mu_1^{m-i} \mu_2^{n-j} \Sigma_{11}^{i/2} \Sigma_{22}^{j/2} \left\langle \xi_1^i \xi_2^j \right\rangle_{\mathcal{P}} \end{split}$$

Therefore it suffices to consider the truncated normal distribution with $\mu_i = 0$ and $\Sigma_{ii} = 1$, $\Sigma_{12} = \Sigma_{21} = \rho$.