

Moments of the univariate truncated normal distribution

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The PDF of the truncated normal distribution is¹:

$$f(x; \mu, \sigma, a, b) = \begin{cases} \frac{1}{\sigma} \frac{\phi\left(\frac{x-\mu}{\sigma}\right)}{\Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)} & a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

where μ, σ are the mean and standard deviation of the untruncated normal, and

$$\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad \Phi(x) = \frac{1}{2} \left(1 + \operatorname{erf}\left(x/\sqrt{2}\right)\right)$$

The mean and variance of this distribution are:

$$\begin{aligned} \langle x \rangle &= \mu + \frac{\phi(\alpha) - \phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} \sigma \\ \operatorname{var} x &= \left[1 + \frac{\alpha\phi(\alpha) - \beta\phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} - \left(\frac{\phi(\alpha) - \phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} \right)^2 \right] \sigma^2 \end{aligned}$$

where

$$\alpha = \frac{a - \mu}{\sigma}, \quad \beta = \frac{b - \mu}{\sigma}$$

The numerical evaluation of $\frac{\phi(\alpha) - \phi(\beta)}{\Phi(\beta) - \Phi(\alpha)}$ and $\frac{\alpha\phi(\alpha) - \beta\phi(\beta)}{\Phi(\beta) - \Phi(\alpha)}$ can be problematic. We describe our approach below.

Numerical evaluation

Let us define the functions:

$$F_1(x, y) = \frac{e^{-x^2} - e^{-y^2}}{\operatorname{erf}(y) - \operatorname{erf}(x)}, \quad F_2(x, y) = \frac{xe^{-x^2} - ye^{-y^2}}{\operatorname{erf}(y) - \operatorname{erf}(x)} \quad (1)$$

The problem arises when x, y have the same sign and both are large in magnitude. Then $\operatorname{erf}(y) - \operatorname{erf}(x)$ is the difference of two numbers close to ± 1 , but

¹See https://en.wikipedia.org/wiki/Truncated_normal_distribution.

their difference is very small. Subtraction of two numbers that are both large relative to their difference is a classic source of floating-point errors (known as *catastrophic cancellation*).

Recall a few useful properties of the error and related functions:

$$\operatorname{erf}(-x) = \operatorname{erf}(x), \quad \operatorname{erfc}(x) = 1 - \operatorname{erf}(x) = e^{-x^2} \operatorname{erfcx}(x)$$

The functions $\operatorname{erf}(x)$, $\operatorname{erfc}(x)$, $\operatorname{erfcx}(x)$ are all available in numerical packages. Using these, we can make sure that the error function difference is never catastrophic. Let $\Delta = e^{x^2 - y^2}$. We propose:

$$\begin{aligned} F_1(x, y) &= F_1(y, x), & F_2(x, y) &= F_2(y, x), & \text{if } |x| > |y| \\ &= P_1(x, y - x), & &= P_2(x, y - x), & \text{if } |x - y| < 10^{-7} \\ &= \frac{1 - \Delta}{\Delta \operatorname{erfcx}(-y) - \operatorname{erfcx}(-x)} & &= \frac{x - y\Delta}{\Delta \operatorname{erfcx}(-y) - \operatorname{erfcx}(-x)} & \text{if } x, y \leq 0 \\ &= \frac{1 - \Delta}{\operatorname{erfcx}(x) - \Delta \operatorname{erfcx}(y)} & &= \frac{x - y\Delta}{\operatorname{erfcx}(x) - \Delta \operatorname{erfcx}(y)} & \text{if } x, y \geq 0 \\ &= \frac{(1 - \Delta) e^{-x^2}}{\operatorname{erf}(y) - \operatorname{erf}(x)} & &= \frac{(x - y\Delta) e^{-x^2}}{\operatorname{erf}(y) - \operatorname{erf}(x)} & \text{otherwise} \end{aligned}$$

When $x \approx y$, we employ a Taylor expansion of $F_i(x, x + \epsilon)$ in powers of $\epsilon = y - x$ to fourth-order,

$$\begin{aligned} P_1(x, \epsilon) &= \sqrt{\pi}x + \frac{1}{2}\sqrt{\pi}\epsilon - \frac{1}{6}\sqrt{\pi}x\epsilon^2 - \frac{1}{12}\sqrt{\pi}\epsilon^3 + \frac{1}{90}\sqrt{\pi}x(x^2 + 1)\epsilon^4 \\ P_2(x, \epsilon) &= \frac{1}{2}\sqrt{\pi}(2x^2 - 1) + \sqrt{\pi}x\epsilon - \frac{1}{3}\sqrt{\pi}(x^2 - 1)\epsilon^2 - \frac{1}{3}\sqrt{\pi}x\epsilon^3 + \frac{1}{90}\sqrt{\pi}(2x^4 + 3x^2 - 8)\epsilon^4 \end{aligned}$$

Infinite arguments

The limit of Eq. (1) as $y \rightarrow \pm\infty$ is a well-defined function of x ,

$$F_1(x, \pm\infty) = \frac{\pm 1}{\operatorname{erfcx}(\pm x)}, \quad F_2(x, \pm\infty) = \frac{\pm x}{\operatorname{erfcx}(\pm x)} \quad (2)$$

We define $F_i(\infty, \pm\infty) = \lim_{x \rightarrow \infty} F_i(x, \pm\infty)$ and $F_i(-\infty, \pm\infty) = \lim_{x \rightarrow -\infty} F_i(x, \pm\infty)$. Consistent with these definitions, we use Eq. (2) to evaluate $F_1(x, \pm\infty)$ for all x and $F_2(x, \pm\infty)$ for all $-\infty < x < \infty$. Finally, $F_2(-\infty, \infty) = F_2(\infty, -\infty) = 0$, while $F_2(\infty, \infty) = F_2(-\infty, -\infty) = \infty$.

Mean and variance

From $F_i(x, y)$, we evaluate the mean and variance of the truncated normal distribution as follows:

$$\frac{\phi(\alpha) - \phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} = \sqrt{\frac{2}{\pi}} F_1\left(\frac{\alpha}{\sqrt{2}}, \frac{\beta}{\sqrt{2}}\right), \quad \frac{\alpha\phi(\alpha) - \beta\phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} = \frac{2}{\sqrt{\pi}} F_2\left(\frac{\alpha}{\sqrt{2}}, \frac{\beta}{\sqrt{2}}\right)$$

$$\frac{\alpha\phi(\alpha) - \beta\phi(\beta)}{\Phi(\beta) - \Phi(\alpha)} - \left(\frac{\phi(\alpha) - \phi(\beta)}{\Phi(\beta) - \Phi(\alpha)}\right)^2 = \frac{2}{\sqrt{\pi}} F_2\left(\frac{\alpha}{\sqrt{2}}, \frac{\beta}{\sqrt{2}}\right) - \frac{2}{\pi} \left[F_1\left(\frac{\alpha}{\sqrt{2}}, \frac{\beta}{\sqrt{2}}\right)\right]^2$$