

Bayesian Quantile Regression

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Section 1

Standard Quantile Regression

What is Quantile Regression?

As opposed least squares regression which estimates conditional mean functions, quantile regression focuses on estimating families of conditional quantile functions.

$$\mathbb{E}[y_i|\mathbf{x}_i] = \mathbf{x}_i^T \boldsymbol{\beta}$$

traditional regression

$$q_\tau(y_i|\mathbf{x}_i) = \mathbf{x}_i^T \boldsymbol{\beta}(\tau)$$

quantile regression

- Standard Linear Model: $y_t = \mu(\mathbf{x}_t) + \epsilon_t$
 - $\mu(\mathbf{x}_t) = \mathbf{x}_t' \beta$
- The τ th regression quantile is defined as any solution, $\hat{\beta}(\tau)$, to the quantile regression minimisation problem:

$$\min_{\beta} \sum_t \rho_{\tau}(y_t - \mathbf{x}_t' \beta)$$

- Loss Function: $\rho_{\tau}(u) = u(\tau - I(u < 0))$

Asymmetric Laplace Distribution

- U has an asymmetric Laplace distribution if its probability density is given by:

$$f_{\tau}(u) = \tau(1 - \tau)\exp\{-\rho_{\tau}(u)\}$$

where $0 < \tau < 1$.

- (Note that location and scale parameters can be incorporated into this density)

Bayesian Quantile Regression

- In the Bayesian framework we are interested in the conditional quantile, $q_\tau(y_i|\mathbf{x}_i)$, rather than the conditional expectation.
- Given the observations $\mathbf{y} = (y_1, \dots, y_n)$, the posterior distribution of β is given by:

$$\pi(\beta|\mathbf{y}) \propto L(\mathbf{y}|\beta)p(\beta)$$

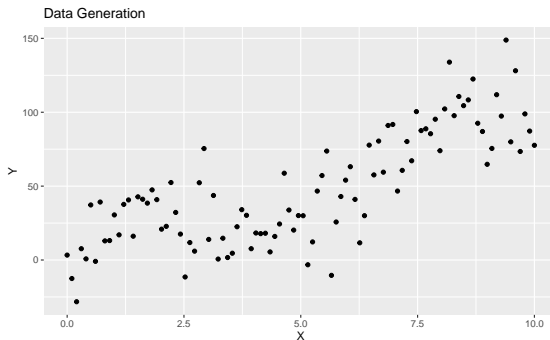
where $p(\beta)$ is the prior distribution of β and the likelihood is:

$$L(\mathbf{y}|\beta) = p^n(1-p)^n \exp\left\{-\sum_i \rho_\tau(y_i - \mathbf{x}_i'\beta)\right\}$$

using the location parameter $\mu_i = \mathbf{x}_i'\beta$

- Note: In the absence of a realistic prior we can use an improper uniform prior for all components of β

Data Generation - Normal error

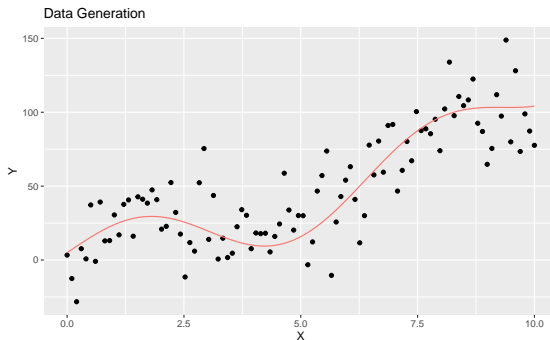


The data was generated via

$$y_i = 5 + x + x^2 + 20 \sin(x) + \epsilon$$

$$\epsilon \sim \mathcal{N}(0, 20^2) \implies SNR \approx 2.42$$

Data Generation - Normal Error

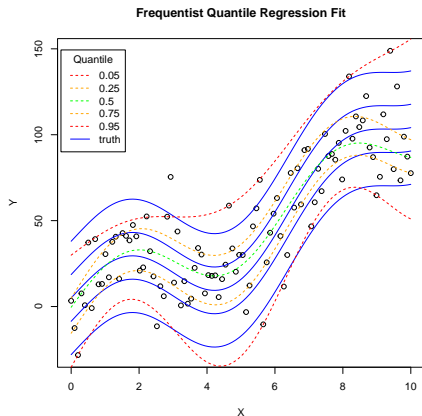


The data was generated via

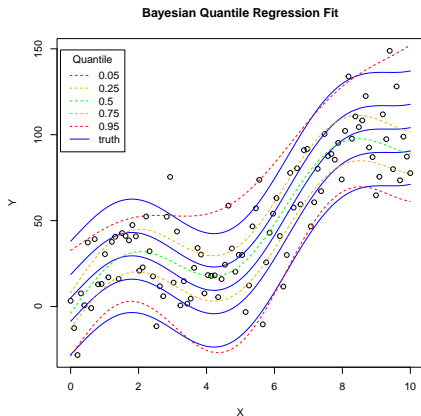
$$y_i = 5 + x + x^2 + 20 \sin(x) + \epsilon$$

$$\epsilon \sim \mathcal{N}(0, 20^2) \implies SNR \approx 3.63$$

Bayesian vs Frequentist - Normal error



(a)



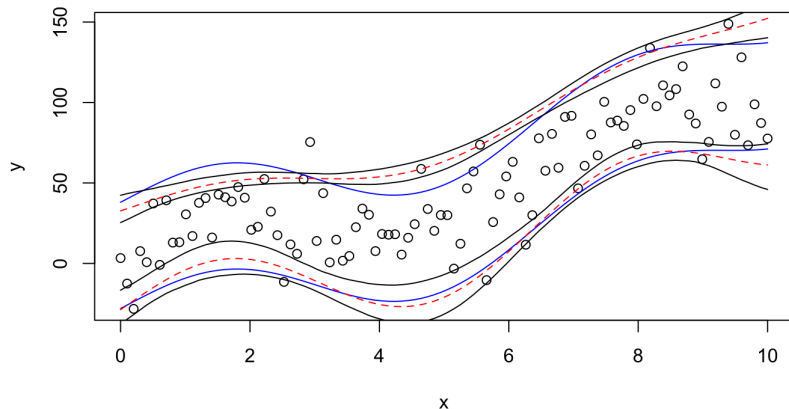
(b)

Figure: (a) Frequentist QR. (b) Bayesian QR.

Quantile Credible Intervals - Normal error

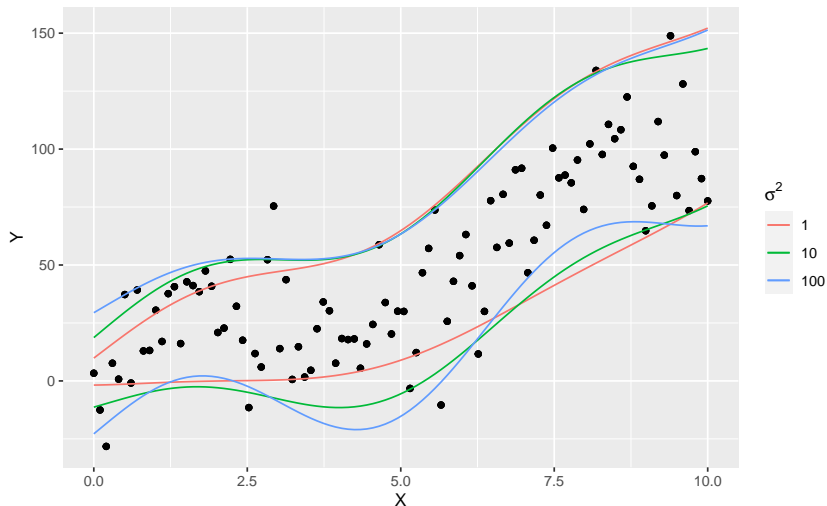
Figure: 0.05, 0.95 Quantiles and their 95% Credible Interval

Scatterplot and Quantile Regression Fit



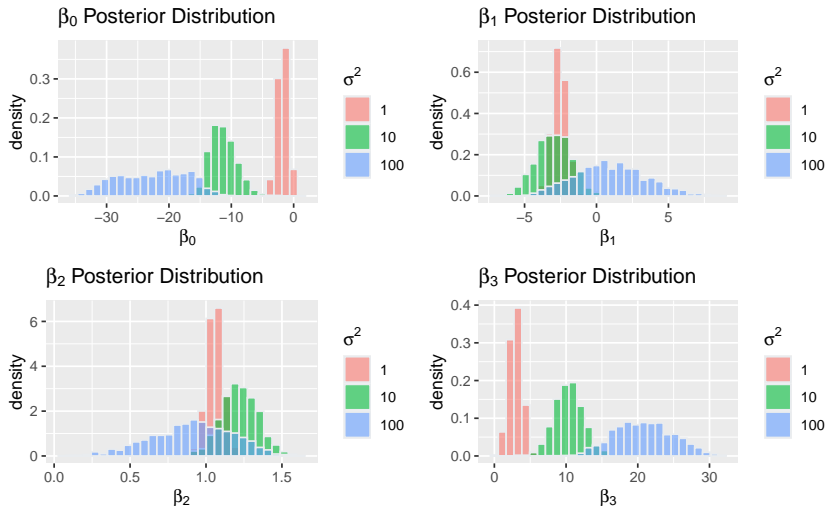
Sensitivity Analysis

Figure: 0.05, 0.95 Quantile Regression Lines



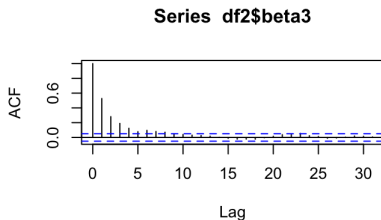
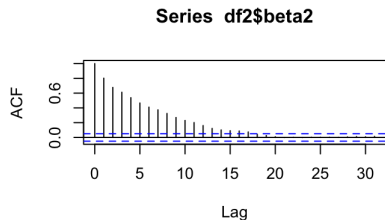
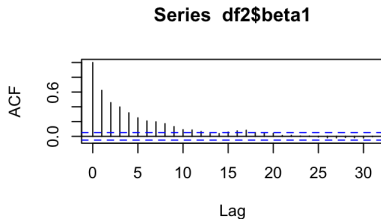
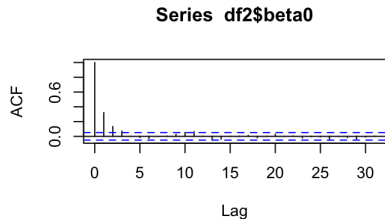
Sensitivity Analysis

Figure: 95% Beta Posteriors

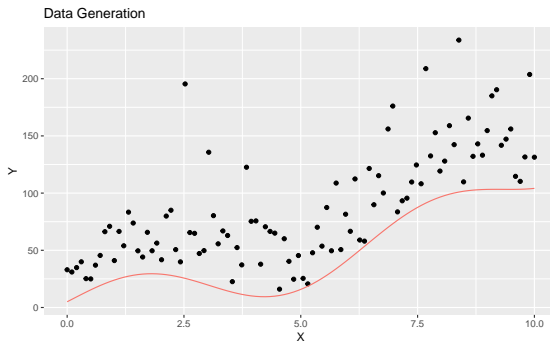


MCMC Diagnostics

Figure: Autocorrelation Plots



Data Generation - χ^2 error



The data was generated via

$$y_i = 5 + x + x^2 + 20 \sin(x) + 10 * \epsilon$$

$$\epsilon \sim \chi_4^2 \implies SNR \approx 3.06$$

Bayesian vs Frequentist - χ^2 error

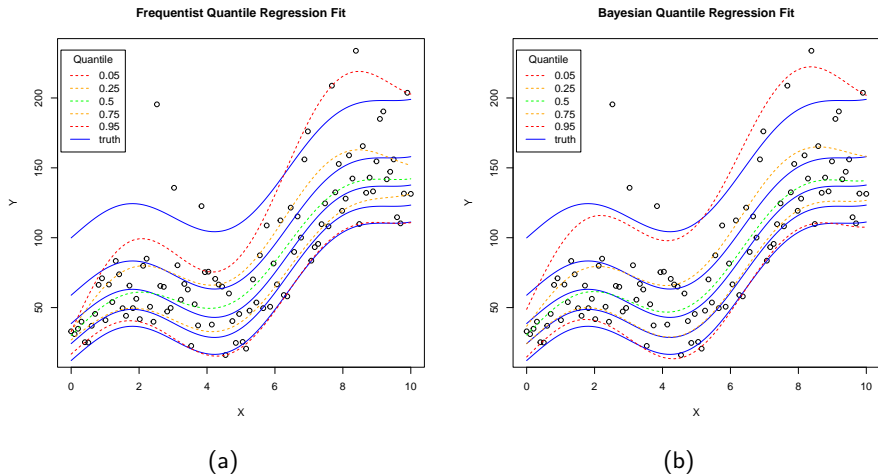
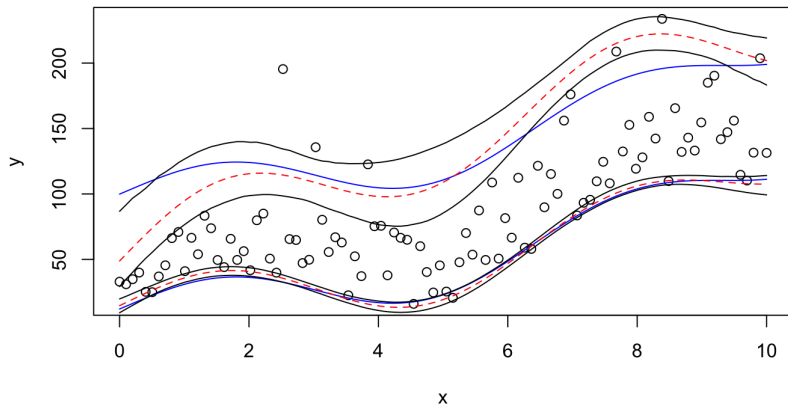


Figure: (a) Frequentist QR. (b) Bayesian QR.

Quantile Credible Intervals - χ^2 error

Figure: 0.05, 0.95 Quantiles and their 95% Credible Interval

Scatterplot and Quantile Regression Fit



Subsection 3

Real Data Example

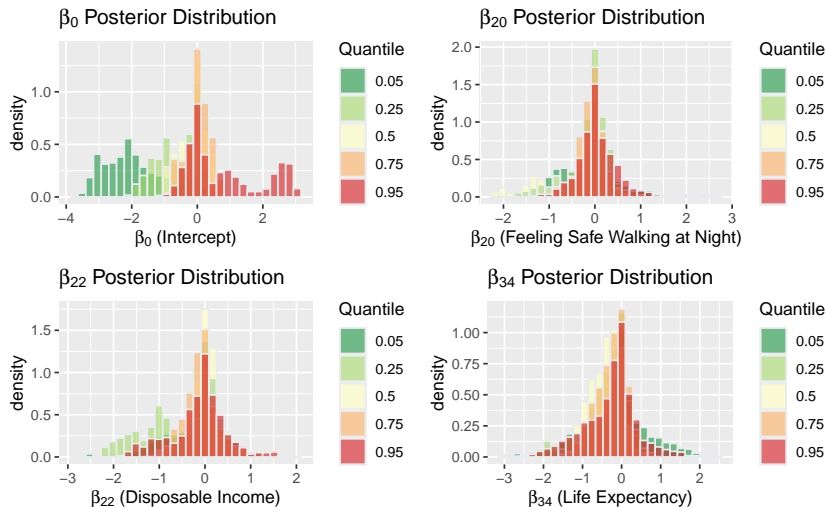
World Development Factors and BMI

Figure: Data Composition

	BMI[, 1]	Age group	Labour market insecurity[, 1]	Stakeholder engagement for developing regulations[, 1]	Dwellings without basic facilit
1	-1.147033255	18-19	-0.43110631	1.141138822	
2	-1.904771749	18-19	-0.62165053	-1.508899200	-0
3	-1.509479344	18-19	-0.59783251	-0.183880189	-0
4	-1.471243637	18-19	NA	0.194696671	0
5	-1.276634097	18-19	-0.26438011	1.519715682	-0
6	-0.923220259	18-19	0.49779679	-1.508899200	0
7	-1.560517393	18-19	NA	-1.319610770	1
8	-1.105922375	18-19	NA	-0.562457050	-0
9	-2.146164838	18-19	-0.62165053	-0.941033910	-0
10	-1.786361680	18-19	-0.09765392	-0.183880189	-0
11	-1.816586709	18-19	0.11670834	1.141138822	0
12	-1.959147603	18-19	-0.64546856	0.194696671	-0
13	-2.160005883	18-19	-0.43110631	0.005408241	-0
14	-1.610007698	18-19	-0.83601279	-0.562457050	-0
15	-1.767856834	18-19	3.99904693	-0.562457050	-0
16	-1.808632729	18-19	-0.26438011	-1.698187631	-0
17	-1.519427236	18-19	-0.93128490	0.005408241	-0
18	-1.213053798	18-19	-0.55019645	-1.508899200	-0
19	-1.465434868	18-19	-0.07383589	0.762561962	
20	-2.121626478	18-19	0.87888524	0.762561962	-0
21	-2.478171094	18-19	-0.52637842	-1.319610770	0
22	-2.293674383	18-19	0.33107059	0.194696671	0
23	-2.174829783	18-19	NA	0.573273531	0
24	-1.898703462	18-19	-0.64546856	-0.751745480	-0
25	-1.010575245	18-19	-0.21674406	2.087580973	2

World Development Factors and BMI

Figure: Beta Posteriors Over Various Quantiles



Section 2

Bayesian Quantile Regression for Longitudinal Data

What is Longitudinal Data?

- Data characterized by repeated measurements on the same subject over time → introduces within-subject dependence/correlation
 - For example: clinical trials, epidemiological studies, etc.
- **Q: How to handle?**
 - **A:** Adding **random effects** to your model can account for over dispersion caused by unobserved heterogeneity or for correlation in longitudinal data.
- **Next:** Extend the ordinary Bayesian quantile regression model from above by introducing two **linear mixed-effects** Bayesian quantile regression models based on work by Luo et al.

Hierarchical Bayesian Quantile Regression Model

- Linear mixed-effects quantile function of the response:

$$Q_{y_{ij}}(\tau | \mathbf{x}_{ij}, \alpha_i) = \mathbf{x}_{ij}^T \beta + \mathbf{z}_{ij}^T \alpha_i,$$

where $Q_{y_{ij}}(\cdot) \equiv F_{y_i}^{-1}(\cdot)$ is the inverse of the cumulative distribution function of y_{ij} conditional on the covariates.

- y_{ij} : response of i th individual measured at time j .
- General hierarchical model

$$y_{ij} \sim \text{ALD}(\mathbf{x}_{ij}^T \beta + \mathbf{z}_{ij}^T \alpha_i, \sigma, \tau)$$

$$\beta \sim \pi(\beta)$$

$$\alpha_i | \Sigma \sim f(\alpha_i | \Sigma)$$

$$\Sigma \sim \pi(\Sigma)$$

$$\sigma \sim \pi(\sigma)$$

Two Proposed Models

Need to compute: $\pi(\beta, \sigma, \Sigma, \alpha | \mathbf{y}) \propto f(\mathbf{y}, \alpha | \beta, \sigma, \Sigma) \pi(\beta) \pi(\sigma) \pi(\Sigma)$, but the posterior densities of both the fixed and random effect parameters are very complicated.

Luo et al. propose two models:

① **BQRMH**: ME Metropolis-Hastings Bayesian Quantile Regression

- Draw posterior samples of β and α via MH, all other parameters can be sampled from well known distributions.
- **Con**: Must tune parameters of the proposal distributions to attain decent acceptance rates. Different for every value of τ !

② **BQRGS**: ME Bayesian Quantile Regression Gibbs Sampler.

- Transforms the likelihood of \mathbf{y} from ALD into a MVN with clever manipulations \rightarrow posteriors are easier to derive with conjugate priors.
- **Plus**: Do not need to deal with any tuning parameters!

Next, we coded up the two algorithms.

Model Comparisons via Simulated Mixed-effects Data

We compare the performance of three Bayesian QR models via a simulated simple linear mixed-effects model.

- ① **OBQR**: Ordinary Bayesian Quantile Regression
- ② **BQRMH**: ME Metropolis-Hastings Bayesian Quantile Regression
- ③ **BQRGS**: ME Bayesian Quantile Regression Gibbs Sampler.

Use the following simple linear mixed-effects model to generate the data:

$$y_{ij} = \beta_0 + \beta_1 x_{ij} + \alpha_{i0} + \alpha_{i1} x_{ij} + \epsilon_{ij}, \quad i = 1, \dots, 20, \quad j = 1, \dots, 5,$$

where $x_{ij} \sim \text{U}(0, 1)$, $\beta = (\beta_0, \beta_1)^T = (1, 5)^T$, $\alpha_i = (\alpha_0, \alpha_1)^T \sim \text{N}_2(\mathbf{0}, \mathbb{I}_2)$, and $\epsilon_{ij} \sim \text{N}(0, 1)$. Assume weak prior information on β :

$$\beta \sim \text{N}_2(\mathbf{0}, 100\mathbb{I}_2).$$

We investigate five different quantiles $\tau = (0.10, 0.25, 0.50, 0.75, 0.90)$.

A Simulated Dataset

example_dataset.png

Model Diagnostics

For a single simulated dataset with 5000 sample draws for $\tau = 0.50$.

trace_ACF.png

Three Algorithms and Their Parameter Estimation

- Obtained via 100 simulated datasets with 5000 sample draws (500 burn-in) for $\tau \in \{0.10, 0.25, 0.50, 0.75, 0.90\}$ for each model.

bias_MSE.png

Longitudinal Studies with Missing Data

A final (third) model: **penalized model** with applications to longitudinal studies with missing data. Want to minimize:

$$\sum_{i=1}^n \sum_{j \in J_{obs}} \rho_{\tau} \left(y_{ij} - \mathbf{x}_{ij}^T \boldsymbol{\beta} - \mathbf{z}_{ij}^T \boldsymbol{\alpha}_i \right) + \frac{1}{2} \sum_{i=1}^n \boldsymbol{\alpha}_i^T \boldsymbol{\Lambda}^{-1} \boldsymbol{\alpha}_i,$$

where $\boldsymbol{\Lambda}$ is a symmetric matrix \rightarrow can be recast into random-effects model.

- Link missing data process with the longitudinal outcome process assuming they share the same random effects $\boldsymbol{\alpha}_i$.
- Adds a new dimension to the likelihood: need to account for likelihood of observing or no observing an observation.
- Need to compute transition probabilities $\pi_{ij}^{(\mathcal{O})}$ and $\pi_{ij}^{(\mathcal{M})}$.

References



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