

# Drink and Thrive? The economics of binge-drinking in Germany

CSSR Final Project

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## 1 Introduction

Alcohol has a history of causing both joy and pain. In our research we want to focus on the latter: Alcohol, if consumed in high amounts and/or regularly over a longer period of time, does severe damage to individuals' health (see e.g. World Health Organization 2014; Miller et al. 2007). It is therefore a policy concern not only to identify reasons and motivation for excessive alcohol consumption but also to develop measures to reduce the public health costs associated to this problem.

It is the aim of this paper is to identify factors interacting with alcohol related health problems and to assess alcohol policy interventions against it. Using the example of all German federal states and time series data for different age groups from 2000 to 2014, we will [1] analyse the possible connection between medical diagnoses of alcohol misuse and possibly explanatory socioeconomic factors like regional economic performance and unemployment rates. For this, we will take both short-term and long-term medical consequences of alcohol misuse into account by examining hospital health records on acute alcohol intoxication as well as on alcoholic liver disease. Besides the socio-economic factors mentioned above, we will [2] test the effect of recent policy measures on the German state level on alcohol consumption, namely the ban on alcohol night sales introduced in 2010 in the state of Baden-Wuerttemberg.

We find no conclusive relationship between alcohol-related hospitalizations and socio-economic indicators. However, the reason for this is not necessarily that the relationship does not exist. It is also possible that

the small number of observations and the limited availability of relevant indicators prevents our data from generating sufficient insights into the matter. Regarding the second topic of investigation, we find that the night sales ban reduces the number of excessive alcohol drinking diagnoses among German youth by about 20%. This finding confirms a prior study by Marcus and Siedler (2015) who found a decrease of about 7%. While their study used more detailed data, it focused only on the short-term effects until 2011. Our study in contrast looked at a longer time-horizon, but used more aggregated data.

## 2 Related Literature

This paper is building on prior research looking at alcohol consumption in the light of socio-economic factors (I) and the effects of policy measures to tackle alcohol misuse (II). The impact of socio-economic factors, e.g. economic performance, for the alcohol abuse was investigated by Popovici and French (2013) and Ettner (1997). Both analysed the effect that a change in employment status has on alcohol consumption of individuals. Focusing on the United States, the two studies found inconsistent results: While one sees “a positive and significant effect of unemployment on drinking behaviors and the findings are robust to numerous sensitivity tests” (Popovici and French 2013) the other argues that “non-employment significantly reduces both alcohol consumption and dependence symptoms, probably due to an income effect.” (Ettner 1997) For the German case, Henkel has been tackling the challenge in extenso (Henkel and Zemlin 2016; Henkel and Schröder 2015; Henkel 2013; Henkel 2000), analysing not only alcohol but also a broad range of other and illegal drugs and their use amongst unemployed and people with low income. People in an economically better situation, he finds, present significantly fewer cases of alcohol misuse and addiction (Henkel and Schröder 2015). An extensive overview on the international literature can be found in Henkel (2011).

In terms of the effectiveness of policy measures to tackle these issues, the present study is inspired by the work of Marcus and Siedler (2015) who analysed “*The Effect of a Ban on Late-Night Off-premise Alcohol Sales on Alcohol-Related Hospital Stays in Germany*.” They find that the introduction of the alcohol night sales ban reduces alcohol-related hospitalizations among adolescents and young adults by about seven percent. In their paper on a similar policy change in the canton of Geneva, Switzerland, Wicki and Gmel (2011) come to a confirming result: The “hospitalisation rate among adolescents and young adults could be reduced by between 25% and 40% due to restricted hours of off-premise sales, and a ban on the sale of alcoholic beverages in gas stations and video stores.” Green, Hollingsworth, and Navarro (2015) found a similar effect in their study on the extension of pub opening hours in England and Wales, introduced in 2005: “increased alcohol availability in England and Wales led to increased consumption, heavy drinking and led to poorer physical and mental health outcomes.” Looking at the effect of Minimum Legal Drinking Ages, beer taxes, and “Zero Tolerance” laws, Carpenter et al. (2007) find that “a variety of types of government intervention have been and can be effective at reducing youth alcohol use.” In their broader, related literature review on alcohol-related harm, Wilkinson, Livingston, and Room (2016) develop a corresponding result stating that “reducing the hours during which on-premise alcohol outlets can sell alcohol late at night can substantially reduce rates of violence. Increasing trading hours tends to result in higher rates of harm, while restricting trading hours tends to reduce harm.”

## 3 Data

Our research health data origins from the German hospital diagnosis statistics from 2000 to 2014, obtained via the Information System of the Federal Health Monitoring (GBE). The data is reported by hospitals to the statistical bureaus of the respective German states and then aggregated by the Statistische Bundesamt Destatis. The data contains aggregated numbers of hospital diagnoses for each German state by age group, gender and year. The diagnoses are published according to the WHO International Statistical Classification of Diseases and Related Health Problems (ICD-10).

The data is gathered via the online-interface of the GBE. Unfortunately, the data provider neither provides an API nor a web-scrappable interface. Therefore we download the base tables manually by searching for the

respective ICD-10 code, using the malleable tables to gather as much information in a single table as possible and then export them.

To identify alcohol-related health problems we use the diagnose category *F10* (Mental and behavioural disorders due to use of alcohol), more specifically: its subdivision *F10.0* (Acute intoxication) and *F10.2* (Dependence syndrome), and *K70* (Alcoholic liver disease). The diagnoses category *F10.0* essentially captures short term effects of excessive drinking, and can be used as an indicator for binge drinking.<sup>1</sup> The categories *F10.2* and *K70* capture long-term effects of drinking as these diagnoses are consequences of regular drinking.

The used data has been downloaded on 01 November 2016 and, due to the data platform's table size limitations, consists of three csv files, one for each diagnose category, containing the German states, the former area of Western Germany as well as Eastern Germany as columns and years (2000-2014), gender (male/female/unknown), and age groups (less than 1 year; 1 to less than 100 years in 5-year groups; older than 100 years).

The advantages of this data is that they reflect a complete survey of all civilian hospitals in Germany for the time-frame of our investigation. Furthermore, the data is not self-reported by patients in interviews but instead consists of the professional diagnoses by a third party, i.e. doctors, which eliminates any problems of possible self-reporting biases. For the case study of Baden-Wuerttemberg's night sale prohibition we do also benefit from the availability of a longer time series as compared to the prior work on the issue by Marcus and Siedler (2015). Finally, the state-year aggregation level does allows us to supplement to our analysis a wide variety of other freely available data, e.g. unemployment rates etc..

However, there are several serious limitations to the data we are using. First of all, the data set can only be exploited in limits with more sophisticated methods like panel data analysis. We only have access to data which is reported annually and on the state level.<sup>2</sup> Hence, for a panel data analysis our scope is limited to 14 years and 16 states, which makes a total of 224 data points for each combination of diagnoses, age group and gender. We still think it is meaningful to work with the data at hand: On the one hand we can employ simpler but still insightful methods like multiple linear regression. On the other hand, our approach can easily be extended to more fine-grained data if such data becomes available to us. A second shortcoming is the lack of further information on the patients beyond their age and gender, e.g. socio-economic characteristics. Factors like available income and medical history will matter for alcohol-related health problems, but for good reasons (data protection) they are not recorded in the respective statistics.

As already mentioned, our analysis will be supplemented by data from other sources. We are using state level indicators. They include the respective population, population density, unemployment rates of different age groups, the state GDP, and beer sales by state. We are aware of the fact that other alcoholic beverages besides beer probably play a role in the cases under scrutiny. However, with these numbers not available, the beer consumption serves as a proxy for general alcohol consumption in the state. We are collecting the supplementary data from different statistics provided by *Destatis*.

## 4 Empirical Strategy

We want to make a two stage analysis of short and long-term alcohol-related hospitalizations (ARH)<sup>3</sup>. We will start with a multiple regression for a better understanding of factors correlated to high rates of hospitalization for alcohol-related reasons. This analysis shall show how socio-economic characteristics of states are correlated to the number of ARH. To keep it simple, the first multiple regression model only includes observations of one year:

$$ARH_s = \alpha_0 + \alpha_1 \cdot GDP_s + \alpha_2 \cdot UR_s + \alpha_3 \cdot PD_s + \epsilon_s \quad (1)$$

<sup>1</sup>In contrast to Marcus and Siedler (2015) and Wicki and Gmel (2011) we do not include T51 (Intoxication due to alcohol) as its covers the consumption of pure ethanol and its number are relatively small.

<sup>2</sup>In principle, there is more fine-grained data available, even down to the individual level. However, the access to the data requires using the paid services by forschungsdatenzentrum.de which is out of our students' budget.

<sup>3</sup>ARH is normalized by the state population and denoted in hospitalizations per 100,000 inhabitants.

where  $GDP$  stands for the GDP level,  $UR$  for the unemployment rate and  $PD$  for the population density. All variables are recorded on the state level as indicated by the indices. The regression will be conducted for the aggregated number for all age groups. We will run this first regression separately for the years 2000, 2007, and 2014.<sup>4</sup>

Based on the idea that individuals adapt to their circumstances and are more reactive to changes than to actual levels we will then analysis ARH difference over time:

$$\Delta ARH_{s,t} = \beta_1 \cdot \Delta GDP_{s,t} + \beta_2 \cdot UR_{s,t} + \epsilon_{s,t} \quad (2)$$

In contrast to the multiple regression on levels, the difference approach highlights how *changes* in hospitalizations are correlated to *changes* in the socio-economic factors used as independent variables. We eliminated the population density because its changes are not fine-grained recorded and the Zensus 2011 led to a jump due to measurement adjustments in the data series. To check for further robustness of the results, we are comparing short-term ARH (F10.0), which should be affected by short-term changes and long-term ARH (K70/F10.2), which should not be affected.

In the second step of our analysis, we are looking at the effect of the night sales ban on alcohol in Baden-Wuerttemberg. To measure the effect of the ban we use the difference-in-difference (DD) approach. As the ban was only introduced in Baden-Wuerttemberg, this state will be the treatment group. All other states are in the control group. This distinction can be justified by the fact that most alcohol regulation is done on the federal level with the exception of sales hour regulation and campaigns.<sup>5</sup> The ban was introduced in 2010, which is captured by the treatment dummy. We further assume there are no dynamic effects, which is reasonable as patients are registered when their treatment begins:<sup>6</sup>

$$ARH_{s,t} = \gamma_0 + \gamma_1 \cdot dBW_s + \gamma_2 \cdot dPOST_t + \gamma_4 \cdot dBAN_{s,t} + \epsilon_{s,t} \quad (3)$$

where  $dBW$  is a dummy variable to indicate the treatment group (here the state of Baden-Wuerttemberg),  $dPOST$  a time dummy indicating the post-treatment period and  $dBAN$  a dummy for the night-sale-prohibition (= 1 for Baden-Wuerttemberg from 2010 on). Since we are re-engineering Marcus and Siedler (2015) here, we will limit the DD application to ARH by young people (15-25). To check for further robustness of the results, we use different control group compositions: All states but Baden-Wuerttemberg and all territorial states, excluding the city states Berlin, Bremen and Hamburg.<sup>7</sup>

As the DD approach crucially depends on the common trend assumption, it is advisable to include control variables to capture possible trends affecting treatment and control group differently. We control for youth unemployment ( $YUR$ ) and  $GDP$  as a proxy variable for the economic situation.

$$ARH_{s,t} = \delta_0 + \delta_1 \cdot dBW_s + \delta_2 \cdot dPOST_t + \delta_3 \cdot dBAN_{s,t} + \delta_4 \cdot YUR_{s,t} + \delta_5 \cdot GDP_{s,t} + \epsilon_{s,t} \quad (4)$$

Finally, we are refining our DD approach further by turning it into a panel and allowing for heterogeneity in the control group. This is the final version of the model, but we expect it to generate less insightful results as we (with all states separately) have comparably few data points (due to comparably little money to invest in data...).

$$ARH_{s,t} = \theta_0 + \theta_1 \cdot dBAN_{s,t} + \theta_2 \cdot YUR_{s,t} + \theta_3 \cdot GDP_{s,t} + c_s + z_t + \epsilon_{s,t} \quad (5)$$

<sup>4</sup>We also computed all other years to check robustness. The results can be found in the appendix.

<sup>5</sup>The night-sale ban in Baden-Wuerttemberg has been the most prominent alcohol policy in 2010. There have also been general changes to opening hours as during this time the competency was transferred to the states and media campaigns. As we only have annual data we are not fully able to distinguish between these treatments, but we assume that they are outweighed by the night-sales ban.

<sup>6</sup>Even if they were registered in consecutive reporting periods, the stay of patients for short-term ARH (F10.0) would be in average 2,1 days and for long-term ARH 10,7 days (K70) and 11,4 days (F10.2) in 2014, making the dynamic effect insignificantly small.

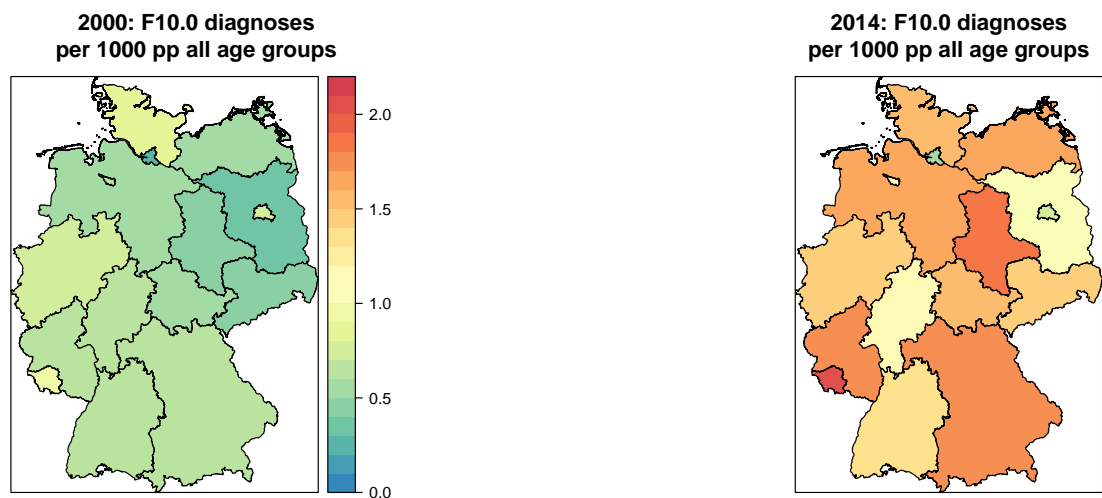
<sup>7</sup>The hospitalization statistics has revealed that city states followed a slightly different trend than the territorial states, see section data overview.

where  $c_s$  now is capturing unobservable and time-invariant heterogeneity of states. In addition, we are replacing  $dPOST$  with  $z_t$  to capture time fixed effects affecting all states, e.g. a change in a federal law. This model can be estimated by using dummies for the time-periods and first-differencing or fixed-effects to address the unobserved heterogeneity of states.

## 5 Data Overview

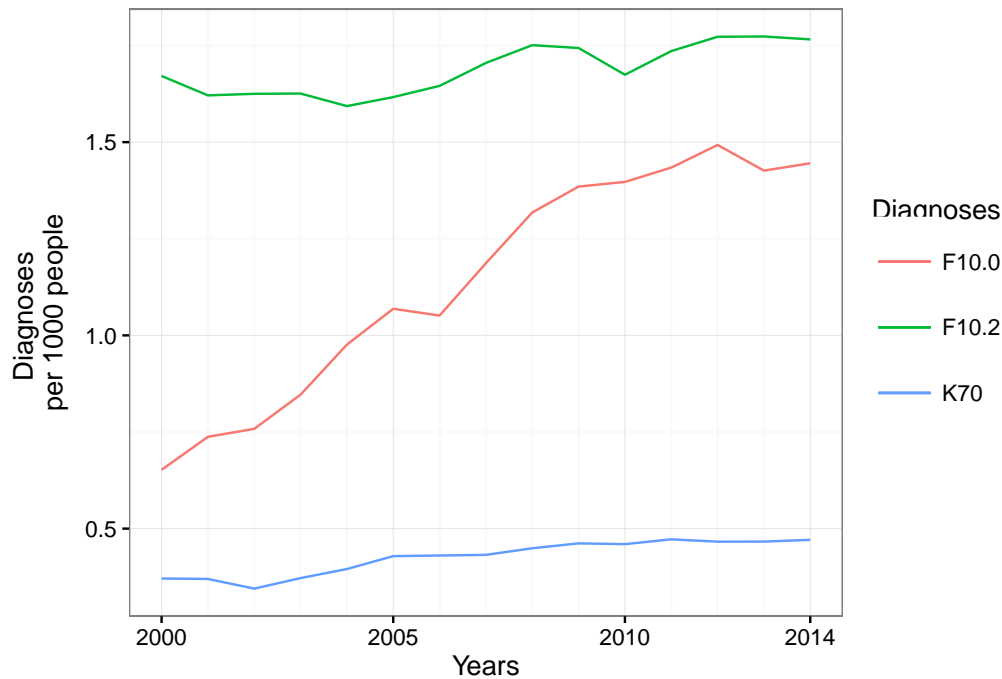
Looking at the dependend variables, especially the short-term alcohol related hospitalizations (ARH-ST) show a clear upward trend over time. The numbers of acute intoxication per 1000 people only moderately increase in city states like Berlin (BE), Bremen (HB), and Hamburg (HH). Schleswig-Holstein and other states, on the contrary, double their rates in the time span from 2000 to 2014.

The development in German states from the beginning of the present dataset to the end (2000 - 2014) shows that apparently only the big city states have been able to retain their relatively low level of short-term ARH.



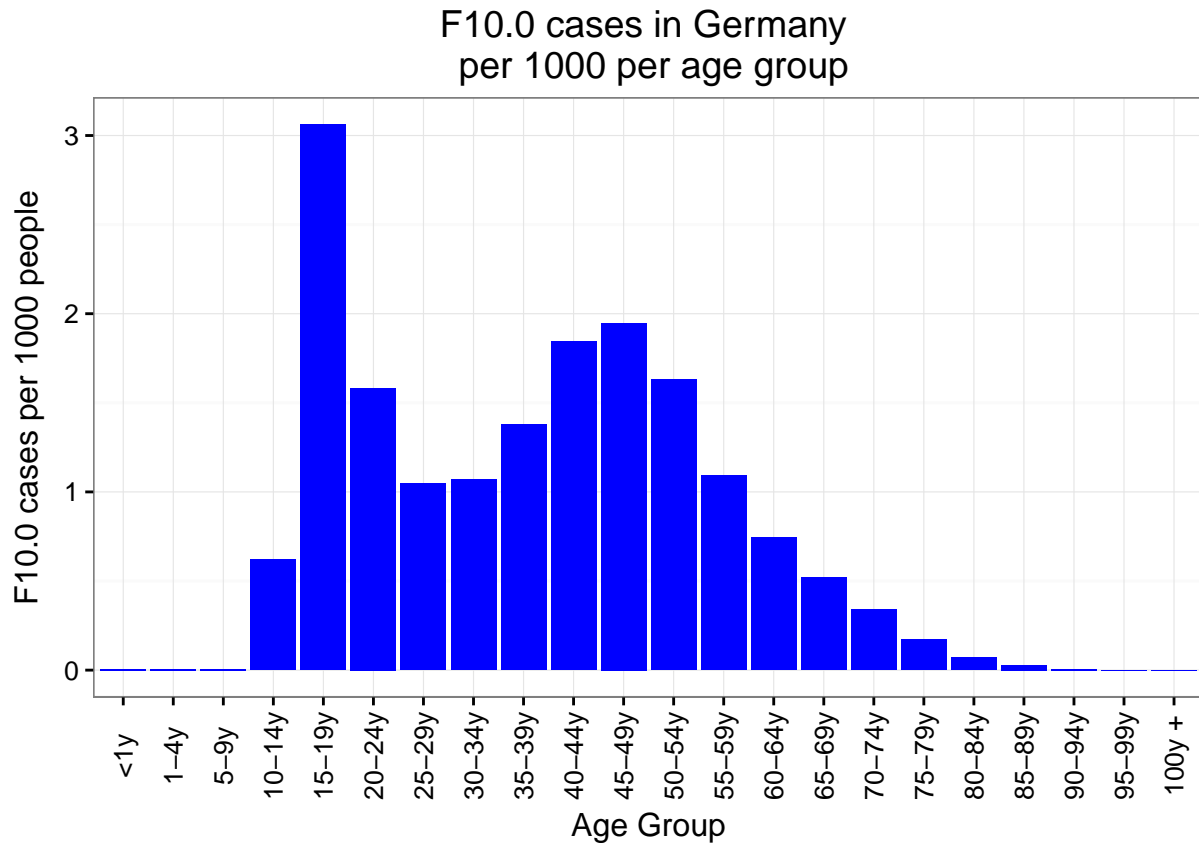
The long-term ARH values also slightly increase, but this increase is a rather moderate one. This is a reasonable finding, as already the description “long-term” suggests that any change is very unlikely to happen in a short period of time like the one captured by our sample.

F10.0, F10.2 and K70 diagnoses in Germany 2000–2014

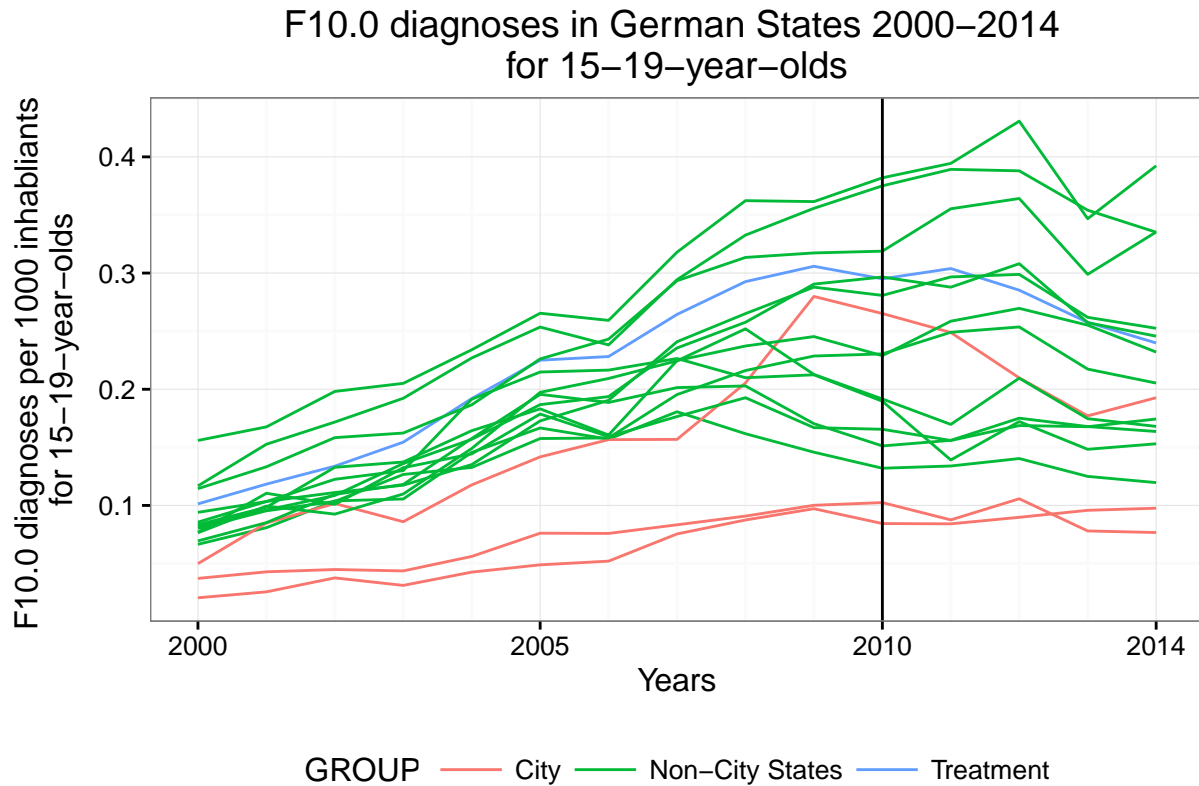


With regards to the different age groups, the distribution of cases is already somewhat revealing: long-term ARH numbers as an indicator for strong alcoholism show that it rather is a problem for best-agers than for younger people. The numbers begin to rise at the 35-39 years age group and start declining with the 65-69 years group – probably due to a higher mortality (although the data does not offer such interpretation).

In Germany, the two age-thresholds for buying alcoholic beverages are 16 years (for beer and wine) and 18 years (for brandy). This helps with the interpretation of the short-term ARH bar plot with its two peaks: The first is a very strong one at the 15-19 years age group. It appears reasonable to interpret this one as a sign of youngsters making first experiences with alcoholic drinks – and overdoing it. The second “mid-life crisis” peak corresponds with the peak of long-term alcohol abuse shown before.



An alarming development is illustrated by the high increase in 15-19-year-olds admitted to hospitals with acute alcohol intoxication. From 2000 to 2011, Bavaria more than tripled its quota in this regard. The vertical line is indicating the year of the night sale introduction in Baden-Wuerttemberg.



## 6 Results

The first model shown in *Table 1* is simple and its results are indicative, at most. It does reveal to which state characteristics short-term alcohol related hospitalizations (ARH-ST) are correlated. We find that ARH-ST are higher in states with a lower GDP. This captures the basic intuition that binge drinking is more common in regions that are economically less well-off. But this view is not confirmed by our second indicator for economic well-being: the unemployment rate. To the contrary, states with higher unemployment rates have less cases of binge drinking. These results are robust over all years of our investigation. But this is not surprising as the variables we are considering are not strongly fluctuating during the period under investigation. Most of them follow more or less the same trend. As the contradicting interpretation already hints at, we are most likely not seeing what is happening below the aggregated surface.

The major drawback of this regression is the low number of observations. With one observation per state the standard errors are *huge*. In the second step, we increased our data set by looking at difference over time, which increases the number of available observations to 224 (*Table 2*).

The difference show a correlation between GDP growth and short-term ARH. A superficial interpretation could be that individuals in years of strong growth tend to party more. But again, this should be seen with care. For short-term ARH the other variables are not significant. The effects of changes on long-term ARH are unsurprisingly inconclusive. In our case each coefficient is directed exactly in the opposite direction of our two long-term ARH indicators. But a clear result would again have been implausible, as long-term ARH result from excessive drinking behaviour over years. Short-term macro developments should not have any impact on this.

The next step in refining our model (*Model 3*) takes us to analyzing a policy measure, namely the introduction of a ban on night sales of alcohol in the state of Baden-Wuerttemberg. Contrary to the prior analysis, we



Table 1: Regression results for Model 1 for 2000, 2007 and 2014

	<i>Dependent variable:</i>		
	F10.0 Diagnoses per 1000 capita 2000	2007	2014
GDP per capita	−0.02** (0.01)	−0.01 (0.01)	−0.01 (0.01)
Unemployment rate	−0.04*** (0.01)	−0.02 (0.03)	−0.03 (0.04)
Population density	0.0001* (0.0001)	−0.0002 (0.0001)	−0.0002* (0.0001)
(Intercept)	1.63*** (0.31)	1.95*** (0.58)	2.23*** (0.59)
Observations	16	16	16
R <sup>2</sup>	0.49	0.52	0.59
Adjusted R <sup>2</sup>	0.36	0.40	0.48
Residual Std. Error (df = 12)	0.15	0.28	0.28
F Statistic (df = 3; 12)	3.79**	4.35**	5.64**

*Note:* \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

Table 2: Regression results for Model 2 with first differenced data

	<i>Dependent variable:</i>		
	Change in F10.0 F10.0	Change in F10.2 F10.2	Change in K70 K70
GDP Change	0.02*** (0.01)	−0.01 (0.01)	0.002 (0.002)
Unemployment Change	−0.01 (0.01)	−0.03** (0.01)	0.002 (0.003)
Observations	224	224	224
R <sup>2</sup>	0.09	0.02	0.004
Adjusted R <sup>2</sup>	0.08	0.02	−0.005
Residual Std. Error (df = 222)	0.11	0.18	0.04
F Statistic (df = 2; 222)	10.88***	2.72*	0.48

*Note:* \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

now look at the age group of 15 to 19 year olds. The results are shown in *Table 3*. While we are introducing a dummy for the state in consideration, one for the post-treatment period, and the interaction term of both, the only significant variable is the post-treatment dummy. Even more counterintuitively, the post-treatment dummy has is *positively* correlated with short-term ARHs. This hints to the fact that we might witness a common trend over time, increasing short-term ARHs, that is similar for all states and that is currently captured by the time dummy. That would explain why the time dummy (which essentially splits the whole set of observations into two consecutive phases) is highly significant but neither the state dummy nor the interaction term are of any significance.

Table 3: Model 3 and 4 - simple differences-in-differences without and with controls

	<i>Dependent variable:</i>			
	F10.0 cases per 1000 PP among 15-19 year olds			
	(3) All states	(3) non-city states	(4) All states	(5) non-city states
BW night sale ban	0.04 (0.03)	0.02 (0.02)	-0.03 (0.02)	-0.04 (0.02)
Post-ban dummy	0.06*** (0.01)	0.07*** (0.01)	0.03** (0.01)	0.03** (0.01)
GDP per capita			-0.003*** (0.001)	0.01*** (0.001)
Youth unemployment			-0.01*** (0.001)	-0.002 (0.002)
Interaction dummy	0.01 (0.05)	0.01 (0.04)	0.04 (0.04)	0.003 (0.04)
(Intercept)	0.16*** (0.01)	0.18*** (0.01)	0.40*** (0.03)	0.06 (0.06)
Observations	255	210	240	195
R <sup>2</sup>	0.14	0.18	0.38	0.40
Adjusted R <sup>2</sup>	0.13	0.16	0.36	0.39
Residual Std. Error	0.08 (df = 251)	0.07 (df = 206)	0.07 (df = 234)	0.06 (df = 189)

*Note:*

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01

In case such a common trend is an economic one, we can capture it and take it out of the diff-in-diff model by including two more control variables that not only allow for development over time but which also capture varying developments among states. *Model 4* therefore includes state GDP per capita and the youth employment rate on state level (*Table 3*).

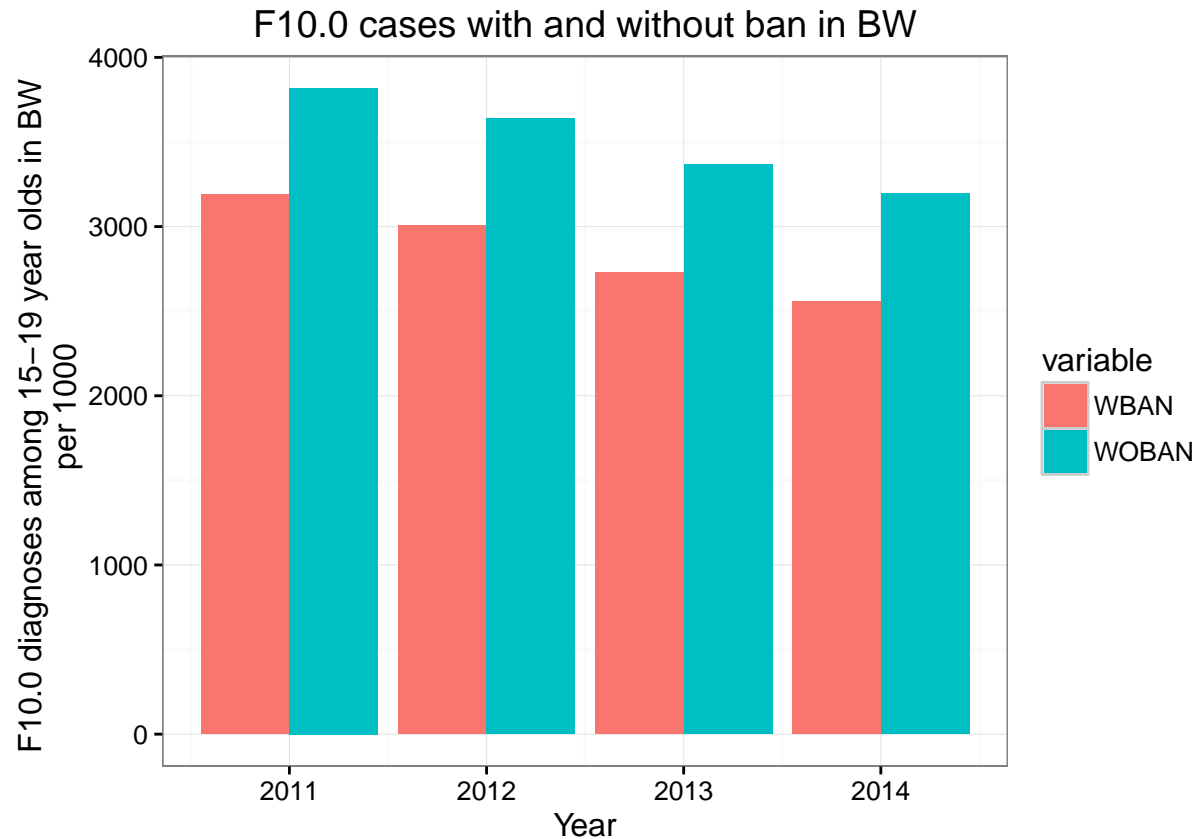
The results do not have any effect on the significance of treatment-related dummy variables. Also, the “wrong” direction of the post-treatment dummy’s effect stays. However, the two newly introduced control variables are not only highly significant. In the case of GDP per capita, the effect at first sight is also pointing into a plausible direction: Increasing GDP per capita correlates with decreasing short-term ARHs. The reasoning behind this might be that in times of economic prosperity, alcohol as remedy for despair is not needed. But the youth unemployment rate points into the opposite direction, with increasing youth unemployment rates being correlated with decreasing ARHs. An explanation for this effect might be that for young people unemployment comes with more severe financial constraints than it comes for older members of the workforce, therefore quicker closing the tap for younger people.

But these explanations are tentatively, at best. Therefore, it is advisable to take the final step in refining the model by introducing a panel data approach which best fits the fact that we are analysing 15 years of the very same sample of German states.

Table 4: Model 5 - multi-period panel differences-in-differences

	<i>Dependent variable:</i>		
	All age/states	F10.0 cases per 1000 PP 15-19y all states	15-19y + non-city state
BW night sale ban	-0.06 (0.09)	-0.03* (0.02)	-0.06*** (0.01)
Youth unemployment rate	-0.02** (0.01)	0.02*** (0.002)	0.02*** (0.002)
GDP per capita	-0.01 (0.02)	0.003 (0.003)	0.01*** (0.003)
Observations	240	240	195
R <sup>2</sup>	0.03	0.34	0.49
Adjusted R <sup>2</sup>	0.02	0.29	0.42
F Statistic	2.00 (df = 3; 207)	35.01*** (df = 3; 207)	53.28*** (df = 3; 165)
<i>Note:</i>		*p<0.1; **p<0.05; ***p<0.01	

The final regression indicate a negative impact of the ban on ARH related short-term hospitalizations. The estimated relation of introducing a night-sales ban is a decrease of 0.06 hospitalizations per 1000 inhabitants, which makes roughly 600 cases less per year based on the population of Baden-Wuerttemberg. This would imply a decrease of 20%, which is more than Marcus and Siedler (2015) found (7%). To illustrate the result, we calculated the hypothetical number of hospitalizations without the ban:



But again the result should be treated with care. The calculations are based on 195 observations, which are due to reduced control group to territorial states. Moreover, the exact estimate is sensitive to adjustments in the control group. However, the general direction and the magnitude are meaningful and point in the same direction as Marcus and Siedler (2015).

## 7 Conclusion

Overall, our research supports the general conclusion of the literature on drinking and socio-economic factors. In particular, the most sophisticated approach (model 5) we used finds that higher youth unemployment is correlated with more alcohol-related diagnoses among young people and that the night sales ban on alcohol has a measurable negative effect on binge-drinking. But due to the limitations of our data this finding is not fully consistent over the models we used. However, this inconsistency should not be overevaluated as it mostlikely steams from the limitations of the used model and the limited data available.

As our study reengenerated Marcus and Siedler (2015) the confirmation of their general result deserves particular attention. While their study used more fain-grained data, our research included a longer treatment horizon. Hence, the effect (???) find prevailed over the time horizon of their research according to our findings.

The research could be further improved by using more diaggregated data, in particular by using monthly numbers and the county level. This would increase the number of observations tremendously and allow therefore for more substantial results. It would also allow the inclusion of more controls as in we only relied on a small number of socio-economic controls. Many more controls exist, but we did not include them as they are not available as a time-series and on the regional level.

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