

ASEN 5044, Fall 2018

Statistical Estimation for Dynamical Systems

Lecture 27: Steady state KF Behavior; KF Consistency Evaluation

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Thurs 11/8/2018

Announcements

- **HW 7 due today**
- **Midterm 2: posted, due Nov 15 at 11 am on Canvas**
 - Will focus on HWs 5-7 (HW 7 solutions posted)
- **Final project partner sign up sheet on “Assignments” tab in Canvas**
 - Google docs sheet (**Due: Wed 11/14**) – please read + follow all instructions!!
 - Folks with a partner: enter group names
 - Folks without partner: start a new group or email each other to find a match
 - Preview system descriptions posted (will pick for HW 8 and final assignment)
 - **You must stay in your group for HW 8 (group hw) and final assignment!!!**
(group gets same grade on both!)

Final Spec. topic lecture tomorrow.

- **No Lecture on Tues 11/13, but will have class on Thurs 11/15**

Last Time...

- **The Kalman Filter (KF):** dynamic predictor-corrector state estimator
(“dynamic RLLS”: combine state prediction with RLLS updates at each time k – i.e. handle dynamics + process noise w_k + noisy measurements y_k)
 - Algorithm
 - Example
 - Important/useful properties of the Kalman Gain

Today...

- One more generally useful property of KFs: **steady state behavior**
 - Riccati equation, steady state error covariance, and steady state Kalman gain
- **How to tell if your (linear) KF is actually working correctly???**
 - Want to avoid GIGO systems (Garbage Input, Garbage Output)



- **KF dynamic consistency analysis and “Truth Model Testing” (TMT)**
- **Chi-square tests (NEES/NIS)** – check if KF’s state errors/measurement residuals make sense for given system + measurement + noise models
 - **Do actual state errors/meas. residuals agree with KF’s estimated error covariances?**
 - Formal statistical tests to examine this question

Stopped here 11/6/18.

Steady-state Properties of the KF

- KF gives state estimate along with update of estimation error covariance (P_k^+)
- What is the "smallest/best possible" covariance? (i.e. what is smallest possible cost $J(k) = \text{tr}(P_k^+)$?)
(How much "juice" can be squeezed from yik?)

Recall: KF pred: $P_{k+1}^- = F P_k^+ F^T + Q$

KF meas. update: $P_{k+1}^+ = P_{k+1}^- - P_{k+1}^- H^T [H P_{k+1}^- H^T + R]^{-1} H P_{k+1}^-$

→ but: @ time k : $P_k^+ = P_k^- - P_k^- H^T [H P_k^- H^T + R]^{-1} H P_k^-$
Sub in def of KF gain K_{k+1}

→ so plug P_k^+ expansion into P_{k+1}^- equation:

$$P_{k+1}^- = F [P_k^- - P_k^- H^T [H P_k^- H^T + R]^{-1} H P_k^-] F^T + Q$$

→ Non-linear DT "one-step" update to get P_{k+1}^- from P_k^- (assuming meas. update occurs @ time k)
known as → DT Matrix Riccati Equation (MRE)

The Algebraic Riccati Equation (ARE)



Jacopo Riccati (1676 - 1754)

- Special case for LTI DT systems:
 - if process noise $w(k)$ hits every state
 - AND if (F, H) is observable

then DT MRE implies convergence to a steady state a priori $P_{\infty}^{-} > 0$ (posdef)

AND obtain the **DT Algebraic Riccati Equation (ARE)**, which is easier to solve (though still non-linear in P_{∞}^{-}):

$$P_{\infty}^{-} = F(P_{\infty}^{-} - P_{\infty}^{-}H^T[HP_{\infty}^{-}H^T + R]^{-1}HP_{\infty}^{-})F^T + Q$$

Algebraic Riccati Equation (ARE)

Handwritten scribble in purple ink.

Steady-State KF Gain

- Suppose the conditions for the ARE hold, so that

$$P_{\infty}^{-} = F(P_{\infty}^{-} - P_{\infty}^{-}H^T[HP_{\infty}^{-}H^T + R]^{-1}HP_{\infty}^{-})F^T + Q$$

- Since the Kalman gain is

$$K_{k+1} = P_{k+1}^{-}H^T[HP_{k+1}^{-}H^T + R]^{-1}$$

→ it follows that there must also exist a steady state Kalman gain K_{∞}

$$K_{\infty} = P_{\infty}^{-}H^T[HP_{\infty}^{-}H^T + R]^{-1}$$

→ K_{∞} often used in practice to save computation at each time step

(since $K_{k+1} \rightarrow K_{\infty}$ quickly anyway, there is generally little performance loss)

*In Matlab: can use the “**dlqe.m**” (discrete linear quadratic estimator) command to find steady state KF gain, along with steady state a priori and a posteriori covariances

$$[K_{\infty}, P_{\infty}^{-}, P_{\infty}^{+}] = \text{dlqe}[F, \text{eye}(n), H, Q, R]$$

→ $tr(P_{\infty}^{+}) = \text{smallest possible cost } J(k)$

Example: 1D Robot: Part Trois: Régime Permanent

- Same DT model as before:

$$x(k) = [\xi(k), \dot{\xi}(k)]^T$$

$$u(k) = 2 \cos(0.75t_k) \text{ (ZOH)}$$

$$w(k) \sim \mathcal{N}(0, Q) \text{ (AWGN)}$$

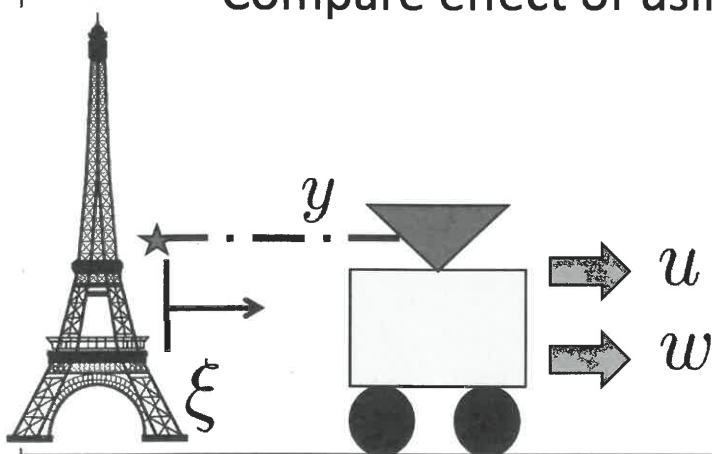
$$x(0) \sim \mathcal{N}(\mu_0, P_0), \text{ where } \mu_0 = [0, 0]^T, P_0 = I_{2 \times 2}$$

$$x(k+1) = Fx(k) + Gu(k) + w(k)$$

$$F = \begin{bmatrix} 1 & \Delta t \\ 0 & 1 \end{bmatrix} \quad G = \begin{bmatrix} 0.5\Delta t^2 \\ \Delta t \end{bmatrix} \quad \Delta t = 0.1 \text{ sec}$$

$$Q = 1 \text{ (m/s)}^2, R = 0.5 \text{ m}^2$$

- Compare effect of using dynamic KF gain to steady-state KF gain



$$\rightarrow K_{\infty} = \begin{bmatrix} 0.1548 \\ 0.1300 \end{bmatrix}$$

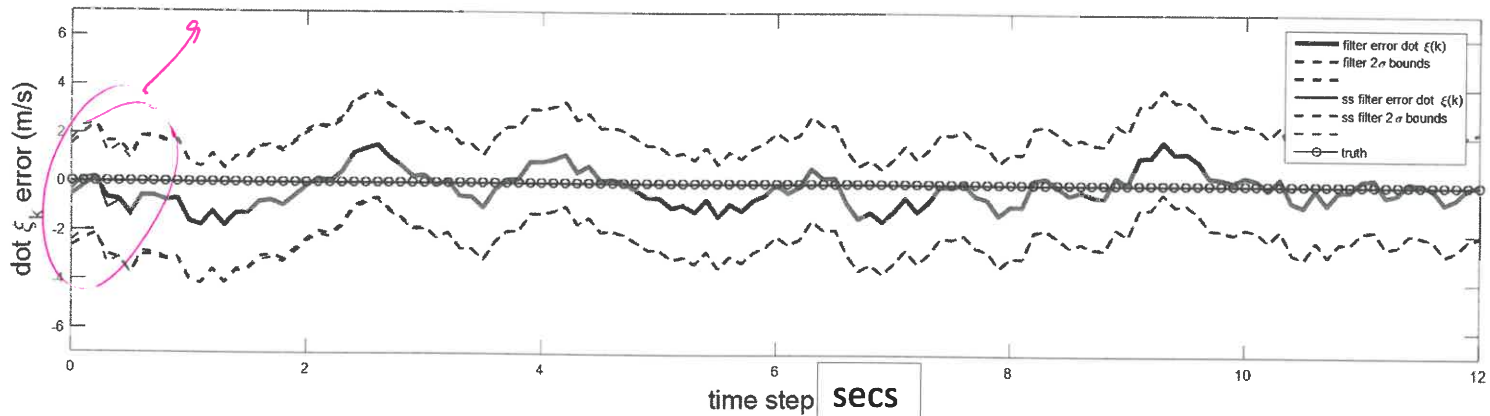
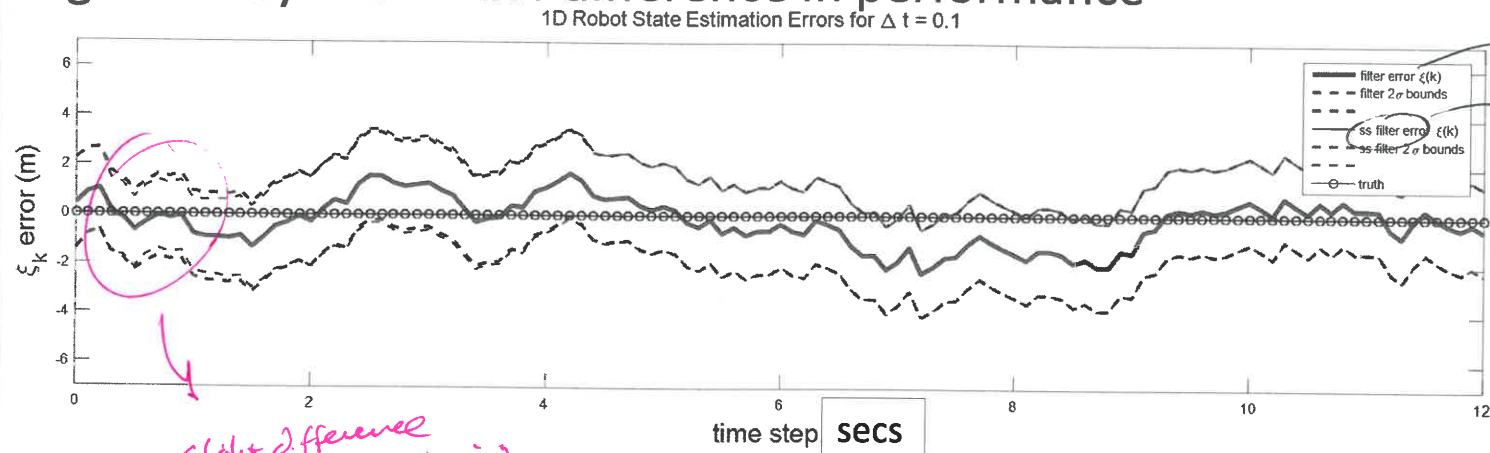
$$P_{\infty}^{-} = \begin{bmatrix} 0.9157 & 0.7691 \\ & 1.2406 \end{bmatrix}$$

$$P_{\infty}^{+} = \begin{bmatrix} 0.7740 & 0.6501 \\ & 1.1406 \end{bmatrix}$$

computed w/dlqe

Results: Dynamic KF Gain vs. Steady-state KF Gain

- Not a significantly noticeable difference in performance



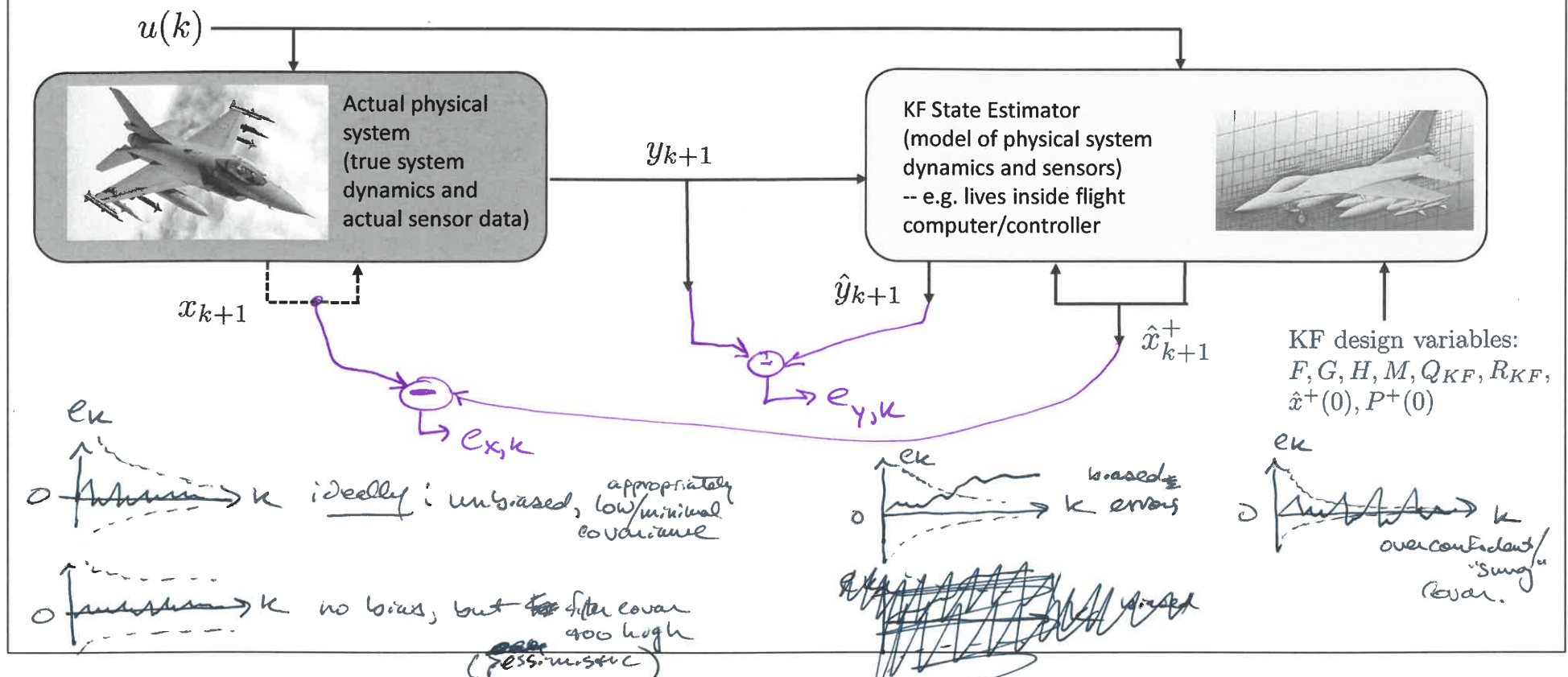
usual time varying KF gain
steady state KF gain

KF Consistency Analysis

- Like batch LLS and RLLS: estimates produced by the KF are random vectors
- For KF: this is due to uncertainties from process and measurement noise
- KF recursively assesses estimation uncertainty via error covariance matrix
- BUT: first need to set tuning parameters in Q_{KF} – generally not obvious!
- Other possible latent practical issues: approximate (F,G,H,M) model, unmodeled state dynamics, non-white noise, ...
- So: how to know if KF estimates & covariances are correct for given system?
i.e. do error statistics provided by KF reflect the actual error statistics?

Evaluating Errors in a KF State Estimator

- KF is itself a dynamical system (defined in software) that tries to track an actual physical system:



What Does “Minimal Error Estimation” Mean for the KF?

- In a perfect universe, we would like estimation errors to eventually vanish completely, i.e.

if $e_{x,k} = x_k - \hat{x}_k^+$, then $E[e_{x,k}] = 0$ and $P_k^+ = E[e_{x,k}e_{x,k}^T] = 0$ as $k \rightarrow \infty$

(ie ideally: we would like perfect certainty in x_k as more measurements obtained)

- BUT, THIS DOES NOT HOLD AS EVIDENCED BY THE FACT THAT (IN MOST CASES) $P_\infty^- \neq 0$

(The Matrix Riccati & Algebraic Riccati Equations say so!!!)

- What is responsible for this?: Random Process Noise inputs incessantly disturb the true state $x_k \rightarrow \therefore e_{x,k} \not\rightarrow 0$ as $k \rightarrow \infty$!
(generally speaking)

“Proper KF Error Characteristics”: Dynamic Filter Consistency

- Because some finite/non-zero error will always exist, we instead say that the KF is “working properly” (for a given DT state space model and noise specs) if it satisfies the following **3 demands for dynamic filter consistency**:

1) Unbiasedness: $E[e_{x,k}] = 0$ for all k

2) Efficiency: $E[e_{x,k}e_{x,k}^T] = P_k^+$ (true state errors match filter covariance)

3) KF measurement residuals/innovations are a white Gaussian sequence:

$$e_{y,k} \sim \mathcal{N}(0, S_k), \quad E[e_{y,k}e_{y,j}^T] = S_k \cdot \delta(k, j)$$

where $e_{y,k} = y_k - \hat{y}_k = y_k - H\hat{x}_k^-$,

$$S_k = HP_k^-H^T + R$$

(S_k : “innovation covariance matrix”
 $\in \mathbb{R}^{p \times p}$)

Analyzing KF Estimation Errors and Measurement Innovations

- How do we analyze the two types of random error vectors in a KF?

- State estimation errors (w.r.t. ground truth x_k): $e_{x,k} = x_k - \hat{x}_k^+ \in \mathbb{R}^n \sim \mathcal{N}(0, P_k^+)$
- Measurement innovations/residuals (w.r.t. observations y_k): $e_{y,k} = y_k - \hat{y}_k^- \in \mathbb{R}^p \sim \mathcal{N}(0, S_k)$

- Simplest check is to look at the normalized magnitude of these vectors over time:

⊗ $\epsilon_{x,k} = e_{x,k}^T (P_k^+)^{-1} e_{x,k} \rightarrow$ Normalized estimation error squared (NEES) at time k

⊗ $\epsilon_{y,k} = e_{y,k}^T (S_k)^{-1} e_{y,k} \rightarrow$ Normalized innovation squared (NIS) at time k

⊗ \rightarrow NEES and NIS are both **positive random scalars**: (weighted 2-norms)² of $e_{x,k}$ and $e_{y,k}$!

Key question: if dynamic consistency conditions (1)-(3) on previous slide hold, then what pdfs ought to describe our expected NEES & NIS outcomes?

ie what are
 $p(\epsilon_{x,k})$ & $p(\epsilon_{y,k})$?

Useful Fact #1: Squared 2-norms of Gaussian Random Vectors

- Suppose we are given some random vector $e \sim \mathcal{N}_e(0, \mathcal{P}_e)$, $e \in \mathbb{R}^n$
 - let $Z Z^T = (\mathcal{P}_e)^{-1}$, where $Z \in \mathbb{R}^{n \times n}$ is Cholesky factor of $(\mathcal{P}_e)^{-1}$ [matrix square root]
 - Also define: $g = Z^T e \in \mathbb{R}^n$ (linear transformation of e via Z^T)
 - Now since $e \sim \mathcal{N}_e(0, \mathcal{P}_e)$
 - & since $\mathcal{N}_e(0, \mathcal{P}_e) \stackrel{\text{pdf def.}}{=} (\text{const.}) \cdot \exp(-\frac{1}{2} e^T (\mathcal{P}_e)^{-1} e)$
 - (apply Chol. fact.) $= (\text{const.}) \cdot \exp(-\frac{1}{2} e^T [Z Z^T] e)$ (apply def. of g) $= (\text{const.}) \cdot \exp(-\frac{1}{2} g^T g) = \mathcal{N}_g(0, I)$
 - it follows that $g \sim \mathcal{N}_g(0, I)$ [std. multivariate normal pdf] \Leftrightarrow "decorrelated"/"whitened" Gaussian random vector
 - $g_i \sim \mathcal{N}(0, 1)$ for $i=1, \dots, n$
- ⇒ so $\epsilon = e^T (\mathcal{P}_e)^{-1} e = e^T [Z Z^T] e = g^T g = \|g\|^2 = \sum_{i=1}^n g_i^2 = \text{Sum of squares of } n \text{ i.i.d. Gaussian random scalar vars!}$
- ↑ must have the same pdf!

$$P(\epsilon) = P\left(\sum_{i=1}^n g_i^2\right)$$

Example:

Correlated and Decorrelated Gaussian Random Vectors

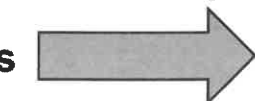
- Apply vector transform via Cholesky decomposition

$$e \sim \mathcal{N}(0, P_e) \quad P_e = \begin{bmatrix} 3 & 2 \\ 2 & 6 \end{bmatrix}$$

$$ZZ^T = P_e^{-1}$$

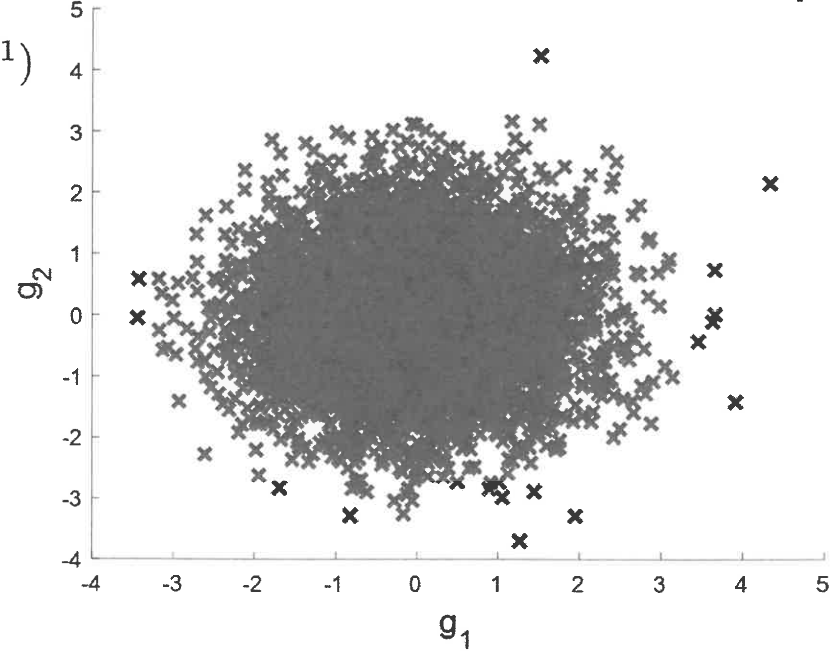
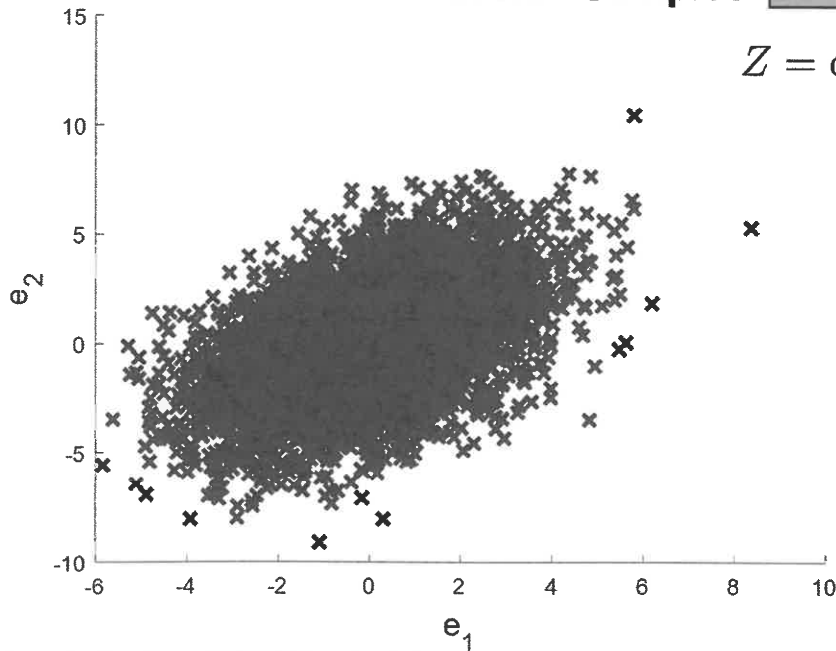
$$g = Z^T e, \text{ where } g \sim \mathcal{N}(0, I)$$

Correlated Gaussian Samples



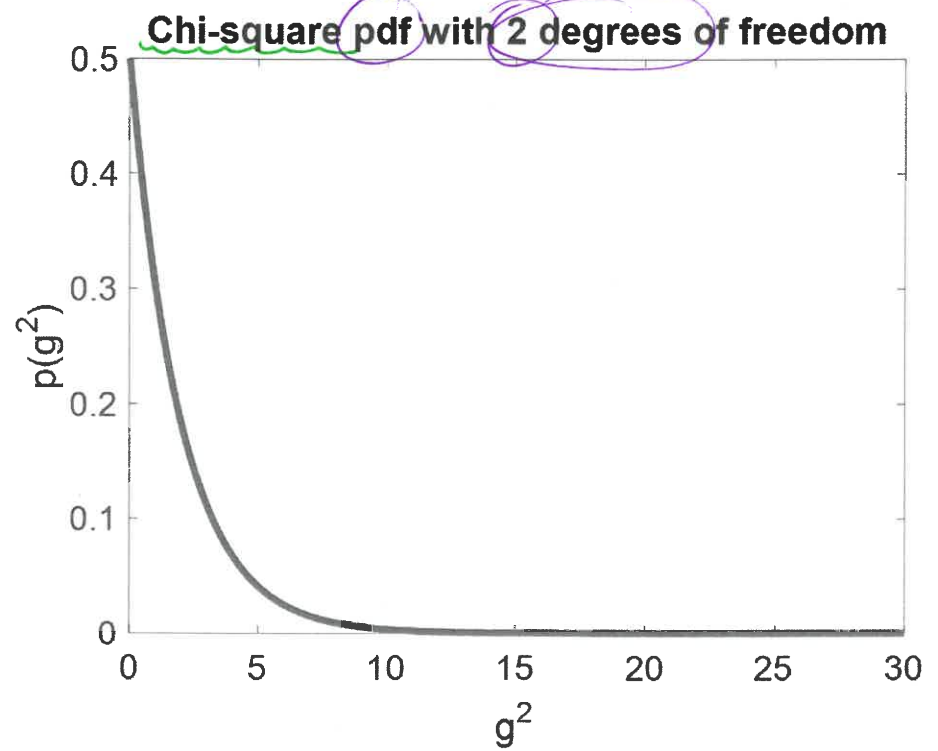
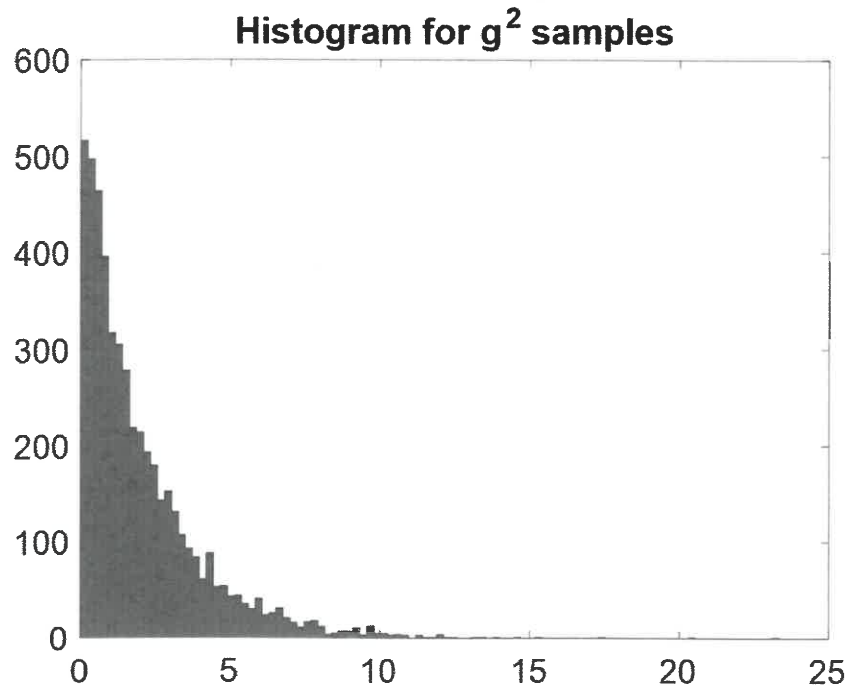
Decorrelated 'Whitened' Gaussian Samples

$$Z = \text{chol}(P_e^{-1})$$



Example: Distribution of Gaussian RV Squared Magnitudes

- What does pdf look like for $g^T g = [\text{norm}(g)]^2 = g_1^2 + g_2^2$?



Useful Fact #2: The Chi-square Distribution

Suppose we have scalar i.i.d. random variables g_1, \dots, g_n where $g_i \sim \mathcal{N}(0, 1)$ for $i = 1, \dots, n$.

Define: random variable $q = \sum_{i=1}^n g_i^2 = \vec{g}^T \vec{g}$, where $\vec{g} = [g_1, \dots, g_n]^T$ (note: $\vec{g} \sim \mathcal{N}(\vec{0}, I_{n \times n})$)

\Rightarrow then the pdf $p(q)$ is a **chi-square** (χ^2) distribution with n degrees of freedom:

$$p(q) = \chi_n^2 = \begin{cases} \frac{1}{2^{\frac{n}{2}} \Gamma(\frac{n}{2})} q^{\frac{n-2}{2}} \cdot \exp(-\frac{q}{2}), & \text{for } q \geq 0 \\ 0, & \text{for } q < 0 \end{cases}$$

where the 'gamma function' is

$$\Gamma(\frac{1}{2}) = \sqrt{\pi}$$

$$\Gamma(1) = 1$$

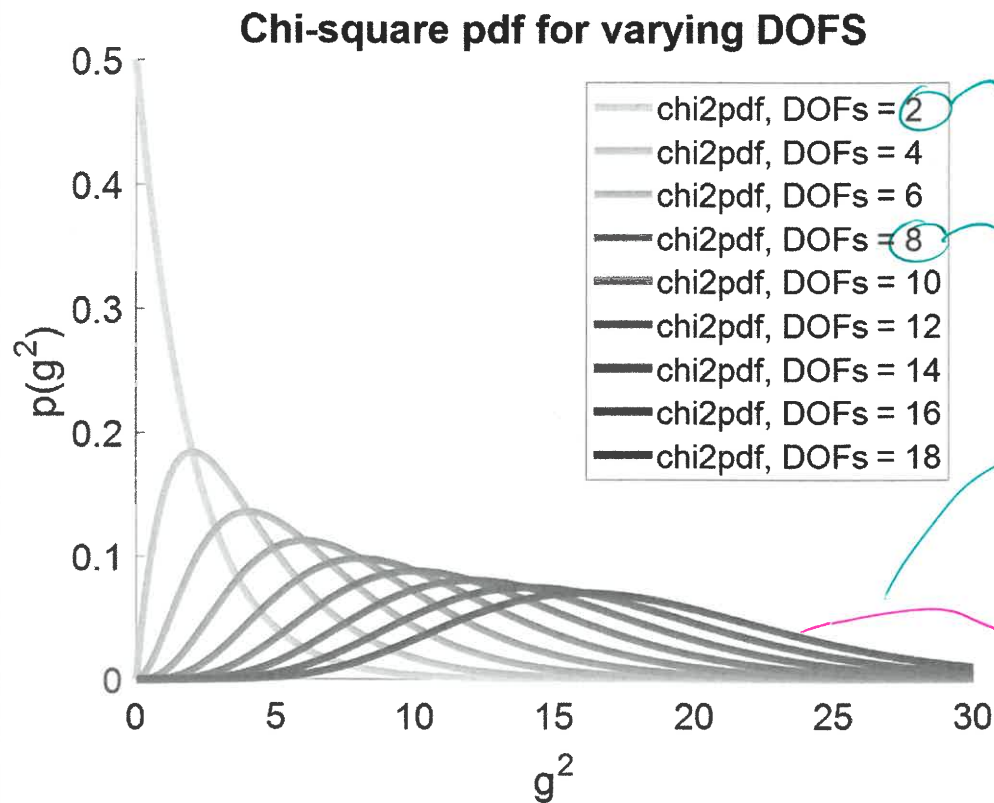
$$\Gamma(m+1) = m \cdot \Gamma(m) = m! \text{ for integer } m$$

* easy to show that: $E[q] = n$, and $\text{var}(q) = 2n$

* also can show: if $q_1 \sim \chi_{n_1}^2$ and $q_2 \sim \chi_{n_2}^2$, then $q_3 = q_1 + q_2 \Rightarrow q_3 \sim \chi_{n_1+n_2}^2$,
(i.e. $n_3 = n_1 + n_2$, so DOFs add!)

Example: Chi-square Distributions

- What does pdf look like for $[\text{norm}(g)]^2$ for different vector lengths $n=\text{length}(e)$?



$$g = \begin{bmatrix} g_1 \\ g_2 \end{bmatrix} \sim \mathcal{N}(0, I_{2 \times 2})$$

$$g = \begin{bmatrix} g_1 \\ \vdots \\ g_8 \end{bmatrix} \sim \mathcal{N}(0, I_{8 \times 8})$$

$$P(\|g\|^2) = P(g^T g) \\ = P\left(\sum_{i=1}^n g_i^2\right) = P(\epsilon)$$

Note: Starts to look like Gaussian pdf for large n/DOFs ! (by the CLT!)

Upshot: Theoretical KF NEES and NIS Error Distributions

- So, combining Facts #1 and #2, we deduce the following must be true:

If the KF works properly as per our DT state space model and noise specs

(i.e. if it meets the consistency criteria #1-#3 laid out earlier), then we must have:

I. if $e_{x,k}(= x_k - \hat{x}_k^+) \sim \mathcal{N}(0, P_k^+)$ and $\epsilon_{x,k} = e_{x,k}^T (P_k^+)^{-1} e_{x,k}$ (NEES)

→ then $\epsilon_{x,k} \sim \chi_n^2 \forall k$, where $E[\epsilon_{x,k}] = n$, $\text{var}(\epsilon_{x,k}) = 2n$

II. if $e_{y,k}(= y_k - \hat{y}_k^-) \sim \mathcal{N}(0, S_k)$ and $\epsilon_{y,k} = e_{y,k}^T (S_k)^{-1} e_{y,k}$ (NIS)

→ then $\epsilon_{y,k} \sim \chi_p^2 \forall k$, where $E[\epsilon_{y,k}] = p$, $\text{var}(\epsilon_{y,k}) = 2p$

We can use **“truth model testing” (TMT)** with NEES and use real/simulated sensor data with NIS to see if these pdfs actually show up!

→ if NOT, then we did something wrong!! (necessary but not sufficient conditions)