Neuro Dynamic Programming

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1 Introduction

2 Dynamic Programming

Definition 1. (Proper stationary policy). (Reach termination state 0 w.p.1)

$$\rho^{\pi} = \max_{i=1,\dots,n} P^{\pi} \left\{ s_n \neq 0 | i_0 = i \right\} < 1$$

In stochastic shortest path problems, we have two assumptions:

- There exists at least one proper policy;
- For every improper policy π , the corresponding cost-to-oo $J^{\pi}(i)$ is infinite for at least one state i.

Policy Iteration as an Actor-Critic System

- Critic: policy evaluation;
- actor: policy improvement.

3 Neural Network Architectures and Training

Trivial. Using some function to approximate $V^{\pi}, V^*, Q^{\pi}, Q^*$. This book uses neural network.

4 Stochastic Iterative Algorithms

Suppose that we are interested in solving a system of equations of the form

$$Hr = r$$

where H is a function from \mathbb{R}^n into itself. If $Hr = r - \nabla f(r)$, the solution of the system Hr = r is of the form

$$\nabla f(r) = 0,$$

Then it's sometime minimize the cost function f.

One possible algorithm for solving the system Hr=r is provided by the iteration

$$r_{t+1} = Hr_t$$
, or $r_{t+1} = (1 - \gamma)r_t + \gamma Hr_t$.

the second method reduces to the gradient method if $Hr = r - \nabla f(r)$.

Sometimes an exact evaluation of Hr is difficult but that we have access to a random variable s of the form s = Hr + w, where w is a random noise term. Then we obtain stochastic iterative or stochastic approximation algorithm

$$r_{t+1} = (1 - \gamma)r + \gamma(Hr + w).$$

A more concrete setting is obtained as follows. Let v be a random variable with a known probability distribution p(v|r) that depends on r. Suppose that we are interested in solving:

$$\mathbb{E}_{v \sim p(v|r)} \left[g(r, v) \right] = r,$$

where g is a known function. We can use preceeding algorithm:

$$r_{t+1} = (1 - \gamma)r_t + \gamma \mathbb{E}_{v \sim p(v|r)} \left[g(r, v) \right].$$

We can estimate $\mathbb{E}_{v \sim p(v|r)} \left[g(r,v) \right] \approx \frac{1}{k} \sum_{i=1}^{k} g(r,\tilde{v}_i)$. We get Robbins-Monro stochastic approximation algorithm (k=1),

$$r_{t+1} = (1 - \gamma)r_t + \gamma g(r, \tilde{v}),$$

which is a special case of the algorithm $r_{t+1} = (1 - \gamma)r_t + \gamma (Hr_t + w)$, where $Hr = \mathbb{E}_{v \sim p(v|r)} [g(r,v)]$, and $w = g(r,\tilde{v}) - \mathbb{E} [g(r,v)]$.

4.1 THE BASIC MODEL

Let T^i be the set of times at which r(i) updates:

$$r_{t+1}(i) = \begin{cases} r_t(i), & t \notin T^i \\ (1 - \gamma_t(i))r_t(i) + \gamma_t(i)\left((Hr_t)(i) + w_t(i)\right), & t \in T^i \end{cases}$$

Assumption: $\sum_{t=0}^{\infty} \gamma_t(i) = \infty$ and $\sum_{t=0}^{\infty} \gamma_t^2(i) < \infty$.

4.2 CONVERGENCE BASED ON A SMOOTH POTHEN-TIAL FUNCTION

$$r_{t+1} = r_t + \gamma_t \delta_t, \quad \delta_t = Hr_t - r_t + w_t.$$

Let \mathcal{H}_t denote the history of the algorithm

$$\mathcal{H}_t = \{r_0, \dots, r_t, \delta_0, \dots, \delta_{t-1}, \gamma_0, \dots, \gamma_t\}.$$

Assumption 1. Exist function $f : \mathbb{R}^n \to \mathbb{R}$, with the following properties:

- 1. $\forall \mathbb{R}^n, f(r) \geq 0;$
- 2. $\|\nabla f(r_1) \nabla f(r_2)\| \le L \|r_1 r_2\|_2$;
- 3. (Pseudogradient property) $c\|\nabla f(r_t)\|_2^2 + \langle \nabla f(r_t), \mathbb{E}\left[\delta_t | \mathcal{H}_t\right] \rangle \leq 0$
- 4. $\mathbb{E}\left[\|\delta_t\|_2^2|\mathcal{H}_t\right] \le K_1 + K_2\|\nabla f(r_t)\|_2^2$

Proposition 1. Consider the algorithm $r_{t+1} = r_t + \gamma_t s_t$, if $\sum_{t=0}^{\infty} \gamma_t = \infty$ and $\sum_{t=0}^{\infty} \gamma_t^2 < \infty$. Under preceding assumption, the following hold with probability 1:

- The sequence $f(r_t)$ converges;
- $\lim_{t\to\infty} \nabla f(r_t) = 0$;
- Every limit point of r_t is a stationary point of f.

Example 1. (Stochastic Gradient Algorithm).

$$r_{t+1} = r_t + \gamma_t \delta_t, \quad \delta_t = -(\nabla f(r_t) + w_t)$$

Assumption:

1.
$$\sum_{t=0}^{\infty} \gamma_t = \infty$$
, $\sum_{t=0}^{\infty} \gamma_t^2 < \infty$;

2. f is nonnegative and has a Lipschitz continuous gradient;

3.
$$\mathbb{E}[w_t | \mathcal{H}_t] = 0$$
, $\mathbb{E}[\|w_t\|^2 | \mathcal{H}_t] \le A + B \|\nabla f(r_t)\|_2^2$;

We proof Assumption 1 is satisfied.

$$\langle \nabla f(r_t), \mathbb{E} \left[\delta_t | \mathcal{H}_t \right] \rangle = \langle \nabla f(r_t), -\nabla f(x_t) - \mathbb{E} \left[w_t | \mathcal{H}_t \right] \rangle = -\|\nabla f(r_t)\|_2^2$$

$$\mathbb{E}\left[\|\delta_{t}\|_{2}^{2}|\mathcal{H}_{t}\right] = \|\nabla f(r_{t})\|_{2}^{2} + \mathbb{E}\left[\|w_{t}\|_{2}^{2}|\mathcal{H}_{t}\right] + \langle 2\nabla f(r_{t}), \mathbb{E}\left[w_{t}|\mathcal{H}_{t}\right]\rangle$$

$$= \|\nabla f(r_{t})\|_{2}^{2} + A + B\|\nabla f(r_{t})\|_{2}^{2}$$

$$= A + (B+1)\|\nabla f(r_{t})\|_{2}^{2}$$

Example 2. (Estimate of an Unknown Mean). For random variables v with unknow mean μ and unit variance.

$$r_{t+1} = (1 - \gamma_t)r_t + \gamma_t v_t.$$

with assumption

1.
$$\sum_{t=0}^{\infty} \gamma_t = \infty$$
 and $\sum_{t=0}^{\infty} \gamma_t^2 < \infty$;

Proof.

$$r_{t+1} = r_t - \gamma_t(r_t - \mu) + \gamma_t(v_t - \mu)$$

where $f(r) = (r - \mu)^2/2$, $\nabla f(r_t) = (r_t - \mu)$. (The other assumptions in stochastic gradient algorithm are sastified naturally.)

Example 3. (Euclidean Norm Pseudo-Contractions).

$$r_{t+1} = (1 - \gamma_t)r_t + \gamma_t(Hr_t + w_t),$$

Assuming:

1.
$$||Hr - r^*||_2 \le \beta ||r - r^*||_2, \forall r \in \mathbb{R}^n, \ 0 \le \beta < 1;$$

2.
$$\mathbb{E}[w_t|\mathcal{H}_t] = 0$$
;

3.
$$\mathbb{E}\left[\|w_t\|_2^2 |\mathcal{H}_t|\right] \le A + B\|r_t - r^*\|_2^2$$

The potential function is $f(r) = \frac{1}{2} ||r - r^*||_2^2$, $\delta_t = -r_t + Hr_t + w_t$, then $\mathbb{E} [\delta_t | \mathcal{H}_t] = Hr_t - r_t$.

$$\langle Hr - r^*, r - r^* \rangle \le ||Hr - r^*||_2 ||r - r^*||_2 \le \beta ||r - r^*||_2^2$$
$$\langle Hr - r, r - r^* \rangle \le -(1 - \beta) ||r - r^*||_2^2$$
$$\langle \mathbb{E} \left[\delta_t | \mathcal{H}_t \right], \nabla f(r_t) \rangle < -(1 - \beta) ||\nabla f(r_t)||_2^2$$

$$\mathbb{E}\left[\delta_t^2|\mathcal{H}_t\right] = \mathbb{E}\left[\left(-r_t + Hr_t\right)^2|\mathcal{H}_t\right] + \mathbb{E}\left[\|w_t\|^2|\mathcal{H}_t\right] \leq \left(Hr_t - r_t\right)^2 + A + B\|r_t - r^*\|_2^2$$

4.2.1 Convergence Proofs

In this section, we discarded a suitable set of measure zero, and don't keep repeating the qualification "with probability 1".

Theorem 1. (Supermartingale Convergence Theorem). Here is three sequences of random variables $\{X_t\}$, $\{Y_t\}$ and $\{Z_t\}$. And let \mathcal{F}_t be set of random variables and $\mathcal{F}_t \subset \mathcal{F}_{t+1}$. Suppose that

- 1. X_t, Y_t, Z_t are nonegative, and are functions of the random variables in \mathcal{F}_t ;
- 2. $\forall t$, we have $\mathbb{E}[Y_{t+1}|\mathcal{F}_t] \leq Y_t X_t + Z_t$;
- 3. $\sum_{t=0}^{\infty} Z_t < \infty.$

Then we have $\sum_{t=0}^{\infty} X_t < \infty$, and the sequence Y_t converges to a nonegative random variable Y, w.p.1.

Theorem 2. (Martigale Convergence Theorem) Let $\{X_t\}$ be a sequence of random variables and let \mathcal{F}_t be set of random variables such that $\mathcal{F}_t \subset \mathcal{F}_{t+1}$. Suppose that:

- 1. The random variable X_t is a function of the random variable in \mathcal{F}_t ;
- 2. $\mathbb{E}[X_{t+1}|\mathcal{F}_t] = X_t$,
- 3. $\exists M < \infty \text{ such that } \mathbb{E}[|X_t|] \leq M$.

Then, the sequence X_t converges to a random variable X, w.p.1. Now we begin proof the preceding section.

Proof. By assumption, we have $\|\nabla f(r_1) - \nabla f(r_2)\|_2 \le L\|r_1 - r_2\|$, we have

$$f(r_{t+1}) \le f(r_t) + \gamma_t \langle \nabla f(r), \delta_t \rangle + \frac{L}{2} \gamma_t^2 ||\delta_t||_2^2$$

$$\mathbb{E}\left[f(r_{t+1}|\mathcal{F}_{t})\right] \leq f(r_{t}) + \gamma_{t} \langle \nabla f(r_{t}), \mathbb{E}\left[\delta_{t}|\mathcal{F}_{t}\right] \rangle + \frac{L}{2} \gamma_{t}^{2} \left(K_{1} + K_{2} \|\nabla f(r_{t})\|_{2}^{2}\right)$$

$$\leq f(r_{t}) - \gamma_{t} \left(c - \frac{LK_{2}\gamma_{t}}{2}\right) \|\nabla f(r_{t})\|_{2}^{2} + \frac{LK_{1}\gamma_{t}^{2}}{2}$$

$$= f(r_{t}) - X_{t} + Z_{t},$$

where

$$X_t = \begin{cases} \gamma_t \left(c - \frac{LK_2 \gamma_t}{2} \right) \|\nabla f(r_t)\|_2^2, & if \ LK_2 \gamma_t \le 2c, \\ 0, & otherwise. \end{cases}$$

and

$$Z_t = \begin{cases} \frac{LK_1\gamma_t^2}{2}, & if \ LK_2\gamma_t \le 2c, \\ \frac{LK_1\gamma_t^2}{2} - \gamma_t \left(c - \frac{LK_2\gamma_t}{2}\right) \|\nabla f(r_t)\|_2^2, & otherwise \end{cases}$$

Because $\sum_{t=0}^{\infty} \gamma_t^2 < \infty$, so after some finite time $LK_2\gamma_t \leq 2c$, and $Z_t = LK_1\gamma_t^2/2$, and therefore $\sum_{t=0}^{\infty} Z_t < \infty$. Thus, the supermartingale convergence theorem applies and shows that $f(r_t)$ converges and $\sum_{t=0}^{\infty} X_t < \infty$.

Because $X_t = \gamma_t \left(c - \frac{LK_2\gamma_t}{2}\right) \|\nabla f(r_t)\|_2^2 \ge \frac{c}{2}\gamma_t \|\nabla f(r_t)\|_2^2$ after some finite time. Hence

$$\sum_{t=0}^{\infty} \gamma_t \|\nabla f(r_t)\|_2^2 < \infty$$

Because $\sum_{t=0}^{\infty} \gamma_t = \infty$, $\liminf_{t \to \infty} \|\nabla f(r_t)\|_2 = 0$ Let us denote $\bar{s}_t = \mathbb{E}\left[s_t | \mathcal{F}_t\right]$ and $w_t = s_t - \bar{s}_t$, then

$$\|\bar{s}_t\|_2^2 + \mathbb{E}\left[\|w_t\|_2^2 |\mathcal{F}_t\right] = \mathbb{E}\left[\|s_t\|_2^2 |\mathcal{F}_t\right] \le K_1 + K_2 \|\nabla f(r_t)\|_2^2$$

We need take a break and proof another lemma

Lemma 1. $u_t = \sum_{\tau=0}^{t-1} \chi_\tau \gamma_\tau w_\tau$, converges w.p.1. where $\chi_t = \mathbb{1}_{\left[\|\nabla f(r_t)\|_2 \le \epsilon\right]}$.

Proof. We start the assumption $\sum_{t=0}^{\infty} \gamma_t^2 \leq A < \infty$.

$$\mathbb{E}\left[\chi_t \gamma_t w_t | \mathcal{F}_t\right] = \chi_t \gamma_t \mathbb{E}\left[w_t | \mathcal{F}_t\right] = 0 \Rightarrow \mathbb{E}\left[u_{t+1} | \mathcal{F}_t\right] = u_t$$

If $\chi_t = 0$, then $\mathbb{E}\left[\|u_{t+1}\|_2^2|\mathcal{F}_t\right] = \|u_t\|^2$. If $\chi_t = 1$, we have

$$\mathbb{E}\left[\|u_{t+1}\|_{2}^{2}|\mathcal{F}_{t}\right] = \|u_{t}\|_{2}^{2} + \gamma_{t}^{2}\mathbb{E}\left[\|w_{t}\|_{2}^{2}|\mathcal{F}_{t}\right] \le \|u_{t}\|_{2}^{2} + \gamma_{t}^{2}(K_{1} + K_{2}\epsilon^{2})$$

$$\mathbb{E}\left[\left\|u_{t}\right\|_{2}^{2}\right] \leq (K_{1} + K_{2}\epsilon^{2})\mathbb{E}\left[\sum_{\tau=0}^{t-1} \gamma_{\tau}^{2}\right] \leq (K_{1} + K_{2}\epsilon^{2})A$$

$$\sup_{t} \mathbb{E}\left[\left\|u_{t}\right\|^{2}\right] \leq \sup_{t} \mathbb{E}\left[1 + \left\|u_{t}\right\|_{2}^{2}\right] < \infty$$

Then we can use Martigale convergence theorem to u_t and get that u_t converges, w.p.1.

We can assume that $\sum_{\tau=0}^{t-1} \gamma_{\tau}^2 \leq A < \infty$ and get the same result.

I give up today. □

5 Sinulation Methods for a Lookup Table Representation

5.1 SOME ASPECTS OF MONTE CARLO SIMULATION

Suppose that v is a random variable with an unknown mean m that we wish to estimate. Monte-Carlo simulation is to generate a number of samples $\{v_1, \ldots, v_N\}$, and estimate the mean of v by forming the sample mean

$$M_N = \frac{1}{N} \sum_{k=1}^{N} v_k = M_{N-1} + \frac{1}{N} (v_N - M_N).$$

(The Case of i.i.d. Samples):

• $\mathbb{E}[M_N] = \frac{1}{N} \sum_{k=1}^N \mathbb{E}[v_k] = m;$

• $Var(M_N) = \frac{1}{N^2} \sum_{k=1}^{N} Var(v_k) = \frac{\sigma^2}{N}$.

(The Case of Dependent Samples):

- The estimator of mean remains unbiased;
- We can use a weighted average to lower the variance.

(The Case of a Random Sample Size):

In this case, the number of samples N is itself a random variable.

If N is correlated with $\{v_1, \ldots, v_N\}$. Then the sample mean is unbiased, and the variance is $\sigma^2 \mathbb{E}[1/N]$.

If N is correlated with $\{v_1, \ldots, v_N\}$, the sample mean maybe biased.

Theorem 3. If $\{v_1, \ldots, v_N\}$ is i.i.d. with finite mean, N depends upon given sequence, and $\mathbb{E}[N] < \infty$.

$$\mathbb{E}\left[\sum_{k=1}^{N} v_{k}\right] = \sum_{k=1}^{\infty} P(N \geq k) \mathbb{E}\left[v_{k} | N \geq k\right] = \mathbb{E}\left[v_{1}\right] \sum_{k=1}^{\infty} P\left(N \geq k\right)$$
$$= \mathbb{E}\left[v_{1}\right] \sum_{k=1}^{\infty} \sum_{k=1}^{\infty} P(N = n) = \mathbb{E}\left[v_{1}\right] \sum_{k=1}^{\infty} \sum_{k=1}^{n} P(N = n) = \mathbb{E}\left[v_{1}\right] \mathbb{E}\left[N\right].$$

5.2 POLICY EVALUATION BY MONTE CARLO SIM-ULATION

- State space: $\{0, 1, \dots, n\}$, where 0 is a cost-free absorbing state;
- The policy π is proper;
- For mth trajectory is $(s_0^m, s_1^m, \dots, s_N^m)$;
- Cost of trajectory is $c(s_0^m, m) = \sum_{t=1}^{N-1} g(s_t^m, s_{t+1}^m);$
- Policy value: $J^{\pi}(s) = \mathbb{E}^{\pi} [c(s, m)].$

Algorithm1:

$$\tilde{J}(i) = \frac{1}{K} \sum_{m=1}^{K} c(i, m)$$

$$\tilde{J}^m(i) = \tilde{J}^{m-1}(i) + \frac{1}{m}(c(i,m) - \tilde{J}^{m-1}(i)), \quad s.t.J^0(i) = 0.$$

Algorithm2: Use the full trajectory.

$$J(i_k) = J(i_k) + \gamma(i_k)(g(i_k, i_k + 1) + \dots + g(i_{N-1}, i_N) - J(i_k))$$

Every-visit method provides a biased estimator.

Consistency of the Every-Visit Method

Let c(s, k, m) mean, in mth trajectory, the cost after visiting state s kth times. And n_s^m means the total number of state s in mth trajectory. Then every-visit method estimator is

$$\tilde{J}(s) = \frac{\sum_{m=1}^{K} \sum_{k=1}^{n_s^m} c(s, k, m)}{\sum_{m=1}^{K} n_s^m}$$

When K is fixed, the extimator is biased. But if $K \to \infty$, the estimator is unbiased:

$$\mathbb{E}\left[\tilde{J}(s)\right] = \frac{\mathbb{E}\left[\sum_{k=1}^{n_k^m} c(s,k,m) | n_k \geq 1\right]}{\mathbb{E}\left[n_k | n_k \geq 1\right]} = \mathbb{E}\left[c(s,1,m) | n_k \geq 1\right] = J^{\pi}(s)$$

The First-Visit Method

$$\tilde{J}(s) = \frac{\sum_{m:n_s^m \ge 1} c(s, 1, m)}{\sum_{m=1}^K 1_{[n_s^m \ge 1]}}$$

which is unbiased.

The significance of the comparison of the every-visit and first-visit methods should not be overemphasized. For problems with large state space, the likelihood of a trajectory visiting the same state twice is usually quite small.

5.2.1 Q-Factors and Policy Iteration

$$Q^{\pi}(s, a) = \sum_{s' \in S} p_{s, s'}(a) (g(s, a, s') + J^{\pi}(s'))$$

5.3 TEMPORAL DIFFERENCE METHODS

TD: policy evaluation.

$$\begin{split} J_{k+1}^m(s_k^m) &= J_k^m(s_k^m) + \gamma(g(s_k^m, s_{k+1}^m) + \dots + g(s_{N-1}^m, s_N^m) - J_k^m(i_k)) \\ J_{k+1}^m(s_k^m) &= J_k^m(s_k^m) + \gamma[(g(s_k^m, s_{k+1}^m) + J_{k+1}^m(s_{k+1}^m) - J_k^m(s_k^m)) \\ &\quad + (g(s_{k+1}^m, s_{k+2}^m) + J_{k+2}^m(s_{k+2}^m) - J_{k+1}^m(s_{k+1}^m)) \\ &\quad + \dots \\ &\quad + (g(s_{N-1}^m, s_N^m) + J_N^m(s_N^m) - J_{N-1}^m(s_{N-1}^m))] \end{split}$$

where we have made use of $J(s_N) = 0$.

$$J_{k+1}^m(s_k^m) = J_k^m(s_k^m) + \gamma(d_k^m + d_{k+1}^m + \dots + d_{N-1}^m),$$

$$d_k^m = g(s_k^m, s_{k+1}^m) + J_{k+1}^m(s_{k+1}^m) - J_k^m(s_k^m)$$

Make explanation complex, maybe approximation TD is more easy to understanding.

6 Approximate DP with Cost-to-Go Function Approximation

This chapter is the core target I read this book.

The lookup table representation can be viewed as a limiting form of an appproximate representation.

- $\begin{array}{ccc} \textbf{6.1} & \textbf{GENERRIC ISSUES-FROM PARAMETERS TO POLICIES} \\ \end{array}$
- 6.2 APPROXIMATE POLICY ITERATION