

# 2019-2019 概率统计理工 A 卷参考解答

## 一、填空 (3\*6=18 分)

1. 2/9;    2. 1/2;    3. 0.6;    4. 0.3085;    5. 4/9;    6.  $\pm\sqrt{2}/2$

## 二、解答题 (82 分)

1. (12 分) 记  $A$  为该生知道答案,  $B$  为他答对该题, 则

$$(1) P(B) = P(A)P(B|A) + P(\bar{A})P(B|\bar{A}) = 0.4 \times 1 + 0.6 \times 0.25 = 0.55 \quad (2)$$

$$(2) P(A|B) = \frac{P(AB)}{P(B)} = \frac{P(A)P(B|A)}{P(B)} = \frac{0.4 \times 1}{0.55} = \frac{8}{11} \quad (2)$$

2. (10 分) (1) 根据概率分布求和为 1,  $9a + 0.1 = 1 \Rightarrow a = 0.1$  (3)

$$(2) Y \text{ 的概率分布为 } Y \sim \begin{pmatrix} -1 & 0 & 3 & 8 \\ 0.3 & 0.2 & 0.3 & 0.2 \end{pmatrix} \quad (4)$$

$$\text{分布函数为 } F_Y(y) = \begin{cases} 0, & y < -1 \\ 0.3, & -1 \leq y < 0 \\ 0.5, & 0 \leq y < 3 \\ 0.8, & 3 \leq y < 8 \\ 1, & y \geq 8 \end{cases} \quad (3)$$

3. (10 分) (1) “第二个学生拿到自己的帽子”等价于“第一个学生没有拿到第二个的帽子,

然后第二个人拿到自己的帽子”, 故概率为  $\frac{49}{50} \times \frac{1}{49} = \frac{1}{50}$  (4)

$$(2) \text{ 记随机变量 } X_i = \begin{cases} 1, & \text{第 } i \text{ 个学生拿到自己的帽子} \\ 0, & \text{第 } i \text{ 个学生没有拿到自己的帽子} \end{cases}, \text{ 则 } X = \sum_{i=1}^{50} X_i \quad (2)$$

$$E(X_i) = P(X_i = 1) = \frac{C_{49}^{i-1}}{C_{50}^{i-1}} \times \frac{1}{50 - (i-1)} = \frac{1}{50} \quad (2) \quad E(X) = \sum_{i=1}^{50} E(X_i) = 1 \quad (2)$$

$$4. (15 \text{ 分}) (1) f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_x^1 3y dy = \frac{3}{2}(1-x^2), & 0 < x < 1 \\ 0, & \text{其它} \end{cases} \quad (2)$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_0^y 3y dx = 3y^2, & 0 < y < 1 \\ 0, & \text{其它} \end{cases} \quad (2)$$

$$f(x, y) \neq f_X(x)f_Y(y), \text{ 不独立} \quad (1)$$

$$(2) \text{ 给定 } y \in (0,1), f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \begin{cases} \frac{3y}{3y^2} = \frac{1}{y}, & 0 < x < y \\ 0, & \text{其它} \end{cases} \quad (2)$$

$$P\left(X \geq \frac{1}{2} \middle| Y = \frac{3}{4}\right) = \int_{-\infty}^{+\infty} f_{X|Y}(x|\frac{3}{4}) dx = \int_{\frac{1}{2}}^{\frac{3}{4}} \frac{4}{3} dx = \frac{1}{3} \quad (2)$$

$$(3) \quad Z = Y - X, Z \in (0,1), \forall Z \in (0,1), \quad (2)$$

$$F_Z(z) = P(Y - X \leq z) = 1 - \int_z^1 dy \int_0^{y-z} 3y dx = \frac{3}{2}z - \frac{1}{2}z^3 \quad (2)$$

$$f_Z(z) = F'_Z(z) = \begin{cases} \frac{3}{2}(1-z^2), & 0 < z < 1 \\ 0, & \text{其它} \end{cases} \quad (2)$$

$$5. \quad (8 \text{ 分}) \quad (1) \quad X_i \text{ 的密度函数为 } f_X(x) = \begin{cases} 1, & 0 < x < 1 \\ 0, & \text{其它} \end{cases}$$

$$X_i \text{ 独立同分布} \Rightarrow g(X_i) \text{ 独立同分布} \quad (1)$$

$$E(g(X_i)) = \int_{-\infty}^{+\infty} g(x)f_X(x)dx = \int_0^1 g(x)dx = I \quad (2)$$

$$E(g^2(X_i)) = \int_0^1 g^2(x)dx \text{ 存在, 故 } g(X_i) \text{ 方差存在}$$

根据独立同分布的大数定律, 随机变量序列  $\{Y_n\}$  依概率收敛到  $I$

(2)  $Y_{100}$  为 100 个独立同分布的随机变量的平均, 根据中心极限定理,  $Y_{100}$  近似服从

$$\text{正态分布, } I = \int_0^1 x^{1.5} dx = 0.4, \quad E(g(X_i)) = \int_{-\infty}^{+\infty} g(x)f_X(x)dx = \int_0^1 x^{1.5} dx = 0.4, \quad (1)$$

$$E(g^2(X_i)) = \int_{-\infty}^{+\infty} g^2(x)f_X(x)dx = \int_0^1 x^3 dx = 0.25 \quad (1)$$

$$D(g(X_i)) = E(g^2(X_i)) - E^2(g(X_i)) = 0.25 - 0.4^2 = 0.09 \quad (1)$$

$$E(Y_{100}) = I = 0.4, D(Y_{100}) = D(g(X_i))/100 = 0.0009, Y_{100} \sim N(0.4, 0.0009) \quad (1)$$

$$\begin{aligned} P(|Y_{100} - I| < 0.01) &= P(0.39 < Y_{100} < 0.41) = \Phi\left(\frac{0.41 - 0.4}{0.03}\right) - \Phi\left(\frac{0.39 - 0.4}{0.03}\right) \\ &= \Phi(1/3) - \Phi(-1/3) = 2\Phi(1/3) - 1 = 2 \times 0.6304 - 1 = 0.2608 \quad (1) \end{aligned}$$

$$6. (15 \text{ 分}) (1) E(X) = \int_0^{\theta} x \frac{3x^2}{\theta^3} dx = \frac{3}{4}\theta \Rightarrow \theta = \frac{4}{3}E(X) \Rightarrow \hat{\theta}_1 = \frac{4}{3}\bar{X} \quad (2)$$

$$(2) L(\theta) = \frac{3x_1^2}{\theta^3} \frac{3x_2^2}{\theta^3} \dots \frac{3x_n^2}{\theta^3}, \ln L(\theta) = n \ln 3 - 3n \ln \theta + 2 \sum_{i=1}^n \ln x_i \quad (2)$$

$$[\ln L(\theta)]' = -\frac{3n}{\theta} < 0 \Rightarrow L(\theta) \text{ 单调递减} \quad (2)$$

$$X_i \leq \theta \Rightarrow \theta \geq \max\{X_i\} \Rightarrow \hat{\theta}_2 = \max\{X_i\} \quad (2)$$

$$(3) E(\hat{\theta}_1) = E\left(\frac{4}{3}\bar{X}\right) = \frac{4}{3}E(X) = \frac{4}{3} \times \frac{3}{4}\theta = \theta, \text{ 故 } \hat{\theta}_1 \text{ 为无偏估计} \quad (2)$$

$$X \text{ 的分布函数 } F(x) = \begin{cases} 0, & x < 0 \\ \frac{x^3}{\theta^3}, & 0 \leq x \leq \theta \\ 1, & x > \theta \end{cases}, F_{\hat{\theta}_2}(x) = F^n(x)$$

$$f_{\hat{\theta}_2}(x) = nF^{n-1}(x)f(x) = \begin{cases} 3n \frac{x^{3n-1}}{\theta^{3n}}, & 0 \leq x \leq \theta \\ 0, & \text{其它} \end{cases} \quad (1)$$

$$E(\hat{\theta}_2) = \int_0^{\theta} x(3n \frac{x^{3n-1}}{\theta^{3n}}) dx = \frac{3n}{3n+1}\theta \neq \theta, \text{ 故 } \hat{\theta}_2 \text{ 不是无偏估计} \quad (2)$$

$$7. (12 \text{ 分}) (1) \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1), 1 - \frac{\alpha}{2} = 0.975, t_{0.975}(24) = 2.0639 \quad (2)$$

$$95\% CI: (\bar{x} - \frac{s}{\sqrt{n}} t_{0.975}(24), \bar{x} + \frac{s}{\sqrt{n}} t_{0.975}(24)) = (7.69, 8.11) \quad (2)$$

$$(2) H_0: \mu \geq \mu_0 = 8, H_1: \mu < \mu_0, \quad (2)$$

$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}, \text{ 在原假设下, } P(T < t_{0.05}(24)) \leq 0.05, \quad (2)$$

$$t_{0.05}(24) = -t_{0.95}(24) = -1.7109, \text{ 故拒绝域为: } T < -1.7109 \quad (2)$$

$$\text{代入数据得 } T = \frac{7.9 - 8}{0.5/\sqrt{25}} = -1 > -1.7109, \quad (2)$$

没有落入拒绝域, 故不能拒绝原假设, 即不能认为这批弹壳直径明显低于标准值