

《习题册》参考答案及解答

第 1-2 页 《样本空间与事件》

一. 1. D 2. C 二. 1. A 与 B 恰有一个事件发生 2. B

三. 1. (1) $\Omega = \{(x, y) : 0 \leq x \leq 1, 0 \leq y \leq 1\},$

$$A = \left\{ (x, y) : x^2 + y^2 < \frac{1}{4}, x \geq 0, y \geq 0 \right\}$$

$$B = \left\{ (x, y) : 0 \leq x < y, \frac{1}{3} < y \leq 1 \right\}$$

$$(2) \Omega = \{(i, j) : i, j = 1, 2, 3, 4, 5, 6\}$$

$$A = \{(i, j) : i = 1, 3, 5, j = 1, 2, 3, 4, 5, 6\} \quad B = \{(1, 3), (3, 1), (2, 2)\}$$

$$2. (1) A\bar{B}\bar{C} = \left\{ \omega : 0 \leq \omega < \frac{1}{4} \right\} \quad (2) \bar{A}\bar{B}\bar{C} = \{\omega : \omega = 1\} = \{1\}$$

$$(3) \bar{A} \cup \bar{B} \cup \bar{C} = \overline{ABC} = \{\omega : 0 \leq \omega \leq 1\} = \Omega$$

$$(4) \bar{A}(B \cup C) = \left\{ \omega : \frac{1}{3} \leq \omega < 1 \right\} \quad (5) A \cup B \cup C = \{\omega : 0 \leq \omega < 1\}$$

3. (1) A, B, C 恰有一个发生 (2) 至少两个发生

(3) A 发生, 且 B, C 至少一个不发生 (4) 至多一个发生

第 3-4 页 《概率的性质与古典概率》

一. 1. D

$$P(\bar{A}) = 0.3 \Rightarrow P(A) = 0.7 \Rightarrow P(AB) = P(A) - P(A\bar{B}) = 0.3$$

$$\Rightarrow P(\bar{A} \cup \bar{B}) = P(\overline{AB}) = 1 - P(AB) = 1 - 0.3 = 0.7$$

2. C 3. D

二. 1. $\frac{2}{9}$

$$P(AC) = 0 \Rightarrow P(ABC) = 0$$

$$\Rightarrow P(\bar{A}\bar{B}\bar{C}) = P(\overline{A \cup B \cup C}) = 1 - P(A \cup B \cup C)$$

$$= 1 - P(A) - P(B) - P(C) + P(AB) + P(AC) + P(BC) - P(ABC)$$

$$= 1 - \frac{1}{3} - \frac{1}{3} - \frac{1}{3} + \frac{1}{9} + \frac{1}{9} + 0 - 0 = \frac{2}{9}$$

$$2. \frac{1}{12} \quad 3. 1 - q$$

三. 1. 教材习题一(A)三第 5 题

解答: 因 $P(A - B) = P(A) - P(AB)$, $P(A \cup B) = P(A) + P(B) - P(AB)$,

则

$$P(A) - P(B) \leq P(A) - P(AB) \leq P(A) \leq P(A \cup B) \leq P(A) + P(B),$$

即

$$P(A) - P(B) \leq P(A - B) \leq P(A) \leq P(A \cup B) \leq P(A) + P(B).$$

2. 证明:

$$\begin{aligned} P(A) &\geq P(A(B \cup C)) = P(AB \cup AC) = P(AB) + P(AC) - P(ABC) \\ &\geq P(AB) + P(AC) - P(BC), \end{aligned}$$

故

$$P(AB) + P(AC) - P(BC) \leq P(A)$$

$$3. (1) \frac{A_8^5}{8^5} = 0.205 \quad (2) 1 - \frac{A_8^5}{8^5} = 0.795 \quad (3) \frac{C_5^2 C_8^1 (A_7^3 + A_7^1)}{8^5} = 0.5298$$

4. 教材习题一(A)三第 10 题

$$(1) \text{ 有放回抽取时: } P(A) = \frac{2^2}{6^2} = \frac{1}{9}; P(B) = \frac{4 \times 2 + 2 \times 4}{6^2} = \frac{4}{9}; \text{ 因 } C = A \cup B$$

$$\text{且 } AB = \emptyset, \text{ 所以 } P(C) = P(A) + P(B) = \frac{5}{9}$$

(2) 无放回抽取时:

$$\text{用排列计算: } P(A) = \frac{A_2^2}{A_6^2} = \frac{1}{15}; P(B) = \frac{A_4^1 A_2^1 + A_2^1 A_4^1}{A_6^2} = \frac{8}{15}; \text{ 因 } C = A \cup B$$

$$\text{且 } AB = \emptyset, \text{ 所以 } P(C) = P(A) + P(B) = \frac{9}{15} = \frac{3}{5}$$

用组合计算： $P(A) = \frac{C_2^2}{C_6^2} = \frac{1}{15}$ ； $P(B) = \frac{C_4^1 C_2^1}{C_6^2} = \frac{8}{15}$ ；因 $C = A \cup B$ 且

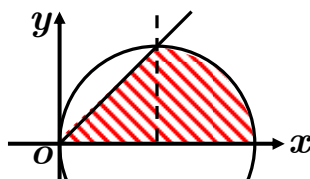
$$AB = \emptyset, \text{ 所以 } P(C) = P(A) + P(B) = \frac{9}{15} = \frac{3}{5}$$

$$5. p = \frac{C_4^1 C_{13}^7 C_{13}^3 C_{13}^1 C_{13}^2}{C_{52}^{13}}$$

第 5-6 页《几何概率、条件概率及乘法公式》

一. 1. B 2. D 3. D

二. 1. $\frac{2 + \pi}{2\pi}$



这是一个二维几何概型问题：如图

样本空间为中心在 $(a, 0)$ 处，半径为 a 的上半圆，其面积为 $m(\Omega) = \frac{1}{2} \pi a^2$ ，设 A 表事

件“原点与该点的连线与 x 轴的夹角小于 $\frac{\pi}{4}$ ”，则 $m(A) = \frac{1}{4} \pi a^2 + \frac{1}{2} a^2$ ，所以所求概

$$\text{率为 } P(A) = \frac{m(A)}{m(\Omega)} = \frac{\frac{1}{4} \pi a^2 + \frac{1}{2} a^2}{\frac{1}{2} \pi a^2} = \frac{2 + \pi}{2\pi}.$$

2. $\frac{3}{8}$

因事件 A 发生导致事件 B 发生，则 $A \subset B$ 或 $AB = A$ ；事件 B 与事件 C 互斥，则

$BC = \emptyset$ 或 $B \subset \bar{C}$ ；从而有 $AB\bar{C} = A$ ， $B\bar{C} = B$ ，于是

$$P(A|B\bar{C}) = \frac{P(AB\bar{C})}{P(B\bar{C})} = \frac{P(A)}{P(B)} = \frac{0.3}{0.8} = \frac{3}{8}$$

3. $\frac{6}{7}$

设 $A =$ “至少有一个女孩”， $B =$ “至少有一个男孩”，则 $\bar{A} =$ “三个孩子全是男孩”

$\overline{AB} =$ “三个孩子全是男孩或全是女孩”，从而有 $P(\bar{A}) = \frac{1}{8}$ ， $P(\overline{AB}) = \frac{2}{8}$ ，故所求

$$\text{概率为 } P(B|A) = \frac{P(AB)}{P(A)} = \frac{1 - P(\overline{AB})}{1 - P(\overline{A})} = \frac{1 - \frac{2}{8}}{1 - \frac{1}{8}} = \frac{6}{7}$$

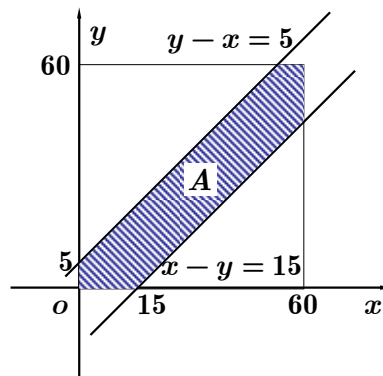
三. 1. 0.2986

设王同学于 9 点 X 分到达，张同学于 9 点 Y 分到达，如图，则

$$\Omega = \{(X, Y) : 0 \leq X \leq 60, 0 \leq Y \leq 60\}$$

设 A = “两同学能见面”，则

$$A = \{(X, Y) : 0 \leq Y - X \leq 5\} \cup \{(X, Y) : 0 \leq X - Y \leq 15\}$$



则所求概率为

$$p = \frac{m(A)}{m(\Omega)} = 1 - \frac{1}{2} \frac{55^2 + 45^2}{60^2} = 0.2986$$

2. 0.7283

设 A = “该种动物活到 10 岁”， B = “该种动物活到 15 岁”，由已知条件得所求

$$\text{概率为 } p = P(B|A) = \frac{P(AB)}{P(A)} = \frac{P(B)}{P(A)} = \frac{0.67}{0.92} = 0.7283$$

3. 0.2333, 0.4651

因 $A\bar{B} = A - B$ ，所以

$$P(A|\bar{B}) = \frac{P(A\bar{B})}{P(\bar{B})} = \frac{P(A - B)}{1 - P(B)} = \frac{0.14}{1 - 0.4} = 0.2333$$

$$\overline{\overline{A \cup B}} = \overline{\bar{A}\bar{B}} = A\bar{B} = A - B$$

$$\Rightarrow P(\bar{A \cup B}) = 1 - P(\overline{\bar{A}\bar{B}}) = 1 - P(A - B) = 1 - 0.14 = 0.86$$

$$\Rightarrow P(B|\bar{A \cup B}) = \frac{P(B(\bar{A \cup B}))}{P(\bar{A \cup B})} = \frac{P(B)}{P(\bar{A \cup B})} = \frac{0.4}{0.86} = 0.4651$$

$$4. P(B|A) = \frac{P(AB)}{P(A)} = \frac{P(A) - P(A\bar{B})}{P(A)} \geq \frac{P(A) - P(\bar{B})}{P(A)} = 1 - \frac{1 - p_2}{p_1}$$

5. 0.24, 0.424

设 A = “甲机第一次攻击并击落乙机”， B = “乙机第一次攻击并击落甲机”，
 C = “甲机第二次攻击并击落乙机”，则

$$(1) P(B) = P(AB) + P(\bar{A}B) = 0 + P(\bar{A})P(B|\bar{A}) = 0.8 \times 0.3 = 0.24;$$

$$\begin{aligned} (2) P(A \cup C) &= P(A) + P(C) - P(AC) = 0.2 + P(\bar{A}BC) - 0 \\ &= 0.2 + P(\bar{A})P(B|\bar{A})P(C|\bar{A}B) \\ &= 0.2 + 0.8 \times 0.7 \times 0.4 = 0.424. \end{aligned}$$

第 7-8 页 《全概率与贝叶斯公式、事件的独立性与贝努利概型》

一. 1. C

$$\begin{aligned} P(\bar{B}|\bar{A} \cup B) &= \frac{P(\bar{B}(\bar{A} \cup B))}{P(\bar{A} \cup B)} = \frac{P(\bar{A}\bar{B})}{P(\bar{A} \cup B)} = \frac{P(\bar{A})P(\bar{B})}{P(\bar{A}) + P(B) - P(\bar{A}B)} \\ &= \frac{P(\bar{A})P(\bar{B})}{P(\bar{A}) + P(B) - P(\bar{A})P(B)} = \frac{0.2 \times 0.6}{0.2 + 0.4 - 0.2 \times 0.4} = 0.2308 \end{aligned}$$

2. D

3. C

$$\begin{aligned} P(A|B) + P(\bar{A}|\bar{B}) &= 1 \Rightarrow P(A|B) = 1 - P(\bar{A}|\bar{B}) = P(A|\bar{B}) = \frac{P(A\bar{B})}{P(\bar{B})} \\ \Rightarrow P(A|B) &= \frac{P(A) - P(AB)}{1 - P(B)} \Rightarrow P(A|B)[1 - P(B)] = P(A) - P(AB) \\ \Rightarrow P(A|B) &= P(A) \Rightarrow A \text{ 与 } B \text{ 相互独立.} \end{aligned}$$

二. 1. $p^3(2-p)$

2. $\frac{48}{60}$

因 A 与 B 互斥，故 $A\bar{C}$ 与 $B\bar{C}$ 互斥，从而有

$$\begin{aligned} P(A \cup B|\bar{C}) &= \frac{P((A \cup B)\bar{C})}{P(\bar{C})} = \frac{P(A\bar{C} \cup B\bar{C})}{P(\bar{C})} = \frac{P(A\bar{C}) + P(B\bar{C})}{P(\bar{C})} \\ &= \frac{P(A)P(\bar{C}) + P(B)P(\bar{C})}{P(\bar{C})} = P(A) + P(B) = 0.8 \end{aligned}$$

思考题：一般情况下， A 与 C 独立， B 与 C 独立，则 $A \cup B$ 与 C 也独立吗？

3. 49

设同学数为 n ，则由题意有 $1 - (1 - p)^n = 1 - 0.94^n \geq 0.95 \Rightarrow n \geq 49$

三. 1. 0.15，最可能乘火车

设 A_1, A_2, A_3, A_4 分别表他乘火车，轮船，汽车，飞机去上海参加会议，则 A_1, A_2, A_3, A_4

构成一个完备事件组， B 表他开会迟到，由题目已知条件可得

$$(1) P(B) = \sum_{i=1}^4 P(A_i)P(B|A_i) = 0.3 \times \frac{1}{4} + 0.2 \times \frac{1}{3} + 0.1 \times \frac{1}{12} + 0.4 \times 0 = 0.15$$

$$(2) P(A_1|B) = \frac{P(B|A_1)P(A_1)}{\sum_{i=1}^4 P(A_i)P(B|A_i)} = \frac{0.3 \times \frac{1}{4}}{0.15} = 0.5$$

$$P(A_2|B) = \frac{P(B|A_2)P(A_2)}{\sum_{i=1}^4 P(A_i)P(B|A_i)} = \frac{0.2 \times \frac{1}{3}}{0.15} = 0.4444$$

$$P(A_3|B) = \frac{P(B|A_3)P(A_3)}{\sum_{i=1}^4 P(A_i)P(B|A_i)} = \frac{0.1 \times \frac{1}{12}}{0.15} = 0.0556$$

$$P(A_4|B) = \frac{P(B|A_4)P(A_4)}{\sum_{i=1}^4 P(A_i)P(B|A_i)} = \frac{0.4 \times 0}{0.15} = 0$$

因在诸 $P(A_i|B) (i=1, 2, 3, 4)$ 中， $P(A_1|B)$ 最大，所以，若他迟到了，他最可能是乘火车去的。

2. 0.0171

设 A 表“被检验者经检验认为没有患关节炎”， B 表“被检验者患关节炎”，由贝叶斯公式有

$$P(B|A) = \frac{P(B)P(A|B)}{P(B)P(A|B) + P(\bar{B})P(A|\bar{B})} = \frac{0.15 \times 0.1}{0.15 \times 0.1 + 0.9 \times 0.96} = 0.0171$$

3. (1) 0.2157, (2) 0.4095, 0.7678

(1) 设 A_i 表“从甲箱中取出的两件产品中有 $i (i=0, 1, 2)$ 件次品”， B 表“从乙箱中取得次品”，由全概率公式有

$$P(B) = \sum_{i=0}^2 P(A_i)P(B|A_i) = \sum_{i=0}^2 \frac{C_{10}^{2-i} C_5^i}{C_{15}^2} \frac{C_{3+i}^1}{C_{17}^1} = \frac{11}{51} = 0.2157$$

(2) 设 C_1 为选自甲箱, C_2 为选自乙箱, B_i 表第 i ($i=1,2$) 次取出正品, 由全概率公式

$$\begin{aligned} P(B_1 \bar{B}_2 \cup \bar{B}_1 B_2) &= P(C_1)P(B_1 \bar{B}_2 \cup \bar{B}_1 B_2 | C_1) + P(C_2)P(B_1 \bar{B}_2 \cup \bar{B}_1 B_2 | C_2) \\ &= \frac{1}{2} \frac{C_5^1 C_{10}^1}{C_{15}^2} + \frac{1}{2} \frac{C_3^1 C_{12}^1}{C_{15}^2} = \frac{43}{105} = 0.4095 \end{aligned}$$

由条件概率公式及全概率公式有

$$\begin{aligned} P(B_1 | \bar{B}_2) &= \frac{P(B_1 \bar{B}_2)}{P(\bar{B}_2)} = \frac{P(C_1)P(B_1 \bar{B}_2 | C_1) + P(C_2)P(B_1 \bar{B}_2 | C_2)}{P(C_1)P(\bar{B}_2 | C_1) + P(C_2)P(\bar{B}_2 | C_2)} \\ &= \frac{\frac{1}{2} \times \frac{10}{15} \times \frac{5}{14} + \frac{1}{2} \times \frac{12}{15} \times \frac{3}{14}}{\frac{1}{2} \times \frac{5}{15} + \frac{1}{2} \times \frac{3}{15}} = \frac{43}{56} = 0.7679 \end{aligned}$$

4. 0.2098, 0.0621

$$(1) p = \frac{C_{40}^3 C_{10}^2}{C_{50}^5} = 0.2098$$

$$(2) p = 1 - \frac{C_{40}^5}{C_{50}^5} = 0.0621$$

第 9-10 页 《第一章综合练习》

一. 1. B 2. A

二. 1. 0.1837

设 A, B 分别表甲乙击中靶子, 则所求概率为

$$\begin{aligned} P(A\bar{B} | A \cup B) &= \frac{P(A\bar{B})}{P(A \cup B)} = \frac{P(A)P(\bar{B})}{P(A) + P(B) - P(A)P(B)} \\ &= \frac{0.9 \times (1 - 0.8)}{0.9 + 0.8 - 0.9 \times 0.8} = 0.1837 \end{aligned}$$

2. $\frac{1}{4}$

三. 1. 教材习题一(B)三第 1 题

$$\begin{aligned}
 \text{证明: } P(\bar{A}\bar{B}) &= P(\overline{A \cup B}) = 1 - P(A \cup B) = 1 - P(A) - P(B) + P(AB) \\
 &= 1 - p - (1 - \sqrt{p}) + P(AB) = \sqrt{p} - p + P(AB) \\
 &\geq \sqrt{p} - p (\because P(AB) \geq 0) > 0 (\because 0 < p < 1, \therefore \sqrt{p} > p)
 \end{aligned}$$

2. 教材习题一(B)三第4题

解答: 设 A_1, A_2, A_3 分别表在 100, 150, 200 米处击中动物, 由 $P(A_1) = 0.6 = \frac{k}{100}$

$$\begin{aligned}
 \text{得 } k &= 60, \text{ 从而得 } P(A_1 \cup \bar{A}_1 A_2 \cup \bar{A}_1 \bar{A}_2 A_3) = P(A_1) + P(\bar{A}_1)P(A_2|\bar{A}_1) \\
 &+ P(\bar{A}_1)P(\bar{A}_2|\bar{A}_1)P(A_3|\bar{A}_1\bar{A}_2) = 0.6 + 0.4 \times \frac{60}{150} + 0.4 \times \frac{90}{150} \times \frac{60}{200} \\
 &= 0.832
 \end{aligned}$$

3. 教材习题一(B)三第5题

解答: A 表“选正确答案”, B 表“知道正确答案”, 由贝叶斯公式得

$$\begin{aligned}
 P(B|A) &= \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|\bar{B})P(\bar{B})} \\
 &= \frac{p \times 1}{p \times 1 + (1-p) \times \frac{1}{m}} = \frac{mp}{mp + 1 - p}
 \end{aligned}$$

4. 教材习题一(B)三第3题

解答: 设 A 为“甲系统有效”, B 为“乙系统有效”, 则由题意有

$$P(A) = 0.92, P(B) = 0.93, P(B|\bar{A}) = 0.85,$$

从而有

$$P(AB) = P(B) - P(\bar{A}B) = P(B) - P(\bar{A})P(B|\bar{A}) = 0.962.$$

$$(1) P(A \cup B) = P(A) + P(B) - P(AB) = 0.988,$$

$$(2) P(A|\bar{B}) = \frac{P(\bar{A}\bar{B})}{P(\bar{B})} = \frac{P(A) - P(AB)}{1 - P(B)} = 0.8286.$$

5. 教材习题一(B)三第8题

解答：每个能出厂的概率为 $p = 0.7 + 0.3 \times 0.8 = 0.94$ ，所以

(1) 全部都能出厂的概率为 0.94^n ；

(2) 恰有两个不能出厂的概率为 $C_n^2 0.06^2 0.94^{n-2}$ 。

第 11-12 页 《分布函数及离散型随机变量》

一. 1. B (利用分布函数的性质判断) 2. D

二. 1. 3 或 7 (即 $P(X=2) = \frac{7}{30} = \frac{m}{10} \cdot \frac{10-m}{9}$) 2. $\frac{11}{24}$

三. 1. 教材习题二(A)三第 1 题

① 当 $x < 0$ 时, 显然有 $F(x) = P(X \leq x) = P(\emptyset) = 0$;

② 当 $0 \leq x \leq a$ 时, 由题意有 $P(X \in [0, x]) = \delta x^3$ 且 $P(X \in [0, a]) = 1$, 联立两

式解得 $\delta = a^{-3}$, 从而 $F(x) = P(X \leq x) = P(X \in [0, x]) = a^{-3}x^3$;

③ 当 $x > a$ 时, 有 $F(x) = P(X \leq x) = P(\Omega) = 1$;

$$\text{从而分布函数为 } F(x) = \begin{cases} 0, & x < 0 \\ a^{-3}x^3, & 0 \leq x \leq a, \\ 1 & x > a \end{cases}$$

$$P\left(\frac{a}{3} < X \leq \frac{2}{3}a\right) = F\left(\frac{2}{3}a\right) - F\left(\frac{a}{3}\right) = a^{-3} \left[\left(\frac{2}{3}a\right)^3 - \left(\frac{a}{3}\right)^3 \right] = \frac{7}{27}.$$

2. 教材习题二(A)三第 2 题

由分布函数在 $x=1$ 和 $x=2$ 处的右连续性有 $0=a, 1=a+b$, 解之得 $a=0, b=1$

$$P\left(X > \frac{3}{2}\right) = 1 - F\left(\frac{3}{2}\right) = 1 - \left[a + b \left(\frac{3}{2} - 1 \right)^2 \right] = 1 - \left(\frac{3}{2} - 1 \right)^2 = \frac{3}{4}$$

3. 教材习题二(A)三第 3 题

显然, X 的可能取值为 1, 2, 3 且 $P(X=1) = \frac{C_3^1}{C_4^2} = \frac{1}{2}$; $P(X=2) = \frac{C_2^1}{C_4^2} = \frac{1}{3}$;

$$P(X=3) = \frac{C_1^1}{C_4^2} = \frac{1}{6}; \text{ 故 } X \sim \begin{bmatrix} 2 & 3 & 4 \\ 1 & 1 & 1 \\ 2 & 3 & 6 \end{bmatrix}, \text{ 分布函数为 } F(x) = \begin{cases} 0, & x < 1 \\ \frac{1}{2}, & 1 \leq x < 2 \\ \frac{5}{6}, & 2 \leq x < 3 \\ 1, & x \geq 3 \end{cases}$$

4. 教材习题二(A)三第4题

显然 X 可能取 $0, 1, 2, 3$; 设 A_i 表“在第 $i (i = 1, 2, 3)$ 路口遇到红灯”, 则

$$P(A_i) = \frac{1}{2}, i = 1, 2, 3 \text{ 且 } A_1, A_2, A_3 \text{ 相互独立, 所以有}$$

$$P(X = 0) = P(A_1) = \frac{1}{2};$$

$$P(X = 1) = P(\bar{A}_1 A_2) = P(\bar{A}_1)P(A_2) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4};$$

$$P(X = 2) = P(\bar{A}_1 \bar{A}_2 A_3) = P(\bar{A}_1)P(\bar{A}_2)P(A_3) = \frac{1}{8};$$

$$P(X = 3) = P(\bar{A}_1 \bar{A}_2 \bar{A}_3) = P(\bar{A}_1)P(\bar{A}_2)P(\bar{A}_3) = \frac{1}{8}.$$

$$\text{故有 } X \sim \begin{bmatrix} 0 & 1 & 2 & 3 \\ \frac{1}{2} & \frac{1}{4} & \frac{1}{8} & \frac{1}{8} \end{bmatrix}.$$

第 13-14 页《常见离散型分布》

一. 1. B 2. D (用泊松分布近似计算)

二. 1. e^{-1} (即为 $P(\Delta > 0) = P(4 - 4X > 0) = P(X < 1) = P(X = 0) = \frac{1^0}{0!} e^{-1} = e^{-1}$)

2. $C_3^1 \times 0.18 \times 0.82^2$ (设对 X 的 3 次取值中取到 1 的次数为 ξ , 而每次取到 1 的概率为

$$P(X = 1) = C_2^1 0.1^1 0.9^1 = 0.18, \text{ 从而有 } \xi \sim B(3, 0.18), \text{ 所以所求概率为}$$

$$P(\xi = 1) = C_3^1 \times 0.18^1 \times 0.82^2 = C_3^1 \times 0.18 \times 0.82^2)$$

三. 1. 教材习题二(A)三第6题

$$(1) X \sim H(6, 4, 20), \text{ 所以 } P(X = k) = \frac{C_{16}^{6-k} C_4^k}{C_{20}^6}, k = 0, 1, 2, 3, 4, \text{ 即}$$

$$X \sim \begin{bmatrix} 0 & 1 & 2 & 3 & 4 \\ \frac{C_{16}^6 C_4^0}{C_{20}^6} & \frac{C_{16}^5 C_4^1}{C_{20}^6} & \frac{C_{16}^4 C_4^2}{C_{20}^6} & \frac{C_{16}^3 C_4^3}{C_{20}^6} & \frac{C_{16}^2 C_4^4}{C_{20}^6} \end{bmatrix}$$

$$(2) Y \sim B(6, 0.2), \text{ 所以 } P(Y = k) = C_6^k 0.2^k 0.8^{6-k}, \quad k = 0, 1, 2, 3, 4, 5, 6$$

2. 教材习题二(A)三第7题

$$\text{显然 } X \sim B(15, 0.2)$$

$$(1) P(X = 3) = C_{15}^3 0.2^3 0.8^{12} = 0.2501$$

$$(2) P(X \geq 2) = 1 - P(X = 0) - P(X = 1)$$

$$= 1 - 0.8^{15} - 15 \times 0.8^{14} \times 0.2 = 0.8329$$

$$(3) P(1 \leq X \leq 3) = P(X = 3) + P(X = 2) + P(X = 1)$$

$$= C_{15}^3 0.2^3 0.8^{12} + C_{15}^2 0.2^2 0.8^{13} + C_{15}^1 0.2^1 0.8^{14} = 0.6130$$

$$(4) P(X \leq 1) = 1 - P(X \geq 2) = 1 - 0.8329 = 0.1671$$

3. 教材习题二(A)三第9题

$$\text{设 } X \text{ 为任意时刻同时出故障的车床台数, 则 } X \sim B(300, 0.01)$$

$$P(X = 4) = C_{300}^4 0.01^4 0.99^{296} = 0.1689$$

$$\text{由泊松定理近似地有 } X \sim P(3), \text{ 所以 } P(X = 4) \doteq \frac{3^4}{4!} e^{-3} = \frac{27}{8} e^{-3} = 0.1680$$

$$\text{相对误差为 } \frac{0.1689 - 0.1680}{0.1689} = 0.533\%$$

4. 教材习题二(A)三第10题

$$(1) P(X > r) = \sum_{k=r+1}^{\infty} (1-p)^{k-1} p = \frac{p(1-p)^r}{1-(1-p)} = (1-p)^r$$

$$(2) \text{ 由(1)的结论知 } P(X > t+r) = (1-p)^{r+t}, \quad P(X > t) = (1-p)^t \text{ 从而}$$

$$\begin{aligned}\mathbf{P}(X > r+t | X > r) &= \frac{\mathbf{P}(X > r+t, X > r)}{\mathbf{P}(X > r)} = \frac{\mathbf{P}(X > r+t)}{\mathbf{P}(X > r)} \\ &= \frac{(1-p)^{r+t}}{(1-p)^r} = (1-p)^t = \mathbf{P}(X > t).\end{aligned}$$

第 15-16 页 《连续型随机变量》

一. 1. **C** (用密度函数的特征(非负性和归一性)进行检验)

2. **C**

3. **A** 因密度函数为偶函数, 则必有

$$1 = \int_{-\infty}^{+\infty} f(x) dx = 2 \int_0^{+\infty} f(x) dx = 2[\mathbf{F}(+\infty) - \mathbf{F}(0)] = 2 - 2\mathbf{F}(0),$$

从而 $\mathbf{F}(0) = \frac{1}{2}$; 所以

$$\begin{aligned}\mathbf{P}(|X| > a) &= 1 - \mathbf{P}(|X| \leq a) = 1 - \int_{-a}^a f(x) dx = 1 - 2 \int_0^a f(x) dx \\ &= 1 - 2[\mathbf{F}(a) - \mathbf{F}(0)] = 1 - 2\mathbf{F}(a) + 2\mathbf{F}(0) = 2[1 - \mathbf{F}(a)]\end{aligned}$$

二. 1. $\frac{2}{\pi}$

$$\text{由归一性得 } 1 = \int_{-\infty}^{+\infty} f(x) dx = \int_0^1 \frac{A}{\sqrt{1-x^2}} dx = A \arcsin x \Big|_0^1 = \frac{\pi}{2} A \Rightarrow A = \frac{2}{\pi}$$

2. $\frac{15}{8}\sqrt{\pi}$

$$\begin{aligned}\int_{-\infty}^{+\infty} x^{\frac{3}{2}} f(x) dx &= \int_0^{+\infty} x^{\frac{3}{2}} \frac{1^2}{\Gamma(2)} x^{2-1} e^{-x} dx = \int_0^{+\infty} x^{\frac{7}{2}-1} e^{-x} dx = \Gamma\left(\frac{7}{2}\right) \\ &= \frac{5}{2} \times \frac{3}{2} \times \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = \frac{15}{8} \sqrt{\pi}\end{aligned}$$

3. **B**(3, e^{-1})

$$X \sim e(1), \text{ 故 } X \text{ 的分布函数为 } \mathbf{F}(x) = \begin{cases} 1 - e^{-x}, & x > 0 \\ 0, & x \leq 0 \end{cases},$$

电子元件寿命大于 1 万小时的概率为 $p = \mathbf{P}(X \geq 1) = 1 - \mathbf{F}(1) = e^{-1}$

所以有 $\mathbf{Y} \sim \mathbf{B}(\mathbf{3}, e^{-1})$.

三. 1. 教材习题二(A) 三第 12 题

首先函数 $\varphi(x)$ 满足非负性;

其次证明存在 c 使得函数 $\varphi(x)$ 满足归一性: 由 $\int_{-\infty}^{+\infty} \varphi(x) dx = \int_0^{+\infty} \frac{x^2}{c^2} e^{-\frac{x^3}{c}} dx$
 $= \frac{1}{3c} \int_0^{+\infty} e^{-\left(\frac{x^3}{c}\right)} d\left(\frac{x^3}{c}\right)$ (此处应需 $c > 0$) $= \frac{1}{3c} \Gamma(1) = \frac{1}{3c}$, 这说明当 $c = \frac{1}{3}$ 时函数 $\varphi(x)$ 满足归一性;

所以, 当 $c = \frac{1}{3}$ 时函数 $\varphi(x)$ 为某连续型随机变量的密度函数. 此时

$$P(X \leq 1) = \int_0^1 9x^2 e^{-3x^3} dx = \int_0^1 e^{-(3x^3)} d(3x^3) = -e^{-3x^3} \Big|_{x=0}^{x=1} = 1 - e^{-3}$$

2. 教材习题二(A) 三第 14 题

$$\begin{aligned} (1) \text{ 由归一性有 } 1 &= \int_{-\infty}^{+\infty} f(x) dx = \int_{-A}^A \frac{2}{\pi(1+x^2)} dx = \frac{4}{\pi} \int_0^A \frac{1}{1+x^2} dx \\ &= \frac{4}{\pi} \arctan x \Big|_0^A = \frac{4}{\pi} \arctan A, \text{ 所以 } \arctan A = \frac{\pi}{4}, \text{ 从而 } A = 1 \end{aligned}$$

$$(2) F(x) = \int_{-\infty}^x f(t) dt = \begin{cases} 0, & x \leq -1 \\ \int_{-1}^x \frac{2}{\pi(1+t^2)} dt = \frac{2}{\pi} \arctan x + \frac{1}{2}, & -1 < x < 1 \\ 1, & x \geq 1 \end{cases}$$

3. 教材习题二(A) 三第 15 题

$$(1) \text{ 由归一性有 } 1 = \int_{-\infty}^{+\infty} f(x) dx = \int_0^2 \frac{1}{8} dx + \int_2^4 kx dx = \frac{1}{4} + 6k, \text{ 所以 } k = \frac{1}{8};$$

$$(2) F(x) = \int_{-\infty}^x f(t) dt = \begin{cases} 0, & x \leq 0 \\ \int_0^x \frac{1}{8} dt = \frac{1}{8}x, & 0 < x < 2 \\ \int_0^2 \frac{1}{8} dt + \int_2^x \frac{1}{8} t dt = \frac{1}{16}x^2, & 2 \leq x < 4 \\ 1, & x \geq 4 \end{cases}$$

4. 教材习题二(A)三第 16 题

$$\text{显然 } X \text{ 的分布函数为 } F(x) = \begin{cases} 1 - e^{-\frac{x}{1000}}, & x > 0 \\ 0, & x \leq 0 \end{cases}, \text{ 每只元件寿命不超过 400 小时的概}$$

$$\text{率为 } p = P(X \leq 400) = F(400) = 1 - e^{-\frac{1}{1000} \times 400} = 1 - e^{-0.4}$$

设在仪器使用的最初 400 小时内元件的损坏数, 则 $Y \sim B(6, 1 - e^{-0.4})$, 从而有

$$(1) P(Y = 1) = C_6^1 \times (1 - e^{-0.4}) e^{-0.4 \times 5} = 6e^{-2} (1 - e^{-0.4});$$

$$(2) P(Y \geq 1) = 1 - P(Y = 0) = 1 - e^{-0.4 \times 6} = 1 - e^{-2.4}$$

第 17-18 页《随机变量函数的分布》

$$\text{一. 1. D } (F_Y(y) = P(Y \leq y) = P(3X - 1 \leq y) = P(X \leq \frac{y+1}{3}) = F(\frac{y+1}{3}))$$

2. C

$$\text{二. 1. } f_Y(y) = \frac{1}{4} y^{-\frac{1}{2}}, \quad 0 < y < 4$$

$$\text{对任意 } y \in (0, 4), \quad F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y})$$

$$= P(0 \leq X \leq \sqrt{y}) = \int_0^{\sqrt{y}} \frac{1}{2} dx = \frac{1}{2} \sqrt{y}, \text{ 所以 } f_Y(y) = F'_Y(y) = \frac{1}{4} y^{-\frac{1}{2}}$$

也可以直接利用平方变换的公式求解

$$2. P(Y \geq 2) = P(1 - \sqrt[3]{X} \geq 2) = P(X \leq -1) = F_X(-1) = 0$$

三. 1. 教材习题二(A)三第 18 题

$$(1) X \sim \begin{bmatrix} -2 & -1 & 2 & 3 \\ 0.2 & 0.3 & 0.2 & 0.3 \end{bmatrix} \quad (2) Y \sim \begin{bmatrix} 1 & 4 & 9 \\ 0.3 & 0.4 & 0.3 \end{bmatrix}$$

2. 教材习题二(A) 三第 19 题

$$(1) F_Y(y) = P(Y \leq y) = P(2X - 1 \leq y) = P\left(X \leq \frac{y+1}{2}\right) = F_X\left(\frac{y+1}{2}\right)$$

$$f_Y(y) = F'_Y(y) = \frac{1}{2} f_X\left(\frac{y+1}{2}\right) = \begin{cases} \frac{1}{2} e^{-\frac{y+1}{2}}, & y > -1, \\ 0, & y \leq -1 \end{cases}$$

$$(2) F_Y(y) = P(Y \leq y) = P(e^X \leq y) = \begin{cases} 0, & y \leq 1 \\ P(X \leq \ln y) = F_X(\ln y), & y > 1 \end{cases}$$

$$f_Y(y) = F'_Y(y) = \begin{cases} 0, & y \leq 1 \\ F'_X(\ln y) = \frac{1}{y} e^{-\ln y} = \frac{1}{y^2}, & y > 1 \end{cases}$$

$$(3) F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = \begin{cases} P(X \in [-\sqrt{y}, \sqrt{y}]), & y > 0 \\ 0, & y \leq 0 \end{cases}$$

$$= \begin{cases} F_X(\sqrt{y}) - F_X(-\sqrt{y}), & y > 0 \\ 0, & y \leq 0 \end{cases}$$

$$f_Y(y) = F'_Y(y) = \begin{cases} F'_X(\sqrt{y}) - F'_X(-\sqrt{y}), & y > 0 \\ 0, & y \leq 0 \end{cases}$$

$$= \begin{cases} \frac{1}{2\sqrt{y}} [f_X(\sqrt{y}) + f_X(-\sqrt{y})] = \frac{1}{2\sqrt{y}} e^{-\sqrt{y}}, & y > 0 \\ 0, & y \leq 0 \end{cases}$$

3. 教材习题二(A) 三第 22 题

$$F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = \begin{cases} 0, & y \leq 0 \\ P(X \in [-\sqrt{y}, \sqrt{y}]), & y \in (0, 9) \\ 1, & y \geq 9 \end{cases}$$

$$= \begin{cases} 0, & y \leq 0 \\ F_X(\sqrt{y}) - F_X(-\sqrt{y}), & y \in (0, 9) \\ 1, & y \geq 9 \end{cases}$$

$$f_Y(y) = F'_Y(y) = \begin{cases} \frac{1}{2\sqrt{y}} [f_X(\sqrt{y}) + f_X(-\sqrt{y})], & y \in (0, 9) \\ 0, & y \notin (0, 9) \end{cases}$$

$$= \begin{cases} \frac{1}{2\sqrt{y}} \left[\frac{1}{4} + \frac{1}{4} \right] = \frac{1}{4\sqrt{y}}, & y \in (0, 1] \\ \frac{1}{2\sqrt{y}} \left[0 + \frac{1}{4} \right] = \frac{1}{8\sqrt{y}}, & y \in (1, 9) \\ 0, & y \notin (0, 9) \end{cases}$$

第 19-20 页 《第二章综合练习》

一. 1. C (用分布函数的特征验证. 注意第二个答案, 若 $a=2, b=-1$ 不能保证

$F(x) = aF_1(x) + bF_2(x)$ 的非负性)

2. B 教材习题二(B)一第 4 题

$X \sim e(\lambda), Y = \min(X, 2)$, 显然可见 Y 的有效值域为 $R(Y) = (0, 2]$, 所以, 当

$y \in (0, 2)$ 时, " $Y \leq y$ " \Leftrightarrow " $\min(X, 2) \leq y$ " \Leftrightarrow " $X \leq y$ ", 从而

$F_Y(y) = P(X \leq y) = F_X(y) = 1 - e^{-\lambda y}$, 于是綜上有

$$F_Y(y) = \begin{cases} 0, & y \leq 0 \\ 1 - e^{-\lambda y}, & y \in (0, 2) \\ 1, & y \geq 2 \end{cases}$$

显然可见, $F_Y(y)$ 在 $y=0$ 处连续, 在 $y=2$ 处间断.

(本题中的随机变量 Y 是非离散非连续型随机变量)

二. 1. $\frac{9}{64}$

因 $P(X \leq 1/2) = \int_0^{1/2} 2x dx = \frac{1}{4}$, 则 $Y \sim B(3, 1/4)$, 故

$$P(Y=2) = C_3^2 \left(\frac{1}{4} \right)^2 \frac{3}{4} = \frac{9}{64}$$

$$2. \frac{16\sqrt{2}}{3\sqrt{\pi}}$$

由归一性得

$$1 = \int_{-\infty}^{+\infty} f(x) dx = A \int_0^{+\infty} x^{3/2} e^{-2x} dx = \frac{A}{2^{5/2}} \int_0^{+\infty} (2x)^{3/2} e^{-2x} d(2x) = \frac{A}{2^{5/2}} \Gamma(5/2)$$

$$\text{所以 } A = \frac{2^{\frac{5}{2}}}{\Gamma\left(\frac{5}{2}\right)} = \frac{4\sqrt{2}}{\frac{3}{4}\sqrt{\pi}} = \frac{16\sqrt{2}}{3\sqrt{\pi}}$$

$$\text{三. 1. (1)} \frac{9}{10} \quad (2) 1 - 0.9^{10} - C_{10}^1 0.9^9 \times 0.1 = 1 - 1.9 \times 0.9^9 \quad (3) 1 - 11e^{-10}$$

$$(1) P(|X| < 1.8) = \frac{3.6}{4} = 0.9$$

(2) 设 Y 表 10 次测量中误差绝对值大于 1.8 的次数, 则 $Y \sim B(10, 0.1)$, 从而

$$P(Y \geq 2) = 1 - P(Y = 0) - P(Y = 1) = 1 - 0.9^{10} - C_{10}^1 0.9^9 \times 0.1$$

(3) 设 Z 表 100 次测量中误差绝对值大于 1.8 的次数, 则 $Z \sim B(100, 0.1)$, 故近似地有

$$Z \sim P(10), \text{ 从而 } P(Z \geq 2) = 1 - P(Z = 0) - P(Z = 1) = 1 - 11e^{-10}$$

2. 教材习题二(B)三第 4 题

显然密度函数是偶函数, 所以

$$(1) \text{ 由归一性得 } 1 = \int_{-\infty}^{+\infty} f(x) dx = 2A \int_0^{+\infty} \frac{1}{e^x + e^{-x}} dx, \text{ 作变量代换 } y = e^x \text{ 可得}$$

$$1 = 2A \int_1^{+\infty} \frac{1}{1+y^2} dy = 2A \arctan y \Big|_1^{+\infty} = \frac{1}{2} \pi A \Rightarrow A = \frac{2}{\pi}$$

$$(2) \text{ 对任意 } x \in R, \quad F_X(x) = \int_{-\infty}^x f(t) dt = \frac{2}{\pi} \int_{-\infty}^x \frac{1}{e^t + e^{-t}} dt, \text{ 作变量代换 } y = e^t \text{ 可得}$$

$$F_X(x) = \frac{2}{\pi} \int_{-\infty}^x \frac{1}{e^t + e^{-t}} dt = \frac{2}{\pi} \arctan y \Big|_{y=0}^{y=e^x} = \frac{2}{\pi} \arctan e^x,$$

$$\text{于是有 } P(0 \leq X \leq 1) = F_X(1) - F_X(0) = \frac{2}{\pi} \arctan e - \frac{1}{2}$$

$$(3) \quad F_Y(y) = P(Y \leq y) = P(e^{-|X|} \leq y) = \begin{cases} 0, & y \leq 0 \\ P(|X| \geq -\ln y), & y \in (0,1) \\ 1, & y \geq 1 \end{cases}$$

$$= \begin{cases} 0, & y \leq 0 \\ \frac{4}{\pi} \int_{-\ln y}^{+\infty} \frac{1}{e^t + e^{-t}} dt, & y \in (0,1) \\ 1, & y \geq 1 \end{cases}$$

$$\text{从而有 } f_Y(y) = F'_Y(y) = \begin{cases} \frac{4}{\pi(1+y^2)}, & y \in (0,1) \\ 0, & y \notin (0,1) \end{cases}$$

3. 教材习题二(B)三第8题

$$\text{显然, } X \text{ 的分布函数为 } F_X(x) = \begin{cases} 0, & x \leq 0 \\ 1 - e^{-\frac{1}{60}x}, & x > 0 \end{cases};$$

设 A_1, A_2, A_3 分别表年龄在 15 岁以下, 15 到 50 岁, 50 岁以上, 则 A_1, A_2, A_3 构成一完备事件组且

$$P(A_1) = F_X(15) = 1 - e^{-\frac{1}{4}},$$

$$P(A_2) = F_X(50) - F_X(15) = \left(1 - e^{-\frac{50}{60}}\right) - \left(1 - e^{-\frac{15}{60}}\right) = e^{-\frac{1}{4}} - e^{-\frac{5}{6}}$$

$$P(A_3) = 1 - F_X(50) = 1 - \left(1 - e^{-\frac{50}{60}}\right) = e^{-\frac{5}{6}}$$

再设 B 为某人得重病, 则

$$(1) \quad P(B) = \sum_{i=1}^3 P(A_i)P(B|A_i) = \left[1 - e^{-\frac{1}{4}}\right] \times 0.1 + \left[e^{-\frac{1}{4}} - e^{-\frac{5}{6}}\right] \times 0.02 + e^{-\frac{5}{6}} \times 0.2$$

$$= 0.1 - 0.08e^{-\frac{1}{4}} + 0.18e^{-\frac{5}{6}} \approx 0.1159$$

$$(2) \quad P(A_1|B) = \frac{P(A_1)P(B|A_1)}{P(B)} = \frac{\left[1 - e^{-\frac{1}{4}}\right] \times 0.1}{0.1159} = \frac{0.02212}{0.1159} = 0.1908$$

$$P(A_2|B) = \frac{P(A_2)P(B|A_2)}{P(B)} = \frac{\left[e^{-\frac{1}{4}} - e^{-\frac{5}{6}}\right] \times 0.02}{0.1159} = 0.0444$$

$$P(A_3|B) = 1 - 0.1908 - 0.0444 = 0.7648$$

所以若某人得病，他的年龄最可能的是 50 岁以上.

第 21-22 页 《二维随机变量》

一. 1. B 2. C

二. 1. $F(b, c) - F(a, c)$ 2. 3 3. $p = \frac{10!}{1!5!3!1!} 0.07 \times 0.43^5 \times 0.35^3 \times 0.15$

三. 1. 教材习题三(A)三第 1 题

显然 X_1, X_2 都只能取 0, 1, 且

$$P(X_1 = 0, X_2 = 0) = P(Y \leq 1, Y \leq 2) = P(Y \leq 1) = F_Y(1) = 1 - e^{-1};$$

$$P(X_1 = 0, X_2 = 1) = P(Y \leq 1, Y > 2) = P(\emptyset) = 0;$$

$$\begin{aligned} P(X_1 = 1, X_2 = 0) &= P(Y > 1, Y \leq 2) = P(1 < Y \leq 2) \\ &= F_Y(2) - F_Y(1) = e^{-1} - e^{-2}; \end{aligned}$$

$$P(X_1 = 1, X_2 = 1) = P(Y > 1, Y > 2) = P(Y > 2) = e^{-2};$$

故 (X_1, X_2) 的联合分布律为

$X \setminus Y$	0	1
0	$1 - e^{-1}$	0
1	$e^{-1} - e^{-2}$	e^{-2}

2. 教材习题三(A)三第 3 题

显然, X 可取 0, 1, 2, 3, Y 可取 0, 1, 2, 且

$$P(X = 0, Y = 0) = P(\emptyset) = 0; \quad P(X = 0, Y = 1) = P(\emptyset) = 0;$$

$$P(X = 0, Y = 2) = \frac{C_3^0 C_2^2 C_2^2}{C_7^4} = \frac{1}{35}; \quad P(X = 1, Y = 0) = P(\emptyset) = 0;$$

$$P(X = 1, Y = 1) = \frac{C_3^1 C_2^1 C_2^2}{C_7^4} = \frac{6}{35}; \quad P(X = 1, Y = 2) = \frac{C_3^1 C_2^2 C_2^1}{C_7^4} = \frac{6}{35};$$

$$P(X = 2, Y = 0) = \frac{C_3^2 C_2^0 C_2^2}{C_7^4} = \frac{3}{35}; \quad P(X = 2, Y = 1) = \frac{C_3^2 C_2^1 C_2^1}{C_7^4} = \frac{12}{35};$$

$$P(X=2, Y=2) = \frac{C_3^2 C_2^2 C_2^0}{C_7^4} = \frac{3}{35}; \quad P(X=3, Y=0) = \frac{C_3^3 C_2^0 C_2^1}{C_7^4} = \frac{2}{35};$$

$$P(X=3, Y=1) = \frac{C_3^3 C_2^1 C_2^0}{C_7^4} = \frac{2}{35}; \quad P(X=3, Y=2) = P(\emptyset) = 0.$$

则 (X, Y) 的联合分布律为

$X \setminus Y$	0	1	2
0	0	0	1/35
1	0	6/35	6/35
2	3/35	12/35	3/35
3	2/35	2/35	0

3. 教材习题三(A) 三第 4 题

$$(1) \quad 1 = \iint_{R^2} f(x, y) dx dy = \int_0^{+\infty} dx \int_0^{+\infty} c e^{-(2x+4y)} dy = \frac{c}{8} \Rightarrow c = 8$$

$$(2) \quad P(X > 2) = P(X > 2, Y < +\infty) = \int_2^{+\infty} dx \int_{-\infty}^{+\infty} f(x, y) dy$$

$$= 8 \int_2^{+\infty} dx \int_0^{+\infty} e^{-(2x+4y)} dy = \int_2^{+\infty} 2e^{-2x} dx \int_0^{+\infty} 4e^{-4y} dy = e^{-4};$$

$$P(X > Y) = \int_0^{+\infty} 2e^{-2x} dx \int_0^x 4e^{-4y} dy = \int_0^{+\infty} (1 - e^{-4x}) 2e^{-2x} dx$$

$$= \int_0^{+\infty} 2e^{-2x} dx - \frac{1}{3} \int_0^{+\infty} e^{-6x} dx = 1 - \frac{1}{3} = \frac{2}{3};$$

$$P(X + Y < 1) = \int_0^1 2e^{-2x} dx \int_0^{1-x} 4e^{-4y} dy = \int_0^1 (1 - e^{-4(1-x)}) 2e^{-2x} dx$$

$$= \int_0^1 e^{-2x} d(2x) - e^{-4} \int_0^1 e^{2x} d(2x) = 1 - 2e^{-2} + e^{-4};$$

(3) 显然, 当 $x \leq 0$ 或 $y \leq 0$ 时, 必有 $F(x, y) = P(X \leq x, Y \leq y) = 0$;

$$\text{而当 } x > 0, y > 0 \text{ 时, } F(x, y) = \int_{-\infty}^x dt \int_{-\infty}^y f(t, s) ds$$

$$= \int_0^x 2e^{-2t} dt \int_0^y 4e^{-4s} ds = (1 - e^{-2x})(1 - e^{-4y})$$

$$\text{综上得 } F(x, y) = \begin{cases} (1 - e^{-2x})(1 - e^{-4y}), & x > 0, y > 0 \\ 0, & \text{others} \end{cases}.$$

第 23-24 页 《边缘分布、边缘密度及独立性》

一. 1. **D** (利用分布函数的性质判断) 2. **B** (利用密度函数的特征判断) 3. **C**

二. 1. **0** 2. $F(x)(1 - F(y))$ 3. $\frac{1}{2}$

三. 1. 教材习题三(A)三第 6 题

$X \setminus Y$	y_1	y_2	y_3	$P(X = x_i) = p_{i.}$
x_1	1/24	1/8	1/12	1/4
x_2	1/8	3/8	1/24	3/4
$P(Y = y_j) = p_{.j}$	1/6	1/2	1/3	1

2. 教材习题三(A)三第 8 题

$$\begin{aligned} (1) \quad 1 &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^{\pi/2} dx \int_0^1 \frac{A \cos x}{\sqrt{1-y^2}} dy = A \int_0^{\pi/2} \cos x dx \int_0^1 \frac{1}{\sqrt{1-y^2}} dy \\ &= A(\pi/2) \Rightarrow A = \frac{2}{\pi}; \end{aligned}$$

$$\begin{aligned} (2) \quad f_X(x) &= \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \frac{2}{\pi} \cos x \int_0^1 \frac{1}{\sqrt{1-y^2}} dy, & 0 < x < \frac{\pi}{2} \\ 0, & \text{others} \end{cases} \\ &= \begin{cases} \cos x, & 0 < x < \pi/2 \\ 0, & \text{others} \end{cases}; \end{aligned}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \frac{2}{\pi \sqrt{1-y^2}} \int_0^{\pi/2} \cos x dy, & 0 < y < 1 \\ 0, & \text{others} \end{cases}$$

$$= \begin{cases} \frac{2}{\pi\sqrt{1-y^2}}, & 0 < y < 1; \\ 0, & \text{others} \end{cases}$$

$$(3) \quad P\left(X \leq \frac{\pi}{3}\right) = \int_{-\infty}^{\pi/3} f_X(x) dx = \int_0^{\pi/3} \cos x dx = \frac{\sqrt{3}}{2};$$

$$P\left(Y \geq \frac{1}{2}\right) = \int_{1/2}^{+\infty} f_Y(y) dy = \int_{1/2}^1 \frac{2}{\pi\sqrt{1-y^2}} dy = \frac{2}{3};$$

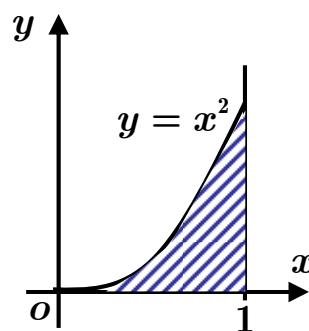
(4) 由 3 个密度函数可知 $f(x, y) = f_X(x)f_Y(y), \forall x, y \in R$, 所以 X 与 Y 相互独立!

3. 教材习题三(A) 三第 9 题

如图, 密度函数不为零的区域即图中阴影部分

$$\Omega = \{(x, y) : 0 < x < 1, 0 < y < x^2\},$$

其面积为 $m(\Omega) = \int_0^1 x^2 dx = \frac{1}{3}$, 所以



(1) 联合密度函数为

$$f(x, y) = \begin{cases} 3, & 0 < x < 1, 0 < y < x^2 \\ 0, & \text{others} \end{cases}$$

(2) 边缘密度函数为

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^{x^2} 3 dy = 3x^2, & 0 < x < 1 \\ 0, & \text{others} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{\sqrt{y}}^1 3 dx = 3(1 - \sqrt{y}), & 0 < y < 1 \\ 0, & \text{others} \end{cases}$$

(3) 在公共连续点 $(1/2, 1/8)$ 处, $f_X(x)f_Y(y) \neq f(x, y)$, 所以 X 与 Y 不相互独立!

第 25-26 页 《条件分布》

1. 教材习题三(A) 三第 10 题

$X \setminus Y$	0	1	2
0	0	0	1/35
1	0	6/35	6/35
2	3/35	12/35	3/35
3	2/35	2/35	0

前面已经得到

, 由此易得

$$(1) \quad X \sim \begin{bmatrix} 0 & 1 & 2 & 3 \\ 1/35 & 12/35 & 18/35 & 4/35 \end{bmatrix} \quad Y \sim \begin{bmatrix} 0 & 1 & 2 \\ 5/35 & 20/35 & 10/35 \end{bmatrix}$$

$$(2) \quad P(X=1|Y=2) = \frac{P(X=1, Y=2)}{P(Y=2)} = \frac{6/35}{10/35} = \frac{3}{5};$$

$$P(Y=1|X=2) = \frac{P(X=2, Y=1)}{P(X=2)} = \frac{12/35}{18/35} = \frac{2}{3};$$

$$\begin{aligned} P(Y=1|X \neq 2) &= \frac{P(X \neq 2, Y=1)}{P(X \neq 2)} = \frac{P(Y=1) - P(X=2, Y=1)}{1 - P(X=2)} \\ &= \frac{20/35 - 12/35}{1 - 18/35} = \frac{8}{17}; \end{aligned}$$

$$(3) \quad X|Y=2 \sim \begin{bmatrix} 0 & 1 & 2 & 3 \\ 1/10 & 6/10 & 3/10 & 0 \end{bmatrix} = \begin{bmatrix} 0 & 1 & 2 \\ 1/10 & 6/10 & 3/10 \end{bmatrix}$$

(4) 由于 $P(X=2, Y=3) \neq P(X=2)P(Y=3)$, 所以 X 与 Y 不独立!

2. 教材习题三(A)三第 11 题

由前面的计算结果知,

$$\text{当 } 0 < y < 1 \text{ 时, } f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{3}{3(1-\sqrt{y})} = \frac{1}{1-\sqrt{y}}, & \sqrt{y} < x < 1 \\ 0, & \text{others} \end{cases}$$

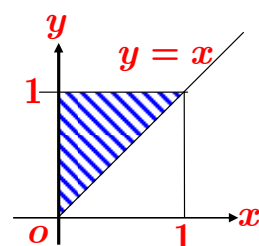
$$\text{当 } 0 < x < 1 \text{ 时, } f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} \frac{3}{3x^2} = \frac{1}{x^2}, & 0 < y < x^2 \\ 0, & \text{others} \end{cases}$$

$$\text{显然 } f_{X|Y}(x|1/4) = \frac{f(x, 1/4)}{f_Y(1/4)} = \begin{cases} \frac{3}{3(1-\sqrt{1/4})} = 2, & 1/2 < x < 1 \\ 0, & \text{others} \end{cases}, \text{ 所以}$$

$$P\left(X < 2/3 \mid Y = 1/4\right) = \int_{-\infty}^{2/3} f_{X|Y}\left(x \mid 1/4\right) dx = \int_{1/2}^{2/3} 2 dx = \frac{1}{3};$$

$$P\left(X < 2/3 \mid Y > 1/4\right) = \frac{P\left(X < 2/3, Y > 1/4\right)}{P\left(Y > 1/4\right)} = \frac{\int_{-\infty}^{2/3} dx \int_{1/4}^{+\infty} f(x, y) dy}{\int_{1/4}^{+\infty} f_Y(y) dy}$$

$$= \frac{\int_{1/2}^{2/3} dx \int_{1/4}^{x^2} 3 dy}{\int_{1/4}^1 3(1 - \sqrt{y}) dy} = \frac{\frac{5}{108}}{\frac{1}{2}} = \frac{5}{54}$$



3. 教材习题三(A)三第 14 题

如图,

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \begin{cases} \int_x^1 3y dy = \frac{3}{2}(1 - x^2), & 0 < x < 1 \\ 0, & \text{others} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \begin{cases} \int_0^y 3y dx = 3y^2, & 0 < y < 1 \\ 0, & \text{others} \end{cases}$$

显然可见, $f_X(x)f_Y(y) \neq f(x, y)$, 所以 X 与 Y 不相互独立!

$$\text{当 } 0 < x < 1 \text{ 时, } f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \begin{cases} \frac{3y}{3(1 - x^2)/2} = \frac{2y}{1 - x^2}, & x < y < 1 \\ 0, & \text{others} \end{cases}$$

$$\text{当 } 0 < y < 1 \text{ 时, } f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{3y}{3y^2} = \frac{1}{y}, & 0 < x < y \\ 0, & \text{others} \end{cases}$$

4. 教材习题三(A)三第 12 题

由题意知,

$$\text{当 } 0 < x < 1 \text{ 时, } f_{Y|X}(y|x) = \begin{cases} 1/x, & 0 < y < x \\ 0, & \text{others} \end{cases},$$

从而有

$$f(x, y) = f_X(x) f_{Y|X}(y|x) = \begin{cases} 3x, & 0 < x < 1, 0 < y < x \\ 0, & \text{others} \end{cases}$$

故有

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \begin{cases} \int_y^1 3x dx = \frac{3}{2}(1 - y^2), & 0 < y < 1 \\ 0, & \text{others} \end{cases}$$

$$\text{于是有 } P(Y < 1/2) = \int_{-\infty}^{1/2} f_Y(y) dy = \int_0^{1/2} \frac{3}{2}(1 - y^2) dy = \frac{11}{16}$$

第 27-28 页 《二维随机变量函数的分布》

一. 1. 0.64 2. $p^2 + (1 - p)^2$

二. 1. 教材习题三(A) 三第 13 题

前面已经得到

$X \setminus Y$	0	1	2
0	0	0	1/35
1	0	6/35	6/35
2	3/35	12/35	3/35
3	2/35	2/35	0

, 由此易得

(X, Y)	(0,2)	(1,1)	(1,2)	(2,0)	(2,1)	(2,2)	(3,0)	(3,1)
p_{ij}	1/35	6/35	6/35	3/35	12/35	3/35	2/35	2/35

(X, Y)	(0,2)	(1,1)	(1,2)	(2,0)	(2,1)	(2,2)	(3,0)	(3,1)
p_{ij}	1/35	6/35	6/35	3/35	12/35	3/35	2/35	2/35
$\Rightarrow Z = X + Y$	2	2	3	2	3	4	3	4
$U = \max(X, Y)$	2	1	2	2	2	2	3	3
$V = \min(X, Y)$	0	1	1	0	1	2	0	1

$$\Rightarrow Z \sim \begin{bmatrix} 2 & 3 & 4 \\ 10/35 & 20/35 & 5/35 \end{bmatrix} \qquad U \sim \begin{bmatrix} 1 & 2 & 3 \\ 6/35 & 25/35 & 4/35 \end{bmatrix}$$

$$V \sim \begin{bmatrix} 0 & 1 & 2 \\ 6/35 & 26/35 & 3/35 \end{bmatrix}$$

2. 教材习题三(A)三第 14(3) 题

显然, $R(Z) = (0, 1)$, 对任意 $z \in (0, 1)$, 有

$$\begin{aligned} F_Z(z) &= P(Z \leq z) = P(Y - X \leq z) = \iint_{y-x \leq z} f(x, y) dx dy \\ &= \int_0^z dy \int_0^y 3y dx + \int_z^1 dy \int_{y-z}^y 3y dx = \int_0^z 3y^2 dy + \int_z^1 3yz dy \\ &= z^3 + \frac{3}{2} z (1 - z^2) = \frac{3}{2} z - \frac{1}{2} z^3 \end{aligned}$$

此时, Z 的密度函数为 $f_Z(z) = F'_Z(z) = \frac{3}{2}(1 - z^2)$, 综上得 Z 的密度函数为

$$f_Z(z) = \begin{cases} \frac{3}{2}(1 - z^2), & z \in (0, 1) \\ 0 & z \notin (0, 1) \end{cases}.$$

3. 教材习题三(A)三第 15 题

显然, $R(Z) = (0, 1)$, 对任意 $z \in (0, 1)$, 有

$$\begin{aligned} F_Z(z) &= P(Z \leq z) = P(XY \leq z) = \iint_{xy \leq z} f(x, y) dx dy \\ &= \int_0^z dx \int_0^1 dy + \int_z^1 dx \int_0^{z/x} dy = \int_0^z dx + \int_z^1 \left(z/x\right) dx = z - z \ln z \end{aligned}$$

此时, Z 的密度函数为 $f_Z(z) = F'_Z(z) = -\ln z$, 综上得 Z 的密度函数为

$$f_Z(z) = \begin{cases} -\ln z, & z \in (0, 1) \\ 0 & z \notin (0, 1) \end{cases}.$$

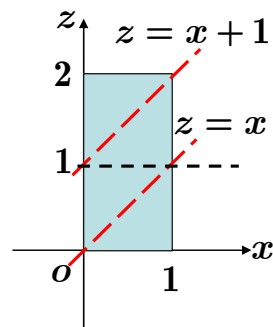
4. 教材习题三(A)三第 16 题

显然, $R(Z) = [0, 2]$, 对任意 $z \in (0, 2)$, 有 $f_Z(z) = \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx$, 要使

被积函数不为零, 需 $\begin{cases} 0 \leq x \leq 1 \\ 0 \leq z-x \leq 1 \end{cases}$, 即 $\begin{cases} 0 \leq x \leq 1 \\ x \leq z \leq x+1 \end{cases}$,

如图, 从而有

$$\begin{aligned}
 (1) \quad & \text{当 } z \in [0, 1] \text{ 时, } f_Z(z) = \int_0^z 1 dx = z; \\
 (2) \quad & \text{当 } z \in [1, 2] \text{ 时, } f_Z(z) = \int_{z-1}^1 1 dx = 2 - z \\
 \text{综上得 } Z \text{ 的密度函数为 } f_Z(z) &= \begin{cases} z, & z \in [0, 1] \\ 2 - z, & z \in [1, 2] \\ 0, & \text{others} \end{cases}
 \end{aligned}$$



5. 教材习题三(A)三第 17 题

显然, $R(Z) = (0, +\infty)$, 对任意 $z > 0$, 有

$$\begin{aligned}
 F_Z(z) &= P(Z \leq z) = P(X/Y \leq z) = \iint_{x/y \leq z} f(x, y) dx dy \\
 &= \int_0^{+\infty} dy \int_0^{yz} e^{-x-y} dx = \int_0^{+\infty} e^{-y} dy \int_0^{yz} e^{-x} dx = \int_0^{+\infty} e^{-y} (1 - e^{-yz}) dy \\
 &= \int_0^{+\infty} e^{-y} dy - \int_0^{+\infty} e^{-y(z+1)} dy = 1 - \frac{1}{z+1}
 \end{aligned}$$

此时, Z 的密度函数为 $f_Z(z) = F'_Z(z) = (1+z)^{-2}$, 综上得 Z 的密度函数为

$$f_Z(z) = \begin{cases} (1+z)^{-2}, & z > 0 \\ 0, & z \leq 0 \end{cases}.$$

6. 教材习题三(A)三第 21 题

$$\text{显然 } T_i \text{ 的分布函数为 } F_i(t) = F(t) = \begin{cases} 1 - e^{-0.2t}, & t > 0 \\ 0, & t \leq 0 \end{cases}, \quad i = 1, 2, \dots, 5.$$

(1) 并联时, 系统的寿命 $T_{\text{并}} = \max_{1 \leq i \leq 5} \{X_i\}$, 其分布函数为

$$F_{\text{并}}(t) = (F(t))^5 = \begin{cases} (1 - e^{-0.2t})^5, & t > 0 \\ 0, & t \leq 0 \end{cases},$$

$$\text{从而其密度函数为 } f_{\text{并}}(t) = \begin{cases} e^{-0.2t} (1 - e^{-0.2t})^4, & t > 0 \\ 0, & t \leq 0 \end{cases};$$

使用寿命大于 1 万小时的概率为

$$P(T_{\text{并}} > 1) = 1 - F_{\text{并}}(1) = 1 - (1 - e^{-0.2})^5 = 0.9998;$$

(2) 串联时, 系统的寿命 $T_{\text{串}} = \min_{1 \leq i \leq 5} \{X_i\}$, 其分布函数为

$$F_{\text{串}}(t) = 1 - (1 - F(t))^5 = \begin{cases} 1 - e^{-t}, & t > 0 \\ 0, & t \leq 0 \end{cases},$$

$$\text{从而其密度函数为 } f_{\text{串}}(t) = \begin{cases} e^{-t}, & t > 0 \\ 0, & t \leq 0 \end{cases};$$

使用寿命大于 1 万小时的概率为

$$P(T_{\text{串}} > 1) = 1 - F_{\text{串}}(1) = 1 - (1 - e^{-1}) = e^{-1} = 0.3679.$$

第 29-30 页 《第三章综合练习》

一. 1. 教材习题三(B)一第 2 题 (A)

2. 教材习题三(B)一第 3 题 (D)

$$\text{显然, } X \text{ 的分布函数为 } F_X(x) = \begin{cases} 1 - e^{\lambda x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

$$Y \text{ 的分布函数为 } F_Y(y) = \begin{cases} 0, & x < 0 \\ 1/2, & 0 \leq x < 1 \\ 1, & x \geq 1 \end{cases}$$

从而 N 的分布函数为

$$F_N(z) = 1 - [1 - F_X(z)][1 - F_Y(z)] = \begin{cases} 0, & z < 0 \\ \frac{1}{2}, & z = 0 \\ 1 - \frac{e^{-\lambda z}}{2}, & 0 < z < 1 \\ 1, & z \geq 1 \end{cases}$$

二. 1. 教材习题三(B)二第 3 题 $\frac{1}{2}$

(X, Y)	$(0, 0)$	$(0, 1)$	$(1, 0)$	$(1, 1)$
p_{ij}	$1/4$	$1/4$	$1/4$	$1/4$
M	0	1	1	1
N	0	0	0	1

2. 类似教材习题三(B)三第5题 $f_Z(z) = \frac{1}{2}[f_Y(z) + f_Y(z-1)]$

三. 1. (1) 由题意有 $Z|Y=n \sim B(n, 0.2)$, 所以 $P(Z=k|Y=n) = C_n^k 0.2^k 0.8^{n-k}$,

$$k = 0, 1, 2, \dots, n;$$

$$\begin{aligned} (2) \text{ 当 } k \leq n \text{ 时, } P(Z=k, Y=n) &= P(Y=n)P(Z=k|Y=n) \\ &= \frac{30^n}{n!} e^{-30} C_n^k 0.2^k 0.8^{n-k}, \quad k = 0, 1, 2, \dots, n \end{aligned}$$

$$\text{当 } k > n \text{ 时, } P(Z=k, Y=n) = 0$$

2. 教材习题三(B)三第2题

先计算 X 与 Y 的(边缘)密度函数为

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_{-1}^{+1} \frac{1}{4}(1+xy) dy = \frac{1}{2}, & |x| < 1 \\ 0, & |x| \geq 1 \end{cases} \\ f_Y(y) &= \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{-1}^{+1} \frac{1}{4}(1+xy) dx = \frac{1}{2}, & |y| < 1 \\ 0, & |y| \geq 1 \end{cases} \end{aligned}$$

易见, 在三个密度函数的公共连续点 $\left(\frac{1}{2}, \frac{1}{2}\right)$ 处, $f(x, y) = \frac{5}{16} \neq \frac{1}{4} = f_X(x)f_Y(y)$,

所以 X 与 Y 不独立!

令 $U = X^2, V = Y^2$, 显然 $R(U) = R(V) = [0, 1]$, 当 $0 < u < 1, 0 < v < 1$ 时,

$$\begin{aligned} F(u, v) &= P(U \leq u, V \leq v) = P(X^2 \leq u, Y^2 \leq v) \\ &= P(-\sqrt{u} \leq X \leq \sqrt{u}, -\sqrt{v} \leq Y \leq \sqrt{v}) \\ &= \int_{-\sqrt{u}}^{\sqrt{u}} \left(\int_{-\sqrt{v}}^{\sqrt{v}} \frac{1}{4}(1+xy) dy \right) dx = \sqrt{uv} \end{aligned}$$

从而 (U, V) 的联合密度函数为

$$\psi(u, v) = \frac{\partial^2 F(u, v)}{\partial u \partial v} = \begin{cases} \frac{1}{4\sqrt{uv}}, & 0 < u < 1, 0 < v < 1 \\ 0, & \text{others} \end{cases}$$

而 U 与 V 的(边缘)密度函数为

$$\psi_U(u) = \int_{-\infty}^{+\infty} \psi(u, v) dv = \begin{cases} \int_0^1 \frac{dv}{4\sqrt{uv}} = \frac{1}{2\sqrt{u}}, & 0 < u < 1, \\ 0, & \text{others} \end{cases}$$

$$\psi_V(v) = \int_{-\infty}^{+\infty} \psi(u, v) du = \begin{cases} \int_0^1 \frac{du}{4\sqrt{uv}} = \frac{1}{2\sqrt{v}}, & 0 < v < 1, \\ 0, & \text{others} \end{cases}$$

显然可见, 对任意的 u, v , 都有 $\psi(u, v) = \psi_U(u)\psi_V(v)$, 所以 U 与 V 相互独立,

即 X^2 与 Y^2 相互独立!

3. 教材习题三(B)三第6题

$$(X, Y) \text{ 的联合密度为 } f(x, y) = \begin{cases} 1/4, & (x, y) \in [1, 3] \times [1, 3] \\ 0, & \text{others} \end{cases}$$

显然 $U = |X - Y|$ 的值域为 $R(U) = [0, 2]$, 对任意的 $u \in (0, 2)$, 有

$$F_U(u) = P(U \leq u) = P(|X - Y| \leq u) = \iint_{|x-y| \leq u} f(x, y) dx dy$$

$$= \frac{4 - (2 - u)^2}{4} = u - \frac{1}{4}u^2,$$

$$\text{从而密度函数为 } p_U(u) = \begin{cases} 1 - \frac{1}{2}u, & u \in (0, 2) \\ 0, & u \notin (0, 2) \end{cases}$$

第 31 - 32 页

一. 1. C 2. B 3. B

二. 1. $a = 2, b = 8$

2. 780 设抽得 3 张奖券的总金额为 ξ , 则 $\xi \sim \left[\begin{array}{ccc} 600 & 900 & 1200 \\ C_8^3/C_{10}^3 & C_8^2 C_2^1/C_{10}^3 & C_8^1 C_2^2/C_{10}^3 \end{array} \right]$, 从而

$$\text{有 } E(\xi) = 600 \times \frac{C_8^3}{C_{10}^3} + 900 \times \frac{C_8^2 C_2^1}{C_{10}^3} + 1200 \times \frac{C_8^1 C_2^2}{C_{10}^3} = 780$$

3. $-\frac{67}{45}, 9\frac{8}{15}$

三. 1. 教材习题四(A)三第 2 题

显然, X 可取 1, 2, 3, 4, 5, 7, 8, 9, 10, 11, 12, 其分布律为

X	1	2	3	4	5	7	8	9	10	11	12
p_i	1/6	1/6	1/6	1/6	1/6	1/36	1/36	1/36	1/36	1/36	1/36

$$\text{平均得分为 } E(X) = 1/6 \sum_{i=1}^5 i + 1/36 \sum_{i=7}^{12} i = \frac{15}{6} + \frac{57}{36} = \frac{147}{36} = \frac{49}{12}$$

2. 教材习题四(A)三第 4 题

由分布函数的右连续性得 $A - \frac{B}{25} = 0(1)$, 再由 $F(+\infty) = 1$ 得 $A = 1(2)$, 联立(1)(2)

两式即得 $A = 1, B = 25$, 从而密度函数为 $f(x) = F'(x) = \begin{cases} 0, & x \leq 5 \\ 50x^{-3}, & x > 5 \end{cases}$, 于是动

$$\text{物的平均寿命为 } E(X) = \int_{-\infty}^{+\infty} xf(x)dx = \int_5^{+\infty} x50x^{-3}dx = 50 \int_5^{+\infty} x^{-2}dx = 10$$

3. 教材习题四(A)三第 5 题

设工厂售出一台设备获利为 Y , 则 $Y = g(X) = \begin{cases} 600, & X < 1 \\ 1000, & X \geq 1 \end{cases}$, 从而有

$$\begin{aligned} E(Y) &= \int_{-\infty}^{+\infty} f(x)g(x)dx = \int_0^1 0.25e^{-0.25x} \times 600dx + \int_1^{+\infty} 0.25e^{-0.25x} \times 1000dx \\ &= 600(1 - e^{-0.25}) + 1000e^{-0.25} = 600 + 400e^{-0.25} \end{aligned}$$

4. 教材习题四(A)三第 7 题

显然 $E(X) = 1, D(X) = 1, E(Y) = 2, D(Y) = 2$, 从而有

$$(1) \mathbf{E}(X - Y) = \mathbf{E}(X) - \mathbf{E}(Y) = 1 - 2 = -1;$$

$$(2) \mathbf{E}(2X^2 + 3Y^2) = 2\mathbf{E}(X^2) + 3\mathbf{E}(Y^2) \\ = 2\left[\mathbf{D}(X) + (\mathbf{E}(X))^2\right] + 3\left[\mathbf{D}(Y) + (\mathbf{E}(Y))^2\right] = 22;$$

$$(3) \text{ 若 } X \text{ 与 } Y \text{ 相互独立, 则 } \mathbf{E}(XY) = \mathbf{E}(X) \mathbf{E}(Y) = 2.$$

5. Y 表每次检验的次品数, 则 $Y \sim B(10, 0.1)$

$$p = P(Y > 1) = 1 - 0.9^{10} - C_{10}^1 0.1^1 0.9^{10} = 0.2639$$

$$Z \sim B(4, 0.2639), \quad \mathbf{E}(X) = 4 \times 0.2639 = 1.06$$

第 33-34 页

一. 1. A 2. C 3. B

二. 1. 0.9 2. $2a^2$ 3. $k! \frac{1}{2^k}$

三. 1. 教材习题四(A)三第 10 题

$$\mathbf{E}(X) = \int_{-\infty}^{+\infty} xf(x)dx = \int_{-1}^1 x \frac{1}{\pi\sqrt{1-x^2}} dx = 0;$$

$$\mathbf{E}(X^2) = \int_{-\infty}^{+\infty} x^2 f(x)dx = \int_{-1}^1 x^2 \frac{1}{\pi\sqrt{1-x^2}} dx = \frac{1}{2};$$

$$\left(\int \frac{x^2 dx}{\sqrt{a^2 - x^2}} = -\frac{x}{2} \sqrt{a^2 - x^2} + \frac{a^2}{2} \arcsin \frac{x}{a} + C \right)$$

$$\mathbf{D}(X) = \mathbf{E}(X^2) - (\mathbf{E}(X))^2 = \frac{1}{2}.$$

2. 教材习题四(A)三第 13 题

$$\mathbf{E}(Y) = \int_0^{1/2} 2x^2 \times 2dx = \frac{1}{6}; \quad \mathbf{E}(Y^2) = \int_0^{1/2} (2x^2)^2 \times 2dx = \frac{1}{20};$$

$$\mathbf{D}(Y) = \mathbf{E}(Y^2) - (\mathbf{E}(Y))^2 = \frac{1}{20} - \frac{1}{36} = \frac{1}{45}$$

3. 教材习题四(A)三第 12 题

由于 $X \sim U(-1, 3)$, $Y \sim e(2)$, $Z \sim \Gamma(2, 2)$, 所以

$$E(X) = 1, E(Y) = \frac{1}{2}, E(Z) = 1, D(X) = \frac{4}{3}, D(Y) = \frac{1}{4}, D(Z) = \frac{1}{2},$$

且 X, Y, Z 相互独立, 从而有

$$\begin{aligned} (1) E(U) &= E(3X - 2XY + 4YZ - 2) \\ &= 3E(X) - 2E(X)E(Y) + 4E(Y)E(Z) - 2 \\ &= 3 \times 1 - 2 \times 1 \times \frac{1}{2} + 4 \times \frac{1}{2} \times 1 - 2 = 2 \end{aligned}$$

$$\begin{aligned} (2) D(V) &= D(X - 2Y + 3Z - 2) \\ &= D(X) + 4D(Y) + 9D(Z) = \frac{4}{3} + 4 \times \frac{1}{4} + 9 \times \frac{1}{2} = \frac{41}{6} \end{aligned}$$

$$4. E(X - c)^2 = D(X - c) + [E(X - c)]^2 = D(X) + [E(X - c)]^2 \geq D(X)$$

5. 1.5

设 $X_i = \begin{cases} 1, & \text{第 } i \text{ 个部件需要调整} \\ 0 & \text{第 } i \text{ 个部件不需调整} \end{cases}$, $i = 1, 2, 3, 4, 5$, 由题意知 X_1, X_2, X_3, X_4, X_5 相

互独立且 $X = \sum_{i=1}^5 X_i$, $X_i \sim \begin{bmatrix} 1 & 0 \\ i/10 & 1 - i/10 \end{bmatrix}$, $i = 1, 2, 3, 4, 5$, 从而有

$$E(X_i) = \frac{i}{10}, E(X_i^2) = \frac{i}{10}, D(X_i) = \frac{i(10-i)}{100}, i = 1, 2, 3, 4, 5,$$

所以有

$$\begin{aligned} E(X) &= E\left(\sum_{i=1}^5 X_i\right) = \sum_{i=1}^5 E(X_i) = \frac{1+2+3+4+5}{10} = 1.5, \\ D(X) &= D\left(\sum_{i=1}^5 X_i\right) = \sum_{i=1}^5 D(X_i) = \sum_{i=1}^5 \frac{i(10-i)}{100} = \frac{9+16+21+24+25}{100} = 0.95 \end{aligned}$$

第 35-38 页 《协方差与相关系数》

一. 1. B 2. A 3. C

二. 1. $\frac{4}{3} + \frac{1}{2\sqrt{3}}$ 2. $\frac{1}{4}$ 3. $\frac{2}{3}\sigma^2$

三. 1. 教材习题四(A) 三第 15 题

$$\mathbf{E}(XY) = \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} xyf(x, y) dy = \int_0^1 x dx \int_0^1 y(2-x-y) dy = \frac{1}{6}$$

边缘密度函数为

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^1 (2-x-y) dy = \frac{3}{2} - x, & 0 \leq x \leq 1 \\ 0, & x \notin [0, 1] \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_0^1 (2-x-y) dx = \frac{3}{2} - y, & 0 \leq y \leq 1 \\ 0, & y \notin [0, 1] \end{cases}$$

从而有

$$\mathbf{E}(X) = \int_{-\infty}^{+\infty} xf_X(x) dx = \int_0^1 x(3/2 - x) dx = \frac{5}{12}$$

$$\mathbf{E}(X^2) = \int_{-\infty}^{+\infty} x^2 f_X(x) dx = \int_0^1 x^2 (3/2 - x) dx = \frac{1}{4}$$

$$\mathbf{D}(X) = \mathbf{E}(X^2) - (\mathbf{E}(X))^2 = \frac{1}{4} - \frac{5^2}{12^2} = \frac{36}{144}$$

$$\mathbf{E}(Y) = \int_{-\infty}^{+\infty} yf_Y(y) dy = \int_0^1 y(3/2 - y) dy = \frac{5}{12}$$

$$\mathbf{E}(Y^2) = \int_{-\infty}^{+\infty} y^2 f_Y(y) dy = \int_0^1 y^2 (3/2 - y) dy = \frac{1}{4}$$

$$\mathbf{D}(Y) = \mathbf{E}(Y^2) - (\mathbf{E}(Y))^2 = \frac{1}{4} - \frac{5^2}{12^2} = \frac{36}{144}$$

$$\text{Cov}(X, Y) = \mathbf{E}(XY) - \mathbf{E}(X)\mathbf{E}(Y) = \frac{1}{6} - \frac{5^2}{12^2} = -\frac{1}{144}$$

$$\text{协方差阵为 } V = \begin{bmatrix} D(X) & \text{Cov}(X, Y) \\ \text{Cov}(X, Y) & D(Y) \end{bmatrix} = \begin{bmatrix} \frac{36}{144} & -\frac{1}{144} \\ -\frac{1}{144} & \frac{36}{144} \end{bmatrix}$$

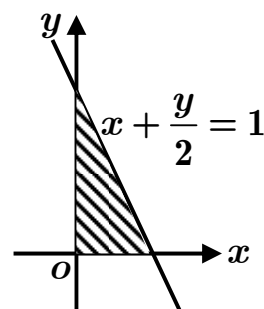
$$\text{相关系数为 } R(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-\frac{1}{144}}{\sqrt{\frac{36}{144}}\sqrt{\frac{36}{144}}} = -\frac{1}{36}$$

2. 教材习题四(A)三第 16 题

如图, 联合密度为

$$f(x, y) = \begin{cases} 1, & 0 < x < 1, 0 < y < 2 - 2x \\ 0, & \text{others} \end{cases}$$

$$\begin{aligned} E(XY) &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} xyf(x, y) dy \\ &= \int_0^1 x dx \int_0^{2-2x} y dy = \frac{7}{6} \end{aligned}$$



边缘密度函数为

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^{2-2x} 1 dy = 2 - 2x, & 0 \leq x \leq 1 \\ 0, & x \notin [0, 1] \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_0^{1-y/2} 1 dx = 1 - \frac{y}{2}, & 0 \leq y \leq 2 \\ 0, & y \notin [0, 2] \end{cases}$$

从而有

$$(1) E(X) = \int_{-\infty}^{+\infty} xf_X(x) dx = \int_0^1 x(2 - 2x) dx = \frac{1}{3}$$

$$E(Y) = \int_{-\infty}^{+\infty} yf_Y(y) dy = \int_0^2 y(1 - y/2) dy = \frac{2}{3}$$

$$(2) E(X^2) = \int_{-\infty}^{+\infty} x^2 f_X(x) dx = \int_0^1 x^2(2 - 2x) dx = \frac{1}{6}$$

$$D(X) = E(X^2) - (E(X))^2 = \frac{1}{6} - \frac{1}{9} = \frac{1}{18}$$

$$E(Y^2) = \int_{-\infty}^{+\infty} y^2 f_Y(y) dy = \int_0^2 y^2 (1 - y/2) dy = \frac{2}{3}$$

$$D(Y) = E(Y^2) - (E(Y))^2 = \frac{2}{3} - \frac{4}{9} = \frac{2}{9}$$

$$(3) \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{1}{6} - \frac{1}{3} \times \frac{2}{3} = -\frac{1}{18}$$

$$R(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-\frac{1}{18}}{\sqrt{\frac{1}{18}}\sqrt{\frac{2}{9}}} = -\frac{1}{2}$$

(4) 因 $R(X, Y) = -1/2 \neq 0$, 即 X 与 Y 相关, 当然不独立.

3. 教材习题四(A)三第 17 题

$$\begin{aligned} \text{Cov}(X+Y, X-Y) &= \text{Cov}(X, X) + \text{Cov}(Y, X) - \text{Cov}(X, Y) - \text{Cov}(Y, Y) \\ &= D(X) - D(Y) \end{aligned}$$

4. 教材习题四(A)三第 18 题

直接由上题的结论有

$$\begin{aligned} \text{Cov}(U, V) &= \text{Cov}(\alpha X + \beta Y, \alpha X - \beta Y) = D(\alpha X) - D(\beta Y) \\ &= \alpha^2 D(X) - \beta^2 D(Y) = \alpha^2 \sigma^2 - \beta^2 \sigma^2 = (\alpha^2 - \beta^2) \sigma^2 \end{aligned}$$

因 X 与 Y 相互独立, 故

$$D(U) = D(\alpha X + \beta Y) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2) \sigma^2$$

$$D(V) = D(\alpha X - \beta Y) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2) \sigma^2$$

从而有

$$R(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{(\alpha^2 - \beta^2) \sigma^2}{\sqrt{(\alpha^2 + \beta^2) \sigma^2} \sqrt{(\alpha^2 + \beta^2) \sigma^2}} = \frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2}$$

5. 教材习题四(A)三第 14 题

由联合分布律易得

$$X \sim \begin{bmatrix} 1 & 2 & 3 \\ 4/9 & 1/3 & 2/9 \end{bmatrix}, Y \sim \begin{bmatrix} 0 & 1 & 2 \\ 5/9 & 1/3 & 1/9 \end{bmatrix}, XY \sim \begin{bmatrix} 0 & 1 & 2 \\ 5/9 & 2/9 & 2/9 \end{bmatrix}$$

所以

$$(1) E(X) = 1 \times 4/9 + 2 \times 3/9 + 3 \times 2/9 = 16/9$$

$$E(X^2) = 1^2 \times 4/9 + 2^2 \times 3/9 + 3^2 \times 2/9 = 34/9$$

$$D(X) = E(X^2) - [E(X)]^2 = 34/9 - 16^2/9^2 = 50/81$$

$$E(Y) = 0 \times 5/9 + 1 \times 3/9 + 2 \times 1/9 = 5/9$$

$$E(Y^2) = 0^2 \times 5/9 + 1^2 \times 3/9 + 2^2 \times 1/9 = 7/9$$

$$D(Y) = E(Y^2) - [E(Y)]^2 = 7/9 - 5^2/9^2 = 38/81$$

$$(2) E(XY) = 0 \times 5/9 + 1 \times 2/9 + 2 \times 2/9 = 2/3$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{2}{3} - \frac{16}{9} \times \frac{5}{9} = -\frac{26}{81}$$

$$R(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-26/81}{\sqrt{50/81}\sqrt{38/81}} = -\frac{13}{5\sqrt{19}}$$

$$(3) D(X - 3Y) = D(X) + 9D(Y) - 6\text{Cov}(X, Y)$$

$$= \frac{50}{81} + 9 \times \frac{38}{81} - 6 \times \left(-\frac{26}{81}\right) = \frac{548}{81}$$

第 39-40 页 《第四章综合练习》

一. 1. $\frac{8}{9}$ 2. $\sqrt{\frac{m}{m+n}}$

二. 1. 教材习题四(B)三第 1 题

设应组织 y 吨货源, 才能使平均收益最大, 则必有 $y \in (2000, 4000)$.

设收益为 Y , 由题意有

$$Y = g(X) = \begin{cases} 3y, & X > y \\ 3X - (y - X) = 4X - y, & X \leq y \end{cases}$$

于是平均收益为

$$\begin{aligned}\mathbf{E}(Y) &= \mathbf{E}(g(X)) = \int_{-\infty}^{+\infty} g(x) f_X(x) dx = \int_{2000}^y \frac{4x-y}{2000} dx + \int_y^{4000} \frac{3y}{2000} dx \\ &= \frac{1}{1000} [-y^2 + 7000y - 2000^2]\end{aligned}$$

显然可见, 当组织的货源为 $y = 3500$ 吨时, 平均收益 $\mathbf{E}(Y)$ 最大.

2. 教材习题四(B)三第2题

设 X 为停车次数, 并设

$$X_i = \begin{cases} 1, & \text{第 } i \text{ 站有人下车} \\ 0, & \text{第 } i \text{ 站无人下车} \end{cases}, \quad i = 1, 2, \dots, 10, \quad \text{则 } X = \sum_{i=1}^{10} X_i$$

由题意有

$$X_i \sim \begin{bmatrix} 0 & 1 \\ 0.9^{20} & 1 - 0.9^{20} \end{bmatrix}, \quad i = 1, 2, \dots, 10, \quad \text{则 } \mathbf{E}(X_i) = 1 - 0.9^{20}, \quad i = 1, 2, \dots, 10$$

从而有平均停车次数为 $\mathbf{E}(X) = \mathbf{E}\left(\sum_{i=1}^{10} X_i\right) = \sum_{i=1}^{10} \mathbf{E}(X_i) = 10(1 - 0.9^{20})$.

3. 教材习题四(B)三第6题

$$(1) \quad \mathbf{E}(X) = \int_{-\infty}^{+\infty} xf(x)dx = \int_{-1}^1 x|x|dx = 0,$$

$$\mathbf{E}(X^2) = \int_{-\infty}^{+\infty} x^2 f(x)dx = \int_{-1}^1 x^2|x|dx = 2 \int_0^1 x^2 x dx = \frac{1}{2},$$

$$\mathbf{D}(X) = \mathbf{E}(X^2) - [\mathbf{E}(X)]^2 = \frac{1}{2}$$

$$(2) \quad \text{因 } \mathbf{E}(X|X) = \int_{-\infty}^{+\infty} x|x|f(x)dx = \int_{-1}^1 x|x|^2 dx = 0, \quad \mathbf{E}(X) = 0, \quad \text{所以}$$

$$\mathbf{Cov}(X, |X|) = \mathbf{E}(X|X) - \mathbf{E}(X)\mathbf{E}(|X|) = 0;$$

(3) $\mathbf{Cov}(X, |X|) = 0$ 表明 $R(X, |X|) = 0$, 从而 X 与 $|X|$ 不相关!

但 X 与 $|X|$ 不是相互独立的. 这是因为: 对任意 $x \in (0, 1/2)$, 有 $\mathbf{P}(X < x) < 1$,

从而

$$P(X < x, |X| \leq x) = P(|X| \leq x) \neq P(X < x)P(|X| \leq x)$$

所以 X 与 $|X|$ 不相互独立!

4. 教材习题四(B)三第8题

证明 由题意知 (X, Y) 的联合分布律为

$$P(X=1, Y=1) = P(AB), \quad P(X=-1, Y=-1) = P(\bar{A}\bar{B})$$

$$P(X=-1, Y=1) = P(\bar{A}B), \quad P(X=1, Y=-1) = P(A\bar{B})$$

从而有

$$XY \sim \begin{bmatrix} 1 & -1 \\ P(AB) + P(\bar{A}\bar{B}) & P(\bar{A}B) + P(A\bar{B}) \end{bmatrix},$$

故

$$\begin{aligned} E(XY) &= 1 \times [P(AB) + P(\bar{A}\bar{B})] + (-1) \times [P(\bar{A}B) + P(A\bar{B})] \\ &= P(AB) + P(\bar{A}\bar{B}) - P(\bar{A}B) - P(A\bar{B}) \\ &= P(AB) + P(\overline{A \cup B}) - [P(B) - P(AB)] - [P(A) - P(AB)] \\ &= 3P(AB) + [1 - P(A \cup B)] - P(B) - P(A) \\ &= 4P(AB) + 1 - 2P(B) - 2P(A) \end{aligned}$$

边缘分布律为

$$X \sim \begin{bmatrix} 1 & -1 \\ P(A) & P(\bar{A}) \end{bmatrix}, \quad Y \sim \begin{bmatrix} 1 & -1 \\ P(B) & P(\bar{B}) \end{bmatrix},$$

故

$$E(X) = 1 \times P(A) + (-1) \times P(\bar{A}) = P(A) - P(\bar{A}) = 2P(A) - 1$$

$$E(Y) = 1 \times P(B) + (-1) \times P(\bar{B}) = P(B) - P(\bar{B}) = 2P(B) - 1$$

$$\text{于是 } X \text{ 与 } Y \text{ 不相关} \Leftrightarrow E(XY) = E(Y)E(X)$$

$$\Leftrightarrow 4P(AB) + 1 - 2P(B) - 2P(A) = [2P(A) - 1][2P(B) - 1]$$

$$\Leftrightarrow P(AB) = P(A)P(B) \Leftrightarrow A \text{ 与 } B \text{ 相互独立}$$

一. 1. A

$$p_1 = P(X \leq \mu - 2) = \Phi\left(\frac{\mu - 2 - \mu}{2}\right) = \Phi(-1) = 1 - \Phi(1)$$

$$p_2 = P(Y \geq \mu + 3) = 1 - P(Y \leq \mu + 3) = 1 - \Phi\left(\frac{\mu + 3 - \mu}{3}\right) = 1 - \Phi(1)$$

2. C

$$\begin{aligned} P(|X - \mu| < q) > P(|Y - \mu| < q) &\Rightarrow P\left(\left|\frac{X - \mu}{\sigma_1}\right| < \frac{q}{\sigma_1}\right) > P\left(\left|\frac{Y - \mu}{\sigma_2}\right| < \frac{q}{\sigma_2}\right) \\ &\Rightarrow 2\Phi\left(\frac{q}{\sigma_1}\right) - 1 > 2\Phi\left(\frac{q}{\sigma_2}\right) - 1 \Rightarrow \Phi\left(\frac{q}{\sigma_1}\right) > \Phi\left(\frac{q}{\sigma_2}\right) \Rightarrow \sigma_1 < \sigma_2 \end{aligned}$$

3. C

$$\text{由 } X \sim f(x) = \frac{1}{\sqrt{3\pi}} e^{\frac{1}{3}(2x-x^2-1)} = \frac{1}{\sqrt{2\pi}\sqrt{3/2}} e^{-\frac{(x-1)^2}{2(\sqrt{3/2})^2}}, \text{ 即 } X \sim N(1, 3/2), \text{ 从而}$$

$$E(X) = 1, D(X) = 3/2$$

二. 1. 0.2

因 $X \sim N(\mu, \sigma^2)$ 且 $P(X > 2) = 0.5$, 则 $\mu = 2$, $P(X \leq 2) = 0.5$, 从而

$$P(0 < X < 2) = P(2 < X < 4) = 0.3, \text{ 所以}$$

$$P(X < 0) = P(X \leq 2) - P(0 < X < 2) = 0.5 - 0.3 = 0.2$$

2. 0.12

因 $X \sim N(2, 4)$ 且 $P(|X - 2| > a) \geq 0.95$, 则 $P\left(\left|\frac{X - 2}{2}\right| \leq \frac{a}{2}\right) \leq 0.05$, 即

$$2\Phi(a/2) - 1 \leq 0.05, \text{ 所以 } \Phi(a/2) \leq 0.525, \text{ 查表得 } a/2 \leq 0.06, \text{ 因而 } a \leq 0.12.$$

3. 0

因 X 的密度函数是偶函数, 所以 $E(X^{2n+1}) = 0 = E(X)$, 从而有 $\text{Cov}(X, Y) =$

$$E(XY) - E(X)E(Y) = E(X^{2n+1}) - E(X)E(X^{2n}) = 0, \text{ 故 } R(X, Y) = 0.$$

三. 1. (1) 0.9525 (2) 0.3707 (3) 0.3115 (4) 0.7714

2. 教材习题五(A)三第2题

每个新生婴儿体重小于 **2719g** 的概率为

$$P(X < 2719) = \Phi\left(\frac{2719 - 3315}{575}\right) = \Phi(-1.0365) = 1 - \Phi(1.0365) = 0.15$$

设所选的 **100** 个新生婴儿中体重小于 **2719g** 的个数为 Y ，则 $Y \sim B(100, 0.15)$ ，从而

$$\begin{aligned} \text{所求概率为 } P(Y \geq 2) &= 1 - P(Y = 0) - P(Y = 1) = 1 - C_{100}^0 0.15^0 0.85^{100} - \\ &- C_{100}^1 0.15^1 0.85^{99} = 1 - 0.85^{99} \times 15.85 = 1 \end{aligned}$$

3. 教材习题五(A)三第3题

显然 $R(Y) = [0, +\infty)$. 对任意的 $y > 0$, $F_Y(y) = P(Y \leq y) = P(|X| \leq y) =$

$$2\Phi(y) - 1, \text{ 此时 } f_Y(y) = F'_Y(y) = 2\varphi(y) = \frac{2}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}, \text{ 于是綜上有}$$

$$f_Y(y) = \begin{cases} \frac{2}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}, & y > 0 \\ 0, & y \leq 0 \end{cases}$$

4. 教材习题五(A)三第5题

$$(1) \text{ 联合密度函数为 } f(x, y) = \begin{cases} \frac{1}{2} e^{-\frac{1}{2}y}, & x \in [0, 1], y > 0 \\ 0, & \text{others} \end{cases}$$

(2) $a^2 + 2aX + Y = 0$ 有实根 $\Leftrightarrow X^2 \geq Y$, 故所求概率为

$$\begin{aligned} P(X^2 \geq Y) &= \int_0^1 dx \int_0^{x^2} \frac{1}{2} e^{-\frac{1}{2}y} dy = 1 - \int_0^1 e^{-\frac{1}{2}x^2} dx \\ &= 1 - \sqrt{2\pi} \int_0^1 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} dx = 1 - \sqrt{2\pi} (\Phi(1) - \Phi(0)) \\ &= 1 - 0.3413\sqrt{2\pi} = 0.1445 \end{aligned}$$

第 43-44 页 二维正态分布及自然指数分布族

一. 1. B 由二维正态密度定义, 该密度函数可改写为

$$f(x, y) = \frac{1}{\sqrt{3\pi}} \exp\left[-\frac{2}{3}(x^2 + xy + y^2)\right]$$

$$= \frac{1}{2\pi\sqrt{1-(-1/2)^2}} \exp \left[-\frac{1}{2[1-(-1/2)^2]} (x^2 - 2(-1/2)xy + y^2) \right]$$

2. C 由上题显然可见 $(X, Y) \sim N(0, 0, 1, 1, -1/2)$, 故 $X, Y \sim N(0, 1)$ 但不独立.

3. B (U, V) 服从二维正态分布 (因 U, V 的任一维线性组合服从一维正态), 故: U, V 独

$$\text{立} \Leftrightarrow \text{Cov}(U, V) = D(X) - D(Y) = 0 \Leftrightarrow \sigma_1^2 = \sigma_2^2$$

二. 1. $\frac{1}{3}, \quad 7, \quad N(1/3, 7)$

2. $V(m) = 10p(1-p) = m - \frac{m^2}{10}$ (因为 $V(m) = D(X), m = E(X)$)

3. $\frac{1}{2a}\sqrt{2\pi}$

三. 1. 教材习题五(A)三第4题

显然, $R(Z) = [0, +\infty)$; 对任意的 $z > 0$, $F_Z(z) = P(Z \leq z) = P(\frac{1}{2}m(X^2 + Y^2) \leq z)$

$$= P(X^2 + Y^2 \leq \frac{2z}{m}) = \iint_{x^2+y^2 \leq \frac{2z}{m}} \frac{1}{2\pi\sigma^2} e^{-\frac{x^2+y^2}{2\sigma^2}} dx dy$$

$$\xrightarrow[\substack{x=r\cos\theta \\ y=r\sin\theta}]{\substack{x=r\cos\theta \\ y=r\sin\theta}} \int_0^{2\pi} d\theta \int_0^{\sqrt{\frac{2z}{m}}} \frac{1}{2\pi\sigma^2} e^{-\frac{r^2}{2\sigma^2}} r dr = 1 - e^{-\frac{z}{m\sigma^2}}; \text{ 所以}$$

$$F_Z(z) = \begin{cases} 1 - e^{-\frac{z}{m\sigma^2}}, & z > 0, \\ 0, & z \leq 0 \end{cases}, \quad f_Z(z) = \begin{cases} \frac{1}{m\sigma^2} e^{-\frac{z}{m\sigma^2}}, & z > 0 \\ 0, & z \leq 0 \end{cases}$$

2. 教材习题五(A)三第6题

$$C_v = 1 \Rightarrow \frac{\sqrt{D(X_i)}}{E(X_i)} = 1 \Rightarrow \frac{\sigma_i}{i} = 1 \Rightarrow \sigma_i = i, \quad i = 1, 2, 3, 4$$

于是, $X = \sum_{i=1}^4 X_i \sim N\left(\sum_{i=1}^4 \mu_i, \sum_{i=1}^4 \sigma_i^2\right) = N(10, 30)$, 从而有

$$P(2 < Z < 18) = \Phi\left(\frac{18-10}{\sqrt{30}}\right) - \Phi\left(\frac{2-10}{\sqrt{30}}\right) = 2\Phi\left(\frac{2-10}{\sqrt{30}}\right) - 1 = \dots$$

3. 教材习题五(A)三第7题(略)

第 45-46 页 极限定理

一. 1. C 2. C 3. A

二. 1. $\frac{1}{n\lambda}$ 2. $\frac{1}{12}$

三. 1. 教材习题六(A)三第1题

$$E(X) = \int_0^{+\infty} x \frac{x^m e^{-x}}{m!} dx = \frac{\Gamma(m+2)}{m!} = \frac{(m+1)!}{m!} = m+1$$

$$E(X^2) = \int_0^{+\infty} x^2 \frac{x^m e^{-x}}{m!} dx = \frac{\Gamma(m+3)}{m!} = \frac{(m+2)!}{m!} = (m+1)(m+2)$$

$$D(X) = E(X^2) - [E(X)]^2 = m+1$$

$$P(0 < X < 2(m+1)) = P(|X - E(X)| < (m+1)) > 1 - \frac{D(X)}{(m+1)^2}$$

$$= 1 - \frac{m+1}{(m+1)^2} = \frac{m}{m+1}$$

2. 教材习题六(A)三第2题

$$(1) \frac{\theta}{2}; (2) E(X^4) = \int_0^\theta x^4 \frac{1}{\theta} dx = \frac{1}{5} \theta^4; (3) \theta$$

3. 教材习题六(A)三第4题

设 X_i 为第 i 个灯泡的寿命, 则 $E(X_i) = 0.2, D(X_i) = 0.04, i = 1, 2, \dots, 200$.

由独立同分布中心极限定理, 近似地有 $\sum_{i=1}^{200} X_i \sim N(40, 8)$, 则灯泡的平均寿命

$\bar{X} \sim N(0.2, 0.0002)$, 于是, 所求概率为

$$P(\bar{X} > 0.21) = 1 - \Phi\left(\frac{0.21 - 0.2}{\sqrt{0.0002}}\right) = 1 - \Phi(0.71) = 0.2389$$

4. 教材习题六(A)三第6题

设 X 为正常工作的元件个数, 由题意知, $X \sim B(100, 0.9)$

(1) 由二项分布以正态分布为极限的中心极限定理, 近似地有 $X \sim N(90, 9)$,

从而系统正常工作的概率为 $P(X \geq 85) = 1 - \Phi\left(\frac{85 - 90}{3}\right) = \Phi\left(\frac{5}{3}\right) = 0.952$

(2) 此时, $X \sim B(n, 0.9)$, 由二项分布以正态分布为极限的中心极限定理, 近似地

有 $X \sim N(0.9n, 0.09n)$, 由题意应有 $P(X \geq 80\%n) \geq 0.95$, 即

$$\Phi\left(\frac{0.9n - 0.8n}{0.3\sqrt{n}}\right) \geq 0.95, \text{ 所以 } \frac{0.9n - 0.8n}{0.3\sqrt{n}} \geq 1.645, \quad n \geq 25$$

第 47-48 页 χ^2, t, F 分布

一. 1. D 2. C 3. D 4. C

二. 1. $\frac{1}{3}, 2$ 2. 0.7 3. $t(9)$ 4. 0.4234

三. 1. 教材习题七(A)三第1题

$$\bar{x} = 50.56, s^2 = 4.28, b_2 = 3.80$$

2. 教材习题七(A)三第3题

$$\frac{Y_n}{6.4} = \frac{\sum_{i=1}^n X_i^2}{6.4} \sim \chi^2(n), \quad P(Y_n > 200) \leq 0.01 \Rightarrow P\left(\frac{Y_n}{6.4} \leq \frac{200}{6.4}\right) \geq 0.9$$

$$\Rightarrow \frac{200}{6.4} \geq \chi_{0.9}^2(n) \Rightarrow 31.25 \geq \chi_{0.9}^2(n) \Rightarrow n \leq 22$$

3. 教材习题七(A)三第5题

$$\bar{X} \sim N\left(52, \frac{6.3^2}{36}\right)$$

$$\begin{aligned}
 (1) P(50.8 < \bar{X} < 54.8) &= \Phi\left(\frac{54.8 - 52}{6.3/6}\right) - \Phi\left(\frac{50.8 - 52}{6.3/6}\right) \\
 &= \Phi(2.67) - 1 + \Phi(1.14) = 0.8691
 \end{aligned}$$

$$\begin{aligned}
 (2) E|\bar{X} - 52|^2 \leq 0.05 &\Rightarrow D(\bar{X}) + (E(\bar{X} - 52))^2 \leq 0.05 \Rightarrow D(\bar{X}) \leq 0.05 \\
 &\Rightarrow \frac{6.3^2}{n} \leq 0.5 \Rightarrow n \geq 79.38 \Rightarrow n \geq 80
 \end{aligned}$$

4. 教材习题七(A)三第6题

$$(1) \bar{X} - \bar{Y} \sim N\left(30 - 40, \frac{3^2}{9} + \frac{2^2}{12}\right) = N\left(-10, \frac{4}{3}\right)$$

$$\begin{aligned}
 P(|\bar{X} - \bar{Y}| < 12) &= \Phi\left(\frac{12 - (-10)}{\sqrt{4/3}}\right) - \Phi\left(\frac{-12 - (-10)}{\sqrt{4/3}}\right) \\
 &= \Phi(11\sqrt{3}) - 1 + \Phi(\sqrt{3}) = \Phi(\sqrt{3}) = 0.9583
 \end{aligned}$$

$$(2) \frac{\bar{X} - 30}{S/3} \sim t(8), P\left(\frac{\bar{X} - 30}{S/3} < C_1\right) = 0.975 \Rightarrow C_1 = t_{0.975}(8) = 2.3060$$

$$\frac{S_2^2/S_1^2}{\sigma_2^2/\sigma_1^2} = \frac{S_2^2/S_1^2}{4/9} \sim F(11, 8) \Rightarrow P\left(\frac{S_2^2}{S_1^2} < C_2\right) = P\left(\frac{S_2^2/S_1^2}{4/9} < \frac{C_2}{4/9}\right) = 0.05$$

$$\Rightarrow \frac{9}{4}C_2 = F_{0.05}(11, 8) = \frac{1}{F_{0.95}(8, 11)} \Rightarrow C_2 = \frac{4}{9} \frac{1}{F_{0.95}(8, 11)} = \frac{4}{9} \cdot \frac{1}{2.95} = 0.151$$

第 49-50 页 《矩估计》

一. 1. C 2. C

二. 1. $(\bar{X} - 1/2)^3$ 2. $\bar{X}^2/3$

三. 1. 教材习题八(A)三第1题

由于有两个未知参数，所以考虑总体的一、二阶原点矩：

$$m_1 = E(X) = \mu, \quad m_2 = E(X^2) = D(X) + [E(X)]^2 = \mu^2 + \sigma^2$$

反解得

$$\mu = m_1, \sigma^2 = m_2 - m_1^2;$$

用样本的一二阶原点矩替换总体的一二阶原点矩，即得 μ, σ^2 的矩估计为：

$$\hat{\mu} = \bar{X}, \quad \hat{\sigma}^2 = A_2 - \bar{X}^2 = B_2,$$

由样本算得 $\hat{x} = 6.20$, $b_2 = 0.38$, 所以 μ, σ^2 的矩估计值为：

$$\hat{\mu} = 6.2, \quad \hat{\sigma}^2 = 0.38.$$

因函数 $\sigma = \sqrt{\sigma^2}$ ($\sigma > 0$) 连续，所以 $\hat{\sigma} = \sqrt{\hat{\sigma}^2} = \sqrt{0.38} = 0.6164$.

干燥时间小于 6.5 小时的概率为 $P(X < 6.5) = \Phi(\frac{6.5-\mu}{\sigma})$ ，它是关于 μ, σ 的连续函数，所以

$$\text{其矩估计值为 } \hat{P}(X < 6.5) = \hat{\Phi}(\frac{6.5-\hat{\mu}}{\hat{\sigma}}) = \Phi(\frac{6.5-\hat{\mu}}{\hat{\sigma}}) = \Phi(\frac{6.5-6.2}{0.6164}) = \Phi(0.49) = 0.6879.$$

2. 教材习题八(A) 三第 3 题

由于仅一个未知参数，故只需总体期望

$$(1) E(X) = \int_0^1 x(1-\theta)x^{-\theta} dx = \frac{1-\theta}{2-\theta}, \text{ 解之得 } \theta = \frac{1-2E(X)}{1-E(X)}, \text{ 用样本均值代替总体均值}$$

$$\text{即得 } \theta \text{ 的矩估计量为 } \hat{\theta} = \frac{1-2\bar{X}}{1-\bar{X}};$$

$$(2) E(X) = \int_0^{+\infty} x \frac{x}{\theta^2} e^{-\frac{x}{\theta}} dx = 2\theta, \text{ 解之得 } \theta = \frac{E(X)}{2}, \text{ 用样本均值代替总体均值即得 } \theta \text{ 的}$$

$$\text{矩估计量为 } \hat{\theta} = \frac{\bar{X}}{2}.$$

3. 教材习题八(A) 三第 4 题

$$X \sim f_X(x) = F'_X(x) = \begin{cases} \beta x^{-\beta-1}, & x > 1 \\ 0, & x \leq 1 \end{cases}$$

由于仅一个未知参数，故只需总体期望

$$E(X) = \int_1^{+\infty} \beta x^{-\beta} dx = \frac{\beta}{1-\beta}, \text{ 解之得 } \beta = \frac{E(X)}{1+E(X)}, \text{ 用样本均值代替总体均值即得 } \beta \text{ 的矩}$$

$$\text{估计量为 } \hat{\beta} = \frac{\bar{X}}{1+\bar{X}}.$$

$$4. \quad \hat{\beta} = \bar{X}/B_2, \quad \hat{\alpha} = \bar{X}^2/B_2 \quad (\text{上课已讲})$$

5. 教材习题八(A) 三第 2 题

由于有两个未知参数，所以考虑总体的一、二阶原点矩：

$$m_1 = E(X) = \int_{\theta_1}^{+\infty} x \frac{1}{\theta_2} e^{\frac{x-\theta_1}{-\theta_2}} dx = \theta_1 + \theta_2,$$

$$m_2 = E(X^2) = \int_{\theta_1}^{+\infty} x^2 \frac{1}{\theta_2} e^{\frac{x-\theta_1}{-\theta_2}} dx = 2\theta_2^2 + \theta_1^2 + 2\theta_1\theta_2 = \theta_2^2 + (\theta_1 + \theta_2)^2$$

反解得

$$\theta_1 = m_1 - \sqrt{m_2 - m_1^2}, \quad \theta_2 = \sqrt{m_2 - m_1^2};$$

用样本的一二阶原点矩替换总体的一二阶原点矩，即得 θ_1, θ_2 的矩估计量为：

$$\hat{\theta}_1 = \bar{X} - \sqrt{B_2}, \quad \hat{\theta}_2 = \sqrt{A_2 - \bar{X}^2} = \sqrt{B_2}.$$

第 51-52 页

一. 1. D 2. A

二. 1. $\Phi\left(\frac{t - \bar{x}}{\sqrt{b_2}}\right)$ 2. $\min\{X_1, X_2, \dots, X_n\}$

3. 首先求的极大似然估计：

$$L(\beta) = \prod_{i=1}^n f(x_i, \beta) = \prod_{i=1}^n \frac{1}{2} \beta^3 x_i^2 e^{-\beta x_i} = \frac{1}{2^n} \beta^{3n} \left[\prod_{i=1}^n x_i^2 \right] e^{-\beta \sum_{i=1}^n x_i}$$

$$\ln L(\beta) = -n \ln 2 + 3n \ln \beta + 2 \sum_{i=1}^n \ln x_i - \beta \sum_{i=1}^n x_i$$

$$\frac{d \ln L(\beta)}{d\beta} = 0 \Rightarrow \frac{3n}{\beta} - \sum_{i=1}^n x_i = 0$$

解之即得 β 的极大似然估计值为 $\hat{\beta} = \frac{3}{\bar{x}}$ ；所以 β 的极大似然估计量为 $\hat{\beta} = \frac{3}{\bar{X}}$ ；

因

$$m_3 = E(X^3) = \int_0^{+\infty} x^3 \frac{\beta^3}{\Gamma(3)} x^{3-1} e^{-\beta x} dx = \frac{1}{2\beta^3} \int_0^{+\infty} (\beta x)^5 e^{-\beta x} d(\beta x) = \frac{60}{\beta^3}$$

是 β 的函数，该函数具有单值的反函数，所以其极大似然估计量为

$$\hat{m}_3 = \frac{60}{\hat{\beta}^3} = \frac{60}{\left(\frac{3}{\bar{X}}\right)^3} = \frac{20}{9} \bar{X}^3;$$

因 $V(m) = D(X) = \frac{3}{\beta^2} (\beta > 0)$ 具有单值反函数, 所以 $V(m)$ 的极大似然估计量为

$$\hat{V}(m) = \frac{3}{(\bar{X}/3)^2} = \frac{\bar{X}^2}{3}$$

三. 1. 教材习题八(A)三第3题

$$(1) \text{ 似然函数为 } L(\theta) = \prod_{i=1}^n f(x_i, \theta) = \prod_{i=1}^n (1 - \sqrt{\theta}) x_i^{-\sqrt{\theta}} = (1 - \sqrt{\theta})^n \left[\prod_{i=1}^n x_i \right]^{-\sqrt{\theta}}$$

$$\Rightarrow \ln L(\theta) = n \ln(1 - \sqrt{\theta}) - \sqrt{\theta} \sum_{i=1}^n \ln x_i$$

从而有

$$\frac{d \ln L(\theta)}{d\theta} = 0 \Rightarrow -\frac{n}{2(\sqrt{\theta} - \theta)} - \frac{1}{2\sqrt{\theta}} \sum_{i=1}^n \ln x_i = 0,$$

$$\text{解之即得 } \theta \text{ 的极大似然估计值为 } \hat{\theta} = \left[1 + \frac{n}{\sum_{i=1}^n \ln x_i} \right]^2, \text{ 其极大似然估计量为}$$

$$\hat{\theta} = \left[1 + \frac{n}{\sum_{i=1}^n \ln X_i} \right]^2;$$

$$(2) \text{ 似然函数为 } L(\theta) = \prod_{i=1}^n f(x_i, \theta) = \prod_{i=1}^n \frac{x_i}{\theta^2} e^{-\frac{x_i}{\theta}} = \theta^{-2n} \left[\prod_{i=1}^n x_i \right] e^{-\frac{1}{\theta} \sum_{i=1}^n x_i}$$

$$\Rightarrow \ln L(\theta) = -2n \ln \theta + \sum_{i=1}^n \ln x_i - \frac{1}{\theta} \sum_{i=1}^n x_i$$

从而有

$$\frac{d \ln L(\theta)}{d\theta} = 0 \Rightarrow -\frac{2n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n x_i = 0,$$

$$\text{解之即得 } \theta \text{ 的极大似然估计值为 } \hat{\theta} = \frac{1}{2n} \sum_{i=1}^n x_i = \frac{1}{2} \bar{x}, \text{ 其极大似然估计量为}$$

$$\hat{\theta} = \frac{1}{2} \bar{X}.$$

2. 教材习题八(A)三第4题

$$\text{由分布函数得密度函数为 } f_X(x, \alpha, \beta) = \begin{cases} \alpha^\beta \beta x^{-\beta-1}, & x \geq \alpha \\ 0, & x < \alpha \end{cases}$$

$$(1) \text{ 当 } \alpha = 1 \text{ 时, 似然函数为 } L(\beta) = \prod_{i=1}^n f(x_i, \beta) = \prod_{i=1}^n \beta x_i^{-\beta-1} = \beta^n \left[\prod_{i=1}^n x_i \right]^{-\beta-1}$$

$$\Rightarrow \ln L(\beta) = n \ln \beta - (\beta + 1) \sum_{i=1}^n \ln x_i$$

从而有

$$\frac{d \ln L(\beta)}{d\beta} = 0 \Rightarrow \frac{n}{\beta} - \sum_{i=1}^n \ln x_i = 0,$$

$$\text{解之即得 } \beta \text{ 的极大似然估计值为 } \hat{\beta} = \frac{n}{\sum_{i=1}^n \ln x_i}, \text{ 其极大似然估计量为}$$

$$\hat{\beta} = \frac{n}{\sum_{i=1}^n \ln X_i};$$

$$(2) \text{ 当 } \beta = 2 \text{ 时, 似然函数为 } L(\alpha) = \prod_{i=1}^n f(x_i, \alpha) = \prod_{i=1}^n 2\alpha^2 x_i^{-3} = 2^n \alpha^{2n} \left[\prod_{i=1}^n x_i \right]^{-3}$$

显然可见, 似然函数 $L(\alpha)$ 关于 α 严格单增; 与此同时, 对任意 $i = 1, 2, \dots, n$,

$$\alpha \leq x_i, \text{ 故 } \alpha \leq \min \{x_1, x_2, \dots, x_n\}, \text{ 从而 } L(\alpha) \text{ 在 } \alpha = \min \{x_1, x_2, \dots, x_n\} \text{ 处}$$

取得最大值, 所以 α 的极大似然估计值为 $\hat{\alpha} = \min \{x_1, x_2, \dots, x_n\}$, 其极大似然

估计量为 $\hat{\alpha} = \min \{X_1, X_2, \dots, X_n\}$.

3. 教材习题八(A)三第5题

$$\text{似然函数为 } L(\sigma^2) = \prod_{i=1}^n f(x_i, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{x_i^2}{2\sigma^2}} = (2\pi)^{-\frac{n}{2}} (\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n x_i^2}$$

$$\Rightarrow \ln L(\sigma^2) = -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n x_i^2$$

从而有

$$\frac{d \ln L(\sigma^2)}{d\sigma^2} = 0 \Rightarrow -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n x_i^2 = 0,$$

解之即得 σ^2 的极大似然估计值为 $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2$.

4. 教材习题八(A)三第6题

$$\begin{aligned} \text{似然函数为 } L(\mu, \sigma^2) &= \prod_{i=1}^n f(x_i, \mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(\ln x_i - \mu)^2} \\ &= (2\pi)^{-\frac{n}{2}} (\sigma^2)^{-\frac{n}{2}} \left[\prod_{i=1}^n \frac{1}{x_i} \right] e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (\ln x_i - \mu)^2} \end{aligned}$$

$$\Rightarrow \ln L(\mu, \sigma^2) = -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) - \sum_{i=1}^n \ln x_i - \frac{1}{2\sigma^2} \sum_{i=1}^n (\ln x_i - \mu)^2$$

从而有

$$\begin{cases} \frac{\partial \ln L(\mu, \sigma^2)}{\partial \mu} = 0 \\ \frac{\partial \ln L(\mu, \sigma^2)}{\partial \sigma^2} = 0 \end{cases} \Rightarrow \begin{cases} \frac{1}{\sigma^2} \sum_{i=1}^n (\ln x_i - \mu) = 0 \\ -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (\ln x_i - \mu)^2 = 0 \end{cases},$$

解之即得 μ, σ^2 的极大似然估计值为 $\begin{cases} \hat{\mu} = \frac{1}{n} \sum_{i=1}^n \ln x_i \\ \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \left(\ln x_i - \frac{1}{n} \sum_{j=1}^n \ln x_j \right)^2 \end{cases}$, 极大似然估计

$$\text{计量为 } \begin{cases} \hat{\mu} = \frac{1}{n} \sum_{i=1}^n \ln X_i \\ \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \left(\ln X_i - \frac{1}{n} \sum_{j=1}^n \ln X_j \right)^2 \end{cases}.$$

第 53-54 页

一. 1. B 2. B 3. B

二. 1. $\frac{1}{6}$ 2. $\frac{1}{2(n-1)}$

三. 1. $E(\overline{X^2}) = \frac{\sigma^2}{n} + \mu^2 \rightarrow \mu^2$

2. 教材习题八(A)三第7题

$$\text{因 } E(|X - \mu_0|) = \int_{-\infty}^{+\infty} |x - \mu_0| f_X(x) dx = \int_{-\infty}^{+\infty} |x - \mu_0| \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu_0)^2}{2\sigma^2}} dx, \text{ 令}$$

$$y = \frac{x - \mu_0}{\sigma}, \text{ 再因被积函数是偶函数得 } E(|X - \mu_0|) = 2\sigma \int_0^{+\infty} y \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy, \text{ 用 } \Gamma$$

$$\text{函数计算可得 } E(|X - \mu_0|) = \frac{2\sigma}{\sqrt{2\pi}} \int_0^{+\infty} e^{-\frac{y^2}{2}} d\frac{y^2}{2} = \frac{2\sigma}{\sqrt{2\pi}} \Gamma(1) = \frac{2\sigma}{\sqrt{2\pi}}, \text{ 从而有}$$

$$E(\hat{\sigma}) = \frac{1}{n} \sqrt{\frac{\pi}{2}} \sum_{i=1}^n E(|X_i - \mu_0|) = \frac{1}{n} \sqrt{\frac{\pi}{2}} \sum_{i=1}^n E(|X - \mu_0|) = \sqrt{\frac{\pi}{2}} E(|X - \mu_0|)$$

$$= \sqrt{\frac{\pi}{2}} \cdot \frac{2\sigma}{\sqrt{2\pi}} = \sigma, \text{ 所以 } \hat{\sigma} \text{ 为 } \sigma \text{ 的无偏估计量.}$$

3. 教材习题八(A)三第8题

$$\text{由《习题册》第54页第3题知 } \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2.$$

(1) 因 $E(X) = 0$, 所以 $E(X^2) = D(X) = \sigma^2$, 从而有

$$E(\hat{\sigma}^2) = \frac{1}{n} \sum_{i=1}^n E(X_i^2) = \frac{1}{n} \sum_{i=1}^n E(X^2) = \frac{1}{n} \sum_{i=1}^n \sigma^2 = \sigma^2,$$

所以 $\hat{\sigma}^2$ 为 σ^2 的无偏估计量.

(2) 因总体 $X \sim N(0, \sigma^2)$, 所以

$$\chi_0^2 = \sum_{i=1}^n \frac{X_i^2}{\sigma^2} \sim \chi^2(n), \chi^2 = \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1),$$

从而有

$$D(\hat{\sigma}^2) = D\left(\frac{1}{n} \sum_{i=1}^n X_i^2\right) = D\left(\frac{\sigma^2}{n} \chi_0^2\right) = \frac{\sigma^4}{n^2} D(\chi_0^2) = \frac{\sigma^4}{n^2} 2n = \frac{2\sigma^4}{n};$$

$$D(S^2) = D\left(\frac{\sigma^2}{n-1} \chi^2\right) = \frac{\sigma^4}{(n-1)^2} D(\chi^2) = \frac{\sigma^4}{(n-1)^2} 2(n-1) = \frac{2\sigma^4}{n-1};$$

显然可见 $D(\hat{\sigma}^2) = \frac{2\sigma^4}{n} < \frac{2\sigma^4}{n-1} = D(S^2)$, 所以 $\hat{\sigma}^2$ 比 S^2 更有效.

4. 教材习题八(A)三第9题 略
5. 教材习题八(A)三第10题 上课已讲!

第 55-56 页 《区间估计》

一. 1. B 2. $4u^2 \frac{\sigma_0^2}{1-\frac{\alpha}{2}} l^2$

二. 1. 教材习题八(A)三第11题

计算得 $\bar{x} = 575.2$; $s = 8.7025$

(1) 因 $1 - \alpha = 0.95$, $\sigma^2 = \sigma_0^2 = 25$, 故 μ 的 95% 置信区间为

$$\left(\hat{\mu}_1, \hat{\mu}_2 \right) = \left(\bar{x} - u_{1-\frac{\alpha}{2}} \frac{\sigma_0}{\sqrt{n}}, \bar{x} + u_{1-\frac{\alpha}{2}} \frac{\sigma_0}{\sqrt{n}} \right) = (572.101, 578.299)$$

(2) 因 $\sigma = 5$ 时, $C_v = \frac{\sigma}{|\mu|} = \frac{5}{\mu} (\mu > 0)$ 关于 μ 严格单减, 所以 C_v 的 95% 置信区间

$$\text{为 } \left(\hat{C}_v^1, \hat{C}_v^2 \right) = \left(C_v(\hat{\mu}_2), C_v(\hat{\mu}_1) \right) = \left(\frac{5}{\hat{\mu}_2}, \frac{5}{\hat{\mu}_1} \right) = (0.00865, 0.00874);$$

(3) 因 $1 - \alpha = 0.95$, $t_{1-\frac{\alpha}{2}}(n-1) = t_{0.975}(9) = 2.2622$, 又 σ^2 未知, 故 μ 的 95% 置信区间为

$$\left(\bar{x} - t_{1-\frac{\alpha}{2}}(n-1) \frac{s}{\sqrt{n}}, \bar{x} + t_{1-\frac{\alpha}{2}}(n-1) \frac{s}{\sqrt{n}} \right) = (568.974, 581.426).$$

2. 教材习题八(A)三第12题

经计算可得 $\bar{x} = 10.08333$, $s^2 = 0.06333$, $s = 0.83467$,

查表可得 $t_{1-\frac{\alpha}{2}}(n-1) = t_{0.95}(11) = 1.7959$, $\chi_{\frac{\alpha}{2}}^2(n-1) = \chi_{0.05}^2(11) = 4.575$,

$\chi_{1-\frac{\alpha}{2}}^2(n-1) = \chi_{0.95}^2(11) = 19.675$.

因总体均值 μ 及方差 σ^2 未知, 又 $n = 12$, 故 μ, σ^2, σ 的置信度为 90% 的置信区间为

$$\left(\bar{x} - t_{1-\frac{\alpha}{2}}(n-1) \frac{s}{\sqrt{n}}, \bar{x} + t_{1-\frac{\alpha}{2}}(n-1) \frac{s}{\sqrt{n}} \right) = (9.6506, 10.5160)$$

$$\left(\frac{(n-1)s^2}{\chi^2_{1-\frac{\alpha}{2}}(n-1)}, \frac{(n-1)s^2}{\chi^2_{\frac{\alpha}{2}}(n-1)} \right) = (0.0354, 0.1523)$$

$$\left(\sqrt{\frac{(n-1)s^2}{\chi^2_{1-\frac{\alpha}{2}}(n-1)}}, \sqrt{\frac{(n-1)s^2}{\chi^2_{\frac{\alpha}{2}}(n-1)}} \right) = (0.1881, 0.3902)$$

3. 教材习题八(A)三第 13 题 略

4. 教材习题八(A)三第 14 题 略

第 57-58 页

一. 1. B 2. D 教材习题七(B)一第 2 题 3. C

二. 1. 16 2. σ^2 3. $\frac{1}{X}$ $\frac{1}{\bar{X}}$

三. 1. 教材习题五(B)三第 2 题

$$\begin{aligned} E(M) &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} \max(x, y) f_X(x) f_Y(y) dy \\ &= \int_{-\infty}^{+\infty} dx \int_x^{+\infty} y \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy + \int_{-\infty}^{+\infty} dx \int_{-\infty}^x x \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy \\ &= \frac{1}{2\pi} \left[\int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx \int_x^{+\infty} e^{-\frac{y^2}{2}} d\frac{y^2}{2} + \int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy \int_y^{+\infty} e^{-\frac{x^2}{2}} d\frac{x^2}{2} \right] \\ &= \frac{1}{2\pi} \left[\int_{-\infty}^{+\infty} e^{-x^2} dx + \int_{-\infty}^{+\infty} e^{-y^2} dy \right] = \frac{2}{\pi} \int_{-\infty}^{+\infty} e^{-x^2} dx \\ &= \frac{1}{\pi} \int_0^{+\infty} (x^2)^{-\frac{1}{2}} e^{-x^2} dx^2 = \frac{1}{\pi} \Gamma\left(\frac{1}{2}\right) = \frac{1}{\sqrt{\pi}} \end{aligned}$$

$$\begin{aligned} E(N) &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} \min(x, y) f_X(x) f_Y(y) dy \\ &= \int_{-\infty}^{+\infty} dx \int_x^{+\infty} x \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy + \int_{-\infty}^{+\infty} dx \int_{-\infty}^x y \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{2\pi} \left[\int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy \int_{-\infty}^y x e^{-\frac{x^2}{2}} dx + \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx \int_{-\infty}^x y e^{-\frac{y^2}{2}} dy \right] \\
&= \frac{1}{\pi} \int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy \int_{-\infty}^y x e^{-\frac{x^2}{2}} dx = \frac{1}{\pi} \int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy \int_{-\infty}^y e^{-\frac{x^2}{2}} d\frac{x^2}{2} \\
&= -\frac{1}{\pi} \int_{-\infty}^{+\infty} e^{-y^2} dy = -\frac{2}{\pi} \int_0^{+\infty} e^{-y^2} dy \\
&= -\frac{1}{\pi} \int_0^{+\infty} (y^2)^{-\frac{1}{2}} e^{-y^2} dy^2 = -\frac{1}{\pi} \Gamma\left(\frac{1}{2}\right) = -\frac{1}{\sqrt{\pi}}
\end{aligned}$$

2. 教材习题八(B)三第2题

矩估计法：总体的一阶原点矩为

$$m_1 = E(X) = \int_0^1 x\theta dx + \int_1^2 x(1-\theta) dx = \frac{1}{2}\theta + \frac{3}{2}(1-\theta) = \frac{3}{2} - \theta$$

反解得 $\theta = \frac{3}{2} - m_1$ ，用样本的一阶原点矩（即样本均值）代替总体的一阶原点矩得

未知参数 θ 的矩估计量为 $\hat{\theta}_M = \frac{3}{2} - \bar{X}$ ，其矩估计值为 $\hat{\theta}_M = \frac{3}{2} - \bar{x}$ ；

极大似然估计法： $L(\theta) = \prod_{i=1}^n f(x_i; \theta) = \theta^N (1-\theta)^{n-N}$

$\Rightarrow \ln L(\theta) = N \ln \theta + (n-N) \ln(1-\theta)$ ，由此得

$$\frac{d \ln L(\theta)}{d\theta} = 0 \Leftrightarrow \frac{N}{\theta} - \frac{n-N}{1-\theta} = 0, \text{ 解之即得 } \theta \text{ 的矩估计值为 } \hat{\theta}_L = \frac{N}{n}.$$

3. 教材习题八(B)三第3题

此题即为在一个正态总体下，总体均值未知时对总体总体标准差的区间估计。

$$\begin{aligned}
&\left(\sqrt{\frac{(n-1)s^2}{\chi_{1-\alpha/2}^2(n-1)}}, \sqrt{\frac{(n-1)s^2}{\chi_{\alpha/2}^2(n-1)}} \right) = \left(\sqrt{\frac{0.17985}{\chi_{0.975}^2(9)}}, \sqrt{\frac{0.17985}{\chi_{0.025}^2(9)}} \right) \\
&= \left(\sqrt{\frac{0.17985}{19.023}}, \sqrt{\frac{0.17985}{2.7}} \right) = (0.0972, 0.2581), \text{ 显然可见, 在 } 95\% \text{ 的置信度下, 这批}
\end{aligned}$$

零件长度的标准差符合设计要求。