

§ 5.3 二维正态分布

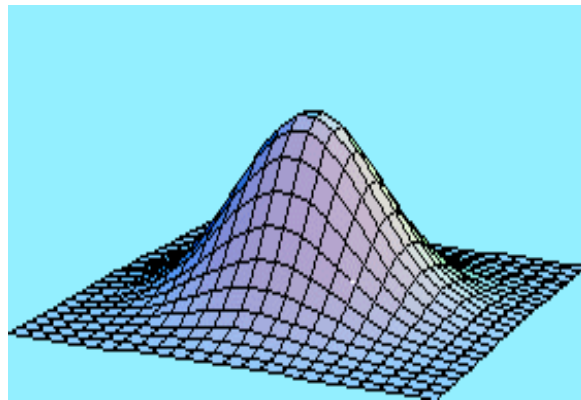
定义5.3 若二维连续型随机变量 (X,Y) 的密度函数为

$$f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-r^2}} e^{-\frac{1}{2(1-r^2)}\left[\frac{(x-u_1)^2}{\sigma_1^2} + \frac{(y-u_2)^2}{\sigma_2^2} - 2r\frac{(x-u_1)(y-u_2)}{\sigma_1\sigma_2}\right]},$$

$-\infty < x, y < +\infty$, (where $u_1, u_2 \in R, \sigma_1, \sigma_2 > 0, |r| < 1$),

则称 (X,Y) 服从参数为 $u_1, u_2, \sigma_1^2, \sigma_2^2, r$ 的二维正态分布分布，记为 $(X,Y): N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; r)$ 。

二维正态分布密度
函数图像



二维正态分布的边缘分布

定理5.6 设 $(X, Y) : N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; r)$ 则

$$X : N(\mu_1, \sigma_1^2), \quad Y : N(\mu_2, \sigma_2^2).$$

证明：记 $t = \frac{x - \mu_1}{\sigma_1}$, 并作变量代换 $s = \frac{y - \mu_2}{\sigma_2}$, 则
 $dy = \sigma_2 ds$, 从而有

$$\begin{aligned} & -\frac{1}{2(1-r^2)} \left[\frac{(x - \mu_1)^2}{\sigma_1^2} + \frac{(y - \mu_2)^2}{\sigma_2^2} - 2r \frac{(x - \mu_1)(y - \mu_2)}{\sigma_1 \sigma_2} \right] \\ &= -\frac{1}{2(1-r^2)} (t^2 + s^2 - 2rst) = -\frac{(s - rt)^2}{2(1-r^2)} - \frac{t^2}{2} \end{aligned}$$

$$\begin{aligned}
 f_X(x) &= \int_{-\infty}^{+\infty} f(x, y) dy = \int_{-\infty}^{+\infty} \frac{1}{2\pi\sigma_1\sqrt{1-r^2}} e^{-\frac{(s-rt)^2}{2(1-r^2)} - \frac{t^2}{2}} ds \\
 &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{t^2}{2}} \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sqrt{1-r^2}} e^{-\frac{(s-rt)^2}{2(1-r^2)}} ds \\
 &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{t^2}{2}} \\
 &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}, -\infty < x < +\infty,
 \end{aligned}$$

$N(rt, 1-r^2)$ 的密度函数

所以 $X : N(\mu_1, \sigma_1^2)$. 同理可证 $Y : N(\mu_2, \sigma_2^2)$.

例： 设二维随机变量 $(X, Y) \sim N(u_1, u_2, \sigma_1^2, \sigma_2^2, r)$ ，求 (X, Y) 的协方差矩阵以及相关系数。

解： 因 $(X, Y) \sim N(u_1, u_2, \sigma_1^2, \sigma_2^2, r)$ $X \sim N(u_1, \sigma_1^2)$,
 $Y \sim N(u_2, \sigma_2^2)$, 从而 $E(X) = u_1$, $E(Y) = u_2$, $D(X) = \sigma_1^2$,
 $D(Y) = \sigma_2^2$;

$$\begin{aligned} \text{Cov}(X, Y) &= E(XY) - E(X)E(Y) = E(XY) - u_1u_2 \\ &= E[(X - u_1)(Y - u_2)] \end{aligned}$$

$$\text{令 } \mu = \frac{x - u_1}{\sigma_1}, \nu = \frac{y - u_2}{\sigma_2} \quad d\mu = \frac{dx}{\sigma_1}, \quad d\nu = \frac{dy}{\sigma_2}.$$

$$Cov(X, Y)$$

$$= E[(X - u_1)(Y - u_2)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x - u_1)(y - u_2) f(x, y) dx dy$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{\mu \nu \sigma_1 \sigma_2}{2\pi \sqrt{1-r^2}} e^{-\frac{\mu^2 - 2r\mu\nu + \nu^2}{2(1-r^2)}} d\mu d\nu$$

$$= \int_{-\infty}^{+\infty} \frac{\sigma_1 \sigma_2}{\sqrt{2\pi}} \mu e^{-\frac{\mu^2}{2}} \left[\int_{-\infty}^{+\infty} \frac{\nu}{\sqrt{2\pi} \sqrt{1-r^2}} e^{-\frac{(\nu-r\mu)^2}{2(1-r^2)}} d\nu \right] d\mu$$

$$= \int_{-\infty}^{+\infty} \frac{\sigma_1 \sigma_2}{\sqrt{2\pi}} \mu e^{-\frac{\mu^2}{2}} r \mu d\mu = \frac{2r\sigma_1\sigma_2}{\sqrt{2\pi}} \int_0^{+\infty} \mu^2 e^{-\frac{\mu^2}{2}} d\mu$$

$$= \frac{2r\sigma_1\sigma_2}{\sqrt{\pi}} \int_0^{+\infty} \left(\frac{\mu^2}{2} \right)^{\frac{1}{2}} e^{-\left(\frac{\mu^2}{2} \right)} d\left(\frac{\mu^2}{2} \right) = \frac{2r\sigma_1\sigma_2}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right) = r\sigma_1\sigma_2.$$

从而, (X,Y) 的相关系数为 r ,协方差阵为

$$V = \begin{bmatrix} \sigma_1^2 & r\sigma_1\sigma_2 \\ r\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}.$$

协方差的性质告诉我们

若 X 与 Y 相互独立, 则 $Cov(X,Y) = 0$.

定理5.7: 设随机变量 $(X,Y) : N(u_1, u_2, \sigma_1^2, \sigma_2^2, r)$, 则:
 X 与 Y 相互独立 $\Leftrightarrow r = 0$ 。

思想: X 与 Y 相互独立 \Leftrightarrow 在 $f(x,y)$, $f_X(x)$, $f_Y(y)$ 的公共连续点 (x,y) 处成立 $f(x,y) = f_X(x)f_Y(y)$

二维正态分布的条件分布

当 $\mathbf{X}=\mathbf{x}$ 条件下, \mathbf{Y} 的密度函数可如下计算:

$$f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)} \quad \text{详见Page142,此时}$$

$$Y | X = x \sim N\left(\mu_2 + r \frac{\sigma_2}{\sigma_1} (x - \mu_1); \sigma_2^2 (1 - r^2)\right)$$

定理5.8 二维随机变量 (\mathbf{X}, \mathbf{Y}) 服从二维正态分布的充要条件是 \mathbf{X} 与 \mathbf{Y} 的任意线性组合 $\mathbf{Z}=\mathbf{aX}+\mathbf{bY}$ 服从一维正态分布, 即: $\mathbf{Z} \sim N(\mathbf{E}(\mathbf{Z}), \mathbf{D}(\mathbf{Z}))$.

例5.8 设 $(X, Y) \sim N(2, 3; 4, 9; \frac{1}{2})$, $Z = \frac{1}{2}X - \frac{1}{3}Y$, 求 $E(|Z|)$ 。

解: $X \sim N(2, 4)$, $Y \sim N(3, 9)$, 相关系数 $r = 1/2$.

$$\text{Cov}(X, Y) = r \cdot \sqrt{DX} \cdot \sqrt{DY} = 3, \quad E(Z) = \frac{1}{2}EX - \frac{1}{3}EY = 0$$

$$D(Z) = \frac{1}{4}DX + \frac{1}{9}DY - 2 \cdot \frac{1}{2} \cdot \frac{1}{3} \text{Cov}(X, Y) = 1$$

由定理**5.8** $\therefore Z \sim N(0,1)$

$$E(|Z|) = \int_{-\infty}^{\infty} |z| f(z) dz$$

$$= \int_{-\infty}^{\infty} |z| \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = 2 \int_0^{\infty} z \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} d(-e^{-\frac{z^2}{2}}) = \frac{\sqrt{2}}{\sqrt{\pi}}$$

习题: X, Y 独立同分布 $N(\mu, \sigma^2)$.

$Z_1 = \alpha X + \beta Y, Z_2 = \alpha X - \beta Y, \alpha, \beta \in R$, 求 $\rho_{Z_1 Z_2}$

解: $E(Z_1) = \alpha E(X) + \beta E(Y) = (\alpha + \beta) \mu$

$E(Z_2) = \alpha E(X) - \beta E(Y) = (\alpha - \beta) \mu$

$D(Z_1) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2) \sigma^2$

$D(Z_2) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2) \sigma^2$

$E(Z_1 Z_2) = E((\alpha X + \beta Y)(\alpha X - \beta Y))$

$= E(\alpha^2 X^2 - \beta^2 Y^2) = \alpha^2 E(X^2) - \beta^2 E(Y^2)$

$= \alpha^2 (D(X) + E^2(X)) - \beta^2 (D(Y) + E^2(Y))$

$= \alpha^2 (\sigma^2 + \mu^2) - \beta^2 (\sigma^2 + \mu^2) = (\alpha^2 - \beta^2)(\sigma^2 + \mu^2)$

$$\begin{aligned}\rho_{Z_1 Z_2} &= \frac{(\alpha^2 - \beta^2)(\sigma^2 + \mu^2) - \mu^2(\alpha^2 - \beta^2)}{\sigma^2(\alpha^2 + \beta^2)} \\ &= \frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2}\end{aligned}$$

习题： X, Y 独立同分布 $N(\mu, \sigma^2)$. $U = \max(X, Y)$, $V = \min(X, Y)$, 求 U, V 的期望.

解：直接按照确定积分区域的方法计算比较烦琐.

令： $Z = X - Y$. X, Y 互相独立, 故 Z 服从正态分布. 即 $Z \sim N(0, 2\sigma^2)$.

$$U = \frac{1}{2}(X + Y + |X - Y|), V = \frac{1}{2}(X + Y - |X - Y|)$$

$$E(|X - Y|) = E(|Z|) = 2 \int_0^{+\infty} \frac{z}{\sqrt{2\pi} \cdot \sqrt{2}\sigma} e^{-\frac{z^2}{2 \times 2\sigma^2}} dz$$

$$= 2 \frac{1}{2\sigma\sqrt{\pi}} \int_0^{+\infty} z e^{-\frac{z^2}{4\sigma^2}} dz = \frac{2\sigma}{\sqrt{\pi}}$$

$$E(U) = \frac{1}{2}(EX + EY + E(|Z|)) = \mu + \frac{\sigma}{\sqrt{\pi}}, E(V) = \mu - \frac{\sigma}{\sqrt{\pi}}$$

练习一：设 $\xi \sim N(\mu, \sigma^2)$, $\eta = a + b\xi$ ($b \neq 0$). 求 $f_\eta(y)$.

解：
$$f_\xi(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad x \in R$$

设 η 的分布函数 $F_\eta(y)$.

(1) 当 $b > 0$ 时,

$$\begin{aligned} F_\eta(y) &= P\{\eta \leq y\} = P\{a + b\xi \leq y\} = P\left\{\xi \leq \frac{y-a}{b}\right\} \\ &= F_\xi\left(\frac{y-a}{b}\right) \end{aligned}$$

$$f_\eta(y) = F'_\eta(y) = f_\xi\left(\frac{y-a}{b}\right) \cdot \frac{1}{b}$$

$$= \frac{1}{\sqrt{2\pi}\sigma b} e^{-\frac{(y-a-b\mu)^2}{2\sigma^2 b^2}}, \quad y \in R$$

(2) 当**b<0**时,

$$\begin{aligned} F_{\eta}(y) &= P\{\eta \leq y\} = P\{a + b\xi \leq y\} = P\{\xi \geq \frac{y-a}{b}\} \\ &= 1 - P\{\xi < \frac{y-a}{b}\} = 1 - F_{\xi}(\frac{y-a}{b}) \end{aligned}$$

$$f_{\eta}(y) = -\frac{1}{\sqrt{2\pi\sigma b}} e^{-\frac{(y-a-b\mu)^2}{2\sigma^2 b^2}} \quad y \in R$$

综合起来, 对于 $\eta = a + b\xi (b \neq 0)$, 有

$$f_{\eta}(y) = \frac{1}{\sqrt{2\pi\sigma |b|}} e^{-\frac{(y-a-b\mu)^2}{2\sigma^2 b^2}}, \quad y \in R$$

练习二：

1、设随机变量 $X \sim N(0,1)$, $Y \sim U(0,1)$, $Z \sim B(5,0.5)$,且 X, Y, Z 独立, 求随机变量 $U=(2X+3Y)(4Z-1)$ 的数学期望.

答: $E(U) = E(2X + 3Y)E(4Z - 1) = \frac{27}{2}$

2、设随机变量 X_1, X_2, \dots, X_n 相互独立, 且均服从 $N(\mu, \sigma^2)$ 分布, 求随机变量 $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ 的数学期望.

答: $E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \mu$