

Traditional Probability Theory

Opposition : $P(A) + P(\bar{A}) = 1$ **Exclusive** : $A \cap B = \emptyset \Rightarrow P(AB) = 0$

Independent : $P(AB) = P(A)P(B)$ **Equal** : $A = B \Rightarrow P(A) = P(B)$

$A - B = A\bar{B} = A \cap \bar{B} \Rightarrow P(A - B) = P(A) - P(AB)$

$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(BC) - P(AC) - P(AB) + P(ABC)$

$P(B|A) = \frac{P(AB)}{P(A)}$ $P(A_1 A_2 \dots A_n) = P(A_1)P(A_2|A_1)P(A_3|A_1 A_2) \dots P(A_n|A_1 \dots A_{n-1})$

Bayes $P(B_j|A) = \frac{P(A|B_j)P(B_j)}{\sum_{i=1}^n P(A|B_i)P(B_i)}$

Random Variable Digital Properties

Distr	Mark	EX	DX	Addition
Bin	$B(n, p)$	np	$np(1 - p)$	$P\{X = k\} = \mathbf{C}_n^k (1 - p)^{n-k} p^k$
Poi	$\pi(\lambda)$	λ	λ	$P\{X = k\} = \frac{\lambda^k}{k!} e^{-\lambda}, k = 0, 1, 2, \dots$
Geo	$G(p)$	$\frac{1}{p}$	$\frac{1-p}{p^2}$	$P\{X = k\} = (1 - p)^{k-1} p, k = 1, 2, \dots$
Hyp	$H(n, M, N)$	$\frac{nM}{N}$	$\frac{nM}{N} (1 - \frac{M}{N}) (\frac{N-n}{N-1})$	$P\{X = i\} = \frac{\mathbf{C}_M^i \mathbf{C}_{N-M}^{n-i}}{\mathbf{C}_N^n}$
Uni	$U(a, b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	
Exp	$E(\lambda)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	
Nor	$N(\mu, \sigma^2)$	μ	σ^2	

$$\begin{aligned}
\mathbf{Uni} : f(x) &= \begin{cases} \frac{1}{b-a}, a \leq x \leq b \\ 0, \text{others} \end{cases} & F(x) &= \begin{cases} 0, x < a \\ \frac{x-a}{b-a}, a \leq x < b \\ 1, x \geq b \end{cases} \\
\mathbf{Exp} : f(x) &= \begin{cases} \lambda e^{-\lambda x}, x > 0 \\ 0, x \leq 0 \end{cases} & F(x) &= \begin{cases} 1 - e^{-\lambda x}, x > 0 \\ 0, x \leq 0 \end{cases} \\
\mathbf{Nor} : f(x) &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, F(x) = \int_{-\infty}^x f(t)dt, x \in (-\infty, +\infty) \\
f(\mu+x) &= f(\mu-x), F(\mu+x) + F(\mu-x) = 1, F(\mu) = \frac{1}{2} \\
X &\sim N(0, 1), \phi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, \Phi(x) = \int_{-\infty}^x \phi(t)dt \\
\phi(-x) &= \phi(x), \quad \Phi(a) + \Phi(-a) = 1, \Phi(0) = \frac{1}{2}, \quad F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right) \\
f(x, y) &\geq 0 \quad \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1 \\
f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy > 0 \quad f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx > 0 \\
f_{X|Y}(x|y) &= \frac{f(x, y)}{f_Y(y)}, f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} \\
\mathbf{Discrete} : P\{Z = g(x_i, y_j)\} &= P\{X = x_i, Y = y_j\} = P_{ij} \\
F_Z(z) &= P\{Z \leq z\} = P\{g(X, Y) \leq z\} = \sum_{g(x_i, y_j) \leq z} P\{X = x_i, Y = y_j\} \\
\mathbf{Continuous} : Z = g(X, Y) \quad F_Z(z) &= P\{Z \leq z\} = P\{g(X, Y) \leq z\} = \iint_{g(x, y) \leq z} f(x, y) dx dy \\
Z = \max(X, Y), F_{\max}(z) &= F_X(z)F_Y(z) \quad Z = \min(X, Y), F_{\min}(z) = 1 - [1 - F_X(z)][1 - F_Y(z)] \\
Z_1 = \max(X, Y) &= \frac{X + Y + |X - Y|}{2} \quad Z_2 = \min(X, Y) = \frac{X + Y - |X - Y|}{2} \quad Z_1 Z_2 = XY \\
(X, Y) &\sim U(D), f(x, y) = \begin{cases} \frac{1}{S_D}, (x, y) \in D \\ 0, \text{others} \end{cases} \\
(X, Y) &\sim N(\mu_1, \mu_2; \sigma_1, \sigma_2, \rho), f(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}[(\frac{x-\mu_1}{\sigma_1})^2 - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + (\frac{y-\mu_2}{\sigma_2})^2]}
\end{aligned}$$

$$\begin{aligned}
F(x) &= P\{X \leq x\} = \int_{-\infty}^x f(x)dx \\
F(x, y) &= P\{X \leq x, Y \leq y\} = \int_{-\infty}^x \int_{-\infty}^y f(u, v)dvdu \\
1 &= \sum_{i=1}^{\infty} x_i p_i \quad 1 = \int_{-\infty}^{+\infty} f(x)dx \\
E(X) &= \sum_{i=1}^{\infty} x_i p_i \quad E(X) = \int_{-\infty}^{+\infty} x f(x)dx \\
E[g(X)] &= \sum_{i=1}^{\infty} g(x_i) p_i \quad E[g(X)] = \int_{-\infty}^{+\infty} g(x) f(x)dx \\
E[g(X, Y)] &= \sum_{j=1}^{\infty} \sum_{i=1}^{\infty} g(x_i, y_j) P_{ij} \quad E[g(X, Y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f(x, y) dx dy \\
\exists X \sim F(x), \forall X > 0, E(X) &= \int_0^{+\infty} (1 - F(x)) dx = \int_0^{+\infty} P\{X > x\} dx \\
E(C) &= C \quad E(CX) = CE(X) \quad E(X + C) = E(X) + C \quad E(X + Y) = E(X) + E(Y) \\
D(X) &= \sum_{i=1}^n [x_i - E(X)]^2 p_i \quad D(X) = \int_{-\infty}^{+\infty} [x - E(X)]^2 f(x) dx \\
D(X) &= E(X^2) - [E(X)]^2 \quad D(C) = 0 \quad D(CX) = C^2 D(X) \quad D(X + C) = D(X) \\
Cov(X, Y) &= E(XY) - E(X)E(Y) \quad D(X \pm Y) = D(X) + D(Y) \pm 2Cov(X, Y) \\
Cov(X, Y) &= Cov(Y, X) \quad Cov(X, X) = D(X) \quad Cov(X, c) = 0 \\
Cov(aX, bY) &= abCov(X, Y) \quad Cov(X_1 + X_2, Y) = Cov(X_1, Y) + Cov(X_2, Y) \\
\rho_{XY} &= \frac{E(XY) - E(X)E(Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{Cov(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} \\
X &\sim N(0, 1) \quad E(X^{2k}) = (2k - 1)!! \quad E(X^{2k-1}) = 0, k \in \{1, 2, 3, \dots\}
\end{aligned}$$

Large Number Law Central Limit Theorem

$$P\{|X - E(X)| \geq \epsilon\} \leq \frac{D(X)}{\epsilon^2}$$

$$\lim_{n \rightarrow \infty} P\{|\frac{X_n}{n} - p| < \epsilon\} = 1$$

$$\lim_{n \rightarrow \infty} P\{|\frac{1}{n} \sum_{k=1}^n X_k - \mu| < \epsilon\} = 1$$

$$\lim_{n \rightarrow \infty} P\{|\frac{1}{n} \sum_{k=1}^n X_k - \frac{1}{n} \sum_{k=1}^n E(X_k)| < \epsilon\} = 1$$

$$\sum_{i=1}^n X_i \sim N(n\mu, n\sigma^2), \lim_{n \rightarrow \infty} P\{\frac{\sum_{i=1}^n X_i - n\mu}{\sqrt{n}\sigma} \leq x\} = F_n(x) = \Phi(x)$$

$$X \sim N(np, np(1-p)), \lim_{n \rightarrow \infty} P\{\frac{X_n - np}{\sqrt{np(1-p)}} \leq x\} = \Phi(x)$$

Mathematical Statistics

$$\bar{X} \sim N(\mu, \frac{\sigma^2}{n}), \frac{\bar{X} - u}{\sigma/\sqrt{n}} \sim N(0, 1)$$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} [\sum_{i=1}^n X_i^2 - n\bar{X}^2]$$

$$E(S^2) = \sigma^2, D(S^2) = \frac{2\sigma^4}{n-1}$$

$$\frac{\sum_{i=1}^n (X_i - \mu)^2}{\sigma^2} \sim \chi^2(n)$$

$$\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma^2} \sim \chi^2(n-1)$$

$$\chi^2 \sim \chi^2(n), E(\chi^2(n)) = n, D(\chi^2(n)) = 2n$$

$$T \sim t(n), T = \frac{X}{\sqrt{Y/n}} \sim \frac{N(0, 1)}{\sqrt{\chi^2(n)/n}}$$

$$F \sim F(n_1, n_2), F = \frac{X/n_1}{Y/n_2} \sim \frac{\chi^2(n_1)/n_1}{\chi^2(n_2)/n_2}$$

$$(X, Y) \sim N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2, 0)$$

$$\frac{\frac{(n_1-1)S_1^2}{\sigma_1^2}/(n_1-1)}{\frac{(n_2-1)S_2^2}{\sigma_2^2}/(n_2-1)} \sim F(n_1-1, n_2-1)$$

$$\frac{\frac{\sum_{i=1}^{n_1} (X_i - \mu_1)^2}{\sigma_1^2}/n_1}{\frac{\sum_{i=1}^{n_2} (X_i - \mu_2)^2}{\sigma_2^2}/n_2} \sim F(n_1, n_2)$$

$$\frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1)$$

$$\exists \sigma_1 = \sigma_2, \frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{S_\omega \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t(n_1 + n_2 - 2)$$

$$S_\omega = \sqrt{\frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1 + n_2 - 2}}$$