# The Bayesian Blocks algorithm from time series analysis to histogram representation

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#### Outline

- The bayesian block representation is a non-parametric representation of data derived with a bayesian statistical procedure
- Invented by Jeffrey D. Scargle and applied in the context of astronomical time series analysis (GRBs hunting and characterization etc.)

This presentation is mainly based on the following works:

- [1] J. D. Scargle *et al.*, Astrophys. J. 764 (2013) 167
- [2] B. Pollack et al. (2017), arXiv:1708.00810
- [3] J. D. Scargle, Astrophys. J. 504 (1998) 405

#### Features in the wish list

- Non-parametric: generic representation of data (not fitting!). Another famous technique on the market is e.g. kernel density estimation (KDE)
- Discover local structures in background data exploiting the full information brought by the data
- Impose few preconditions as possible on signal and background shapes
- Handle arbitrary sampling and dynamic ranges of data
- Operate in a bayesian framework and work with posterior probabilities:

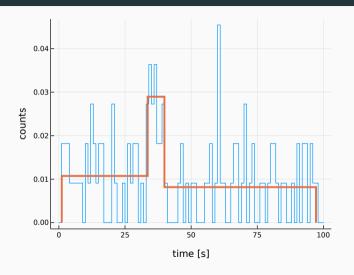
$$P(M|D) \propto P(D|M)P(M)$$

## The algorithm

#### The idea

Segmentation of the data interval into variable-sized blocks, each block containing consecutive data elements satisfying some well-defined criterion.

The *optimal segmentation* is the one that maximizes some quantification of this criterion.

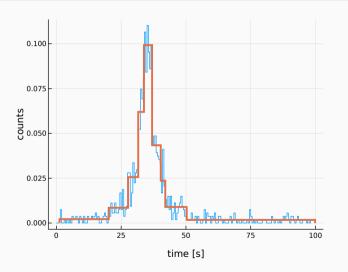


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# The fitness function (a.k.a. the likelihood) -P(D|M)P(M)

We'll consider here, for simplicity, 1-dim data and the *Piecewise Constant Model* (a.k.a. constant representation of data within a block)

- · key property: a block's fitness depends on its data only!
- We are interested only in the length spanned by the block, e.g. the height is just a nuisance parameter
- Recipe: build your fitness function and then marginalize (or maximize/minimize) wrt the nuisance parameters
- The fitness of the entire partition will be the product of each block's fitness

#### The fitness function – Cash statistics

There's a considerable freedom in choosing the fitness function (rely on sufficient statistics). Let's use here the Cash statistics (Cash 1979). With a model  $M(t, \vartheta)$ , the unbinned log-likelihood reads:

$$\log L(\vartheta) = \sum_{n} \log M(t_n, \vartheta) - \int M(t, \vartheta) dt$$

our model is constant:  $M(t, \lambda) = \lambda$ , so, for block k:

$$\log L^{(k)}(\lambda) = N^{(k)} \log \lambda - \lambda T^{(k)}$$

Now we maximize wrt the nuisance parameter  $\lambda$  (height of the block)

$$\log L_{\max}^{(k)} = N^{(k)} (\log N^{(k)} - \log T^{(k)}) + N^{(k)}$$

#### The fitness function - Cash statistics

The  $N^{(k)}$  term sums up to N so it can be dropped because it's independent of the partition

$$\log L_{\max}^{(k)} = N^{(k)} (\log N^{(k)} - \log T^{(k)}) + \mathcal{N}^{(k)}$$

Has nice features:

- · it's simple...
- ...and scale invariant! (try to  $T \rightarrow \alpha T$ )

The fitness of the entire partition will be then:

$$\log L = \sum_{k} \log L_{\max}^{(k)}$$

## The prior for the number of blocks -P(D|M)P(M)

A flat prior on the number of blocks is unreasonable, most of the times one looks for a representation where  $N_{blocks} \ll N$  rather than  $N_{blocks} \approx N$ . For example the "geometric" prior:

$$P(N_{\text{blocks}}) = \begin{cases} P_0 \gamma^{N_{\text{blocks}}} & 0 \le N_{\text{blocks}} \le N \\ 0 & \text{else} \end{cases}$$

has well-understood properties ( $\gamma$  < 1) and is simply implemented in the algorithm.

The value of  $\gamma$  affects the representation (note that, however, sharply defined structures are retained). But wait, there's an objective procedure to select it...

# The prior for $N_{blocks}$ — fixing the parameter

It's a tradeoff between a conservative choice and a liberal choice — it's always a matter of fixing the rate of Type-I errors!

#### General rule

Running the algorithm with a few different values can be enough. In general, the number of change-points is insensitive to a large range of reasonable values of your "steepness" parameter

**Rigorous approach:** calibrate the prior as a function of the number of data points N and the *false-positive rate*  $p_0$  on toy pure-noise experiments. A calibration of this type performed in [1] yields:

$$\log P(N, p_0) = \log(73.53p_0N^{-0.478}) - 4$$

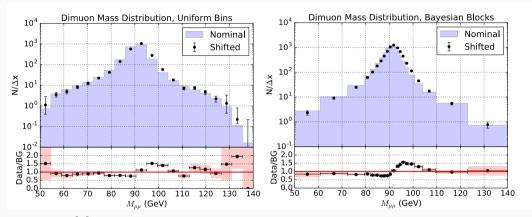
## Applications — binning histograms

The idea was developed to be mainly applied to time series analyses (e.g. to spot light flux changes from astrophysical objects), but has advantages also in binning histograms:

- More objective way to present data, avoid ad-hoc binning (my guess: forget about systematic uncertainties from binning choice?)
- · binning-dependent features can be more objectively spotted
- · differently from other rules (Knuth's, Scott's...) it doesn't use fixed bin width
- attractive when data spans different orders of magnitude
- effective hypothesis testing with the hybrid scheme proposed in [2]

Though can still seem crude and blocky to someone...

# Applications — binning histograms



Taken from [2], note the definiteness of the distortion pattern in the residuals with the bayesian block representation.

#### Implementation

Investigating the optimality of  $2^N$  data partitions isn't a quick task for a computer  $\rightarrow$  dynamic programming approach following the spirit of mathematical induction:

- Sort the data and start from the first one, the only possible partition is trivially optimal
- The optimal partition is updated at each step using the information from the previous ones  $\rightarrow \mathcal{O}(N^2)$

```
for k in 1:N
# fitness function + prior
   F(r) = logfitness(N_k, T_k) + log_prior
   # compute all possible configurations
   A = [F(r) + (r == 1 ? 0 : best[r-1]) for r in 1:k]
   # save best configuration
   push!(last, indmax(A))
   push!(best, maximum(A))
```

#### Implementation

An implementation in Python exists, provided by the AstroML package<sup>1</sup>, but it's targeted to time series analyses and it's slow when applied to large datasets.

I implemented a faster version in Julia<sup>2</sup> available here:

https://tinyurl.com/bayesian-blocks-jl

that can also rebin histograms.

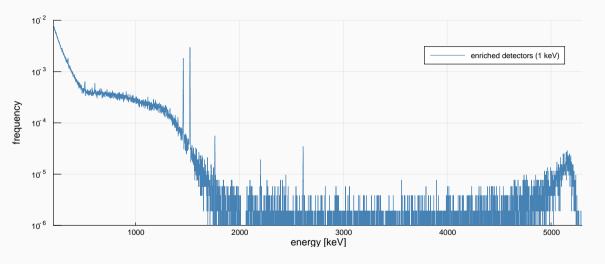
http://www.astroml.org

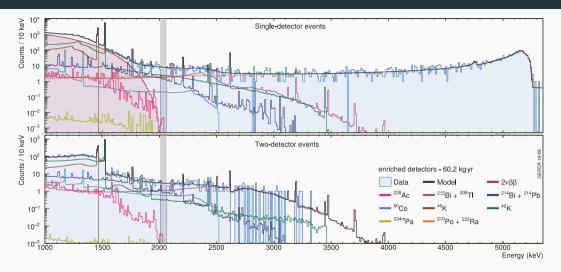
<sup>2</sup>https://julialang.org

# \_\_\_\_

Application: GERDA's energy

spectrum





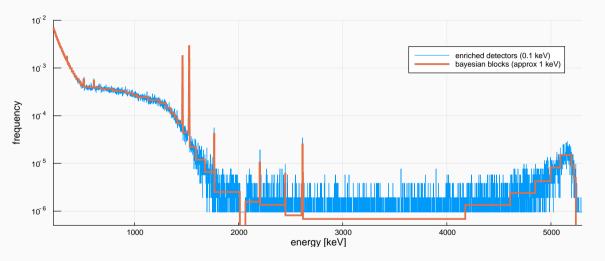
#### **Problem**

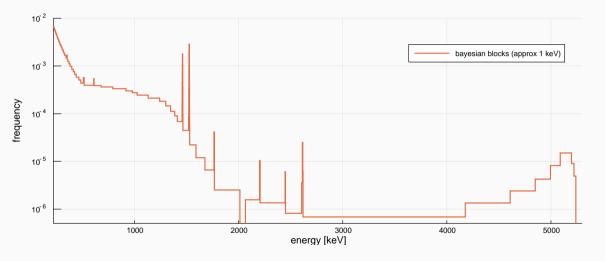
The spectrum consists of  $\sim 10^6$  entries, even with  $\mathcal{O}(N^2)$  instead of  $\mathcal{O}(2^N)$  and fast code the computational effort is quite large!

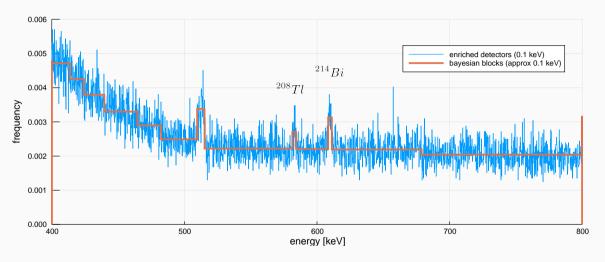
#### **Approximation**

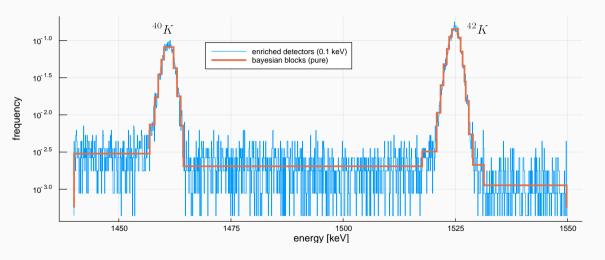
adopted for high statistics: bin the data and treat each bin as a data point of multiplicity = bincontent in the algorithm.

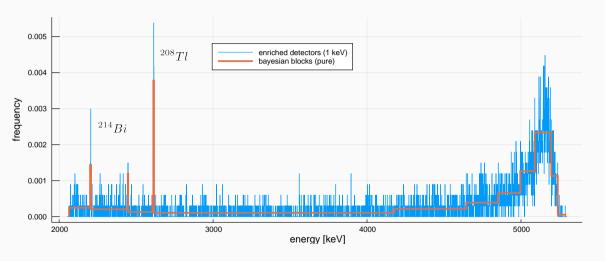
**NB:** I adopted the calibrated prior computed in [1] with  $p_0 = 0.1$  in all the following plots.

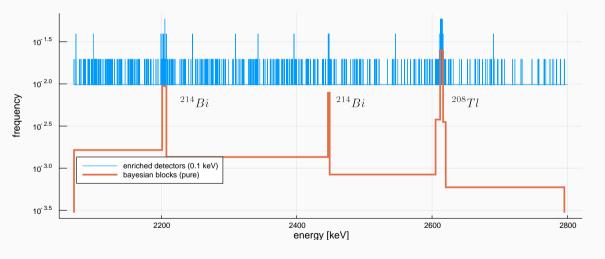












#### Conclusions

The bayesian block representation provides an objective way to enlighten the key features of a data set by imposing few preconditions as possible.

#### Tips:

- · choose a fitness function that's suitable for your data
- · experiment different priors or...
- · ...run the algorithm on toy data to calibrate it

#### Future:

 How is it to fit data with this binning? Is this data representation as informative as a fine, 1 keV, fixed-width binning? → verify statement in [2]

#### References

- [1] Studies in Astronomical Time Series Analysis. VI. Bayesian Block Representations, J. D. Scargle et al., Astrophys. J. 764 (2013) 167
- [2] Bayesian Blocks in High Energy Physics: Better Binning made easy! B. Pollack et al. (2017), arXiv:1708.00810
- [3] Studies in astronomical time series analysis: 5. Bayesian blocks, a new method to analyze structure in photon counting data, J. D. Scargle, Astrophys. J. 504 (1998) 405