

Sampling: Latin Hypercube & Markov Chain Monte Carlo

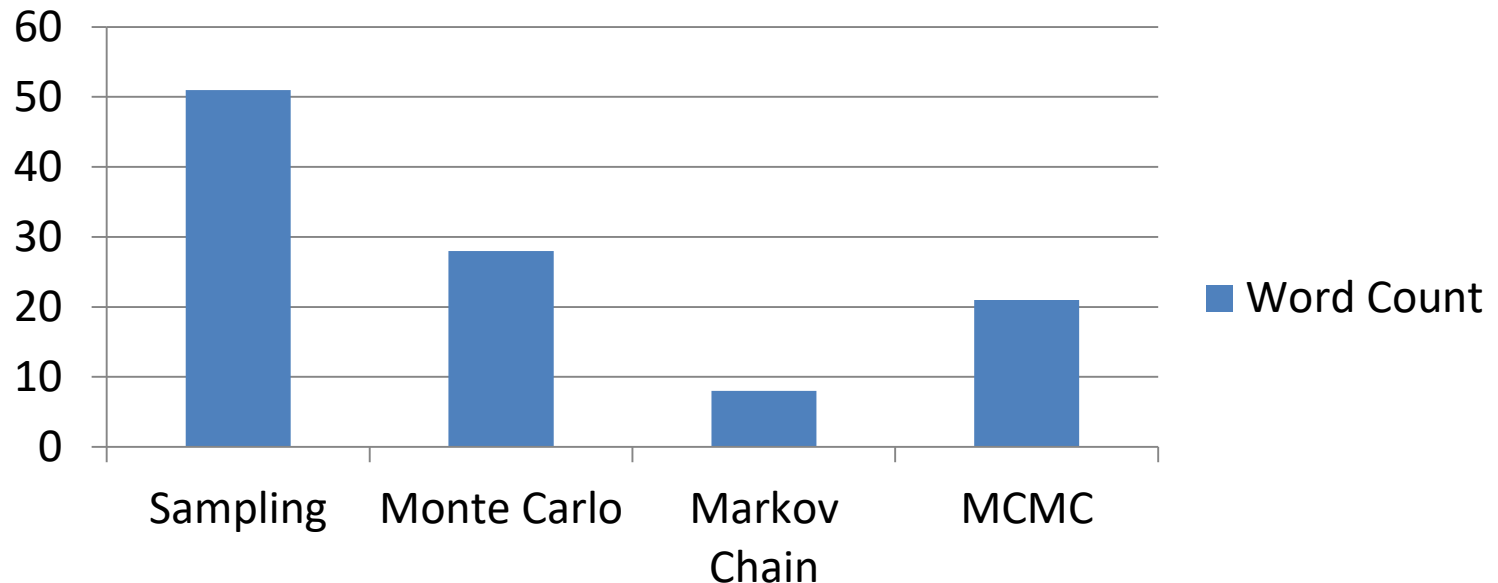
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Motivation

Some statistics on the previous reading assignments...



→ We should learn something about these topics

Talk Outline

- Sampling Problem
- Monte Carlo Idea
- Markov Chains
- Markov Chain Monte Carlo Methods
- Latin Hypercube Sampling
- Comparison of the presented Methods

The Sampling Problem

- D : Distribution over finite set X
- Given: **Black-box** access to the probability distribution function $p(x)$
- Goal: Output a sample of elements drawn according to $p(x)$

What can we do with the samples ?

- Analyze intractable posterior distribution
 - Remember the Roughier Paper ?

$$\Pr(y_f | z = \tilde{z}) = \int \Pr(y_f | x^*, z = \tilde{z}) \Pr(x^* | z = \tilde{z}) dx^*$$

- Integration

$$I = \int_{\theta} g(\theta) p(\theta) d\theta$$

$$I_M = \frac{1}{M} \sum_i^M g(\theta^{(i)})$$

Monte Carlo Methods



- First experiments with Monte Carlo like methods by Enrico Fermi in the 1930s
- Modern version developed in the late 1940s by Stanislaw Ulam, while working on nuclear weapons
- Further work by John von Neuman

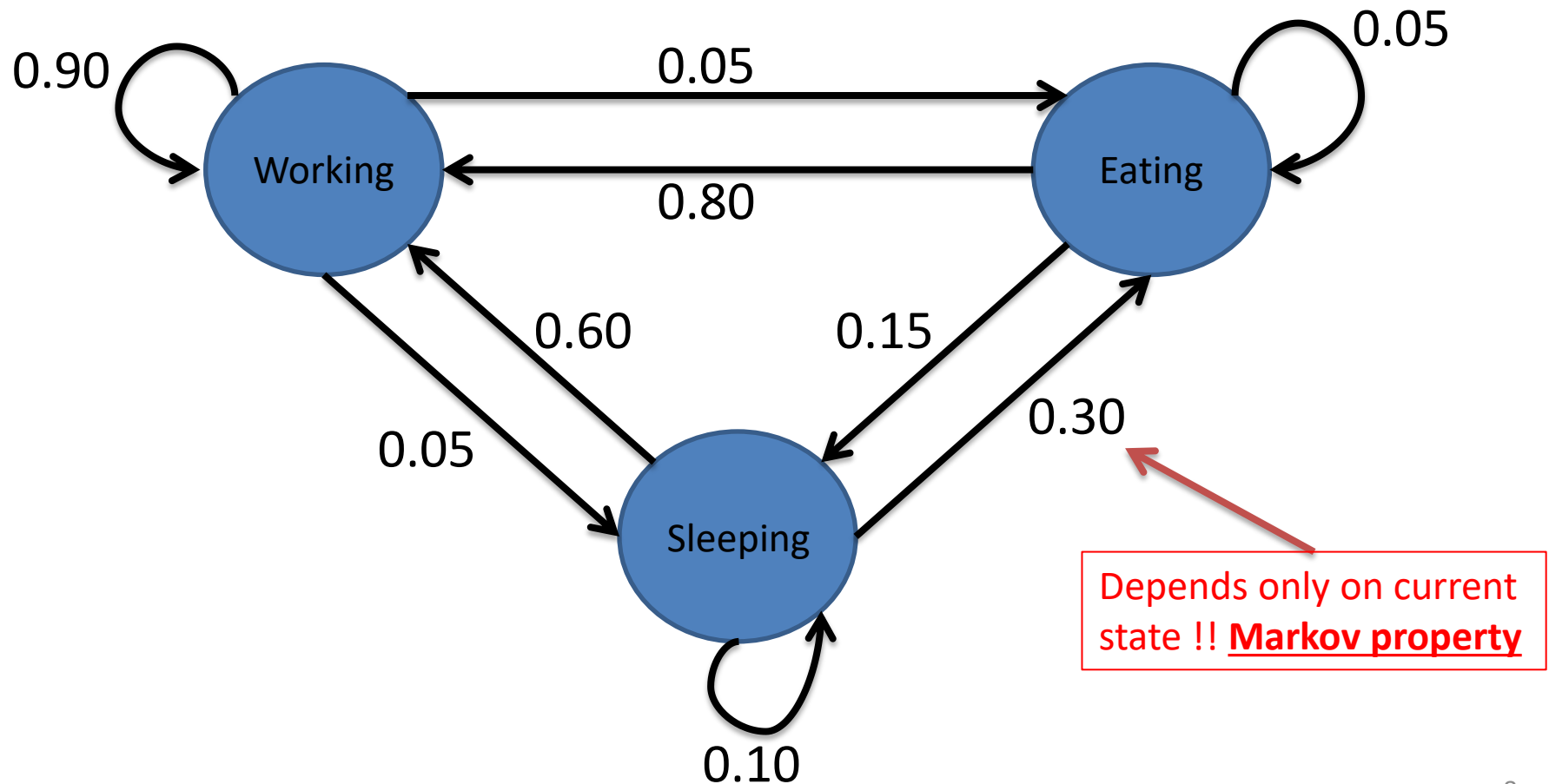
Monte Carlo Methods

General Monte Carlo pattern:

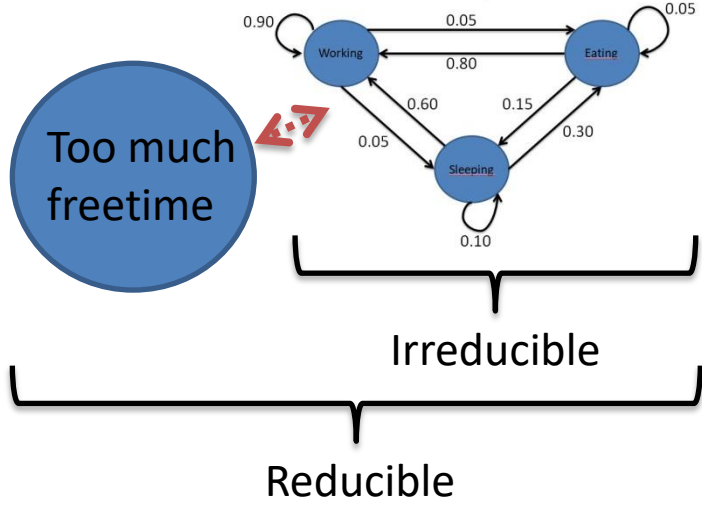
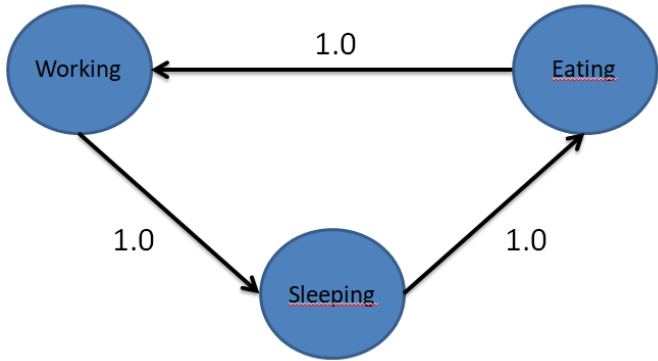
- Define a domain of possible inputs
- Generate inputs **randomly** from a probability distribution
- Perform a **deterministic computation** on the inputs
- Aggregate the results

Markov Chains

Markov state diagram of a student



Markov Chains

Irreducible	<p>Possible to go from every state to every other state in one or more steps.</p>	 <p>Too much freetime</p> <p>Irreducible</p> <p>Reducible</p>
Periodic	<p>One can return to a state only at regular intervals</p>	 <p>Periodic</p>

Markov Chains

- Irreducible and aperiodic Markov Chains have unique stationary distributions!

-> find this stationary distribution for the example using the Transition matrix

$$T = \begin{pmatrix} 0.9 & 0.05 & 0.05 \\ 0.6 & 0.10 & 0.30 \\ 0.8 & 0.15 & 0.05 \end{pmatrix}$$

$P(\text{Eating}|\text{Working})$
↓

Markov Chains

-> **State probabilities and Transition matrix**

$$p = (p_{\text{work}}, p_{\text{eat}}, p_{\text{sleep}})$$

$$p^{j+1} = p^j T$$

-> **Stationary distribution :**

$$p = pT \Leftrightarrow p^T T^T = p^T$$

Find eigenvector
with eigenvalue 1

$$p = (p_{\text{work}}, p_{\text{eat}}, p_{\text{sleep}}) \approx (0.88, 0.06, 0.06)$$

Markov Chain Monte Carlo Methods (MCMC)

- MCMC sampling sets up an **irreducible, aperiodic Markov Chain**.
- The stationary distribution equals posterior of interest **(for infinite chain lengths)**

→ Approximation

MCMC – Metropolis Hastings Algorithm

- Goal: Simulate $g(\theta|y)$ posterior

- Algorithm:

Markov property

1. begin with initial value θ^0

2. candidate value θ^* from proposal dens. $p(\theta^*|\theta^{t-1})$

3. Compute ratio R
$$R = \frac{g(\theta^*)}{g(\theta^{t-1})}$$

4. Compute acceptance prob. $P = \min(R, 1)$

5. Accept θ^* with probability P

→ Sequence of θ will be distributed according to $g(\theta|y)$

MCMC – Metropolis Hastings Algorithm

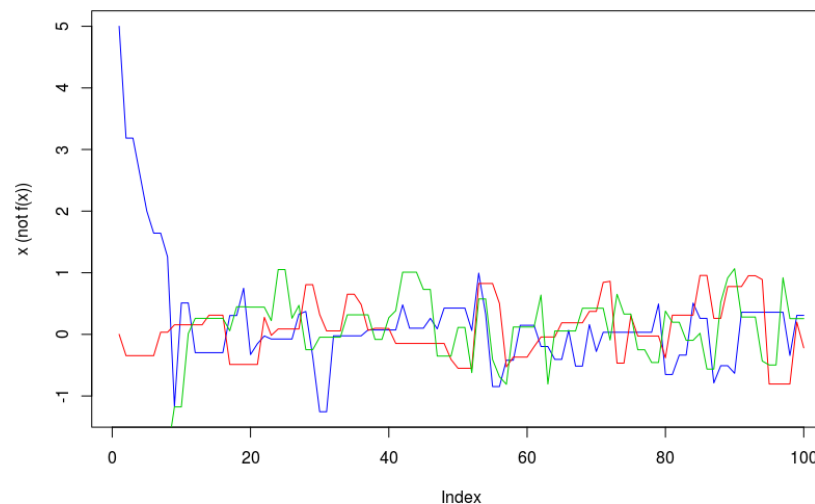
- -> Example (Sampling from unimodal and bimodal Gaussian)

MCMC Problems

- Values in chain must be **representative of the target distribution**
 - Explore full range without getting stuck
- Chain should be of **sufficient size**
 - Estimates accurate and stable
- For Metropolis-Hastings:
 - choose proposal distribution, Ratio, ...
 - Rejection rate

MCMC Representativeness

- Methods to check convergence
 - Visual examination of chain trajectory (trace plot)
 - Burn-in period
 - Gelman-Rubin Test



MCMC Representativeness

- Methods to check accuracy
 - Calculate autocorrelation function (ACF)
 - Calculate effective sample size (ESS)

$$\text{ESS} = N / \left(1 + 2 \sum_{k=1}^{\infty} \text{ACF}(k) \right)$$

- Calculate standard error of sample mean (MCSE)

$$\text{MCSE} = \text{SD} / \sqrt{\text{ESS}}$$

MCMC Representativeness

- -> go back to example (Sampling from unimodal and bimodal Gaussian)

Metropolis Hastings Algorithm – Why it works

- Suppose we have a limited number of states θ_i
- Find Transition Matrix (\rightarrow Markov Chain example)

$$T = \begin{bmatrix} \ddots & p(\theta-2 \rightarrow \theta-1) & \textcircled{0} & 0 & 0 \\ \ddots & p(\theta-1 \rightarrow \theta-1) & p(\theta-1 \rightarrow \theta) & 0 & 0 \\ 0 & p(\theta \rightarrow \theta-1) & p(\theta \rightarrow \theta) & p(\theta \rightarrow \theta+1) & 0 \\ 0 & 0 & p(\theta+1 \rightarrow \theta) & p(\theta+1 \rightarrow \theta+1) & \ddots \\ 0 & 0 & \textcircled{0} & p(\theta+2 \rightarrow \theta+1) & \ddots \end{bmatrix}$$

$p(\theta-2 \rightarrow \theta)$ (arrow to top-right $\textcircled{0}$)
 $p(\theta+2 \rightarrow \theta)$ (arrow to bottom-right $\textcircled{0}$)

Metropolis Hastings Algorithm – Why it works

- What are the elements in T ?

$$p(\theta - 1 \rightarrow \theta) = \text{proposal}(\theta) \times \text{acceptance}(\theta)$$

$$p(\theta - 1 \rightarrow \theta) = \underset{\downarrow}{\text{sal}(\theta)} \times \min \left(\frac{p(\overset{\downarrow}{\theta} - 1)}{p(\theta)}, 1 \right)$$

$$p(\theta + i \rightarrow \theta) = \text{sal}(\theta) \times \min \left(\frac{p(\theta + i)}{p(\theta)}, 1 \right)$$

$$p(\theta \rightarrow \theta) = 1 - \sum_{i \neq 0} p(\theta + i \rightarrow \theta)$$

- We know where we want to go...

$$\boxed{w = wT} \quad (\text{distribution invariant under } T)$$

$$w = 1/Z * (\cdots, p(\theta - 2), p(\theta - 1), p(\theta), p(\theta + 1), \cdots)$$

Metropolis Hastings Algorithm – Why it works

- Evaluate wT for the $p(\theta)$ element (min case-by-case)

→ always reduces to:

$$p(\theta) = p(\theta) * \underbrace{\sum_i \text{sal}(\theta_i)}_{=1}$$

Comment on invariance

Always holds, if T satisfies detailed balance:

$$p(x)T(x \rightarrow x') = p(x')T(x' \rightarrow x)$$

But: detailed balance \Rightarrow invariance \Rightarrow stationary distribution
 \nRightarrow

MCMC – Gibbs Sampling

- special case of the Metropolis–Hastings algorithm
- Parameter vector of interest $\theta = (\theta_1, \dots, \theta_p)$
- **Given:** Set of conditional distributions

$$p(\theta_1 | \theta_2, \theta_3, \dots, \theta_p, \text{data})$$

$$p(\theta_2 | \theta_1, \theta_3, \dots, \theta_p, \text{data})$$

$$p(\theta_p | \theta_1, \dots, \theta_{p-1}, \text{data})$$


- **Goal:** Sample from $p(\theta | \text{data})$

MCMC – Gibbs Sampling

- Algorithm

1. Start with initial vector $\theta^0 = (\theta_1^0, \dots, \theta_p^0)$

2. Calculate new parameters θ_i (Gibbs Cycle)


$$\theta_1^t = p(\theta_1^t | \theta_2^{t-1}, \theta_3^{t-1}, \dots, \theta_p^{t-1})$$

$$\theta_2^t = p(\theta_2^t | \theta_1^t, \theta_3^{t-1}, \dots, \theta_p^{t-1})$$

$$\theta_p^t = p(\theta_p^t | \theta_1^t, \dots, \theta_{p-1}^t)$$

MCMC – Gibbs Sampling

- -> Example (Sample from bivariate gaussian)

$$\begin{pmatrix} \theta_1 \\ \theta_2 \end{pmatrix} = N \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right]$$

$$p(\theta_1 | \theta_2, y) = N(\rho * \theta_2, 1 - \rho^2) = p\theta_2 + \sqrt{1 - p^2} N(0, 1)$$

$$p(\theta_2 | \theta_1, y) = N(\rho * \theta_1, 1 - \rho^2) = p\theta_1 + \sqrt{1 - p^2} N(0, 1)$$

MCMC – Slice Sampling

- Algorithm:

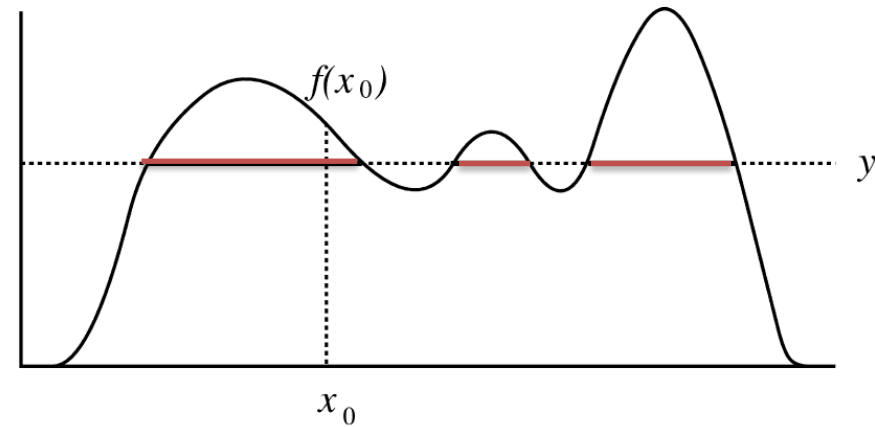
1. Choose starting point x_0

2. Draw a real value y uniformly from $(0, f(x_i))$

3. Define horizontal “**slice**” $S = \{x : y < f(x)\}$

4. Find an interval $I = (L, R)$ around x_i

5. Draw the new point x_{i+1} from the **part of the slice** within this interval

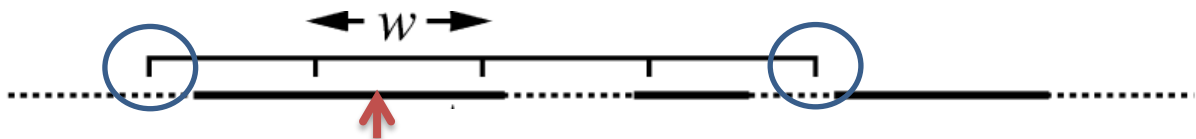


MCMC – Slice Sampling

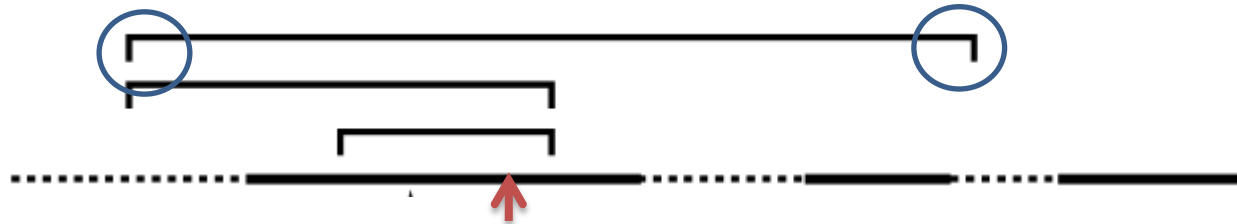
Details to point 4: Find an interval around x_i

- set I to the smallest interval that contains all of $S \rightarrow$ not feasible
- randomly pick an initial interval of size and expand it until both ends lie outside the slice

– "Steeping out"



– "Doubling"



MCMC – Slice Sampling

Details to point 5: Draw new point

Once Interval I has been found:

- Repeatedly sample uniformly from I until a point within S is found
 - **Speed up:** shrink I each time a point is drawn that is not in S

MCMC-Slice Sampling

Why Slice Sampling is a good Method compared to...

- Metropolis Hastings:
 - Less tuning parameters
 - Step sizes can be bigger
- Gibbs Sampling
 - No full conditional distributions for parameters needed

MCMC – Some improvements

- Run parallel chains
- Change parametrization of problem
- Thinning: use only every k th state to reduce autocorrelation
- ...

Latin Hypercube Sampling (LHS)

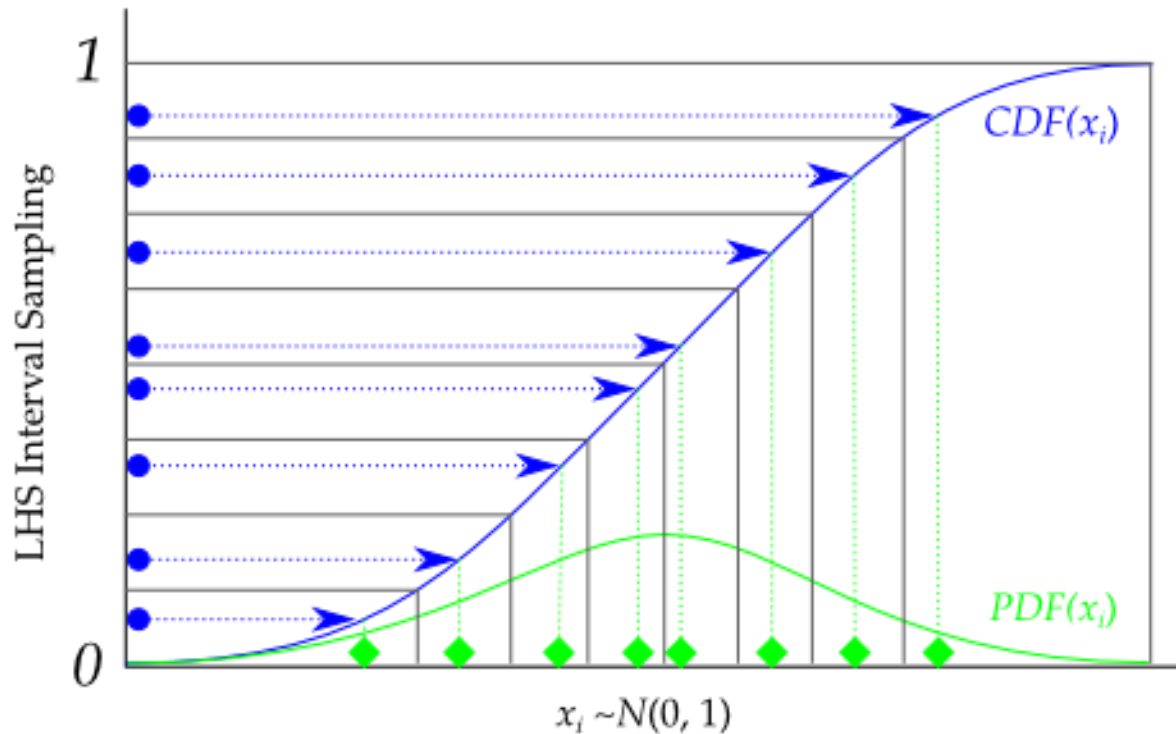
- **Latin square** is an $n \times n$ array filled with n different symbols, each occurring exactly once in each row and exactly once in each column

A	B	C
C	A	B
B	C	A

Latin Hypercube Sampling (LHS)

- How it works:
 - Represent each variable as is Cumulative Distribution Function (CDF)
 - Partition CDF into N regions (\leftrightarrow Latin Hypercube)
 - Take sample from each region
- Full range of the distribution is sampled

Latin Hypercube Sampling (LHS)



https://pythonhosted.org/pyDOE/_images/lhs_custom_distribution.png

Latin Hypercube Sampling (LHS)

- -> Example (Sampling from bivariate Gaussian)

Latin Hypercube vs. Markov Chain Monte Carlo

Why Latin Hypercube Sampling is a good Method...

- produces a clear depiction of each input distribution
- Avoid sampling artefacts

Literature

- Albert, Jim. *Bayesian computation with R*. Springer Science & Business Media, 2009.
- Hoff, Peter D. *A first course in Bayesian statistical methods*. Vol. 580. New York: Springer, 2009.
- Kruschke, John. *Doing Bayesian data analysis: A tutorial with R, JAGS, and Stan*. Academic Press, 2014.
- [https://icme.hpc.msstate.edu/mediawiki/index.php/Latin_Hypercube_Sampling_\(LHS\)](https://icme.hpc.msstate.edu/mediawiki/index.php/Latin_Hypercube_Sampling_(LHS)) (last checked 2019-11-19)
- <https://jeremykun.com/2015/04/06/markov-chain-monte-carlo-without-all-the-bullshit/> (last checked 2019-11-19)
- <http://www.csri.utoronto.ca/pub/radford/slice-aos.pdf> (last checked 2019-11-23)