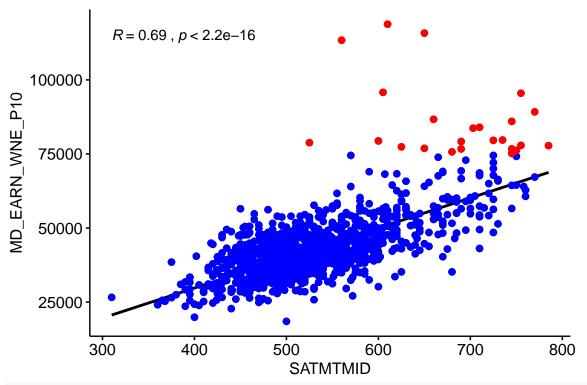
# hw5\_q2

## a) median SAT scores versus the median earnings after 10 years

Here we scatterplot the approximately linear relationship ( $r^2 > 0.6$ ) between these variables. First we removed the rows with missing values. Based on the line of best fit, outliers appeared to have earnings greater than 75000 (red points). It seems the outliers belong to graduates of medical schools or prestigious schools who can earn more than average (printed below). I removed these outliers.



filt[filt\$MD\_EARN\_WNE\_P10 > 75000,"INSTNM"]

- ## [1] California Institute of Technology
- ## [2] Colorado School of Mines
- ## [3] Georgetown University
- ## [4] Rose-Hulman Institute of Technology
- ## [5] Maine Maritime Academy
- ## [6] Babson College
- ## [7] Bentley University
- ## [8] Harvard University
- ## [9] MCPHS University

```
## [10] Massachusetts Institute of Technology
## [11] Kettering University
## [12] St Louis College of Pharmacy
## [13] Princeton University
## [14] Stevens Institute of Technology
## [15] Albany College of Pharmacy and Health Sciences
## [16] Columbia University in the City of New York
## [17] Rensselaer Polytechnic Institute
## [18] Duke University
## [19] Carnegie Mellon University
## [20] Lehigh University
## [21] University of Pennsylvania
## [22] University of the Sciences
## [23] Stanford University
## [24] DigiPen Institute of Technology
## 7535 Levels: A & W Healthcare Educators ...
```

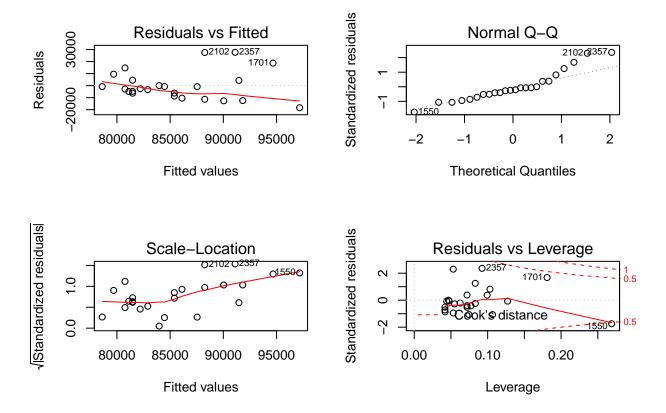
## b) median earnings and SAT math scores

Here we fit an ordinary linear model of median earnings vs median SAT math scores. The diagnostic plots suggest the fit is not very linear:

- The residuals vs fitted plot shows residuals have a downward trend rather than being equally spread relative to the fitted values.
- the QQ plot has many points which are not well aligned on the y=x line indicating some residuals are not normally distributed.
- the plot of residuals vs. leverage shows a couple of potentially problematic outliers with high residuals and/or leverage, lying close to the Cook's distance curves.

Overall the model relationship is not very linear (Adjusted R-squared less than 0.2) and the diagnostic plots show some problems with the fit.

```
##
## Call:
## lm(formula = MD_EARN_WNE_P10 ~ SATMTMID, data = filt_data)
## Residuals:
##
     Min
              1Q Median
                            3Q
                                  Max
  -18374 -6878 -2672
                          4384
                                27693
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 134643.81
                           25133.80
                                      5.357 2.23e-05 ***
## SATMTMID
                  -71.37
                              36.45 -1.958
                                               0.063 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 12320 on 22 degrees of freedom
## Multiple R-squared: 0.1484, Adjusted R-squared: 0.1097
## F-statistic: 3.835 on 1 and 22 DF, p-value: 0.06299
```

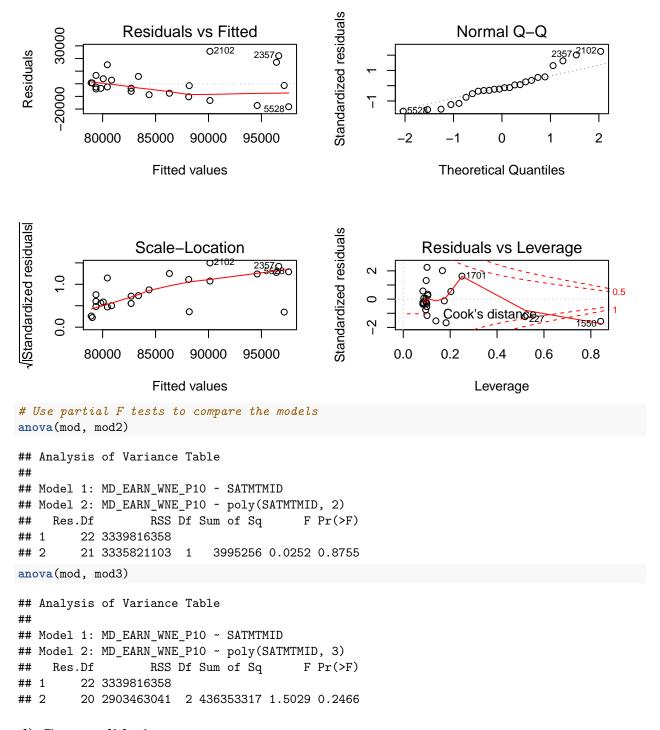


### c) Nonlinear fits

We fit a second-order and third-order polynomial here and find that neither of these provide a statistically significant improvement over the original first order linear model by partial F test (no significant terms at p < 0.05). The diagnostic plots also do not show improvements from previous ones. For example the third order polynomial diagnostic plot of residuals versus leverage (shown below) even see the appearance of an outlier that has moved further outside of the Cook's distance lines. This suggests the nonlinear fits are comparable to the linear fit.

```
filt_data <- data[data$MD_EARN_WNE_P10 >75000,] %>%
  select(SATMTMID, MD_EARN_WNE_P10, INSTNM) %>% na.omit()
mod2 <- lm(MD_EARN_WNE_P10 ~ poly(SATMTMID, 2), data = filt_data)</pre>
summary(mod2)
##
## Call:
## lm(formula = MD_EARN_WNE_P10 ~ poly(SATMTMID, 2), data = filt_data)
##
##
  Residuals:
##
              1Q Median
                             3Q
                                   Max
   -17180
           -7204
                  -2562
                           4307
                                 27509
##
##
##
  Coefficients:
##
                       Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                          85671
                                       2573
                                             33.300
                                                       <2e-16 ***
## poly(SATMTMID, 2)1
                         -24129
                                      12604
                                             -1.914
                                                      0.0693 .
## poly(SATMTMID, 2)2
                          -1999
                                      12604
                                             -0.159
                                                      0.8755
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
```

```
##
## Residual standard error: 12600 on 21 degrees of freedom
## Multiple R-squared: 0.1495, Adjusted R-squared: 0.06846
## F-statistic: 1.845 on 2 and 21 DF, p-value: 0.1827
mod3 <- lm(MD_EARN_WNE_P10 ~ poly(SATMTMID, 3), data = filt_data)</pre>
summary(mod3)
##
## Call:
## lm(formula = MD_EARN_WNE_P10 ~ poly(SATMTMID, 3), data = filt_data)
## Residuals:
##
     Min
             1Q Median
                           3Q
                                 Max
## -18172 -6372 -2014
                        4412 25665
##
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                        85671
                                    2459 34.833 <2e-16 ***
## poly(SATMTMID, 3)1
                       -24129
                                   12049 -2.003
                                                   0.0590 .
## poly(SATMTMID, 3)2
                        -1999
                                   12049 -0.166
                                                   0.8699
                        20793
                                          1.726
                                                   0.0998 .
## poly(SATMTMID, 3)3
                                   12049
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 12050 on 20 degrees of freedom
## Multiple R-squared: 0.2597, Adjusted R-squared: 0.1487
## F-statistic: 2.339 on 3 and 20 DF, p-value: 0.1042
# plot residuals and diagnostics
par(mfrow = c(2, 2))
plot(mod3)
```



#### d) Cross validation

```
library(modelr)
cv <- crossv_kfold(filt_data)

model1 <- map(cv$train, ~lm(MD_EARN_WNE_P10 ~ SATMTMID, data =.))
model2 <- map(cv$train, ~lm(MD_EARN_WNE_P10 ~ poly(SATMTMID, 2), data =.))
model3 <- map(cv$train, ~lm(MD_EARN_WNE_P10 ~ poly(SATMTMID, 3), data =.))</pre>
```

```
# Use cross-validation to estimate the squared-error loss of each of your models.
errs1 <- map2_dbl(model1, cv$test, mse)
errs2 <- map2_dbl(model2, cv$test, mse)
errs3 <- map2_dbl(model3, cv$test, mse)

# print mean mse
mean(as.numeric(errs1))

## [1] 173386251
mean(as.numeric(errs2))

## [1] 265897174
mean(as.numeric(errs3))</pre>
```

## [1] 266892309

Whether or not the F tests suggested the polynomials help, you can also compare the pre- dictive performance of each model. Sometimes variables that are not statistically signifi- cant can improve predictive performance. Use cross-validation to estimate the squared- error loss of each of your models. (Fit a new model to each training set.) Compare the results to what you got using F tests. Why could the results differ?

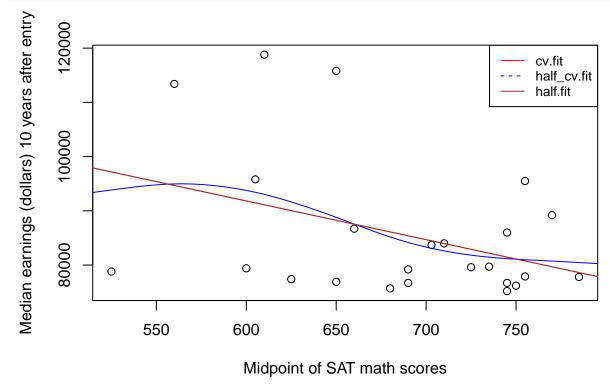
(For K-fold cross-validation, the modelr package has a crossv\_kfold function that auto- matically divides your data up into folds, and gives you lists of the training and test sets.

We checked the cross validation RSS and found that the predictive performance of the original linear model is better than the performance of the order 2 polynomial, which is in turn better than performance of the order 3 polynomial. This is different from our conclusion from the F tests.

# e) Fit a smoothing spline

For the same model as above we compared the following three spline models:

- cv.fit: spar picked by automatic cross-validation
- half cv.fit: spar set to be half as big as R picked
- half.fit: spar set to be halfway between cv.fit and 1



It's clear that whereas half fit and cv.fit are very similar, half cv.fit is a less linear model with higher variance.

#### (f) Use cross-validation to estimate the error of the three splines.

Which is worse, too much bias or too much variance? (Which fit corresponds to high bias, and which corresponds to high variance?)

It seems that high bias wins over high variance in this situation, since the average mse of the half\_cv spline (which has the smallest spar and higher variance) is the largest.

```
spar = spar_fit)
  half_cv.fit <- smooth.spline(trainData$SATMTMID,
                              trainData$MD_EARN_WNE_P10,
                               spar = spar_fit/2.0)
  half.fit <- smooth.spline(trainData$SATMTMID,
                           trainData$MD_EARN_WNE_P10,
                           spar = (spar_fit-1.0)/2.0 + 1.0)
  # extract model performance on test fold
  y <- testData$MD_EARN_WNE_P10
  yhat1 <- predict(cv.fit, testData$SATMTMID)$y</pre>
  yhat2 <- predict(half_cv.fit,testData$SATMTMID)$y</pre>
  yhat3 <- predict(half.fit, testData$SATMTMID)$y</pre>
  # get the mse
  mse1[i] <- mean((y-yhat1)^2)</pre>
  mse2[i] <- mean((y-yhat2)^2)</pre>
  mse3[i] <- mean((y-yhat3)^2)</pre>
# Compare the average errors of the three.
mean(as.numeric(mse1))
## [1] 177254822
mean(as.numeric(mse2))
## [1] 210967799
mean(as.numeric(mse3))
## [1] 177272169
```