# ECS256 - Homework III

Olga Prilepova, Christopher Patton, Alexander Rumbaugh, John Chen, Thomas Provan

February 17, 2014

#### Problem 1.a

First, we'll derive  $\pi_i$ . The definition of the tree searching markov model leads to the following set of balance equations for the long-run state probabilities:

$$\pi_i = \pi_{i-1}q_{i-1} = \pi_0 \prod_{j=0}^{i-1} q_j$$
 for  $i \ge 1$ , and 
$$\pi_0 = \sum_{j=0}^{\infty} \pi_i (1 - q_i)$$
 for  $i = 0$ .

This definition for  $\pi_0$  is a bit unwelldy. We can also think of this quantity as one over the expected recurrence time, as in eq. (10.63) in the book:

$$\pi_0 = \frac{1}{E(T_{0,0})}$$

$$E(T_{0,0}) = 1 + \sum_{k \neq 0} p_{0,k} E(T_{k,0})$$

$$= 1 + p_{0,1} E(T_{1,0})$$

$$= 1 + p_{0,1} (1 + \sum_{k \neq 0} p_{1,k} E(T_{k,0}))$$

$$= 1 + p_{0,1} (1 + p_{1,2} E(T_{2,0}))$$

$$= 1 + p_{0,1} (1 + p_{1,2} (1 + \sum_{k \neq 0} p_{2,k} E(T_{k,0})))$$

$$= 1 + p_{0,1} (1 + p_{1,2} (1 + p_{2,3} E(T_{3,0})))$$

and so on. This unravels into a familiar closed form:

$$E(T_{0,0}) = 1 + q_0(1 + q_1(1 + q_2(1 + \dots)))$$

$$= 1 + q_0 + q_0q_1 + q_0q_1q_2 + \dots$$

$$= 1 + \sum_{i=1}^{\infty} \left[ \prod_{j=0}^{i-1} q_j \right]$$

If the model is positive recurrent, then there exists some value R such that

$$R = \sum_{i=1}^{\infty} \left[ \prod_{j=0}^{i-1} q_j \right] < \infty.$$

Thus,

$$\pi_i = \frac{\prod_{j=0}^{i-1} q_j}{1+R}$$
 for  $i \ge 0$ .

Next,  $E(T_{i,0})$  follows a similar pattern.

$$E(T_{i,0}) = 1 + \sum_{k \neq 0} p_{i,k} E(T_{k,0})$$

$$= 1 + p_{i,i+1} E(T_{j+1,0})$$

$$= 1 + q_i + q_i q_{i+1} + q_i q_{i+1} q_{i+2} + \dots$$

$$= 1 + \sum_{j=i}^{\infty} \left[ \prod_{k=i}^{j} q_k \right].$$

#### Problem 1.b

If  $q_i = 0.5$  for all i, then R is a geometric series that indeed converges.

$$\pi_2 = \frac{0.5 \cdot 0.5}{1 + \sum_{i=1}^{\infty} 0.5^{i-1}} = \frac{0.25}{1+2} \approx 0.083.$$

$$E(T_{2,0}) = 1 + \sum_{j=2}^{\infty} 0.5^{j-2} = 1 + \sum_{j=1}^{\infty} 0.5^{j-1} = 1 + 2 = 3.$$

#### Problem 1.c

The rate of backtracking, in terms of the stationary probabilities  $\pi_i$ , is simply

$$\sum_{i=1}^{\infty} \pi_i (1 - q_i).$$

#### Problem 3

Given a hazard function, h(t), the density function, f(t), can be found as follows:

$$f(t) = h(t) \cdot e^{-\int_0^t h(s) \, ds}$$

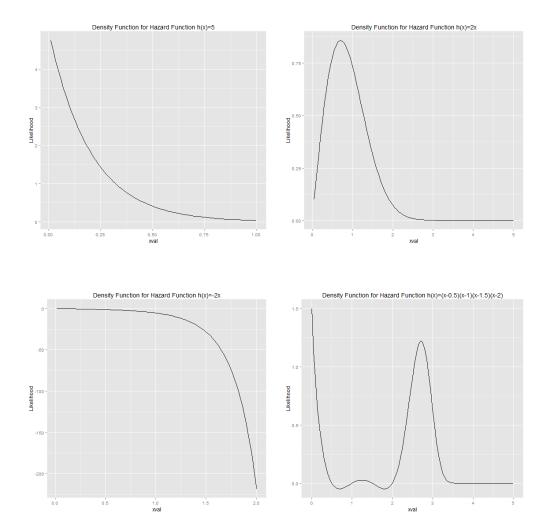
We looked at the following hazard functions to explore what their density would look like:

$$h(t) = 5$$

$$h(t) = 2t$$

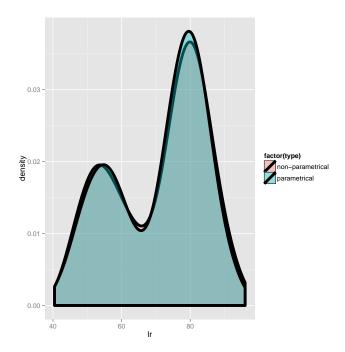
$$h(t) = -2t$$

$$h(t) = (t - 0.5)(t - 1)(t - 1.5)(t - 2)$$



# Problem 4

## 4.a-b



### Appendix

```
Problem 4
```

```
1 #install.packages("ggplot2")
2 #install.packages("mixtools")
3 library (mixtools)
4 library (ggplot2)
7 p<−ggplot (data.frame(faithful))
  p+geom_density(aes(x=faithful$waiting))
_{11} simulateFromDist <- function(n,p1,m1,s1,m2,s2){
           k1 \leftarrow p1*n \#proportion of type 1
           k2 \leftarrow n-k1+1 \#proportion \ of \ type \ 2
           x1 \leftarrow rnorm(k1, mean=m1, sd=s1)
15
           x2 \leftarrow rnorm(k2, mean=m2, sd=s2)
           \mathbf{c}(\mathtt{x1},\mathtt{x2}) \ \textit{\#order} \ \textit{of} \ \textit{events} \ \textit{doesn't} \ \textit{matter} \ \textit{for} \ \textit{histogram} \, , \ \textit{so} \ \textit{simply} \ \textit{concatenate}
16
17
18
19 ###from mixtools simulation
20 mixout <-normalmixEM (faithful $waiting, lambda = 0.5, mu=c (55,80), sigma = 10,k=2)
  str (mixout)
22 # $ lambda
                   : num [1:2] 0.361 0.639
_{23} \# \$ mu
                  : num [1:2] 54.6 80.1
24 # $ sigma
                  : num [1:2] 5.87 5.87
25
27 # Is it necessary to simulate this? Can we plot the function directly?
28 sim_waiting<-simulateFromDist(length(faithful$waiting), 0.361, 54.6, 5.87, 80.1, 5.87)
30 data <- rbind( data.frame(type="non-parametrical", lr=faithful$waiting), data.frame(type="parametrical")
```