# Assignment 2 (ML for TS) - MVA 2022/2023

Clémence Grislain clemence.grislain@eleves.enpc.fr Roman Plaud roman.plaud@eleves.enpc.fr

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### 1 Introduction

**Objective.** The goal is to better understand the properties of AR and MA processes, and do signal denoising with sparse coding.

### Warning and advice.

- Use code from the tutorials as well as from other sources. Do not code yourself well-known procedures (e.g. cross validation or k-means), use an existing implementation.
- The associated notebook contains some hints and several helper functions.
- Be concise. Answers are not expected to be longer than a few sentences (omitting calculations).

#### Instructions.

- Fill in your names and emails at the top of the document.
- Hand in your report (one per pair of students) by Monday 27<sup>th</sup> February 11:59 PM.
- Rename your report and notebook as follows:
   FirstnameLastname1\_FirstnameLastname1.pdf and
   FirstnameLastname2\_FirstnameLastname2.ipynb.
   For instance, LaurentOudre\_CharlesTruong.pdf.
- Upload your report (PDF file) and notebook (IPYNB file) using this link: .

# 2 General questions

A time series  $\{y_t\}_t$  is a single realisation of a random process  $\{Y_t\}_t$  defined on the probability space  $(\Omega, \mathcal{F}, P)$ , i.e.  $y_t = Y_t(w)$  for a given  $w \in \Omega$ . In classical statistics, several independent realisations are often needed to obtain a "good" estimate (meaning consistent) of the parameters of the process. However, thanks to a stationarity hypothesis and a "short-memory" hypothesis, it is still possible to make "good" estimates. The following question illustrates this fact.

### **Question 1**

An estimator  $\hat{\theta}_n$  is consistent if it converges in probability when the number n of samples grows to  $\infty$  to the true value  $\theta \in \mathbb{R}$  of a parameter, i.e.  $\hat{\theta}_n \stackrel{\mathcal{D}}{\longrightarrow} \theta$ .

- Recall the rate of convergence of the sample mean for i.i.d. random variables with finite variance.
- Let  $\{Y_t\}_{t\geq 1}$  a wide-sense stationary process such that  $\sum_k |\gamma(k)| < +\infty$ . Show that the sample mean  $\bar{Y}_n = (Y_1 + \cdots + Y_n)/n$  is consistent and enjoys the same rate of convergence as the i.i.d. case. (Hint: bound  $\mathbb{E}[(\bar{Y}_n \mu)^2]$  with the  $\gamma(k)$  and recall that convergence in  $L_2$  implies convergence in probability.)

### **Answer 1**

• Let's  $\{X_i\}_{\mathbb{N}}$  be iid random variable with finite variance  $\mu$  and  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^{\infty} X_i$ .

From the Central limit theorem, we have that

$$\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{n \to \infty} \mathcal{N}(0, \sigma^2)$$

Thus, the rate converge of  $\bar{X}_n$  is  $\sqrt{n}$ .

•  $\{Y_t\}_{\mathbb{N}}$  is wide sense stationnary:

$$- \forall t \quad \mathbb{E}(Y_t) = \mu = \lim_{t \to \infty} \frac{1}{t} \sum_{i=1}^t Y_i$$

$$- \forall (k, t) \quad \mathbb{E}(Y_t Y_{t+k}) = \gamma_Y(k) = \lim_{t \to \infty} \frac{1}{t} \sum_{i=1}^t Y_i Y_{i+k}$$

We have:

$$\mathbb{E}((\bar{Y}_t - \mu)^2) = \mathbb{E}((\frac{1}{t} \sum_{i=1}^t Y_i - \mu)^2)$$

$$= \mathbb{E}((\frac{1}{t} \sum_{i=1}^t Y_i)^2 + \mu^2 - 2\frac{1}{t} \sum_{i=1}^t Y_i \mu)$$

$$= \mathbb{E}((\frac{1}{t} \sum_{i=1}^t Y_i)^2) + \mu^2 - 2\frac{1}{t} \sum_{i=1}^t \mathbb{E}(Y_i) \mu$$

$$= \mathbb{E}((\frac{1}{t} \sum_{i=1}^t Y_i)^2) + \mu^2 - 2\frac{1}{t} \sum_{i=1}^t \mu^2$$

$$= \frac{1}{t^2} (\sum_{i=1}^t \mathbb{E}(Y_i^2) + 2\sum_{i=1}^t \sum_{j=1}^t \mathbb{E}(Y_i Y_j)) + \mu^2 - 2\mu^2$$

$$= \frac{1}{t^2} (t \gamma_Y(0) + 2\sum_{i=1}^t \sum_{j=i+1}^t \gamma_Y(j-i)) - \mu^2$$

with k = j - i

$$= \frac{\gamma_Y(0)}{t} + \frac{2}{t^2} \sum_{k=1}^t \sum_{j=k+1}^{t+k} \gamma_Y(k) - \mu^2$$

$$= \frac{\gamma_Y(0)}{t} + \frac{2}{t} \sum_{k=1}^t \gamma_Y(k) - \mu^2$$

$$= \leq \frac{\gamma_Y(0)}{t} + \frac{2}{t} \sum_{k=1}^t |\gamma_Y(k)|$$

$$\leq \frac{\gamma_Y(0)}{t} + \frac{2}{t} \sum_{k=1}^\infty |\gamma_Y(k)| = \frac{1}{t} (\gamma_Y(0) + 2l)$$

Hence we have,  $\mathbb{E}((\bar{Y}_t - \mu)^2) \xrightarrow{t \to \infty} 0$  in a rate convergence  $\frac{1}{t}$ 

 $\implies \bar{Y}_t \xrightarrow{\mathcal{L}_2} \mu$  in a rate convergence  $\frac{1}{\sqrt{t}}$ 

 $\Longrightarrow \bar{Y}_t \xrightarrow{\mathbb{P}} \mu$  in a rate convergence  $\frac{1}{\sqrt{t}}$ ,  $\bar{Y}_t$  is a consistent estimator and enjoys the same rate of convergence as the i.i.d. case.

## 3 AR and MA processes

**Question 2** *Infinite order moving average MA*( $\infty$ )

Let  $\{Y_t\}_{t\geq 0}$  be a random process defined by

$$Y_t = \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \dots = \sum_{k=0}^{\infty} \psi_k \varepsilon_{t-k}$$
 (1)

where  $(\psi_k)_{k\geq 0} \subset \mathbb{R}$  ( $\psi=1$ ) are square summable, i.e.  $\sum_k \psi_k^2 < \infty$  and  $\{\varepsilon_t\}_t$  is a zero mean white noise of variance  $\sigma_{\varepsilon}^2$ . (Here, the infinite sum of random variables is the limit in  $L_2$  of the partial sums.)

- Derive  $\mathbb{E}(Y_t)$  and  $\mathbb{E}(Y_tY_{t-k})$ . Is this process weakly stationary?
- Show that the power spectrum of  $\{Y_t\}_t$  is  $S(f) = \sigma_{\varepsilon}^2 |\phi(e^{-2\pi i f})|^2$  where  $\phi(z) = \sum_j \psi_j z^j$ . (Assume a sampling frequency of 1 Hz.)

The process  $\{Y_t\}_t$  is a moving average of infinite order. Wold's theorem states that any weakly stationary process can be written as the sum of the deterministic process and a stochastic process which has the form (1).

### **Answer 2**

• 
$$-\mathbb{E}(Y_{t}) = \mathbb{E}(\sum_{i=0}^{\infty} \psi_{i} \varepsilon_{t-i}) = \sum_{i=0}^{\infty} \psi_{i} \mathbb{E}(\varepsilon_{t-i}) = 0$$

$$-\mathbb{E}(Y_{t}Y_{t-k}) = \mathbb{E}\left((\sum_{i=0}^{\infty} \psi_{i} \varepsilon_{t-i}) \times (\sum_{j=0}^{\infty} \psi_{j} \varepsilon_{t-k-j})\right) = \mathbb{E}\left(\sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_{i} \psi_{j} \varepsilon_{t-i} \varepsilon_{t-k-j}\right) = \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_{i} \psi_{j} \mathbb{E}\left(\varepsilon_{t-i} \varepsilon_{t-k-j}\right)$$

$$\text{With } \mathbb{E}\left(\varepsilon_{t-i} \varepsilon_{t-k-j}\right) = \begin{cases} \sigma_{\varepsilon}^{2} & \text{if } i = k+j \\ 0 & \text{otherwise} \end{cases} \implies \mathbb{E}(Y_{t}Y_{t-k}) = \sum_{i=0}^{\infty} \psi_{i} \psi_{i-k} \sigma_{\varepsilon}^{2}$$

Note: The expected value and the infinte sums were inverted thanks to the Fubini theorem as the square-summability condition on the coefficients  $\psi_k$  guarantees that the series converges in the mean square sense.

 $\mathbb{E}(Y_t)$  is independent of t, the auto-covariance  $\mathbb{E}(Y_tY_{t-k})$  only depends k and  $Var(Y_t) = \mathbb{E}(Y_t^2) = \sigma_\epsilon^2 \sum_{i=1}^\infty \psi^2 = \sigma_\epsilon^2 \times const$  is a constant. Thus,  $\{Y_t\}_{\mathbb{N}}$  is weakly stationary.

• With  $f_s = 1$ ,

$$S(f) = \sum_{k=-\infty}^{\infty} \gamma_Y(k) e^{-2i\pi \times f \times k}$$

$$= \sigma_{\epsilon}^2 \sum_{k=-\infty}^{\infty} \sum_{j=0}^{\infty} \psi_j \psi_{j-k} e^{-2i\pi \times f \times k}$$

$$= \sigma_{\epsilon}^2 \sum_{j=0}^{\infty} \psi_j \sum_{k=-\infty}^{\infty} \psi_{j-k} e^{-2i\pi \times f \times k}$$

With l=j-k

$$= \sigma_{\epsilon}^{2} \sum_{j=0}^{\infty} \psi_{j} \sum_{l=-\infty}^{\infty} \psi_{l} e^{-2i\pi \times f \times (j-l)}$$

$$= \sigma_{\epsilon}^{2} \sum_{j=0}^{\infty} \psi_{j} \sum_{l=-\infty}^{\infty} \psi_{l} e^{-2i\pi \times f \times (j-l)}$$

$$= \sigma_{\epsilon}^{2} \sum_{j=0}^{\infty} \psi_{j} e^{-2i\pi \times f \times j} \sum_{l=-\infty}^{\infty} \psi_{l} e^{+2i\pi \times f \times l}$$

With  $\psi_l = 0$  if l < 0

$$= \sigma_{\epsilon}^{2} \sum_{j=0}^{\infty} \psi_{j} e^{-2i\pi \times f \times j} \sum_{l=0}^{\infty} \psi_{l} e^{+2i\pi \times f \times l}$$

$$= \sigma_{\epsilon}^{2} |\sum_{l=0}^{\infty} \psi_{l} e^{-2i\pi \times f \times l}|^{2} = \sigma_{\epsilon}^{2} |\phi(e^{-2i\pi \times f})|^{2} \quad \Box$$

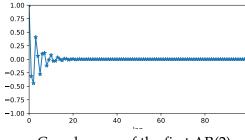
## **Question 3** AR(2) process

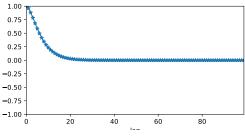
Let  $\{Y_t\}_{t\geq 1}$  be an AR(2) process, i.e.

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t \tag{2}$$

with  $\phi_1, \phi_2 \in \mathbb{R}$ . The associated characteristic polynomial is  $\phi(z) := 1 - \phi_1 z - \phi_2 z^2$ . Assume that  $\phi$  has two distinct roots (possibly complex)  $r_1$  and  $r_2$  such that  $|r_i| > 1$ . Properties on the roots of this polynomial drive the behaviour of this process.

- Express the autocovariance coefficients  $\gamma(\tau)$  using the roots  $r_1$  and  $r_2$ .
- Figure 1 shows the correlograms of two different AR(2) processes. Can you tell which one has complex roots and which one has real roots?
- Express the power spectrum S(f) (assume the sampling frequency is 1 Hz) using  $\phi(\cdot)$ .
- Choose  $\phi_1$  and  $\phi_2$  such that the characteristic polynomial has two complex conjugate roots of norm r = 1.05 and phase  $\theta = 2\pi/6$ . Simulate the process  $\{Y_t\}_t$  (with n = 2000) and display the signal and the periodogram (use a smooth estimator) on Figure 2. What do you observe?





Correlogram of the first AR(2)

Correlogram of the second AR(2)

Figure 1: Two AR(2) processes

### **Answer 3**

• With independence property of  $Y_{\tau}$  and  $\varepsilon_t$ , we can show that

$$\gamma(\tau) = \phi_1 \gamma(\tau - 1) + \phi_2 \gamma(\tau - 2)$$

.  $\gamma$  is then the solution of a second order sequence whose characteriscal polynomial roots are  $\frac{1}{r_1}$  and  $\frac{1}{r_2}$  so that.

- If  $r_1, r_2 \in \mathbb{R}^2$ ,  $\exists (\lambda, \mu) \in \mathbb{R}^2$  such that :

$$\gamma(\tau) = \frac{\lambda}{r_1^{\tau}} + \frac{\mu}{r_2^{\tau}}$$

– If  $r_1 = r_2 \in \mathbb{R}$ ,  $\exists (\lambda, \mu) \in \mathbb{R}^2$  such that :

$$\gamma( au) = rac{\lambda}{r_1^ au} + rac{\mu au}{r_2^ au}$$

- If  $r_1, r_2 \in \mathbb{C}^2$ ,  $r_1 = r \exp^{i\theta}$ ,  $r_1 = r \exp^{-i\theta} \exists (\lambda, \mu) \in \mathbb{R}^2$  such that :

$$\gamma(\tau) = \frac{1}{r^{\tau}} (\lambda \cos(\tau \theta) + \mu \sin(\tau \theta))$$

- Observing the general form of  $\gamma$  thanks to previous question we can say that the oscillating correlogram correspond to the complex roots whereas the non-increasing one correspond to the real roots.
- To express the power spectrum of S(f), we will show that an AR(2) process can be expressed as a MA( $\infty$ ) process with characteristic polynomial  $\frac{1}{\phi(.)}$ .

Let's L be the Lag operator such that  $L^k X_t = X_{t-k}$ . We can note that a MA( $\infty$ ) process  $Y_t = \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \cdots = \sum_{k=0}^{\infty} \psi_k \varepsilon_{t-k}$  can be expressed thanks to the Lag operator as

$$Y_t = \sum_{k=0}^{\infty} \psi_k L^k \varepsilon_t = \phi(L) \varepsilon_t$$

where  $\phi(z) = \sum_j \psi_j z^j$  is the characteristic polynomial, and the power spectrum (with  $f_s = 1$ ) is  $S(f) = \sigma_{\epsilon}^2 |\phi(e^{-2i\pi \times f})|^2$ .

AR(2) process  $Y_t = \phi_1 Y_{t-1} +_2 Y_{t-2}$  can be expressed using the Lag operator  $Y_t = (\phi_1 L + \phi_2 L^2) Y_t + \epsilon_t \implies \phi(L) Y_t = \epsilon_t$ 

$$\implies Y_t = \frac{1}{\phi(L)} \epsilon_t$$

where  $\phi(z) = 1 - \phi_1 z - \phi_2 z^2$ .

The polynomial  $\phi(z)$  can be factorized thanks to its roots  $r_1$  and  $r_2$ :  $\phi(z) = (1 - \frac{1}{r_1}z)(1 - \frac{1}{r_2}z)$ 

$$\implies Y_t = \frac{1}{(1 - \frac{1}{r_1}z)(1 - \frac{1}{r_2}z)}\epsilon_t$$

$$\implies Y_t = \left(\sum_{i=0}^{\infty} \frac{1}{r_1^i} L^i\right) \left(\sum_{j=0}^{\infty} \frac{1}{r_2^j} L^j\right) \epsilon_t$$

as  $\left|\frac{1}{r_1}\right| < 1$  and  $\left|\frac{1}{r_2}\right| < 1$ 

$$\implies Y_t = \left(\sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \frac{1}{r_1^i} \frac{1}{r_2^j} L^{i+j}\right) \epsilon_t$$

with k = i + j

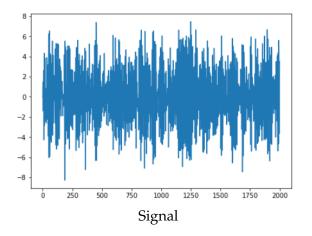
$$\implies Y_t = \left(\sum_{i=0}^{\infty} \sum_{k=i}^{\infty} \frac{1}{r_1^i} \frac{1}{r_2^{k-i}} L^k\right) \epsilon_t$$

$$\implies Y_t = \left(\sum_{k=0}^{\infty} \underbrace{\left(\sum_{i=0}^k \frac{1}{r_1^i} \frac{1}{r_2^{k-i}}\right)}_{a_k} L^k\right) \epsilon_t$$

$$\implies Y_t = \left(\sum_{k=0}^{\infty} a_k L^k\right) \epsilon_t$$

An AR(2) process  $Y_t$  can be expressed as a MA( $\infty$ ) process with characteristic polynomial  $\frac{1}{\phi(.)}$  thus, its power spectrum with fs=1 is:

$$S(f) = \sigma_{\epsilon}^2 \left| \frac{1}{\phi(e^{-2i\pi \times f})} \right|^2$$



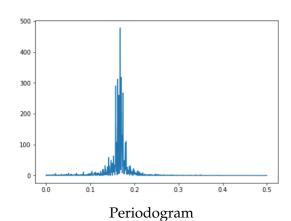


Figure 2: AR(2) process

# 4 Sparse coding

The modulated discrete cosine transform (MDCT) is a signal transformation often used in sound processing applications (for instance to encode a MP3 file). A MDCT atom  $\phi_{L,k}$  is defined for a length 2L and a frequency localisation k (k = 0, ..., L - 1) by

$$\forall u = 0, \dots, 2L - 1, \quad \phi_{L,k}[u] = w_L[u] \sqrt{\frac{2}{L}} \cos\left[\frac{\pi}{L} \left(u + \frac{L+1}{2}\right) (k + \frac{1}{2})\right]$$
 (3)

where  $w_L$  is a modulating window given by

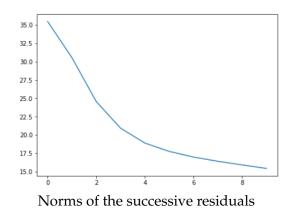
$$w_L[u] = \sin\left[\frac{\pi}{2L}\left(u + \frac{1}{2}\right)\right]. \tag{4}$$

## **Question 4** Sparse coding with OMP

For the signal provided in the notebook, learn a sparse representation with MDCT atoms. The dictionary is defined as the concatenation of all shifted MDCDT atoms for scales L in [32,64,128,256,512,1024].

- For the sparse coding, implement the Orthogonal Matching Pursuit (OMP). (Use convolutions to compute the correlations coefficients.)
- Display the norm of the successive residuals and the reconstructed signal with 10 atoms.

### **Answer 4**



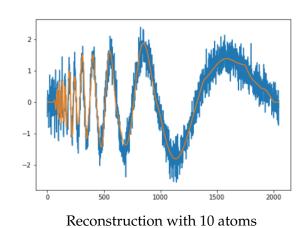


Figure 3: Question 4