#### Tufts CS 136 - 2024s - HW1 Submission

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<b>Collaboration Statement:</b>
Total hours spent: 20
I discussed ideas with these individuals:
• Patrick Feeney
I consulted the following resources:
<ul> <li>https://www.youtube.com/watch?v=emnfq4txDuI</li> </ul>
<ul><li>https://www.youtube.com/watch?v=3OgCcnpZtZ8</li></ul>
By submitting this assignment, I affirm this is my own original work that abides by the course collaboration policy.
Links: [HW1 instructions] [collab. policy]
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1a: Problem Statement

Let  $\rho \in (0.0, 1.0)$  be a Beta-distributed random variable:  $p \sim \text{Beta}(a, b)$ . Show that  $\mathbb{E}[\rho] = \frac{a}{a+b}$ .

Hint: You can use these identities, which hold for all a > 0 and b > 0:

$$\Gamma(a) = \int_{t=0}^{\infty} e^{-t} t^{a-1} dt \tag{1}$$

$$\Gamma(a+1) = a\Gamma(a) \tag{2}$$

$$\int_{0}^{1} \rho^{a-1} (1-\rho)^{b-1} d\rho = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$
 (3)

1a: Solution

The expected value is the total of all potential decision outcomes multiplied by each possibility's probability:

$$E(\rho) = \int \rho \cdot f(\rho) \, dx. \tag{4}$$

The probability density function of the beta distribution is the following:

$$f(\rho) = Beta(\mu|a,b) = \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \cdot \rho^{a-1} \cdot (1-\rho)^{b-1}, 0 \le x \le 1$$
 (5)

Then, we combine the expressions (4) and (5) and perform algebraic manipulation:

$$E(\rho) = \int_0^1 x \cdot \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \rho^{a-1} (1-\rho)^{b-1} dx \tag{6}$$

(7)

$$= \int_0^1 \frac{\Gamma(a+b)}{\Gamma(a) \cdot \Gamma(b)} \rho^{(a+1)-1} (1-\rho)^{b-1} dx$$
 (8)

(9)

$$= \frac{\Gamma(a+b)}{\Gamma(a)} \int_0^1 \frac{\rho^{(a+1)-1} (1-\rho)^{b-1}}{\Gamma(b)} dx$$
 (10)

(11)

$$= \frac{\Gamma(a+b)}{\Gamma(a)} \cdot \frac{\Gamma(a+1)}{\Gamma(a+b+1)} \int_0^1 \frac{\Gamma(a+b+1)}{\Gamma(a+1) \cdot \Gamma(b)} \cdot \rho^{(a+1)-1} (1-\rho)^{b-1} dx$$
(12)

Then, we use the identity  $\Gamma(a+1)=a\cdot\Gamma(a)$  to the coefficient and simplify it.

$$= \frac{\Gamma(a+b)}{\Gamma(a)} \cdot \frac{a \cdot \Gamma(a)}{(a+b) \cdot \Gamma(a+b)} \int_0^1 \frac{\Gamma(a+b+1)}{\Gamma(a+1) \cdot \Gamma(b)} \cdot \rho^{(a+1)-1} (1-\rho)^{b-1} dx$$
(13)

$$= \frac{a}{(a+b)} \int_0^1 \frac{\Gamma(a+b+1)}{\Gamma(a+1) \cdot \Gamma(b)} \cdot \rho^{(a+1)-1} (1-\rho)^{b-1} dx$$
 (14)

(15)

Recognize that the part in blue is a Beta distribution and substitute  $Beta(\mu|(a+1),b)$  in it:

$$= \frac{a}{(a+b)} \int_0^1 Beta(\mu|(a+1), b) \, dx \tag{16}$$

Since  $\int_0^1 Beta(\mu|(a+1),b) dx = 1$ , we have:

$$=\frac{a}{a+b}\tag{17}$$

#### 1b: Problem Statement

Let  $\mu$  be a Dirichlet-distributed random variable:  $\mu \sim \text{Dir}(a_1, \dots a_V)$ .

Show that  $\mathbb{E}[\mu_w] = \frac{a_w}{\sum_{v=1}^V a_v}$ , for any integer w that indexes a vocabulary word.

**Hint: You can use the identity:** 

$$\int \mu_1^{a_1 - 1} \mu_2^{a_2 - 1} \dots \mu_V^{a_V - 1} d\mu = \frac{\prod_{v=1}^V \Gamma(a_v)}{\Gamma(a_1 + a_2 \dots + a_V)}$$
(18)

1b: Solution

$$E[\mu_w] = \int \mu_w \cdot \frac{\Gamma(a_0)}{\Gamma(a_1) \dots \Gamma(a_V)} \cdot \prod_{v=1}^V \mu_v^{a_v - 1} d\mu$$
 (19)

(20)

Rearrange the terms with  $\mu$  in red for easier integration.

$$= \int \frac{\Gamma(a_0)}{\Gamma(a_1)\dots\Gamma(a_V)} \cdot \mu_w \cdot \mu_w^{a_w-1} \cdot \prod_{\substack{v=1\\v \neq w}}^V \mu_v^{a_v-1} d\mu$$
 (21)

$$= \int \frac{\Gamma(a_0)}{\Gamma(a_1)\dots\Gamma(a_V)} \cdot \mu_w^{(a_w+1)-1} \cdot \prod_{\substack{v=1\\v\neq w}}^V \mu_v^{a_v-1} d\mu$$
 (22)

(23)

Define a set of new parameters:

$$b_W = a_W + 1 \tag{24}$$

$$b_i = a_i, for 1 \le i \le V \tag{25}$$

For these parameters, we have these identities:

$$\Gamma(b_W) = \Gamma(a_W + 1) = a_W \cdot \Gamma(a_W) \tag{26}$$

$$\Gamma(b_i) = \Gamma(a_i) \tag{27}$$

$$b_0 = 1 + \sum_{v=1}^{V} a_i = 1 + a_0 \tag{28}$$

## Combining (26) and (27), we derive the following:

$$\Gamma(b_1)\dots\Gamma(b_v) = a_w \cdot \Gamma(a_1)\dots\Gamma(a_v)$$
(29)

$$\Gamma(a_1)\dots\Gamma(a_v) = \frac{\Gamma(b_1)\dots\Gamma(b_v)}{a_w}$$
(30)

(31)

## Using (28), we derive the following:

$$\Gamma(b_0) = \Gamma(1 + a_0) \tag{32}$$

$$= a_0 \cdot \Gamma(a_0) \tag{33}$$

$$\Gamma(a_0) = \frac{\Gamma(b_0)}{a_0} \tag{34}$$

(35)

### Substitute (30) and (31) in the the equation at (22):

$$= \int \frac{a_w}{a_0} \cdot \frac{\Gamma(b_0)}{\Gamma(b_1) \dots \Gamma(b_V)} \cdot \mu_w^{(a_w+1)-1} \cdot \prod_{\substack{v=1\\v \neq w}}^V \mu_v^{a_v-1} d\mu$$
 (36)

(37)

# Substitute (24) and (24) in the equation at (36):

$$= \int \frac{a_w}{a_0} \cdot \frac{\Gamma(b_0)}{\Gamma(b_1) \dots \Gamma(b_V)} \cdot \mu_w^{b_w - 1} \cdot \prod_{\substack{v=1\\v \neq w}}^V \mu_v^{b_v - 1} d\mu$$
 (38)

(39)

# Then, take the coefficients out oand simplify the integral:

$$= \frac{a_w}{a_0} \cdot \frac{\Gamma(b_0)}{\Gamma(b_1) \dots \Gamma(b_V)} \int \prod_{v=1}^V \mu_v^{b_v - 1} d\mu$$
 (40)

(41)

Realize that you can rewrite the given identity at (12) in question as:

$$\int \prod_{v=1}^{V} \mu_v^{b_v - 1} d\mu = \int \mu_1^{b_1 - 1} \dots \mu_V^{b_V - 1} d\mu \tag{42}$$

$$= \frac{\prod_{v=1}^{V} \Gamma(b_v)}{\Gamma(b_1) + \ldots + b_V} = \frac{\Gamma(b_1) \ldots \Gamma(b_V)}{\Gamma(b_0)}$$

$$\tag{43}$$

Substitute the simplified version of the give identity at (43) in the integral:

$$= \frac{a_w}{a_0} \cdot \frac{\Gamma(b_0)}{\Gamma(b_1) \dots \Gamma(b_V)} \cdot \frac{\Gamma(b_1) \dots \Gamma(b_V)}{\Gamma(b_0)}$$
(44)

$$=\frac{a_w}{a_0}\tag{46}$$

(45)

$$= \frac{a_w}{a_0}$$

$$= \frac{a_w}{\sum_{v=1}^{V} a_v}$$

$$(46)$$

2a: Problem Statement

Show that the likelihood of all N observed words can be written as:

$$p(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N | \mu) = \prod_{v=1}^V \mu_v^{n_v}$$
(48)

Hint: It may be helpful to recall the definition of the Categorical PMF using indicator notation:

$$p(X_n = x_n | \mu) = \prod_{v=1}^{V} \mu_v^{[x_n = v]}$$
(49)

Also, remember the relationship between this bracket notation and the count of how often vocabulary term v appears in the training data:  $n_v = \sum_{n=1}^{N} [x_n = v]$ 

2a: Solution

Because we assume that each word is conditionally independent of the other words given a parameter vector, the likelihood of all N observed words is:

$$p(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N | \mu) = \prod_{n=1}^N p(X_n = x_n | \mu)$$
 (50)

Furthermore, since each individual word is identically distributed according to  $\mu$ , we can write:

$$p(X_n = x_n | \mu) = \mu_{x_n} \tag{51}$$

Then, the likelihood can be written as:

$$p(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N | \mu) = \prod_{n=1}^N \mu_{x_n}$$
 (52)

We can reformulate the likelihood as a product over the vocabulary, where each vocabulary probability  $\mu_v$  is raised to the power of its corresponding

count.

$$L(\mu) = \prod_{n=1}^{N} \mu_{x_n} = \prod_{v=1}^{V} \mu_v^{n_v}$$
 (53)

The Iverson bracket  $[x_n = v]$  counts occurrences. It equals 1 if  $x_n$  equals v (the n-th word is the v-th vocabulary term) and 0 otherwise. Therefore,  $n_v$ , can be expressed using the Iverson bracket as:

$$n_v = \sum_{n=1}^{N} [x_n = v]$$
 (54)

**2b: Problem Statement** 

Derive the next-word posterior predictive, after integrating away parameter  $\mu$ .

That is, show that after seeing the N training words, the probability of the next word  $X_*$  being vocabulary word v is:

$$p(X_* = v | X_1 = x_1 \dots X_N = x_N) = \int p(X_* = v, \mu | X_1 = x_1 \dots X_N = x_N) d\mu$$

$$= \frac{n_v + \alpha}{N + V\alpha}$$
(55)

Hint: You will use the expectation of a Dirichlet-distributed random variable that we proved in 1b

2b: Solution

The posterior predictive for a new observation  $X^*$  given the data D is the integral over all possible parameter vectors  $\mu$ :

$$p(X^* = v|D) = \int p(X^* = v|\mu, D)p(\mu|D)d\mu$$
 (56)

Since  $X^*$  is conditionally independent of D given  $\mu$ , we can simplify the first term in the integral as following:

$$p(X^* = v|D) = \int p(X^* = v|\mu)p(\mu|D)d\mu$$
 (57)

From the problem statement, we know that the posterior  $p(\mu|D)$  is Dirichlet with parameters  $\alpha+m$ . So, we use the equation for the Dirichlet posterior (Eq. 2.41) and substitute  $p(\mu|D)$ :

$$p(X^* = v|D) = \int \mu_v \frac{\Gamma(\alpha_0 + N)}{\prod_{k=1}^K \Gamma(\alpha_k + m_k)} \prod_{k=1}^K \mu_k^{\alpha_k + m_k - 1} d\mu$$
 (58)

Recognize that the integral in (57) is the expected value of  $\mu_v$  under the Dirichlet distribution, and therefore we don't need to perform an explicit

integration:

$$E[\mu_v] = \frac{\alpha_v + m_v}{\sum_{v=1}^{V} (\alpha_v + m_v)}$$
(59)

We know that, in our Dirichlet distribution,  $\ln \alpha_k$  are equal to  $\alpha$ , so  $\alpha_v = \alpha$  for all v, and  $m_v = n_v$ , which is the count of the v-th word. So, we substitute these in.

$$p(X^* = v|D) = E[\mu_v] = \frac{\alpha_v + n_v}{\sum_{v=1}^{V} (\alpha_v + n_v)}$$
 (60)

Since  $\alpha_v=\alpha$  for all v, and  $\sum_{v=1}^V(\alpha_v+n_v)=V\alpha+N$  because the sum of all  $n_v$  is N, the total count of words, we have:

$$p(X^* = v|D) = \frac{n_v + \alpha}{V\alpha + N} \tag{61}$$

#### 2c: Problem Statement

Derive the marginal likelihood of observed training data, after integrating away the parameter  $\mu$ .

That is, show that the marginal probability of the observed N training words has the following closed-form expression:

$$p(X_1 = x_1 \dots X_N = x_N) = \int p(X_1 = x_1, \dots X_N = x_N, \mu) d\mu$$
 (62)

$$= \frac{\Gamma(V\alpha) \prod_{v=1}^{V} \Gamma(n_v + \alpha)}{\Gamma(N + V\alpha) \prod_{v=1}^{V} \Gamma(\alpha)}$$
(63)

2c: Solution

Using the product rule, we can rewrite the marginal probability of the observed N training words as follows:

$$p(X_1 = x_1, \dots, X_N = x_N) \tag{64}$$

$$= \int p(X_1 = x_1, \dots X_N = x_N, \mu) d\mu$$
 (65)

$$= \int p(X_1 = x_1, \dots, X_N = x_N | \mu) p(\mu) d\mu$$
 (66)

First, let's evaluate likelihood of the data given the parameter  $\mu$ . We know that

$$p(X_1 = x_1, \dots, X_N = x_N | \mu) = \prod_{v=1}^V \mu_v^{n_v}$$
(67)

because of conditional independence from 2(a).

Second, we evaluate the Dirichlet prior  $\rho(\mu)$ . Dirichlet prior is

$$p(\mu) = \frac{1}{B(\alpha)} \prod_{v=1}^{V} \mu_v^{\alpha_v - 1}$$
 (68)

where  $B(\alpha)$  is the beta function (or the normalization constant for the Dirichlet distribution), and  $\alpha_v$  are the parameters of the Dirichlet distribution.

### Combine the likelihood and the prior:

$$p(X_1 = x_1, \dots, X_N = x_N) = \int \left(\prod_{v=1}^V \mu_v^{n_v}\right) \left(\frac{1}{B(\alpha)} \prod_{v=1}^V \mu_v^{\alpha_v - 1}\right) d\mu$$
 (69)

(70)

Simplify the expression and recognize the new Dirichlet distribution.

$$= \int \frac{1}{B(\alpha)} \prod_{v=1}^{V} \mu_v^{n_v + \alpha_v - 1} d\mu \tag{71}$$

(72)

$$=\frac{1}{B(\alpha)}B(n+\alpha)\tag{73}$$

(74)

Now, we can use the properties of the gamma function and the definition of the beta function in terms of gamma functions:

$$B(\alpha) = \frac{\prod_{v=1}^{V} \Gamma(\alpha_v)}{\Gamma\left(\sum_{v=1}^{V} \alpha_v\right)}$$
$$B(n+\alpha) = \frac{\prod_{v=1}^{V} \Gamma(n_v + \alpha_v)}{\Gamma\left(\sum_{v=1}^{V} (n_v + \alpha_v)\right)}$$

Combining these properties, we get the marginal likelihood:

$$= \frac{\Gamma\left(\sum_{v=1}^{V} \alpha_v\right)}{\prod_{v=1}^{V} \Gamma(\alpha_v)} \frac{\prod_{v=1}^{V} \Gamma(n_v + \alpha_v)}{\Gamma\left(\sum_{v=1}^{V} (n_v + \alpha_v)\right)}$$
(75)

(76)

$$= \frac{\Gamma(V\alpha) \prod_{v=1}^{V} \Gamma(n_v + \alpha)}{\Gamma(N + V\alpha) \prod_{v=1}^{V} \Gamma(\alpha)}$$
(77)