

Logistic Regression

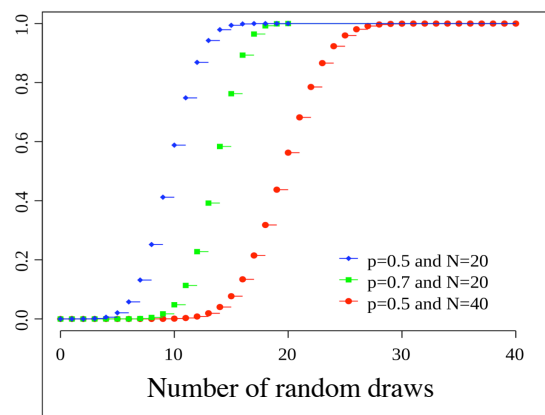
Lecture 6

CLPS 2908 Multivariate Statistical Techniques
February 14, 2019

Categorical DVs?

Multiple Regression

- Assumes normal error distributions — but binary outcomes have binomial error distributions
 - violates homoscedasticity



Categorical DVs?

Multiple Regression

- Assumes normal error distributions — but binary outcomes have binomial error distributions
 - violates homoscedasticity
- MR → impossible predicted values: < 0 (no), > 1 (yes)

One option: Discriminant Function Analysis; later

Logistic regression

- No distributional assumptions about Y or ϵ
- Bounded between 0 and 1 (probability of event)
- No constraints on predictors [unlike logit; later]
- Fits curved regression line to $p(event) \leftarrow predictors$
- Maximum Likelihood estimation instead of least-squares (no linear variance decomposition)

Logistic Regression Logic

X_1, X_2, \dots

$$u = B_0 + B_1X_1 + \dots + B_pX_p$$

→ e^u

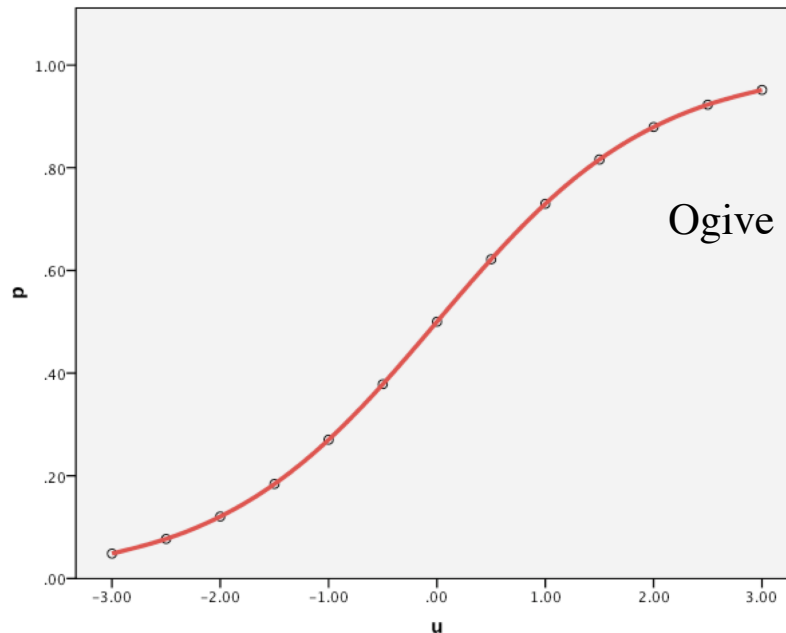
allows us to estimate:

$$p(event) = \frac{e^u}{1 + e^u}$$

$$p = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}}$$

u	e^u	p
3	20.10	.95
2	7.40	.88
1	2.72	.70
.5	1.65	.62
0	1.0	.50
-.5	.61	.38
-1	.37	.27
-2	.14	.12
-3	.05	.05

p as a function of ln-transformed u $p = \frac{e^u}{1 + e^u}$



Model

- ML estimation: **B values are chosen** so as to minimize total **prediction errors** (over all X_j), subject to this constraint:

$$\sum_{i=1}^n (y_i - \hat{p}_i) x_{ij} = 0$$

= no error-predictor correlation

- Iterative process of re-estimating covariance matrix and B values (\rightarrow *best possible B parameters*)
- Log-likelihood (total remaining error \sim summed residuals)

$$\sum_{i=1}^n (y_i \ln(\hat{p}_i) + (1 - y_i) \ln(1 - \hat{p}_i))$$

What contributes to LL being maximally negative (large error):
predicted p is extreme (e.g., .05 or .95) but incorrect.

Example: If $p = .05$ for a given person whose real value y_i is 1, then the LL term is: $1 * (-3.03)$ [right term drops out because $1 - y = 0$ in this case].

Comparisons among Predictors in LR

- Weights B_1 , B_2 , etc. are again **semi-partial weights**

- The **Wald** statistic tests their significance (slightly conservative)
$$Wald_i = \frac{B_i^2}{SE_i^2}$$

- Semi-partial r**
$$r_{semi-partial} = \sqrt{\frac{Wald - 2df}{-2LL_0}}$$

- Don't have 1 as their maximum, so they are typically **smaller** than what a normal correlation would give you.
- Nevertheless, allow us to compare various predictors to each other
- Familiar:** Compare with zero-order r s to gain a sense of shared/suppressed predictive power.
- SPSS does not display these
- Compute them using this formula or the EXCEL file in the canvas folder.
- Null model $-2LL_0$ requires full output option.



Semi-partial calculator

Calculating semi-partial r's in logistic regression	
WALD	2 [Wald must be greater than 2 for the computation to work]
df	1 For continuous predictors, df = 1; for categorical predictors with k categories, df = k-1
-2LL(null)	266.11
r	0.0000

Comparisons (cont' d)

- Bs** are a little difficult to interpret in LR because they are “stuck” in the power of e

$$p(\text{event}) = \frac{e^u}{1 + e^u}$$

- Pull them out: $u = \ln(\text{odds of event}) = \text{“log-odds”} = \text{logit}$

$$u = \ln\left(\frac{p}{1-p}\right) = b_0 + b_1X_1 + \dots + b_jX_j$$

- So every B-fold unit change of X goes along with a change of the log-odds (still not particularly intuitive...).

Odds in statistics = “odds for” = $f(\text{event})$ vs. $f(\text{nonevents})$
e.g., odds of rolling a “4” are 1/5

Odds in gambling = “odds against” (5/1)

Exp(B)

- **Exp(B)**, for a given X , is somewhat more interpretable: it is the factor by which the **actual odds** change with every B-fold unit change in X .
- **Exp(B) around 1** is unimpressive;
 - if > 1 , predictor **increases** $p(\text{event})$;
 - if < 1 , predictor **decreases** $p(\text{event})$.
- $\frac{p}{1-p} = e^{b_0 + b_1 X_1 + \dots + b_j X_j}$
- **Exp(B)** is just the \ln -transformation of raw B (e^B);
B is semi-partial, so **exp(B)** must be semi-partial too.
- That is, **exp(B_i)** represents the impact of its X_i on Y **while holding constant all other predictors**.



Exponent B effects.xls

Overall Success

- R^2 **analog** $R^2_{Cox-Snell} = 1 - \left(\frac{l(\text{null})}{l(\text{current})} \right)^{\frac{2}{w}}$
- Menard suggests to compute **R^2 of predicted & observed Y_i values**: $[0;1] * 0,1$ (SPSS can save predicted values, making it easy to compute the correlation)
- Most common:
Case classification (“confusion matrix”)

		Predicted Y	
Actual Y	32	5	
	11	22	

Confusion Matrix

Predicted \ Real	"Present"	"Absent"	
Present	30	20	50
Absent	10	40	50
	40	60	

Two types of errors:

False alarm = $10/50 = 20\%$

Misses = $20/50 = 40\%$

Diagnosticity:
30/40

Two Ways to Take Control

- SPSS allows you to define contrasts for categorical predictors in this subcommand:
/CONTRAST (predictor) = [HELMERT, POLYNOMIAL, etc.]
(Read about it in the SPSS syntax reference file.)
- You can change classification cut-off:
/CRITERIA = CUT(0.5)

Additional Literature:

Hosmer, D. W., & Lemeshow, S. (2000). *Applied logistic regression*. New York: Wiley.

Menard, S. W. (2002). *Applied Logistic Regression Analysis*. Sage.

Model Building and Testing

- -2LL ratios for models with/without X_j

$$LR = -2 \ln \left(\frac{l(\text{reduced})}{l(\text{full})} \right) \quad LR = -2(LL(\text{reduced}) - LL(\text{full}))$$

- Define predictor sets, guided by hypotheses
- Best subsets (time consuming, out of fashion)

- Goodness of fit (χ^2).
$$\sum_{i=1}^n \frac{(y_i - \hat{p}_i)^2}{\hat{p}_i(1 - \hat{p}_i)}$$
 - Actually, **badness of fit**

- Run the analysis with and without the predictor and compare the classification results.

Sample Output

LOGREG output.pdf
(with annotations)

```
LOGISTIC REGRESSION VARIABLES natur6  
/METHOD=ENTER degree6 contr6 source6  
/CLASSPLOT  
/CASEWISE  
/PRINT = all.
```