Generalized Linear Models(Lecture 5)

Zhengfan Wang September 26, 2017

Variance/Covariance "wrap up"

$$\beta^{(t+1)} = \beta^{(t)} + (I^{(t)})^{-1} * \mu^{(t)},$$

where I is variance/covariance matrix of estimate using Fischer scoring.

- Characterizing uncertainty
- Approximation to the likelihood or posterior surface
- Frequentist Method/Fisher Scoring
 - relying on large sample approximation to likelihood surface
 - relies on distributional assumptions
- Other options
 - more computationly intensive
 - hard to use w/black box
 - sampling based (Bayesian MCMC; Gibbs sampling)

Inference Miscellany

Setting:logistic regression and the model is:

$$logit(\pi_i) = \alpha + \beta x_i$$

In large sample and $x = x_k$, the SE for

 $logit(\hat{\pi_k})$

is described as:

$$\sqrt{var(\hat{\alpha} + \hat{\beta}x_k)}$$

And the variance can be written as

$$var(\hat{\alpha} + \hat{\beta}x_k) = var(\hat{\alpha}) + x_k^2 var(\hat{\beta}) + 2x_k cov(\hat{\alpha}, \hat{\beta})$$

- Confidence Interval

95% C.I. for

$$logit(\hat{\pi_k}) = logit(\hat{\pi_k}) \pm 2 * SE(logit(\hat{\pi_k}))$$

If the C.I. is (a,b), 95% C.I. for

$$\hat{\pi_k} \in (\frac{e^a}{1+e^a}, \frac{e^b}{1+e^b})$$

R function is

predict(mymodel,type="response")

Model Checking and Building

Checking: does the model fit data well?

- residual plots/analysis
- predicted v.s. observed
- instability in model estimate (large SEs?) colinearity?

Building: process by which you arrive at a "chosen" model

- systematic/pre-specificed analysis plan
- corvariate choices
- selection criteria. e.g. AIC(Akaike information criterion),BIC(Bayesian information criterion), Likelihood-ratio test
- describe/explain analysis
- specify type 1 error
- specify correction for multiple testing
- specify validation SE
 - iterative/subject
- use residual plots or other G.O.F measure to modify model
- incorporate domain-specific knowledge, e.g. other corvariates
- vigilant about "gorden of forking paths"
- important to use validation samples
- penalized regression (LASSO(least absolute shrinkage and selection operator, often used in Bioinformatics), GAMs,...)
- ensemble of "reasonable" model

Poisson GLMs

$$y_i \sim poisson(\lambda_i)$$

$$\eta_i = X_i \beta$$

$$g(E(y_i)) = log(\lambda_i) = \eta_i = X_i \beta$$

Loglink

• implies covariates have multiplicate effect.

$$\lambda_i = e^{\beta_0} * e^{\beta_1 x_1} * e^{\beta_2 x_2} * \dots$$

• relative risk/rate interpretation for e^{β_k}

$$\begin{split} \log \lambda_i &= \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} \\ (\log \lambda_i \mid x_1 = k+1, x_2 = j) &= \beta_0 + \beta_1 (k+1) + \beta_2 j \\ (\log \lambda_i \mid x_1 = k, x_2 = j) &= \beta_0 + \beta_1 k + \beta_2 j \\ \beta_1 &= (\log \lambda_i \mid x_1 = k+1, x_2 = j) - (\log \lambda_i \mid x_1 = k, x_2 = j) = \log (\frac{\lambda_i \mid x_1 = k+1}{\lambda_i \mid x_1 = k}) = \log (RR) \\ Relative Risk &= e^{\beta_1} \end{split}$$

Holding all other variables constant, the expected value of y change $[(e^{\beta}-1)*100]\%$, e.g. $(0.8-1)*100 \rightarrow -20\%$ decrease; $(1.2-1)*100 \rightarrow 20\%$ increase.

Exposure/Offset

outcome of interest e.g.1 rank of hospitalizations by county

 $y_i = \#$ of times of outcome occurred

 $u_i = \text{offset or exposure}$

e.g.2

 $y_i = \#$ of case of flu in a population

 $u_i = \text{Population}$

e.g.3

 $y_i = \#$ of traffic accident at intersection i in one day

 $u_i = 1$.average number of vichicles at intersection today. 2.average number of vichicles at intersection yesterday

$$y_i \sim poisson(\mu_i \lambda_i)$$

 $E(y_i) = \mu_i \lambda_i$

$$log(E(y_i)) = log\mu_i + log\lambda_i = log\mu_i + X\beta$$

which $log\mu_i$ is the offset and $X\beta$ is the linear predictor.

$$log(\lambda_i) = \eta_i = x_i \beta$$

And in R

 $glm(y \sim x_1 + x_2, offset = log(a), family = poisson, ...)$ offset is not aleardy on log scale.