

# CS 236 Autumn 2019/2020 Homework 1

SUNet ID: 06009508

Name: Brandon McKinzie

Collaborators: N/A

By turning in this assignment, I agree by the Stanford honor code and declare that all of this is my own work.

## Problem 1: MLE and KL Divergence

$$D_{KL}(\hat{p}(y | x) || p_{\theta}(y | x)) = \mathbb{E}_{\hat{p}(y|x)} \left[ \log \frac{\hat{p}(y | x)}{p_{\theta}(y | x)} \right] \quad (1)$$

$$= \mathbb{E}_{\hat{p}(y|x)} [\log \hat{p}(y | x)] - \mathbb{E}_{\hat{p}(y|x)} [\log p_{\theta}(y | x)] \quad (2)$$

$$\arg \min_{\theta \in \Theta} \mathbb{E}_{\hat{p}(x)} [D_{KL}(\hat{p}(y | x) || p_{\theta}(y | x))] = \arg \min_{\theta \in \Theta} \mathbb{E}_{\hat{p}(x)} \left[ \mathbb{E}_{\hat{p}(y|x)} [\log \hat{p}(y | x)] - \mathbb{E}_{\hat{p}(y|x)} [\log p_{\theta}(y | x)] \right] \quad (3)$$

$$= \arg \min_{\theta \in \Theta} \mathbb{E}_{\hat{p}(y,x)} [\log \hat{p}(y | x)] - \mathbb{E}_{\hat{p}(y,x)} [\log p_{\theta}(y | x)] \quad (4)$$

$$= - \arg \min_{\theta \in \Theta} \mathbb{E}_{\hat{p}(y,x)} [\log p_{\theta}(y | x)] \quad (5)$$

$$= \arg \max_{\theta \in \Theta} \mathbb{E}_{\hat{p}(y,x)} [\log p_{\theta}(y | x)] \quad (6)$$

## Problem 2: Logistic Regression and Naive Bayes

We are asked to show that  $\forall \theta, \exists \gamma$  such that

$$p_\theta(y \mid x) = p_\gamma(y \mid x) \quad (7)$$

First, note that we can rewrite  $p_\theta(y \mid x)$  using Bayes' rule and basic probability:

$$p_\theta(y \mid x) = \frac{p_\theta(x \mid y)p_\theta(y)}{\sum_{y'} p_\theta(x \mid y')p_\theta(y')} \quad (8)$$

Then, substituting in the formulas we were provided:

$$p_\theta(y \mid x) = \frac{\mathcal{N}(x \mid \mu_y, \sigma^2 I) \pi_y}{\sum_{y'} \mathcal{N}(x \mid \mu_{y'}, \sigma^2 I) \pi_{y'}} \quad (9)$$

$$= \frac{\exp(-(x - \mu_y)^2 / (2\sigma^2)) \pi_y}{\sum_{y'} \exp(-(x - \mu_{y'})^2 / (2\sigma^2)) \pi_{y'}} \quad (10)$$

where I've used compact notation to denote  $(x - \mu_y)^T(x - \mu_y)$  simply as  $(x - \mu_y)^2$ . In what follows, let  $\alpha = -\frac{1}{2\sigma^2}$ . I'll also be writing  $x^T x$  as  $x^2$ , etc. We can expand this out as follows:

$$\frac{\exp(-(x - \mu_y)^2 / (2\sigma^2)) \pi_y}{\sum_{y'} \exp(-(x - \mu_{y'})^2 / (2\sigma^2)) \pi_{y'}} = \frac{\pi_y e^{\alpha x^2} e^{\alpha \mu_y^2} e^{-\alpha 2x^T \mu_y}}{\sum_{y'} \pi_{y'} e^{\alpha x^2} e^{\alpha \mu_{y'}^2} e^{-\alpha 2x^T \mu_{y'}}} \quad (11)$$

$$= \frac{\pi_y e^{\alpha \mu_y^2} e^{-\alpha 2x^T \mu_y}}{\sum_{y'} \pi_{y'} e^{\alpha \mu_{y'}^2} e^{-\alpha 2x^T \mu_{y'}}} \quad (12)$$

$$(13)$$

Therefore, for any  $\theta$ , if we define  $\gamma$  using:

$$b_y := \log(\pi_y e^{\alpha \mu_y^2}) = \log(\pi_y) + \alpha \mu_y^2 \quad (14)$$

$$w_y := -2\alpha \mu_y \quad (15)$$

then we will satisfy  $p_\theta(y \mid x) = p_\gamma(y \mid x)$ .

### Problem 3: Conditional Independence and Parameterization

1. The total number of independent parameters is  $(\prod_{i=1}^n k_i) - 1$ .
2. Our network factorizes the joint distribution as

$$p(X_1, X_2, \dots, X_n) = p(X_1) \left[ \prod_{i=2}^m p(X_i \mid X_{i-1}, \dots, X_1) \right] \left[ \prod_{i=m+1}^n p(X_i \mid X_{i-1}, \dots, X_{i-m}) \right] \quad (16)$$

- $p(X_1)$  requires  $k_1 - 1$  parameters.
- For  $1 < i \leq m$ ,  $p(X_i \mid X_{i-1}, \dots, X_1)$  requires  $(\prod_{j=1}^i k_j) - 1$  parameters.
- For  $i > m$ ,  $p(X_i \mid X_{i-1}, \dots, X_{i-m})$  requires  $(\prod_{j=i-m}^i k_j) - 1$  parameters.

In total, the number of parameters is thus

$$(k_1 - 1) + \sum_{i=2}^m \left[ \left( \prod_{j=1}^i k_j \right) - 1 \right] + \sum_{i=m+1}^n \left[ \left( \prod_{j=i-m}^i k_j \right) - 1 \right] \quad (17)$$

3. To represent a Bayesian network over  $X_1, \dots, X_n$  with  $\sum_{i=1}^n (k_i - 1)$  parameters, you'd need to impose  $X_i \perp X_j \forall i, j \neq i$ .

## Problem 4: Autoregressive Models

Given any choice of  $\{\mu_i, \sigma_i\}_{i=1}^n$ , does there always exist a choice of  $\{\hat{\mu}_i, \hat{\sigma}_i\}_{i=1}^n$  such that  $p_f = p_r$ ?

As suggested by the hint, consider the case where  $n = 2$ . Then we have

$$p_f(x_1, x_2) = \mathcal{N}(x_1 \mid \mu_1, \sigma_1^2) \mathcal{N}(x_2 \mid \mu_2(x_1), \sigma_2^2(x_1)) \quad (18)$$

$$p_r(x_1, x_2) = \mathcal{N}(x_1 \mid \hat{\mu}_1(x_2), \hat{\sigma}_1^2(x_2)) \mathcal{N}(x_2 \mid \hat{\mu}_2, \hat{\sigma}_2^2) \quad (19)$$

which reveals the answer that *no, these models do not cover the same hypothesis space of distributions.*

For example, if  $\mu_1 = 0$  and  $\sigma_1^2$  is sufficiently close to zero, then for negligibly small  $\epsilon$ :

$$p_f(x_1, x_2) \approx \begin{cases} 0 & |x_1| < \epsilon \\ \mathcal{N}(x_2 \mid \mu_2(x_1), \sigma_2^2(x_1)) & |x_1| > \epsilon \end{cases} \quad (20)$$

The only way to ensure  $p_r = p_f$  in this case is for  $\mu_2(x_1)$  and  $\sigma_2^2(x_1)$  to be constants (since  $\hat{\mu}_2$  and  $\hat{\sigma}_2^2$  are constants). Otherwise,  $p_f(x_1, x_2)$  is essentially a gaussian over  $x_2$  with moving mean and variance as a function of  $x_1$ , which  $p_r$  is not able to model.

## Problem 5: Monte Carlo Integration

1. Show that  $A$  is an unbiased estimator of  $p(x)$ .

$$\mathbb{E}_{p(z)} [A] = \frac{1}{k} \sum_{i=1}^k \mathbb{E}_{p(z)} [p(x \mid z^{(i)})] \quad (21)$$

$$= \frac{1}{k} \sum_{i=1}^k p(x) \quad (22)$$

$$= p(x) \quad (23)$$

2. Is  $\log A$  an unbiased estimator of  $\log p(x)$ ? Explain why or why not. No,  $\log A$  is not an unbiased estimator of  $\log p(x)$ , because  $\mathbb{E}_{p(z)} [\log(f(z))] \neq \log \mathbb{E}_{p(z)} [f(z)]$ <sup>1</sup>.

---

<sup>1</sup>Note that if  $A$  were not a function of the  $z^{(i)}$  variables, then (trivially)  $\mathbb{E}_{p(z)} [\log(A)] = \log \mathbb{E}_{p(z)} [A] = \log A$

## Problem 6: Programming Assignment

1. Suppose we wish to find an efficient bit representation for the 50257 tokens. That is, every token is represented as  $(a_1, \dots, a_n)$ , where  $a_i \in \{0, 1\}, \forall i = 1, 2, \dots, n$ . What is the minimal  $n$  that we can use?

Since we are confined to use a fixed-length code of size  $n$ , we cannot exploit regularities/patterns in the data distribution. As such, the minimal  $n$  we can use is  $\lceil \lg 50257 \rceil = 16$ .

2. If the number of possible tokens increases from 50257 to 60000, what is the increase in the number of parameters? Give an exact number and explain your answer.

The only layer that is impacted by this is the final fully-connected layer, which projects the transformer state to the output vocabulary space. The FC layer has a kernel with  $768 \times V$  parameters (where  $V$  is vocabulary size)<sup>2</sup>. Therefore, increasing  $V$  from 50257 to 60000 results in a parameter increase of

$$768 \times (60000 - 50257) \tag{24}$$

which is equal to **7,482,624** additional parameters.

---

<sup>2</sup>GPT2 does not use a bias vector in their final projection. See line 171 of their model.py file.