

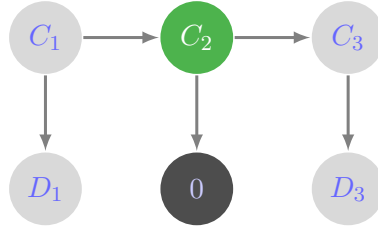
Homework 7: Car Tracking

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PROBLEM 1: BAYESIAN NETWORK BASICS

(a) Suppose we have a sensor reading for the second timestep, $D_2 = 0$. Compute the posterior distribution $\mathbb{P}(C_2 = 1 \mid D_2 = 0)$.

Below is the Bayesian network, where we've observed $D_2 = 0$:



$$\Pr [C_2 = 1 \mid D_2 = 0] \propto \Pr [C_2 = 1, D_2 = 0] \quad (1)$$

$$= \sum_{c_1} \Pr [C_2 = 1, D_2 = 0, c_1] \quad (2)$$

$$= \sum_{c_1} \Pr [c_1] \Pr [C_2 = 1 \mid c_1] \Pr [D_2 = 0 \mid C_2 = 1] \quad (3)$$

$$= 0.5 \sum_{c_1} \Pr [C_2 = 1 \mid c_1] \Pr [D_2 = 0 \mid C_2 = 1] \quad (4)$$

$$= 0.5\eta \sum_{c_1} \Pr [C_2 = 1 \mid c_1] \quad (5)$$

$$= 0.5\eta(\epsilon + (1 - \epsilon)) \quad (6)$$

$$= 0.5\eta \quad (7)$$

$$\Pr [D_2 = 0] = \sum_{c_2} \Pr [D_2 = 0, c_2] \quad (8)$$

$$= \sum_{c_2} \Pr [D_2 = 0 \mid c_2] \sum_{c_1} \Pr [c_2 \mid c_1] \Pr [c_1] \quad (9)$$

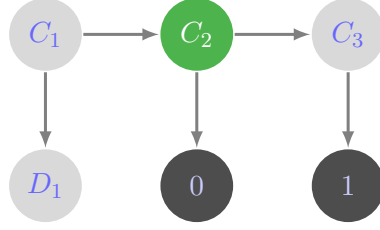
$$= (1 - \eta) \cdot (0.5(1 - \epsilon) + 0.5\epsilon) + \eta \cdot (0.5\epsilon + 0.5(1 - \epsilon)) \quad (10)$$

$$= 0.5 \quad (11)$$

$$\therefore \Pr [C_2 = 1 \mid D_2 = 0] = \frac{\Pr [C_2 = 1, D_2 = 0]}{\Pr [D_2 = 0]} = \eta \quad (12)$$

(b) Compute $\mathbb{P}(C_2 = 1 \mid D_2 = 0, D_3 = 1)$

Now our Bayesian network looks like:



$$p(C_2, D_2, D_3) = \sum_{c_1} \sum_{c_3} p(c_1) p(c_2 \mid c_1) p(d_2 \mid c_2) p(c_3 \mid c_2) p(d_3 \mid c_3) \quad (13)$$

$$= 0.5 \sum_{c_1} \sum_{c_3} p(c_2 \mid c_1) p(d_2 \mid c_2) p(c_3 \mid c_2) p(d_3 \mid c_3) \quad (14)$$

$$= 0.5 \cdot p(d_2 \mid c_2) \sum_{c_3} p(c_3 \mid c_2) p(d_3 \mid c_3) \quad (15)$$

$$p(C_2 = 1, D_2 = 0, D_3 = 1) = 0.5\eta (\epsilon\eta + (1 - \epsilon)(1 - \eta)) \quad (16)$$

$$p(D_2 = 0, D_3 = 1) = \sum_{c_2} p(c_2, D_2 = 0, D_3 = 1) \quad (17)$$

$$= 0.5\eta (\epsilon\eta + (1 - \epsilon)(1 - \eta)) + 0.5(1 - \eta) ((1 - \epsilon)\eta + \epsilon(1 - \eta)) \quad (18)$$

$$p(C_2 = 1 \mid D_2 = 0, D_3 = 1) = \frac{0.5\eta (\epsilon\eta + (1 - \epsilon)(1 - \eta))}{0.5\eta (\epsilon\eta + (1 - \epsilon)(1 - \eta)) + 0.5(1 - \eta) ((1 - \epsilon)\eta + \epsilon(1 - \eta))} \quad (19)$$

$$= \frac{\epsilon\eta^2 + \eta(1 - \epsilon)(1 - \eta)}{\epsilon\eta^2 + 2(1 - \eta)(1 - \epsilon)\eta + \epsilon(1 - \eta)^2} \quad (20)$$

(c)

i.

$$P(C_2 = 1 \mid D_2 = 0) = \eta = 0.2 \quad (21)$$

$$P(C_2 = 1 \mid D_2 = 0, D_3 = 1) = \frac{(0.1)(0.2)^2 + 0.2(0.9)(0.8)}{(0.1)(0.2)^2 + 2(0.8)(0.9)(0.2) + (0.1)(0.8)^2} \quad (22)$$

$$\approx 0.4157 \quad (23)$$

ii. The second reading ($D_3 = 1$) makes it more likely that $C_2 = 1$. Informally, after we observe $D_3 = 1$ while realizing that it's unlikely the car has actually moved (relative to the probability that our sensor was wrong), it becomes more probable that our previous reading of $D_2 = 0$ was just a bad reading.

iii. We can compute the value of ϵ by equating the two formulas and solving:

$$\eta = \frac{\epsilon\eta^2 + \eta(1 - \epsilon)(1 - \eta)}{\epsilon\eta^2 + 2(1 - \eta)(1 - \epsilon)\eta + \epsilon(1 - \eta)^2} \quad (24)$$

$$1 = \frac{\epsilon\eta + (1 - \epsilon)(1 - \eta)}{\epsilon\eta^2 + 2(1 - \eta)(1 - \epsilon)\eta + \epsilon(1 - \eta)^2} \quad (25)$$

$$\epsilon\eta + (1 - \epsilon)(1 - \eta) = \epsilon\eta^2 + 2(1 - \eta)(1 - \epsilon)\eta + \epsilon(1 - \eta)^2 \quad (26)$$

$$\epsilon\eta = \epsilon\eta^2 + (2\eta - 1)(1 - \eta)(1 - \epsilon) + \epsilon(1 - \eta)^2 \quad (27)$$

$$\epsilon(\eta - \eta^2 - (1 - \eta)^2) = (2\eta - 1)(1 - \eta)(1 - \epsilon) \quad (28)$$

$$\epsilon((1 - \eta)(\eta - (1 - \eta))) = (2\eta - 1)(1 - \eta)(1 - \epsilon) \quad (29)$$

$$\epsilon(1 - \eta)(2\eta - 1) = (2\eta - 1)(1 - \eta)(1 - \epsilon) \quad (30)$$

$$\epsilon = 1 - \epsilon \quad (31)$$

$$\epsilon = \frac{1}{2} \quad (32)$$

Intuitively, we'd have to set $\epsilon = \frac{1}{2}$ (car equally likely to be in 0 or 1 independent of previous location) since that would make additional observations essentially useless.

PROBLEM 5: WHICH CAR IS IT?

- $C_{ti} \in \mathbb{R}^2$: location of i th car at time t , where $1 \leq i \leq K$, and $1 \leq t \leq T$.
- $D_t = \{D_{t1}, \dots, D_{tK}\}$, where each $D_{ti} \in \mathbb{R}$ is noisy distance measurement of i th car at time t .

(a) Write an expression for $\Pr[C_{11}, C_{12} \mid E_1 = e_1]$ as a function of $\mathcal{N}(v; \mu, \sigma^2)$ and the priors $p(c_{11})$ and $p(c_{12})$.

Since $T = 1$, I'm going to drop the time index in the following calculations to avoid confusing myself. Then, since $K = 2$, I'll denote $E_1 \equiv E \equiv (E_1, E_2)$ (again, I've dropped the time index since it is always 1 for this problem). So we need an expression for $\Pr[C_1, C_2 \mid E_1, E_2]$.

First, we can rewrite

$$\Pr[C_1, C_2 \mid E_1, E_2] \propto \Pr[E_1, E_2 \mid C_1, C_2] \Pr[C_1, C_2] \quad (33)$$

$$= \Pr[E_1, E_2 \mid C_1, C_2] \Pr[C_1] \Pr[C_2] \quad (34)$$

where we've taken advantage of the independence of the two cars. Now we need to break down $\Pr[E_1, E_2 \mid C_1, C_2]$. Since E is sampled uniformly at random from the possible permutations in D , we can think of a given instantiation $E = (e_1, e_2)$ as an event whose probability is proportional to the union of all possible readings (d_1, d_2) that could result in (e_1, e_2) . Formally,

$$\Pr[E_1 = e_1, E_2 = e_2 \mid C_1, C_2] \propto \Pr[D_1 = e_1, D_2 = e_2 \mid C_1, C_2] + \Pr[D_1 = e_2, D_2 = e_1 \mid C_1, C_2] \quad (35)$$

$$= \Pr[D_1 = e_1 \mid C_1] \Pr[D_2 = e_2 \mid C_2] + \Pr[D_1 = e_2 \mid C_1] \Pr[D_2 = e_1 \mid C_2] \quad (36)$$

where again I've used the independence of the two cars in the final line. We can replace the $D_i \mid C_i$ conditionals with the provided normal distribution to get the final expression (with time index included). Let $\mathcal{N}_{C_{ti}}(x) = \mathcal{N}(x; \|a_t - C_{ti}\|; \sigma^2)$ (for brevity's sake).

$$\Pr[C_{11}, C_{12} \mid E_1 = e_1] \propto \Pr[C_{11}] \Pr[C_{12}] [\mathcal{N}_{C_{11}}(e_{11}) \mathcal{N}_{C_{12}}(e_{12}) + \mathcal{N}_{C_{11}}(e_{12}) \mathcal{N}_{C_{12}}(e_{11})] \quad (37)$$

(b) Assuming the prior $p(c_{1i})$ is the same for all i , show that the number of assignments for all K cars (c_{11}, \dots, c_{1K}) that obtain the maximum value of $\mathbb{P}(C_{11} = c_{11}, \dots, C_{1K} = c_{1K} \mid E_1 = e_1)$ is at least $K!$.

Intuitively, by introducing the constraint that $p(c_{1i})$ is the same for all i , combined with the fact that we don't know which car each element of E_1 corresponds to, we can take any of the assignments that maximize $\Pr [C_{11}, \dots, C_{1K} \mid E_1]$ and permute the locations of the cars without changing the probability. Since there are $K!$ such permutations, there are at least $K!$ assignments of the cars that obtain the maximum value.

(c) For general K , what is the treewidth corresponding to the posterior distribution over all K car locations at all T time steps conditioned on all the sensor readings:

$$\mathbb{P}(C_{11} = c_{11}, \dots, C_{1K} = c_{1K}, \dots, C_{T1} = c_{T1}, \dots, C_{TK} = c_{TK} \mid E_1 = e_1, \dots, E_T = e_T)$$

Given the factor graph, we'll first have to condition on the evidence e_1, \dots, e_T . This will create factors f_t for all timesteps $1 \leq t \leq T$, each with arity K . Since we want a variable ordering that starts with variables with the least amount of neighbors, we'll order along timesteps (from 1 to T). Since the Markov blanket of each variable we eliminate along timesteps has K variables (corresponding to the $K - 1$ other cars plus the transition), the resultant factors will have arities no larger than K . By definition, then, the treewidth is K .

(d) Now suppose you change your sensors so that at each time step t , they return the list of exact positions of the K cars, but shifted (with wrap around) by a random amount. For example, if the true car positions at time step 1 are $c_{11} = 1, c_{12} = 3, c_{13} = 8, c_{14} = 5$, then e_1 would be $[1, 3, 8, 5], [3, 8, 5, 1], [8, 5, 1, 3]$, or $[5, 1, 3, 8]$, each with probability $1/4$. Describe an efficient algorithm for computing $p(c_{ti} \mid e_1, \dots, e_T)$ for any time step t and car i . Your algorithm should not be exponential in K or T .